II Università di Roma, Tor Vergata Dipartimento d'Ingegneria Civile e Ingegneria Informatica LM in Ingegneria dell'Informazione e dell'Automazione Complementi di Probabilità e Statistica - Advanced Statistics Instructors: Roberto Monte & Massimo Regoli Solved Problems on Hypothesis Tests 2023-01-17

Problem 1 Let X be a Gaussian random variable with unknown mean μ_X and variance σ_X^2 representing a population. Assume that testing the sample mean \bar{X}_n and the sample standard deviation S_n of a simple random sample X_1, \ldots, X_n of size $n \equiv 9$ drawn from X we obtain the value $\bar{X}_n(\omega) \equiv \bar{x}_n = 251.50$ cm and $S_n(\omega) \equiv s_n = 2.30$ cm.

- 1. Considering both the rejection region method and the p-value method, should the null hypothesis $H_0: \mu_X = 250 cm$ be rejected against the alternative $H_a: \mu_X \neq 250 cm$ at the significance level $\alpha = 0.1$?
- 2. Considering both the rejection region method and the p-value method, should the null hypothesis $H_0: \sigma_X^2 = 4$ be rejected against of the alternative $H_a: \sigma_X^2 > 4$ at the significance level $\alpha = 0.05$? Calculate the probability β (5) of a II type error.

Solution.

1. Since X is Gaussian distributed with unknown mean and variance and the size of the sample is small, the statistic to be used is

$$\frac{\bar{X}_n - \mu_X}{S_n / \sqrt{n}}.\tag{1}$$

Consider testing the null hypothesis $H_0: \mu_X = \mu_0$, where $\mu_0 \equiv 250cm$, against the alternative $H_1: \mu_X \neq \mu_0$ at the significance level $\alpha = 0.1$. Under the assumption that the null hypothesis is true the statistic (1) with $\mu_X = \mu_0$ has the Student distribution with n-1 degree of freedom, that is

$$\frac{\bar{X}_n - \mu_0}{S_n / \sqrt{n}} \sim T_{n-1}.$$

Moreover, the stucture of the alternative hypothesis calls for a rejection region of the form

$$R = \left\{ T_{n-1} < t_{n-1, 1-\alpha/2} \right\} \cup \left\{ T_{n-1} > t_{n-1, \alpha/2} \right\}.$$

where,

$$t_{n-1,\alpha/2} = t_{8,0.05} = 1.860$$
 and $t_{n-1,1-\alpha/2} = -t_{n-1,\alpha/2} = -t_{8,0.05} = -1.860$.

Hence,

$$R = (-\infty, -1.860) \cup (1.860, +\infty)$$

Computing the realization of the statistic, we have

$$\frac{\bar{X}_n(\omega) - \mu_X}{S_n(\omega)/\sqrt{n}} = \frac{\bar{x}_n - \mu_0}{s_n/\sqrt{9}} = \frac{251.50 - 250}{2.30/3} = 1.96 \in R.$$

This implies a rejection of the null hypothesis in favor of the alternative. Adopting the p-value method, we recall that, on account of the alternative hypothesis, the p-value is the probability that the absolute value of the test statistic under the null assumption yields a value not less than

the realization of the statistic. We will reject the null hypothesis when the computed p-value is smaller that the given significance level α . In symbols,

$$p = \mathbf{P}\left(|T_{n-1}| \ge \frac{\bar{x}_n - \mu_0}{s_n/\sqrt{9}} \mid H_0 = T\right) = \mathbf{P}\left(T_8 \le -1.96\right) + \mathbf{P}\left(T_8 \ge 1.96\right) = 0.087 < 0.1,$$

which confirms the rejection of the null hypothesis.

2. Since we are interested in testing a hypothesis on the variance of X, which is normally distributed with unknown variance and the size of the sample is small, the statistic to be used is

$$\frac{(n-1)S_n^2}{\sigma_V^2}. (2)$$

Consider testing the null hypothesis $H_0: \sigma_X^2 = \sigma_0^2$, where $\sigma_0^2 \equiv 4$, against the alternative $H_1: \sigma_X^2 > \sigma_0^2$ at the significance level $\alpha \equiv 0.05$. Under the assumption that the null hypothesis is true the statistic (2) with $\sigma_X^2 = \sigma_0^2$ has the chi-square distribution with n-1 degrees of freedom, that is

$$\frac{(n-1)S_n^2}{\sigma_0^2} \sim \chi_{n-1}^2$$
.

Hence, the upper tail rejection region is given by

$$R = \left\{ \chi_{n-1}^2 > \chi_{n-1,\alpha}^2 \right\}$$

where, the upper $\alpha = 0.05$ critical value $\chi^2_{n-1,\alpha}$ of the chi-square distribution with n-1=8 degrees of freedom is given by

$$\chi^2_{8.0.05} \simeq 15.51.$$

Hence,

$$R = (15.51, +\infty)$$

Computing the realization of the statistic, we have

$$\frac{\left(n-1\right)S_{n}^{2}\left(\omega\right)}{\sigma_{X}^{2}}=\frac{\left(n-1\right)s_{n}^{2}}{\sigma_{0}^{2}}=\frac{8\cdot2.30^{2}}{4}=10.58\notin R$$

Hence, the realization of the statistic does not belong to the rejection region. This implies that H_0 cannot be rejected.

In terms of p-value we have to compute

$$\mathbf{P}\left(\chi_{n-1}^{2} \geq \frac{(n-1)\,s_{n}^{2}}{\sigma_{0}^{2}} \mid H_{0} \text{ true}\right) = \mathbf{P}\left(\chi_{8}^{2} \geq 10.58\right) = 1 - \mathbf{P}\left(\chi_{8}^{2} \leq 10.58\right) = 0.227 > 0.05.$$

The p-value method confirms that H_0 cannot be rejected.

With regard to the evaluation of β (5), setting $\sigma_1^2 \equiv 5$, we have

$$\beta(5) = \mathbf{P} \left(\operatorname{accept} H_0 \mid \sigma_X^2 = \sigma_1^2 \right)$$

$$= \mathbf{P} \left(\frac{(n-1)S_n^2}{\sigma_0^2} \notin R \mid \sigma_X = \sigma_1^2 \right)$$

$$= \mathbf{P} \left(\frac{(n-1)S_n^2}{\sigma_0^2} \frac{\sigma_1^2}{\sigma_1^2} \notin R \mid \sigma_X = \sigma_1^2 \right)$$

$$= \mathbf{P} \left(\frac{(n-1)S_n^2}{\sigma_1^2} \frac{\sigma_1^2}{\sigma_0^2} \notin R \mid \sigma_X = \sigma_1^2 \right)$$

$$= \mathbf{P} \left(\frac{(n-1)S_n^2}{\sigma_1^2} \frac{\sigma_1^2}{\sigma_0^2} \notin R \right)$$

$$= \mathbf{P} \left(\chi_{n-1}^2 \frac{\sigma_1^2}{\sigma_0^2} \notin R \right)$$

$$= \mathbf{P} \left(\chi_{8}^2 \frac{\sigma_1^2}{\sigma_0^2} \le 15.51 \right)$$

$$= \mathbf{P} \left(\chi_{8}^2 \le 15.51 \cdot \frac{\sigma_0^2}{\sigma_1^2} \right)$$

$$= \mathbf{P} \left(\chi_{8}^2 \le 8.46 \right)$$

$$= 0.61.$$

Solution.

Problem 2 Let X be a normal random variable with unknown mean μ and variance σ^2 , which represents a certain characteristic of a population and let X_1, \ldots, X_n be a simple random sample of size n drawn from X. Assume that n = 25 and the realizations x_1, \ldots, x_{25} of the sample give the information summarized by

$$\sum_{k=1}^{25} x_k = 100$$
 and $\sum_{k=1}^{25} x_k^2 = 550$

- Considering both the rejection region method and the P-value method, should the null hypothesis
 H₀: σ² = 4 be rejected against of the alternative H₁: σ² > 4 with a significance level α = 0.05?
 Calculate the probability β(5) of a II type error. What should have been the size n of the sample to achieve β(5) = 0.5?
- 2. Considering both the rejection region method and the P-value method, should the null hypothesis $H_0: \sigma^2 = 4$ be rejected against of the alternative $H_1: \sigma^2 \neq 4$ with a significance level $\alpha = 0.05$? Calculate the probability $\beta(5)$ of a II type error. What should have been the size n of the sample to achieve $\beta(5) = 0.5$?

Solution. Note that the given information allows the knowledge of the realization of sample variance,

which is given by

$$\begin{split} s_n^2 &= \frac{1}{n-1} \sum_{k=1}^n \left(x_k - \frac{1}{n} \sum_{\ell=1}^n x_\ell \right)^2 \\ &= \frac{1}{n-1} \sum_{k=1}^n \left(x_k^2 - \frac{2}{n} x_k \sum_{\ell=1}^n x_\ell + \frac{1}{n^2} \left(\sum_{\ell=1}^n x_\ell \right)^2 \right) \\ &= \frac{1}{n-1} \left(\sum_{k=1}^n x_k^2 - \frac{2}{n} \left(\sum_{k=1}^n x_k \right) \left(\sum_{\ell=1}^n x_\ell \right) + \frac{1}{n^2} \sum_{k=1}^n \left(\sum_{\ell=1}^n x_\ell \right)^2 \right) \\ &= \frac{1}{n-1} \left(\sum_{k=1}^n x_k^2 - \frac{2}{n} \left(\sum_{k=1}^n x_k \right)^2 + \frac{1}{n^2} n \left(\sum_{\ell=1}^n x_\ell \right)^2 \right) \\ &= \frac{1}{n-1} \left(\sum_{k=1}^n x_k^2 - \frac{1}{n} \left(\sum_{k=1}^n x_k \right)^2 \right). \end{split}$$

Thus, in the present case, we have

$$s_n^2 = \frac{1}{24} \left(550 - \frac{1}{25} 100^2 \right) = \frac{25}{4} = 6.25.$$

Now, since we are interested in testing a hypothesis on the variance of X, which is normally distributed, the standard statistic to be considered is

$$\frac{(n-1)S_n^2}{\sigma^2} \sim \chi_{n-1}^2.$$

1. By virtue of the above considerations, the upper tail rejection region is given by

$$R = \left\{ \frac{(n-1)S_n^2}{\sigma^2} > \chi_{n-1,\alpha}^2 \right\}$$

where, the upper $\alpha = 0.05$ critical value of the chi-square distribution with n-1=24 degrees of freedom is given by

$$\chi^2_{24,0.05} = 36.4150.$$

Now, we have

$$\frac{(n-1)s_n^2}{\sigma^2} = \frac{24}{4} \cdot \frac{25}{4} = 37.50 \in R \tag{3}$$

Hence, the realization of the statistic belongs to the rejection region. This implies that the null hypotesis H_0 has to be rejected in favor of H_1 .

In terms of P-value we have to compute

$$\mathbf{P}\left(\frac{(n-1)S_n^2}{\sigma^2} \ge \frac{(n-1)s_n^2}{\sigma^2} \mid H_0 = T\right) = \mathbf{P}\left(\frac{(n-1)S_n^2}{\sigma^2} \ge \frac{(n-1)s_n^2}{\sigma^2} \mid \sigma = \sigma_0\right)$$

$$= \mathbf{P}\left(\chi_{24}^2 \ge 37.50\right)$$

$$= 1 - \mathbf{P}\left(\chi_{24}^2 \le 37.50\right)$$

$$= 0.039 < 0.050.$$

As expected, the P-value method confirms the rejection of H_0 in favor of H_1 .

With regard to the evaluation of β (5), $\sigma_0^2 \equiv 4$, $\sigma_1^2 \equiv 5$, and n=25, we have that

$$\beta(5) = \mathbf{P} (\text{II type error}) = \mathbf{P} (\text{accept } H_0 \mid H_0 = F) = \mathbf{P} (\text{accept } H_0 \mid \sigma = \sigma_1^2)$$

$$= \mathbf{P} \left(\frac{(n-1) S_n^2}{\sigma_0^2} \notin R \mid \sigma = \sigma_1^2 \right)$$

$$= \mathbf{P} \left(\frac{(n-1) S_n^2}{\sigma_0^2} \frac{\sigma_1^2}{\sigma_1^2} < 36.4150 \mid \sigma = \sigma_1^2 \right)$$

$$= \mathbf{P} \left(\frac{(n-1) S_n^2}{\sigma_1^2} < 36.4150 \frac{\sigma_0^2}{\sigma_1^2} \mid \sigma = \sigma_1^2 \right)$$

$$= \mathbf{P} \left(\frac{(n-1) S_n^2}{\sigma^2} < 36.4150 \frac{4}{5} \right)$$

$$= \mathbf{P} \left(\chi_{n-1}^2 \le 29.132 \right)$$

$$= 0.7848.$$

In the end, to achieve $\beta(5) = 0.5$ we have to solve the equation

$$0.5 = \mathbf{P} \left(\chi_{n-1}^2 \le 29.132 \right)$$

in terms of the smallest n. A computer aided computation yields

$$\mathbf{P}\left(\chi_{n-1}^2 \le 29.132\right) = \begin{cases} 0.54 & \text{if } n = 29\\ 0.49 & \text{if } n = 30 \end{cases},$$

which implies that n = 30 is the desired n.

2. In this case the rejection region is given by

$$R = \left\{\frac{\left(n-1\right)S_n^2}{\sigma^2} < \chi_{n-1,1-\alpha/2}^2\right\} \cup \left\{\frac{\left(n-1\right)S_n^2}{\sigma^2} > \chi_{n-1,\alpha/2}^2\right\}$$

where, the lower $1 - \alpha/2 = 0.975$ and the upper $\alpha/2 = 0.025$ critical values of the chi-square distribution with n - 1 = 24 degrees of freedom are given by

$$\chi^2_{24,0.975} = 12.40$$
 and $\chi^2_{24,0.025} = 39.36$

respectively. In this case the values 37.50 of the statistics (see 3) does not belong to the rejection region. Therefore, H_0 cannot be rejected against H_1 . In terms of P-value we have to compute

$$2\mathbf{P}\left(\frac{(n-1)S_n^2}{\sigma^2} \ge \frac{(n-1)s_n^2}{\sigma^2} \mid H_0 = T\right) = 2\mathbf{P}\left(\frac{(n-1)S_n^2}{\sigma^2} \ge \frac{(n-1)s_n^2}{\sigma^2} \mid \sigma = \sigma_0^2\right)$$

$$= 2\mathbf{P}\left(\chi_{24}^2 \ge 37.50\right)$$

$$= 2\left(1 - \mathbf{P}\left(\chi_{24}^2 < 37.50\right)\right)$$

$$= 0.078 > 0.050.$$

Hence, the P-value method confirms that H_0 cannnot be rejected against H_1 .

With regard to the evaluation of β (5), setting $\sigma_0^2 \equiv 4$, $\sigma_1^2 \equiv 5$, and n = 25, we have that

$$\beta(5) = \mathbf{P} \text{ (II type error)} = \mathbf{P} \text{ (accept } H_0 \mid H_0 = F) = \mathbf{P} \text{ (accept } H_0 \mid \sigma = \sigma_1^2)$$

$$= \mathbf{P} \left(\frac{(n-1) S_n^2}{\sigma_0^2} \notin R \mid \sigma = \sigma_1^2 \right)$$

$$= \mathbf{P} \left(12.40 \le \frac{(n-1) S_n^2}{\sigma_0^2} \frac{\sigma_1^2}{\sigma_1^2} \le 39.36 \mid \sigma = \sigma_1^2 \right)$$

$$= \mathbf{P} \left(12.40 \frac{\sigma_0^2}{\sigma_1^2} \le \frac{(n-1) S_n^2}{\sigma_1^2} \le 39.36 \frac{\sigma_0^2}{\sigma_1^2} \mid \sigma = \sigma_1^2 \right)$$

$$= \mathbf{P} \left(12.40 \frac{4}{5} \le \frac{(n-1) S_n^2}{\sigma^2} \le 39.36 \frac{4}{5} \right)$$

$$= \mathbf{P} \left(9.92 \le \chi_{n-1}^2 \le 31.49 \right)$$

$$= \mathbf{P} \left(\chi_{n-1}^2 \le 31.49 \right) - \mathbf{P} \left(\chi_{n-1}^2 \le 9.92 \right)$$

$$= 0.860 - 0.005 = 0.855.$$