lecture 3 presentation

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- Statistik opsumering
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- 3 Auto regressive Processes opsumering
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Section 1

Statistik opsumering

Expected value

$$E(aX + b) = aE(X) + b$$
$$E(X + Y) = E(X)E(Y)$$

$$E(XY) = E(X)E(Y)$$

- Hvis X og Y er uafhængige
- Bevis for dette behøves i ik at kende

Variance

$$Var(X) = E[X - E(X)]^2 = E(XX) - E(X)E(X)$$

- Vi kan se hvis E(X) = 0 er $Var(X) = E(X^2)$
- Bevis ovenstående

$$Var(aX + b) = a^{2}Var(X)$$
 $Var(X + Y) = Var(X) + Var(Y) + 2Cov(X, Y)$

Bevis ovenstående

$$Var(X + Y) = Var(X) + Var(Y)$$

• Hvis X og Y er uafhængige og dermed Cov(X, Y) = 0

Bevis for Var(X)

$$Var(X) = E[X - E(X)]^2 = E(XX) - E(X)E(X)$$

• Løs på tavlen

Bevis for Var(X+Y)

$$Var(X + Y) = Var(X) + Var(Y) + 2Cov(X, Y)$$

Covariance

$$Cov(X, Y) = E[X - E(X)][Y - E(Y)] = E(XY) - E(X)E(Y)$$

Bevis ovenstående

$$Cov(X, X) = Var(X)$$

$$Cov(aX + b, cY + d) = acCov(X, Y)$$

$$Cov(X + Y, Z) = Cov(X, Z) + Cov(Y, Z)$$

Bevis Cov(X,Y)

Forsøg at bevis nedenstående:

$$Cov(X, Y) = E[X - E(X)][Y - E(Y)] = E(XY) - E(X)E(Y)$$

Løs på tavlen

Section 2

Moving Average process opsumering:

Mean, varians og Covariance

$$Y_t = \mu + \varepsilon_t + \alpha \varepsilon_{t-1}$$
$$\varepsilon_t \sim IID(0, \sigma^2)$$

• Calculate mean of Y_t

$$E[Y_t] = E[\mu]E[\varepsilon_t] + E[\alpha\varepsilon_{t-1}]$$

• Vi ved $E[\varepsilon_t] = E[\varepsilon_{t-1}] = 0$

$$E[Y_t] = \mu$$

Mean, varians og Covariance

$$Y_t = \mu + \varepsilon_t + \alpha \varepsilon_{t-1}$$
$$\varepsilon_t \sim IID(0, \sigma^2)$$

Calculate the variance

$$Var[Y_t] = E[(Y_t - \mu)^2]$$

$$= E[(\varepsilon_t + \alpha \varepsilon_{t-1})^2]$$

$$= E[\varepsilon_t^2 + 2\alpha \varepsilon_t \varepsilon_{t-1} + \alpha^2 \varepsilon_{t-1}^2]$$

$$= E[\varepsilon_t^2] + 2\alpha E[\varepsilon_t \varepsilon_{t-1}] + \alpha^2 E[\varepsilon_{t-1}^2]$$

- Hvorfor ved vi $2\alpha E[\varepsilon_t \varepsilon_{t-1}] = 0$?
- Hvad sker der med $E[\varepsilon_t^2]$ og $E[\varepsilon_{t-1}^2]$, og hvorfor?

$$= E[\varepsilon_t^2] + \alpha^2 E[\varepsilon_{t-1}^2]$$
$$(1 + \alpha^2)\sigma^2$$

Mean, varians og Covariance

$$Y_t = \mu + \varepsilon_t + \alpha \varepsilon_{t-1}$$
$$\varepsilon_t \sim IID(0, \sigma^2)$$

• Calculate autocovariance between Y_t and Y_{t-1}

$$Cov[Y_t, Y_{t-1}] = E[(Y_t - \mu) * (Y_{t-1} - \mu)]$$
$$E[(\varepsilon_t + \alpha \varepsilon_{t-1})(\varepsilon_{t-1} + \alpha \varepsilon_{t-2})]$$

• Forsøg selv at tage de sidste steps!

$$\alpha E[\varepsilon_{t-1}]$$

 $\alpha \sigma^2$

Mean, Variance og Covariance

• Calculate autocovariance between Y_t and Y_{t-2}

$$Cov[Y_t, Y_{t-2}] = E[(Y_t - \mu) * (Y_{t-2} - \mu)]$$
$$= E[(\varepsilon_t + \alpha \varepsilon_{t-1})(\varepsilon_{t-2} + \alpha \varepsilon_{t-3})]$$

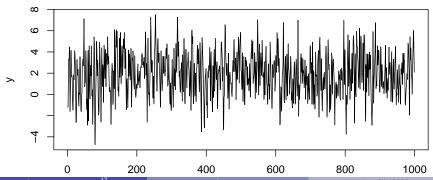
• Samme metode som vist på tavlen før:

$$= 0$$

MA(1) Simulation

$$Y_t = 2 + \varepsilon_t + 0.9\varepsilon_{t-1}$$
$$\varepsilon_t \sim IID(0, 1.5^2)$$

MA(1) process with α =0.9



MA(1) Simulation

```
mean(y)# 1.919655
var(y) # 3.991872
cov(y[-length(y)],y[-1])# 1.922769
```

$$E(y_t) = \mu = 2$$
 $V(y_t) = (1 + \alpha^2)\sigma^2$
 $V(y_t) = (1 + 0.9^2)1.5^2$
 $V(y_t) = (1 + 0.9^2)1.5^2 = 4.0725$
 $Cov = \alpha * \sigma^2$
 $Cov = 0.9 * 1.5^2 = 2.025$

Section 3

Auto regressive Processes opsumering

Auto regressive Processes opsumering

$$Y_t = \mu + \theta Y_{t-1} + \varepsilon_t$$
$$\varepsilon_t \sim IID(0, \sigma^2)$$

- Værktøjer vi skal bruge til properties af AR-modeller:
- Geometriske serier.
- Difference ligninger.

Geometriske serier review - Eksempler på serier

Eksempel 1:

$$\sum_{n=1}^{n} x^{n} = x + x^{2} + x^{3} + x^{4} + \dots + x^{n-1}$$

Eksempel 2:

$$x + \sqrt{x} + 1 + \frac{1}{\sqrt{x}} + \frac{1}{x}$$
...

Typer af geometriske serier

Note til senere:

- \bullet $a=\mu$
- $k = \theta$

Endelig serie

$$\sum_{n=1}^{n} ak^n = a * \frac{1-k^n}{1-k}, \ k \neq 1$$

Uendelig serie

$$\sum_{n=1}^{n}ak^{n}=rac{a}{1-k}$$
, $|k|<1$
 $\sum_{n=1}^{n}ak^{n}=n$ a, $|k|=1$

Udregning

Endelig serie

$$S_{n} = \alpha + \alpha * k + \alpha * k^{2} + \alpha * k^{3} + \dots + \alpha * k^{n-1}$$

$$k * S_{n} = \alpha * k + \alpha * k^{2} + \alpha * k^{3} + \alpha * k^{4} + \dots + \alpha * k^{n}$$

$$S_{n} - k * S_{n} = \alpha + (\alpha * k - \alpha * k) + (\alpha * k^{2} - \alpha * k^{2}) + \dots + (\alpha * k^{n-1} - \alpha * k^{n-1}) - \alpha * k^{n}$$

$$S_{n} - k * S_{n} = \alpha - \alpha * k^{n}$$

$$S_{n}(1 - k) = \alpha(1 - k^{n})$$

$$S_{n} = \alpha * \frac{1 - k^{n}}{1 - k}$$

Udregning

Uendelig serie

Hvis |k| < 1 når $n \to \infty$ vil udtrykket gå mod:

$$S_n = \alpha * \frac{1}{1-k}$$

Dermed kan vi skrive:

$$\sum_{n=1}^{\infty} \alpha k^n = \frac{\alpha}{1-k}$$

Difference equations: Math to econometrics

Looking at a 1. order difference equations

In 2. semester math seen like this:

$$x_t = ax_{t-1} + b$$

In time series econometrics you have seen it like this (AR(1) process):

$$y_t = \mu + \theta y_{t-1} + \varepsilon_t$$

What are the differences?

Lets first see what they got in common:

- ullet Both equations got a constant μ and b
- ullet Both equations got a coefficient heta and a
- Both equations got a variable that changes over time (discrete) y_t and x_t

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Difference equations: Math to econometrics

The difference is the error term ε_t with the differition:

$$\varepsilon_t \sim IID(0, \sigma^2)$$

Identical, Independent Distributed med mean = 0 og $Var = \sigma^2$

Later we take a look at how this changes things!

Lets solve the difference equation we saw before:

$$x_t = ax_{t-1} + b_t$$

We can start from a given point x_0

$$x_1 = ax_0 + b_1$$

$$x_2 = ax_1 + b_2 = a(ax_0 + b_1) + b_2 = a^2x_0 + ab_1 + b_2$$

$$x_3 = ax_2 + b_3 = a(a^2x_0 + ab_1 + b_2) + b_3 = a^3x_0 + a^2b_1 + ab_2 + b_3$$

We can already see the pattern:

$$x_t = a^t x_0 + \sum_{k=1}^t a^{t-k} b_k$$

We now assume the case when $b_k = b$ so we now got a constant as in the AR(1) model (fixed over time).

So we can now write the last term as:

$$\sum_{k=1}^{t} a^{t-k} b$$

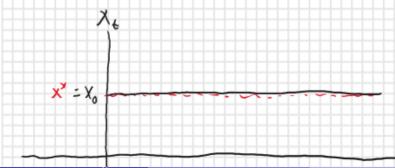
Which is a geometric series! that we just covered!

$$\sum_{k=1}^{t} a^{t-k} b = b(a^{t-1} + a^{t-2} + a^{t-3} + \dots + a + 1) = \frac{(b - ba^t)}{(1 - a)}$$

Therefor we can now write:

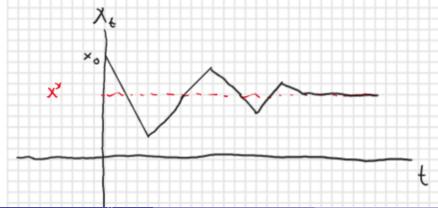
$$x_t = a^t(x_0 - \frac{b}{1-a}) + \frac{b}{1-a}$$

We can see that if $x_0 = \frac{b}{1-a}$ we get that $x_t = \frac{b}{1-a}$ which I have illustrated down below:



In fact if just x_s at any point hits $\frac{b}{1-a}$ we wont get away from it as:

$$x_{s+1} = a \frac{b}{1-a} + b = \frac{b}{1-a}$$



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Difference equations (stability)

Case 1

We then see that a^t goes towards 0 as $t \to \infty$ in:

$$x_t = a^t(x_0 - \frac{b}{1-a}) + \frac{b}{1-a}$$

And we will end up with

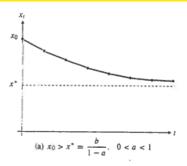
$$x_t = \frac{b}{1-a}$$

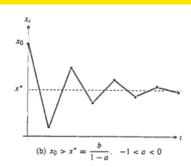
Difference equations (stability)

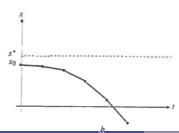
Case 2

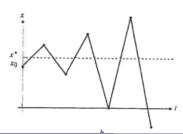
- We then see that a^t goes towards ∞ as $t \to \infty$ and will explode.
- Lets look at the different scenarios

Difference equations (stability)









Lets look at the AR(1) process again:

$$y_t = \mu + \theta y_{t-1} + \varepsilon_t$$

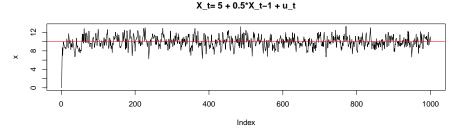
Where the only difference was the error term: ε_t lets see some examples and what the difference is

To give an example we use the model:

$$y_t = 5 + 0.5y_{t-1} + \varepsilon_t$$

With start value $y_0 = 0$

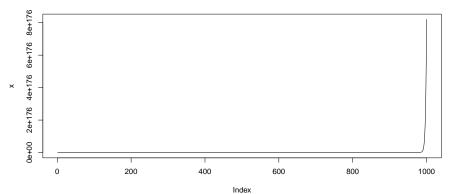
Lad is bruge løsningen for en difference ligning for når $|{\it a}| < 1$



We see the shocks from ε_t does so we never stay in $\frac{b}{1-a}$, så istedet regner vi mean!

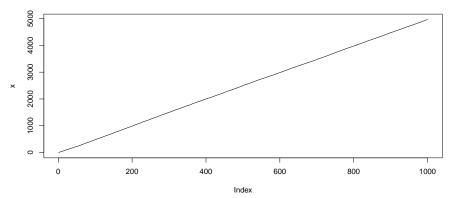
$$y_t = 5 + 1.5y_{t-1} + \varepsilon_t$$

$$X_t = 5 + 1.5 \times X_t - 1 + u_t$$



$$y_t = 5 + 1y_{t-1} + \varepsilon_t$$

$$X_t = 5 + 1*X_t - 1 + u_t$$



Plottet vi så før ligner bare en lineær funktion! Hvilket det er!

Husk løsning til vores difference eq.

$$x_t = a^t x_0 + \sum_{k=1}^t a^{t-k} b_k$$

Her brugte vi løsningen til den geometriske serie til at substituere ind for $\sum_{k=1}^t a^{t-k}b_k$ Men vi antog |a|<1

Gå tilbage til geometriske serier!

Vi kan derfor nu indsætte $\sum_{k=1}^{t} a^{t-k} b_k = ta$

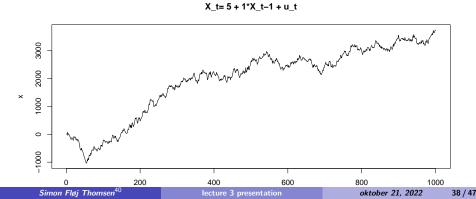
$$x_t = a^t x_0 + ta$$

Vi ved a=1 dermed starter vi i x_0 og vokser linært som $t\to\infty$

Tilbage til AR(1) processen! Hvad er det vi kalder det når |a|=1 aka | heta|=1

En Random walk med drift!

Lad os lave det lidt mere tydeligt ved at øge standard deviation i fejlledet!



AR(1) model Exonometrics (mean)

We saw before that we never stay the value $\frac{b}{1-a}$ in the AR(1) model, but what if we calculate the mean?

$$E[y_t] = E[\mu + \theta y_{t-1} + \varepsilon_t]$$

$$= \mu + \theta E[y_{t-1}] + E[\varepsilon]$$

$$= \mu + \theta E[\mu + \theta y_{t-2} + \varepsilon_{t-1}]$$

$$= \mu + \mu \theta + \theta^2 E[y_{t-2}]$$

$$= \mu + \mu \theta + \theta^2 E[\mu + \theta y_{t-3} + \varepsilon_{t-2}]$$

$$= \mu(1 + \theta + \theta^2 + \theta^3 + \dots + \theta^\infty)$$

So back to the geometric series if $|\theta|<1$ we get $\frac{\mu}{1-\theta}$

AR(1) model Exonometrics (mean)

Lets try calculating the mean using the example from before

$$y_t = 5 + 0.5y_{t-1} + \varepsilon_t$$

 $E[y_t] = \frac{5}{1 - 0.5} = 10$

Lets look at the plot again!

AR(1) model Exonometrics (Variance)

Da μ er en konstant vil denne ikke påvirke variancen og vi kan fjerne denne fra start.

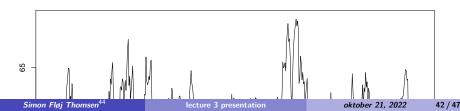
```
set.seed(213)
n=500
e <- rnorm(n, mean = 0, sd = 1.5)
Y <- 2/(1-0.9)
for (t in 2:n) {
Y[t] <- 2+ 0.9*Y[t-1]+e[t]
}
ts.plot(Y, main=expression(paste("AR(1) process with ", theta</pre>
```

AR(1) process with θ =0.1

AR(1) model Exonometrics (Variance)

```
set.seed(213)
n=500
e <- rnorm(n, mean = 0, sd = 1.5)
Y <- 6/(1-0.9)
for (t in 2:n) {
Y[t] <- 6+ 0.9*Y[t-1]+e[t]
}
ts.plot(Y, main=expression(paste("AR(1) process with ", theta</pre>
```

AR(1) process with θ =0.1



AR(1) model Exonometrics (Variance)

$$y_t = \mu + \theta y_{t-1} + \varepsilon_t$$
$$\varepsilon_t \sim IID(0, \sigma^2)$$

Vi antager derfor $\mu = 0$

$$V(y_t) = E[(y_t - E[y_t])^2]$$

$$V(y_t) = E[(\theta y_{t-1} + \varepsilon_t)^2]$$

$$V(y_t) = E[\varepsilon_t^2] + \theta^2 E[y_{t-1}^2]$$

$$V(y_t) = \sigma^2 + \theta^2 E[y_{t-1}^2]$$

$$V(y_t) = \sigma^2 + \theta^2 E[(\theta y_{t-2} + \varepsilon_{t-1})^2]$$

• Vi kan nu indsætte y_{t-2} og gøre nøjagtigt de samme steps:

AR(1) model Exonometrics (Auto covariance)

$$cov(Y_t, Y_{t-1}) = cov(\mu + \theta Y_{t-1} + \varepsilon_t, Y_{t-1})$$

• fra statistik ved vi at Cov(X + Y, Z) = Cov(X, Z) + Cov(Y, Z)

$$cov(\mu, Y_{t-1}) + \theta cov(Y_{t-1}, Y_{t-1}) + cov(\varepsilon_t, Y_{t-1})$$

- vi ved at $cov(\mu, Y_{t-1}) = cov(\varepsilon_t, Y_{t-1}) = 0$
- Og hvad er det nu $cov(Y_{t-1}, Y_{t-1})$ er?

$$cov(Y_t, Y_{t-1}) = \theta \frac{\sigma^2}{1 - \theta^2}$$

- Og antagelsen fra variance skal nu bruges: | heta| < 1

AR(1) model Exonometrics (Auto covariance)

$$cov(Y_t, Y_{t-2}) = cov(\mu + \theta Y_{t-1} + \varepsilon_t, Y_{t-2})$$

$$cov(\mu, Y_{t-2}) + \theta cov(Y_{t-1}, Y_{t-2}) + cov(\varepsilon_t, Y_{t-1})$$

- vi ved at $cov(\mu, Y_{t-1}) = cov(\varepsilon_t, Y_{t-1}) = 0$
- Og vi kender $cov(Y_{t-1}, Y_{t-2})$ som vi fandt på sidste slide.

$$cov(Y_t, Y_{t-2}) = \theta^2 \frac{\sigma^2}{1 - \theta^2}$$

- ullet Da vi bruger covariancen med antagelsen, gælder den stadig: | heta| < 1
- Derfor ACF aftager over tid når i plotter en AR-model.

Section 4

Done

Cheatsheet

AR(1)		AR(1) (RW1)	AR(1) (RW2)	AR(1) (RW3)
Name $ heta < 1$	MA(1)	heta $= 1$	heta $= 1$	$ \theta $ = 1
Mean $\frac{\mu}{1-\theta}$	μ	<i>Y</i> ₀	$Y_0 + T\mu$	<i>Y</i> ₀ +
				$T\mu + t$
Var $\frac{\sigma^2}{1-\theta^2}$	$(1+\alpha^2)\sigma^2$	$T\sigma^2$	$T\sigma^2$	$T\sigma^2$
Cov $\theta \frac{\sigma^2}{1-\theta^2}$	$\alpha \sigma^2$	$T\sigma^2$	$T\sigma^2$	$T\sigma^2$
Stationary? YES!	YES!	NO!	NO!	NO!