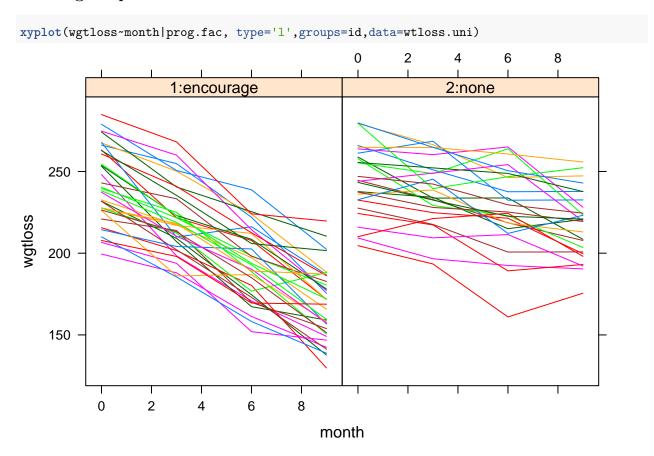
Complex Data - lab3

Stanisław Wilczyński

Data input and Calculating Means

```
wtloss <- read.table("../data/weightloss.dat",
header=F)
## Give names to variables
names(wtloss) <- c("id", paste("y", 1:4, sep=""), "program")
## Univariate format
wtloss.uni <- data.frame(id=rep(wtloss$id, each=4),
wgtloss=as.numeric(t(as.matrix(wtloss[,2:5]))),
program=rep(wtloss$program, each=4),
month=seq(0,9,3),
time.cat=rep(1:4))
wtloss.uni$prog.fac <- factor(wtloss.uni$program, labels=c("1:encourage", "2:none"))
attach(wtloss.uni)
wgt.mean <- tapply(wgtloss, list(month, program), mean)
wgt.sd <- tapply(wgtloss, list(month, program), sd)
detach(wtloss.uni)</pre>
```

Plotting Response Profiles



Choice of covariance structure

```
## CATEGORICAL time, UN, REML
wtloss.un.cat <- gls(wgtloss~factor(month)*prog.fac,</pre>
correlation=corSymm(form= ~1 | id),
weights=varIdent(form= ~1 | month),
data=wtloss.uni)
#summary(wtloss.un.cat)
## CATEGORICAL time, CS, REML
wtloss.cs.cat <- gls(wgtloss~factor(month)*prog.fac,</pre>
correlation=corCompSymm(form= ~1 | id),
weights=varIdent(form= ~1),
data=wtloss.uni)
#summary(wtloss.cs.cat)
## CATEGORICAL time, AR1, REML
wtloss.ar1.cat <- gls(wgtloss~factor(month)*prog.fac,</pre>
correlation=corAR1(form= ~1 | id),
weights=varIdent(form= ~1),
data=wtloss.uni)
#summary(wtloss.ar1.cat)
```

Unstructured vs compound symmetry

 H_0 : compound symmetry is adequate for the data

 H_1 : unstructured is required

```
un_vs_cs <- anova(wtloss.un.cat, wtloss.cs.cat)</pre>
```

	CS vs. UN	
	-2 REML	Number of
Structure	Log-Likelihood	Cov. Parameters
Compound Symmetry	1905.8232573	10
Unstructured	1900.8766893	18
Difference	4.9465681	8

LRT yields $G^2 = 4.9465681$ with 8 df (p-value = 0.7632715) so we do not reject the null hypothesis at $\alpha = 0.05$ and conclude that the assumption of compound symmetry covariance structure is adequate for the data.

Unstructured vs autoregressive

 H_0 : autoregressive is adequate for the data H_1 : unstructured is required

un_vs_ar <- anova(wtloss.un.cat, wtloss.ar1.cat)</pre>

	AR-1 vs. UN	
	-2 REML	Number of
Structure	Log-Likelihood	Cov. Parameters
Autoregressive	1931.246837	10
Unstructured	1900.8766893	18
Difference	30.3701477	8

LRT yields $G^2 = 30.3701477$ with 8 df (p-value = 1.8178034×10^{-4}), so we reject the null hypothesis at $\alpha = 0.05$ and conclude that the assumption of autoregressive covariance structure is not adequate when compared to unstructured.

Autoregressive vs compound symmetry

```
cs_vs_ar1 <- anova(wtloss.cs.cat, wtloss.ar1.cat)
```

Since CS and AR-1 have the same number of parameters = 10, no LRT is necessary. We can directly compare their likelihoods, or -2*log(likelihood):

- -2 * log(likelihood) for CS = 1905.8232573
- -2 * log(likelihood) for AR-1 = 1931.246837

Since -2 * log(likelihood) for CS is smaller than for AR-1, CS has a higher likelihood and we conclude that CS is an adequate model for the covariance structure when compared to AR-1.

The most adequate covariance structure is compound symmetry model.

Are all these tests correct?

No the last test is theoretically not correct, because for non nested models we can't just compare log-likelihoods or perform LRT. We have to compare AIC or BIC instead. However, in this case the number of parameters is the same, so comparing likelihoods is equivalent to comparing AICs.

	-2 REML	Number of	
Structure	Log-Likelihood	Parameters	AIC
Compound Symmetry	1905.8232573	10	1925.8232573
Autoregressive	1931.246837	10	1951.246837
Unstructured	1900.8766893	18	1936.8766893

Thus, we will use a compound symmetry covariance structure for the remainder of the lab.

Single Degree of Freedom Contrasts

AUC - test for equality of the area under the curve in two groups.

```
t <- wtloss.uni$month
L2 <- 0.5*c(t[1] + t[2] - 2*t[4], t[3] - t[1], t[4]-t[2], t[4]-t[3])
coefs <- summary(wtloss.cs.cat)$coefficients
AUC_enc <- sum(L2 * wgt.mean[,1])
AUC_none <- sum(L2 * wgt.mean[,2])
y_mean_enc <- mean(wtloss.uni$wgtloss[which(wtloss.uni$program==1)])</pre>
```

The obtained values of AUCs are:

Estimated mean AUC in encouragement program is -319.5705882. Estimated mean AUC in no encouragement program is -117.7038462.

As written in the textbook to test for the equality of AUCs, we test for the contrast $(-L_2, L_2)$.

```
L <- c(-L2, L2)
anova(wtloss.cs.cat, L=L)
```

```
## Denom. DF: 232
   F-test for linear combination(s)
                                                 factor(month)3
##
                     (Intercept)
                             7.5
                                                            -3.0
##
##
                  factor(month)6
                                                 factor(month)9
##
                             -3.0
                                                            -1.5
##
                  prog.fac2:none factor(month)3:prog.fac2:none
##
## factor(month)6:prog.fac2:none factor(month)9:prog.fac2:none
##
                              3.0
                                                             1.5
##
     numDF F-value p-value
         1 759.3707 <.0001
```

The p-value is almost 0, so we reject the null hypothesis that the response profile is the same for two treatments.

Parametric curves

Quadratic time trend

```
## QUADRATIC time, CS, ML,
wtloss.cs.quad <- gls(wgtloss~month*prog.fac + I(month^2)*prog.fac,
correlation=corCompSymm(form= ~1 | id),
weights=varIdent(form= ~1),
data=wtloss.uni,
method="ML")
#summary(wtloss.cs.quad)

## CATEGORICAL time, CS, ML
wtloss.cs.cat.ml <- gls(wgtloss~factor(month)*prog.fac,
correlation=corCompSymm(form= ~1 | id),
weights=varIdent(form= ~1),
data=wtloss.uni,
method="ML")
#summary(wtloss.cs.cat.ml)</pre>
```

 H_0 : quadratic model H_1 : saturated mode

qm_vs_sm <- anova(wtloss.cs.cat.ml, wtloss.cs.quad)

Testing the Quadratic trend

	-2 Log	Number of
Structure	Likelihood	Parameters
Quadratic model	1935.7000029	8
Saturated model	1935.3313128	10
Difference	0.3686901	2

LRT yields $G^2 = 0.3686901$ with 2 df (p-value = 0.8316488), so we fail to reject the null hypothesis at $\alpha = 0.05$ and conclude that the model with Month as a quadratic effect seems to fit the data adequately. (Note: $\chi^2_{2,0.95} = 5.99$)

Linear time trend

```
wtloss.cs.lin <- gls(wgtloss~month*prog.fac,
correlation=corCompSymm(form= ~1 | id),
weights=varIdent(form= ~1),
data=wtloss.uni,
method="ML")
#summary(wtloss.cs.lin)</pre>
```

 H_0 : linear model H_1 : quadratic model

qm_vs_lm <- anova(wtloss.cs.quad, wtloss.cs.lin)

Linear vs. Quadratic		
-	-2 Log	Number of
Structure	Likelihood	Parameters
Linear model	1936.295409	6
Quadratic model	1935.7000029	8
Difference	0.5954061	2

LRT yields $G^2 = 0.5954061$ with 2 df (p-value = 0.7425218), so we fail to reject the null hypothesis at $\alpha = 0.05$ and conclude that the model with Month as a linear effect fits the data better than the model with Month as a quadratic effect. (Note: $chi_{2,0.95}^2 = 5.99$)

Testing for intersections

month

prog.fac

month:prog.fac

Since the p-value for month*program is < 0.0001, we reject the null hypothesis and conclude that there is an interaction between month and program.

Thus, our final model is the linear model with interaction.

1

What can you conclude about the two weight programs in terms of their effectiveness?

1 1011.995 <.0001

1 184.368 <.0001

24.301 < .0001

Interpreting quadratic trends:

 $\hat{\beta_1}=??$ expected mean response at baseline of subjects in program 1. $\hat{\beta_1}=??$ change in expected response at baseline of subjects in program 2 vs. program 1. Rate of change in program 1 is $\beta_2+2\beta_3time_{ij}$. Plugging in the above estimates,

	Progran	n 1
	Rate of	Expected
$_{ m time}$	Change	Response
0		
1		
2		
_3		

Thus, the mean response for Program 1 ?? (increases/decreases) over time. What about Program 2?