Education Choices in Mexico: Using a Structural Model and a Randomized Experiment to Evaluate PROGRESA

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Central Questions

- Analyze the impact of monetary incentives on education choices in rural Mexico.
- Discuss the design of effective policies.

Methodological Point

- Structural estimation methods can be combined with experimental data to give better answers.
- Experimental data provides more variation than observational data
- Structural methods help
 - lacktriangleright interpret experimental results (mechanisms) ightarrow external validity
 - General Equilibrium effects
 - Evaluation of counterfactuals/ policy experiments
- Todd and Wolpin (2006)

Outline

- PROGRESA and data
- Impact of the program
- Model and Estimation
- Results and Simulations
- Conclusion

Overview of PROGRESA

- Started in 1997.
- Key provision: 'Eligible' mothers were given grants to keep their children in school.
- A household is eligible for the grant if it is deemed 'poor'.
- 506 low-income rural localities in Mexico were selected.
- 186 of these were randomized out of the program forming the control group to receive program later.
- Grants increasing with grades; higher for girls than for boys in secondary.

Grant Structure

TABLE 1
The PROGRESA grants

PROGRESA bimonthly monetary benefits					
Type of benefit	1998 1st semester	1998 2nd semester	1999 1st semester	1999 2nd semester	
Nutrition support	190	200	230	250	
Primary school					
3	130	140	150	160	
4	150	160	180	190	
5	190	200	230	250	
6	260	270	300	330	
Secondary school					
First year					
Boys	380	400	440	480	
Girls	400	410	470	500	
Second year					
Boys	400	400	470	500	
Girls	440	470	520	560	
Third year					
Boys	420	440	490	530	
Girls	480	510	570	610	
Maximum support	1170	1250	1390	1500	

Data

- Baseline survey for all households in 1997.
- Program implemented first for school year 1998/99.
- Surveys conducted in October 1998, March 1999 and November 1999 for all households.
- This paper mainly uses data from October 1998 and also from baseline survey.
- Only look at boys.

Average impact: Treatment - Control

TABLE 2 Experimental results October 1998

Difference estimates of the impact of PROGRESA on boys school enrolment				
Age group	Enrolment rates in control villages (eligible)	Impact on Poor 97	Impact on Poor 97–98	Impact on. non-eligible
10	0.951	0-0047	0.0026	0.0213
		(0.013)	(0.011)	(0.021)
11	0.926	0-0287	0.0217	-0.0195
		(0.016)	(0.015)	(0.019)
12	0.826	0.0613	0.0572	0.0353
		(0.024)	(0.022)	(0.043)
13	0.780	0.0476	0.0447	0.0588
		(0.030)	(0.027)	(0.060)
14	0.584	0.1416	0.1330	0.0672
		(0.039)	(0.035)	(0.061)
15	0.455	0.0620	0.0484	0.1347
		(0.042)	(0.039)	(0.063)
16	0.292	0.0304	0.0355	0.1063
		(0.038)	(0.036)	(0.067)
12-15	0.629	0.0655	0.0720	0.0668
		(0.027)	(0.024)	(0.022)
10-16	0.708	0.0502	0.0456	0.0810
		(0.018)	(0.015)	(0.026)

Note: standard errors in parentheses are clustered at the locality level.

Average impact: Diff-in-diff

TABLE A1
Difference in difference estimates—August 1997 to October 1998

The impact of PROGRESA on boys school enrolment				
Age group	Impact on Poor 97	Impact on Poor 97–98	Impact on non-eligible	
10	0.0291	-0.0007	0.0723	
	(0.027)	(0.026)	(0.060)	
11	0.0240	0-0176	-0.0111	
	(0.016)	(0.014)	(0.019)	
12	0.0478	0.0420	0.0194	
	(0.021)	(0.020)	(0.043)	
13	0.0391	0.0396	0.0411	
	(0.028)	(0.025)	(0.049)	
14	0.0838	0.0731	-0.0460	
	(0.032)	(0.027)	(0.051)	
15	0.0963	0.0816	0.0617	
	(0.035)	(0.033)	(0.062)	
16	0.0350	0.0472	0.0517	
	(0.036)	(0.030)	(0.059)	
12–15	0.056	0.043	0.015	
	(0.019)	(0.012)	(0.023)	
10–16	0.048	0.053	0.006	
	(0.013)	(0.012)	(0.032)	

Note: Standard errors in parentheses are clustered at the locality level.

The ideal experiment

- Conditional grant can have differential benefit from other subsidies.
- The available variation only identifies effect of existence of grant.
- No variation in grant, conditional on class BUT variation in ages within class.
- Model (assume something) about behavior (conditional on age) in order to use this variation.

Overview

- Dynamic model
 - Opportunity cost of work
 - ► Habit formation/ State dependence
- Estimated using cross sectional data

Overview

- Child chooses to go to school or work → (random) instantaneous utility.
- If work, get a wage.
- Habits → initial condition will matter.
- Possibility of going to school until age 17 → terminal value.
- Time invariant heterogeneity between children.
- Uncertainty about graduating a grade.
- Anticipation effects no evidence.

Utility from schooling:

$$u_{it}^{s} = Y_{it}^{s} + \alpha g_{it}$$

$$Y_{it}^{s} = \mu_{i}^{s} + a^{s'} z_{it} + b^{s} \operatorname{ed}_{it} + \mathbb{I}(p_{it} = 1) \beta^{p} x_{it}^{p} + \mathbb{I}(s_{it} = 1) \beta^{s} x_{it}^{s} + \epsilon_{it}^{s}$$

- z_{it} : household taste shifter subsumes income
- mu^s_i: heterogeneity
- ed_{it}: 'habit' variable
- ϵ_{it}^s : extreme value, iid across i and t

Utility from work:

$$u_{it}^{w} = Y_{it}^{w} + \delta w_{it}$$

$$Y_{it}^{w} = \mu_{i}^{w} + a^{w'} z_{it} + b^{w} \operatorname{ed}_{it} + \epsilon_{it}^{w}$$

- δ : differential impact of g_{it}
- w_{it} : (potential) earnings

Summarizing,

$$u_{it}^{s} = \alpha g_{it} + \mu_{i}^{s} + a^{s'} z_{it} + b^{s} \operatorname{ed}_{it} + \mathbb{I}(p_{it} = 1) \beta^{p} x_{it}^{p} + \mathbb{I}(s_{it} = 1) \beta^{s} x_{it}^{s} + \epsilon_{it}^{s}$$

$$u_{it}^{w} = \delta w_{it} + \mu_{i}^{w} + a^{w'} z_{it} + b^{w} \operatorname{ed}_{it} + \epsilon_{it}^{w}$$

Summarizing,

$$u_{it}^{s} = \alpha g_{it} + \mu_{i}^{s} + a^{s\prime} z_{it} + b^{s} \operatorname{ed}_{it} + \mathbb{I}(p_{it} = 1) \beta^{p} x_{it}^{p} + \mathbb{I}(s_{it} = 1) \beta^{s} x_{it}^{s} + \epsilon_{it}^{s}$$

$$u_{it}^{w} = \delta w_{it} + \mu_{i}^{w} + a^{w\prime} z_{it} + b^{w} \operatorname{ed}_{it} + \epsilon_{it}^{w}$$

ullet Observe schooling choice \Longrightarrow only the difference between parameters is identified.

$$\begin{aligned} u_{it}^s &= \alpha g_{it} + \mu_i^s + a^{s\prime} z_{it} + b^s \mathrm{ed}_{it} + \mathbb{I}(p_{it} = 1) \beta^p x_{it}^p + \mathbb{I}(s_{it} = 1) \beta^s x_{it}^s + \epsilon_{it}^s \\ u_{it}^w &= \delta w_{it} + \mu_i^w + a^{w\prime} z_{it} + b^w \mathrm{ed}_{it} + \epsilon_{it}^w \end{aligned}$$

Therefore, WLOG

$$u_{it}^{s} = \gamma \delta g_{it} + \mu_{i} + a' z_{it} + b \operatorname{ed}_{it} + \mathbb{I}(p_{it} = 1) \beta^{p} x_{it}^{p} + \mathbb{I}(s_{it} = 1) \beta^{s} x_{it}^{s} + \epsilon_{it}$$

$$u_{it}^{w} = \delta w_{it}$$

$$u_{it}^s = \gamma \delta g_{it} + \mu_i + a' z_{it} + b \operatorname{ed}_{it} + \mathbb{I}(p_{it} = 1) \beta^p x_{it}^p + \mathbb{I}(s_{it} = 1) \beta^s x_{it}^s + \epsilon_{it}$$

$$u_{it}^w = \delta w_{it}$$

- parameters are differences
- ullet ϵ_{it} : logistic distribution (difference between two extreme value)
- \bullet γ : impact of grant proportional to impact of wage
- μ_i : discrete random variable, estimated

$$u_{it}^s = \gamma \delta g_{it} + \mu_i + a' z_{it} + b \operatorname{ed}_{it} + \mathbb{I}(p_{it} = 1) \beta^p x_{it}^p + \mathbb{I}(s_{it} = 1) \beta^s x_{it}^s + \epsilon_{it}$$

$$u_{it}^w = \delta w_{it}$$

- parameters are differences
- \bullet ϵ_{it} : logistic distribution (difference between two extreme value)
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- μ_i : discrete random variable, estimated
- Add Dummy for eligibility

$$u_{it}^s = \gamma \delta g_{it} + \mu_i + a' z_{it} + b \operatorname{ed}_{it} + \mathbb{I}(p_{it} = 1) \beta^p x_{it}^p + \mathbb{I}(s_{it} = 1) \beta^s x_{it}^s + \epsilon_{it}$$
$$u_{it}^w = \delta w_{it}$$

- parameters are differences
- \bullet ϵ_{it} : logistic distribution (difference between two extreme value)
- ullet γ : impact of grant proportional to impact of wage
- μ_i : discrete random variable, estimated
- Add Dummy for eligibility
- Add Dummy for treatment+ineligible

Aside: Income pooling

Consider static model

$$U^{s} = \beta^{s} Y + \theta^{s} g$$

$$U^{w} = \beta^{w} Y + \theta^{w} w - \alpha$$

- $\beta^{s} = \theta^{s}$ and $\beta^{w} = \theta^{w}$: Income pooling
- Alternative estimation strategies:
 - Assume $\beta^s = \theta^s$ and $\beta^w = \theta^w$ no variation in g
 - Assume $\theta^s = \theta^w$ no variation in g
 - Use variation in g

Terminal Value function

- At age 18, could measure returns to education but limited data.
- So,

$$V(\mathsf{ed}_{i,18}) = \frac{\alpha_1}{1 + \mathsf{exp}(-\alpha_2 \mathsf{ed}_{i,18})}$$

where $\alpha_1, \alpha_2 > 0$.

 Assumption: Only education determines returns, everything else affects education.

Value functions

Value from schooling (V^s) and work (V^w)

$$\begin{split} V_{it}^s(\mathsf{ed}_{it}|\Gamma_{it}) &= u_{it}^s + \beta \left\{ p_t^s(\mathsf{ed}_{it}) E \max \left[V_{it+1}^s(\mathsf{ed}_{it}+1), V_{it+1}^w(\mathsf{ed}_{it}+1) \right] \right. \\ &\left. + (1 - p_t^s(\mathsf{ed}_{it})) E \max \left[V_{it+1}^s(\mathsf{ed}_{it}), V_{it+1}^w(\mathsf{ed}_{it}) \right] \right\} \\ V_{it}^w(\mathsf{ed}_{it}|\Gamma_{it}) &= u_{it}^w + \beta E \max \left[V_{it+1}^s(\mathsf{ed}_{it}), V_{it+1}^w(\mathsf{ed}_{it}) \right] \end{split}$$

- $p_t^s(ed_{it})$ calculated from data
- Given this, E over ϵ_{it}

Wages

- Need to estimate a child wage equation
 - wage unobserved for school-goers
 - predict future wage for dynamic programming problem
 - GE effects (local)
- Estimate an equation of the following form

$$\mathsf{In}\ \mathsf{w}_{ij} = \mathsf{q}_j + \mathsf{a}_1 \mathsf{age}_i + \mathsf{a}_2 \mathsf{educ}_i + \omega_{ij}$$

Will use point estimates for prediction

Wages I - endogeneity

Education as a regressor

- Selection issue: more able children, higher wage
 - Tobit selection model
- Education choice is a function of wage
 - almost flat relationship between education and wages within village
 - ► Child wage + returns to education mostly through migration

Wages II - Modeling q_j

• Consider the village economy

[labor supply]:
$$H_k = L_k w_k^{\gamma k}$$
 [production]: $Q = \left[\delta H_{child}^{\sigma} + (1 - \delta) H_{adult}^{\sigma}\right]^{\frac{1}{\sigma}}$ $\sigma = \frac{\rho - 1}{\rho}, \rho > 0$

Wages II - Modeling q_j

In eqilibrium,

$$\ln w_{\textit{child}} = \frac{\rho + \gamma_{\textit{adult}}}{\rho + \gamma_{\textit{child}}} \ln w_{\textit{adult}} - \left[\frac{1}{\rho + \gamma_{\textit{child}}} \ln \left(\frac{L_{\textit{child}}}{L_{\textit{adult}}} \right) + \kappa \right]$$

- PROGRESA $\implies \left(\frac{L_{child}}{L_{adult}}\right)$ goes down.
- Identification: assume all parameters are constant across localities, conditional on treatment.
- w_{adult} is a sufficient statistic for w_{child} .
- Also, w_{adult} is excluded from the selection equation.

Estimated Wage Equation

$$\begin{split} \ln w_{ij} &= -0.983 + 0.0605 P_j + 0.883 \ln w_j^{ag} + 0.066 \text{age}_i \\ &\quad + 0.0116 \text{educ}_i - 0.056 \text{Mills}_i + \bar{\omega}_{ijt} \end{split}$$

- Mills insignificant
- Small coefficient on educ
- At individual level: age
- At village level: adult wage ($\ln w_j^{ag}$) and P_j



Initial Conditions problem

- The issue: ed_{it} term in utility function, but only observe cross section.
- Specify a reduced form

$$\operatorname{ed}_{it} = h'_i \lambda + \xi \mu_i$$

- μ_i captures the initial conditions (endogeneity) problem.
- h_i instrument distance to school changes between the two time periods.
- Ordered probit with age dependent thresholds.

Estimation

- Estimate 3 specifications:
 - Without dummy for treatment + ineligible
 - With dummy
 - Only control group without variation in grants
- Wage equation estimated separately.
- MLE for the main model

Estimation

To obtain the log-likelihood function:

- Guess parameters and use data to
- Solve the dynamic programming problem for full state space
 - Initial conditions equation enters value function
- Calculate log-likelihood
 - Value functions enter as latent variables
 - ▶ 3 point support for μ_i distribution
 - Logistic distribution allows for parametrization

Maximize likelihood iteratively.

Results - μ_i

TABLE 3
The distribution of unobserved heterogeneity

	A	В	C
Point of Support 1	-9.706	-8.327	-4.290
	1.041	1.101	2.46
Point of Support 2	-14.466	-13.287	-17.62
	1.173	1.208	<i>3.144</i>
Point of Support 3	-5.933	-4.301	-0.267
••	0.850	0.941	2.45
Probability of 1	0.513	0.518	0.490
•	0.024	0.023	0.032
Probability of 2	0.342	0.335	0.270
•	0.022	0.021	0.017
Probability of 3	0.145	0.147	0.240
Load factor for initial condition	0.108	0.102	0.068
	0.016	0.014	0.013

Notes: Column A: eligible dummy only; B: eligible dummy and non-eligible in treatment village dummy. C: model estimated on control sample only. Asymptotic standard errors in italics.

TABLE 4
Equation for initial conditions

	Α	В	C
Poor	-0.275	-0.243	-0.280
	0.030	0.046	0.051
Ineligible individual in a PROGRESA village	_	0.057	_
	_	0.055	_
Father's education			
Primary	0.180	0.181	0.218
•	0.025	0.025	0-04262
Secondary	0.262	0.264	0.281
,	0.030	0.030	0.05302
Preparatoria	0.559	0.558	0.499
	0.0160	0.057	0.09107
Mother's education			
Primary	0.159	0.158	0.231
,	0.026	0.026	0.04446
Secondary	0.316	0.314	0.398
,	0.030	0.030	0.05139
Preparatoria	0.301	0.301	0.334
	0.061	0.061	0.09740
Indigenous	-0.005	0.006	0.133
	0.036	0.026	0.0461
Availability of Primary 1997	0.373	0.372	0.691
,,	0.073	0.073	0.19003
Availability of Secondary 1997	0.808	0.804	-0.568
, ,	0.188	0.188	0.349
Kilometer to closest secondary school 97	0.00004	0.00004	-0.0002
	0.00024	0.00003	0.00007
Availability of Primary 1998	-0.261	-0.264	-0.449
,,	0.127	0.126	0.235
Availability of Secondary 1998	-0.845	-0.841	0.516
	0.187	0.187	0.348
Kilometer to closest secondary school 98	-0.0001	-0.0001	0.00015
your	0.00003	0.00003	0-00007
Cost of attending secondary	0.00006	0.0001	-0.00019
	0.00024	0.00024	0.00037

Notes: As in Table 3. State dummies included. Availability means school in the village.

TABLE 5
Parameter estimates for the education choice model

	A	В	С
Wage	0.134	0-168	0.357
	0.043	0.045	0.100
PROGRESA grant	3.334	2-794	_
-	1.124	0.796	_
Parameter in terminal function $ln(a_1)$	5-876	5-886	6-59
	0.115	0.113	0.175
Parameter in terminal function $ln(a_2)$	-1.276	-1-286	-1.62
1 27	0.025	0-024	0.089
Poor	0.676	0-105	0.431
	0.154	0.215	0.274
Ineligible individual in a PROGRESA village		-1-079	
•		0.261	
Father's Education - Default is less than primary			
Primary	-0.462	-0-509	-0.486
	0.120	0.123	0.217
Secondary	-0.746	-0.803	-0.959
*	0.147	0.150	0.261
Preparatoria	-1.794	-1-819	-2.176
	0.323	0-328	0-558
Mother's Education - Default is less than primary			
Primary	-0.488	-0.488	-0.870
,	0.123	0.126	0.233
Secondary	-0.624	-0-613	-1.119
	0.143	0.145	0.254
Preparatoria	-1.576	-1-681	-2.158
	0.351	0.355	0.645
Indigenous	-0.783	-0.777	-1.018
	0.132	0-135	0.241
Availability of Primary 1998	3-600	3-765	3.092
	0.285	0.295	0.499
Availability of Secondary 1998	-0.030	-0.074	0.789
* *	0.193	0-197	0.425
Kilometer to closest secondary school 98	0.0003	0-0003	0.00078
	0.00005	0.00005	0.00014
Cost of attending secondary	0.007	0.007	0.013
• •	0.001	0-001	0.0033
Age	2.291	2-249	2.903
=	0.160	0-157	0.354
Prior years of education	-2.785	-2-896	-3.621
-	0.256	0.261	0.621
Discount rate	0.95	0.96	0.975
Log-Likelihood	-23.403.98	-23.395-31	-8862-34

18-66

0.0

16-14

0.0

State dummies included *Notes* as in Table 3. LR test: equal effect of grant and wage χ_1^2 p-value

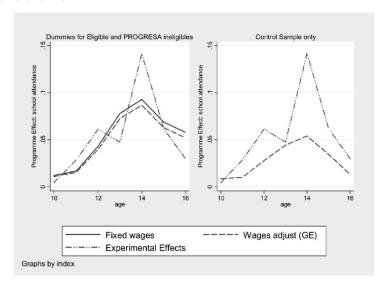
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Simulations - I

- Two scenarios
 - grant (actual)
 - no grant (counterfactual)
 - ★ GE wage
 - ★ unchanged wage
- For full sample and only for control

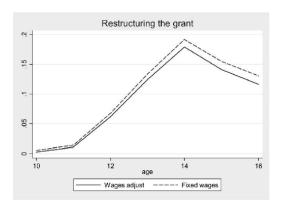
Simulations - I



- Age pattern, GE effect
- Only control sample

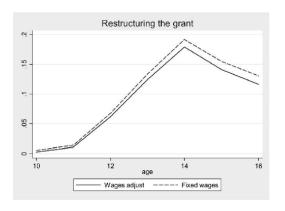


Simulations - II



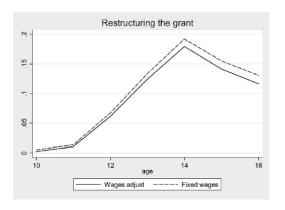
Experiment: Revenue neutral restructuring of grant profile

Simulations - II



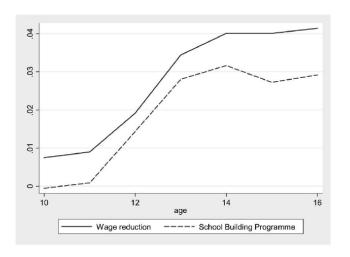
Impact: Almost double of original scheme

Simulations - II



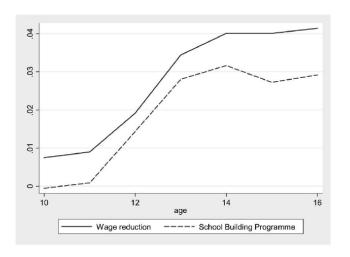
But: Not appropriate if credit constrained or early age health investments

Simulations - III



Experiment: Reduce wage by average grant size

Simulations - III



Experiment: Limit max school distance to 3 km



Conclusion

- Study the effect of PROGRESA on school choice.
- Structural modeling helps
 - Grants vs wages
 - General equilibrium
 - Anticipation effect
 - Counterfactuals
- Usefulness of using full variation in data.