

Hypothesis testing  $\rightarrow$  Hypotheses on the coefficients of the regression

$H_0$ : null hypothesis:  $R\beta = c$   $R$  and  $c$  matrix and vector  
of coefficient

$n \times k$   $k \times 1$   $r \times 1$   
 generic  
 number of  
 rows of  $r$   
 (is the universal  
 restriction on one history)  
 for sure because it has to be  
 compatible with  $p$

$$B = \begin{bmatrix} B_1 \\ B_2 \\ \vdots \\ B_5 \end{bmatrix}$$

Ho:  $\begin{cases} \beta_1 + \beta_2 = 5 \\ \beta_5 - \beta_4 = 2\beta_3 \end{cases} \rightarrow \text{example of linear restriction}$   
 $\rightarrow \beta_5 - \beta_4 - 2\beta_3 = 0$

R in this case is

$$R = \begin{bmatrix} 1 & 1 & 0 & 0 & 0 \\ 0 & 0 & -2 & -1 & 1 \end{bmatrix}$$

2, number of aqueous

$C$  in this case is  $\begin{bmatrix} 5 \\ 0 \end{bmatrix}$

hypothesis are made on the three parameters of the model

$$(2 \times 1) \quad \boxed{\mathcal{E}NN(0, \sigma^2 \mathbf{I}_N)}$$

$$\beta | X \sim N(\beta^0; \sigma_e^2 (X'X)^{-1})$$

→ this holds independently to the sample size

↓  
Multivariate normal distribution

because we are assuming normality, then,  
if we are assuming normality with only 10  
observations, maybe the assumption is wrong

$$(R\hat{\beta}_1 - c) \sim ?$$

Gaussian is closed under linear operations

$$\hat{\beta} - c \sim N(\beta^0 - c; R \sigma_E^2 (X'X)^{-1} R')$$

$\beta$  is normal also  
this linear transformation is

## LINEARITY OF EX. VALUE

Quadratic operator, so  $R$  is squared

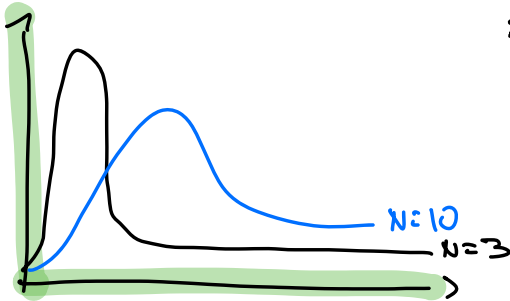
## NOTABLE DISTRIBUTIONS

$(X_1 \dots X_n)$  iid random variables, normal with  $\mu=0$  and  $\sigma^2=1$

$$Y = \sum_{i=1}^n X_i^2 \text{ sum of squares of standard gaussians}$$

$\chi^2(n)$  degrees of freedom

• we see summing squares, the distributions can only be positive



Adding degrees of freedom shifts the distribution more to the right

$\chi^2$  distribution is only positive!

## (II) Student's t

$(X_1 \dots X_n)$  iid normal  $(0, 1)$

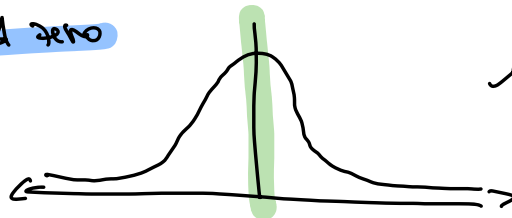
$W_1 \sim$  independent from the others and  $N(0, 1)$

and  $Y \sim t$  student

$$Y = \frac{W_1}{\sqrt{\frac{\sum_{i=2}^n X_i^2}{n}}} \sim t(n)$$

$\rightarrow$  normal of the numerator  
 $\rightarrow$  degrees of freedom  
 $\rightarrow$  it's like a quadratic mean of normals of the denominator

t distribution takes values on the entire real line, and it is symmetric around zero



looks like this

• F-distribution

$W_1$  is  $N \chi^2(N)$

$W_1$  is independent from  $\tau_1$   $W_1 \perp \tau_1$

$Z_1 \sim N \chi^2(M)$

$$Y = \frac{\frac{W_1}{N}}{\frac{\tau_1}{M}}$$

$\sim F(N, M) \rightarrow$  Positive  
degrees of freedom of numerator and denominator

two chi-squared distributions' ratio

$$R\hat{\beta} - c | X \sim N(R\beta^0 - c; R\sigma^2 (X'X)^{-1}R')$$

Suppose we had a  $\hat{\beta}$  very close to the constraint, the difference  $R\hat{\beta} - c$  will be close to zero

we want a r.v. as close as possible to zero when the difference is very small and large when we are far from zero

$$(R\hat{\beta} - c)' [R\sigma^2 (X'X)^{-1}R']^{-1} (R\hat{\beta} - c) \quad (1)$$

we're doing to be  $\approx 0$  because continuous random variables

$$\sum w^2$$

maybe we are not subtracting because under the null hyp. the means are also zero

under the null hypothesis (1) is zero and all the + statistics are calculated under the null hypothesis  $\rightarrow$  null hypo is

dividing by the variance, not the std to standardize, because of the variance we have squares

this is  $N \chi^2(r)$

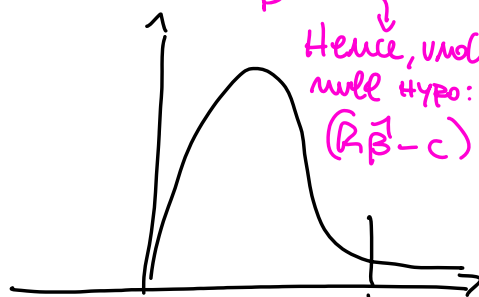
$\hookrightarrow$  sum of standardized squares

we square this because we are interested in distance, both positive and negative, from the value we want

$$R\hat{\beta} = c$$

Hence, under null hypo:

$$(R\hat{\beta} - c) = 0$$



distance from the value  
ex 8.74

i can choose an  $\alpha\%$  level and for example  $\alpha\%$  is 9.78 and since 8.74 is smaller i do not reject the results

or i can compute the p-value and compare it directly with  $\alpha$ , if the probability is less than  $\alpha$  we do not reject the null hypothesis

But we do not know  $\sigma^2_\epsilon$ , hence we cannot do computations we can use  $s^2$ , but we are substituting a scalar with a random variable, hence we are changing the distribution

$$(R\hat{\beta} - c)' (R\hat{\sigma}^2_\epsilon (X'X)^{-1} R')^{-1} (R\hat{\beta} - c) \quad (R - \hat{\beta}c)' [R\hat{\sigma}^2_\epsilon (X'X)^{-1} R']^{-1} (R - \hat{\beta}c)$$

↳ this is standard normal

$\frac{\hat{\epsilon}'\hat{\epsilon}}{N-k} \rightarrow$  it is a scalar and it is an unbiased, so it is going to be denominator

$$\frac{(R\hat{\beta} - c)' (R\hat{\sigma}^2_\epsilon (X'X)^{-1} R')^{-1} (R\hat{\beta} - c)}{\frac{\hat{\epsilon}'\hat{\epsilon}}{N-k}} \rightarrow \sim \chi^2(r)$$

↳ not a normal anymore

$$\frac{(\hat{\epsilon}'\hat{\epsilon})}{(N-k)\hat{\sigma}^2_\epsilon} \rightarrow \text{then are not standardize} \rightarrow \sim \chi^2(N-k)$$

$\rightarrow \sim F(r, N-k)$

$\rightarrow$  the p-value is  $= 1 - \text{cdf}(\hat{F}, r, N-k)$

**F test statistic for the existence of the regression**

$$Y = \beta_1 + \beta_2 x_1 + \dots + \beta_k x_{k-1} + \epsilon$$

**test that all the coefficients are equal to zero**

$Y = \beta_1 + \epsilon$  is the model under the null hypothesis

↳ by testing this we are testing if some of the  $x_i$  have explanatory power for  $y$

$H_0: \beta_2 = 0 \wedge \beta_3 = 0 \dots \wedge \beta_k = 0$   
the  $\beta$ s are **jointly** zero

$H_1$ : the opposite of the null, is that at least one is  $\neq 0$   $\beta_i \neq 0$   $i=2, \dots, k$

$\downarrow$   
 $R\beta = c$

we are imposing  $k-1$  restrictions  $\rightarrow$   $k-2$  zeros, one is  $\beta_1$  one is the  $\beta=0$

$R_{(k-1) \times k} = \begin{bmatrix} 0 & 1 & 0 & 0 & \dots \\ 0 & 0 & 1 & \dots & \\ 0 & 0 & 0 & 1 & 0 & \dots \\ \vdots & \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & 0 & \dots & 0 & 1 \end{bmatrix} = \begin{bmatrix} 0 \\ \vdots \\ 0 \end{bmatrix}_{(k-1) \times 1} \quad I_{(k-1)}$

$R_{(k-1) \times k}$  always equal to the row of  $\beta$

$c = 0_{(k-1) \times 1}$

Why this is an F distribution here?  
 we not dividing for the variance of the numerator!

$$\frac{(R\hat{\beta} - c)' (R(X'X)^{-1}R')^{-1} (R\hat{\beta} - c) / (k-1)}{\frac{\hat{\epsilon}'\hat{\epsilon}}{N-k}}$$

$\sim F(k-1, N-k)$

CAREFUL TO NOT SWAP THESE!

we can also do univariate test, on each of the coefficients

$Y = \beta_1 + \beta_2 X + \epsilon$

test  $\beta_2 = 0 \rightarrow H_0: \beta_2 = 0$

$H_1: \beta_2 \neq 0$

can use the F-test, it works also with a single coefficient

but we can use the t-test for a single restriction

t-test =  $\frac{\hat{\beta}_2}{S.E.(\hat{\beta}_2)} \sim t\text{-distributed}(N-k)$

WHAT IS

$SE(\hat{\beta}_2) = \sqrt{s_e^2 (X'X)^{-1}_{22}}$

t-test is related to signal/noise ratio

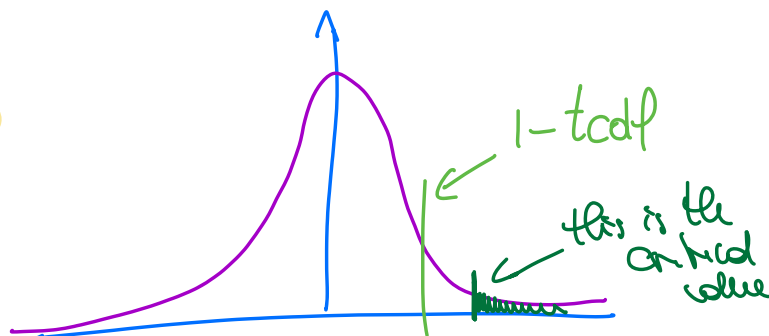
in Models we computed values for each  $\beta$ , then we have computed the CDF

t-test

How likely it is that I got the t-value that I have calculated if the null hypothesis is true?

if  $H_0: \beta = 0$  and the t-test is low

the prob. of getting that number GIVEN that the null hypothesis is true, may be very small, hence I reject the hypothesis



t-test for our hypothesis is that is not  $\beta_0 = 0$  is  $\frac{\hat{\beta} - \text{value to test}}{\hat{\beta} \text{ S.E.}}$

t-test is used to test single parameters

if I want to test multiple parameters (test of existence of the model) I use the F-test

$$H_0 = \beta_i = 0 \quad i=1, \dots, k$$

$H_1 = \beta_i \neq 0 \quad i=1, \dots, k$  → if also one parameter is significantly different from zero I reject the null hypothesis

think also to the F-test as a ratio between RSS

$$\frac{RSS_{restricted} - RSS_{unrestricted} / df_1}{RSS_{unrestricted} / df_2}$$

$RSS_{restricted}$  is bigger than  $RSS_{unrestricted}$

of course if I have the t-test I already know if some parameters are significantly  $\neq 0$

Exercise - test the hypothesis  $\beta_1 = \beta_5 = \beta_6 = 0$

$$R\beta = c$$

$(3 \times 6) \quad (6 \times 1) \quad (3 \times 1)$

the model has 6 parameters

$$R = \begin{bmatrix} 0 & 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 0 & 1 \end{bmatrix} \quad C = \begin{bmatrix} 0 \\ 0 \\ 0 \end{bmatrix}$$

point 5 under Global  $\Rightarrow \beta_{1,5,6} = 0$

$$\text{test } \beta_3 = 1 - \beta_2 \Rightarrow \beta_2 + \beta_3 = 1$$

$$R = \begin{bmatrix} 0 & 1 & 1 \end{bmatrix} \begin{bmatrix} \beta_1 \\ \beta_2 \\ \beta_3 \end{bmatrix} \quad C = \begin{bmatrix} 1 \end{bmatrix}$$