

# Assignment 1.

2020249.

## Section 1 theory.

Part a)

given  $\bar{X}$  &  $\bar{Y}$  are arithmetic mean of  $X$  &  $Y$  respectively.

$X \rightarrow$  independent variable.

$Y \rightarrow$  Dependent variable.

$$\bar{Y} = \omega_0 + \omega_1 \bar{X}.$$

With noise. (generalisation)

$$Y_i = \omega_0 + \omega_1 X_i + \epsilon$$

$n =$  number of samples in data

$$\sum_{i=1}^n Y_i = n(\omega_0 + \omega_1 X_i + \epsilon_i)$$

$$= n\omega_0 + \omega_1 \sum_{i=1}^n X_i + \underbrace{\left( \sum_{i=1}^n \epsilon_i \right)}_{\text{noise considered zero.}}$$

$$\frac{1}{n} \sum_{i=1}^n \frac{Y_i}{1} = \omega_0 + \frac{\omega_1}{n} \sum_{i=1}^n X_i$$

$$\bar{Y} = \omega_0 + \omega_1 \bar{X}$$

$$\sum_{i=1}^n \frac{Y_i}{n} = \text{arithmetic mean of } Y.$$

$$\frac{1}{n} \sum_{i=1}^n X_i = \text{arithmetic mean of } X.$$



b) No, (Not always)

example.

let  $X, Y, Z$  are Random variables.

$X$  &  $Y$  are positively correlated  $\Rightarrow$

$Y$  &  $Z$  are positively correlated.

~~$$C(X, Y) = C(Y, Z) * C(Z, X) - \sqrt{1 - C(Y, Z)^2} * \sqrt{1 - C(Z, X)^2}$$~~

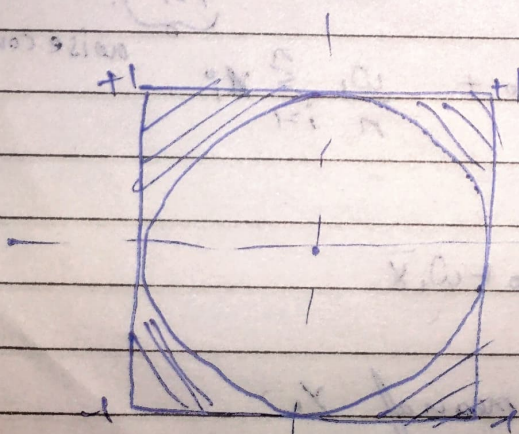
$$C(X, Y) = C(Y, Z) * C(Z, X) - \sqrt{1 - C(Y, Z)^2} * \sqrt{1 - C(Z, X)^2}$$

for  $C(X, Y)$  to be more than zero.

$$C(Y, Z) * C(Z, X) > \sqrt{1 - C(Y, Z)^2} * \sqrt{1 - C(Z, X)^2}$$

Squaring both sides.

$$C(Y, Z)^2 + C(Z, X)^2 > 1$$



if two correlations are in the circular non-shaded region we can't say precisely anything about the correlation of the 3rd pair.



## c) Proof of LLN (WLLN)

Let  $X_1, \dots, X_n$  are i.i.d. Random variables with mean  $\mu$  where  $n$  is very large.

$$S_n = \frac{X_1 + X_2 + \dots + X_n}{n}$$

$$\lim_{n \rightarrow \infty} P(|S_n - \mu| > \epsilon) = 0.$$

Also

$$E[S_n] = \mu \text{ and variance} = \frac{n\sigma^2}{n^2} = \frac{\sigma^2}{n}.$$

from Chebyshev's inequality

$$P(|S_n - \mu| > \epsilon) \leq \frac{\text{Var}[S_n]}{\epsilon^2} = \frac{\sigma^2}{n\epsilon^2}$$

$$\lim_{n \rightarrow \infty} P(|S_n - \mu| \geq \epsilon) = 0$$

Example.

Let's consider a fair die distribution.

$$S = \{1, 2, 3, 4, 5, 6\}.$$

$$S_n = 3.5.$$

When we roll the die for a very large number ( $n$ ) times.



The average value approaches 3.5.

Pseudo code

```
def FLNDie(n)
    result = []
    for i in Range(1, n+1)
        result.append(random.choice(1, 2, 3, 4, 5, 6))
    return result
```

```
def LLN(n)
    result result = LLPDie(n)
    average = []
    for i in Range(len(result)):
        average.append((np.cumsum(result)[i+1]) / (i+1))
    return average.
```



Q.4 MAP Solution for linear regression from Bayes th.

$$P(w|D) = \underbrace{P(D|w)}_{P(D)} \times P(w)$$

← unknown.

$$P(w|D) \propto P(D|w) \times P(w)$$

take log

$$\log(P(w|D)) = \log P(D|w) + \log(P(w)) \quad \text{--- (1)}$$

assume a gaussian distribution for  $w$  with mean 0 and variance  $\sigma^2$ ,  $N(0, \sigma^2)$

$$\log(P(w)) = \log \left( \frac{1}{\sqrt{2\pi\sigma^2}} e^{\left(\frac{-\sigma}{2} w^T w\right)} \right) = \log(p(w))$$

$$\log(P(D|w)) = \frac{1}{2\sigma^2} (y - Dw)^T (y - Dw)$$

$= \left( \frac{1}{2\sigma^2} w^T w \right)$

put  $P(D|w)$  &  $P(w)$  in eq (1) & differentiate w.r.t  $w$ .



$$\frac{d}{dw} \left[ \frac{1}{2\sigma^2} (y - Dw)^T (y - Dw) \right] + \frac{d}{dw} \left( \frac{1}{2\sigma^2} w^T w \right)$$

$$= \frac{1}{\sigma^2} (w^T D^T D - y^T D) + \frac{1}{\sigma^2} w^T$$

$$\frac{1}{\sigma^2} (w^T D^T D - y^T D) + \frac{1}{\sigma^2} w^T = 0$$

$$\frac{w^T}{\sigma^2} (D^T D + I) = \frac{y^T D}{\sigma^2}$$

$$w^T = (y^T D) (D^T D + I)^{-1}$$

$$w_{reg}^T = y^T D (D^T D + \lambda I)^{-1}$$

↑  
regularizer.