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Inferring individual-level processes from population-level patterns in cultural evolution


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Our species is characterized by a great degree of cultural variation, both within and between populations. Understanding how group-level patterns of culture emerge from individual-level behaviour is a long-standing question in the biological and social sciences. We develop a simulation model capturing demographic and cultural dynamics relevant to human cultural evolution, focusing on the interface between population-level patterns and individual-level processes. The model tracks the distribution of variants of cultural traits across individuals in a population over time, conditioned on different pathways for the transmission of information between individuals. From these data, we obtain theoretical expectations for a range of statistics commonly used to capture population-level characteristics (e.g. the degree of cultural diversity). Consistent with previous theoretical work, our results show that the patterns observed at the level of groups are rooted in the interplay between the transmission pathways and the age structure of the population. We also explore whether, and under what conditions, the different pathways can be distinguished based on their group-level signatures, in an effort to establish theoretical limits to inference. Our results show that the temporal dynamic of cultural change over time retains a stronger signature than the cultural composition of the population at a specific point in time. Overall, the results suggest a shift in focus from identifying the one individual-level process that likely produced the observed data to excluding those that likely did not. We conclude by discussing the implications for empirical studies of human cultural evolution.

1. Introduction

Our species is characterized by a great degree of cultural variation, both within and between populations. This variation can be observed in domains as different as material culture (e.g. [1]), linguistic features (e.g. [2]) or social norms (e.g. [3]). These population-level patterns are the aggregate product of underlying individual strategies. Understanding how those patterns of culture emerge from individual-level behaviour is a long-standing question in the biological and social sciences [4].

The field of cultural evolution encompasses efforts to answer this question, through a variety of theoretical and empirical tools (see [5] for a recent review). Cultural evolution is the process of change in the frequency of different variants of a cultural trait over time. Cultural traits comprise the knowledge, ideas, beliefs, skills, attitudes or any other form of information that can be socially transmitted between individuals, for example, through teaching or imitation [6]. Seminal early contributions to the field focused on human culture (e.g. [7,8]), but the scope now extends also to non-human culture [9].

Overall, this body of work has identified a number of factors shaping cultural variation within and across human groups, including the pathways for the transmission of information between individuals, demography, shared population history and adaptation to environmental conditions. However, we lack a systematic understanding of the effect of these different factors and their interactions. In particular, we lack a conceptual framework explicitly mapping the patterns observed at the level of groups onto the underlying processes occurring at the level of individuals [5]. For example, a large body of empirical work in cultural evolution builds on the notion of the relative ‘conservativeness’ of vertical transmission (i.e. parent to child) compared to other transmission pathways (e.g. horizontal transmission, i.e. between peers) [7]. This body of work includes field-based investigations (e.g. [10–12]) and cross-cultural studies (e.g. [12–15]). But how exactly does the group-level ‘signature’ of vertical transmission differ from that of other transmission pathways? And is it sufficiently different that it can be mapped, unequivocally, onto vertical transmission?

To address these and related questions, we develop a simulation model capturing demographic and cultural dynamics relevant to human cultural evolution, focusing on the interface between population-level patterns and individual-level processes (see [16–19] for other modelling frameworks). By design, the simulation model is the simplest it can be. For example, we focus on neutral cultural traits (i.e. not linked to fitness) and constant population sizes; traits only differ in the way they are transmitted between individuals. In this respect, our work differs from related recent analyses by Fogarty *et al.* [18], who focused on more complex demographic scenarios. They explored the effect of trait transmission on the age-structure of a population, and the extent to which cultural traits affecting demographic change can spread. The general modelling framework, however, is similar to the one introduced in [18].

In other words, our model does not seek to replicate a specific cultural system. Rather, it is used to run ‘artificial experiments’ for various demographic and cultural scenarios, in order to explore and analyse the ranges of possible evolutionary outcomes generated by those scenarios. Specifically, the model tracks the distribution of variants of cultural traits across a population over time, conditioned on different pathways for the transmission of information between individuals. From these simulated data, we obtain theoretical expectations, in the form of probability distributions, for a range of statistics commonly used in the literature to capture population-level characteristics (e.g. the degree of cultural diversity).

Our aim is twofold. First, we aim to derive general insights about the different transmission pathways and their signatures at the level of the group. Second, we aim to establish whether, and under what conditions, the different pathways can be distinguished based on their group-level signatures. Are the signatures sufficiently different that they can be traced back to specific individual-level processes? The rationale is that if we cannot accurately infer underlying scenarios from simulated data, which are generated under known ‘experimental’ conditions, it is unlikely that we will be able to do so based on empirical data.

Some additional context may help grasp the significance of the second aim. Researchers in anthropology and related disciplines often draw inferences from data on variation in cultural traits within a population, but these data are typically sparse in space and/or time. Our analysis aims to establish the theoretical limits to inference—in other words, how much information can in fact be ‘extracted’ from the data under simplified scenarios of cultural change captured by our simulation framework. These theoretical limits represent an upper-bound to inference: we cannot expect to obtain more information from empirical data, which are the product of intricate real-world scenarios.

We begin by describing the mathematical framework and the statistical analysis (§2), followed by presentation of the results (§3). The code used to generate the results is available from

Table 1. Transmission modes.

mode (notation)	description	note
vertical (<i>v</i>)	from parent to offspring	
horizontal (<i>h</i>)	between individuals in the same age group	
oblique (<i>o</i>)	from older to younger individuals	excludes parent–offspring pairs
age-neutral (<i>n</i>)	between any two individuals in the population	equivalent to random drift
mixed (<i>m</i>)	a combination of horizontal and oblique	see text (§2.1.3)

<http://datadryad.org/resource/doi:10.5061/dryad.14q31>. We conclude with a non-technical summary of the key findings, together with discussion of their implications (§4); readers can skip to this part without loss of insight.

2. Methods

Our analysis involves two components. The first is a mathematical framework simulating the temporal dynamic of cultural change in age-structured populations. The second is a set of statistics used to summarize the simulation output; these are used to determine whether the group-level patterns extracted from the simulation output can be reliably distinguished.

The framework includes both discrete and continuous cultural traits: discrete traits assume a finite number of variants, continuous traits assume any value in a given interval. For ease of presentation, this and the next section focus on the discrete case. An outline of key results for the continuous case is in §3.2.4, with additional detail in the electronic supplementary material, section S3.3.4.

2.1. Mathematical framework

Cultural evolution in age-structured populations has been modelled in a variety of contexts, including the spread of the use of birth control [20] and sex-ratio imbalance [21,22]. Furthermore, Fogarty *et al.* [18] investigated the implications of cultural transmission on human demographic changes, in particular changes in population size and age structure, with a framework similar to the one introduced below. Here we keep the demographic properties of the population constant over time and focus on quantifying the effects of different forms of cultural transmission on the cultural composition of the population.

In detail, we consider a simple age-structured population, divided into five age classes, where each individual possesses a variant of five different cultural traits. Each age group comprises an interval of 10 years, i.e. roughly a generation.

2.1.1. Cultural traits and transmission modes

Discrete cultural traits assume one of five possible variants. As the traits are assumed to be neutral, individuals do not benefit from carrying a specific variant over another.

Following [7], we define a set of modes for the transmission of trait variants between individuals (table 1). Broadly, the modes describe the flow of information within and between generations. They delimit the set of potential interaction partners within a population, with interactions between individuals within this set occurring at random.

Additionally, transmission biases may steer interactions towards individuals carrying a particular variant [8]. A large number of biases have been identified in the literature [23]; we focus on conformity bias, i.e. a preference for the most common variant in the population, as detailed in §2.1.4.

We consider a population of N individuals. At each point in time, individual i is described by a variable I_x^i , with $I_x^i \in \{1, \dots, 5\}$ for $x = v, h, o, n, m$ (table 1). This is the individual's current cultural make-up, comprising five cultural traits, each linked to a transmission mode. For all traits, an individual acquires the variant of its parent at birth, i.e. through vertical transmission. For all traits except the one linked to mode v , the variant may change through further transmission events during the individual's lifetime, as described in §2.1.3.

Table 2. Transmission rules for modes h, o, n in the absence of mutation.

mode	possible interaction partner j for individual i	transmission rule
h	$I_{\text{age}}^i = I_{\text{age}}^j$	$I_h^i = I_h^j$
o	$I_{\text{age}}^i > I_{\text{age}}^j$	$I_o^i = I_o^j$
n	any individual	$I_n^i = I_n^j$

2.1.2. Demographic processes

In addition to variable I_x^i , at each point in time individual i is described by a variable I_{age}^i , denoting its current age group. There are five age groups, covering the age range 0–50. The age structure of the population is summarized by $[n_1, \dots, n_5]$, where n_i describes the number of individuals in age group i .

Each time step in the simulation corresponds to 10 years. At each time step, individuals age by moving to the next age group, and they may either die or reproduce. We assume a constant population size N . Therefore, the number of individuals entering the population through reproduction in a given time step is determined by the difference between N and the number of deaths in the previous time step. Individuals die with probability p_{death} . All individuals in age group 5 die with $p_{\text{death}} = 1$ in the following time step.

Reproduction occurs asexually. Individuals in age groups 2 and 3 are chosen at random at each time step to reproduce until population size N is reached. These individuals produce offspring who inherit their trait variants. However, with a small probability μ a mutation occurs and the offspring's variant differs from the parent's, as described in §2.1.3.

These demographic processes produce a pyramid-shaped age structure, with fewer individuals in older generations. In general, it holds $\sum_{i=1}^5 n_i = N$. However, due to the stochastic nature of the framework, it is possible that all potential parents die before reproducing, leading to extinction of the population. The results presented below are based on simulation runs where the population survived for the entire duration of the simulation.

2.1.3. Mutation and transmission events

We define a mutation rate μ , which describes the fidelity of the transmission process and applies to all transmission events. For all traits, individual i acquires the variant of its parent with probability $1 - \mu$, and a randomly selected variant with probability μ . This process describes vertical transmission (table 1).

Additionally, for traits linked to modes h, o, n the variant may change through further transmission events during the individual's lifetime. Specifically, at each time step individual i engages with probability p_w in interaction with a randomly selected individual j . For example, $p_w = 0.5$ corresponds to an average of 2.5 transmission events during the individual's lifetime, $p_w = 1$ to an average of 5 events. During each transmission event, individual i acquires the variant of individual j with probability $1 - \mu$, and a randomly selected variant with probability μ . The transmission rules in table 2 apply to these interaction events; they show that horizontal, oblique and age-neutral transmission differ only in the set of potential interaction partners (§2.1.1).

The 'pure' modes v, h, o, n provide a useful baseline against which to compare more realistic scenarios. Specifically, one of the five traits in an individual's cultural make-up is linked to mode m , whereby possible interaction partners include both peers and older individuals (table 1). We term this mode 'mixed', to indicate that it is effectively the combination of horizontal and oblique transmission. For this trait, individual i engages in horizontal transmission with probability $(1 - p_{\text{mix}})$ and in oblique transmission with probability p_{mix} , following the rules specified in table 2. Qualitatively, the higher the value of p_{mix} , the more likely that transmission is oblique, i.e. that the individual's interaction partner belongs to an older age group, rather than to the same age group.

Across all modes, the interaction probability p_w and the mutation rate μ , combined, determine the potential for cultural change. Broadly, we distinguish between scenarios with low versus high potential for cultural change (e.g. low μ , low p_w versus high μ , high p_w). For instance, the higher p_w and μ , the more opportunities there are for an individual's cultural make-up to change over its lifetime, and hence for the cultural composition of the population to change with each time step.

2.1.4. Conformity bias

As noted in §2.1.1, the transmission modes describe different sets of potential interaction partners, but interactions between individuals within each set occur at random. This implies that the probability p_k that an individual interacts with a partner carrying variant k is proportional to the relative frequency of variant k in the set. It holds

$$p_k = \frac{n_k^*}{N^*},$$

where n_k^* describes the number of individuals in the set of size N^* carrying variant k .

Conformist transmission is defined as the disproportional adoption of common variants [8]. With conformity bias, the probability of an individual interacting with a partner carrying variant k is given instead by [24,25]

$$p_k = \frac{n_k^*}{N^*} + b \left(\hat{k} \frac{n_k^*}{N^*} - 1 \right). \quad (2.1)$$

The parameter $b \geq 0$ describes the strength of the bias and \hat{k} is the number of variants present in the population. Conformity increases the probability that an individual carrying variant k acts as interaction partner if the frequency of variant k exceeds the ‘relative’ majority $1/\hat{k}$.

2.1.5. Simulation set-up

At the beginning of each simulation run, the N individuals in the population are distributed randomly across the five age groups. They all carry the same variants of the cultural traits. In each subsequent time step, the demographic and cultural processes described above take place over time leading to changes in the frequency of the different variants, and thus to the cultural composition of the population.

Individuals update their cultural make-ups asynchronously, i.e. an individual’s interaction partner may have already updated its cultural make-up in the given time step. The order in which individuals engage in transmission events is randomized in every time step.

A single simulation run consists of a burn-in phase followed by 200 time steps. We explore various constellations of the parameters for ranges $N = 25, 50, 100$; $\mu = 0.01, 0.05, 0.1$; $p_w = 0.5, 0.75, 1$; $b = 0, 0.01, 0.02, 0.03$. Note that $b = 0$ for all analyses, except those investigating specifically the effect of conformity on distinguishability of the different transmission modes.

2.2. Statistical analysis

2.2.1. Summary statistics

The mathematical framework tracks the frequencies of trait variants at each time step. We use this information to characterize the cultural composition of the population and the dynamic of cultural change over time, conditioned on the different transmission modes. Specifically, with $x = v, h, o, n, m$ denoting the mode (table 1), for a given trait we derive

- (i) the frequencies of its five variants $[p_1^x, p_2^x, \dots, p_5^x]$ across individuals in the population, where p_k^x describes the relative frequency of variant k ,
- (ii) the frequency of the most common variant in the population, denoted p_{\max}^x ,
- (iii) the total number of variants present in the population, denoted k^x ,
- (iv) the level of cultural diversity, as measured by the Shannon diversity index

$$d_s^x = - \sum_{k=1}^{k^x} p_k^x \log(p_k^x),$$

where k^x describes the number of variants present and p_k^x the relative frequency of variant k , as defined above, and

- (v) the average time a variant stays the most common variant, denoted t_{\max}^x .

The first statistic is the joint probability distribution of the frequencies of the five cultural variants. This captures the most information about the cultural composition of the population at a given point in time, but researchers may not have access to the full data. Accordingly, we also explore how much information can be extracted from ‘partial’ data, namely the frequency of the most common variant, p_{\max}^x , and the total number of variants present, k^x .

Researchers often summarize the cultural composition of a population using a diversity measure such as the Shannon diversity index (e.g. [1,26,27]). We include this statistic in our analysis to investigate how the dimension reduction this involves affects our ability to distinguish between transmission modes.

Finally, we characterize the temporal dynamic of cultural change with t_{\max}^x , which measures the rate of change of the most common variant in the population. Broadly, this gives an indication of how fast the cultural composition of the population can change over time. We use t_{\max}^x to explore whether the temporal dynamic is more informative than the cultural composition of the population at a given point in time, as described by the other statistics.

2.2.2. Distinguishability analysis

We explore the behaviour and inferential power of each statistic by generating probability distributions of the statistic, across simulation runs, for different values of the parameters in the model.

For a given parameter constellation, we determine whether two modes can be distinguished based on the statistic by (i) examining the area of overlap between the corresponding distributions (e.g. vertical versus horizontal transmission) and (ii) determining the probability that a particular transmission mode acted in the population to produce an observed value of the statistic. The approach can be applied more generally to determine the effect of a given parameter. This is done by comparing the distributions for a given mode under different values of the parameter (e.g. oblique transmission with versus without conformity, or vertical transmission with small versus large population size), keeping constant the values of the other parameters. For ease of presentation, we focus here on comparison of two modes for a given parameter constellation. For simplicity, we describe the procedure for one-dimensional probability distributions (i.e. the distributions for p_{\max}^x , k^x , d_s^x and t_{\max}^x), but the same approach applies to the joint distributions $f_{\text{joint},x}$.

We take the area of overlap between the distributions to indicate to what degree the two corresponding modes can be distinguished based on the statistic without knowledge of an empirical estimate. At one extreme, no overlap suggests that the modes can be reliably distinguished; at the other, complete overlap suggests that they cannot be distinguished.

In detail, the area of overlap O_{xy} between probability distributions f_x and f_y with one intersection point only is determined using the Kolmogorov–Smirnov distance K_{xy} between the associated distribution functions F_x and F_y (figure 1a). If z^* denotes the value where $K_{xy} = \sup_z |F_x(z) - F_y(z)|$ is realized, then it holds [28]

$$O_{xy} = 1 - F_y(z^*) + F_x(z^*) = 1 + (F_x(z^*) - F_y(z^*)) = 1 - K_{xy}.$$

Distinguishability is defined using a threshold value \bar{O} : two distributions f_x and f_y are distinguishable if it holds $O_{xy} < \bar{O}$, and we use a value of $\bar{O} = 20\%$, corresponding to the widely accepted 80% power cut-off. An alternative approach, based on the Rayleigh criterion, is introduced in the electronic supplementary material, section S1.

Specifically, based on 30 000 simulation runs, we determine (i) the joint distributions $f_{\text{joint},x}$ of frequencies of the five trait variants and (ii) the probability distributions $f_{p_{\max}^x}$, f_{k^x} , $f_{d_s^x}$, $f_{t_{\max}^x}$ of statistics p_{\max}^x , k^x , d_s^x , t_{\max}^x , conditioned on transmission modes $x = v, h, o, n, m$. For a given statistic, we calculate the areas of overlap O_{xy} for pairs of distributions x and y , and we compare their values against the threshold $\bar{O} = 20\%$ to determine whether the corresponding modes are distinguishable based on the statistic.

This procedure rests on an *a priori* definition of distinguishability, without considering available empirical data. The results provide general expectations about the similarity or dissimilarity of population-level outcomes generated by the different transmission modes. Even in the absence of empirical data, these expectations can inform researchers as to whether the corresponding statistics p_{\max}^x , k^x , d_s^x , t_{\max}^x carry a signature of the underlying transmission modes.

If researchers do have access to empirical estimates of a statistic, then the procedure can be extended to incorporate this information. This is illustrated with an example in figure 1a. Values of the statistic in regions A and C are almost ‘unique’ to modes x and y , respectively, whereas values in region B are ‘shared’. The interpretation is that values of the statistic in regions A and C could only have been produced by one of the two modes, whereas values in region B could have been produced by either mode. It follows that empirical estimates provide no additional information in cases where the area of overlap O_{xy} is close to the extremes 0 and 1 (corresponding to no overlap and complete overlap of the distributions, respectively).

In detail, given the empirical estimate of a statistic, denoted by z , the aim is to determine the probability that a given transmission mode acted in the population to produce this value. We define the set of possible transmission modes $\Gamma = \{v, h, o, n, m\}$ and assume that one of these did produce the

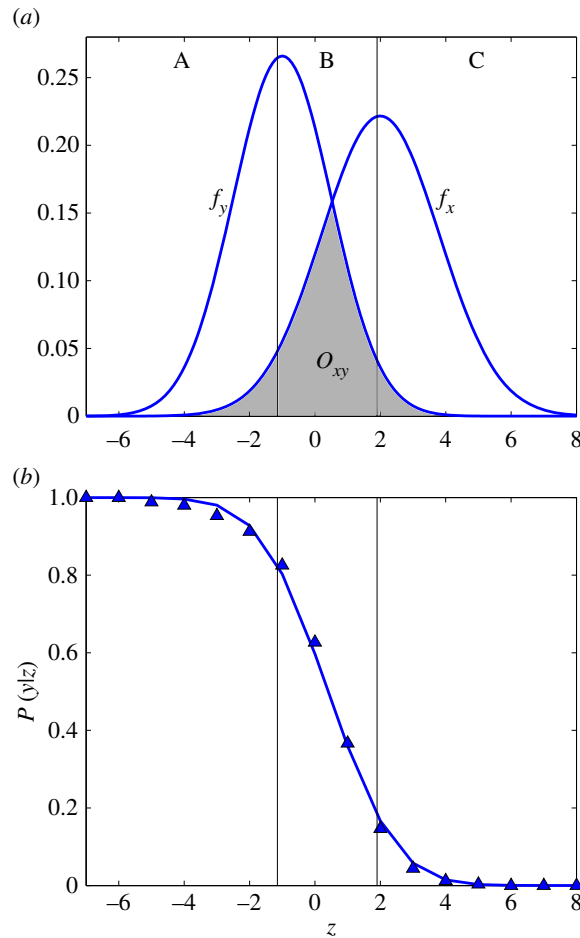


Figure 1. Illustration of (a) the area of overlap O_{xy} between probability distributions f_x and f_y , and (b) the probability $P(y|z)$ that transmission mode y (as opposed to x) acted in the population to generate value z .

observed value of the statistic. Bayes' theorem then results in

$$P(y|z) = \frac{f(z|y)P(y)}{f(z)} \quad \text{with } y \in \Gamma, \quad (2.2)$$

where $P(y|z)$ stands for the probability that mode y is acting in the population given z (see blue triangles in figure 1b). The probability distribution $f(z|y)$ is the distribution generated by the simulation framework; it captures the values of the statistic that can be assumed for mode x . The probability $P(y)$ includes all prior information about the likelihood that mode y acted in the population; in the absence of prior information we assume that $P(y) = 1/|\Gamma|$. The function $f(\cdot)$ describes the probability distribution of the values of the statistic and it holds $f(z) = \sum_{y \in \Gamma} P(y)f(z|y)$.

For comparison of two transmission modes, the probability $P(y|z)$ can be determined by logistic regression (see blue line in figure 1b). Multinomial logistic regression can be used for more than two modes. Probabilities close to 0 or 1 have a clear interpretation: at 0, mode y could not have produced the observed value of the statistic; at 1, mode y is most likely to have produced the observed value, compared to the possible alternatives. Intermediate values indicate that multiple transmission modes could have produced the observed value of the statistic. Alternatively, receiver operating characteristic curve analysis could be applied (see the electronic supplementary material, section S2, for details).

3. Results

3.1. Population-level patterns

In this section, we explore whether the different transmission modes result in characteristic population-level outcomes under scenarios with low versus high potential for cultural change, as defined in §2.1.3.

An example of the analyses underlying the results is presented in the electronic supplementary material, section S3.1.

3.1.1. Pure modes

We begin by investigating the behaviour of the pure modes—vertical, horizontal, oblique and age-neutral transmission. To this end, we study the distributions of statistics p_{\max}^x , k^x , d_s^x , t_{\max}^x , conditioned on transmission modes $x = v, h, o, n$, for parameter constellations $N = 25, 50, 100$; $\mu = 0.01, 0.05, 0.1$; $p_w = 0.5, 0.75, 1$.

Statistics p_{\max}^x , k^x and d_s^x describe the cultural composition of the population at a given point in time (§2.2.1). The different transmission modes result in comparable population-level outcomes for these statistics under scenarios with low potential for cultural change. At the same time, outcomes for oblique transmission are strongly affected by the value of p_w . Specifically, under scenarios with high potential for cultural change oblique transmission leads to more homogeneous cultural compositions than vertical, horizontal and age-neutral transmission.

These differences in behaviour across transmission modes are rooted in the different sets of potential interaction partners. In age-neutral transmission, an individual can interact with any other individual in the population. In horizontal transmission, an individual in age group k can only interact with the $n_k - 1$ individuals in its own age group (corresponding to age-neutral transmission within the age group). In oblique transmission, an individual in age group k can interact with $\sum_{j=k+1}^5 n_j$ older individuals, where n_j is the number of individuals in age group j (corresponding to age-neutral transmission within all age groups older than the individual's). In other words, compared to age-neutral and horizontal transmission, oblique transmission is characterized by a disproportional influence of older individuals. Individuals in age group 5 are potential interaction partners for all individuals in age group 4 and younger, whereas individuals in age group 2 are potential interaction partners only for individuals in age group 1. As a result, a variant at high frequency in age group 5 tends to 'trickle down' to the other age groups; over time, it tends to become increasingly common, eventually leading towards homogenization of the population as a whole.

To illustrate the influence of the age structure on the transmission dynamic, we derive the probability that a mutant variant (i.e. a variant with frequency 1) has frequency $0, 1, 2, \dots, N$ after one time step (electronic supplementary material, section S3.2). For oblique transmission, the spread probability of a mutant variant is highly influenced by the age of the individual introducing it into the population: the older the individual, the higher the probability that the variant is still present in the population after one time step. By contrast, for horizontal transmission the spread probability does not vary greatly with the age of the individual introducing it into the population. By definition, the same is true for age-neutral transmission.

How does this translate into population-level outcomes? Figure 2 shows the level of cultural diversity d_s^x for the different modes in a scenario with high potential for cultural change (parameter constellation $N = 100$, $\mu = 0.1$, $p_w = 1$), separately in each of the five age groups and in the population as a whole. For oblique transmission, the level of diversity within age groups is comparable to the level of diversity in the whole population. This suggests that the age groups are culturally more homogeneous, as expected based on the trickle-down effect. The same pattern applies to age-neutral transmission; by definition, this mode is not affected by the age structure of the population. By contrast, for horizontal transmission the level of diversity within age groups is substantially lower than the level of diversity in the whole population. This suggests that the different age groups sustain different cultural variants.

A final insight relates to the relative rate of cultural change for the different transmission modes. Specifically, the modes can be ranked by ordering the distribution functions of t_{\max}^x (electronic supplementary material, section S3.1.1). As expected, vertical transmission leads to the slowest rate of change, as there are limited opportunities for transmission compared to the other modes (see the electronic supplementary material, figure S1). Furthermore, oblique transmission is characterized by a slower rate of change than horizontal and age-neutral transmission, due to the disproportional influence of older age groups and the consequent homogenization of the population.

3.1.2. Mixed mode

The results for mixed transmission reveal a consistent pattern (electronic supplementary material, section S3.1.2). At low values of p_{mix} , the occasional oblique transmission event may introduce variants otherwise absent in a given age group. This leads to an increase in cultural diversity at the population level compared with pure horizontal transmission. At high values of p_{mix} , the occasional horizontal

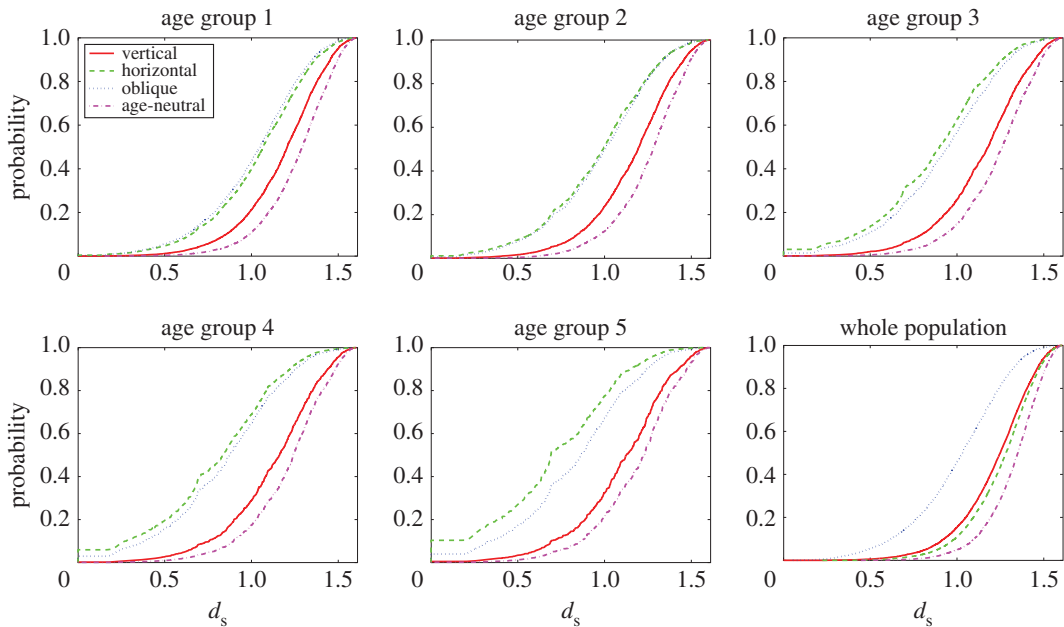


Figure 2. Level of cultural diversity in the five age groups and in the whole population for the pure transmission modes. Shown are the distribution functions of statistic d_s^x for the different modes, under parameter constellation $N = 100$, $\mu = 0.1$, $p_w = 1$.

transmission event effectively dampens the trickle-down effect, weakening the disproportional influence of older age groups and the consequent homogenization of the population. This leads to an increase in cultural diversity at the population level compared with pure oblique transmission. In sum, at both extremes of p_{mix} the mixed mode results in higher levels of cultural diversity in the population than the ‘corresponding’ pure mode.

Similarly, at both extremes of p_{mix} the mixed mode results in a faster rate of cultural change than the ‘corresponding’ pure mode.

3.2. Inference

In the previous section, we have shown that the different transmission modes result in different patterns at the population level. Here, we investigate whether the differences are large enough to ensure distinguishability based on the statistics introduced in §2.2.1.

3.2.1. Pure modes

Following the procedure described in §2.2.2, we compare the probability distributions of the statistics for pairs of transmission modes x and y , with $x, y \in \{v, h, o, n\}$, and parameter constellations $N = 50$; $\mu = 0.01$; $p_w = 0.5$ and $N = 50$; $\mu = 0.1$; $p_w = 1$. Analysis of a larger set of parameters can be found in the electronic supplementary material, section S3.3.1. Qualitatively, comparison of vertical transmission to the other modes captures the effect of interactions that occur during the individuals’ lifetime. Comparison of age-neutral to oblique and horizontal transmission captures the effect of restricting the set of potential interaction partners to individuals within the same age group and to older individuals, respectively. Finally, comparison of horizontal to oblique transmission captures the effect of the age group of the interaction partner.

We start with analysing the area of overlap O_{xy} based on the joint probability distribution of the five variants of a trait (figure 3a). Under scenarios with low potential for cultural change (e.g. low μ , low p_w ; §2.1.3), the different transmission modes result in comparable joint probability distributions: O_{xy} approaches 1 (see the electronic supplementary material, figure S4). O_{xy} tends to decrease as the potential for cultural change increases (i.e. as μ and/or p_w increase). In particular, oblique transmission tends to deviate from the other modes. However, the differences are not large enough to ensure distinguishability based on an arbitrary threshold $\bar{O} = 0.2$ (see the values of O_{xy} for $f_{\text{joint},x}$ in table 3).

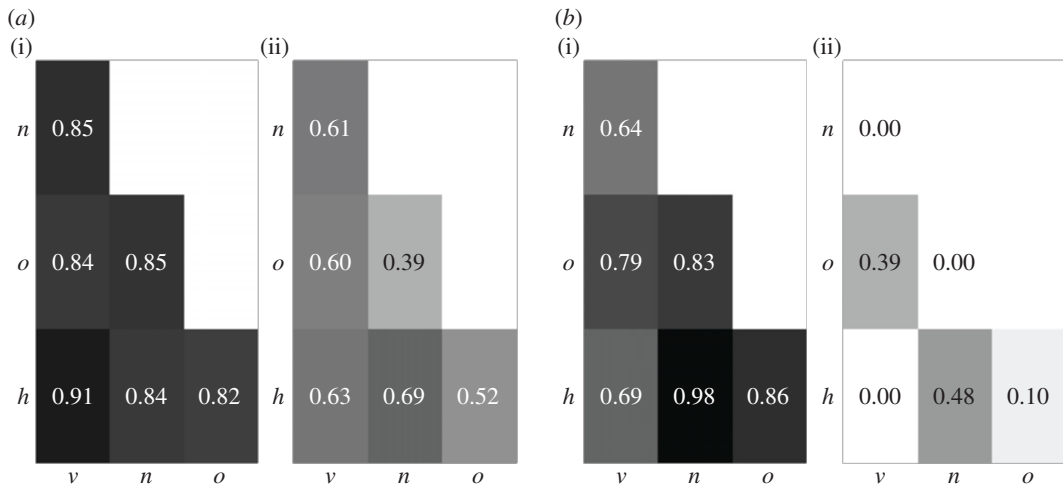


Figure 3. Distinguishability between the pure transmission modes v, h, o, n (table 1) based on (a) the joint probability distribution of the five variants of a trait in the population, $f_{\text{joint},x}$, and (b) the average time a variant stays the most common variant, t_{max}^x . Shown are the values of the area of overlap O_{xy} between the probability distributions of the statistic for pairs of modes, under parameter constellations $N = 50$; $\mu = 0.01$; $p_w = 0.5$ (a(i), b(i)) and $N = 50$; $\mu = 0.1$; $p_w = 1$ (a(ii), b(ii)). The corresponding grey scale ranges from white for $O_{xy} = 0$ to black for $O_{xy} = 1$.

Table 3. Distinguishability between the pure transmission modes based on selected statistics: the joint probability distribution of the five variants of a trait in the population, $f_{\text{joint},x}$, the level of cultural diversity, d_s^x , and the average time a variant stays the most common variant, t_{max}^x . Shown are the values of the area of overlap O_{xy} between the probability distributions of the statistic for pairs of modes, under parameter constellation $N = 50$, $\mu = 0.1$, $p_w = 1$. Values in italics indicate distinguishability based on threshold $\bar{O} = 0.2$.

modes	$f_{\text{joint},x}$	$f_{d_s^x}$	$f_{t_{\text{max}}^x}$
vertical versus horizontal	0.63	0.84	0
vertical versus oblique	0.60	0.73	0.39
vertical versus age-neutral	0.61	0.73	0
age-neutral versus horizontal	0.69	0.96	0.48
age-neutral versus oblique	0.39	0.57	0
horizontal versus oblique	0.52	0.58	0.10

As noted in §2.2.1, the joint probability distribution of the five variants of a trait captures the most information about the system at a given point in time. By contrast, statistics p_{max}^x , k^x rely on partial data, whereas d_s^x provides summary information. Unsurprisingly, then, the distributions of p_{max}^x , k^x , d_s^x carry a weaker signature of the underlying transmission modes than the joint probability distribution (e.g. compare the values of O_{xy} for $f_{\text{joint},x}$ versus $f_{d_s^x}$ in table 3; the corresponding plots are in the electronic supplementary material, section S3.3.1).

Figure 3b shows the behaviour of O_{xy} based on statistic t_{max}^x , the average time a variant stays the most common variant in the population (§2.2.1). Under scenarios with low potential for cultural change (e.g. low μ , low p_w), the different transmission modes result in comparable distributions of t_{max}^x . Under scenarios with high potential for cultural change (e.g. high μ , high p_w), four pairs (vertical versus horizontal, vertical versus age-neutral, age-neutral versus oblique, horizontal versus oblique) produce distributions of t_{max}^x with no overlap (i.e. $O_{xy} = 0$; table 3), indicating that they can be distinguished based on this statistic. In other words, these transmission modes result in temporal dynamics that are sufficiently different to be distinguishable.

Two pairs of modes result in comparable distributions (vertical versus oblique transmission, age-neutral versus horizontal transmission), with areas of overlap substantially greater than 0 (figure 3b). To obtain further insight, we calculate the conditional probabilities $P(y|z)$ given in equation (2.2) for these pairwise comparisons. Figure 4a shows the probabilities based on t_{max}^x . In the case of vertical

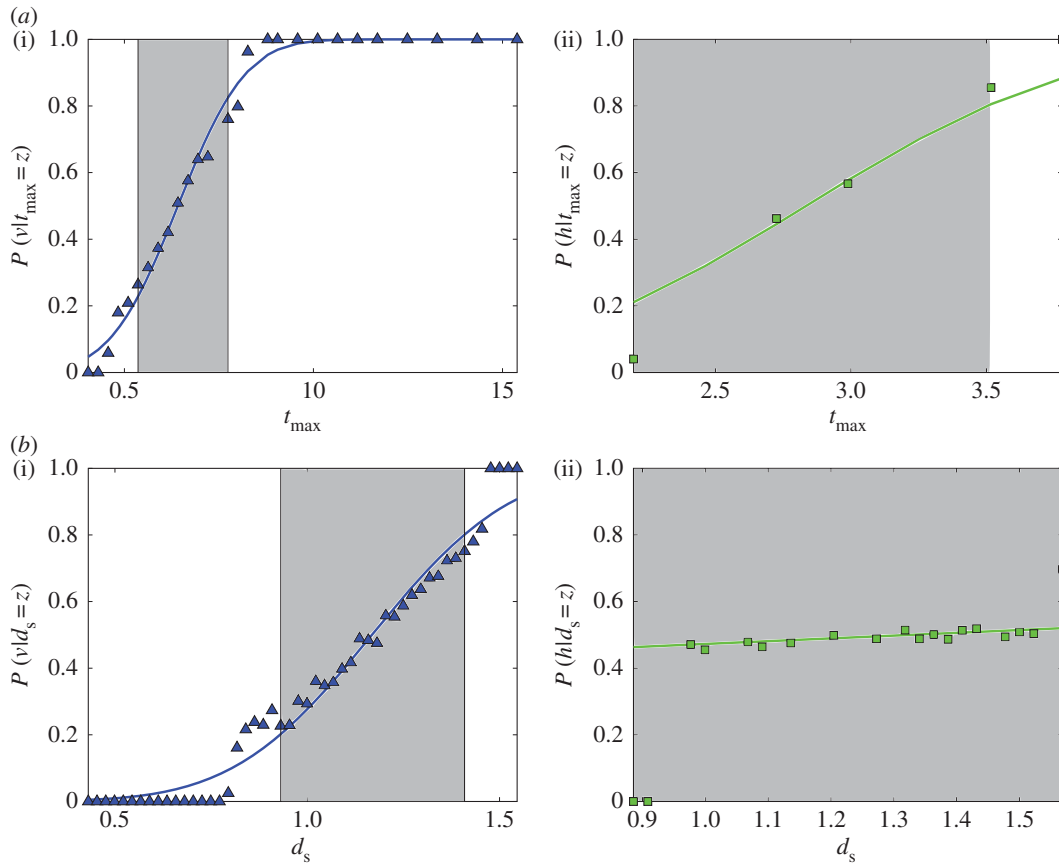


Figure 4. $P(y|z)$ for pairwise comparisons between vertical and oblique transmission (a(i), b(i), blue triangles) and age-neutral and horizontal transmission (a(ii), b(ii), green squares) based on (a) the average time a variant stays the most common variant, t_{\max}^x , and (b) the level of cultural diversity, d_s , under parameter constellation $N = 50$; $\mu = 0.1$; $p_w = 1$. The grey areas indicate the values of the statistic which could have been produced by either transmission mode with a probability greater than 0.2.

versus oblique transmission (figure 4a(i)), smaller values of empirical estimates of t_{\max} point to oblique transmission (cf. $P(v|t_{\max} = z) \approx 0$ for small values of t_{\max}), whereas larger values point to vertical transmission (cf. $P(v|t_{\max} = z) \approx 1$ for large values of t_{\max}). We note that in this context ‘small’ and ‘large’ refer only to the comparison between values generated by vertical and oblique transmission. As discussed in §3.1, oblique transmission produces larger values of t_{\max} than horizontal and age-neutral transmission (see the electronic supplementary material, figure S1d). The distinction between age-neutral and horizontal transmission remains ambiguous, however; almost all values of t_{\max} can be produced by either mode (figure 4a(ii)). Similar conclusions apply for the level of cultural diversity d_s (figure 4b).

In sum, the existence of an empirical estimate greatly improves our ability to distinguish between different transmission modes in situations where *a priori* definitions are not sufficient (i.e. in situations where the area of overlap between the corresponding distributions is larger than 0).

3.2.2. Mixed mode

The results for mixed transmission are in the electronic supplementary material, section S3.3.2. The aim in this case is to compare distributions of a statistic for (i) the mixed mode at a given level of p_{mix} to the pure modes and (ii) the mixed mode at a given level of p_{mix} to the mixed mode at other levels of p_{mix} .

Consistent with results for the pure modes, the temporal dynamic of cultural change carries more information about the underlying process than the cultural composition of the population at a specific point in time, as captured, respectively, by statistics t_{\max}^x and d_s^x . Overall, for both statistics the overlap between distributions tends to decrease as the potential for cultural change increases. However, only a limited number of pairs of modes result in distributions that can be distinguished based on the arbitrary threshold $\bar{O} = 0.2$, and only for t_{\max}^x . For instance, mixed transmission with $p_{\text{mix}} = 0.5$ can only be reliably distinguished from vertical transmission. Further, as expected at both extremes of p_{mix} the mixed mode

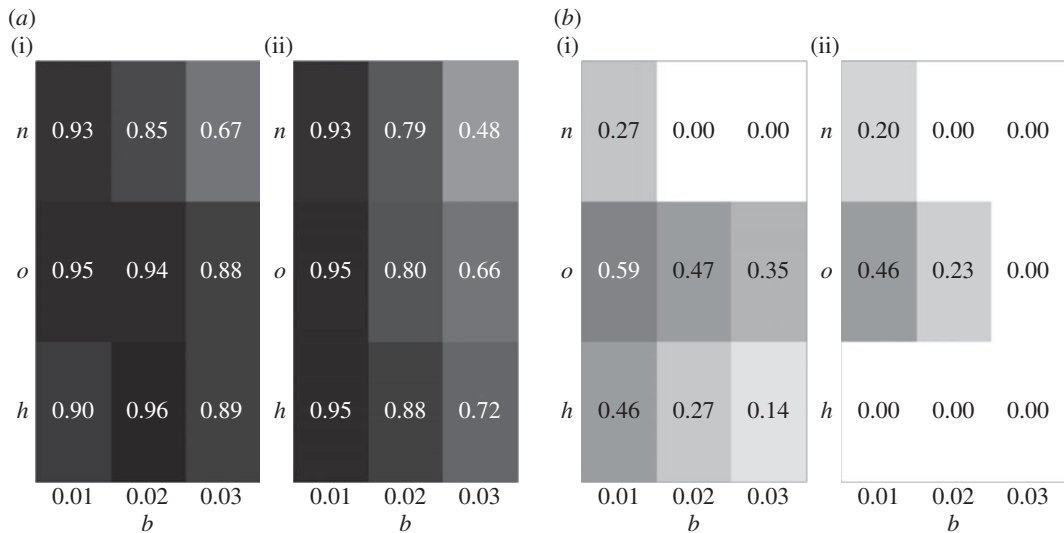


Figure 5. Distinguishability between pure transmission modes h, o, n (table 1) and their conformist counterparts based on (a) the level of cultural diversity, d_s^x , and (b) the average time a variant stays the most common variant, t_{max}^x . Shown are the values of the area of overlap O_{xy} between the probability distributions of the statistic for a given mode with and without conformity, with varying levels of $b = 0.01; 0.02; 0.03$, under parameter constellations $N = 50; \mu = 0.1; p_w = 1$ (a(i), b(i)) and $N = 100; \mu = 0.1; p_w = 1$ (a(ii), b(ii)). The corresponding grey scale ranges from white for $O_{xy} = 0$ to black for $O_{xy} = 1$.

cannot be distinguished from the ‘corresponding’ pure mode (i.e. horizontal transmission for low values of p_{mix} , oblique transmission for high values of p_{mix}).

3.2.3. Conformity bias

Conformist transmission is known to reduce cultural diversity: the cultural composition of a population tends to be more homogeneous with conformity bias than without it (e.g. [8]). But how strong does the bias have to be to produce characteristic population-level outcomes? In this section, we explore whether the presence of conformity bias can be established based on the level of cultural diversity in a population or the temporal dynamic of cultural change.

Without conformity bias, interactions occur at random between individuals in the relevant age group(s). By contrast, with conformity bias the probability of an individual interacting with a partner carrying a given variant increases with the frequency of the variant in the population (see equation (2.1)).

Following the procedure described in §2.2.2, we compare the probability distributions of statistics d_s^x and t_{max}^x for transmission modes $x, y \in \{v, h, o, n\}$ and parameter constellations $N = 50, 100; \mu = 0.1; p_w = 1$. Analysis of a larger set of parameters can be found in the electronic supplementary material, section S3.3.3. The aim in this case is to compare the distributions of each statistic for a given mode without conformity bias to its conformist counterparts with varying levels of b (i.e. $b = 0$ versus $b = 0.01, 0.02, 0.03$; recall that parameter b captures the strength of the bias). Note that we focus on scenarios with interaction probability $p_w = 1$.

Figure 5a shows the areas of overlap O_{xy} for statistic d_s^x , the level of cultural diversity in the population. By definition, vertical transmission is not affected by conformity bias (not shown). For all other modes, as expected there is a trend towards more homogeneous cultural compositions with increasing conformity bias. Thus, as b increases, diversity decreases, the distributions of d_s^x with and without bias become increasingly different, and the values of O_{xy} decrease as a result. However, the differences are typically not large enough to ensure distinguishability based on an arbitrary threshold $\bar{O} = 0.2$. The only exception is age-neutral transmission for a limited subset of parameter values involving a large population size, low to intermediate mutation rates and moderate to strong conformity bias (see the electronic supplementary material, figure S6). As expected, analysis of $P(y|z)$ given in equation (2.2) reveals that low values of d_s are consistent with a hypothesis of conformist transmission.

Figure 5b shows the behaviour of O_{xy} for statistic t_{max}^x , the average time a variant stays the most common variant in the population. We expect this to increase with conformity bias, which by definition sustains the most common variant in the population. Our results show that the ability to reliably detect

the bias increases with population size. At the same time, it is contingent on the number of potential interaction partners. For example, for age-neutral transmission, moderate conformity can be detected for almost all parameter settings (see also the electronic supplementary material, figure S6). By contrast, restricting the set of potential partners to peers (horizontal transmission) or older individuals (oblique transmission) reduces the range of parameter values for which the bias can be detected. In the case of oblique transmission, in fact, only strong bias can be reliably detected, and only for intermediate to high mutation rates (specifically, $N = 100$; $\mu = 0.1$; $b = 0.03$). Analysis of $P(y|z)$ reveals that large values of t_{\max} point to the existence of a conformist bias in the population.

3.2.4. Continuous traits

In the electronic supplementary material we apply the framework described above to continuous cultural traits, i.e. traits that can take any value in a given interval. Overall, we find that the population average of the trait does not carry a detectable signature of the underlying transmission modes. Binning the interval of trait values into discrete variants results in greater inferential power, but the binning must be sufficiently fine-grained, with the traits discretized in more than three variants (see the electronic supplementary material, section S3.3.4 for a detailed analysis).

4. Discussion

We developed a simulation model to explore the interface between individual-level processes and population-level patterns in human cultural evolution. Our first aim was to derive broad expectations about different pathways for the transmission of information between individuals. Our second aim was to establish theoretical limits to inference based on the population-level patterns produced by these pathways. Are the patterns different enough that they can be reliably distinguished based on population-level statistics, providing insight into the underlying individual-level processes?

The model tracks the distribution of variants of cultural traits across individuals in a population over time, conditioned on different modes for the transmission of information between individuals, for a range of parameters capturing demographic and cultural factors. Building on previous work [7], we investigated four ‘pure’ modes: vertical (parent to offspring), horizontal (peer to peer, by age group), oblique (older to younger, excluding parent to offspring), and age-neutral (any individual in the population). A fifth ‘mixed’ mode effectively combined horizontal and oblique transmission. We also investigated the effect of conformity bias, a preference for the most common variant in the population [8].

We used four statistics to summarize the cultural composition of the population at a specific point in time: the frequency distribution of the different variants of the cultural trait, the frequency of the most common variant, the total number of variants present, and the level of cultural diversity. An additional statistic measured the rate of change of the most common variant of a trait. This gave an indication of how fast the cultural composition of the population can change, providing insight into the temporal dynamic of the process.

4.1. Trickle-down effect and relative rates of change

The five transmission modes differ in the set of potential interaction partners (§2.1.1). Previous work suggests that they should therefore produce substantially different evolutionary dynamics, modulated by demographic and cultural factors (e.g. [7]). For example, Fogarty *et al.* [18] explored the effects of vertical, oblique and horizontal transmission on the age-structure of a population, and the extent to which cultural traits affecting demographic change can spread. They showed that a trait that reduces fertility but increases survival can spread to fixation and lead to substantial demographic change, in the form of an increase in population size, under certain modes of cultural transmission. In particular, if different modes of transmission matter for different age groups, then low fertility values can spread in a population and change the population structure via their effects on reproduction.

Our results, based on analysis of neutral traits and constant population sizes, confirm that the five transmission modes lead to differences in evolutionary dynamics. These differences are most pronounced under scenarios with high potential for cultural change (e.g. high mutation rates, high interaction probabilities). In particular, we have shown that oblique transmission is strongly affected by the value of the interaction probability: when this is high, oblique transmission results in more homogeneous cultural compositions compared with the other pure modes. Broadly, a homogeneous

cultural composition is reflected in an uneven frequency distribution of the cultural variants (i.e. a low number of variants present due to the high frequency of the most common variant), resulting in low cultural diversity (§3.1).

Furthermore, we have shown that this effect is rooted in the interplay between the different sets of potential interaction partners and the age structure of the population. In particular, oblique transmission is characterized by a disproportional influence of older individuals, whereby a variant at high frequency in one age group trickles down to all younger age groups; over time, this leads to a more homogeneous population. For example, an individual in the oldest age group is a potential interaction partner for all younger individuals in the population, whereas a ‘middle-aged’ individual is a potential partner for only the younger half of the population. This effect leaves a signature at the population level in terms of (i) lower levels of cultural diversity and (ii) slower rates of cultural change, for oblique transmission compared to horizontal and age-neutral transmission. Horizontal and age-neutral transmission show comparable rates of change.

Vertical transmission leads to the slowest rates of change. In our framework, this mode involves a ‘one-shot’ interaction event between parent and offspring, with opportunities for change limited to the process of mutation. The relative ‘conservativeness’ of vertical transmission is a well-established result in the theory of gene–culture coevolution, starting with seminal work by Cavalli-Sforza & Feldman [7]. We return to this issue below, after outlining our results regarding inference.

4.2. Inferring process from pattern

4.2.1. Summary of results

Questions relating to the indentifiability of transmission modes or other underlying processes are of course not restricted to the field of cultural evolution. For example, a large body of work in population genetics focuses on testing for the presence or the absence of selective forces (e.g. [29–31]). The general idea, similar to the one used here, is to derive expectations for different quantities of interest under a specific evolutionary scenario, and to subsequently compare empirical estimates of the quantities against those expectations. Recently, inferential frameworks have been applied directly to observed data. These frameworks often combine generative modelling of the system under consideration and Bayesian inference techniques. For example, in this way researchers have gained important insights into the transmission dynamics of infectious diseases and other epidemiological processes (e.g. [32–34]), and into human demographic history (e.g. [35–37]). Rather than focusing on a particular cultural dataset (see [11,12,38–40] for examples of such analyses) our aim here was to develop a theoretical understanding of which transmission modes can be distinguished on the basis of population-level data. This relates to the problem of equifinality, i.e. which modes can produce similar population-level patterns.

Our analysis provides two kinds of results: first, general expectations about the distinguishability of population-level outcomes produced by the different transmission modes in the absence of empirical data; second, given an empirical estimate of a statistic, the probability that a given transmission mode acted in the population to produce this estimate.

Beginning with the pure modes, our results show that under scenarios with low potential for cultural change (e.g. low mutation rates, low interaction probabilities) the modes produce outcomes that cannot be reliably distinguished based on any statistic. Yet outcomes tend to diverge as the potential for cultural change increases. In particular, under scenarios with high potential for cultural change they can be reliably distinguished, in most cases, based on the temporal dynamic. The two exceptions are vertical versus oblique transmission, and age-neutral versus horizontal transmission. In these two cases, additional insight can be gained from empirical estimates of the average time a variant stays the most common variant. For vertical versus oblique transmission, small values are indicative of oblique transmission, whereas large values clearly point to oblique transmission. For age-neutral versus horizontal transmission, even an empirical estimate will likely not be able to resolve the distinguishability issue.

These results suggest that even when outcomes are similar in terms of cultural compositions, they can differ substantially in temporal dynamics—in other words, similar distributions of cultural variants at a specific point in time can be reached through substantially different processes. In sum, for the pure modes the temporal dynamic of cultural change over time retains a stronger signature of the underlying processes than a ‘snapshot’ of the relative frequencies of the variants at a given point in time, and the signature is stronger the greater the potential for cultural change.

This general insight also applies to mixed transmission. However, in this case even under scenarios with high potential for cultural change the signature retained by the temporal dynamic tends to be too weak for outcomes to be reliably distinguished.

Finally, we have shown that the signature of conformity bias is, in general, stronger for the temporal dynamic than for the level of cultural diversity. As expected, the level of cultural diversity decreases with increasing conformity bias, while the temporal dynamic slows down. At the same time, with small population sizes the signature of conformist transmission tends to be obscured by random drift. It becomes stronger as population size increases, especially for age-neutral transmission, as the resulting increase in the pool of potential interaction partners is larger in this case than for horizontal or oblique transmission. For horizontal and oblique transmission, only strong conformity bias can be reliably detected, in most cases, and only based on the temporal dynamic with large population sizes.

Empirical estimates can provide further insight also in this case. Specifically, low values of cultural diversity and large average times a variant stays the most common point strongly to the existence of a conformist bias, when compared with the non-conformist situation.

4.2.2. Implications for empirical studies

We conclude by reviewing the implications of our findings for empirical studies of cultural evolution. Broadly, our work complements alternative theoretical approaches used to infer process from pattern in human culture, including efforts based on adoption curves (e.g. [16,41–43]), rank–abundance distributions (e.g. [19,44,45]), levels of diversity (e.g. [1,46,47]) and turnover rates (e.g. [48,49]), as well as model selection frameworks (e.g. [50]).

In terms of inference, our study suggests that if the frequency distribution of the different variants of a trait is available, then inference procedures should rely on these data where possible. Unsurprisingly, using partial data (e.g. the frequency of the most common variant of a trait) or summary statistics (i.e. the level of cultural diversity) results in loss of useful information. For example, a diversity measure such as the Shannon index is often used to summarize the cultural composition of a population (e.g. [1,26, 27]). We have shown that the dimension reduction involved in its calculation obscures population-level differences in the patterns produced by the transmission modes.

As discussed above, our results suggest that the temporal dynamic of cultural change over time is more instructive about the underlying processes than the cultural composition of the population at a given point in time. The approach we used to characterize the temporal dynamic is more parsimonious than other approaches in the literature. For example, Bentley *et al.* [48] use the turnover rate, defined as the number of new variants of a trait that enter the list of variants at the highest frequency in each time step (analogous to the ‘new entries’ in a top-10 or top-100 chart). However, this measure does not readily extend to traits that include only a small number of variants (e.g. five variants, as in our case). Additionally, it requires information about the frequencies of all variants in the population. By contrast, our approach only requires information about the frequency of the most common variant.

At the same time, it should be noted that the results presented here rest on the assumption that we have complete information about the most common variant of the trait at every point in a given time interval—in other words, that estimation of the temporal dynamic of cultural change is exact. Such detailed time-series data are difficult to obtain, however, especially for existing datasets and/or those that rely on historical information. In a related study using the mathematical framework developed here, we investigated how sparse the time-series data can be for the transmission modes to still be distinguishable [51]. Results show that if only incomplete information is available, i.e. the most common variant is known for a sample of time points, then the level of distinguishability depends on the properties of the sample. In particular, the distance between the time points affects how much insight can be obtained from population-level data. In general, if information about the most common variant in the population is only available for a sample of time points, one should not expect levels of distinguishability comparable to those we report here.

In practice, the use of the most appropriate statistic or inference framework is dictated by the available data. Given our focus on cultural evolution, we have analysed statistics that can potentially be derived even from sparse data—a common feature of datasets in anthropology and related disciplines. Where researchers have access to data with high temporal resolution, such as time series describing the frequency change of different cultural variants, then adoption curve analyses, and especially model selection frameworks, can be instructive.

A final set of insights relates to the analysis of continuous cultural traits. Our results show that the mean value of a trait across the population does not carry a detectable signature of the underlying

transmission modes; therefore, it is not a useful statistic for characterizing the cultural composition of a population. Binning the interval of trait values into discrete variants results in greater inferential power, but the binning must be sufficiently fine-grained. For example, discretization into two or three variants (corresponding to, for example, present versus absent or small versus medium versus large) is generally not enough to ensure distinguishability based on the temporal dynamic of cultural change, whereas discretization into five or ten variants produces distinguishability results comparable to the discrete case.

We conclude with a general observation, bearing on the interface between the observed population-level patterns and the underlying individual-level processes. As discussed above, our results show that vertical transmission leads to the slowest rate of change of all the modes, consistent with the notion prevalent in the literature of its relative ‘conservativeness’ [7]. This notion continues to provide the foundation to a large body of empirical work, including field-based investigations (e.g. [10–12]) and cross-cultural studies (e.g. [12–15]). At the same time, our inference results show that vertical transmission can produce temporal dynamics similar to other modes, and in particular to oblique transmission. Empirical estimates of statistics describing the temporal dynamic can provide further crucial information. Still, our findings invite caution in linking population-level patterns to individual-level processes based on data documenting variation in cultural traits within and between populations.

This example illustrates well the theoretical limits to inferring individual-level processes from population-level patterns in human cultural evolution. In particular, we should not expect a one-to-one mapping between population-level statistics and the underlying transmission modes: different scenarios can lead to comparable patterns at the level of groups. Consistency between any one specific scenario and empirical data should be interpreted in this context. However, acknowledging the problem of equifinality does not imply that we cannot extract any information about cultural evolution from these data. Mathematical frameworks similar to the one used here can provide general expectations, in the form of probability distributions, against which to compare empirical estimates. Furthermore, statistical inference procedures that compare simulated data to empirical data can help delimit the amount of information that can be extracted on a case-by-case basis (e.g. [38–40]). Conceptually, this shifts the focus from identifying the one scenario that likely produced the observed data to excluding those that likely did not.

Data accessibility. The code is deposited at <http://datadryad.org/resource/doi:10.5061/dryad.14q31> [52].

Authors' contributions. L.F. and A.K. designed the study and wrote the paper; B.W. and A.K. developed the code and analysed the model; L.F., B.W. and A.K. read and approved the final version of the paper.

Competing interests. We declare we have no competing interests.

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References

- Shennan SJ, Wilkinson JR. 2001 Ceramic style change and neutral evolution: a case study from Neolithic Europe. *Am. Antiq.* **66**, 577–593. (doi:10.2307/2694174)
- Hruschka DJ, Branford S, Smith ED, Wilkins J, Meade A, Pagel M, Bhattacharya T. 2015 Detecting regular sound changes in linguistics as events of concerted evolution. *Curr. Biol.* **25**, 1–9. (doi:10.1016/j.cub.2014.10.064)
- Fortunato L. 2015 Evolution of marriage systems. In *International encyclopedia of the social & behavioral sciences*, vol. 14 (ed. JD Wright), 2nd edn, pp. 611–619. Oxford, UK: Elsevier. (doi:10.1016/B978-0-08-097086-8.81059-4)
- Laland KN, Brown GR. 2011 *Sense and nonsense: evolutionary perspectives on human behaviour*, 2nd edn. Oxford, UK: Oxford University Press.
- Mesoudi A. 2015 Cultural evolution: a review of theory, findings and controversies. *Evol. Biol.* **43**, 481–497. (doi:10.1007/s11692-015-9320-0)
- Richerson PJ, Boyd R. 2005 *Not by genes alone: how culture transformed human evolution*. Chicago, IL: The University of Chicago Press.
- Cavalli-Sforza LL, Feldman MW. 1981 *Cultural transmission and evolution: a quantitative approach*. Princeton, NJ: Princeton University Press.
- Boyd R, Richerson PJ. 1985 *Culture and the evolutionary process*. Chicago, IL: The University of Chicago Press.
- Laland KN, Galef BG. 2009 *The question of animal culture*. Cambridge, MA: Harvard University Press.
- Hewlett BS, Cavalli-Sforza LL. 1986 Cultural transmission among Aka pygmies. *Am. Anthropol.* **88**, 922–934. (doi:10.1525/aa.1986.88.4.02a00100)
- Hewlett BS, Fouts HN, Boyette AH, Hewlett BL. 2011 Social learning among Congo Basin hunter–gatherers. *Phil. Trans. R. Soc. B* **366**, 1168–1178. (doi:10.1098/rstb.2010.0373)
- Tehrani JJ, Collard M. 2009 On the relationship between interindividual cultural transmission and population-level cultural diversity: a case study of weaving in Iranian tribal populations. *Evol. Hum. Behav.* **30**, 286–300. (doi:10.1016/j.evolhumbehav.2009.03.002)
- Guglielmino CR, Viganotti C, Hewlett B, Cavalli-Sforza LL. 1995 Cultural variation in Africa: role of mechanisms of transmission and adaptation. *Proc. Natl Acad. Sci. USA* **92**, 7585–7589. (doi:10.1073/pnas.92.16.7585)
- Hewlett BS, De Silvestri A, Guglielmino CR. 2002 Semes and genes in Africa. *Curr. Anthropol.* **43**, 313–321. (doi:10.1086/339379)
- Moylan JW, Borgerhoff Mulder M, Graham CM, Nunn CL, Håkansson NT. 2006 Cultural traits and linguistic trees: phylogenetic signal in East Africa. In

- Mapping our ancestors: phylogenetic approaches in anthropology and prehistory* (eds CP Lipo, MJ O'Brien, M Collard, SJ Shennan), ch. 3, pp. 33–52. New Brunswick, NJ: Aldine Transaction.
16. Henrich J. 2001 Cultural transmission and the diffusion of innovations: adoption dynamics indicate that biased cultural transmission is the predominate force in behavioral change. *Am. Anthropol.* **103**, 992–1013. (doi:10.1525/aa.2001.103.4.992)
 17. Powell A, Shennan S, Thomas MG. 2009 Late Pleistocene demography and the appearance of modern human behavior. *Science* **324**, 1298–1301. (doi:10.1126/science.1170165)
 18. Fogarty L, Creanza N, Feldman M. 2013 The role of cultural transmission in human demographic change: an age-structured model. *Theor. Popul. Biol.* **88**, 68–77. (doi:10.1016/j.tpb.2013.06.006)
 19. Bentley RA, Hahn MW, Shennan SJ. 2004 Random drift and culture change. *Proc. R. Soc. Lond. B* **271**, 1443–1450. (doi:10.1098/rspb.2004.2746)
 20. Carotenuto L, Feldman MW, Cavalli-Sforza LL. 1989 Age structure in models of cultural transmission. *Morrison Institute Working Paper*, no. 2.
 21. Fogarty L, Feldman MW. 2011 The cultural and demographic evolution of son preference and marriage type in contemporary China. *Biol. Theory* **6**, 272–282. (doi:10.1007/s13752-012-0033-7)
 22. Li N, Feldman MW, Li S. 2000 Cultural transmission in a demographic study of sex ratio at birth in China's future. *Theor. Popul. Biol.* **58**, 161–172. (doi:10.1006/tpbi.2000.1478)
 23. Laland KN. 2004 Social learning strategies. *Anim. Learn. Behav.* **32**, 4–14. (doi:10.3758/BF03196002)
 24. Eriksson K, Enquist M, Ghirlanda S. 2007 Critical points in current theory of conformist social learning. *J. Evol. Psychol.* **5**, 67–87. (doi:10.1556/JEP.2007.1009)
 25. Kandler A, Shennan SJ. 2013 A non-equilibrium neutral model for analysing cultural change. *J. Theor. Biol.* **330**, 18–25. (doi:10.1016/j.jtbi.2013.03.006)
 26. Leonard RD, Jones GT. 1989 *Quantifying diversity in archaeology*. Cambridge, UK: Cambridge University Press.
 27. Cruz-Urbe K. 1988 The use and meaning of species diversity and richness in archaeological faunas. *J. Archaeol. Sci.* **15**, 179–196. (doi:10.1016/0305-4403(88)90006-4)
 28. Kura K, Broom M, Kandler A. 2015 Modelling dominance hierarchies under winner and loser effects. *Bull. Math. Biol.* **77**, 927–952. (doi:10.1007/s11538-015-0070-z)
 29. Ewens WJ. 1972 The sampling theory of selectively neutral alleles. *Theor. Popul. Biol.* **3**, 87–112. (doi:10.1016/0040-5809(72)90035-4)
 30. Lande R. 1976 Natural selection and random genetic drift in phenotypic evolution. *Evolution* **30**, 314–334. (doi:10.1111/j.1558-5646.1976.tb00911.x)
 31. Watterson G. 1978 The homozygosity test of neutrality. *Genetics* **88**, 405–417.
 32. Lintusaari J, Gutmann MU, Kaski S, Corander J. 2016 On the identifiability of transmission dynamic models for infectious diseases. *Genetics* **202**, 911–918. (doi:10.1534/genetics.115.180034)
 33. Tanaka MM, Francis AR, Luciani F, Sisson S. 2006 Using approximate Bayesian computation to estimate tuberculosis transmission parameters from genotype data. *Genetics* **173**, 1511–1520. (doi:10.1534/genetics.106.055574)
 34. Stadler T. 2011 Inferring epidemiological parameters on the basis of allele frequencies. *Genetics* **188**, 663–672. (doi:10.1534/genetics.111.126466)
 35. Pritchard JK, Stephens M, Donnelly P. 2000 Inference of population structure using multilocus genotype data. *Genetics* **155**, 945–959.
 36. Posth C *et al.* 2016 Pleistocene mitochondrial genomes suggest a single major dispersal of non-Africans and a late glacial population turnover in Europe. *Curr. Biol.* **26**, 827–833. (doi:10.1016/j.cub.2016.01.037)
 37. Hofmanová Z *et al.* 2016 Early farmers from across Europe directly descended from Neolithic Aegeans. *Proc. Natl Acad. Sci. USA* **113**, 6886–6891. (doi:10.1073/pnas.1523951113)
 38. Crema E, Edinborough K, Kerig T, Shennan S. 2014 An approximate Bayesian computation approach for inferring patterns of cultural evolutionary change. *J. Archaeol. Sci.* **50**, 160–170. (doi:10.1016/j.jas.2014.07.014)
 39. Kandler A, Shennan SJ. 2015 A generative inference framework for analysing patterns of cultural change in sparse population data with evidence for fashion trends in LBK culture. *J. R. Soc. Interface* **12**, 20150905. (doi:10.1098/rsif.2015.0905)
 40. Kandler A, Powell A. 2015 Inferring learning strategies from cultural frequency data. In *Learning strategies and cultural evolution during the Palaeolithic* (eds A Mesoudi, K Aoki), pp. 85–101. Tokyo, Japan: Springer. (doi:10.1007/978-4-431-55363-2_7)
 41. Reader SM. 2004 Distinguishing social and asocial learning using diffusion dynamics. *Anim. Learn. Behav.* **32**, 90–104. (doi:10.3758/BF03196010)
 42. Franz M, Nunn CL. 2009 Network-based diffusion analysis: a new method for detecting social learning. *Proc. R. Soc. B* **276**, 1829–1836. (doi:10.1098/rspb.2008.1824)
 43. Hoppitt W, Kandler A, Kendal JR, Laland KN. 2010 The effect of task structure on diffusion dynamics: implications for diffusion curve and network-based analyses. *Learn. Behav.* **38**, 243–251. (doi:10.3758/LB.38.3.243)
 44. Hahn MW, Bentley RA. 2003 Drift as a mechanism for cultural change: an example from baby names. *Proc. R. Soc. Lond. B* **270** (Suppl. 1), S120–S123. (doi:10.1098/rsbl.2003.0045)
 45. Mesoudi A, Lycett SJ. 2009 Random copying, frequency-dependent copying and culture change. *Evol. Hum. Behav.* **30**, 41–48. (doi:10.1016/j.evolhumbehav.2008.07.005)
 46. Neiman FD. 1995 Stylistic variation in evolutionary perspective: inferences from decorative diversity and interassemblage distance in Illinois Woodland ceramic assemblages. *Am. Antiq.* **60**, 7–36. (doi:10.2307/282074)
 47. Premo LS. 2014 Cultural transmission and diversity in time-averaged assemblages. *Curr. Anthropol.* **55**, 105–114. (doi:10.1086/674873)
 48. Bentley RA, Lipo CP, Herzog HA, Hahn MW. 2007 Regular rates of popular culture change reflect random copying. *Evol. Hum. Behav.* **28**, 151–158. (doi:10.1016/j.evolhumbehav.2006.10.002)
 49. Acerbi A, Bentley RA. 2014 Biases in cultural transmission shape the turnover of popular traits. *Evol. Hum. Behav.* **35**, 228–236. (doi:10.1016/j.evolhumbehav.2014.02.003)
 50. McElreath R, Bell AV, Efferson C, Lubell M, Richerson PJ, Waring T. 2008 Beyond existence and aiming outside the laboratory: estimating frequency-dependent and pay-off-biased social learning strategies. *Phil. Trans. R. Soc. B* **363**, 3515–3528. (doi:10.1098/rstb.2008.0131)
 51. Wilder B, Kandler A. 2016 Inference of cultural transmission modes based on incomplete information. *Hum. Biol.* **87**, 193–204. (doi:10.13110/humanbiology.87.3.0193)
 52. Kandler A, Wilder B, Fortunato L. 2017 Data from: Inferring individual-level processes from population-level patterns in cultural evolution. Dryad Digital Repository. (doi:10.5061/dryad.14q31)