

# STAT 251 Formula Sheet

## Measures of Center

Mean:  $\bar{x} = \frac{\sum_{i=1}^n x_i}{n} = \frac{x_1 + x_2 + \dots + x_n}{n}$   
 Median: If  $n$  is even then  $\tilde{x} = \frac{(\frac{n}{2})^{\text{th}} \text{ obs.} + (\frac{n+1}{2})^{\text{th}} \text{ obs.}}{2}$   
 If  $n$  is odd then  $\tilde{x} = \frac{n+1}{2}^{\text{th}} \text{ obs.}$

## Measures of Variability

Range:  $R = x_{\text{largest}} - x_{\text{smallest}}$   
 Variance:  $s^2 = \frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n-1} = \frac{\sum_{i=1}^n x_i^2 - n\bar{x}^2}{n-1}$   
 Standard deviation:  $s = \sqrt{\frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n-1}} = \sqrt{\frac{\sum_{i=1}^n x_i^2 - n\bar{x}^2}{n-1}}$   
 IQR:  $\text{IQR} = Q_3 - Q_1$

### Method to compute $Q_{(p)}$ :

- Sort data from smallest to largest:  $x_{(1)} \leq x_{(2)} \leq \dots \leq x_{(n)}$
- Compute the number  $np + 0.5$
- If  $np + 0.5$  is an integer,  $m$ , then:  $Q_{(p)} = x_{(m)}$
- If  $np + 0.5$  is not an integer,  $m < np + 0.5 < m + 1$  for some integer  $m$ , then:  
 $Q_{(p)} = \frac{x_{(m)} + x_{(m+1)}}{2}$

### Outliers:

- Values smaller than  $Q_1 - (1.5 \times \text{IQR})$  are outliers
- Values greater than  $Q_3 + (1.5 \times \text{IQR})$  are outliers

## Discrete Random Variables

Consider a **discrete** random variable  $X$

**Probability Mass Function** (pmf):  $f(x) = P(X = x)$

- $f(x) \geq 0$  for all  $x$  in  $X$
- $\sum_x f(x) = 1$

**Cumulative Distributive Function** (cdf):  $F(x) = P(X \leq x) = \sum_{k \leq x} f(k)$

Mean ( $\mu$ ):  $E(X) = \sum_x x f(x)$   
 Expected value:  $E(g(X)) = \sum_x g(x) f(x)$   
 Variance ( $\sigma^2$ ):  $\text{Var}(X) = \sum_x (x - \mu)^2 f(x) = E(X^2) - [E(X)]^2$   
 SD ( $\sigma$ ):  $SD(X) = \sqrt{\text{Var}(X)}$

## Sets and Probability

### Properties of Probability:

- General Addition Rule:**  $P(A \cup B) = P(A) + P(B) - P(A \cap B)$
- Complement Rule:**  $P(A^c) = 1 - P(A)$
- If  $A \subseteq B$  then  $P(A \cap B) = P(A)$
- If  $A \subseteq B$  then  $P(A) \leq P(B)$
- $P(\emptyset) = 0$  and  $P(S) = 1$
- $0 \leq P(A) \leq 1$  for all  $A$

### Conditional Probability:

- $P(A|B) = \frac{P(A \cap B)}{P(B)}$  and  $P(B|A) = \frac{P(A \cap B)}{P(A)}$
- Multiplication Rule:**  $P(A \cap B) = P(B) \times P(A|B)$  and  $P(A \cap B) = P(A) \times P(B|A)$
- Events  $A$  and  $B$  are **independent** if and only if  $P(A \cap B) = P(A)P(B)$  and thus  $P(A|B) = P(A)$  and  $P(B|A) = P(B)$

**Bayes' Theorem:**  $P(A_i|B) = \frac{P(B|A_i)P(A_i)}{\sum_{i=1}^n P(A_i)P(B|A_i)}$   
 $= \frac{P(B|A_i)P(A_i)}{P(B|A_1)P(A_1) + P(B|A_2)P(A_2) + \dots + P(B|A_n)P(A_n)}$

## Continuous Random Variables

Consider a **continuous** random variable  $X$

**Probability Density Function** (pdf):  $P(a \leq X \leq b) = \int_a^b f(x) dx$

- $f(x) \geq 0$  for all  $x$
- $\int_{-\infty}^{\infty} f(x) dx = 1$

**Cumulative Distributive Function** (cdf):  $F(x) = P(X \leq x) = \int_{-\infty}^x f(t) dt$

Median:  $x$  such that  $F(x) = 0.5$   
 $Q_1$  and  $Q_3$ :  $x$  such that  $F(x) = 0.25$  and  $x$  such that  $F(x) = 0.75$   
 Mean ( $\mu$ ):  $E(X) = \int_{-\infty}^{\infty} x f(x) dx$   
 Expected value:  $E(g(X)) = \int_{-\infty}^{\infty} g(x) f(x) dx$   
 Variance ( $\sigma^2$ ):  $\text{Var}(X) = \int_{-\infty}^{\infty} (x - \mu)^2 f(x) dx = E(X^2) - [E(X)]^2$   
 SD ( $\sigma$ ):  $SD(X) = \sqrt{\text{Var}(X)}$

### Summarizing Main Features of $f(x)$

Consider two random variables  $X, Y$

#### Properties of Probability:

- $E(aX + b) = aE(X) + b$ , for  $a, b \in \mathbb{R}$
- $E(X + Y) = E(X) + E(Y)$ , for all pairs of  $X$  and  $Y$
- $E(XY) = E(X)E(Y)$ , for independent  $X$  and  $Y$
- $Var(aX + b) = a^2 Var(X)$ , for  $a, b \in \mathbb{R}$
- $Var(X + Y) = Var(X) + Var(Y)$   
 $Var(X - Y) = Var(X) + Var(Y)$ , for independent  $X$  and  $Y$

#### Covariance:

- $Cov(X, Y) = E[X - E(X)][Y - E(Y)] = E(XY) - E(X)E(Y)$   
If  $X$  and  $Y$  are independent,  $Cov(X, Y) = 0$
- $Var(X + Y) = Var(X) + Var(Y) + 2Cov(X, Y)$
- $Var(aX + bY + c) = a^2 Var(X) + b^2 Var(Y) + 2abCov(X, Y)$

### Sum and Average of Independent Random Variables

#### Sum of Independent Random Variables:

$Y = a_1X_1 + a_2X_2 + \dots + a_nX_n$ , for  $a_1, a_2, \dots, a_n \in \mathbb{R}$

- $E(Y) = a_1E(X_1) + a_2E(X_2) + \dots + a_nE(X_n)$
- $Var(Y) = a_1^2 Var(X_1) + a_2^2 Var(X_2) + \dots + a_n^2 Var(X_n)$

If  $n$  random variables  $X_i$  have common mean  $\mu$  and common variance  $\sigma^2$  then,

- $E(Y) = (a_1 + a_2 + \dots + a_n)\mu$
- $Var(Y) = (a_1^2 + a_2^2 + \dots + a_n^2)\sigma^2$

#### Average of Independent Random Variables:

$X_1, X_2, \dots, X_n$  are  $n$  independent random variables

- $\bar{X} = \frac{X_1 + X_2 + \dots + X_n}{n}$
- $E[\bar{X}] = \frac{1}{n}[E(X_1) + E(X_2) + \dots + E(X_n)]$
- $Var[\bar{X}] = \frac{1}{n^2}[Var(X_1) + Var(X_2) + \dots + Var(X_n)]$

If  $n$  random variables  $X_i$  have common mean  $\mu$  and common variance  $\sigma^2$  then,

- $E[\bar{X}] = \mu$
- $Var[\bar{X}] = \frac{\sigma^2}{n}$

### Maximum and Minimum of Independent Variables

Given  $n$  independent random variables  $X_1, X_2, \dots, X_n$ .

For each  $X_i$ , cdf  $F_X(x)$  and pdf is  $f_X(x)$ .

#### Maximum of Independent Random Variables:

Consider  $V = \max\{X_1, X_2, \dots, X_n\}$

##### cdf of V

$$\begin{aligned} F_V(v) &= P(V \leq v) = P(X_1 \leq v, X_2 \leq v, \dots, X_n \leq v) \\ &= P(X_1 \leq v)P(X_2 \leq v) \dots P(X_n \leq v) = F_{X_1}(v)F_{X_2}(v) \dots F_{X_n}(v) \\ &= [F_X(v)]^n ; \text{ if } X_i\text{'s are all identically distributed} \end{aligned}$$

##### pdf of V

$$\begin{aligned} f_V(v) &= F'_V(v) = \frac{d}{dv} F_V(v) = \frac{d}{dv} [F_X(v)]^n = n[F_X(v)]^{n-1} \frac{d}{dv} F_X(v) \\ &= n[F_X(v)]^{n-1} f_X(v) \end{aligned}$$

#### Minimum of Independent Random Variables:

Consider  $U = \min\{X_1, X_2, \dots, X_n\}$

##### cdf of U

$$\begin{aligned} F_U(u) &= P(U \leq u) = 1 - P(U > u) = 1 - P(X_1 > u, X_2 > u, \dots, X_n > u) \\ &= 1 - P(X_1 > u)P(X_2 > u) \dots P(X_n > u) \\ &= 1 - [1 - F_{X_1}(u)][1 - F_{X_2}(u)] \dots [1 - F_{X_n}(u)] \\ &= 1 - [1 - F_X(u)]^n ; \text{ if } X_i\text{'s are all identically distributed} \end{aligned}$$

##### pdf of U

$$\begin{aligned} f_U(u) &= F'_U(u) = \frac{d}{du} \{1 - [1 - F_X(u)]^n\} = 0 - n[1 - F_X(u)]^{n-1} \frac{d}{du} (-F_X(u)) \\ &= n[1 - F_X(u)]^{n-1} f_X(u) \end{aligned}$$

### Some Continuous Distributions

#### Uniform Distribution: $X \sim U(a, b)$

Mean:  $\mu = E(X) = \frac{a+b}{2}$

Variance:  $\sigma^2 = Var(X) = \frac{(b-a)^2}{12}$

##### pdf of X

$$f(x) = \begin{cases} \frac{1}{b-a} & a \leq x \leq b \\ 0 & \text{otherwise} \end{cases}$$

##### cdf of X

$$F(x) = \begin{cases} 0 & x < a \\ \frac{x-a}{b-a} & a \leq x \leq b \\ 1 & x > b \end{cases}$$

#### Exponential Distribution: $X \sim Exp(\lambda)$

Mean:  $\mu = E(X) = \frac{1}{\lambda}$

Variance:  $\sigma^2 = Var(X) = \frac{1}{\lambda^2}$

##### pdf of X

$$f(x) = \begin{cases} \lambda e^{-\lambda x} & x \geq 0 \\ 0 & x < 0 \end{cases}$$

##### cdf of X

$$F(x) = \begin{cases} 1 - e^{-\lambda x} & x \geq 0 \\ 0 & x < 0 \end{cases}$$

## Normal Distribution

**Normal Distribution:**  $X \sim N(\mu, \sigma^2)$

Standardized Normal:  $Z \sim N(0, 1)$  where  $Z = \frac{X - \mu}{\sigma}$

68-95-99.7 Rule:

- approximately 68% of observations fall within  $\sigma$  of  $\mu$
- approximately 95% of observations fall within  $2\sigma$  of  $\mu$
- approximately 99.7% of observations fall within  $3\sigma$  of  $\mu$

## Bernoulli and Binomial Random Variables

**Bernoulli Random Variable:**

Bernoulli random variable  $X$  has only two outcomes, success and failure.

$P(\text{Success}) = p$  and  $P(\text{Failure}) = 1 - p$

**Bernoulli Distribution:**

$X \sim \text{Bernoulli}(p)$

pmf:  $P(X = x) = p^x(1 - p)^{1-x}$  for  $x = 0, 1$

Mean:  $\mu = E(X) = p$

Variance:  $\sigma^2 = \text{Var}(X) = p(1 - p)$

**Binomial Random Variable:**

Binomial random variable  $X$  is the number of successes for  $n$  independent trials and each trial has the same probability of success  $p$ .

**Binomial Distribution:**

$X \sim \text{Bin}(n, p)$

pmf:  $P(X = x) = \binom{n}{x} p^x (1 - p)^{n-x}$  for  $x = 0, 1, 2, \dots, n$

cdf:  $P(X \leq x) = \sum_{i=0}^x \binom{n}{i} p^i (1 - p)^{n-i}$  for  $x = 0, 1, 2, \dots, n$

Note:  $\binom{n}{x} = \frac{n!}{x!(n-x)!}$

Mean:  $\mu = E(X) = np$

Variance:  $\sigma^2 = \text{Var}(X) = np(1 - p)$

## Geometric Distribution

**Geometric Random Variable:**

Geometric random variable  $X$  is the number of independent trials needed until the first success occurs.

**Geometric Distribution:**

$X \sim \text{Geo}(p)$  where  $p$  is the probability of success

pmf:  $P(X = x) = p(1 - p)^{x-1}$  for  $x = 1, 2, 3, \dots$

cdf:  $P(X \leq x) = 1 - (1 - p)^x$  for  $x = 1, 2, 3, \dots$

Mean:  $\mu = E(X) = \frac{1}{p}$

Variance:  $\sigma^2 = \text{Var}(X) = \frac{1-p}{p^2}$

## Poisson Distribution

**Poisson Process:**

Random variable  $X$  is the number of occurrences in a given interval.

**Poisson Distribution:**

$X \sim \text{Poisson}(\lambda)$  where  $\lambda$  is the rate of occurrences

pmf:  $P(X = x) = \frac{\lambda^x e^{-\lambda}}{x!}$  for  $x = 0, 1, 2, 3, \dots$

cdf:  $P(X \leq x) = \sum_{i=0}^x \frac{\lambda^i e^{-\lambda}}{i!}$  for  $x = 0, 1, 2, 3, \dots$

Mean:  $\mu = E(X) = \lambda$

Variance:  $\sigma^2 = \text{Var}(X) = \lambda$

- Let  $T \sim \text{Exp}(\lambda)$  be the time between two consecutive occurrences of events. (Can also be the waiting time for first event.)

## Poisson Approximation to the Binomial Distribution

Let  $X \sim \text{Bin}(n, p)$  be a binomial random variable. If  $n$  is large ( $n \geq 20$ ) and  $p$  or  $1 - p$  is small ( $np < 5$  or  $n(1 - p) < 5$ ), then we can use a Poisson random variable with rate  $\lambda = np$  to approximate the probabilistic behaviour of  $X$ .

$X \sim \text{Poisson}(np)$ , approx. for  $x = 0, 1, 2, \dots, n$

## Central Limit Theorem

Let  $X_1, X_2, \dots, X_n$  be a random sample from an arbitrary population/distribution with mean  $\mu$  and variance  $\sigma^2$ . When  $n$  is large ( $n \geq 20$ ) then

$\bar{X} = \frac{X_1 + X_2 + \dots + X_n}{n} \sim N(\mu, \frac{\sigma^2}{n})$ , approx.

When dealing with sum, the CLT can still be used. Then

$T = X_1 + X_2 + \dots + X_n = n\bar{X}$

$T \sim N(n\mu, n\sigma^2)$ , approx.

## Normal Approximation to the Binomial Distribution

Let  $X \sim \text{Bin}(n, p)$ . When  $n$  is large so that both  $np \geq 5$  and  $n(1 - p) \geq 5$ . We can use the normal distribution to get an approximate answer. Remember to use **continuity correction**.

$X \sim N(np, np(1 - p))$ , approx.

## Normal Approximation to the Poisson Distribution

Let  $X \sim \text{Poisson}(\lambda)$ . When  $\lambda$  is large ( $\lambda \geq 20$ ) then the Normal distribution can be used to approximate the Poisson distribution. Remember to use **continuity correction**.

$X \sim N(\lambda, \lambda)$ , approx.

### Continuity Correction

Consider continuous random variable  $Y$  and discrete random variable  $X$ .

- $P(X > 4) = P(X \geq 5) = P(Y \geq 4.5)$
- $P(X \geq 4) = P(Y \geq 3.5)$
- $P(X < 4) = P(X \leq 3) = P(Y \leq 3.5)$
- $P(X \leq 4) = P(Y \leq 4.5)$
- $P(X = 4) = P(3.5 \leq Y \leq 4.5)$

### Point Estimators

Suppose that  $X_1, X_2, \dots, X_n$  are random samples from a population with mean  $\mu$  and variance  $\sigma^2$ .

- $\bar{x}$  is an unbiased estimator of  $\mu$   

$$\bar{x} = \frac{\sum_{i=1}^n x_i}{n}$$
- $s^2$  is an unbiased estimator of  $\sigma^2$   

$$s^2 = \frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n-1} = \frac{\sum_{i=1}^n x_i^2 - n\bar{x}^2}{n-1}$$
- $\theta$  is the parameter,  $\hat{\theta}$  is the point estimator. When  $E(\hat{\theta}) = \theta$ ,  $\hat{\theta}$  is an unbiased estimator. The bias of an estimator is  $bias(\hat{\theta}) = E(\hat{\theta}) - \theta$ .

### Confidence Interval

$(1 - \alpha)100\%$  **Confidence Interval for population mean  $\mu$ :**  
(point estimator of  $\mu$  is  $\bar{x}$ )

*General Form:* point estimate  $\pm$  margin of error

When  $\sigma^2$  is **known**:  $\bar{x} \pm z_{\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}}$

When  $\sigma^2$  is **unknown**:  $\bar{x} \pm t_{\frac{\alpha}{2}, n-1} \frac{s}{\sqrt{n}}$

Typical  $z$  values of  $\alpha$ :

$$\alpha = 0.1 \quad 90\% \quad z_{\frac{\alpha}{2}} = z_{0.05} = 1.645$$

$$\alpha = 0.05 \quad 95\% \quad z_{\frac{\alpha}{2}} = z_{0.025} = 1.96$$

$$\alpha = 0.01 \quad 99\% \quad z_{\frac{\alpha}{2}} = z_{0.005} = 2.575$$

$(1 - \alpha)100\%$  **Confidence Interval for  $\mu_1 - \mu_2$ :**

$$(\bar{x}_1 - \bar{x}_2) \pm t_{\frac{\alpha}{2}, n_1+n_2-2} s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}$$

### Pooled Standard Deviation

Requires assumptions that population variances are equal:  $\sigma_1^2 = \sigma_2^2 = \sigma^2$

The pooled standard deviation  $s_p$  estimates the common standard deviation  $\sigma$ .

$$s_p = \sqrt{\frac{(n_1-1)s_1^2 + (n_2-1)s_2^2}{n_1+n_2-2}}$$

### Testing of Hypotheses about $\mu$

$H_o$ : Null hypothesis is a tentative assumption about a population parameter.

$H_a$ : Alternative hypothesis is what the test is attempting to establish.

- $H_o : \mu \geq \mu_o$  vs  $H_a : \mu < \mu_o$  (one-tail test, lower-tail)
- $H_o : \mu \leq \mu_o$  vs  $H_a : \mu > \mu_o$  (one-tail test, upper-tail)
- $H_o : \mu = \mu_o$  vs  $H_a : \mu \neq \mu_o$  (two-tail test)

#### Test Statistic:

Case 1:  $\sigma^2$  is known

$$z = \frac{\bar{x} - \mu_o}{\frac{\sigma}{\sqrt{n}}} \sim N(0, 1)$$

Case 2:  $\sigma^2$  is unknown

$$t = \frac{\bar{x} - \mu_o}{\frac{s}{\sqrt{n}}} \sim t_{n-1}$$

#### Type I and Type II errors:

Type I error: rejecting  $H_o$  when  $H_o$  is true

Type II error: not rejecting  $H_o$  with  $H_o$  is false

$$P(\text{Type I error}) = \alpha$$

$$P(\text{Type II error}) = \beta$$

**Power** is the probability of rejecting  $H_o$ , when  $H_o$  is false.

$$\text{Power} = 1 - \beta$$

#### Comparison of two means:

Two independent populations with means  $\mu_1$  and  $\mu_2$ .

Assume random samples, normal distributions, and equal variances ( $\sigma_1^2 = \sigma_2^2$ ).

- $H_o : \mu_1 - \mu_2 \geq \Delta_o$  vs  $H_a : \mu_1 - \mu_2 < \Delta_o$  (lower-tail)
- $H_o : \mu_1 - \mu_2 \leq \Delta_o$  vs  $H_a : \mu_1 - \mu_2 > \Delta_o$  (upper-tail)
- $H_o : \mu_1 - \mu_2 = \Delta_o$  vs  $H_a : \mu_1 - \mu_2 \neq \Delta_o$  (two-tail)

#### Test Statistic:

$$t = \frac{(\bar{x}_1 - \bar{x}_2) - \Delta_o}{s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} \sim t_{n_1+n_2-2}$$

$s_p$  is the pooled standard deviation

#### Rejection Rules:

Consider test statistic  $z$ , and significance value  $\alpha$ .

- Lower-tail test: Reject  $H_o$  if  $z \leq z_\alpha$
- Upper-tail test: Reject  $H_o$  if  $z \geq z_\alpha$
- Two-tail test: Reject  $H_o$  if  $|z| \geq z_{\frac{\alpha}{2}}$

## Analysis of Variance (ANOVA)

### One-way ANOVA:

$k$  = number of populations or treatments being compared

$\mu_1$  = mean of population 1 or true average response when treatment 1 is applied.

...

$\mu_k$  = mean of population  $k$  or true average response when treatment  $k$  is applied.

### Assumptions:

- For each population, response variable is normally distributed
- Variance of response variable,  $\sigma^2$  is the same for all the populations
- The observations must be independent

### Hypotheses:

$$H_o : \mu_1 = \mu_2 = \dots = \mu_k$$

$$H_a : \mu_i \neq \mu_j \text{ for } i \neq j$$

### Notation:

$y_{ij}$  is the  $j^{th}$  observed value from the  $i^{th}$  population/treatment.

$$\text{Total mean: } \bar{y}_{i.} = \frac{y_{i.}}{n_i} = \frac{\sum_{j=1}^{n_i} y_{ij}}{n_i}$$

$$\text{Total sample size: } n = n_1 + n_2 + \dots + n_k$$

$$\text{Grand total: } y_{..} = \sum_{i=1}^k \sum_{j=1}^{n_i} y_{ij}$$

$$\text{Grand mean: } \bar{y}_{..} = \frac{y_{..}}{n} = \frac{\sum_{i=1}^k \sum_{j=1}^{n_i} y_{ij}}{n}$$

$$s^2 = \frac{\sum_{j=1}^k (n_i - 1) s_i^2}{n - k} = \text{MSE, where } s_i^2 = \frac{\sum_{j=1}^{n_i} (y_{ij} - \bar{y}_{i.})^2}{n_i - 1}$$

### The ANOVA Table:

Source of Variation	df	Sum of Squares	Mean Square	F-ratio
Treatment	$k - 1$	SSTr	$\text{MSTr} = \frac{\text{SSTr}}{k - 1}$	$\frac{\text{MSTr}}{\text{MSE}}$
Error	$n - k$	SSE	$\text{MSE} = \frac{\text{SSE}}{n - k}$	
Total	$n - 1$	SST		

$$\text{SST} = \text{SSTr} + \text{SSE}$$

$$\text{SST} = \sum_{i=1}^k \sum_{j=1}^{n_i} (y_{ij} - \bar{y}_{..})^2 = \sum_{i=1}^k \sum_{j=1}^{n_i} y_{ij}^2 - \frac{1}{n} y_{..}^2$$

$$\text{SSTr} = \sum_{i=1}^k \sum_{j=1}^{n_i} (\bar{y}_{i.} - \bar{y}_{..})^2 = \sum_{i=1}^k \frac{1}{n_i} y_{i.}^2 - \frac{1}{n} y_{..}^2$$

$$\text{SSE} = \sum_{i=1}^k \sum_{j=1}^{n_i} (y_{ij} - \bar{y}_{i.})^2 = \sum_{i=1}^k \sum_{j=1}^{n_i} y_{ij}^2 - \sum_{i=1}^k \frac{y_{i.}^2}{n_i} = \sum_{i=1}^k (n_i - 1) s_i^2$$

### Test Statistic:

$$F_{obs} = \frac{\text{MSTr}}{\text{MSE}} \sim F_{v_1, v_2}$$

$$v_1 = df(\text{SSTr}) = k - 1$$

$$v_2 = df(\text{SSE}) = n - k$$

Reject  $H_o$  if  $F_{obs} \geq F_{\alpha, v_1, v_2}$

## Covariance and Correlation Coefficient

On a scatter plot, each observation is represented as a point with x-coord  $x_i$  and y-coord  $y_i$ .

### Sample Covariance: $Cov(x, y)$

$$\begin{aligned} Cov(x, y) &= \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y}) \\ &= \frac{1}{n-1} \left[ \sum_{i=1}^n x_i y_i - \frac{\sum_{i=1}^n x_i \sum_{i=1}^n y_i}{n} \right] \\ &= \frac{1}{n-1} [\sum_{i=1}^n x_i y_i - n \bar{x} \bar{y}] \end{aligned}$$

- If  $x$  and  $y$  are positively associated, then  $Cov(x, y)$  will be large and positive
- If  $x$  and  $y$  are negatively associated, then  $Cov(x, y)$  will be large and negative
- If the variables are not positively nor negatively associated, then  $Cov(x, y)$  will be small

### Sample Correlation Coefficient: $r$

$$r = \frac{1}{n-1} \sum_{i=1}^n \left( \frac{x_i - \bar{x}}{s_x} \right) \left( \frac{y_i - \bar{y}}{s_y} \right), \text{ where } s_x = \sqrt{\frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n-1}} \text{ and } s_y = \sqrt{\frac{\sum_{i=1}^n (y_i - \bar{y})^2}{n-1}}$$

$$r = \frac{Cov(x, y)}{s_x s_y}$$

- Always falls between -1 and +1
- A positive  $r$  value indicates a positive association
- A negative  $r$  value indicates a negative association
- $r$  value close to +1 or -1 indicates a strong linear association
- $r$  value close to 0 indicates a weak association

## Simple Linear Regression

### Regression Line:

Simple linear regression model:  $y = \beta_o + \beta_1 x + \varepsilon$

$\beta_o$ ,  $\beta_1$ , and  $\sigma^2$  are parameters,  $y$  and  $\varepsilon$  are random variables.  $\varepsilon$  is the error term.

True regression line:  $E(y) = \beta_o + \beta_1 x$

Least squares regression line:  $\hat{y} = \hat{\beta}_o + \hat{\beta}_1 x$

$\hat{y}$ ,  $\hat{\beta}_o$ , and  $\hat{\beta}_1$  are point estimates for  $y$ ,  $\beta_o$ , and  $\beta_1$ .

Residual:  $\varepsilon_i = y_i - \hat{y}_i$

$$\hat{\beta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2} = \frac{\sum_{i=1}^n x_i y_i - n \bar{x} \bar{y}}{\sum_{i=1}^n x_i^2 - n \bar{x}^2} = r \frac{s_y}{s_x}$$

$$\hat{\beta}_o = \frac{\sum_{i=1}^n y_i - \hat{\beta}_1 \sum_{i=1}^n x_i}{n} = \bar{y} - \hat{\beta}_1 \bar{x}$$

### Coefficient of Determination: $r^2$

The proportion of observed  $y$  variation that can be explained by the simple linear regression model.

### Estimating $\sigma^2$ (SLR)

$$\hat{\sigma}^2 = s^2 = \frac{SSE}{n-2} = \frac{\sum_{i=1}^n (y_i - \hat{y}_i)^2}{n-2}$$

#### Error Sum of Squares (SSE):

$$SSE = \sum_{i=1}^n \varepsilon_i^2 = \sum_{i=1}^n (y_i - \hat{y}_i)^2 = \sum_{i=1}^n [y_i - (\hat{\beta}_0 + \hat{\beta}_1 x_i)]^2$$

SSE is a measure of variation in  $y$  left unexplained by linear regression model.

#### Total Sum of Squares (SST):

$$SST = \sum_{i=1}^n (y_i - \bar{y})^2$$

SST is sum of squared deviations about sample mean of observed  $y$  values.

#### Regression Sum of Squares (SSR):

$$SSR = \sum_{i=1}^n (\hat{y}_i - \bar{y})^2$$

SSR is total variation explained by the linear regression model.

$$SST = SSR + SSE$$

#### Coefficient of Determination from SST, SSR, and SSE:

$$r^2 = 1 - \frac{SSE}{SST}$$

or

$$r^2 = \frac{SSR}{SST}$$

### Slope Parameter $\beta_1$ (SLR)

When  $\beta_1 = 0$  there is no linear relationship between the two variables.

#### Hypotheses:

$$H_o : \beta_1 = 0$$

$$H_a : \beta_1 \neq 0$$

#### Test statistic:

$$t_{\text{obs}} = \frac{\hat{\beta}_1}{s_{\hat{\beta}_1}} \sim t_{n-2}, \text{ where } s_{\hat{\beta}_1} = \frac{s}{s_x \sqrt{n-1}}$$

Reject  $H_o$  if  $|t_{\text{obs}}| \geq t_{\frac{\alpha}{2}, n-2}$

$(1 - \alpha)100\%$  **Confidence Interval for  $\beta_1$ :**

$$\hat{\beta}_1 \pm t_{\frac{\alpha}{2}, n-2} s_{\hat{\beta}_1}$$