

$$\hat{\theta}_N = \underset{\theta \in \Theta \subseteq \mathbb{R}^L}{\operatorname{argmax}}$$

estimator is implicit
function of the sample size

• θ is an $L \times 1$ vector of
unknown parameters

refers to
scalar

$$E[l_i(y_i | x_i; \theta)]$$

Average log-likelihood

$$E_N[\cdot] = \frac{1}{N} \sum_{i=1}^N [\cdot]$$

in this case
log-likelihood function

$$E_{\theta_0}[l_i(\theta_0; y_i)]$$

\hookrightarrow is the true log-likelihood function

$$= \int l_i(y_i | x_i; \theta) p(y_i | x_i; \theta_0) dy$$

Example

$\{x_i\}_{i=1}^N$ iid $\sim \exp(1)$
 scalar, i
 can draw on
 a continuous
 axis

$$f_x(x_i, 1) = 1 \exp(-1x_i)$$

\downarrow

$$l_x(x_i, 1) = \log 1 - 1x_i$$

log-likelihood for each individual
chromosome

$$E_N[l_x(x_i, 1)] = \frac{1}{N} \sum [\log 1 - 1x_i] = \log 1 - 1 \left[\frac{1}{N} \sum x_i \right]$$

$$= \log 1 - 1\bar{x}$$

$$E_{\theta_0}[x_i] = \frac{1}{\lambda_0}$$

$$\hat{\theta} = \underset{\lambda > 0}{\operatorname{argmax}} [\log 1 - 1\bar{x}]$$

$$\frac{\partial E_N[x, 1]}{\partial \lambda} = \frac{1}{\lambda} - \bar{x} = 0$$

$$\hat{\lambda} = \frac{1}{\bar{x}} = \frac{N}{\sum x_i}$$

$$E_{\theta} [l(x_i, \lambda)] = \log 1 - \frac{\lambda}{\lambda_0}$$

does not depend on θ
 because we
 took expectation
 (we really we do not know
 this quantity)

argmax
 $\lambda > 0$ $\left[\log 1 - \frac{\lambda}{\lambda_0} \right]$

true pop. exp.
 $E_{\theta} [l(\lambda; \theta_0)]$
 so we know
 the data

↓
 but at least i can study
 if theoretically the model is sound

$$\frac{\partial E_{\theta} [l(\lambda, \lambda_0)]}{\partial \lambda} = \frac{1}{\lambda} - \frac{1}{\lambda_0} = 0 \Rightarrow \boxed{\lambda = \lambda_0}$$

Exactly what we wanted

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 the solution is unique!

$$E_N [l(y_i | x_i, \theta)] \xrightarrow{P} E_{\theta_0} [l(y_i | x_i; \theta)]$$

(sample average) convergence pointwise,
 for each value of θ (population moment)

⇓
 $\hat{\theta}_N \rightarrow \theta_0$

$$E_N [l(y_i; \theta)] \stackrel{?}{\underset{\text{pointwise}}{>}} E_{\theta_0} [l(y_i; \theta)]$$

ASYMPTOTIC NORMALITY

$$\sqrt{N}(\hat{\theta} - \theta_0) \xrightarrow{d} N_L(0; \Sigma_{\theta_0})$$

(LxL)
multidimensional

$$\Sigma_{\theta_0} = -E_{\theta_0} \left[\frac{\partial^2 \ell(\theta)}{\partial \theta \partial \theta'} \bigg|_{\theta=\theta_0} \right]$$

inverse of the variance is sort of a measure of the precision

$$E_N \left[\frac{\partial \ell(\theta)}{\partial \theta} \bigg|_{\theta=\hat{\theta}} \right] = 0$$

by construction the empirical score evaluated in $\hat{\theta}$ is zero

the more curve is the flatter in θ_0 the more we are precise with estimation, the steeper is the curvature, the

EXPANSION OF THE SCORE AROUND θ_0

$$E_N \left[\frac{\partial \ell(\theta)}{\partial \theta} \bigg|_{\theta=\theta_0} \right] + E_N \left[\frac{\partial^2 \ell(\theta)}{\partial \theta \partial \theta} \bigg|_{\theta=\bar{\theta}} \right] (\hat{\theta} - \theta_0)$$

less the data are informative
less curvature, less precision, more variance and worse use

RIS is to have a precise estimation and to not care about the other terms of the Taylor polynomial

MEAN VALUE THEOREM

Now apply the CLT

$f(x) = f(x_0) + f'(x_0)(x - x_0)$

$$\sqrt{N}(\hat{\theta} - \theta_0) = -E_N \left[\frac{\partial^2 \ell(\theta)}{\partial \theta \partial \theta} \bigg|_{\theta=\bar{\theta}} \right]^{-1} \sqrt{N} E_N \left[\frac{\partial \ell(\theta)}{\partial \theta} \bigg|_{\theta=\theta_0} \right]$$

$$\sqrt{N} E_N \left[\frac{\partial \ell(\theta)}{\partial \theta} \bigg|_{\theta=\theta_0} \right] \xrightarrow{d} N_L \left(0; -E_{\theta_0} \left[\frac{\partial^2 \ell(\theta)}{\partial \theta \partial \theta} \bigg|_{\theta=\theta_0} \right] \right)$$

Lx1
WHY THE MEAN OF THIS IS ZERO?

$E_{\theta_0} \left(\frac{\partial \ell(\theta)}{\partial \theta} \bigg|_{\theta=\theta_0} \right)$

$$\text{Var} \left(\frac{\partial \ell(\theta)}{\partial \theta} \bigg|_{\theta=\theta_0} \right) = E_{\theta_0} \left[\frac{\partial \ell(\theta)}{\partial \theta} \cdot \frac{\partial \ell(\theta)}{\partial \theta'} \bigg|_{\theta=\theta_0} \right]$$

outer product

$$= -E_{\theta_0} \left[\frac{\partial^2 \ell(\theta)}{\partial \theta \partial \theta'} \Big|_{\theta=\theta_0} \right] \text{ Fisher's information}$$

$L \times L$

$$\bullet \text{ plim}_{N \rightarrow \infty} E_N \left[\frac{\partial^2 \ell(\theta)}{\partial \theta \partial \theta'} \Big|_{\theta=\bar{\theta}} \right] = E_{\theta_0} \left[\frac{\partial^2 \ell(\theta)}{\partial \theta \partial \theta'} \Big|_{\theta=\bar{\theta}} \right]$$

true expected Hessian
evaluated at $\bar{\theta}$

$$E_N[\cdot] \xrightarrow[\text{wLLN or vLLN}]{P} E_{\theta_0}[\cdot]$$

$$\text{if } \hat{\theta} \xrightarrow{P} \theta_0, \quad \bar{\theta} = w \hat{\theta} + (1-w) \theta_0$$

$w(0,1)$

$$\downarrow$$

$$\bar{\theta} \rightarrow \theta_0$$

we call the expected Hessian H and the expected gradient g

$$\sqrt{N}(\hat{\theta} - \theta_0) \xrightarrow{d} -H^{-1}(\sqrt{N} \bar{g}) \rightarrow N(0; -H)$$

$\sim N(0; H^{-1} H H^{-1})$

Outer product of the gradient estimator to estimate the inverse of the Hessian

$$H = E_{\theta_0} \left[\frac{\partial^2 \ell(\theta)}{\partial \theta \partial \theta'} \Big|_{\theta=\theta_0} \right]$$

$$g = E_{\theta_0} \left[\frac{\partial \ell(\theta)}{\partial \theta} \Big|_{\theta=\theta_0} \right]$$

$I \rightarrow$ Fisher information

↓
measure of the information we have about a parameter

↓
it is the variance of the score

$$I = \text{Var} \left(\frac{\partial \ell_i(y_i; \theta)}{\partial \theta} \middle| \theta = \theta_0 \right)$$

$$I = -E \left[\frac{\partial^2 \ell_i(y_i; \theta)}{\partial \theta \partial \theta'} \middle| \theta = \theta_0 \right]$$

↓
measure of curvature of the log-likelihood of model

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the more the function is curved, less is the uncertainty about the parameter

If we have a sample of 100 obs. the log-likelihood for that sample will be a function of θ

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we want to maximize it, we do the first derivative and we pose it = 0

↓
if $k > 1$ its derivative is a vector (just direct)
and the variance is a matrix (Hessian)