

Bayesian Inference in Gravitational Wave Astronomy

Mukesh Kumar Singh
ICTS Astrophysical Relativity Group



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Outline

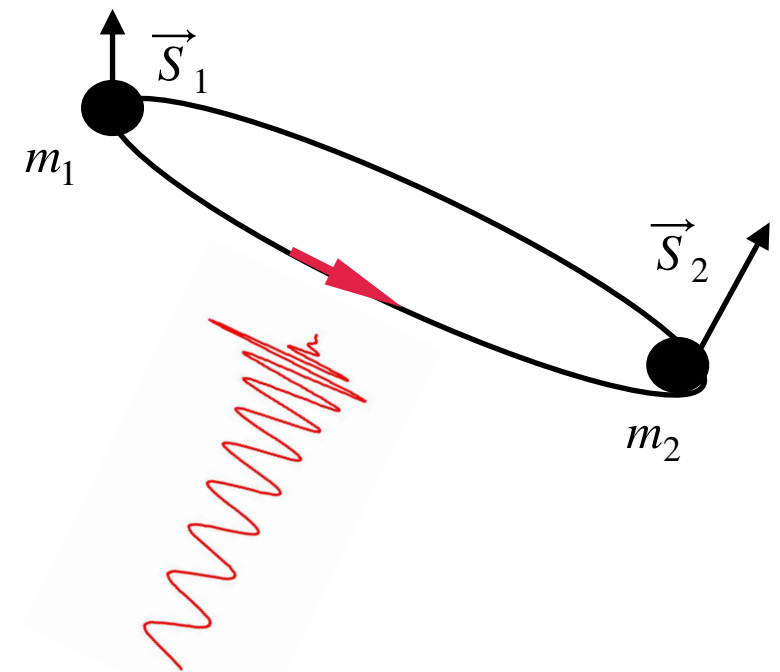
- Introduction
- Parameter estimation
- Conditional probability & Bayes' theorem
- Bayesian inference
- Stochastic sampling methods
- GW parameter estimation: Bilby
- Summary

Introduction

- Characteristic shape of the gravitational-wave signal encodes the information about the **astrophysical properties** of the source.
- In a **compact binary merger** with a quasicircular orbit will have:

Intrinsic parameters: $m_1, m_2, \vec{S}_1, \vec{S}_2$

Extrinsic parameters: $\begin{cases} d_L, \alpha, \delta & \text{location} \\ l, \psi & \text{orientation} \\ \varphi_0, t_c & \text{arrival time \& phase} \end{cases}$

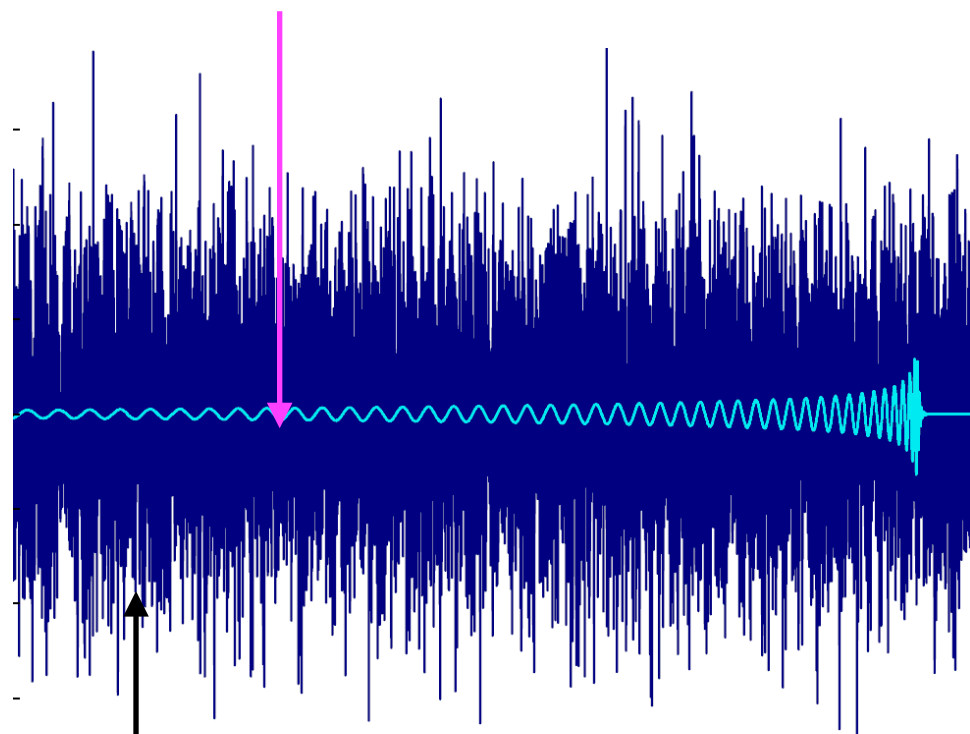


- The 15 dimensional parameter space.
- How to infer the these complex set of parameters?

Problem in hand

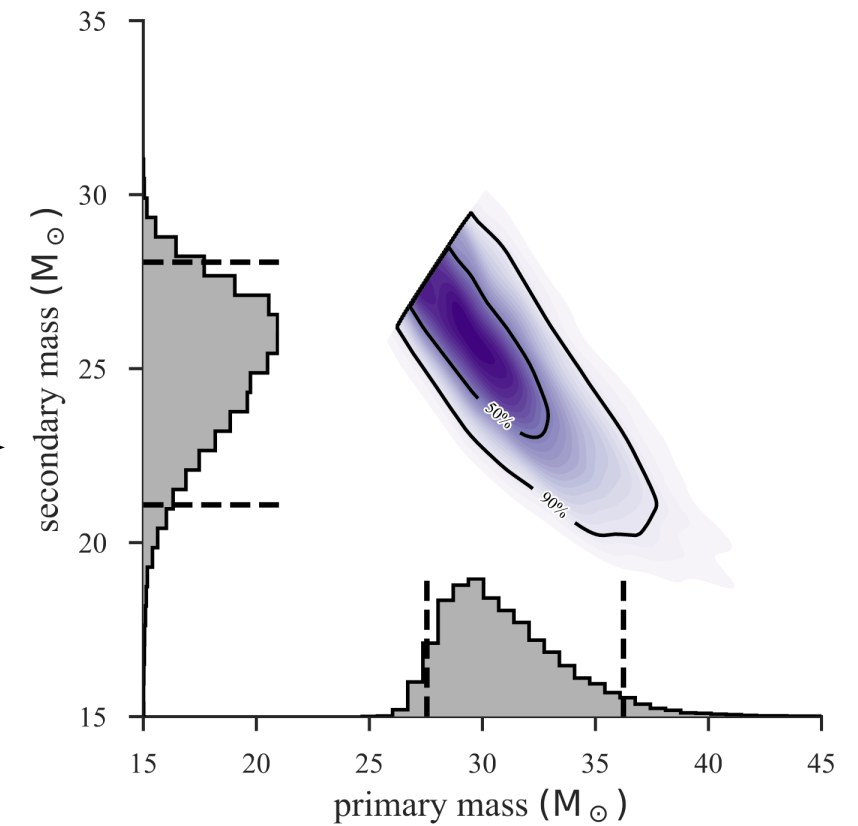
Credit: LSC

Gravitational-Wave Signal



Credit: H Gabbard et al

Noise

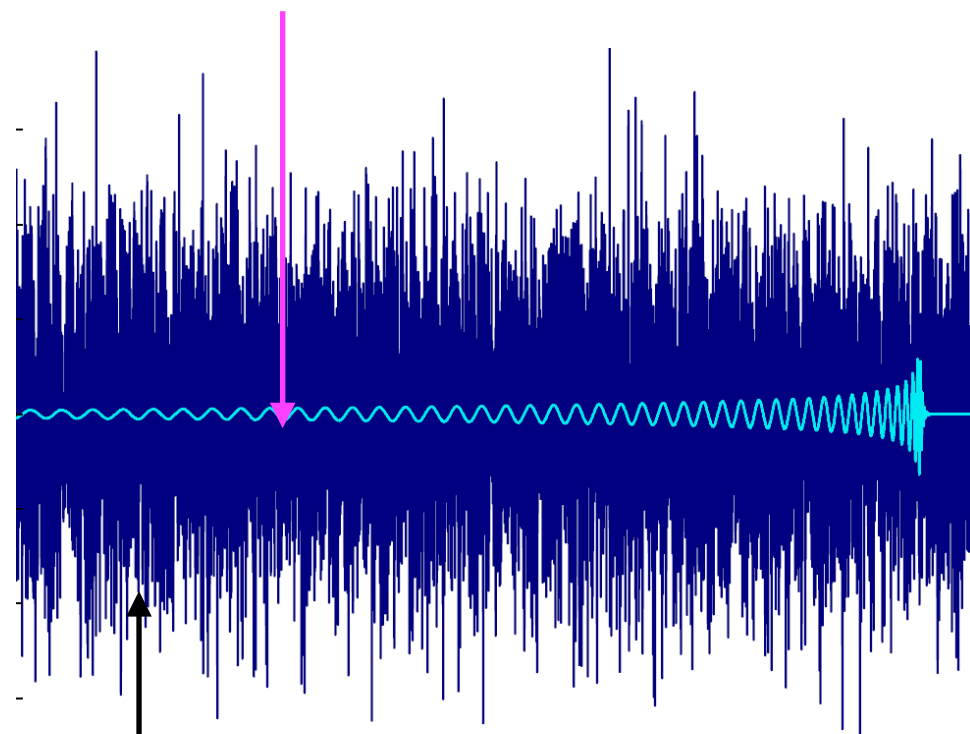


Probability distribution for component masses of the binary

Problem in hand

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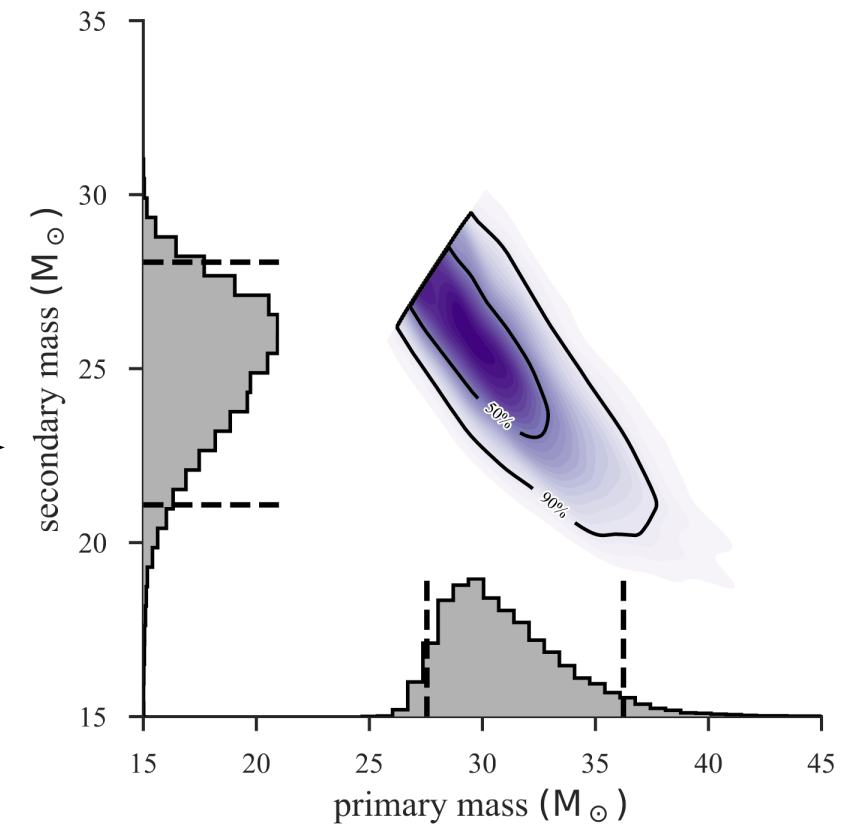
Gravitational-Wave Signal



Noise

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Parameter
Estimation

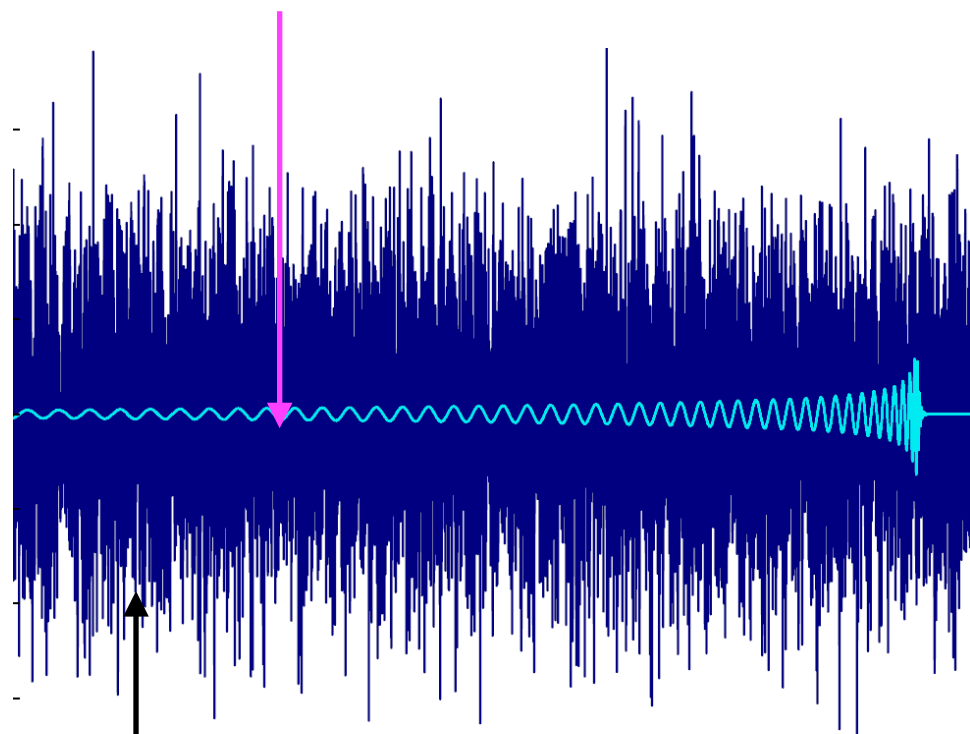


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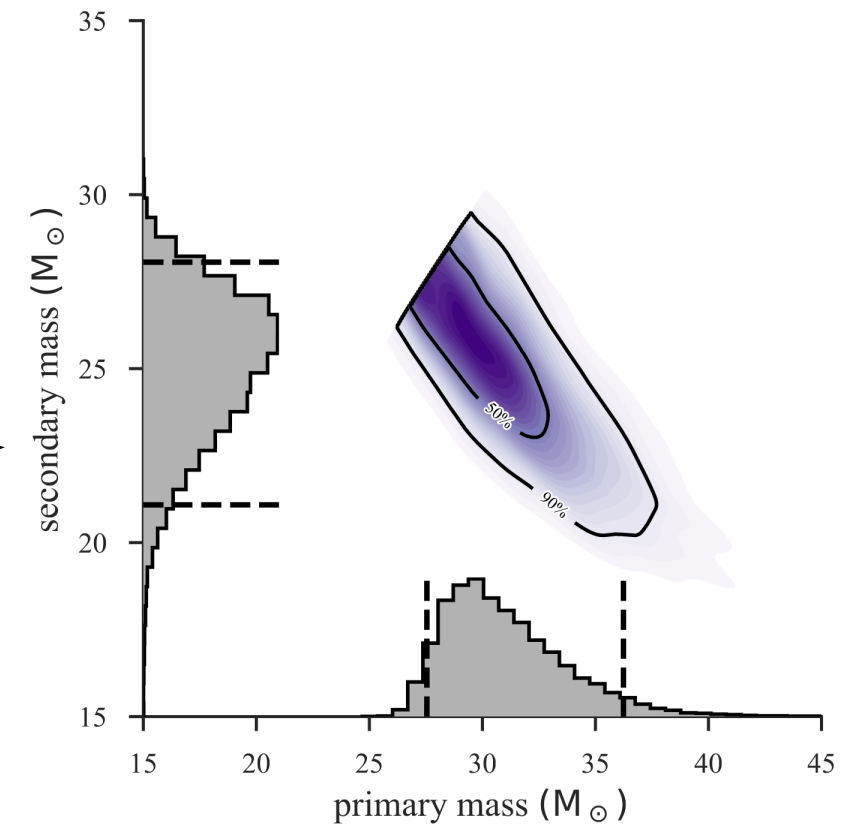


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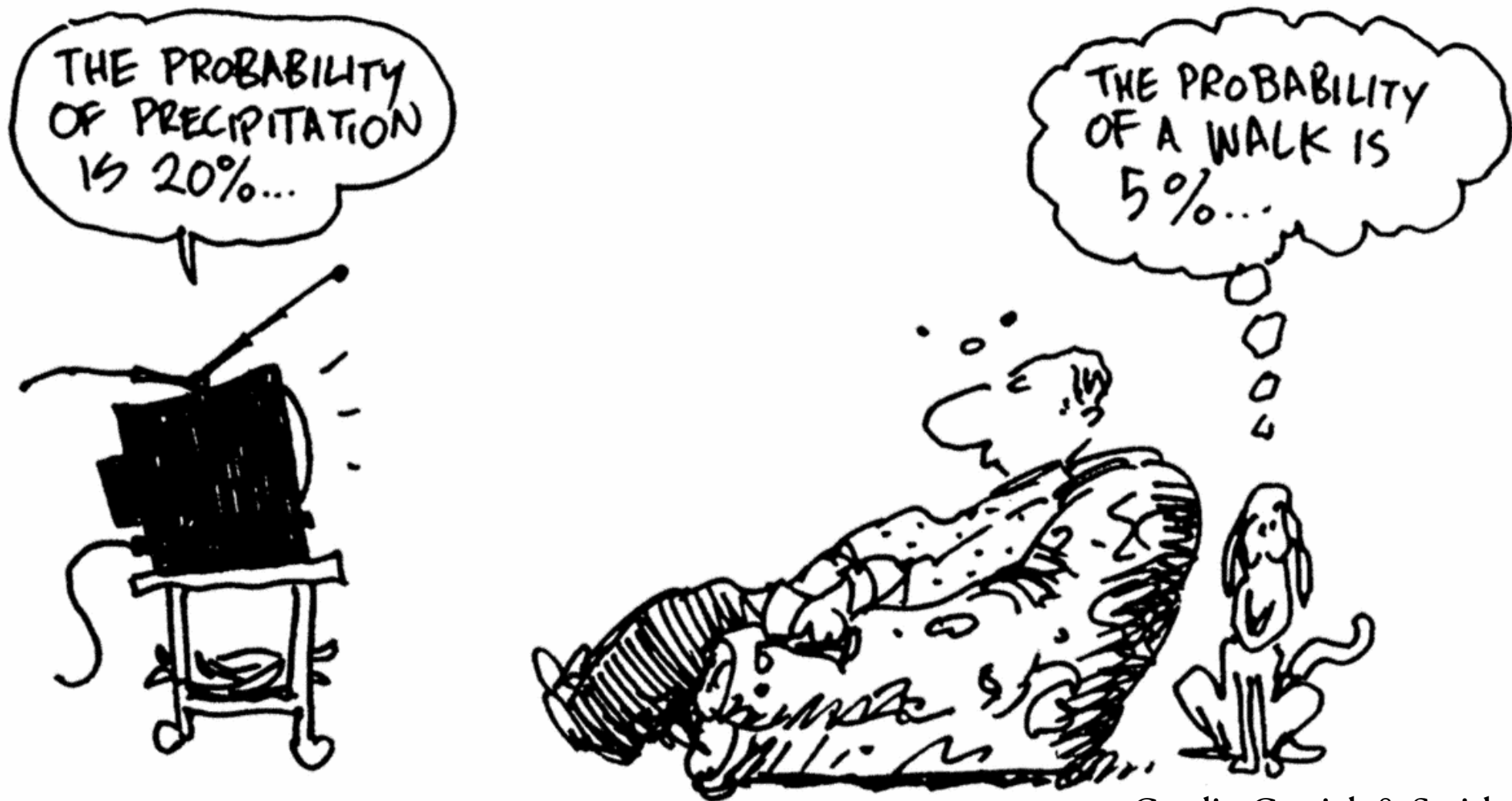


Probability distribution for component masses of the binary

- How can we infer the source parameters of an unknown signal buried in noise and with what accuracy?

Conditional Probability

- $p(A | B)$: conditional probability i.e. the probability of A **given** B.



Credit: Gonick & Smith

Bayes Theorem

- Conditional probability

$$p(A, B | I) = p(A | B, I)p(B | I)$$

- We can also write the same thing as

$$p(A, B | I) = p(B | A, I)p(A | I)$$

- Rearranging the terms in above two equations

$$p(A | B, I) = \frac{p(B | A, I)p(A | I)}{p(B | I)}$$

Bayesian Inference

- Inference means figuring something out from the data (d).
- Inference can be made two ways:
 1. **Parameter estimation:** finding out the parameters (θ) of a model (M_A) that best fits the data (d).
 2. **Model selection:** finding out which model, M_A or M_B , fits the data more effectively.

Parameter estimation

- Figuring out the model parameters θ , given the data d and the model M_A

$$\underbrace{p(\theta | d, M_A)}_{\text{Posterior}} = \frac{\overbrace{p(d | \theta, M_A)}^{\text{Likelihood}} \overbrace{p(\theta | M_A)}^{\text{Prior}}}{\underbrace{p(d | M_A)}_{\text{Evidence}}}$$

Used when comparing models

- Evidence is just a normalisation. So..

$$\underbrace{p(\theta | d, M_A)}_{\text{Posterior}} \propto \underbrace{p(d | \theta, M_A)}_{\text{Likelihood}} \underbrace{p(\theta | M_A)}_{\text{Prior}}$$

The degree of belief After the experiment

The degree of belief before the experiment

One Dimensional Example

- Given some specific observation y at time t , the data d

$$d(t) = y(t) + n \implies n = d(t) - y(t)$$

where n is random noise drawn from a gaussian distribution $\mathcal{N}(0, \sigma)$

- If the model is given by $M_A : y_A(t) = \sin(\omega t)$ with only parameter ω , the likelihood

$$\mathcal{L}(d | \omega, M_A) = p(d | \omega, y_A) \equiv p(n) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp \left[-\frac{(d - y_A(t, \omega))^2}{2\sigma^2} \right]$$

- A good idea to work with log-likelihood for the sake of stability

$$\ln \mathcal{L} = -\frac{1}{2} \left(\frac{(d(t) - y_A(t, \omega))^2}{\sigma^2} + \ln(2\pi\sigma^2) \right)$$

One Dimensional Example

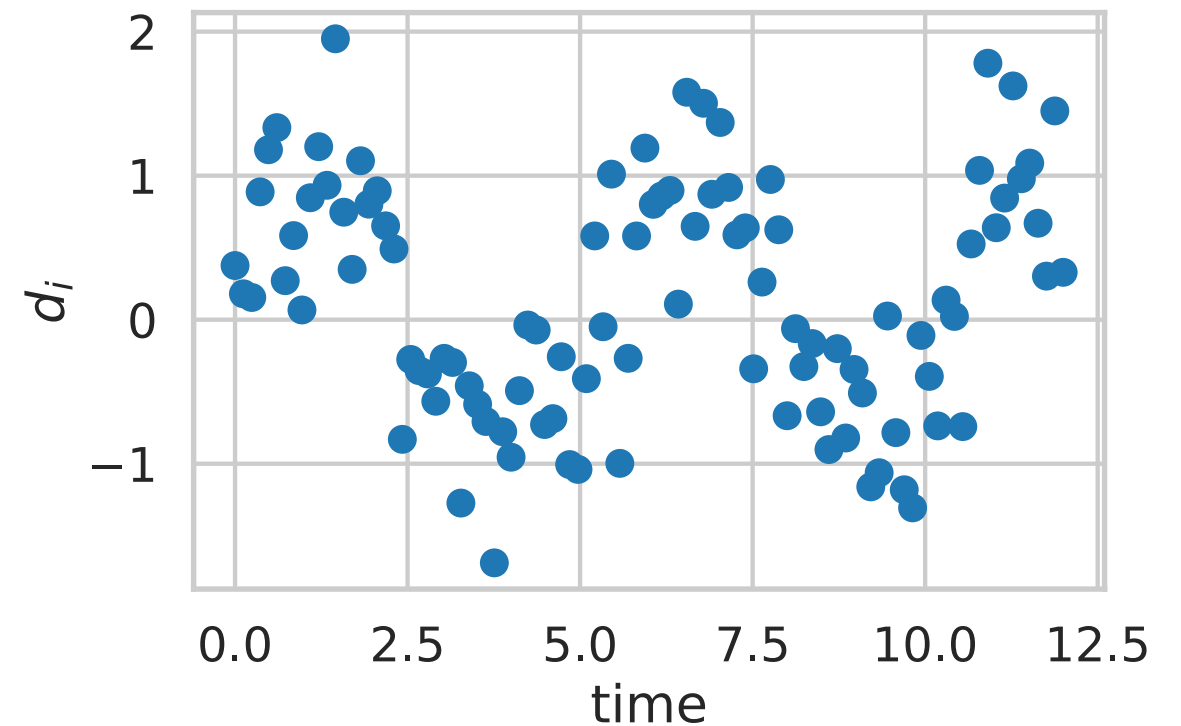
- For multiple observations,

$$\mathcal{L}(\mathbf{d} | \omega, M_A) = \prod_i \mathcal{L}(\mathbf{d}_i | \omega, M_A)$$

Or,

$$\ln \mathcal{L} = -\frac{1}{2} \sum_i \left(\frac{(d_i - y_A(t_i, \omega))^2}{\sigma^2} + \ln(2\pi\sigma^2) \right)$$

- Simulate $d_i = y(t_i, \omega_{\text{true}}) + n_i$, with say $\sigma = 0.1$ and $\omega_{\text{true}} = 1.2$.



One Dimensional Example

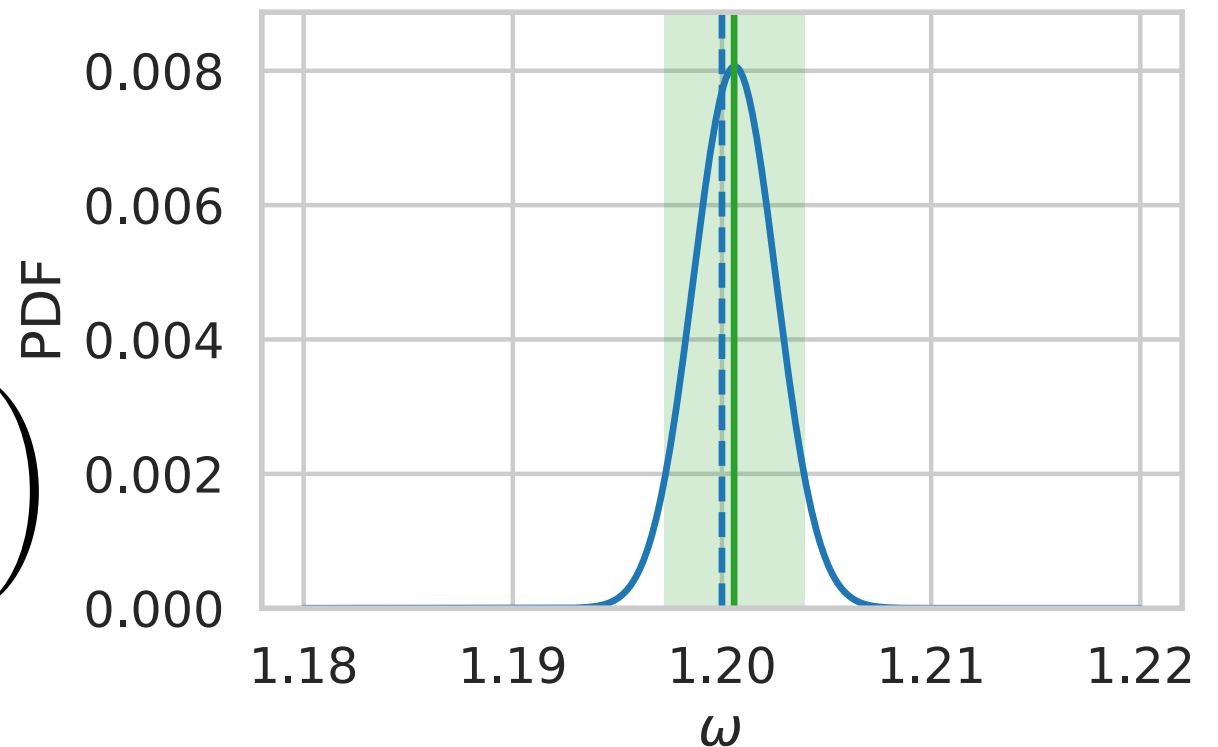
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Posterior probability distribution

- Compute the likelihood on grid of ω . The **posterior \propto likelihood** if prior is constant.
- Why is the **peak (median) not consistent with the true value?**

One Dimensional Example

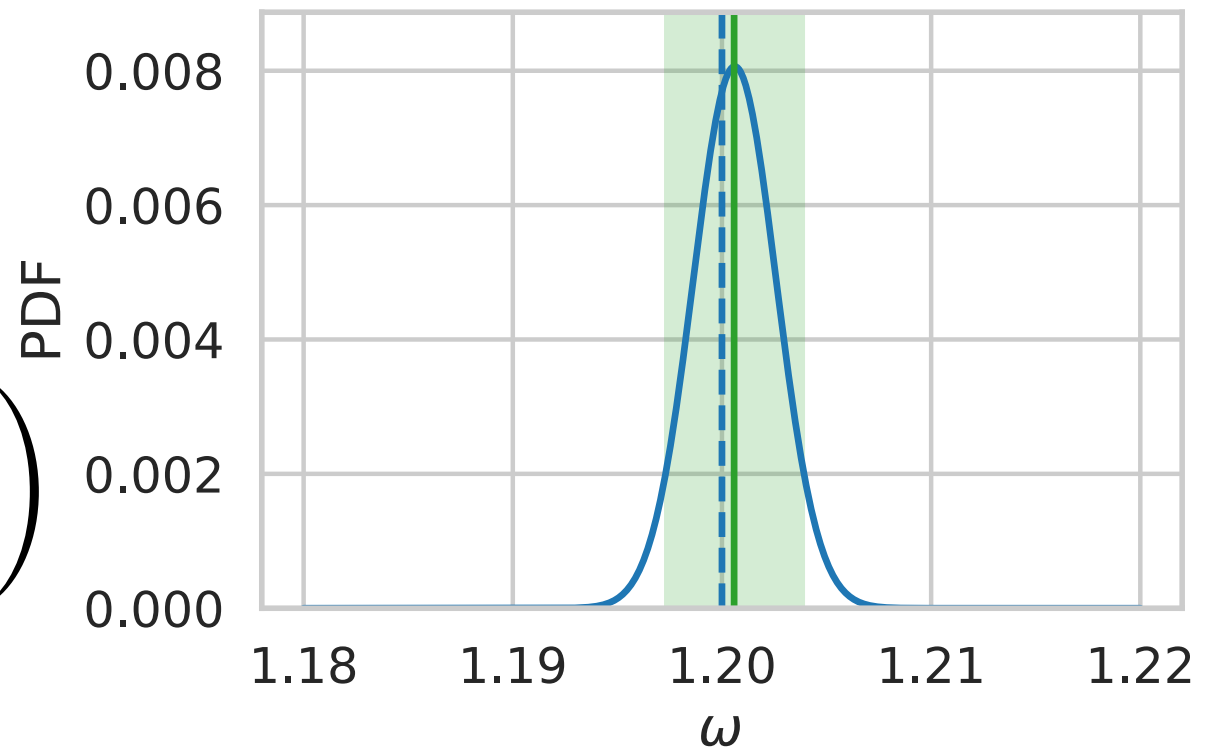
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Posterior probability distribution

- Compute the likelihood on grid of ω . The **posterior \propto likelihood** if prior is constant.
- Why is the **peak (median) not consistent with the true value?** **Noise!**
- How to quantify this? Bayesian answer is **credible interval**.
- We should always report inferences as ω has a median of XX and lies between YY and ZZ with 90% probability.

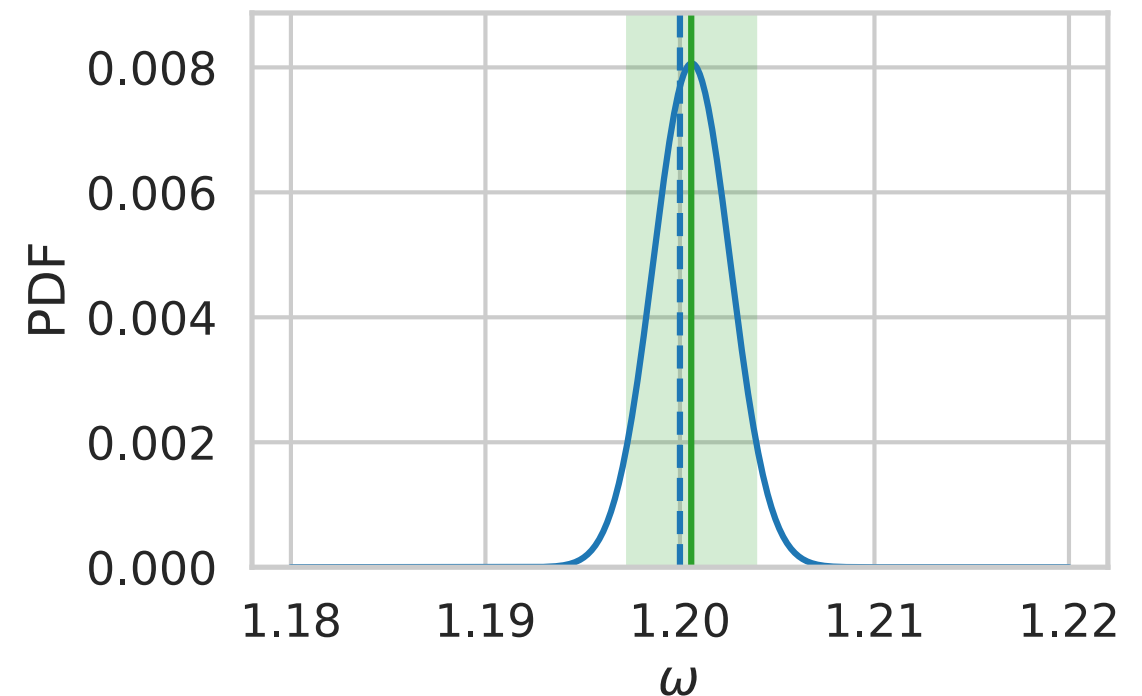
One Dimensional Example

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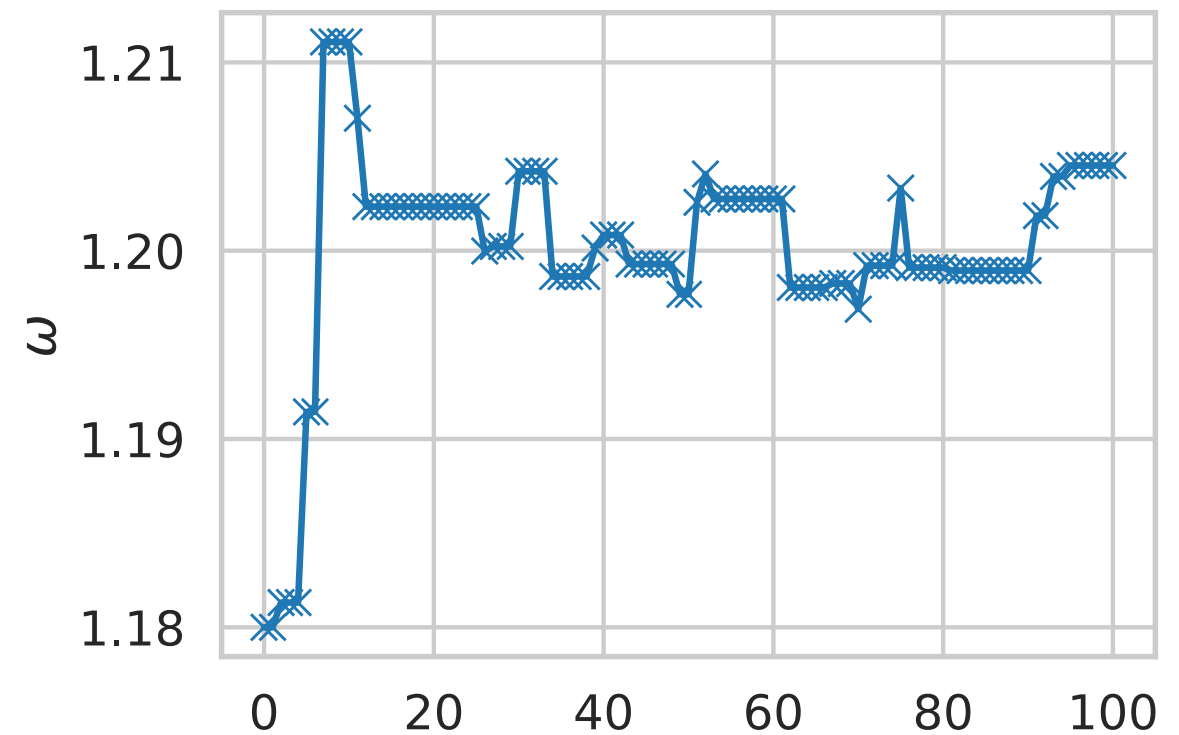
$$\ln \mathcal{L} = -\frac{1}{2} \sum_i \left(\frac{(d_i - y_A(t_i, \omega))^2}{\sigma^2} + \ln(2\pi\sigma^2) \right)$$



- Choose, say $\sigma = 0.1$ and $\omega_{\text{true}} = 1.2$, to simulate the data $d_i = y(t_i, \omega_{\text{true}}) + n_i$.
- Compute the likelihood on grid of ω . The **posterior \propto likelihood** if prior is constant.
- Why is the **peak (median) not consistent with the true value?** **Noise!**
- How to quantify this? Bayesian answer is **credible interval**.
- What if the dimensionality (D) is reasonably high, then number of computations \sim (no. of grid points)^D. Not feasible! Stochastic methods can come handy?!

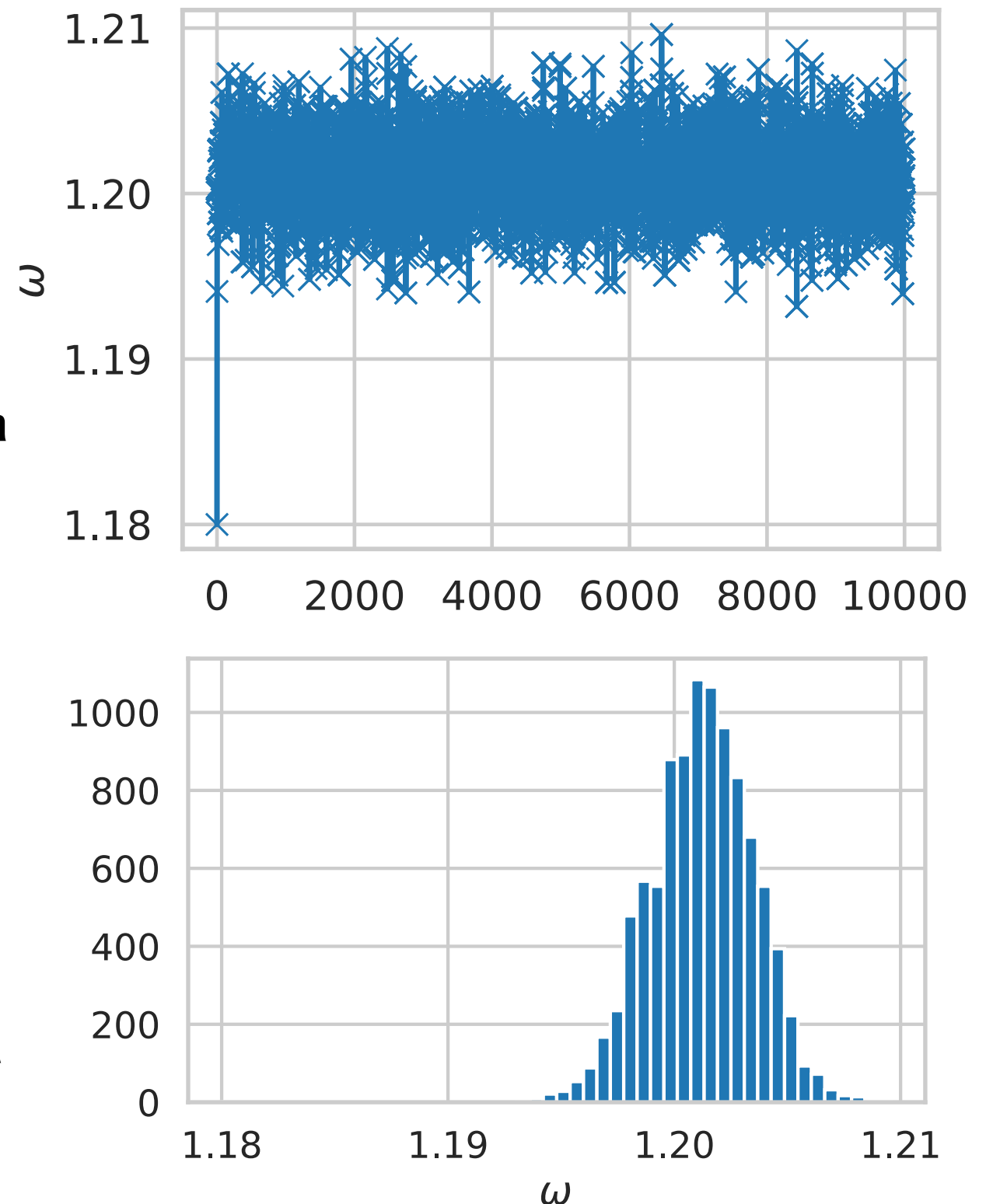
Metropolis-Hastings Sampler

- Let us add some “randomness” in the steps
 - Choose a random value, say ω_0 , and evaluate likelihood \mathcal{L}_0
 - Find a new $\omega_1 = \omega_0 + \delta\omega$ (random shift) and calculate likelihood \mathcal{L}_1 .
 - Store the new ω and likelihood \mathcal{L}_1 if $\mathcal{L}_1 > \mathcal{L}_0 \alpha$, where α is a random number $\in [0,1]$.
 - The samplers will walk both uphill and downhill.



Metropolis-Hastings Sampler

- Let us add some “**randomness**” in the steps
 - Choose a random value, say ω_0 , and evaluate likelihood \mathcal{L}_0
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 - Store the new ω and likelihood \mathcal{L}_1 if $\mathcal{L}_1 > \mathcal{L}_0 \alpha$, where α is a random number $\in [0,1]$.
 - The samplers will **walk both uphill and downhill**.
- Limitations: (i) not efficient if multimodal posterior.



Nested Sampling

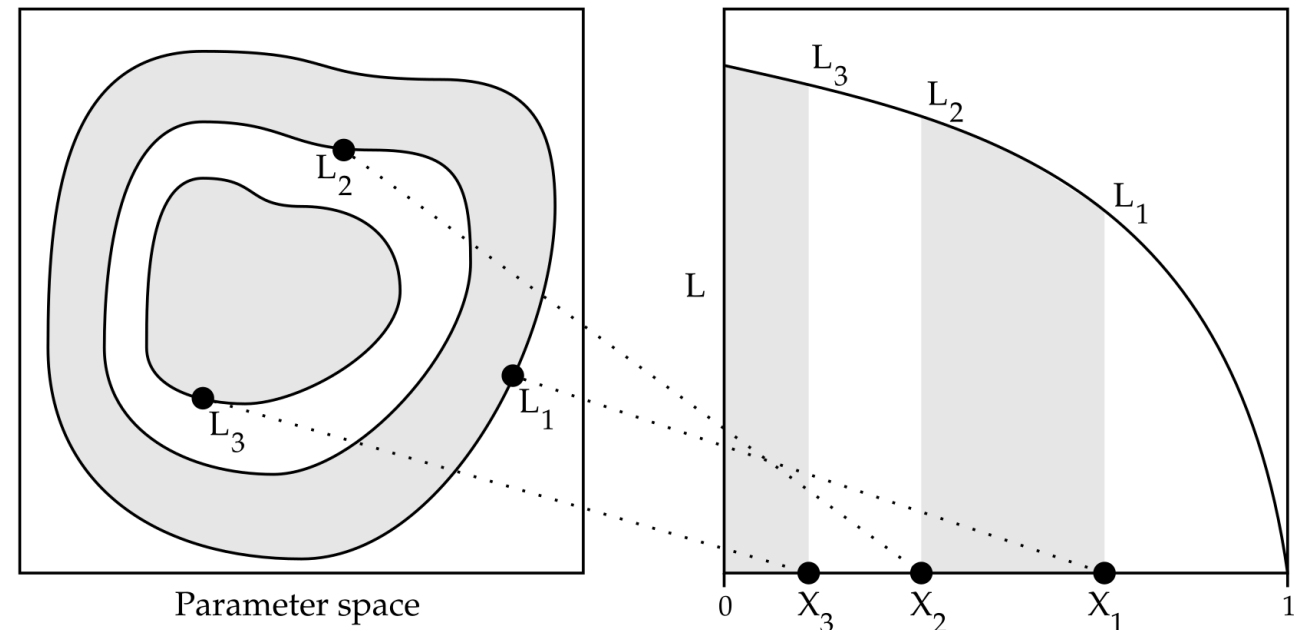
- Defining the **prior volume** as X such that $dX = p(\theta | M_A) d\theta$ where

$$X(\mathcal{L}) = \int_{p(d|\theta, M_A) > \mathcal{L}} d\theta p(\theta | M_A)$$

The total probability volume within a likelihood contour defined by $p(d | \theta, M_A) = \mathcal{L}$.

- The **evidence**,

$$Z \equiv p(d | M_A) = \int_0^1 \mathcal{L}(X) dX$$

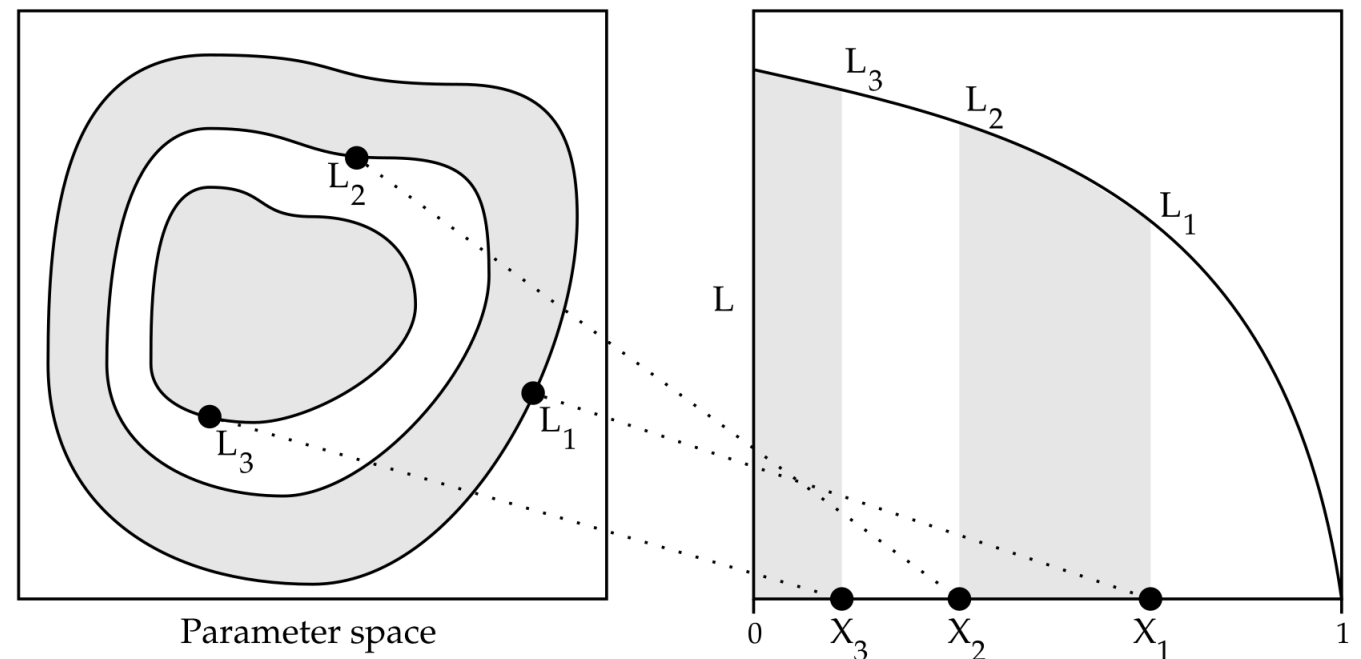


- Evaluating the likelihoods $\mathcal{L}_i = \mathcal{L}(X_i)$ associated with **monotonically decreasing sequence of prior volumes** $X_i: 0 < X_N < \dots < X_2 < X_1 < X_0 = 1$

$$Z = \sum_{i=1}^N \frac{1}{2} (X_{i+1} - X_i) \mathcal{L}_i \implies p(\theta | d, M_A) = \frac{\frac{1}{2} (X_{i+1} - X_i) \mathcal{L}_i}{Z}$$

Nested Sampling

- Select a set of initial **live points** sampled from the prior.
- The point with the **lowest likelihood** is replaced with a new sample with **higher likelihood**.



- Iterate this until reaching the stopping condition $\mathcal{L}_{\max} X_i / Z_i > e^{0.1}$ with \mathcal{L}_{\max} is the maximum likelihood value.
- Checking whether **the evidence estimate would change by more than a factor of ~ 0.1** if all the prior support were at the maximum likelihood.

GW Parameter Estimation: Bilby

- A generic **Bayesian Inference Library**.
- Special support to gravitational-wave transients.
- Structure

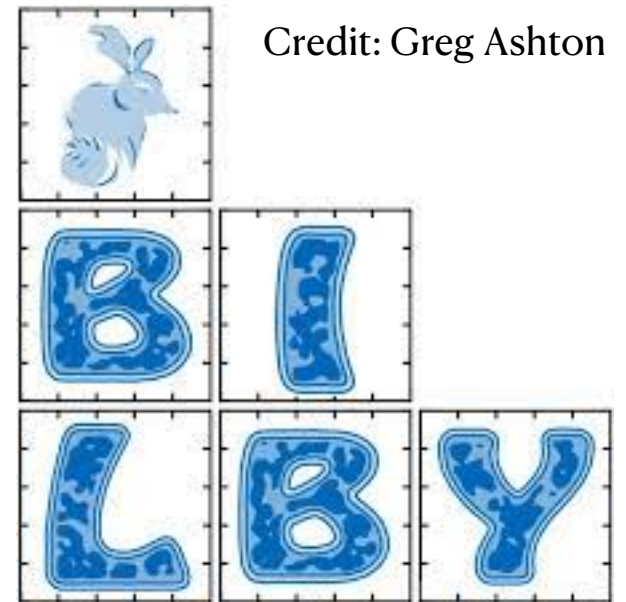
- Likelihood object

```
likelihood = bilby.gw.likelihood.base.GravitationalWaveTransient(  
    interferometers, waveform_generator, priors, ...  
)
```

- Priors as python dictionaries

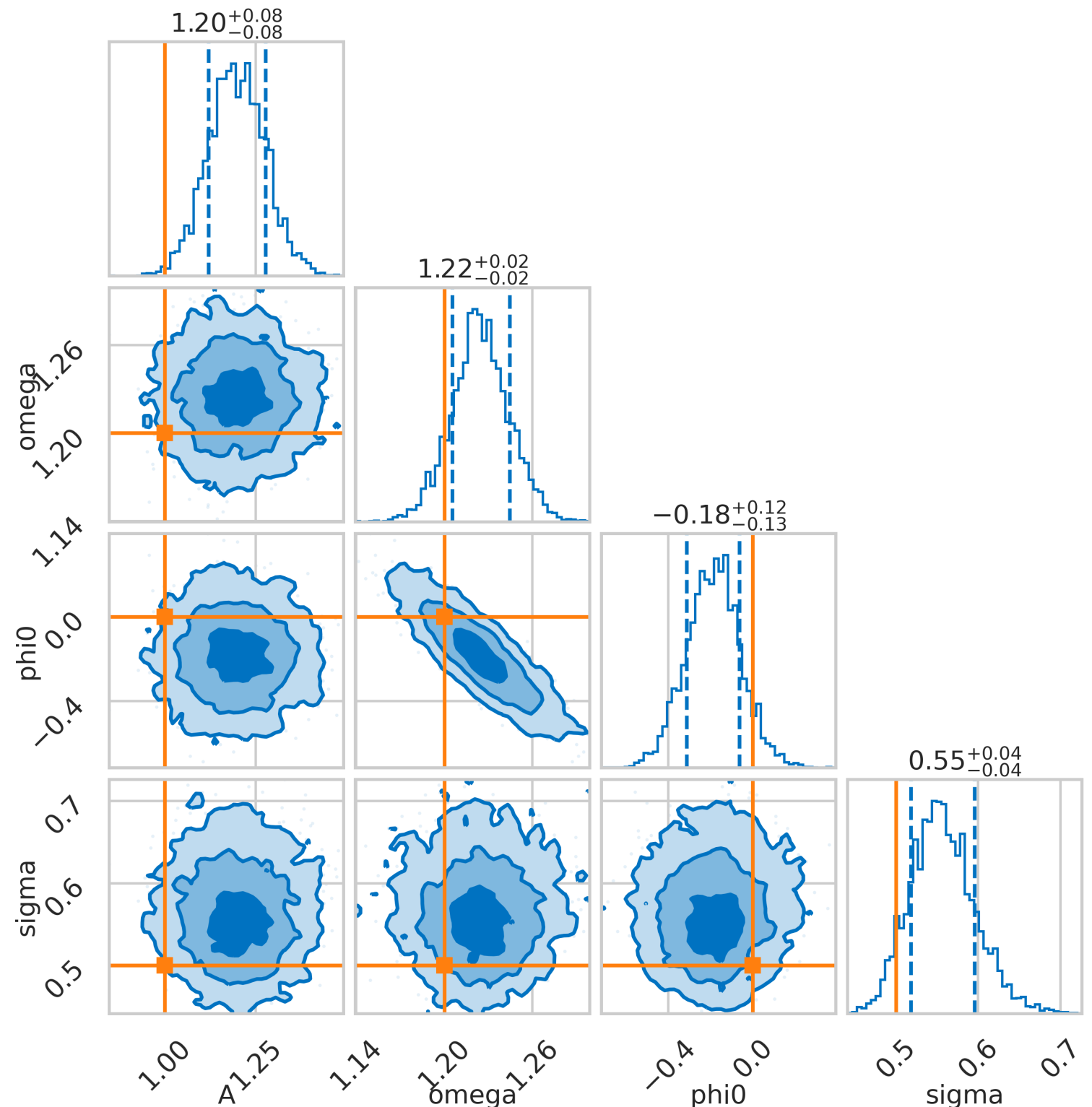
- Samplers: dynesty, pymultinest, ..., etc.

```
result = bilby.run_sampler(  
    likelihood, prior, sampler="dynesty", outdir="outdir",  
    label="GW150914",  
    nlive=500, dlogz=0.1, ...  
)
```

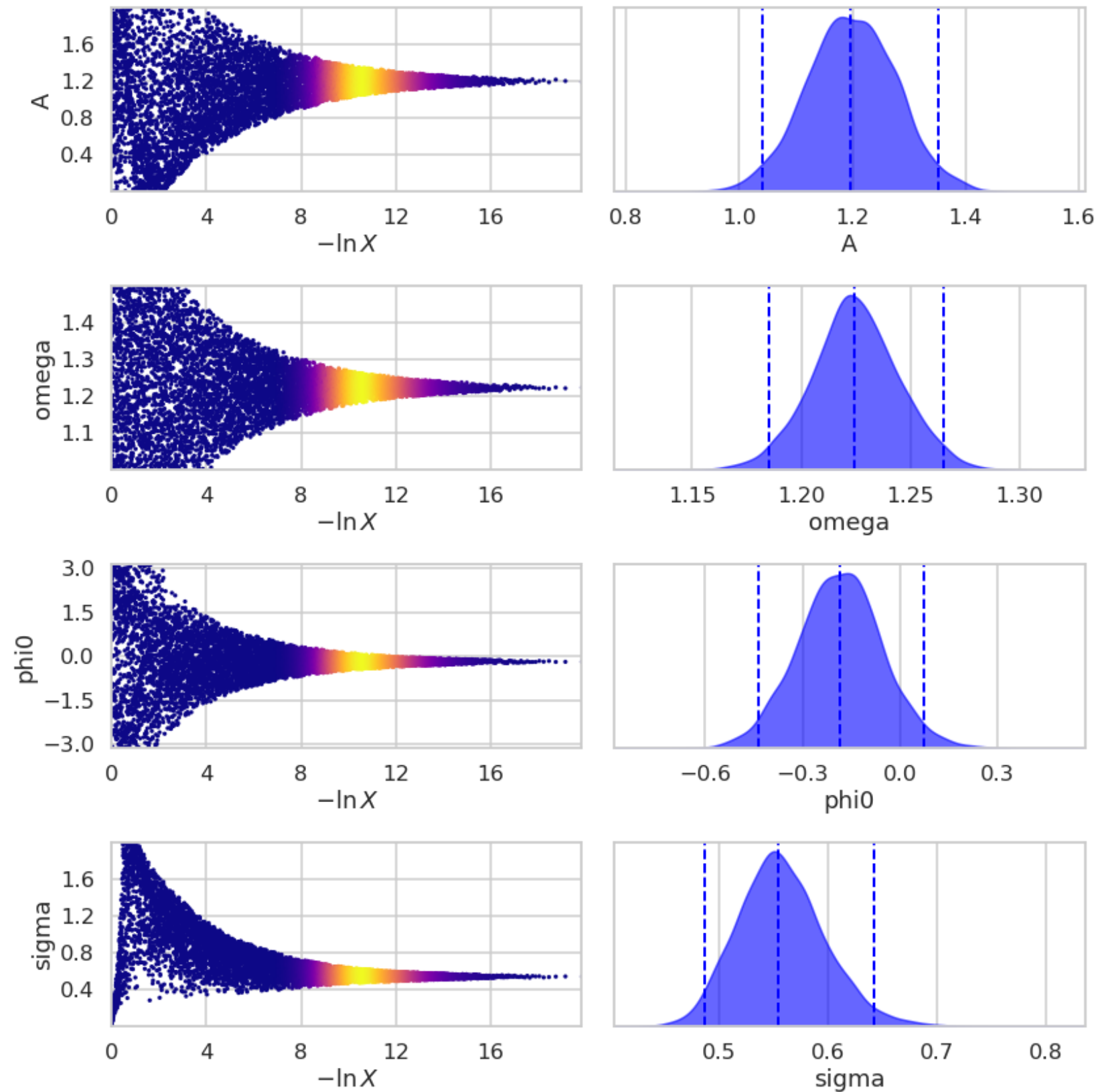


Bilby Output

- Let the model M_A
$$y(t) = A \sin(\omega t + \phi_0)$$
parameters A , ω , and ϕ_0 .
- Result object contains information about posteriors, priors, and likelihood, ..., etc.
- Just `result.plot_corner()` will give us



Bilby Trace Plots



Conclusion

- Parameter estimation of a compact binary merger in GW is a high dimensionality problem.
- Need stochastic samplers to sample the likelihood in such case.
- Output is probability distributions of the parameters due to noise uncertainty in the data.
- Bayesian inference is key to the parameter estimation in GW sources, especially for compact binary mergers.
- Bilby is one such Bayesian Inference Library to perform parameter estimation.

References

- Bilby: Ashton et al 2018 (<https://lscsoft.docs.ligo.org/bilby/>)
- Data Analysis: A Bayesian Tutorial by D. S. Sivia & J. Skilling
- An Introduction to Bayesian Inference in GW Astronomy,
Thrane & Talbot (2018)
- GWOSC: <https://www.gw-openscience.org/>
- GWpy: <https://gwpy.github.io/>
- PyCBC: <https://pycbc.org/>

**DON'T KNOW WHAT BAYES
THEOREM IS**

Thank you!

**AND AT THIS POINT I'M TO
AFFRAID TO ASK**