# STAT 578 - Advanced Bayesian Modeling - Fall 2019 Assignment 4

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## Solution for Problem (a)

(i)

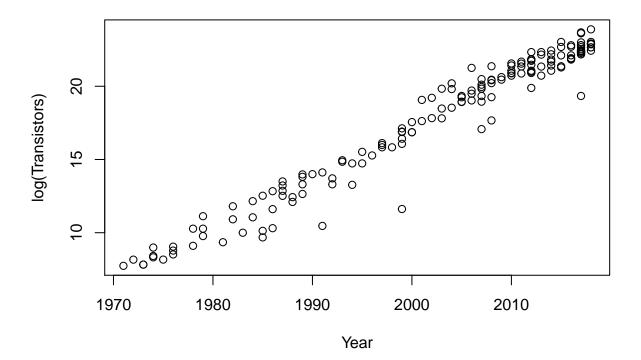
According to the formula for Moore's law and take log on T, we get:

$$log(T) \approx log(C*2^{A/2}) \rightarrow log(T) \approx log(C) + 0.3465736*A$$

This is the form of linear regression that make logT as response variable and A as explanatory variable. log(C) is the intercept and coefficient of A is 0.3465736. Thus we can say logT is roughly follow a simple linear regression on A.

(ii)

```
moores_df = read.csv("mooreslawdata.csv", header=TRUE)
with(moores_df, plot(log(Transistors) ~ Year))
```



#### Solution for Problem (b)

(i)

```
We first use classical linear model to estimate the parameters.
```

```
df_lm <- list(y = log(moores_df$Transistors),</pre>
          x_centered = moores_df$Year - mean(moores_df$Year))
model_lm = lm(y ~ x_centered, data=df_lm)
summary(model_lm)
##
## Call:
## lm(formula = y ~ x_centered, data = df_lm)
## Residuals:
##
      Min
               1Q Median
                               ЗQ
## -5.0598 -0.4443 0.0891 0.6052 2.1820
##
## Coefficients:
               Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) 17.726825
                          0.074044 239.41
                                             <2e-16 ***
                                      62.93
                                             <2e-16 ***
## x_centered
              0.342122
                          0.005436
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.9395 on 159 degrees of freedom
## Multiple R-squared: 0.9614, Adjusted R-squared: 0.9612
## F-statistic: 3960 on 1 and 159 DF, p-value: < 2.2e-16
```

- $\beta$  estimates are about 0.3 to 18, we choose to set initial  $\beta$  values at  $\pm 200$
- Regression error variance  $\sigma^2$  estimate is about  $(0.94^2) \approx 0.9$ , we choose to set initial  $\sigma^2$  values of 0.01 and 100

```
model {
    x_bar <- mean(x)
    for (i in 1:length(y)) {
        y[i] ~ dnorm(beta1 + beta2* (x[i] - x_bar), sigmasqinv)
    }

    beta1 ~ dnorm(0, 1e-06)
    beta2 ~ dnorm(0, 1e-06)
    sigmasqinv ~ dgamma(0.001, 0.001)
    sigmasq <- 1/sigmasqinv
}</pre>
```

```
moores_model <- jags.model("moores.bug", df_jags, initial_vals, n.chains = 4)</pre>
update(moores_model, 1000)
coef_sample <- coda.samples(moores_model, c("beta1","beta2","sigmasq"), n.iter=2000)</pre>
gelman.diag(coef_sample, autoburnin=FALSE)
## Potential scale reduction factors:
##
##
           Point est. Upper C.I.
## beta1
                    1
## beta2
                     1
                                1
## sigmasq
##
## Multivariate psrf
##
## 1
Gelman-Rubin statistic for \beta 1, \beta 2 and \sigma^2 are all 1, thus we can declare convergence of them.
(ii)
summary(coef_sample)
##
## Iterations = 1001:3000
## Thinning interval = 1
## Number of chains = 4
## Sample size per chain = 2000
##
## 1. Empirical mean and standard deviation for each variable,
##
      plus standard error of the mean:
##
##
                          SD Naive SE Time-series SE
           17.7271 0.074781 0.0008361
                                             8.362e-04
## beta1
## beta2
            0.3421 0.005456 0.0000610
                                             6.014e-05
## sigmasq 0.8924 0.099704 0.0011147
                                             1.111e-03
## 2. Quantiles for each variable:
##
##
              2.5%
                                        75%
                                               97.5%
                        25%
                                50%
          17.5780 17.6780 17.7269 17.7779 17.8729
## beta1
## beta2
            0.3313 0.3385 0.3422 0.3457 0.3527
## sigmasq 0.7193 0.8231 0.8833 0.9550 1.1128
(iii)
beta2_sample = as.matrix(coef_sample)[,"beta2"]
Mean of slope,
mean(beta2_sample)
```

## [1] 0.3420985

```
95% posterior credible interval of slope,
```

```
quantile(beta2_sample, c(0.025, 0.975))
##
       2.5%
               97.5%
## 0.331320 0.352727
The interval contains value (0.3465736) determined in part (a).
(iv)
beta1_sample = as.matrix(coef_sample)[,"beta1"]
Mean of intercept,
mean(beta1_sample)
## [1] 17.72707
95% posterior credible interval of intercept,
quantile(beta1_sample, c(0.025, 0.975))
       2.5%
##
               97.5%
## 17.57802 17.87290
Solution for Problem (c)
(i)
model {
  x_bar \leftarrow mean(x)
  for (i in 1:length(y)) {
    y[i] ~ dnorm(beta1 + beta2* (x[i] - x_bar), sigmasqinv)
  beta1 ~ dnorm(0, 1e-06)
  beta2 ~ dnorm(0, 1e-06)
  sigmasqinv ~ dgamma(0.001, 0.001)
  y_predict ~ dnorm(beta1 + beta2*(year_predict - x_bar), sigmasqinv)
  year_start <- x_bar - (beta1 / beta2)</pre>
predict_model <- jags.model("moores_predict.bug",</pre>
                              c(as.list(df_jags), year_predict = 2020),
                              initial_vals, n.chains = 4)
update(predict_model, 1000)
predict_sample <- coda.samples(predict_model, c("y_predict", "year_start"), n.iter=2000)</pre>
gelman.diag(predict_sample, autoburnin=FALSE)
## Potential scale reduction factors:
##
               Point est. Upper C.I.
##
## y_predict
                        1
                                    1
## year_start
                        1
                                    1
##
```

```
## Multivariate psrf
##
## 1
```

Gelman-Rubin statistic for the prediction is 1, thus we can declare convergence of them.

(ii)

```
summary(predict_sample)
##
## Iterations = 1001:3000
## Thinning interval = 1
## Number of chains = 4
## Sample size per chain = 2000
##
## 1. Empirical mean and standard deviation for each variable,
##
      plus standard error of the mean:
##
##
                   Mean
                             SD Naive SE Time-series SE
## y_predict
                  23.84 0.9490
                                 0.01061
                                                 0.010571
## year_start 1950.25 0.8586
                                0.00960
                                                 0.009488
## 2. Quantiles for each variable:
##
##
               2.5%
                        25%
                                 50%
                                          75%
                                                 97.5%
                 22
                       23.2
                                        24.49
                                                 25.72
## y_predict
                               23.84
## year_start 1949 1949.7 1950.26 1950.85 1951.89
(iii)
95% posterior predictive interval for the transistor count, in billions,
exp(quantile(as.matrix(predict_sample)[,"y_predict"], c(0.025, 0.975))) / (10^9)
##
                     97.5%
          2.5%
##
     3.588029 147.620426
(iv)
To explain, we can assume the year when transistor is in invented is the year for transistor count equal to 1.
This means given our model,
log(T) = \beta 1 + \beta 2 * (A_i - \bar{A})
We can set T = 1 to derive A_i as the year when transistor is invented. Therefore, we have,
A_i = \bar{A} - \beta 1/\beta 2
Using the classical linear model, we can estimate this year as,
round(mean(moores_df$Year) - coef(model_lm)[[1]]/coef(model_lm)[[2]])
## [1] 1950
```

The 95% posterior interval for this quantity.

```
round(quantile(as.matrix(predict_sample)[,"year_start"], c(0.025, 0.975)))
## 2.5% 97.5%
## 1949 1952
```

Note: According to Wikipedia, the first working device to be built was a point-contact transistor invented in 1947 by American physicists John Bardeen, Walter Brattain, and William Shockley at Bell Labs. Our estimates is very close.

# Solution for Problem (d)

(i)

```
X <- model.matrix(model_lm)
Nsim <- length(beta1_sample)
error_sim <- matrix(NA, Nsim, nrow(X))
post_beta <- as.matrix(coef_sample)[, c("beta1", "beta2")]
post_simgasq <- as.matrix(coef_sample)[, c("sigmasq")]

for(s in 1:Nsim) {
   error_sim[s,] <- df_jags$y - X %*% cbind(post_beta[s,])
}</pre>
```

(ii)

```
error_rep <- matrix(NA, Nsim, nrow(X))

for(s in 1:Nsim) {
   error_rep[s,] <- rnorm(nrow(X), 0, sqrt(post_simgasq[s]))
}</pre>
```

(iii)

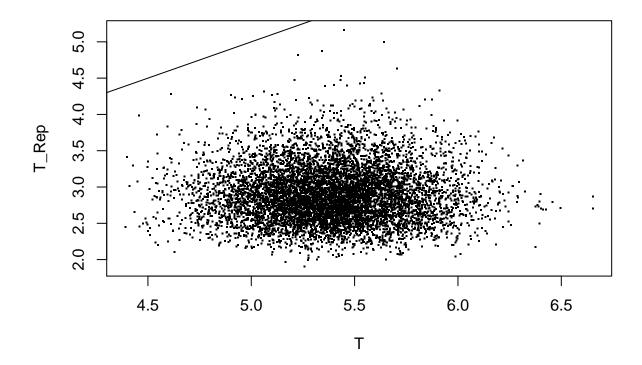
```
error_sim_std <- matrix(NA, Nsim, nrow(X))
error_rep_std <- matrix(NA, Nsim, nrow(X))

for(s in 1:Nsim) {
    error_sim_std[s,] = error_sim[s,] / sqrt(post_simgasq[s])
    error_rep_std[s,] = error_rep[s,] / sqrt(post_simgasq[s])
}

T <- apply(abs(error_sim_std), 1, max)
T_Rep <- apply(abs(error_rep_std), 1, max)</pre>
```

(iv)

```
plot(T_Rep ~ T, pch=".", cex=2)
abline(a=0,b=1)
```



(v)

```
mean(T_Rep >= T)
```

## ## [1] 0

p-value is 0 and this is the strong evidence for an outlier.

(vi)

The most extreme outlier is,

```
moores_df[unique(apply(abs(error_sim_std), 1, which.max)),]
```

```
## Processor Year Transistors
## 49 ARM 9TDMI 1999 111000
```

which is "ARM 9TDMI" produced in 1999 that only has 111,000 transistors. In this time period, other processors have tens of millions of transistors.