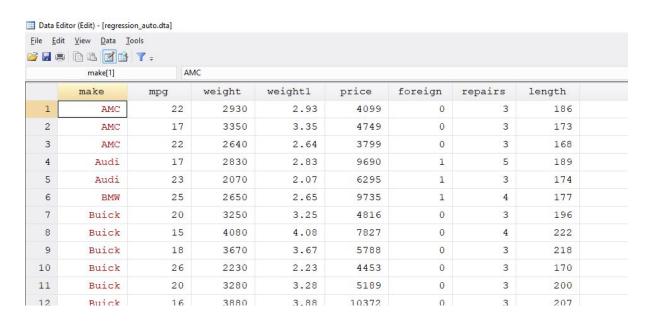
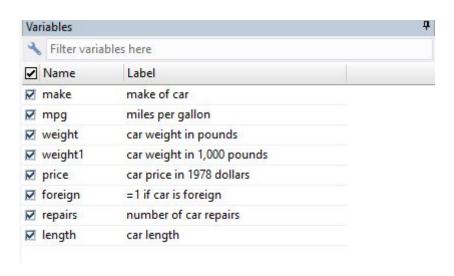
Multiple linear regression

We will use a car data set and our dependent variable is mpg (miles per gallon) vs our independent variables car weight and length





Description of our response variable and predictor variable but let's first create our global variables

```
global ylist mpg
global xlist weightl length
```

| . describe \$yl | list \$xli | st | | | | | | | |
|---|------------------------------------|---|---|---|----------------------------|----------------------------------|------------------|--|--|
| | storage | displa | ay val | ue | | | | | |
| variable name | type | forma | t lab | el | variable label | | | | |
| mpg | byte | %8.0g | | miles per gallon | | | | | |
| weight1 | float | %9.0g | | | car weight in 1,000 pounds | | | | |
| length | int | %8.0g | | | car length | | | | |
| . summarize \$ | ylist \$xl | ist | | | | | | | |
| Variable | Variable 0 | | Mean | Std. | Dev. | Min | Max | | |
| mpg | | 26 | 20.92308 | 4.757 | 504 | 14 | 35 | | |
| | | | | | | | | | |
| | | | | | 794 | 2.02 | 4.33 | | |
| weight1 length . summarize \$5 | /list, de | 26 26 | 3.099231 190.0769 | .6950 18.17 | | 2.02 163 | 4. 33 222 | | |
| weightl length | | 26 26 tail | 3.099231 | .6950 18.17 | | | | | |
| weightl length | | 26 26 tail | 3.099231 190.0769 er gallon | .6950 18.17 | | | | | |
| weightl length . summarize \$y | | 26 26 tail miles po | 3.099231 190.0769 er gallon | .6950 18.17 | | | | | |
| weight1 length summarize \$5 | iles | 26 26 tail miles pe | 3.099231 190.0769 er gallon | .6950 18.17 | | | | | |
| weightl length summarize \$5 Percenti | iles 14 | 26 26 tail miles pe | 3.099231 190.0769 er gallon st 14 | .6950 18.17 | | | | | |
| weightl length summarize \$5 Percenti 1% 5% | iles 14 14 | 26 26 tail miles pe | 3.099231 190.0769 er gallon st 14 14 | .6950 18.17 | 014 | 163 | | | |
| weightl length . summarize \$5 Percenti 1% 5% 10% | iles 14 14 15 | 26 26 tail miles pe | 3.099231 190.0769 er gallon st 14 14 15 | .6950 18.17 | 014 | 26 | | | |
| weightl length summarize \$9 Percenti 1% 5% 10% | iles 14 14 15 17 | 26 26 tail miles pe | 3.099231 190.0769 er gallon st 14 14 15 | .6950 18.17 Obs Sum of W | 014 | 26 26 | | | |
| weightl length summarize \$y Percenti 1% 5% 10% 25% | iles 14 14 15 17 | 26 26 tail miles personales | 3.099231 190.0769 er gallon st 14 14 15 | .6950 18.17 Obs Sum of W | 014 | 26 26 20.92308 | | | |
| weightl length summarize \$y Percenti 1% 5% 10% | iles 14 14 15 17 | 26 26 tail miles personales | 3.099231 190.0769 er gallon st 14 14 15 16 | .6950 18.17 Obs Sum of W | 19t. | 26 26 20.92308 | | | |
| weightl length . summarize \$y Percenti 1% 5% 10% 25% | iles 14 14 15 17 21 | 26 26 tail miles pe Smalle: | 3.099231 190.0769 er gallon st 14 14 15 16 | .6950 18.17 Obs Sum of W Mean Std. Dev | 19t. | 26 26 20.92308 4.757504 | | | |

Question

Can the weight and length statistically significantly predict a car's miles per gallon?

Hypothesis

H0: Car weight and length can not statistically significantly predict a car's miles per gallon Ha: Car weight and length can statistically significantly predict a car's miles per gallon

The level of significance

alpha = 0.05

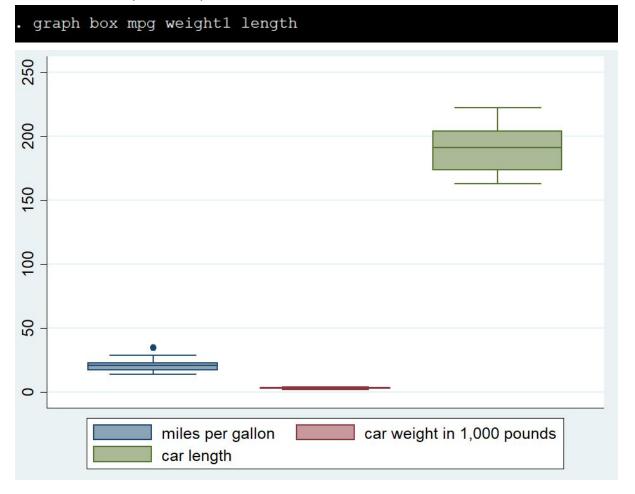
ASSUMPTIONS

Determine if data meets requirements to perform a linear regression.

Assumption #1: Your response variable should be measured on a continuous scale.

Assumption #2: You have two or more independent variables, which should be measured at the continuous or categorical level.

Assumption #3: There should be no significant outliers, high leverage points or highly influential points, which represent observations in your data set that are in some way unusual. Lets first plot a box plot:



As we can see from our box plot, we have an extreme mpg value. Let's find it and drop the observation.

```
egen Q1_mpg= pctile(mpg), p(25)
 egen Q3_mpg= pctile(mpg), p(75)
 egen IC_mpg= iqr(mpg)
gen touse=1 if (mpg< Q1_mpg-1.5*IC_mpg| mpg> Q3_mpg+1.5*IC_mpg) & missing(mpg)==0
(25 missing values generated)
recode touse . =0
(touse: 25 changes made)
tab touse
                            Percent
                  Freq.
                     25
                               96.15
                                           96.15
                               3.85
                                          100.00
                     26
                              100.00
 \star or use extremes command as follows
 extremes mpg, iqr(1.5)
   obs:
           iqr:
                  mpg
    24.
          2.000
 *Drop the outliers
 drop if (mpg< Q1_mpg-1.5*IC_mpg | mpg> Q3_mpg+1.5*IC_mpg)
(1 observation deleted)
```

Assumption #4: You should have independence of observations (i.e., independence of residuals), which you can check in Stata using the Durbin-Watson statistic.

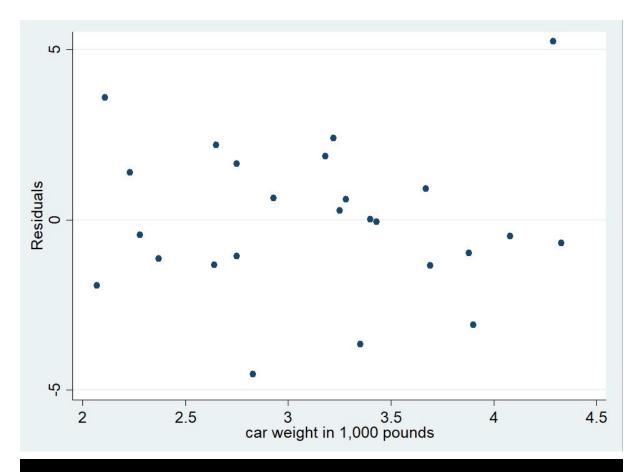
```
//generate time variable
 gen t = n
tsset t
      time variable: t, 1 to 25 delta: 1 unit
 *Run regression first before getting Durbin-Watson statistic as shown below
 * Simple regression
 reg $ylist $xlist
                                                 Number of obs
                                                                          25
                                                 F(2, 22)
                                                                       23.17
     Model
              243.943195
                                 2 121.971597
                                                                     0.0000
               115.816805
                                22 5.26440024
                                                R-squared
                                                                      0.6781
                                                 Adj R-squared =
                                                                     0.6488
                  359.76
                                24
                                         14.99
                                                 Root MSE
                                                                      2.2944
       mpg
              -3.382251 1.582302
    weight1
                                              0.044
                                                       -6.663745
                                                                  -.1007576
                                      -0.92
              -.0552394
                          .0598333
                                              0.366
                                                       -.1793262
                                                                    .0688473
    length
               41.54354
                          7.317217
                                       5.68
                                              0.000
                                                        26.36856
                                                                    56.71852
 dwstat
                                25) = 1.616561
Ourbin-Watson d-statistic( 3,
```

Values of 1.5 < d < 2.5 generally show that there is no autocorrelation in the data while values

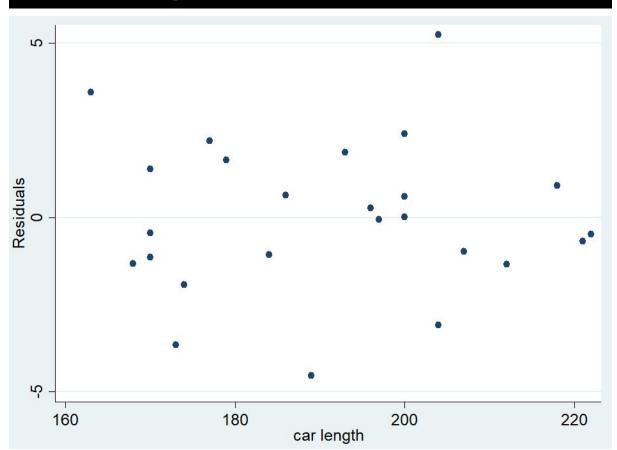
0 to 2< means there is positive autocorrelation and values >2 to 4 means there is negative autocorrelation.

Assumption #5: There needs to be a linear relationship between (a) the dependent variable and each of your independent variables, and (b) the dependent variable and the independent variables collectively. Checking the linearity assumption is not so straightforward in the case of multiple regression as is in simple linear regression. One thing to do is to plot the standardized residuals against each of the predictor variables in the regression model. If there is a clear nonlinear pattern, there is a problem of nonlinearity.

```
    scatter ehat weight1
```

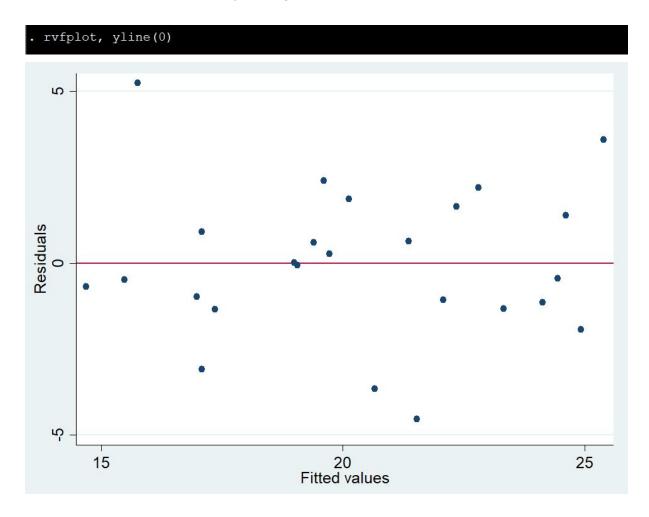


. scatter ehat length



The two residual versus predictor variable plots above do not indicate strongly a clear departure from linearity

Assumption #5: Your data needs to show homoscedasticity, which is where the variances along the line of best fit remain similar as you move along the line. One of the main assumptions for the ordinary least squares regression is the homogeneity of variance of the residuals. A commonly used graphical method is to plot the residuals versus fitted (predicted) values. We do this by issuing the **rvfplot** command.



Now let's look at a couple of commands that test for heteroscedasticity.

```
estat imtest
Cameron & Trivedi's decomposition of IM-test
              Source
                             chi2
                                      df
                                       5
  Heteroskedasticity
                            12.42
                                            0.0294
                             6.69
                                            0.0353
                                       2
            Skewness
            Kurtosis
                             0.12
                                       1
                                            0.7339
               Total
                            19.23
                                       8
                                            0.0137
 estat hettest
Breusch-Pagan / Cook-Weisberg test for heteroskedasticity
         Ho: Constant variance
         Variables: fitted values of mpg
                            0.08
         chi2(1)
         Prob > chi2 =
                          0.7777
```

The first test on heteroskedasticity given by **imest** is the White's test and the second one given by **hettest** is the Breusch-Pagan test. Both test the null hypothesis that the variance of the residuals is homogenous. Therefore, if the p-value is very small, we would have to reject the hypothesis and accept the alternative hypothesis that the variance is not homogenous. So in this case, the evidence is not against the null hypothesis that the variance is homogeneous. These tests are very sensitive to model assumptions, such as the assumption of normality. Therefore it is a common practice to combine the tests with diagnostic plots to make a judgment on the severity of the heteroscedasticity and to decide if any correction is needed for heteroscedasticity.

Assumption #6: Your data must not show multicollinearity, which occurs when you have two or more independent variables that are highly correlated with each other. The term collinearity implies that two variables are near perfect linear combinations of one another. When more than two variables are involved it is often called multicollinearity, although the two terms are often used interchangeably.

We can use the **vif** command after the regression to check for multicollinearity. **vif** stands for *variance inflation factor*.

| . vif | | | |
|-------|----------------|--------------|----------------------|
| Var | iable | VIF | 1/VIF |
| | ength ight1 | 5.17 5.17 | 0.193490 0.193490 |
| Mean | n VIF | 5.17 | |

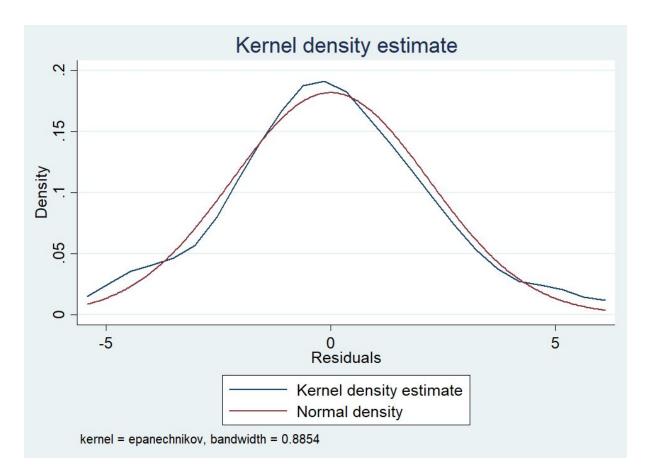
As a rule of thumb, a variable whose VIF values are greater than 10 may merit further investigation. Tolerance, defined as 1/VIF, is used by many researchers to check on the degree of collinearity. A tolerance value lower than 0.1 is comparable to a VIF of 10. It means that the variable could be considered as a linear combination of other independent variables.

Assumption #8: The residuals (errors) should be approximately normally distributed.

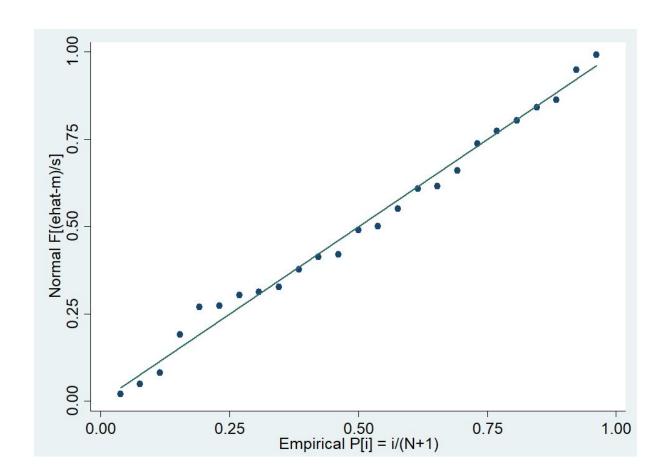
```
reg Şylist Şxlist
                                                Number of obs
                                                                          25
                                                 F(2, 22)
                                                                       23.17
                                2 121.971597
              243.943195
    Model
                                                                      0.0000
              115.816805
                                22 5.26440024
                                                 R-squared
                                                                      0.6781
                                                 Adj R-squared
                                                                      0.6488
                  359.76
                                24
                                         14.99
                                                 Root MSE
                                                                      2.2944
                                                        [95% Conf. Interval]
      mpg
              -3.382251
                          1.582302
                                      -2.14
                                              0.044
                                                       -6.663745
              -.0552394
                          .0598333
                                      -0.92
                                                       -.1793262
    length
                                              0.366
                                                                    .0688473
               41.54354
                          7.317217
                                       5.68
                                              0.000
                                                        26.36856
                                                                    56.71852
predict ehat, resid
```

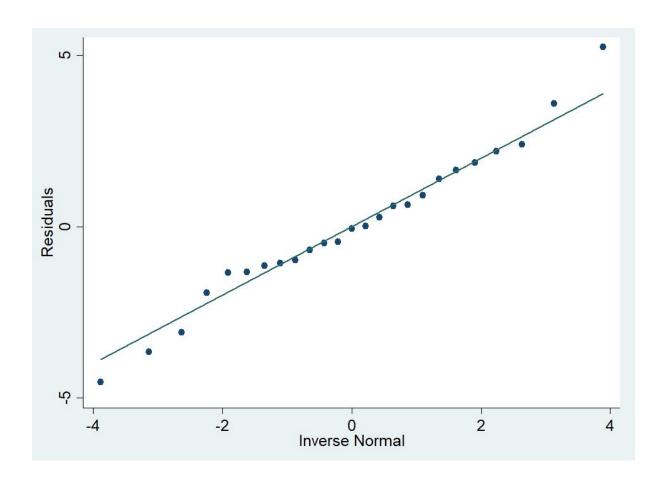
Below we use the **kdensity** command to produce a kernel density plot with the **normal** option requesting that a normal density be overlaid on the plot. **kdensity** stands for kernel density estimate. It can be thought of as a histogram with narrow bins and moving average.

```
. *kdensity command to produce a kernel density plot . kdensity ehat, normal (n\,() set to 25)
```

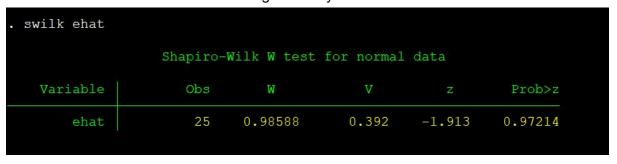


The **pnorm** command graphs a standardized normal probability (P-P) plot while **qnorm** plots the quantiles of a variable against the quantiles of a normal distribution. **pnorm** is sensitive to non-normality in the middle range of data and **qnorm** is sensitive to non-normality near the tails. As you see below, the results from **pnorm** show no indications of non-normality, while the **qnorm** command shows a slight deviation from normal at the upper tail, as can be seen in the **kdensity** above. Nevertheless, this seems to be a minor and trivial deviation from normality. We can accept that the residuals are close to a normal distribution.





There are also numerical tests for testing normality.



As seen above Shapiro-Wilk W test is not significant which means the residuals (errors) are approximately normally distributed.

Multiple Linear Regression

| Source | SS | df | MS | | Number of obs | | 25 |
|----------|------------------|-----------|-----------|--------|----------------------|-----|-----------|
| Model | Model 243.943195 | | 121.97159 | | F(2, 22) Prob > F | | 23.17 |
| Residual | 115.816805 | 2 22 | 5.2644002 | | guared | | 0.6781 |
| | | | | | R-squared | | 0.6488 |
| Total | 359.76 | 24 | 14.9 | 99 Roo | t MSE | | 2.2944 |
| mpg | Coef. | Std. Err. | t | P> t | [95% Cor | nf. | Interval] |
| weight1 | -3.382251 | 1.582302 | -2.14 | 0.044 | -6.66374 | 5 | 1007576 |
| length | 0552394 | .0598333 | -0.92 | 0.366 | 1793262 | 2 | .0688473 |
| cons | 41.54354 | 7.317217 | 5.68 | 0.000 | 26.3685 | 6 | 56.71852 |

A multiple regression was run to predict mpg (miles covered by the car per gallon) from car weight, its price and whether it's foreign or not. These variables statistically significantly predicted mpg (miles covered by the car per gallon), F(2, 22) = 23.17, p < .0005, $R^2 = .678$. But we can see that car weight is the only significant predictor. The regression equation was: predicted miles covered by the car per gallon = 41.544 - 3.382* (weight of the car)-0.055* (Car Length)

```
. correlate $ylist $xlist
(obs=25)

mpg weightl length

mpg 1.0000
weightl -0.8158 1.0000
length -0.7818 0.8981 1.0000
```

