



POLITECNICO
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Bayesian Models for Prediction on Real-Estate Data

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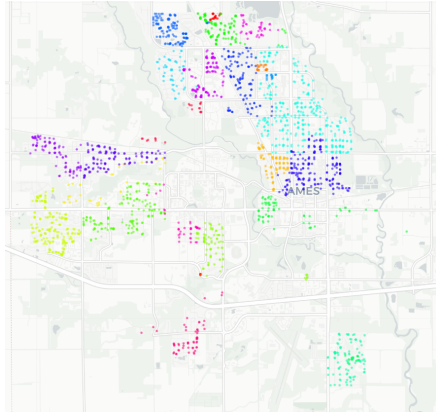


Figure: Houses in the dataset colored by neighbourhood

Model Averaging

- Candidate models: CAR, Linear Mixed Effects and Gaussian Markov Random Fields
- We want to avoid overconfidence in the choice of the model structure
- We follow a principled approach for Model Averaging

Richard McElreath. *Statistical rethinking, a Bayesian course with examples in R and Stan*, pages 195–205.
2015

$$\forall i \in \{\text{CAR}, \text{LME}, \text{GMRF}\}$$

$$\text{dWAIC}_i = \text{WAIC}_i - \min_{j \in \{\text{CAR}, \text{LME}, \text{GMRF}\}} \text{WAIC}_j$$

$$w_i = \frac{\exp\left(-\frac{1}{2}\text{dWAIC}_i\right)}{\sum_{j=1}^m \exp\left(-\frac{1}{2}\text{dWAIC}_j\right)}$$

$$Y_{\text{pred, ENSEMBLE}} = w_{\text{CAR}} Y_{\text{pred, CAR}} + w_{\text{LME}} Y_{\text{pred, LME}} + w_{\text{GMRF}} Y_{\text{pred, GMRF}}$$

Richard McElreath. *Statistical rethinking, a Bayesian course with examples in R and Stan*, pages 195–205.

2015

Conditionally Autoregressive Model

$$\forall i = 1, \dots, p \quad \forall j = 1, \dots, N \quad \forall k = 1, \dots, K$$

$$Y_j | b, \underline{\beta}, \text{Neigh}_j, \underline{x}_j, \underline{\phi}, \nu^2 \sim_{\text{ind}} N(b + \underline{\beta}^T \underline{x}_j + \phi_{\text{Neigh}_j}, \nu^2)$$

$$\beta_i \sim_{\text{iid}} N(0, 10)$$

$$b \sim N(4.5, 0.01)$$

$$\nu^2 \sim \text{Inv-Gamma}(0.1, 0.1)$$

$$\phi_k | \phi_{-k} \sim N \left(\frac{\rho \sum_{l=1}^K a_{kl} \phi_l}{\rho \sum_{l=1}^K a_{kl} + 1 - \rho}, \frac{\tau^2}{\rho \sum_{l=1}^K a_{kl} + 1 - \rho} \right)$$

$$\tau^2 \sim \text{Inv-Gamma}(0.1, 0.1)$$

$$\rho \sim \text{Uniform}(0, 1)$$

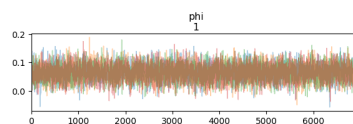
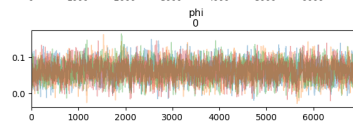
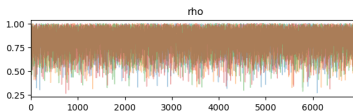
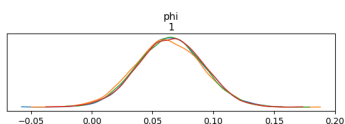
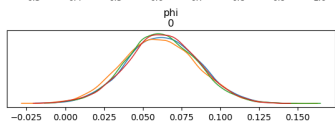
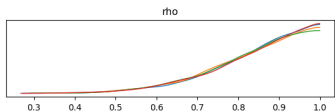
$[a_{kl}]_{k,l \in \{1, \dots, K\}}$ is the neighbourhood adjacency matrix

Conditionally Autoregressive Model

Model fitting and evaluation

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	Mean	StdDev	5%	50%	95%	N _{eff}	\hat{R}
ρ	0.8483	0.1181	0.6164	0.8745	0.9889	16696	1.001
ϕ_0	0.0620	0.0233	0.0239	0.0617	0.1005	2000	1.002
ϕ_1	0.0642	0.0271	0.0199	0.0642	0.1089	3057	1.001



Linear Mixed Effects Model

$$\forall i = 1, \dots, p \quad \forall j = 1, \dots, N \quad \forall k = 1, \dots, K$$

$$Y_j | \underline{X}_j, \text{Neigh}_j, \underline{\beta}, \underline{\gamma}, \sigma^2 \stackrel{\text{ind}}{\sim} \mathcal{N}(\underline{\beta}^T \underline{X}_j + \underline{\gamma}_{\text{Neigh}_j}, \sigma^2)$$

$$\beta_i | \tau \stackrel{\text{iid}}{\sim} \text{half-}\mathcal{N}(0, \tau^2)$$

$$\gamma_k | \xi \stackrel{\text{iid}}{\sim} \mathcal{N}(0, \xi^2)$$

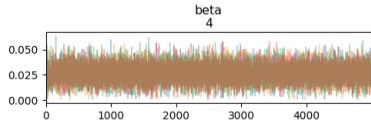
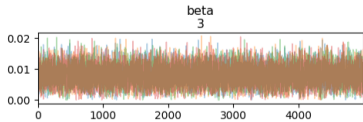
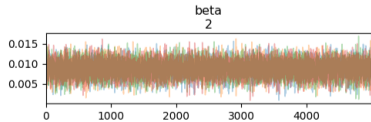
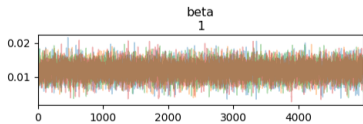
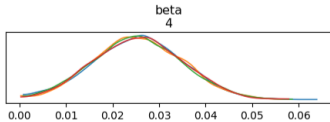
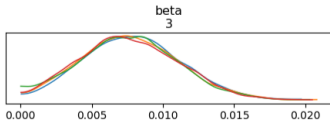
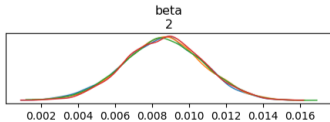
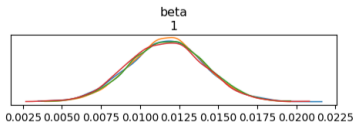
$$\sigma^2 \sim \text{InvGamma}(1, 30)$$

$$\tau^2 \sim \text{InvGamma}(1, 0.05)$$

$$\xi^2 \sim \text{InvGamma}(1, 0.05)$$

Linear Mixed Effects Model Fitting

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Gaussian Markov Random Fields Model

$$Y_j \mid \underline{\beta}, w, \sigma_e^2 \sim \mathcal{N} \left(\underline{X}_j^T \underline{\beta} + w(s), \sigma_e^2 \right)$$

$$\beta_i \sim_{\text{iid}} \mathcal{N}(0, 1)$$

$$\sigma_e^2 \sim \text{Log-Gamma}(0.1, 0.1)$$

$$w(s) \mid \theta \sim \mathcal{N}(0, \Sigma(\theta))$$

$$w(s) \text{ is the solution of } (\kappa^2 - \Delta)^{\alpha/2} (\tau w(s)) = \mathcal{W}(s)$$

$$\log(\kappa) = \log(\kappa_0) - \theta_2$$

$$\log(\tau) = \log(\tau_0) - \theta_1 + \nu\theta_2$$

$$\theta_1 \sim \mathcal{N}(0, 1)$$

$$\theta_2 \sim \mathcal{N}(0, 10)$$

$$\Sigma(s; \theta) = \frac{\sigma^2}{2^{\nu-1}\Gamma(\nu)} (\kappa \|s\|)^{\nu} K_{\nu}(\kappa \|s\|)$$

- We fitted the model using the Integrated Nested Laplace Approximation.
- The calibration of hyperparameters of the covariance function has been done fixing $\alpha = 2$, $\nu = 1$, $\kappa_0 = \frac{\sqrt{8\nu}}{\rho}$ and $\tau_0 = \frac{1}{\sqrt{4\pi\kappa_0}\sigma}$ and choosing $\rho = 4.6$ and $\sigma = 1.5$ in order to minimize the WAIC.
- This method computes an approximation of the marginal posterior distributions for all the parameters and hyperparameters.

Given the gaussian vector $\underline{\xi} = (\underline{\beta}, \underline{w})$ we can approximate the posterior distribution of $\underline{\theta} = (\theta_1, \theta_2)$:

$$\pi(\underline{\theta}|\underline{y}) \propto \frac{\pi(\underline{\theta}, \underline{\xi}, \underline{y})}{\pi(\underline{\xi}|\underline{\theta}, \underline{y})} \bigg|_{\underline{\xi}=\underline{\xi}^*(\underline{\theta})} \approx \frac{\pi(\underline{\theta}, \underline{\xi}, \underline{y})}{\tilde{\pi}(\underline{\xi}|\underline{\theta}, \underline{y})} \bigg|_{\underline{\xi}=\underline{\xi}^*(\underline{\theta})}$$

where $\tilde{\pi}(\underline{\xi}|\underline{\theta}, \underline{y})$ is a gaussian approximation of $\pi(\underline{\xi}|\underline{\theta}, \underline{y})$ and $\underline{\xi}^*(\underline{\theta}) = (\pi(\underline{\xi}|\underline{\theta}, \underline{y}))$.

Using this method we can compute the approximation of the posterior of the parameters:

$$\pi(\theta_k|\underline{y}) \approx \int \tilde{\pi}(\underline{\theta}|\underline{y}) d\underline{\theta}_{-k}$$
$$\pi(\xi_j|\underline{y}) \approx \int \tilde{\pi}(\xi_j|\underline{\theta}, \underline{y}) \tilde{\pi}(\underline{\theta}|\underline{y}) d\underline{\theta}$$

Model Assessment

We computed the ensemble model from the three fitted models

$$\begin{aligned} \text{WAIC}_{\text{CAR}} &= -5717 & \text{WAIC}_{\text{LME}} &= -3929 & \text{WAIC}_{\text{GMRF}} &= -5526 \\ w_{\text{CAR}} &\simeq 1 & w_{\text{LME}} &\simeq 0 & w_{\text{GMRF}} &\simeq 0 \end{aligned}$$

We conclude that the ensemble model coincides with the CAR model; its performance on the test set is

$$\sqrt{\text{MSE}_{\text{ensemble}}} \simeq 27.000 \$$$

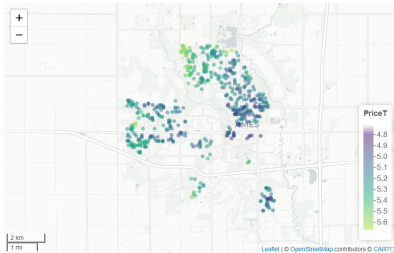


Figure: Prediction on the test set

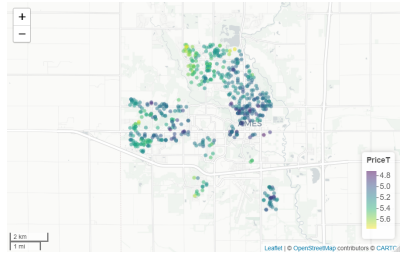


Figure: True $\log(\text{price})$ of the test set

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