

Q1 (25 points)

Theorem 10.4 of Lecture 10 states that the size-biased permutation of the two-parameter Poisson-Dirichlet distribution is given by the two-parameter GEM distribution. Using the notation of the theorem, show that for any integer  $m \geq 1$

$$\mathbb{E} [P_{\tau_3}^m(\alpha, \theta)] = \mathbb{E} [V_3^m(\alpha, \theta)].$$

Using this result to derive the case for one-parameter model. (Justification of all steps is required to get the full credit.)

**Proof:**

For brevity, denote  $P_k(\alpha, \theta)$  as  $P_k$ ,  $\mathbf{P}(\alpha, \theta)$  as  $\mathbf{P}$ ,  $V_k(\alpha, \theta)$  as  $V_k$ ,  $\mathbf{V}(\alpha, \theta)$  as  $\mathbf{V}$ . Fix  $m \in \mathbb{N}$ . We begin by showing that

$$E \left( \left( \frac{P_{\tau_3}}{1 - P_{\tau_1} - P_{\tau_2}} \right)^m \right) = E(U_3^m) :$$

To begin, since  $\mathbf{P} \sim PD(\alpha, \theta)$ , by definition 10.2 (lecture 10), the order statistics of  $\mathbf{V}$  are precisely  $\mathbf{P}$ . We have already shown in class that

$$E(P_{\tau_1}^m) = E(V_1^m).$$

Thus,  $P_{\tau_1} \stackrel{d}{=} V_1$ . Since the order statistics of  $\mathbf{V}$  are equal to  $\mathbf{P}$ , we claim that  $P_{\tau_1} = V_1$ . Indeed, if this was not the case, then there exists  $k \in \mathbb{N} - \{1\}$   $P_{\tau_1} = V_k$ , which implies that  $V_1 = V_k$ . Then

$$E(V_1) = E(V_k) \implies E(U_1) = E((1 - U_1) \dots (1 - U_{k-1}) U_k) = E(1 - U_1) \quad \text{By definition 10.1.}$$

Following by definition 10.1, we have that

$$E(U_1) = \frac{1 - \alpha}{1 + \theta}.$$

And,

$$E(V_k) = E((1 - U_1)(1 - U_2) \dots (1 - U_{k-1}) U_k) = E(1 - U_1) E(1 - U_2) \dots E(1 - U_{k-1}) E(U_k) \quad (\text{By independence}).$$

$$= \frac{\theta + \alpha}{(1 - \alpha) + (\theta + \alpha)} \frac{\theta + 2\alpha}{(1 - \alpha) + (\theta + 2\alpha)} \dots \frac{\theta + (k - 1)\alpha}{(1 - \alpha) + (\theta + (k - 1)\alpha)} \frac{1 - \alpha}{(1 - \alpha) + (\theta + k\alpha)} < \frac{1 - \alpha}{(1 - \alpha) + (\theta + k\alpha)} \quad \text{Since } \alpha \in [0, 1].$$

And so

$$E(V_k) < \frac{1 - \alpha}{\theta + (k - 1)\alpha} \leq \frac{1 - \alpha}{1 + \theta} = E(V_1).$$

Thus, for  $k \in \mathbb{N} - \{1\}$ ,  $E(V_k) < E(V_1)$ , which cannot be the case if  $V_1 \stackrel{d}{=} V_k$ . As a result,  $P_{\tau_1} = V_1$ . It can be shown in a very similar way that since  $P_{\tau_2} \stackrel{d}{=} V_2$  (shown in class),  $P_{\tau_2} = V_2$ :

Suppose, for the sake of a contradiction, that  $P_{\tau_2} \stackrel{d}{=} V_j$  for some  $j \in \mathbb{N} - \{2\}$  which implies that  $V_2 \stackrel{d}{=} V_j$ . We have already shown that  $V_1$  is not equal in distribution to  $V_k$  for  $k \in \mathbb{N} - \{1\}$ , thus  $V_2$  is not equal in distribution to  $V_1$ . So, if  $P_{\tau_2} \stackrel{d}{=} V_j$  for some  $j \in \mathbb{N} - \{2\}$ , it follows that  $j \neq 1$ , so  $j \in \mathbb{N} - \{1, 2\}$ . Note that:

$$E(V_2) = E((1 - U_1) U_2) = E(1 - U_1) E(U_2) = \frac{\theta + \alpha}{(1 - \alpha) + (\theta + \alpha)} \frac{1 - \alpha}{(1 - \alpha) + (\theta + 2\alpha)}.$$

And,

$$\begin{aligned} E(V_j) &= E((1 - U_1) \dots (1 - U_{j-1}) U_j) = E(1 - U_1) \dots E(1 - U_{j-1}) E(U_j) \\ &= \frac{\theta + \alpha}{(1 - \alpha) + (\theta + \alpha)} \dots \frac{\theta + (j - 1)\alpha}{(1 - \alpha) + (\theta + (j - 1)\alpha)} \frac{1 - \alpha}{(1 - \alpha) + (\theta + j\alpha)} < \end{aligned}$$

$$\frac{\theta + \alpha}{(1 - \alpha) + (\theta + \alpha)} \frac{1 - \alpha}{(1 - \alpha) + (\theta + j\alpha)} \quad \text{Since } \alpha \in [0, 1), \text{ and } j \in \mathbb{N} - \{1, 2\}.$$

And so,

$$E(V_j) < \frac{\theta + \alpha}{(1 - \alpha) + (\theta + \alpha)} \frac{1 - \alpha}{(1 - \alpha) + (\theta + j\alpha)} \leq \frac{\theta + \alpha}{(1 - \alpha) + (\theta + \alpha)} \frac{1 - \alpha}{(1 - \alpha) + (\theta + \alpha)} = E(V_2).$$

Thus, for  $j \in \mathbb{N} - \{1, 2\}$ , we conclude that  $E(V_j) \neq E(V_2)$  which cannot be the case if  $V_2 \stackrel{d}{=} V_j$ . As a result,  $P_{\tau_2} = V_2$

So, we have:

$$\begin{aligned} E\left(\left(\frac{P_{\tau_3}}{1 - P_{\tau_1} - P_{\tau_2}}\right)^m\right) &= E_{\mathbf{P}, \tau_1, \tau_2}\left(E_{\tau_3}\left(\left(\frac{P_{\tau_3}}{1 - P_{\tau_1} - P_{\tau_2}}\right)^m \middle| \mathbf{P}, \tau_1, \tau_2\right)\right) \quad \text{By double expectation} \\ &= E\left(\sum_{n=1}^{\infty} \left(\frac{P_n}{1 - P_{\tau_1} - P_{\tau_2}}\right)^m \frac{P_n}{1 - P_{\tau_1} - P_{\tau_2}} I_{n \neq \tau_1, n \neq \tau_2}\right) \\ &= E\left(\sum_{n=1}^{\infty} \left(\frac{P_n}{1 - P_{\tau_1} - P_{\tau_2}}\right)^{m+1} I_{P_n \neq P_{\tau_1}, P_n \neq P_{\tau_2}}\right) \\ &= E\left(\sum_{n=1}^{\infty} \left(\frac{P_n}{1 - V_1 - V_2}\right)^{m+1} I_{P_n \neq V_1, P_n \neq V_2}\right) \quad \text{As we have shown, } V_1 = P_{\tau_1}, V_2 = P_{\tau_2} \\ &= E\left(\sum_{n=3}^{\infty} \left(\frac{V_{(n)}}{1 - V_1 - V_2}\right)^{m+1}\right) \quad \text{Since } \mathbf{P} \text{ are the order statistics of } \mathbf{V} \quad \dagger. \end{aligned}$$

Note that since  $V_k = \prod_{n=1}^{k-1} (1 - U_n) U_k$ , and  $U_i \sim \text{Beta}(1 - \alpha, \theta + i\alpha)$ , we have that  $V_k \in [0, 1]$ , which means that the sequence  $\left(\sum_{n=1}^b V_n^{m+1}, b \in \mathbb{N}\right)$  is a monotone sequence. Moreover, since:

$$\sum_{n=1}^{\infty} V_n = 1 \quad a.s.$$

We can conclude that

$$0 \leq \sum_{n=1}^{\infty} V_n^{m+1} \leq \sum_{n=1}^{\infty} V_n = 1 \quad a.s$$

Thus, the sequence  $\left(\sum_{n=1}^b V_n^{m+1}, b \in \mathbb{N}\right)$  is monotone and bounded, thus by the monotone convergence theorem, the series  $\sum_{n=1}^{\infty} V_n^{m+1}$  converges pointwise. This convergence is absolute since  $V_n \in [0, 1]$  for each  $n \in \mathbb{N}$ , so  $\sum_{n=1}^{\infty} V_n = \sum_{n=1}^{\infty} V_{(n)}$  (since the convergence is absolute, it can be reordered). Thus, plugging this result into  $\dagger$ :

$$\begin{aligned} E\left(\left(\frac{P_{\tau_3}}{1 - P_{\tau_1} - P_{\tau_2}}\right)^m\right) &= E\left(\sum_{n=3}^{\infty} \left(\frac{V_{(n)}}{1 - V_1 - V_2}\right)^{m+1}\right) \\ &= E\left(\sum_{n=1}^{\infty} \left(\frac{V_n}{1 - V_1 - V_2}\right)^{m+1}\right) \end{aligned}$$

Since  $\left(\sum_{n=1}^b V_n^{m+1}, b \in \mathbb{N}\right)$  is a bounded monotone sequence that converges almost surely, the sequence  $\left(\sum_{n=1}^b \left(\frac{V_n}{1 - V_1 - V_2}\right)^{m+1}, b \in \mathbb{N}\right)$  is a bounded monotone sequence that converges almost surely, thus by the monotone convergence theorem:

$$\begin{aligned}
E \left( \left( \frac{P_{\tau_3}}{1 - P_{\tau_1} - P_{\tau_2}} \right)^m \right) &= E \left( \sum_{n=3}^{\infty} \left( \frac{V_n}{1 - V_1 - V_2} \right)^{m+1} \right) \\
&= \sum_{n=3}^{\infty} E \left( \left( \frac{V_n}{1 - V_1 - V_2} \right)^{m+1} \right) \\
&= \sum_{n=3}^{\infty} E \left( \left( \frac{\prod_{i=1}^{n-1} (1 - U_i) U_n}{1 - U_1 - (1 - U_1) U_2} \right)^{m+1} \right) \\
&= \sum_{n=3}^{\infty} E \left( \left( \frac{\prod_{i=1}^{n-1} (1 - U_i) U_n}{(1 - U_1) (1 - U_2)} \right)^{m+1} \right) \\
&= \sum_{n=3}^{\infty} E \left( \left( \prod_{i=3}^{n-1} (1 - U_i) U_n \right)^{m+1} \right) \\
&= \sum_{n=3}^{\infty} E (U_n^{m+1}) \prod_{i=3}^{n-1} E ((1 - U_i)^{m+1}) \quad \text{By independence} \\
&= E (U_3^{m+1}) + \sum_{n=4}^{\infty} E (U_n^{m+1}) \prod_{i=3}^{n-1} E ((1 - U_i)^{m+1})^{\ddagger}.
\end{aligned}$$

Since  $U_n \sim \text{Beta}(1 - \alpha, \theta + n\alpha)$ ,

$$\begin{aligned}
E (U_n^{m+1}) &= \int_0^1 \frac{\Gamma(1 - \alpha + \theta + n\alpha)}{\Gamma(1 - \alpha) \Gamma(\theta + n\alpha)} u^{m+1} u^{1-\alpha+1} (1 - u)^{\theta+n\alpha-1} du \\
&= \frac{\Gamma(1 - \alpha + \theta + n\alpha)}{\Gamma(1 - \alpha) \Gamma(\theta + n\alpha)} \int_0^1 u^{m+2+\alpha-1} (1 - u)^{\theta+n\alpha-1} \\
&= \frac{\Gamma(1 - \alpha + \theta + n\alpha)}{\Gamma(1 - \alpha) \Gamma(\theta + n\alpha)} \frac{\Gamma(m+2 - \alpha) \Gamma(\theta + n\alpha)}{\Gamma(m+2 - \alpha + \theta + n\alpha)} \\
&= \frac{\Gamma(1 - \alpha + \theta + n\alpha)}{\Gamma(1 - \alpha)} \frac{\Gamma(m+1+1 - \alpha)}{\Gamma(m+1+1 - \alpha + \theta + n\alpha)} \\
&= \frac{\Gamma(m+1+1 - \alpha)}{\Gamma(1 - \alpha)} \frac{\Gamma(1 + \theta + (n-1)\alpha)}{\Gamma(m+1+1 + \theta + (n-1)\alpha)} \\
&= \frac{(1 - \alpha)_{(m+1)}}{(1 + \theta + (n-1)\alpha)_{(m+1)}} \clubsuit
\end{aligned}$$

Since  $U_i \sim \text{Beta}(1 - \alpha, \theta + i\alpha)$ ,

$$\begin{aligned}
E\left((1-U_i)^{m+1}\right) &= \int_0^1 \frac{\Gamma(1-\alpha+\theta+i\alpha)}{\Gamma(1-\alpha)\Gamma(\theta+i\alpha)} (1-u)^{m+1} u^{1-\alpha+1} (1-u)^{\theta+i\alpha-1} du \\
&= \frac{\Gamma(1-\alpha+\theta+i\alpha)}{\Gamma(1-\alpha)\Gamma(\theta+i\alpha)} \int_0^1 (1-u)^{m+1+\theta+i\alpha} u^{1-\alpha+1} du \\
&= \frac{\Gamma(1-\alpha+\theta+i\alpha)}{\Gamma(1-\alpha)\Gamma(\theta+i\alpha)} \left( \frac{\Gamma(m+1+1-\alpha+\theta+i\alpha)}{\Gamma(1-\alpha)\Gamma(m+1+\theta+i\alpha)} \right)^{-1} \\
&= \frac{\Gamma(1+\theta+(i-1)\alpha)}{\Gamma(\theta+i\alpha)} \left( \frac{\Gamma(m+1+1+\theta+(i-1)\alpha)}{\Gamma(m+1+\theta+i\alpha)} \right)^{-1} \\
&= \frac{\Gamma(m+1+\theta+i\alpha)}{\Gamma(\theta+i\alpha)} \frac{\Gamma(1+\theta+(i-1)\alpha)}{\Gamma(m+1+1+\theta+(i-1)\alpha)} \\
&= \frac{(\theta+i\alpha)_{(m+1)}}{(1+\theta+(i-1)\alpha)_{(m+1)}}
\end{aligned}$$

Note:

$$\begin{aligned}
\frac{1}{(1+\theta+(i-1)\alpha)_{(m+1)}} &= \frac{1}{(1+\theta+(i-1)\alpha) \cdots (m-1+1+\theta+(i-1)\alpha) (m+1+\theta+(i-1)\alpha)} \\
&= \frac{\theta+(i-1)\alpha}{(\theta+(i-1)\alpha) (1+\theta+(i-1)\alpha) \cdots (m+\theta+(i-1)\alpha) (m+1+\theta+(i-1)\alpha)} \\
&= \frac{\theta+(i-1)\alpha}{(\theta+(i-1)\alpha)_{(m+1)} (m+1+\theta+(i-1)\alpha)}
\end{aligned}$$

So,

$$\begin{aligned}
E\left((1-U_i)^{m+1}\right) &= \frac{(\theta+i\alpha)_{(m+1)}}{(1+\theta+(i-1)\alpha)_{(m+1)}} \\
&= (\theta+i\alpha)_{(m+1)} \frac{\theta+(i-1)\alpha}{(\theta+(i-1)\alpha)_{(m+1)} (m+1+\theta+(i-1)\alpha)} \quad \diamond
\end{aligned}$$

Applying results ♣, ◇ to ‡:

$$E\left(\left(\frac{P_{\tau_3}}{1-P_{\tau_1}-P_{\tau_2}}\right)^m\right) =$$

$$\begin{aligned}
&\frac{(1-\alpha)_{(m+1)}}{(1+\theta+2\alpha)_{(m+1)}} + \sum_{n=4}^{\infty} \frac{(1-\alpha)_{(m+1)}}{(1+\theta+(n-1)\alpha)_{(m+1)}} \prod_{i=3}^{n-1} (\theta+i\alpha)_{(m+1)} \frac{\theta+(i-1)\alpha}{(\theta+(i-1)\alpha)_{(m+1)} (m+1+\theta+(i-1)\alpha)} \\
&= \frac{(1-\alpha)_{(m+1)}}{(1+\theta+2\alpha)_{(m+1)}} + \sum_{n=4}^{\infty} \frac{(1-\alpha)_{(m+1)}}{(1+\theta+(n-1)\alpha)_{(m+1)}} \left( \prod_{i=3}^{n-1} \frac{(\theta+i\alpha)_{(m+1)}}{(\theta+(i-1)\alpha)_{(m+1)}} \right) \left( \prod_{i=3}^{n-1} \frac{\theta+(i-1)\alpha}{m+1+\theta+(i-1)\alpha} \right)
\end{aligned}$$

We simplify  $\prod_{i=3}^{n-1} \frac{(\theta+i\alpha)_{(m+1)}}{(\theta+(i-1)\alpha)_{(m+1)}}$ :

$$\begin{aligned}
\prod_{i=3}^{n-1} \frac{(\theta+i\alpha)_{(m+1)}}{(\theta+(i-1)\alpha)_{(m+1)}} &= \frac{(\theta+3\alpha)_{(m+1)}}{(\theta+2\alpha)_{(m+1)}} \frac{(\theta+4\alpha)_{(m+1)}}{(\theta+3\alpha)_{(m+1)}} \cdots \frac{(\theta+(n-2)\alpha)_{(m+1)}}{(\theta+(n-3)\alpha)_{(m+1)}} \frac{(\theta+(n-1)\alpha)_{(m+1)}}{(\theta+(n-2)\alpha)_{(m+1)}} \\
&= \frac{(\theta+(n-1)\alpha)_{(m+1)}}{(\theta+2\alpha)_{(m+1)}}
\end{aligned}$$

As a result:

$$E \left( \left( \frac{P_{\tau_3}}{1 - P_{\tau_1} - P_{\tau_2}} \right)^m \right) =$$

$$\begin{aligned} & \frac{(1 - \alpha)_{(m+1)}}{(1 + \theta + 2\alpha)_{(m+1)}} + \sum_{n=4}^{\infty} \frac{(1 - \alpha)_{(m+1)}}{(1 + \theta + (n - 1)\alpha)_{(m+1)}} \frac{(\theta + (n - 1)\alpha)_{(m+1)}}{(\theta + 2\alpha)_{(m+1)}} \left( \prod_{i=3}^{n-1} \frac{\theta + (i - 1)\alpha}{m + 1 + \theta + (i - 1)\alpha} \right) \\ &= \frac{(1 - \alpha)_{(m+1)}}{(1 + \theta + 2\alpha)_{(m+1)}} + \frac{(1 - \alpha)_{(m+1)}}{(\theta + 2\alpha)_{(m+1)}} \sum_{n=4}^{\infty} \frac{(\theta + (n - 1)\alpha)_{(m+1)}}{(1 + \theta + (n - 1)\alpha)_{(m+1)}} \left( \prod_{i=3}^{n-1} \frac{\theta + (i - 1)\alpha}{m + 1 + \theta + (i - 1)\alpha} \right) \end{aligned}$$

Note that:

$$\begin{aligned} \frac{(\theta + (n - 1)\alpha)_{(m+1)}}{(1 + \theta + (n - 1)\alpha)_{(m+1)}} &= \frac{\theta + (n - 1)\alpha}{1 + \theta + (n - 1)\alpha} \frac{1 + \theta + (n - 1)\alpha}{2 + \theta + (n - 1)\alpha} \cdots \frac{m + \theta + (n - 1)\alpha}{m + 1 + \theta + (n - 1)\alpha} \\ &= \frac{\theta + (n - 1)\alpha}{m + 1 + \theta + (n - 1)\alpha} \end{aligned}$$

Thus,

$$\begin{aligned} E \left( \left( \frac{P_{\tau_3}}{1 - P_{\tau_1} - P_{\tau_2}} \right)^m \right) &= \frac{(1 - \alpha)_{(m+1)}}{(1 + \theta + 2\alpha)_{(m+1)}} + \frac{(1 - \alpha)_{(m+1)}}{(\theta + 2\alpha)_{(m+1)}} \sum_{n=4}^{\infty} \frac{\theta + (n - 1)\alpha}{m + 1 + \theta + (n - 1)\alpha} \left( \prod_{i=3}^{n-1} \frac{\theta + (i - 1)\alpha}{m + 1 + \theta + (i - 1)\alpha} \right) \\ &= \frac{(1 - \alpha)_{(m+1)}}{(1 + \theta + 2\alpha)_{(m+1)}} + \frac{(1 - \alpha)_{(m+1)}}{(\theta + 2\alpha)_{(m+1)}} \sum_{n=4}^{\infty} \left( \prod_{i=3}^n \frac{\theta + (i - 1)\alpha}{m + 1 + \theta + (i - 1)\alpha} \right) \\ &= \frac{(1 - \alpha)_{(m+1)}}{(1 + \theta + 2\alpha)_{(m+1)}} + \frac{(1 - \alpha)_{(m+1)}}{(\theta + 2\alpha)_{(m+1)}} \sum_{n=4}^{\infty} \frac{\theta + 2\alpha}{m + 1 + \theta + 2\alpha} \left( \prod_{i=4}^n \frac{\theta + (i - 1)\alpha}{m + 1 + \theta + (i - 1)\alpha} \right) \\ &= \frac{(1 - \alpha)_{(m+1)}}{(1 + \theta + 2\alpha)_{(m+1)}} + \frac{(1 - \alpha)_{(m+1)}}{(\theta + 2\alpha)_{(m+1)}} \frac{\theta + 2\alpha}{m + 1 + \theta + 2\alpha} \sum_{n=4}^{\infty} \left( \prod_{i=4}^n \frac{\theta + (i - 1)\alpha}{m + 1 + \theta + (i - 1)\alpha} \right) \heartsuit \\ &= \frac{(1 - \alpha)_{(m+1)}}{(1 + \theta + 2\alpha)_{(m+1)}} + (1 - \alpha)_{(m+1)} \frac{\theta + 2\alpha}{(\theta + 2\alpha)_{(m+1)} (m + 1 + \theta + 2\alpha)} \sum_{n=4}^{\infty} \left( \prod_{i=4}^n \frac{\theta + (i - 1)\alpha}{m + 1 + \theta + (i - 1)\alpha} \right) \heartsuit \end{aligned}$$

Following from direct calculation:

$$\begin{aligned} \frac{\theta + 2\alpha}{(\theta + 2\alpha)_{(m+1)} (m + 1 + \theta + 2\alpha)} &= \frac{\theta + 2\alpha}{(\theta + 2\alpha)(1 + \theta + 2\alpha)(2 + \theta + 2\alpha) \cdots (m + \theta + 2\alpha)(m + 1 + \theta + 2\alpha)} \\ &= \frac{1}{(1 + \theta + 2\alpha)_{(m+1)}} \end{aligned}$$

Plugging this result into  $\heartsuit$ :

$$\begin{aligned} E \left( \left( \frac{P_{\tau_3}}{1 - P_{\tau_1} - P_{\tau_2}} \right)^m \right) &= \frac{(1 - \alpha)_{(m+1)}}{(1 + \theta + 2\alpha)_{(m+1)}} + \frac{(1 - \alpha)_{(m+1)}}{(1 + \theta + 2\alpha)_{(m+1)}} \sum_{n=4}^{\infty} \left( \prod_{i=4}^n \frac{\theta + (i - 1)\alpha}{m + 1 + \theta + (i - 1)\alpha} \right) \end{aligned}$$

$$\begin{aligned}
&= \frac{(1-\alpha)_{(m+1)}}{(1+\theta+2\alpha)_{(m+1)}} \left( 1 + \sum_{n=4}^{\infty} \left( \prod_{i=4}^n \frac{\theta+(i-1)\alpha}{m+1+\theta+(i-1)\alpha} \right) \right) \\
&= \frac{(1-\alpha)_{(m+1)}}{(1+\theta+2\alpha)_{(m+1)}} \left( 1 + \frac{\theta+3\alpha}{m+1+\theta+3\alpha} + \frac{\theta+3\alpha}{m+1+\theta+3\alpha} \frac{\theta+4\alpha}{m+1+\theta+4\alpha} + \dots \right) \\
&= \frac{(1-\alpha)_{(m+1)}}{(1+\theta+2\alpha)_{(m+1)}} \left( 1 + \frac{(\theta+3\alpha)/\alpha}{(m+1+\theta+3\alpha)/\alpha} + \frac{(\theta+3\alpha)/\alpha}{(m+1+\theta+3\alpha)/\alpha} \frac{(\theta+3\alpha+\alpha)/\alpha}{(m+1+\theta+3\alpha+\alpha)/\alpha} + \dots \right) \\
&= \frac{(1-\alpha)_{(m+1)}}{(1+\theta+2\alpha)_{(m+1)}} \left( 1 + \frac{(\theta+3\alpha)/\alpha}{(m+1+\theta+3\alpha)/\alpha} + \frac{(\theta+3\alpha)/\alpha}{(m+1+\theta+3\alpha)/\alpha} \frac{(\theta+3\alpha)/\alpha+1}{(m+1+\theta+3\alpha)/\alpha+1} + \dots \right)
\end{aligned}$$

Let  $a = \frac{\theta+3\alpha}{\alpha}$ ,  $c = \frac{m+1+\theta+3\alpha}{\alpha}$ , then

$$\begin{aligned}
&\frac{(1-\alpha)_{(m+1)}}{(1+\theta+2\alpha)_{(m+1)}} \left( 1 + \frac{(\theta+3\alpha)/\alpha}{(m+1+\theta+3\alpha)/\alpha} + \frac{(\theta+3\alpha)/\alpha}{(m+1+\theta+3\alpha)/\alpha} \frac{(\theta+3\alpha)/\alpha+1}{(m+1+\theta+3\alpha)/\alpha+1} + \dots \right) \\
&= \frac{(1-\alpha)_{(m+1)}}{(1+\theta+2\alpha)_{(m+1)}} \left( 1 + \frac{a}{c} + \frac{a}{c} \frac{a+1}{c+1} + \dots \right) \\
&= \frac{(1-\alpha)_{(m+1)}}{(1+\theta+2\alpha)_{(m+1)}} \left( \frac{c-1}{c-a-1} \right) \quad \text{By Gauss's Hypergeometric theorem} \\
&= \frac{(1-\alpha)_{(m+1)}}{(1+\theta+2\alpha)_{(m+1)}} \left( \frac{c-1}{c-a-1} \right) \\
&= \frac{(1-\alpha)_{(m+1)}}{(1+\theta+2\alpha)_{(m+1)}} \frac{\frac{m+1+\theta+3\alpha}{\alpha} - 1}{\frac{m+1+\theta+3\alpha}{\alpha} - \frac{m+1+\theta+2\alpha}{\alpha} - 1} \\
&= \frac{(1-\alpha)_{(m+1)}}{(1+\theta+2\alpha)_{(m+1)}} \frac{\frac{m+1+\theta+2\alpha}{\alpha}}{\frac{m+1+\theta+2\alpha}{\alpha}} = \frac{(1-\alpha)_{(m+1)}}{(1+\theta+2\alpha)_{(m+1)}} \frac{m+1+\theta+2\alpha}{m+1-\alpha} \\
&= \frac{(1-\alpha)(1+1-\alpha)\dots(m-1+1-\alpha)\cancel{(m+1-\alpha)}}{(1+\theta+2\alpha)(1+1+\theta+2\alpha)(2+1+\theta+2\alpha)\dots(m-1+1+\theta+2\alpha)\cancel{(m+1+\theta+2\alpha)}} \frac{\cancel{m+1+\theta+2\alpha}}{\cancel{m+1-\alpha}} \\
&= \frac{(1-\alpha)_{(m)}}{(1+\theta+2\alpha)_{(m)}}
\end{aligned}$$

Note that:

$$\begin{aligned}
E(U_3^m) &= \int_0^1 u^m \frac{\Gamma(1-\alpha+\theta+3\alpha)}{\Gamma(1-\alpha)\Gamma(\theta+3\alpha)} u^{1-\alpha-1} (1-u)^{\theta+3\alpha-1} du \\
&= \int_0^1 \frac{\Gamma(1-\alpha+\theta+3\alpha)}{\Gamma(1-\alpha)\Gamma(\theta+3\alpha)} u^{m+1-\alpha-1} (1-u)^{\theta+3\alpha-1} du \\
&= \frac{\Gamma(1-\alpha+\theta+3\alpha)}{\Gamma(1-\alpha)\Gamma(\theta+3\alpha)} \int_0^1 u^{m+1-\alpha-1} (1-u)^{\theta+3\alpha-1} du \\
&= \frac{\Gamma(1-\alpha+\theta+3\alpha)}{\Gamma(1-\alpha)\Gamma(\theta+3\alpha)} \frac{\Gamma(m+1-\alpha)\Gamma(\theta+3\alpha)}{\Gamma(m+1-\alpha+\theta+3\alpha)} \\
&= \frac{\Gamma(1+\theta+2\alpha)}{\Gamma(1-\alpha)} \frac{\Gamma(m+1-\alpha)}{\Gamma(m+1+\theta+2\alpha)} \\
&= \frac{\Gamma(m+1-\alpha)}{\Gamma(1-\alpha)} \frac{\Gamma(1+\theta+2\alpha)}{\Gamma(m+1+\theta+2\alpha)} \\
&= \frac{(1-\alpha)_{(m)}}{(1+\theta+2\alpha)_{(m)}} = E\left(\left(\frac{P_{\tau_3}}{1-P_{\tau_1}-P_{\tau_2}}\right)^m\right)
\end{aligned}$$

Thus, we can conclude that

$$E \left( \left( \frac{P_{\tau_3}}{1 - P_{\tau_1} - P_{\tau_2}} \right)^m \right) = E(U_3^m) \quad \triangle$$

To conclude the proof, we now consider:

$$\begin{aligned}
E(P_{\tau_3}^m) &= E \left( (1 - P_{\tau_1} - P_{\tau_2})^m \left( \frac{P_{\tau_3}}{1 - P_{\tau_1} - P_{\tau_2}} \right)^m \right) \\
&= E_{\mathbf{P}, \tau_1, \tau_2} \left\{ E \left[ (1 - P_{\tau_1} - P_{\tau_2})^m \left( \frac{P_{\tau_3}}{1 - P_{\tau_1} - P_{\tau_2}} \right)^m \middle| \mathbf{P}, \tau_1, \tau_2 \right] \right\} \\
&= E_{\mathbf{P}, \tau_1, \tau_2} \left\{ (1 - P_{\tau_1} - P_{\tau_2})^m E \left[ \left( \frac{P_{\tau_3}}{1 - P_{\tau_1} - P_{\tau_2}} \right)^m \middle| \mathbf{P}, \tau_1, \tau_2 \right] \right\} \\
&= E \left( (1 - P_{\tau_1} - P_{\tau_2})^m \sum_{n=1}^{\infty} \left( \frac{P_n}{1 - P_{\tau_1} - P_{\tau_2}} \right)^m \frac{P_n}{1 - P_{\tau_1} - P_{\tau_2}} I_{n \neq \tau_1, n \neq \tau_2} \right) \\
&= E \left( (1 - P_{\tau_1} - P_{\tau_2})^m \sum_{n=1}^{\infty} \left( \frac{P_n}{1 - P_{\tau_1} - P_{\tau_2}} \right)^m \frac{P_n}{1 - P_{\tau_1} - P_{\tau_2}} I_{P_n \neq P_{\tau_1}, P_n \neq P_{\tau_2}} \right) \\
&= E \left( (1 - P_{\tau_1} - P_{\tau_2})^m \sum_{n=1}^{\infty} \left( \frac{P_n}{1 - P_{\tau_1} - P_{\tau_2}} \right)^{m+1} I_{P_n \neq P_{\tau_1}, P_n \neq P_{\tau_2}} \right) \\
&= E \left( (1 - V_1 - V_2)^m \sum_{n=1}^{\infty} \left( \frac{P_n}{1 - V_1 - V_2} \right)^{m+1} I_{P_n \neq V_1, P_n \neq V_2} \right) \quad \text{By the argument on page 2} \\
&= E \left( (1 - V_1 - V_2)^m \sum_{n=3}^{\infty} \left( \frac{V_n}{1 - V_1 - V_2} \right)^{m+1} \right) \quad \text{By the argument on page 2} \\
&= E \left( (1 - U_1 - (1 - U_1)U_2)^m \sum_{n=3}^{\infty} \left( \frac{\prod_{i=1}^{n-1} (1 - U_i) U_n}{1 - U_1 - (1 - U_1)U_2} \right)^{m+1} \right) \\
&= E \left( (1 - U_1)^m (1 - U_2)^m \sum_{n=3}^{\infty} \left( \frac{\prod_{i=1}^{n-1} (1 - U_i) U_n}{(1 - U_1)(1 - U_2)} \right)^{m+1} \right) \\
&= E \left( (1 - U_1)^m (1 - U_2)^m \sum_{n=3}^{\infty} \left( \prod_{i=3}^{n-1} (1 - U_i) U_n \right)^{m+1} \right) \\
&= E((1 - U_1)^m (1 - U_2)^m) E \left( \sum_{n=3}^{\infty} \left( \prod_{i=3}^{n-1} (1 - U_i) U_n \right)^{m+1} \right) \quad \text{As the } U_i' \text{s are independent} \\
&= E((1 - U_1)^m (1 - U_2)^m) E \left( \sum_{n=3}^{\infty} \left( \frac{\prod_{i=1}^{n-1} (1 - U_i) U_n}{(1 - U_1)(1 - U_2)} \right)^{m+1} \right) \\
&= E((1 - U_1)^m (1 - U_2)^m) E \left( \sum_{n=3}^{\infty} \left( \frac{\prod_{i=1}^{n-1} (1 - U_i) U_n}{1 - U_1 - (1 - U_1)U_2} \right)^{m+1} \right) \\
&= E((1 - U_1)^m (1 - U_2)^m) E \left( \sum_{n=3}^{\infty} \left( \frac{V_n}{1 - V_1 - V_2} \right)^{m+1} \right)
\end{aligned}$$

$$\begin{aligned}
&= E((1 - U_1)^m(1 - U_2)^m) E\left(\sum_{n=3}^{\infty} \left(\frac{V_{(n)}}{1 - V_1 - V_2}\right)^{m+1}\right) \\
&= E((1 - U_1)^m(1 - U_2)^m) E\left(\left(\frac{P_{\tau_3}}{1 - P_{\tau_1} - P_{\tau_2}}\right)^m\right) \quad \text{See page 2} \\
&= E((1 - U_1)^m(1 - U_2)^m) E(U_3^m) \quad \text{by } \triangle \\
&= E((1 - U_1)^m(1 - U_2)^m U_3^m) \\
&= E(((1 - U_1)(1 - U_2)U_3)^m) \\
&= E(V_3^m) \quad \text{By the definition of } V_3.
\end{aligned}$$

Thus, since  $m$  was arbitrary, for all  $m \in \mathbb{N}$ ,  $E(P_{\tau_3}^m) = E(V_3^m)$ . The one parameter case follows by setting  $\alpha = 0$ , concluding the proof.