

# Simple Linear Regression

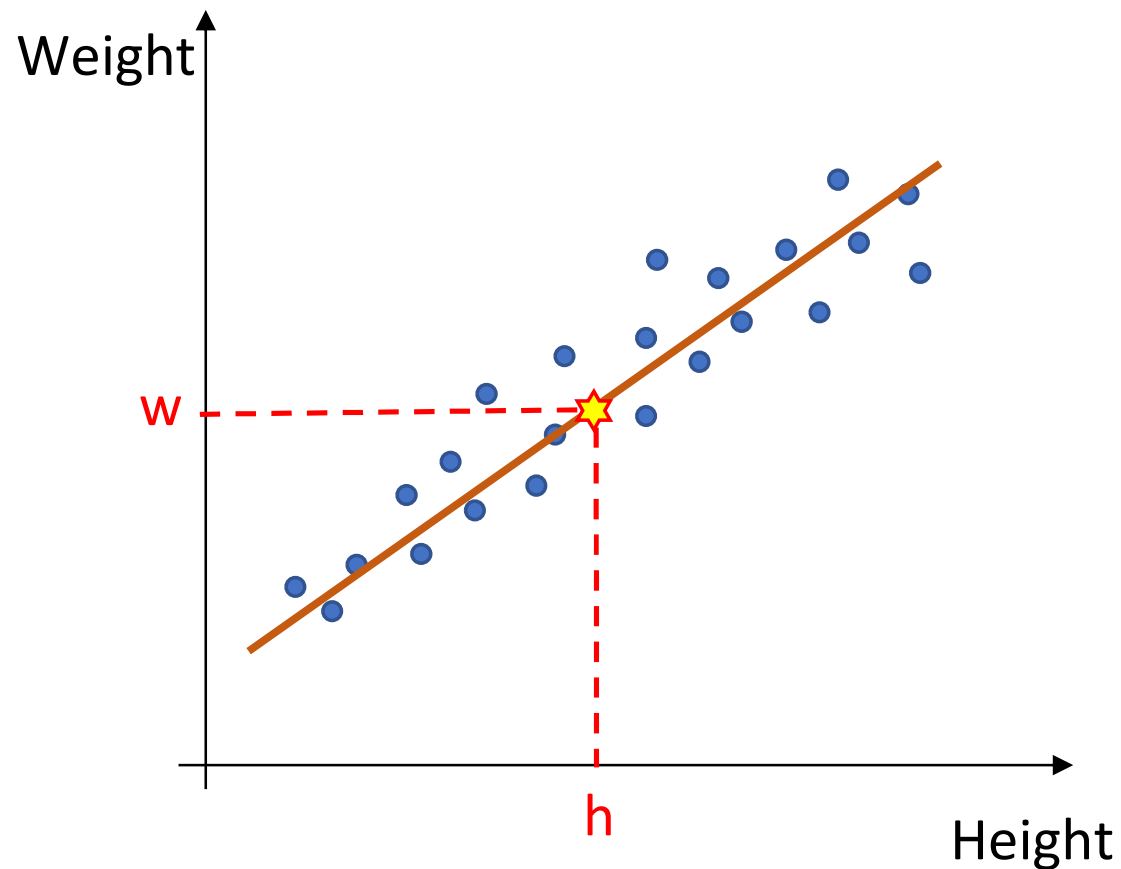
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# OUTLINE

- Simple Linear Regression: Intuition
- Hypothesis function for Simple Linear Regression
- Mean Square Error Loss / Cost function
- Intuition of Cost Function
- Ordinary Least Square Regression

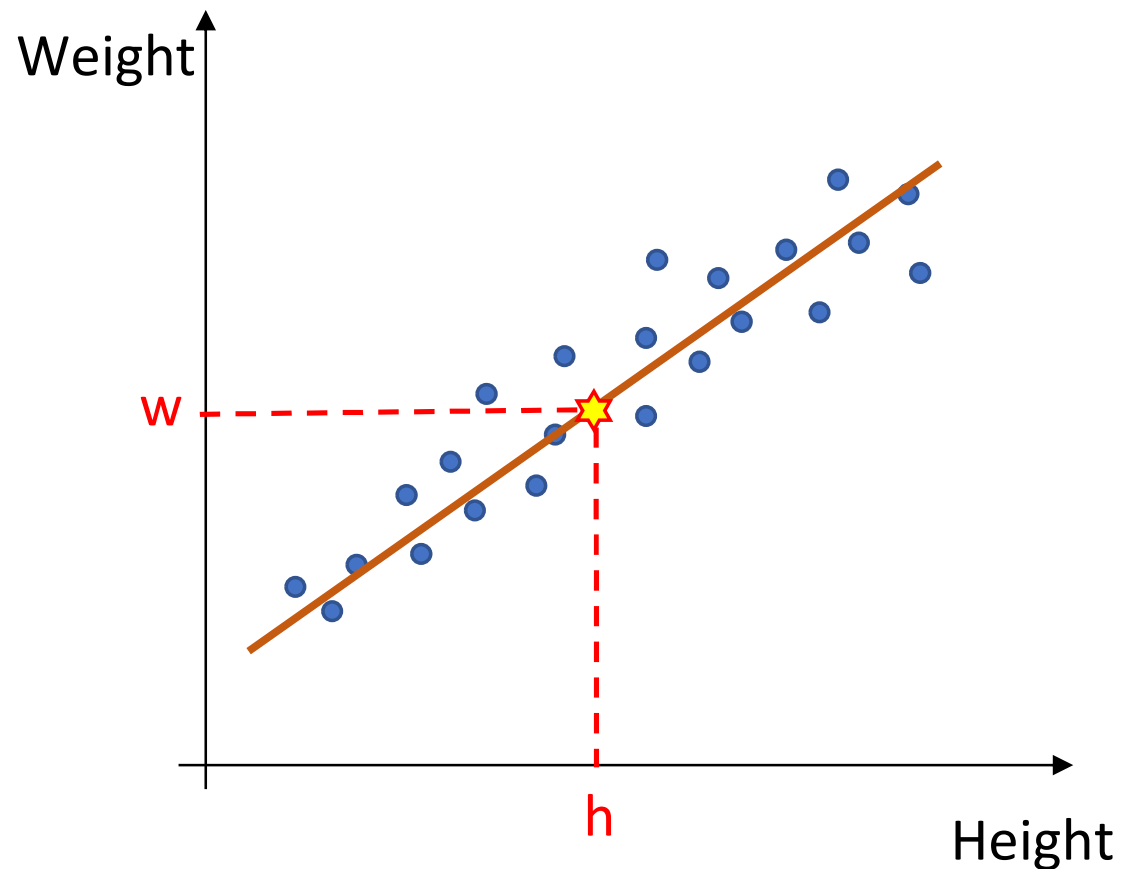
# Simple Linear Regression : Intuition



- Consider the scatter plot of the Weight vs. Height of adults as shown beside.
- The trend or the form of the relationship is strongly positive.
- Now suppose we wish to estimate the **weight** of a person just by knowing his/ her **height**.
- In order to do so we first fit a straight line through our data points.

- Then from the graph, knowing the height we can find the weight of the corresponding person.
- Hence, we are intending to find out the **equation of the straight line** that *best* describes the relationship between Weight and Height.

# Simple Linear Regression : Intuition



- There is only one predictor/input variable (Height) and one target variable (Weight) and we are intending to find out a relationship of the form:

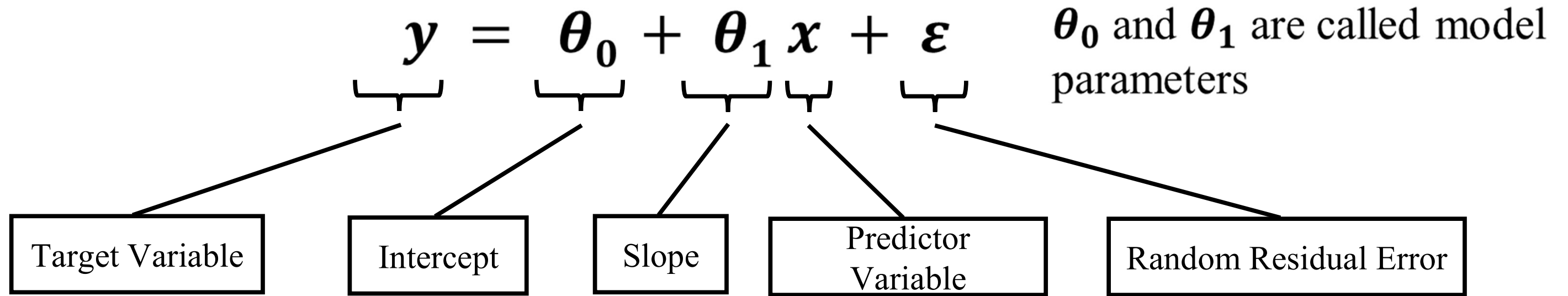
$$y = \theta_0 + \theta_1 x$$

*Here,  $y$  is the target variable and  $x$  is the predictor variable*

- We have to find out  $\theta_0$  and  $\theta_1$ , such that the straight line  $y = \theta_0 + \theta_1 x$  fits into our dataset **best**.
- This is called Simple Linear Regression, because it has only one predictor variable and the relationship among target and predictor variable is linear.

# Simple Linear Regression : Hypothesis

## Simple Linear Regression Model with Single Predictor



- We use our sample data to find estimates for the coefficients/ model parameters  $\theta_0$  and  $\theta_1$  i.e.:  $\widehat{\theta}_0$  and  $\widehat{\theta}_1$ .
- We can then **predict** what the value of  $y$  should be corresponding to a particular value for  $x$  by using the Least Squares Prediction Equation (also known as our **hypothesis function**):

$$\hat{y} = \widehat{\theta}_0 + \widehat{\theta}_1 x \quad \text{Where } \hat{y} \text{ is our prediction for } y$$

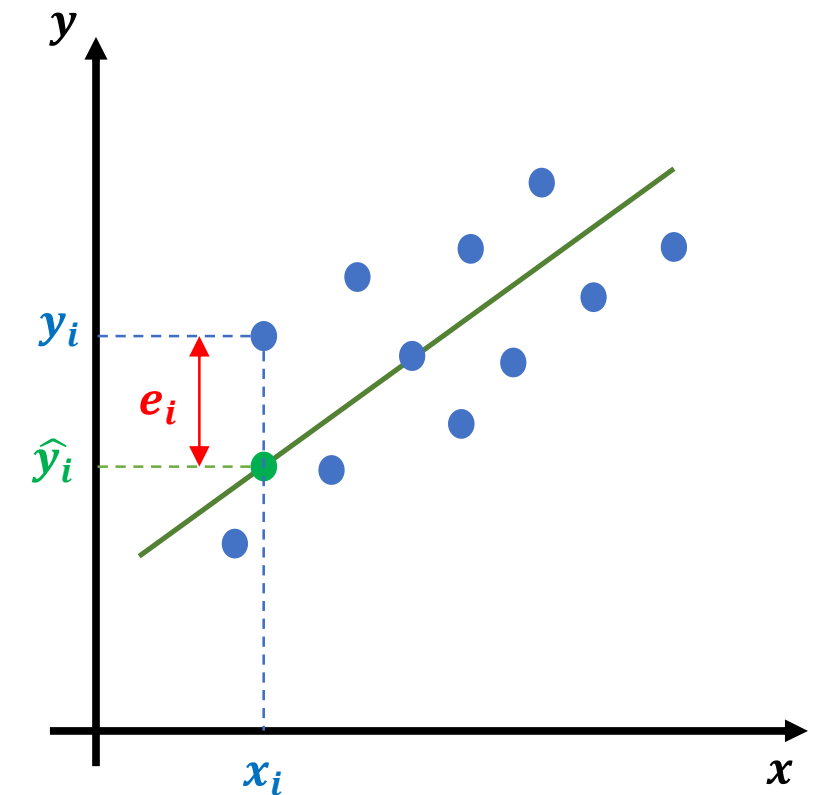
# Simple Linear Regression : Cost Function

## Residuals and Residual Sum of Squares:

- For  $i^{th}$  sample  $\langle \mathbf{x}_i, \mathbf{y}_i \rangle$  the predicted value of  $\mathbf{y}_i$  is  $\hat{\mathbf{y}}_i$ , Which we obtain from the equation  $\hat{\mathbf{y}}_i = \hat{\boldsymbol{\theta}}_0 + \hat{\boldsymbol{\theta}}_1 \mathbf{x}_i$
- Then,  $\mathbf{e}_i = \mathbf{y}_i - \hat{\mathbf{y}}_i$  (actual – predicted) represents the  $i^{th}$  residual.
- We define **Residual Sum of Squares** (RSS) as:

$$RSS = \sum_{i=1}^m \mathbf{e}_i^2 = \sum_{i=1}^m (\mathbf{y}_i - \hat{\mathbf{y}}_i)^2 = \sum_{i=1}^m (\mathbf{y}_i - (\hat{\boldsymbol{\theta}}_0 + \hat{\boldsymbol{\theta}}_1 \mathbf{x}_i))^2$$

There are total  $m$  no. of samples



# Simple Linear Regression : Cost Function

## Mean Square Error Cost Function:

- We can define the cost function as:

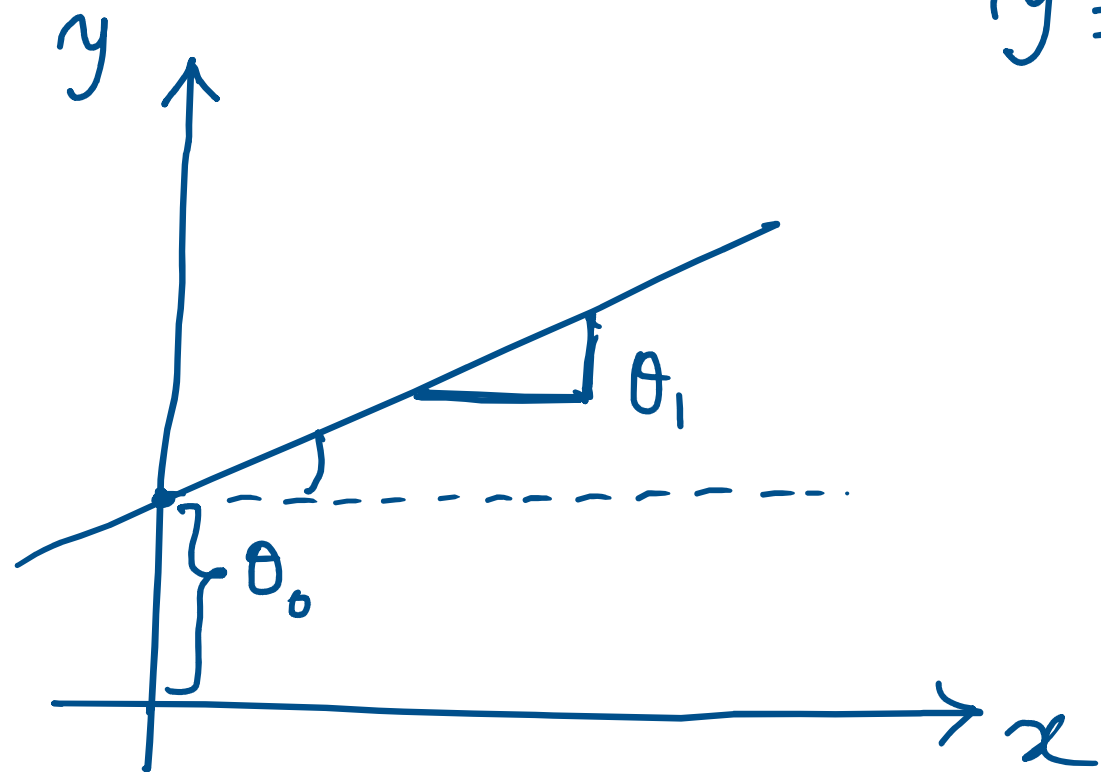
$$J(\widehat{\theta}_0, \widehat{\theta}_1) = \frac{1}{2} \frac{RSS}{\text{Number of training samples}} = \frac{1}{2m} \sum_{i=1}^m (y_i - (\widehat{\theta}_0 + \widehat{\theta}_1 x_i))^2$$

Here a factor  $\frac{1}{2}$  is multiplied just for computational simplicity. Otherwise, the cost function  $J(\widehat{\theta}_0, \widehat{\theta}_1)$  is nothing but **mean or average of the Residual sum of squares**. (also known as **Mean Square Error (MSE)**).

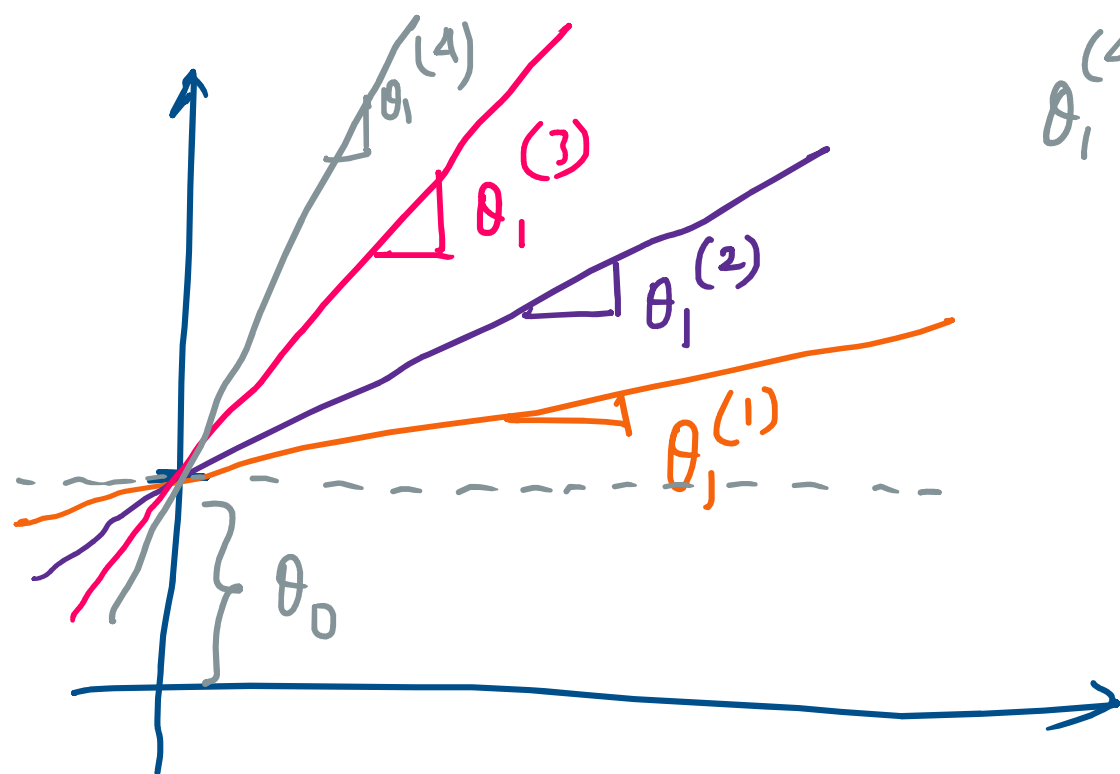
## Our Objective:

- To find the suitable values of  $\widehat{\theta}_0$  and  $\widehat{\theta}_1$  such that the cost function  $J(\widehat{\theta}_0, \widehat{\theta}_1)$  is minimized, in other words the Residual Sum of Square (**RSS**) is minimized. Then the straight line  $\widehat{y} = \widehat{\theta}_0 + \widehat{\theta}_1 x$  will fit our data **best**. This is called **least squares fit**.

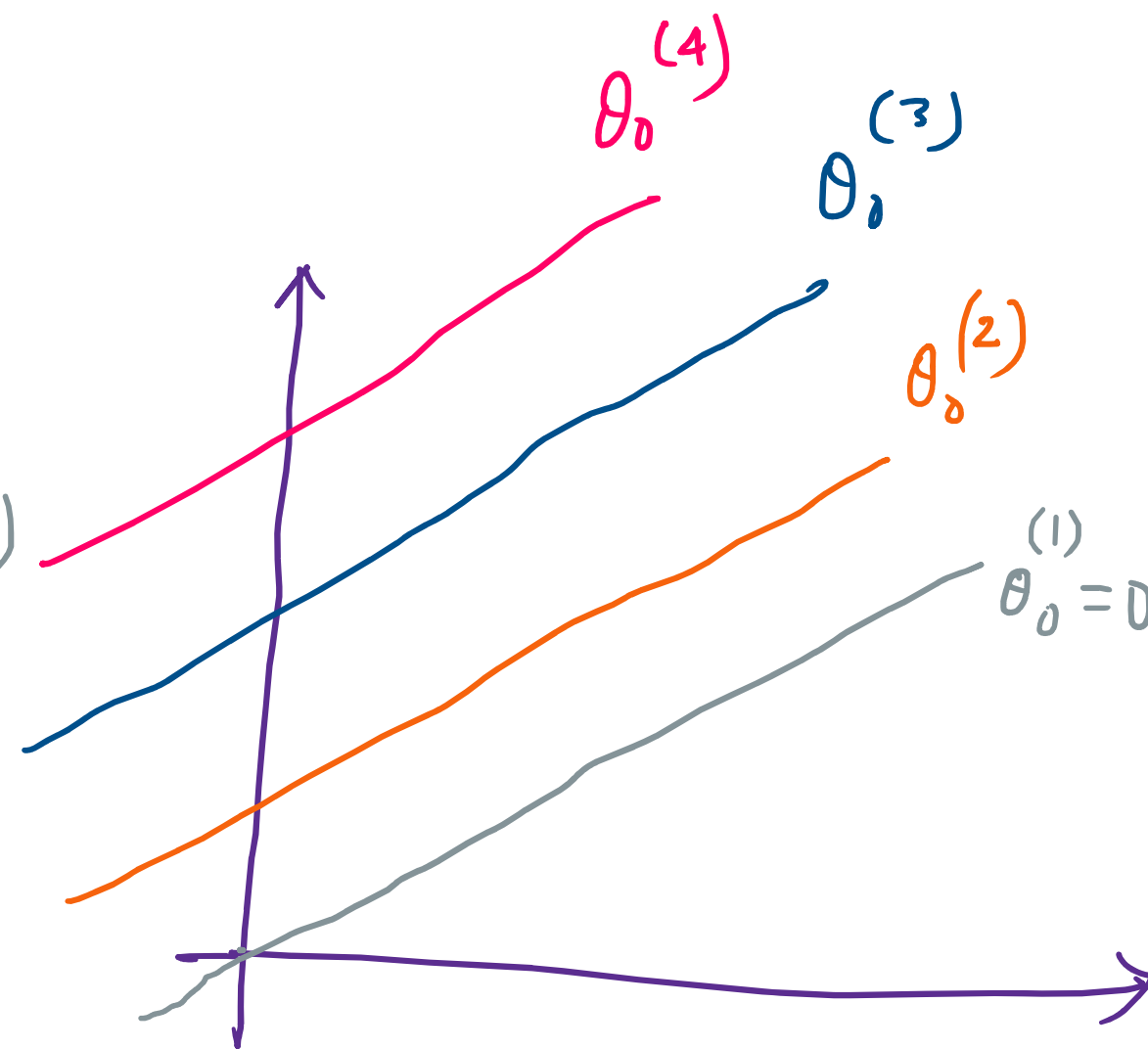




$$y = \underbrace{\theta_0}_{\text{intercept}} + \underbrace{\theta_1 x}_{\text{slope}}$$



$$\theta_1^{(4)} > \theta_1^{(3)} > \theta_1^{(2)} > \theta_1^{(1)}$$



$$\theta_0^{(4)} > \theta_0^{(3)} > \theta_0^{(2)} > \theta_0^{(1)}$$

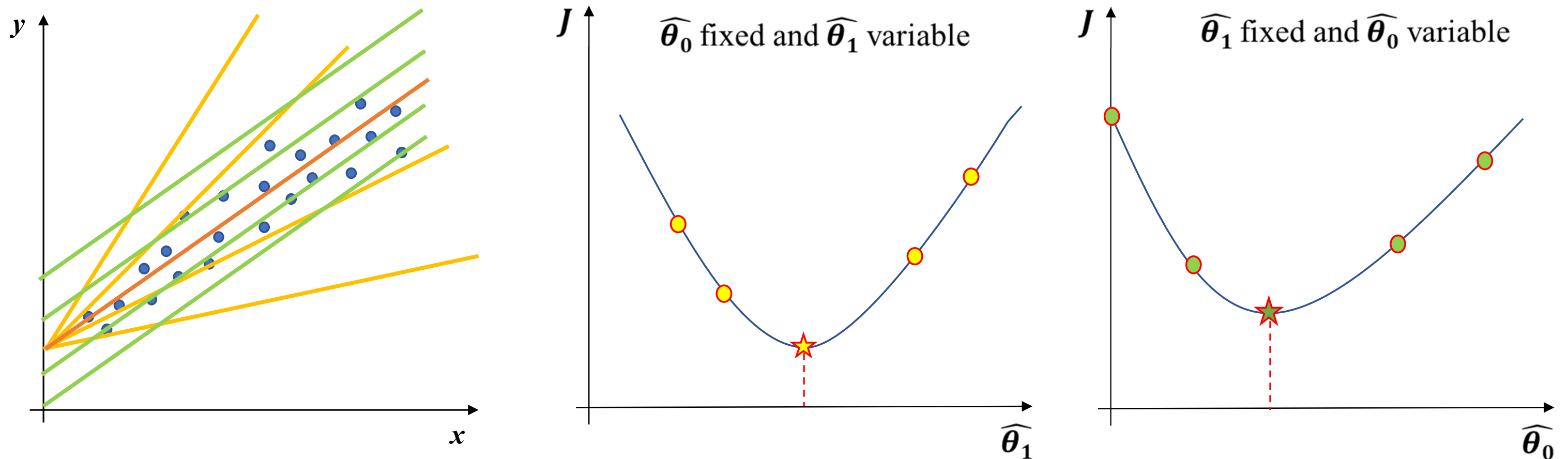


# Simple Linear Regression : Cost Function

## Intuition of Cost Function:

Consider the example of single predictor variable where the hypothesis function is  $\hat{y} = \hat{\theta}_0 + \hat{\theta}_1 x$  and the cost function is  $J(\hat{\theta}_0, \hat{\theta}_1) = \frac{1}{2m} \sum_{i=1}^m (y_i - (\hat{\theta}_0 + \hat{\theta}_1 x_i))^2$ .

Now we keep one parameter fixed and vary other. Let's see how  $J(\hat{\theta}_0, \hat{\theta}_1)$  varies.



- Our objective is to find the values of the parameters for which the cost function is minimized.

# Simple Linear Regression : OLS fit

**Solving for the best fit: Ordinary Least Squares (OLS) Regression:**

- We have to Minimize  $RSS$  or  $J(\widehat{\theta}_0, \widehat{\theta}_1)$  with respect to  $\widehat{\theta}_0$  and  $\widehat{\theta}_1$
- Hence we have to do,  $\frac{\partial}{\partial \widehat{\theta}_0} (RSS) = 0$  and  $\frac{\partial}{\partial \widehat{\theta}_1} (RSS) = 0$
- By solving the above two equations we get the following value of  $\widehat{\theta}_1$  and  $\widehat{\theta}_0$  :

$$\widehat{\theta}_1 = \frac{\sum_{i=1}^m (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^m (x_i - \bar{x})^2} = r_{xy} \frac{\sigma_y}{\sigma_x} \quad \text{and} \quad \widehat{\theta}_0 = \bar{y} - \widehat{\theta}_1 \bar{x}$$

*where,  $\bar{x}$  is the mean of predictor variable  $x$  and  $\bar{y}$  is the mean of target variable  $y$*

*$\sigma_x$  is the standard deviation of  $x$  and  $\sigma_y$  is the standard deviation of  $y$*

*and  $r_{xy}$  is the **correlation coefficient** between  $x$  and  $y$ .*

Calculation of Optimum values of model Parameters  $\hat{\theta}_0$  &  $\hat{\theta}_1$  :-

$$J = \frac{1}{2m} \sum_{i=1}^m (y^{(i)} - \hat{y}^{(i)})^2 = \frac{1}{2m} \sum_{i=1}^m (\hat{y}^{(i)} - y^{(i)})^2 \quad | \quad \hat{y}^{(i)} = \hat{\theta}_0 + \hat{\theta}_1 x^{(i)}$$

$$\therefore J = \frac{1}{2m} \sum_{i=1}^m (\hat{\theta}_0 + \hat{\theta}_1 x^{(i)} - y^{(i)})^2$$

$$\begin{aligned} \frac{\partial J}{\partial \hat{\theta}_0} &= \frac{1}{2m} \sum_{i=1}^m (\hat{\theta}_0 + \hat{\theta}_1 \bar{x} - y^{(i)}) \\ &= \frac{1}{m} \sum_{i=1}^m \hat{\theta}_0 + \hat{\theta}_1 \left[ \frac{1}{m} \sum_{i=1}^m x^{(i)} \right] - \left[ \frac{1}{m} \sum_{i=1}^m y^{(i)} \right] \end{aligned}$$

$$\frac{\partial J}{\partial \hat{\theta}_0} = \hat{\theta}_0 + \hat{\theta}_1 \bar{x} - \bar{y} \quad \left| \quad \frac{\partial J}{\partial \hat{\theta}_0} = 0 \Rightarrow \hat{\theta}_0 + \hat{\theta}_1 \bar{x} - \bar{y} = 0 \right. \\ \Rightarrow \boxed{\hat{\theta}_0 = \bar{y} - \hat{\theta}_1 \bar{x}}$$

$$J = \frac{1}{2m} \sum_{i=1}^m \left( \bar{y} - \hat{\theta}_0 - \hat{\theta}_1 x^{(i)} - y^{(i)} \right)^2$$

$$= \frac{1}{2m} \sum_{i=1}^m \left( \bar{y} + (x^{(i)} - \bar{x}) \hat{\theta}_1 - y^{(i)} \right)^2$$

$$\therefore \frac{\partial J}{\partial \hat{\theta}_1} = \frac{1}{m} \sum_{i=1}^m \left( \bar{y} + (x^{(i)} - \bar{x}) \hat{\theta}_1 - y^{(i)} \right) (x^{(i)} - \bar{x}) \quad \text{Var}(x)$$

$$= \frac{1}{m} \sum_{i=1}^m \left[ (x^{(i)} - \bar{x}) \hat{\theta}_1 - (y^{(i)} - \bar{y}) \right] (x^{(i)} - \bar{x}) = \frac{1}{m} \sum_{i=1}^m (x^{(i)} - \bar{x})^2 \hat{\theta}_1 - \frac{1}{m} \sum_{i=1}^m (x^{(i)} - \bar{x})(y^{(i)} - \bar{y})$$

$$= \text{Var}(x) \cdot \hat{\theta}_1 - \frac{1}{m} \sum_{i=1}^m (x^{(i)} - \bar{x})(y^{(i)} - \bar{y}) \quad \text{Cov}(x, y)$$

$$J = \frac{1}{2m} \sum_{i=1}^m \left( \hat{\theta}_0 + \hat{\theta}_1 x^{(i)} - y^{(i)} \right)^2$$

$$\hat{\theta}_0 = \bar{y} - \hat{\theta}_1 \bar{x}$$

$$\frac{\partial J}{\partial \hat{\theta}_1} = \text{Var}(x) \cdot \hat{\theta}_1 - \text{Cov}(x, y) = 0$$

$$\Rightarrow \hat{\theta}_1 = \frac{\text{Cov}(x, y)}{\text{Var}(x)}$$

$$\left| \begin{array}{l} r_{xy} = \frac{\text{Cov}(x, y)}{\sigma_x \cdot \sigma_y} = \text{Correlation coefficient between } x \text{ \& } y \end{array} \right.$$

$$\Rightarrow \text{Cov}(x, y) = r_{xy} \cdot \sigma_x \cdot \sigma_y$$

$$\text{Var}(x) = \sigma_x^2$$

$$\hat{\theta}_1 = \frac{r_{xy} \cdot \sigma_x \cdot \sigma_y}{\sigma_x^2}$$

$$\boxed{\hat{\theta}_1 = r_{xy} \cdot \frac{\sigma_y}{\sigma_x}}$$

&

$$\boxed{\hat{\theta}_0 = \bar{y} - \hat{\theta}_1 \bar{x}}$$

***Thank You***