

1 Model

The general dynamic linear model (DLM) is a linear, gaussian, state space model and can be written as

$$y_t = F_t \theta_t + v_t \quad v_t \stackrel{\text{ind}}{\sim} N_k(0, V_t) \quad (1)$$

$$\theta_t = G_t \theta_{t-1} + w_t \quad w_t \stackrel{\text{ind}}{\sim} N_p(0, W_t) \quad (2)$$

for $t = 1, 2, \dots, T$, and $v_{1:T}$, $w_{1:T}$ independent. Equation (1) is called the *observation equation* and equation (2) is called the *system equation*. Similarly, $v_{1:T}$ are called the observation errors, $V_{1:T}$ are called the observation variances, $w_{1:T}$ are called the system disturbances and $W_{1:T}$ are called the system variances. The observed data is $y_{1:T}$ while $\theta_{0:T}$ are called the latent states. For each $t = 1, 2, \dots, T$, F_t is a $k \times p$ matrix and G_t is a $p \times p$ matrix. Let ϕ denote the vector of unknown parameters in the model. Then possibly $F_{1:T}$, $G_{1:T}$, $V_{1:T}$, and $W_{1:T}$ are all functions of ϕ .

We will focus our attention on a simpler version of the DLM. Typically additional model structure is used to learn about $V_{1:T}$ and $W_{1:T}$ if time dependence is enforced – e.g. a stochastic volatility prior which would require a statespace model describing the $V_{1:T}$ ’s and $W_{1:T}$ ’s as data. Because of this additional complexity, we focus on the time-constant variances model, though many of our results may be useful in more complicated time-varying variance models. So we set $V_t = V$ and $W_t = W$ for $t = 1, 2, \dots, T$. We will also suppose that F_t and G_t are known matrices for $t = 1, 2, \dots, T$, though this constraint is immaterial since relaxing it will simply add one or more Gibbs steps to the algorithms we explore so long as no parameter that enters any F_t or G_t also enters V or W . Note, however, that in one of the data augmentations that we discuss, the scaled error data augmentation, there is a bit more housekeeping associated with $F_{1:T}$ depending on an unknown parameter (Section ??).

When $\phi = (V, W)$ is our unknown parameter vector and we can write the model as

$$y_t | \theta_{0:T} \stackrel{\text{ind}}{\sim} N(F_t \theta_t, V) \quad (3)$$

$$\theta_t | \theta_{0:t-1} \sim N(G_t \theta_{t-1}, W) \quad (4)$$

To complete the model specification in a Bayesian context, we need priors on θ_0 , V , and W . We’ll use the standard approach and assume that they’re mutually independent a priori and that $\theta_0 \sim N(m_0, C_0)$, $V \sim IW(\Lambda_V, \lambda_V)$ and $W \sim IW(\Lambda_W, \lambda_W)$ where m_0 , C_0 , Λ_V , λ_V , Λ_W , and λ_W are known hyperparameters and $IW(\Lambda, \lambda)$ denotes the inverse Wishart distribution with degrees of freedom λ and positive definite scale matrix Λ . This allows us to write the full joint distribution of $(V, W, \theta_{0:T}, y_{1:T})$ as

$$\begin{aligned} p(V, W, \theta_{0:T}, y_{1:T}) &\propto \exp \left[-\frac{1}{2} (\theta_0 - m_0)' C_0^{-1} (\theta_0 - m_0) \right] \\ &\times |V|^{-(\lambda_V + k + T + 2)/2} \exp \left[-\frac{1}{2} \text{tr}(\Lambda_V V^{-1}) \right] \exp \left[-\frac{1}{2} \sum_{t=1}^T (y_t - F_t \theta_t)' V^{-1} (y_t - F_t \theta_t) \right] \\ &\times |W|^{-(\lambda_W + p + T + 2)/2} \exp \left[-\frac{1}{2} \text{tr}(\Lambda_W W^{-1}) \right] \exp \left[-\frac{1}{2} \sum_{t=1}^T (\theta_t - G_t \theta_{t-1})' W^{-1} (\theta_t - G_t \theta_{t-1}) \right] \end{aligned} \quad (5)$$

where $p = \dim(\theta_t)$, $k = \dim(y_t)$, and $\text{tr}(\cdot)$ is the matrix trace operator.