Teaching the Tax Code: Earnings Responses to an Experiment with EITC Recipients[†]

By Raj Chetty and Emmanuel Saez*

We conducted a randomized experiment with 43,000 EITC recipients at H&R Block. Tax preparers gave simple, personalized information about the EITC schedule to half of their clients. We find no significant effects of information provision on earnings in the subsequent year in the full sample. Further exploration uncovers evidence of heterogeneous treatment effects on both self-employment income and wage earnings across the 1,461 tax preparers involved in the experiment. Providing information about tax incentives does not systematically effect earnings on average. However, tax preparers may influence their clients' earnings decisions by providing advice about how to respond to tax incentives. (JEL H23, H24, H26, J23, J31)

Agrowing body of evidence suggests that individuals are not fully informed about the tax and transfer policies relevant for economic choices (e.g., de Bartolome 1995; Duflo et al. 2006; Chetty, Looney, and Kroft 2009; Bettinger et al. 2009; Jones 2010; Liebman and Luttmer 2011). One natural hypothesis, in light of this evidence, is that policies that provide information about incentives would enable individuals to make better choices. In this paper, we test whether teaching individuals about the tax code affects labor supply choices using a randomized field experiment with Earned Income Tax Credit (EITC) clients at H&R Block. The EITC is the largest cash transfer program for low-income families in the United States, and it generates large marginal subsidies or taxes on the earnings of recipients (Figure 1). Survey evidence shows that the marginal incentive structure of the EITC is not well understood by eligible tax filers. Most low-income families have heard about the EITC and know that working is associated with getting a tax refund check when they file their taxes. But very few recipients know whether working more would increase or reduce their EITC amount (Liebman 1998, Romich and Weisner 2002),

*Chetty: Department of Economics, Harvard University, Littauer Center, 1805 Cambridge Street, Cambridge, MA 02138, and National Bureau of Economic Research (e-mail: chetty@fas.harvard.edu); Saez: Department of Economics, University of California, Berkeley, 530 Evans Hall #3880, Berkeley, CA 94720, and National Bureau of Economic Research (e-mail: saez@econ.berkeley.edu). We are extremely grateful to Joe Cresta, David Hussong, Mike Lammers, Scott McBride, Eileen McCarthy, Robert Weinberger, Jeremy White, Bernie Wilson, and the nearly 1,500 tax professionals at H&R Block for their help in organizing and implementing the experiment. We thank Michael Anderson, David Card, Stefano DellaVigna, Martin Feldstein, Bryan Graham, Caroline Hoxby, Hilary Hoynes, Lawrence Katz, David Laibson, Adam Looney, Erzo Luttmer, Marco Manacorda, Sendhil Mullainathan, Steve Pischke, Karl Scholz, three anonymous referees, and numerous seminar participants for very helpful comments and discussions. Gregory Bruich and Philippe Wingender provided outstanding research assistance. Financial support from CASBS, UC-LERF, NSF Grants SES-0645396 and SES-0850631, the Center for Equitable Growth at UC Berkeley, and the Sloan Foundation is gratefully acknowledged.

[†] To comment on this article in the online discussion forum, or to view additional materials, visit the article page at http://dx.doi.org/10.1257/app.5.1.1.

perhaps because of the program's complexity. The lack of information could potentially explain why the EITC induces small responses along the intensive margin (hours worked and earnings), despite substantially increasing labor force participation (Hotz and Scholz 2003).

We evaluate the impacts of information provision using a randomized experiment that provided information about the EITC to eligible tax filers and tracked the effect of this intervention on their subsequent earnings. The experiment was implemented at 119 H&R Block tax preparation offices in the Chicago metro area in 2007. The experimental population comprised approximately 43,000 tax filers who received EITC payments at one of the 119 H&R Block offices when filing taxes in 2007 and had one or more dependents. Half of these clients were randomly selected to receive a two-minute explanation about how the EITC works from their "tax professional," the H&R Block employee assisting them with their tax returns. Tax professionals were trained to use three tools to explain the EITC to their clients: a verbal description, a graph showing the shape of the EITC as a function of earnings, and a table listing the key EITC parameters. Each client was also given tailored advice emphasizing the implications of his marginal incentives conditional on his location in the EITC schedule. For example, clients in the phase-in region were told, "It pays to work more!"

We view our treatment as changing perceptions of marginal incentives around the tax filer's current location. Survey evidence indicates that most EITC recipients know the size of their current EITC refund, but do not understand the extent to which the EITC varies with their earnings. If the information treatment updates perceptions toward the true EITC schedule and informed tax filers are responsive along the intensive margin, tax filers should change their behavior to increase their EITC refunds. Such behavioral responses should generate a more concentrated earnings distribution around the peak of the EITC schedule.

We analyze the effects of the intervention using data from tax returns filed in 2007 ("year 1") and 2008 ("year 2"). Seventy-two percent of the clients in the treatment and control groups returned to H&R Block to file their taxes in the post-treatment year, allowing us to conduct a panel study of the effects of the information treatment on earnings. We begin with a simple analysis of treatment effects in the full sample. We find weak evidence (p=0.1) that treated clients have larger increases in EITC amounts from year 1 to year 2 relative to control clients. The effect is more pronounced for those with self-employment income in the base year (about 11 percent of the sample), although this effect is imprecisely estimated and still only marginally significant (p=0.1). The information treatment thus had at best a marginal effect on wage earnings behavior overall. We do not find significant effects when we cut the sample by whether the client was in the phase-in, phase-out, or plateau in the base year. Based on this analysis, we conclude that providing information about the tax code does not have a significant impact on labor supply behavior on average.

Next, we analyze heterogeneity of treatment effects across the 1,461 tax professionals who implemented this experiment. Many tax professionals felt that it was in their clients' best interest to work and earn more irrespective of the EITC's incentive effects, and might have framed the phase-out message as an encouragement

to work more because the loss in EITC benefits is relatively small. We first document that there is significant (p < 0.01) heterogeneity across tax professionals in mean treatment effects on EITC amounts using a nonparametric F test. To characterize the nature of the heterogeneity, we follow the methodology of Duflo et al. (2006). We divide tax professionals into two groups that we label "complying" and "noncomplying." To construct these groups, we first define a simple measure of the concentration of the earnings distribution in year 2—the fraction of returning clients with "middle" incomes (between \$7,000 and \$15,400).2 For each tax filer i, we define his tax professional as a "complier" if she has a higher fraction of other clients (excluding client i) with middle income in the treatment group than the control group. Intuitively, from the perspective of client i, complying tax professionals are those who increase the concentration of the earnings distribution for other clients. Critically, because we exclude client i when defining his tax professional's compliance, there is no correlation between client i's outcome and his tax professional's compliance under the hypothesis that all tax professionals have zero treatment effects.

For clients of complying tax professionals, the information treatment increases EITC amounts significantly—by \$58, on average, (p < 0.01), or about 3 percent. The treatment effects are larger for the self-employed, likely due to greater flexibility and reporting effects, as there is no third-party reporting of self-employment income. We also find a significant increase in the concentration of the distribution of wage earnings suggesting that the information intervention induced "real" changes in labor supply behavior for clients treated by complying tax professionals. For clients of noncomplying tax professionals, the information treatment does not lead to significant changes in EITC amounts. However, noncomplying tax professionals increase their treated client's incomes by \$250 (1.5 percent), on average (p < 0.05). Based on our discussions with tax professionals, we speculate that noncompliers may have used the information to simply encourage clients to aim for a high level of earnings rather than maximize their EITC refunds. The heterogeneity in impacts across tax preparers suggests that labor supply behavior may be influenced not just by information but also by the advice that tax professionals provide when helping to explain the incentives.

We conclude that information provision is not a very effective tool for changing earnings behavior, on average, though it might have effects in some subgroups when coupled with advice from tax professionals. The tailored provision of information by expert tax professionals is likely to be a stronger treatment than more easily scalable interventions, such as mailings of informational brochures. Hence, our study suggests that policies that disseminate information are not by themselves likely to change earnings behavior significantly.³ This lesson is consistent with recent

¹During focus groups prior to the experiment, several tax professionals argued that clients should always be encouraged to work more because, "you lose \$2 of EITC benefits for every \$10 you earn, but come out ahead by \$8 and possibly become eligible for other credits, so it still pays to work."

²The upper threshold of \$15,400 is the start of the EITC phase-out range. The lower threshold of \$7,000 is chosen to divide the remaining interval into two equal-sized bins. As we describe in the Appendix, alternative measures of the concentration of the earnings distribution yield similar results.

³ An alternative interpretation of our findings is wage earners may be unable to change their earnings in response to information. However, we find that earnings vary substantially across years within households in our sample,

evidence that information treatments have modest effects in other settings, such as college enrollment (Bettinger et al. 2009) or retirement savings (Beshears et al. 2011). While our results suggest that knowledge about the tax code cannot be easily manipulated with simple information treatments, the spread of knowledge through peer networks or other sources that affect knowledge in more persistent ways could have larger impacts on behavior (Chetty, Friedman, and Saez 2012).

The remainder of the paper is organized as follows. Section I provides background on the EITC and the literature on the effects of the program. Section II describes the experimental design and data. The main results are presented in Section III. Section IV presents results on heterogeneity across tax professionals. Section V concludes. Robustness checks and documentation of the materials used in the experiment are provided in the online Appendix.

I. Background on the EITC

A. Program Structure

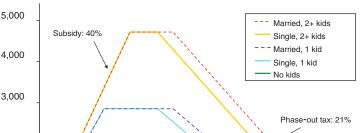
The EITC is a refundable tax credit administered through the income tax system. In 2009, the most recent year for which statistics are available, 25.9 million tax filers received a total of \$57.7 billion in EITC payments (Internal Revenue Service 2011, table 2.5). Eligibility for the EITC depends on earnings—defined as wage and salary income and self-employment income—and the number of qualifying children. Qualifying dependents for EITC purposes are relatives who are under age 19 (24 for full time students) or permanently disabled, and reside with the tax filer for at least half the year.⁴

Figure 1, panel A displays the EITC amount as a function of earnings for single and joint tax filers with zero, one, or two or more qualifying dependents in 2007, the year our experiment was conducted. EITC amounts increase substantially with the number of dependents, but the shape of the schedule as a function of earnings is the same in all three cases. EITC amounts first increase linearly with earnings, then plateau over a short income range, are then reduced linearly, and eventually phased out completely. Since the EITC amounts for tax filers with no children dependents are very small (maximum of \$428), we excluded them from our experiment, focusing only on tax filers with one or more children.

In the phase-in region, the subsidy rate is 34 percent for taxpayers with 1 child and 40 percent for taxpayers with 2 or more children. In the plateau (or peak) region, the EITC is constant and equal to a maximum value of \$2,853 and \$4,716 for tax filers with 1 and 2 or more children, respectively. In the phase-out region, the EITC amount decreases at a rate of 15.98 percent for filers with 1 child, and 21.06 percent for those with 2 or more children. The EITC is entirely phased out at earnings equal

partly because EITC claimants tend to hold many temporary jobs for short periods of time. Hence, we believe that adjustment frictions are unlikely to fully explain the lack of response to our information intervention, though they could certaintly have attenuated its effects.

⁴Only one tax filer can claim an eligible child. For example, in the case of nonmarried parents, only one parent can claim the child.



Panel A. EITC amount as a function of earnings

4,000 EITC amount (\$) 3,000 2,000 Subsidy: 34% 1,000 Phase-out tax: 16% 0 5,000 10,000 15,000 20,000 25,000 30,000 35,000 Earnings (\$)

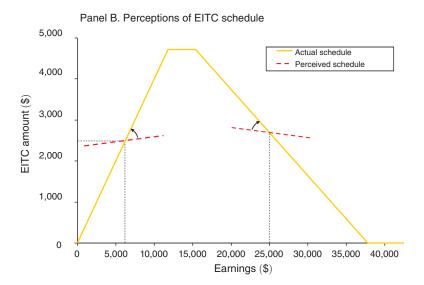


FIGURE 1. THE EARNED INCOME TAX CREDIT SCHEDULE AND PERCEPTIONS

Notes: Panel A depicts the EITC amount as a function of annual earnings in 2007. The EITC amount varies by marital status and number of qualifying children as shown. Panel B contrasts the actual EITC schedule for a single tax filer with two or more children with our model of the perceived schedule based on existing survey evidence. The perceived schedules are drawn for individuals with two levels of earnings, one in the phase-in and one in the phase-out range. Each individual accurately perceives the level of his EITC refund but underestimates the extent to which variations in earnings affect the size of his EITC. If implemented as intended, the information treatment should rotate the perceived EITC schedules (dashed lines) toward the actual EITC schedule (solid line) by clarifying the actual linkage between EITC amounts and earnings.

to \$33,241 and \$37,783 for single filers with 1 and 2 or more children, respectively.⁵ See IRS Publication 596 (Internal Revenue Service 2007) for complete details on program eligibility and rules as of 2007.

B. Claiming the EITC: Administrative Procedures

To claim the EITC, families file an income tax return that includes an EITC schedule between January 1 and April 15 of the following calendar year. The EITC is received in a single payment as part of the tax refund shortly after filing.⁶

According to the 2004 public use microdata on tax returns, 74 percent of families with children receiving the EITC use paid tax preparers to file their returns. The largest company in the market for paid tax preparation in the United States is H&R Block. H&R Block has about 13,000 offices located throughout the United States and employs over 100,000 tax professionals during the tax filing season. H&R Block currently prepares about 12 percent of individual tax returns in the United States. A substantial fraction of these returns are for EITC claimants, as over half of H&R Block's individual clients have an adjusted gross income (AGI) below \$35,000.

To file their tax returns, clients come to an H&R Block office with relevant documents, such as their W-2 wage income forms. The client sits with a "tax professional"—the term used to refer to H&R Block employees who prepare tax returns—in front of a computer running the H&R Block Tax Preparation Software (TPS). TPS consists of a series of screens corresponding to the various steps in tax return preparation. At each screen, the tax professional asks questions or inputs information from the forms brought in by the client. The tax preparation process takes about 30 to 45 minutes to complete for a typical EITC client.

C. Existing Evidence and Perceptions of EITC

There is a large empirical literature estimating the effects of the EITC on labor supply and earnings. Hotz and Scholz (2003) and Eissa and Hoynes (2006) provide comprehensive surveys. A number of studies have found strong evidence that the EITC increases labor force participation—the extensive margin response. However, there is little evidence that the EITC leads to a change in labor supply for those already in the labor market—the intensive margin. Most studies find no effects of the EITC on hours of work (see e.g., Meyer and Rosenbaum 1999 and Rothstein 2010). Using tax return data, Saez (2010) finds clear evidence of bunching of EITC recipients at the first kink of the EITC schedule, where the phase-in ends and the plateau starts, for recipients reporting self-employment income. However, there is

 $^{^{5}}$ For those who are married and file jointly, the plateau and phase-out regions of the EITC are extended by \$2,000 in 2007.

⁶There is an option to receive the EITC in advance during the year through the paycheck, but take-up of this option is extremely low (less than 2 percent). See Government Accountability Office (2007) and Jones (2010).

⁷ See e.g., Eissa and Liebman (1996), Meyer and Rosenbaum (2001). Eissa and Hoynes (2004) present complementary evidence of extensive-margin responses in the opposite direction. The labor force participation rate of married women in the phase-out region of the schedule fell slightly when the EITC was expanded. We expect that this extensive-margin response has a small impact on our results because 91 percent of the tax filers in our sample are single.

no bunching for recipients who do not report any self-employment income, who account for 89 percent of the individuals in our dataset.

The contrast between the strong responses along the extensive margin and small or zero responses along the intensive margin could be explained by a lack of information about the structure of the EITC (Liebman 1998; Hotz and Scholz 2003, 182). To respond along the extensive margin, families only need to know that working is associated with a large tax refund. In contrast, responding along the intensive margin requires knowledge about the nonlinear marginal incentives created by the three ranges of the EITC displayed in Figure 1, panel A. Surveys of low-income families and indepth interviews of EITC claimants show that there is widespread knowledge about the EITC's existence, but little knowledge about the structure of the EITC (Ross Phillips 2001; Olson and Davis 1994; Romich and Weisner 2002; Smeeding, Ross Phillips, and O'Connor 2002; Maag 2005). These interviews indicate that 60–90 percent of low-income families have heard about the EITC and know that it is a tax refund for working. However, less than 5 percent of these families know about the nonlinear pyramid shape of the EITC as a function of earnings and the location of the kink points.⁸

The lack of knowledge about the EITC's structure is striking given that the program parameters have been quite stable since 1996. However, it is not surprising in view of the information currently available about the program. To our knowledge, prior to our experiment, the graphical depiction of the EITC schedule shown in Figure 1, panel A could only be found in academic papers. Official Internal Revenue Service publications provide tables that show exact EITC amounts as a function of income and other characteristics, but do not summarize the EITC phase-in, peak, and phase-out structure in a transparent way. For legal reasons, the IRS only distributes comprehensive documents that cover all possible contingencies, making it impossible to highlight the features of the tax code most relevant for a given taxpayer. In addition, none of the existing commercial tax preparation software describes the EITC structure or marginal incentives explicitly.

We conclude from the existing literature that most EITC recipients know the value of their current EITC refund amount, but do not think about the slope of the EITC schedule when making marginal earnings decisions. For such EITC recipients, the local slope created by the EITC is therefore irrelevant in their labor supply decision. It is natural to assume that EITC recipients who do take into account the EITC when choosing their labor supply have unbiased beliefs about the relevant slope. In this case, the average EITC recipient's perception of the EITC schedule is flatter than the actual schedule. More precisely, let $EITC^p(z)$ denote the individual's perceived EITC refund at an earnings level of z and EITC(z) denote the actual EITC refund at that level of earnings. Let $s^p(z)$ denote the perceived local slope of the EITC schedule and s(z) the actual slope. The existing survey evidence suggests

⁸ Among the 42 families interviewed by Romich and Weisner (2002), 90 percent had heard of the EITC, but only two families knew that they needed to earn a certain amount to maximize their credit. One of those two families aimed at reporting self-employment earnings in order to maximize the credit (Romich and Weisner 2002, 378).

⁹ For example, the official IRS publication on the EITC intended for the public (Internal Revenue Service 2007, Publication 596) is 57 pages long and never explicitly mentions the slope parameters of the credit. The publication simply states the EITC amounts in the form of a 7 page table that has 4,770 entries.

that the representative individual with initial earnings z_0 perceives the relationship between earnings z and his EITC refund to be

(1)
$$EITC^{p}(z) = EITC(z_0) + (1 + s^{p}(z))(z - z_0),$$

where $|s^p(z)| < |s(z)|$. Figure 1, panel B illustrates the perceived budget constraint in (1) for two tax filers, one in the phase-in range and one in the phase-out range. Such misperceptions about marginal incentives motivate our question of whether improving knowledge (updating $s^p(z)$) could amplify the impacts of the EITC on intensive-margin labor supply.¹⁰

II. Experimental Design

We implemented the information-provision experiment in 119 H&R Block offices in the Chicago metropolitan area during the 2007 tax filing season (January 1 to April 15). Clients at these offices, who received an EITC with at least one eligible child, were randomly assigned into the treatment or control group. Assignment was based on the last two digits of the Social Security number of the primary filer. The probability of treatment assignment was 50 percent. The control group followed the standard tax preparation procedure using the TPS software described above. In the standard preparation procedure, a screen notifies the tax filer of his EITC amount if he is eligible for the EITC. This screen does not explain the structure of the EITC.

The new EITC information materials delivered by tax professionals to clients in the treatment group were developed in a series of steps. We began by interviewing 12 single mothers with recent work experience in the welfare office of San Francisco County in early October 2006. All 12 single mothers had filed tax returns in the past and almost all had heard about the EITC, but none knew about or had seen the graphical depiction of how the EITC varies with earnings. The interviewees found the graphical presentation of the EITC reasonably easy to understand and felt that it made the key features of the EITC very salient. Furthermore, most of the individuals recognized the value of this information for their work decisions and found the take-home messages sensible.¹¹

We refined the information materials in a focus group with 15 experienced H&R Block tax professionals and local managers in the Chicago area in late October 2006. Finally, H&R Block's internal staff and legal team edited and approved all the materials used in the experiment. The process described next is the final procedure that resulted from the collaborative effort between the researchers and H&R Block. Note that in all official tax forms, as well as in H&R Block materials, the EITC is referred to as the EIC (Earned Income Credit). We follow this convention in the information treatment materials described below.

¹⁰There is similar evidence that people are not fully informed about many other aspects of income tax schedules. See Fujii and Hawley (1988) for evidence from the United States, Brown (1968) for the United Kingdom, Bises (1990) for Italy, and Brannas and Karlsson (1996) for Sweden.

¹¹ For example, one of the interviewees suggested that we visit her housing complex to distribute this information more widely because her neighbors and friends would find it useful in making overtime and part-time work decisions.

A. Information Treatment

For the treatment group, two special "EIC information" screens are displayed automatically in TPS at the end of the tax preparation process. ¹² The first screen prompts the tax professional to begin the EIC explanation they were trained to provide and introduces the client to the information outreach program. This introductory screen is shown in online Appendix Exhibit I(a) for the case of a single filer with two or more dependents, the case on which we focus below for concreteness. The screen displays the EIC amount the tax filer is getting and describes the goal of the outreach effort, namely to help the client understand how the EIC depends on earnings. The second EIC information screen is displayed in online Appendix Exhibit I(b) for a tax filer in the increasing range of the EIC. This screen provides the key EIC information relevant to the tax filer's case, which the tax professional uses to explain the program to the client.

The central element of the explanation procedure is an "EIC handout" paper form that the tax professional fills out with the client and uses as a visual aid to explain the program. There are four EIC handouts based on the tax filer's marital status and dependents: single versus joint filer and one versus two or more dependents. Exhibit I shows the EIC handout for the case of a single filer with two or more dependents. The tax professional uses the information on the computer screen to fill in the blanks on the form in the following four steps.

First, the tax professional fills in the income that the client earned in 2006 and the corresponding EIC amount the client is receiving. Second, the tax professional draws a dot on the graph illustrating the client's location on the schedule. He then uses the graph to explain the link between earnings and the EIC amount.

In the third step, the tax professional circles the range of the schedule that the client is in—increasing, peak, or decreasing—and provides some advice corresponding to that range. In the increasing range, the take-home message is "Suppose you earn \$10 an hour, then you are really making \$14 an hour. It pays to work more!" In the peak range, the message is "Your earnings are maxing-out the EIC amount." In the decreasing range, the message is "If you earn \$10 more, your EIC is reduced by \$2.10. Earning more reduces your EIC, but you may qualify for additional tax credits."

The decreasing range message deliberately downplays the work disincentive created by the EITC in the phase-out region. The advice took this form because many managers and tax professionals at H&R Block felt strongly that it was in the best interest of tax filers to work and earn more. Indeed, many tax professionals pitched the message verbally as "You lose \$2 of your EIC credit when you earn \$10 more, but you still come out ahead by \$8 and potentially become eligible for other credits, so working more pays off." The fact that some tax professionals advised clients to

¹²This screen appears *after* all the client's tax information has been entered and the tax refund and liability have been calculated. We show that there is no difference in base year earnings across control and treatments groups, implying that treated tax filers did not go back and change their reported earnings in the base year after getting the EIC information.

¹³ In some cases, other credits, such as the nonrefundable portion of the child tax credit, do indeed increase with earnings in the EITC phase-out range, mitigating the implicit tax on work. We chose not to explain all aspects of the tax system in our information handout in the interest of simplicity.

aim for a high level of earnings—irrespective of the EITC's effect on incentives—appears to have important effects on the results, as we will see below.

In the fourth step, the tax professional circles the relevant range in the table that displays the exact parameters for the EITC. This table provides an alternative method of showing exactly how much the client can change his earnings before crossing the threshold for the next range. Tax professionals were trained to spend the most time on whichever of the three methods the client appeared to understand best—the verbal, graphical, or tabular descriptions.

After this information explanation is provided and the tax return process is completed, TPS automatically prints an "EIC printout" page that reproduces the information filled out in the handout. Online Appendix Exhibit II displays an example of the EIC printout. This page is printed at the same time as the tax return and inserted at the top of the packet given to the client to take home. The client is reminded by the tax professional that this information may prove useful when making earnings-related decisions later in the year. The purpose of the printout is to present the EITC information in a clean, accurate format. The temporary handout used to explain the program is kept by the tax professional.

Finally, to reinforce the treatment, H&R Block sent a letter summarizing the EITC information to all treatment-eligible clients in August 2007. Online Appendix Exhibit III displays an example of this letter.

As with most provisions of the tax code, EITC ranges are mechanically indexed for inflation and therefore differ slightly across the base year and subsequent year. Since our goal was to inform tax filers about the EITC parameters relevant for their subsequent labor supply decisions, the table and graph display the EITC parameters for 2007 earnings and the corresponding EITC that would be received when filing in 2008 (the post-treatment year). The classification of tax filers into the three groups—increasing, peak, and decreasing—was also based on the 2007 EITC parameters. As a result, a tax filer who was at the very beginning of the peak range would actually be presented with the increasing scenario that would apply were he to have the same nominal income in 2007. Similarly, a tax filer at the very beginning of the decreasing range would be presented with the peak scenario. Since the IRS inflation rate applied from tax year 2006 to 2007 was relatively small (3.9 percent), only 4 percent of taxpayers were located at a point where their current range differed from their predicted range for the following year. Note that the phase-in and phase-out *rates* were unchanged across the years.

B. Tax Professional Behavior

The effects of the experiment depend critically on the knowledge and behavior of the tax professionals. There were 1,461 tax professionals involved in the experiment, each of whom had 29 clients in our sample, on average (including treatment and control). We trained approximately 100 "office leaders" (senior tax professionals) in November 2006 ourselves, who then trained the rest of the tax professionals during December 2006. The training described the general goal of the outreach effort, why the experimental design required giving information to only half the clients, and explained the changes to the TPS system that would be introduced. A

series of case studies with hypothetical clients were used to illustrate various scenarios and how standardized explanations should be provided in the four steps. ¹⁴ Field observations in January 2007 confirmed that the EIC information screens and printouts were working as planned, and that tax professionals were implementing the experiment as trained.

In pilot sessions, we found that a minimum time of two minutes was required for a coherent explanation of the EITC. To give tax professionals an incentive to administer the information treatment carefully to eligible clients, each tax professional was offered \$5 for each eligible client with whom they spent at least two minutes on the EIC information screens (with time tracked by the software). If the tax professional attempted to exit the information screens before two minutes elapsed, the TPS system displayed a warning, "Does your client understand the explanation of how the EIC impacts their tax return?" The system then allowed the tax professional to go back and continue his explanation, resuming the two minute clock. Tax professionals who spent less than two minutes on the information screens did not receive any compensation for that client. Figure 2 displays a histogram of seconds spent by tax professionals on the EITC screens and shows that there is a clear spike at 120 seconds, implying that most tax professionals understood and responded to the compensation structure. The average time spent on the information screens conditional on reaching 120 seconds is 3.5 minutes.

Overall, 73 percent of tax filers whom we intended to treat were treated for at least two minutes. A substantial fraction of the variance in compliance rates is explained by office fixed effects, presumably due to variations in training. Most offices had very high compliance. However, one large office had a two-minute treatment rate of 6 percent, 11 percentage points below the next lowest office. We believe this exceptionally low-treatment rate arose from a failure to hold the planned training sessions. Since the treatment was effectively not implemented at this office, we exclude it from the analysis. ¹⁵

The decision to offer a 2 minute or longer EITC explanation to eligible clients may have depended on the client's interest in the information. Since a client's interest is not random, we follow standard practice in the experimental literature and estimate "intent-to-treat" effects—comparing outcomes of those eligible and ineligible to receive the information explanation.

To supplement the statistics on compliance rates, we directly assessed the tax professionals' reactions to the experiment using a survey of the tax professionals at the end of the tax season. See online Appendix Exhibit IV for the survey instrument. To obtain candid responses, the surveys identified offices but not individual tax professionals within those offices. Seventy-eight percent of the 119 offices sent back completed surveys, yielding a total of 785 survey responses. Eighty-eight percent of the tax professionals who responded to the survey thought that the EITC information should be offered again in the future. Eighty-one percent of surveyed tax professionals thought that the EITC experiment pilot helped their own understanding

¹⁴The powerpoint slides and case studies used for training are available from the authors upon request.

¹⁵ Including the office does not change our qualitative results but, unsurprisingly, slightly reduces the magnitude and precision of the estimates.

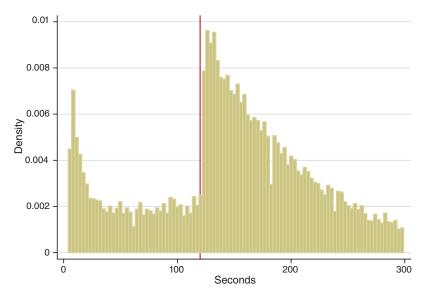


FIGURE 2. TIME SPENT EXPLAINING THE EITC TO CLIENTS ELIGIBLE FOR TREATMENT

Notes: This figure is a histogram of the time spent (in seconds) by tax professionals on explaining the EITC to clients eligible for the information treatment. Time spent was recorded by the tax preparation software. The vertical line at 120 seconds depicts the threshold above which tax professionals received \$5 of compensation (per client) for explaining the EITC. The histogram is based on 20,809 observations. Each bin represents an interval of three seconds.

of how the EITC credit works. This shows that our outreach effort did provide new information about the structure of the EITC beyond what is normally provided in the tax preparation procedure at H&R Block. As an important caveat, note that tax professionals who went through our training process may have offered better explanations on the EITC to tax filers in the control group as well. To minimize such contamination effects, we emphasized repeatedly in training that it was critical not to give any extra information to the clients who were not selected for treatment for the purpose of the study. Any remaining contamination effects would attenuate our treatment effect estimates. Nevertheless, it is important to recognize that the treatment is only the extra advice that trained professionals were willing to provide to treated clients using the guidance from TPS screens. ¹⁶

When asked about client interest, 37 percent of tax professionals said that "most" (>75 percent) of their clients were interested in the information explanation. Thirty-eight percent of the tax professionals said that "many" (25 to 75 percent) clients were interested, while 25 percent of tax professionals felt that few (<25 percent) of their clients were interested. We conclude from these surveys that most tax professionals were enthusiastic about the experiment and thought it was a valuable service for their clients, suggesting that the information treatment was implemented satisfactorily.

¹⁶Unfortunately, we do not have access to data outside of the experimental offices to test whether control clients in experimental offices responded to the experiment as well.

C. Hypothesis

The hypothesis we seek to test is that the provision of information and advice by tax professionals induces clients to change their earnings behavior. More specifically, tax professionals who implement our information treatment as intended should update their clients' perceptions toward the true EITC schedule, shifting s^p toward s in equation (1).¹⁷ This change in perceptions of marginal incentives rotates the perceived budget set as shown in Figure 1, panel B, generating substitution effects but no income effects. Such substitution effects should increase earnings for tax filers who would have been in the phase-in range absent the treatment, leave earnings unchanged for those in the peak, and decrease earnings for tax filers in the phase-out. Hence, in a neoclassical labor supply model, the information provided in the experiment should increase EITC refunds.

It is important to note that we provide information only about the EITC. In practice, other credits such as the Child Tax Credit, or the state and federal income taxes also affect the budget set. Hence, our treatment provides only partial information about the budget set. If individuals react to our information as if it were describing their exact budget set, their decisions might not increase their welfare. Note that if individuals are unable to understand or act upon the information provided in the treatment, then our basic theoretical framework predicts a zero marginal response. More generally, imperfect understanding will attenuate the experimental effects toward zero.

III. Results

Our analysis of the experimental results is based on anonymous statistical compilations prepared by H&R Block in accordance with applicable laws. These compilations were constructed from data extracted from tax returns filed in 2007 and 2008 and from supplemental information collected by H&R Block during the implementation of the experiment in 2007.

A. Descriptive Statistics

Table 1 presents descriptive statistics for the treatment and control groups. Columns 1–3 focus on the full sample, while columns 4–6 focus on the subsample of clients who returned to H&R Block in year 2 and for whom we have data on outcomes of the intervention. Columns 1–3 show that the means of all of the base year variables are similar in the treatment and control groups. None of the differences are significant at the 5 percent level, confirming that randomization was

¹⁷A key limitation of the present study is that we can only speculate about how our treatment changed baseline perceptions because we were unable to collect data on prior beliefs. As a result, we are only able to test the broad null hypothesis that information and advice do not affect behavior. Testing sharper hypotheses about the link between changes in priors and changes in behavior would be a valuable direction for future work.

¹⁸We opted to focus on explaining the EITC because explaining the full tax schedule would have been considerably more complicated, increasing the risk that individuals would not have understood our explanation.

¹⁹Unfortunately, we are unable to obtain tax return data for clients who did not return to H&R Block.

	TABLE 1—MEANS OF	BASE-YEAR	VARIABLES BY	TREATMENT	ELIGIBILITY
--	------------------	-----------	--------------	-----------	-------------

	Panel A. Full base year sample			Panel B. Year 2 returning sample		
Base year variables:	$ \begin{array}{c} \text{Control} \\ [N = 21,193] \\ (1) \end{array} $	Treatment $[N = 20,809]$ (2)	Difference $(2) - (1)$ (3)	$ \begin{array}{c} \text{Control} \\ [N = 15,380] \\ (4) \end{array} $	Treatment $[N = 14,925]$ (5)	Difference (5) - (4) (6)
Income (\$)	16,587 (74.52)	16,624 (79.77)	37.28 (82.35)	17,291 (79.68)	17,397 (84.62)	106.24 (96.26)
Wage earnings (\$)	15,872 (92.76)	15,913 (95.71)	40.25 (93.40)	16,626 (100.03)	16,756 (98.36)	129.75 (107.72)
EITC amount (\$)	2,478 (10.88)	2,465 (12.18)	-13.04 (12.27)	2,533 (11.93)	2,508 (13.19)	-24.41 (14.52)
Percent self-employed	11.40% (0.47)	11.18% (0.45)	-0.21% (0.32)	10.52% (0.49)	10.27% (0.45)	-0.25% (0.35)
Percent low income	14.30% (0.29)	14.69% (0.31)	0.39% (0.35)	11.13% (0.28)	11.62% (0.31)	0.49% (0.38)
Percent middle income	34.28% (0.44)	33.96% (0.47)	-0.32% (0.45)	33.92% (0.46)	33.14% (0.50)	-0.78% (0.53)
Percent upper income	51.41% (0.42)	51.34% (0.46)	-0.07% (0.48)	54.95% (0.46)	55.24% (0.50)	0.29% (0.57)
Percent married	9.53% (0.32)	9.40% (0.32)	-0.14% (0.28)	10.20% (0.37)	9.78% (0.36)	-0.42% (0.33)
Percent with two or more dependents in year 1	59.29% (0.37)	59.29% (0.39)	0.00% (0.48)	61.65% (0.41)	61.86% (0.44)	0.22% (0.55)
Percent return in year 2 percent with two or more	72.57% (0.34)	71.72% (0.37)	-0.85% (0.44)	100.00%	100.00%	0.00%

Notes: All variables are base year (year 1) values except last row. Standard errors clustered by tax professional reported in parentheses. Income is defined as the sum of wage income and self-employment income. Self-employed is a binary variable defined as having positive self-employment income (irrespective of other wage earnings). Low income is defined as income below \$7,000; middle income is defined as income between \$7,000 and \$15,400; and upper income is defined as income above \$15,400. Treatment group includes all tax filers we intended to treat. Columns 1–3 include the full sample in base year, while columns 4–6 include only those returning in year 2 (this is the sample of analysis).

successful. The mean income in the base year (year 1) in the full sample is \$16,600. Income is the sum of wage earnings and self-employment income. Average wage earnings are \$15,900. Average self-employment income is \$700, and 11 percent of tax filers report positive self-employment income.²⁰ The mean EITC amount in the base year is \$2,470. About 59 percent of the claimants have two or more dependents in the base year.

To examine distributional outcomes, throughout the paper we divide the income distribution into three bins: low incomes (below \$7,000), middle incomes (\$7,000 to \$15,400), and high incomes (above \$15,400). The upper threshold of \$15,400 is the start of the EITC phase-out range for single earners. The lower threshold of \$7,000 is chosen to divide the remaining interval into two approximately equal-sized bins. By this classification, 14 percent of the sample is "low income," 34 percent is "middle income," and 51 percent is "high income."

²⁰More precisely, positive self-employment income was measured as having positive self-employment taxes. No self-employment taxes are due if self-employment income is below \$400. Eleven percent of tax filers have self-employment income above \$400.

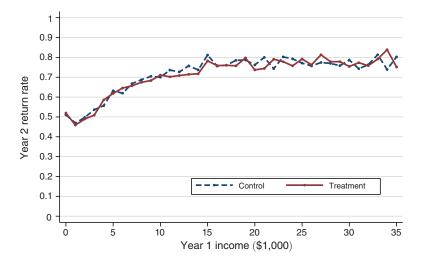


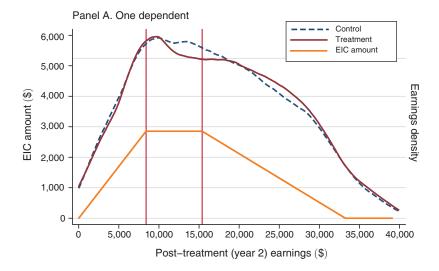
FIGURE 3. RETURN RATES BY BASE-YEAR INCOME

Notes: This figure plots the fraction of base-year clients who returned to H&R Block to file their taxes in year 2. Each point represents the average return rate in a \$1,000 bin. The return rates are plotted separately for the treatment (solid line) and control groups (dashed line).

The bottom row of Table 1 shows the fraction of clients who returned to H&R Block in year 2. The average return rate is around 72 percent. The return rate is 0.85 percent lower in the treatment group, a small but marginally significant difference. We explore the pattern of return rates further in Figure 3, which plots mean return rates by \$1,000 base-year earnings bins in the treatment and control groups. The average return rates track each other very closely, showing that there are no systematic patterns of differential attrition by base-year income. In addition, as shown in columns 4–6 of Table 1, there are no significant differences between the treatment and control groups in the base-year variables for the subsample of clients who return. In view of this evidence, we believe that the comparisons between the treatment and control groups that follow are unlikely to be contaminated by selective attrition.

B. Full Sample Results

We begin our empirical analysis by comparing changes in EITC amounts (from year 1 to year 2) in the treatment and control groups. A nonparametric Kolmogorov-Smirnov (KS) test for differences in the empirical distributions of changes in EITC amounts shows only a marginally significant difference between the treatment and control group (p=0.074), as shown in online Appendix Table A1. Figure 4 plots the density of post-treatment income using a kernel estimator with an Epanechnikov density function and constant bandwidth. The dashed line is for clients in the control group and the solid line is for clients in the treatment group. Panel A considers clients with one dependent, and panel B considers those with two or more dependents. The vertical lines mark the cutoffs for the phase-in and phase-out regions for each



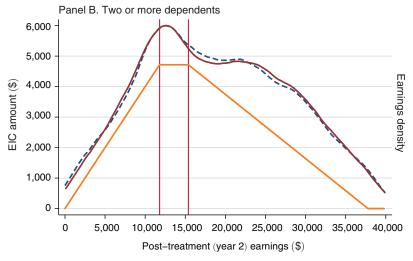


FIGURE 4. YEAR 2 EARNINGS DISTRIBUTIONS: FULL SAMPLE

Notes: These figures plot kernel densities of year 2 (post-treatment) income (sum of wage earnings and self-employment income) for the full sample of individuals filing with a tax professional. The solid curve shows the income distribution for the treatment group. The dashed curve shows the income distribution for the control group. Panel A is for tax filers with one qualifying dependent for EITC purposes in the base year, while panel B is for tax filers with two or more qualifying dependents. Each panel also shows the relevant EITC schedule (on the left y-axis). The vertical lines mark the boundaries between the phase-in, peak, and phase-out ranges of the EITC. Note that the EITC schedule shown in the figure and all subsequent figures is for single filers (91 percent of our sample). The EITC plateau for married filers is extended by \$2,000 (see Figure 1A).

case. Both panels show no discernible effect of the treatment on the earnings density distribution in year 2, confirming the results from the KS-test that the treatment does not have a large effect on EITC amounts.

Dependent variable Sample	Δ EITC amount (1)	Δ EITC amount with controls (2)	Δ Earnings (3)	Δ Earnings with controls (4)
(1) Full sample	24.02	17.17	17.66	29.35
	(14.77)	(14.06)	(84.27)	(83.46)
	30,303	30,303	30,303	30,303
(2) Year 1 in phase-in	3.88	9.47	-259.23	-263.60
	(31.68)	(28.15)	(150.15)	(148.46)
	7,442	7,442	7,442	7,442
(3) Year 1 in plateau	10.39	10.23	151.65	167.11
	(31.96)	(31.33)	(186.40)	(181.29)
	5,687	5,687	5,687	5,687
(4) Year 1 in phase-out	30.63	22.72	76.91	111.34
	(17.82)	(17.34)	(119.19)	(118.51)
	17,174	17,174	17,174	17,174
(5) Self-employed in year 1	72.60	66.10	97.45	80.63
	(45.05)	(43.21)	(247.61)	(242.65)
	3,150	3,150	3,150	3,150
(6) Wage earner in year 1	18.54	11.02	7.96	24.99
	(15.34)	(14.74)	(89.13)	(87.38)
	27,153	27,153	27,153	27,153

TABLE 2—TREATMENT EFFECTS ON EITC AMOUNTS AND EARNINGS

Notes: Standard errors clustered by tax professional reported in parentheses; number of observations is reported below the standard error. Each coefficient is from a separate regression. Columns show treatment effects on various outcomes—columns 1–2: change in EITC amount from year 1 to year 2; columns 3–4: change in earnings from year 1 to year 2. Columns 2 and 4 include the following base year controls: earnings, earnings squared, wage earnings, married filing jointly dummy, and number of qualifying children (one versus two or more).

Row 1 reports coefficients on the treatment indicator from OLS regressions of the form shown in equation (2) in the text for the full sample of tax filers who returned in year 2. Row 2 limits the sample to those with year 1 earnings in the EITC phase-in. Row 3 limits the sample to those with year 1 earnings in the EITC plateau. Row 4 limits the sample to those with year 1 earnings in the EITC phase-out. Row 5 limits the sample to those with positive self-employment income in year 1. Row 6 limits the sample to wage earners in year 1 (defined as not having self-employment income in year 1).

Next, we estimate treatment effects using OLS regressions of the form

(2)
$$y_i = \alpha + \beta \operatorname{treat}_i + \gamma \mathbf{X}_i + \varepsilon_i,$$

where y_i is an outcome (typically a change from year 1 to year 2); $treat_i$ is defined as an indicator for being eligible for the treatment; and \mathbf{X}_i is a vector of year 1 covariates. The coefficient of interest, β , can be interpreted as an intent-to-treat estimate. Estimates of β are presented in Table 2. The columns of Table 2 consider different outcomes or sets of covariates, while the rows consider different subsamples. Hence, each coefficient listed in the table is from a separate regression. We report standard errors clustered by tax professional in parentheses as well as the number of observations below the coefficient.

The dependent variable in columns 1 and 2 is the difference between the client's EITC amount in the post-treatment and pre-treatment years. Columns 3 and 4 consider the change in earnings from year 1 to year 2. In columns 2 and 4, we include the following vector of base-year covariates (\mathbf{X}) : earnings, earnings squared, wage earnings, indicator for married filing jointly, and number of children (one versus two or more).

Row 1 of Table 2 shows treatment effect estimates for the full sample. Consistent with the nonparametric KS test and graphical evidence presented above, we do not detect robust differences in EITC amounts or earnings distribution across the treatment and control groups. Most of the coefficients are small and statistically insignificant. There is weak evidence of a treatment effect on the change in EITC amounts (\$24 higher on average in the treatment group), but the effect is only marginally significant (p < 0.1).

C. Heterogeneity across Subgroups of Individuals

Rows 2–4 of Table 2 divide the sample into subgroups based on whether the filer's income was in the phase-in, plateau, and phase-out region in the base year. Recall from the experimental design that the take-home message varied based on this EITC range (see Exhibit I in the online Appendix). We do not find any significant effects of the information treatments within any of these subgroups.

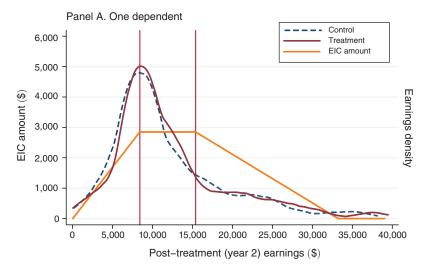
Next, we explore heterogeneity in treatment effect by self-employment status. The self-employed are able to manipulate their income more easily than wage earners, and thus might exhibit more of a response. As in Table 1, the self-employed are defined as the subsample of tax filers with positive self-employment income in the base year. Note that these tax filers may also have additional wage earnings beyond their self-employment income. Wage earners are defined as tax filers who do not have positive self-employment income in base year.

Figure 5 shows the effect of the treatment on the distribution of year 2 earnings for self-employed clients. Panel A is for clients with one dependent, and panel B is for those with two or more dependents. The control group exhibits clear bunching at the first kink point of the EITC schedule, the lowest earnings level at which one obtains the maximum refund. This is consistent with the finding of Saez (2010), who documents bunching at the first kink point among EITC recipients with self-employment income in IRS public use micro-data files. The degree of bunching is slightly amplified in the treatment group, suggesting that the information may have induced some self-employed tax filers to target the refund-maximizing peak more actively following the information treatment.

Rows 5 and 6 of Table 2 compare the impacts of the treatment on EITC amounts and earnings for the self-employed and wage earners. In row 5, the treatment effect on the change in EITC amounts is much larger than in the full sample (\$72.6 instead of \$24), consistent with the view that the self-employed were more responsive to the treatment. However, the effect is imprecisely estimated and remains only marginally significant (p < 0.1) due to the much smaller sample size. As shown in row 6, there is no significant effect on the EITC for wage earners.

²¹Because individuals pay payroll and other taxes on income, the first kink point of the EITC schedule maximizes the size of their net refund from the government.

²²For clients with self-employment income in base year, the treatment increases the probability of reporting earnings in the middle income range significantly in year 2 by 3.93 (standard error 1.57) percentage points.



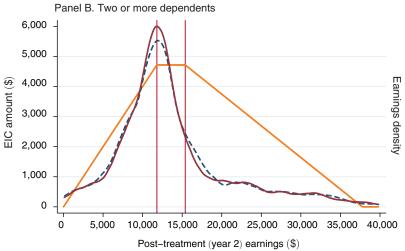


FIGURE 5. YEAR 2 EARNINGS DISTRIBUTIONS: SELF-EMPLOYED IN YEAR 1

Notes: These figures plot kernel densities of year 2 (post-treatment) income (sum of wage income and self-employment income) for tax filers who had positive self-employment earnings in the base year. The solid curve shows the income distribution for the treatment group; the dashed curve shows the income distribution for the control group. Panel A is for the sample of individuals with one dependent, while panel B is for the sample of individuals with two or more dependents. Each panel also shows the relevant EITC schedule for singles (on the left y-axis). The vertical lines mark the boundaries between the phase-in, peak, and phase-out ranges of the EITC.

IV. Heterogeneity across Tax Professionals

We expected that there might be heterogeneity in treatment effects across the 1,461 tax professionals involved in the experiment because of variation in training and willingness to convey the take-home messages we proposed. Such heterogeneity across tax professionals could potentially be masked in the full sample. We begin by implementing an F test for such treatment effect heterogeneity across tax

professionals. Let i = 1, ..., N index clients and p = 1, ..., P index tax professionals. Let $\Delta EITC_i$ denote the change in the EITC amount (from year 1 to year 2) for client i. Let $tp_{i,p}$ denote an indicator variable for whether client i is served by tax professional p, and $treat_i$ denote an indicator for whether the client is in the treatment group. We implement the F test using a regression of the following form:

$$\Delta EITC_i = \sum_{p=1}^{P} \theta_p t p_{i,p} + \sum_{p=1}^{P} \beta_p t reat_i \times t p_{i,p} + \varepsilon_i.$$

In this specification, β_p is tax professional p's treatment effect. The null hypothesis that $\beta_p=0$ for all p is rejected with p=0.0083, implying that some tax professionals generate significant differences in EITC amounts between their treatment and control clients. The hypothesis of constant treatment effects ($\beta_p=\beta_{p'}$ for all p,p') is rejected with p=0.0088, showing the importance of heterogeneity across tax professionals.

The remainder of this section characterizes the magnitudes and patterns of heterogeneity in treatment effects. We begin by developing a method of identifying complying tax professionals who implemented the treatment as planned and thereby induced changes in behavior as we hypothesized, namely increasing the concentration of earnings and EITC amounts. Note that the term "complier" simply refers to compliance with our ex ante intentions for the experiment. It should not be interpreted as a normative judgment about a tax professional, nor confused with the terminology used in the local average treatment effect literature in econometrics.

A. Definition of Compliers

Because we do not observe how tax professionals explained the information to clients, we use an indirect outcome-based method to identify complying tax professionals. For each tax filer i, we define his tax professional as a complier if the tax professional has a higher fraction of other clients (*excluding* client i) with middle income in the treatment group than the control group. Intuitively, from the perspective of a given client i, his tax professional complies with the intention of the experiment if the tax professional increases the concentration of the earnings distribution for her other clients. We define the remaining clients as having noncomplying tax professionals. We use such an outcome-based definition for compliers because we unfortunately do not have any information on tax professionals characteristics (such as experience, ability, or views on the EITC) that could have been used to cut the sample on predetermined characteristics.

²³ Note that *treat_i* is randomized *within* each tax professional's client group because treatment was randomized at the individual client level.

²⁴We also repeated the analysis below defining compliers versus noncompliers at the office level instead of the tax professional level. We do not find any significant treatment heterogeneity with this office-level definition, suggesting that the heterogeneity in treatments occurs primarily at the tax-professional level within offices rather than across offices.

Three important points should be noted about this definition of compliance. First, because client *i* himself is excluded when defining his tax professional's compliance, there is no correlation between client *i*'s outcome and his tax professional's compliance under the hypothesis that all tax professionals had zero treatment effects. A proof of this simple result is given in the online Appendix A1. To see the intuition, suppose a placebo treatment is randomly assigned to individuals, with no information provided to anyone. Define complying and noncomplying tax professionals for each client as above. In this case, complying and noncomplying are effectively randomly assigned, as the placebo treatment has no impact on year 2 earnings. Therefore, the sample of clients with a complying tax professional is simply a random subsample of the initial sample. Within that subsample, individual treatment status remains randomly assigned, and hence should have no impact on outcomes. Hence, we would detect zero treatment effects within the subsample of clients served by complying (or noncomplying) tax professionals if all tax professionals have zero treatment effects.²⁵

Second, the definition of complying tax professionals is client-specific, as excluding a particular client might shift a given tax professional from the complying to the noncomplying category (and vice versa). This creates a correlation in the error terms for clients served by the same tax professional, as similar clients will tend to either all be excluded or included in the complying group. We account for this problem by clustering all standard errors by tax professional. To check this method of computing standard errors, we also calculate p-values for each regression we run using the following permutation method. We first generate a placebo treatment randomly (with 50 percent probability) and recompute complying versus noncomplying tax professional status for each tax filer using this placebo treatment variable. We then estimate the regression specification using the placebo treatment in lieu of the actual treatment to obtain a placebo coefficient. This process is repeated 2,000 times to generate an empirical distribution of placebo coefficients. Finally, the permutation-based p-value is computed using the location of the actual treatment effect in the empirical cumulative distribution function of the placebo coefficients. We find that the difference between the permutation-based p-values and the p-values from regressions with clustered standard errors is less than 0.02 for every regression coefficient reported below.²⁶ This placebo analysis also confirms that our method of identifying complying tax professionals does not induce any artificial correlations between treatment and outcomes.

Third, the definition of compliance above is one of many possible definitions. In our baseline analysis, we define compliance based on the middle income indicator because it provides a simple, nonparametric way of measuring changes in the concentration of the earnings distribution. In online Appendix A2, we show that similar results are obtained when compliance is defined based on treatment effects on EITC amounts, which is effectively a smoother measure of changes in the concentration of

²⁵ As reported in online Appendix Table A2, the differences between the means of the base-year variables in the treatment and control groups are insignificant within the subsamples of clients served by complying and non-complying tax professionals, as in Table 1.

²⁶ Since there is no natural counterpart to clustering for the KS tests in Table 2, we report the permutation-based *p*-values in that table.

the income distribution (see online Appendix Table A4). We also show that controlling for base-year characteristics of clients when classifying tax professionals, and using continuous measures of the degree of compliance instead of a binary classification, yields similar results (see online Appendix Table A5).²⁷

B. Treatment Effects

Graphical Evidence.—Figure 6 plots the density of post-treatment income for clients with complying tax professionals who have one dependent (panel A) and two or more dependents (panel B). In both panels, there is greater mass in the treated group near the first kink point of the EITC schedule than there is in the control group. Conversely, there are fewer treated clients in the phase-out range. The increased concentration in the earnings distribution increases EITC amounts for treated clients. The differences between the treatment and control income distributions in Figure 6 are highly significant. Using a KS-test, the hypothesis that there are no differences in EITC amounts between treated and control clients is rejected with p < 0.01 for complying tax professionals, as shown in column 1 of online Appendix Table A1.

Figure 7 plots the density of post-treatment income for clients with noncomplying tax professionals. The earnings distribution for clients treated by noncompliers is shifted toward the right, placing more clients in the phase-out range, and thereby reducing their EITC refunds.²⁸

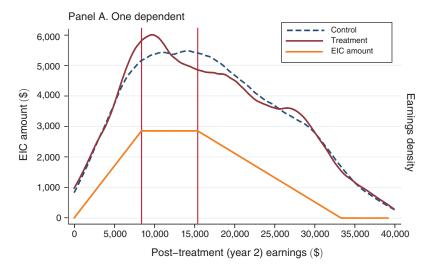
Figures 6 and 7 help explain why we detect no treatment effects in the full sample: the compliers and noncompliers shift the earnings distribution in opposite directions, generating little change in the full sample. The complying tax professionals induce behavioral responses consistent with the two specific hypotheses described in Section IIC. Noncomplying tax professionals did not generate a behavioral response consistent with EITC incentives, instead pushing more of their clients into the phase-out range. One potential explanation for this response is that the noncompliers are tax professionals who framed the EITC incentive effects as being small relative to the benefits of earning a higher income, which we anticipated might occur based on feedback prior to the experiment.

Regression Estimates.—To quantify the size of the behavioral responses, we estimate treatment effects within the complier and noncomplier subgroups, using the OLS specification in (2). The results are reported in Table 3. In all regressions, we control for base-year variables as in Table 2, columns 2 and 4. As a reference, row 1 of Table 3 first presents the estimates pooling compliers and noncompliers, replicating columns 3 and 4 in the first row of Table 2.

Column 1 of Table 3 reports the change in EITC amount. Column 2 reports the change in earnings. Column 3 reports the change in EITC amount among the

²⁷ A more ambitious approach, left for future research, would be to adopt the variable treatment setting of Angrist and Imbens (1995) with the additional difficulty that treatment intensity is not observed.

²⁸This shift in earnings distributions, and hence of the EITC amounts in the non-complying treatment group relative to the control group, is borne out by the KS-tests reported in row 3 of online Appendix Table A1.



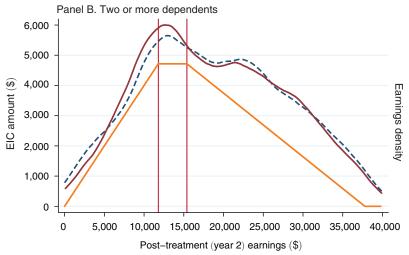
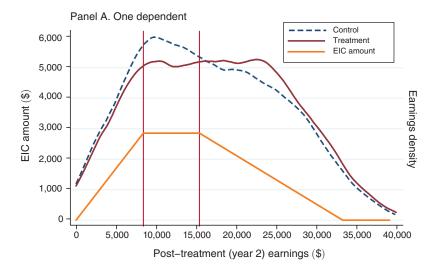


FIGURE 6. YEAR 2 EARNINGS DISTRIBUTIONS: COMPLYING TAX PROFESSIONALS

Notes: These figures plot kernel densities of year 2 (post-treatment) income (sum of wage earnings and self-employment income) for the sample of individuals filing with a "complying" tax professional. A given tax filer's tax professional is defined as a complier if she has a higher fraction of other clients (excluding client i) with middle income (between \$7,000 and \$15,400) in the treatment group than the control group. The solid curve shows the income distribution for the treatment group; the dashed curve shows the income distribution for the control group. Panel A is for tax filers with one qualifying dependent for EITC purposes in the base year, while panel B is for tax filers with two or more qualifying dependents. Each panel also shows the relevant EITC schedule for singles (on the left y-axis). The vertical lines mark the boundaries between the phase-in, peak, and phase-out ranges of the EITC.

self-employed in base year, and column 4 reports the change in EITC amount among the pure wage earners in base year. Finally, column 5 reports the change in EITC amounts computed exclusively using wage earnings (ignoring self-employment income) again for the sample of pure wage earners in base year. This last outcome



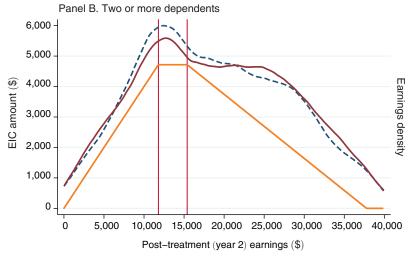


FIGURE 7. YEAR 2 EARNINGS DISTRIBUTIONS: NONCOMPLYING TAX PROFESSIONALS

Notes: These figures plot kernel densities of year 2 (post-treatment) income (sum of wage earnings and self-employment income) for the sample of individuals filing with a non-complying tax professional. A given tax filer's tax professional is defined as a noncomplier if she has a lower fraction of other clients (excluding client *i*) with middle income (between \$7,000 and \$15,400) in the treatment group than the control group. The solid curve shows the income distribution for the treatment group; the dashed curve shows the income distribution for the control group. Panel A is for tax filers with one qualifying dependent for EITC purposes in the base year, while panel B is for tax filers with two or more qualifying dependents. Each panel also shows the relevant EITC schedule for singles (on the left y-axis). The vertical lines mark the boundaries between the phase-in, peak, and phase-out ranges of the EITC.

detects effects on pure wage earnings. Consistent with our preceding results, none of the estimates in row 1 for the full sample are significantly different from zero.

Row 2 of Table 2 shows estimates for the subsample of clients served by complying tax professionals. Column 1 shows that clients treated by complying tax professionals increase their EITC amounts by \$58 (standard error 20.5) more than

Dependent variable	Δ EITC amount	Δ Earnings	Δ EITC amount	Δ EITC amount	Δ Wage based EITC amount
Sample	All (1)	A11 (2)	Year 1 self-employed (3)	Year 1 pure wage earners (4)	Year 1 pure wage earners (5)
(1) Full sample	17.17 (14.06) 30,303	29.35 (83.46) 30,303	66.10 (43.21) 3,150	11.02 (14.74) 27,153	3.72 (15.82) 27,153
(2) Complying tax professionals	58.05 (20.46) 15,395	-172.94 (123.66) 15,395	128.92 (59.69) 1,630	49.38 (21.48) 13,765	54.53 (22.48) 13,765
(3) Noncomplying tax professionals	-32.28 (20.40) 14,534	247.26 (119.87) 14,534	-27.47 (64.87) 1,495	-32.90 (21.21) 13,039	-56.53 (22.76) 13,039
(4) Compliers versus non- compliers: (2)–(3)	90.33 (30.20) 29,929	-420.20 (180.20) 29,929	156.40 (89.25) 3,125	82.29 (31.34) 26,804	111.06 (32.97) 26,804
(5) Compliers versus non- compliers with controls for Heterogeneity	89.78 (30.27) 29,929	-421.74 (180.68) 29,929	161.44 (89.22) 3,125	81.69 (31.35) 26,804	111.01 (33.00) 26,804

TABLE 3—TREATMENT EFFECTS BY TAX PROFESSIONAL COMPLYING STATUS

Notes: Standard errors clustered by tax professional reported in parentheses; *t*-statistics in square brackets; number of observations is reported below the standard error. Each coefficient is from a separate regression. Columns show treatment effects on various outcomes—columns 1, 3, 4: change in EITC amount from year 1 to year 2; column 2: change in earnings from year 1 to year 2; column 5: change in wage-based EITC amount (EITC computed based solely on wage earnings) from year 1 to year 2. All regressions include the following base year controls: earnings, earnings squared, wage earnings, married filing jointly dummy, and number of qualifying children (1 versus 2 or more). Column 3 limits the sample to those with positive self-employment income in year 1. Columns 4 and 5 limit the sample to pure wage earners (no self-employment income in year 1).

Row 1 reports coefficients on the treatment indicator from OLS regressions of the form shown in equation (2) in the text for the full sample of tax filers who returned in year 2. Row 2 limits the sample to complying tax professionals, and row 3 limits the sample to non-complying tax professionals. A given tax filer *i*'s tax professional is defined as a "complier" if she has a higher fraction of other clients (excluding client *i*) with middle income (between \$7,000 and \$15,400) in the treatment group than the control group. Row 4 reports the difference in treatment effects between complying and non-complying tax professionals, which equals the difference in coefficients between rows 2 and 3. In row 4, we regress each outcome variable on the treatment indicator, an indicator for having a complying tax professional, and the interaction of the two indicators.

The coefficient on the interaction is reported. We also include interactions of the base year control variables with the complying tax professional indicator. Row 5 reports the difference in treatment effects between complying and non-complying tax professionals controlling for heterogeneity in treatment effects by client observables. This specification adds interactions of the base year controls with the treatment indicator to the specifications in row 4. The coefficient on the treatment \times complying tax professional interaction is reported.

control group clients of the same tax professionals. Column 2 shows that the treatment does not induce a significant change in mean earnings from year 1 to year 2. The finding is consistent with an increase in concentration rather than a shift of the earnings distribution.

Row 3 considers the noncomplying tax professionals. Clients given the information treatment by these tax professionals experience a statistically insignificant *reduction* of \$32 (column 1) in their EITC amounts relative to their peers in the control group. This is because noncomplying tax professionals shift clients away from the region of the EITC schedule where refunds are maximized (Figure 7). Column 2 shows that the earnings of treated clients of noncompliers rise by \$247 (standard error 120) more, on average, than control clients. These results are consistent with

the density plots in Figure 7. Noncompliers shift the earnings distribution to the right and increase the likelihood of high incomes. The mean of the coefficients in rows 2 and 3 roughly corresponds to the coefficients in row 1, explaining why we do not detect clear treatment effects in the full sample.²⁹

Finally, in rows 4 and 5, we compare the treatment effects for complying and noncomplying tax professionals to test whether the estimates reported in rows 2 and 3 are statistically distinguishable. We estimate a model analogous to (2) on the full sample, interacting all the variables with an indicator for having a complying tax professional. Row 4 reports the coefficient on the interaction of the treatment and complier indicators, which is simply the difference in the coefficients reported in rows 3 and 4. Under the hypothesis of zero treatment effects for all tax professionals, this "difference-in-differences" estimate would be zero. Contrary to the null, all of the coefficients reported in row 4 are statistically significant. Clients treated by complying tax professionals experience a \$90 larger increase in their EITC refund, on average, relative to clients treated by noncomplying tax professionals. Furthermore, clients treated by compliers have, on average, \$420 lower growth in earnings than clients treated by noncompliers. These results highlight the substantial amount of treatment effect heterogeneity across tax professionals.

The heterogeneity in treatment effects that we have documented could come from two potential sources. One natural interpretation—which is the one we have suggested thus far—is that tax professionals implemented the information treatment in different ways, leading to different outcomes. An alternative view is that the variation in treatment effects is not caused by differences in tax professionals' behavior, but instead by variations in the set of clients that different types of tax professionals served. Our experiment randomized the information treatment within tax professional, but did not randomize clients across tax professionals. In row 5 of Table 3, we explore the source of the treatment effect heterogeneity by adding interactions of the vector of base-year controls with the treatment dummy to the specifications in row 4. In this specification, the coefficient on the interaction of the treatment and complier indicators can be interpreted as the effect of having a complying tax professional, holding fixed observable base-year characteristics. We find that all coefficients in row 5 are very similar to the corresponding coefficients in row 4, suggesting that the heterogeneity in treatment effects is not driven by observable heterogeneity in client characteristics.³⁰

²⁹ Online Appendix Table A3 refines this analysis by EITC range in the base year. It shows that most of the differential effects we uncover for compliers and noncompliers come from clients who were in the phase-out region in the base year, consistent with the view that tax professionals explained the phase-out incentives differently.

³⁰The heterogeneity in treatment effects could, however, be driven by *unobservable* heterogeneity in treatment effects across clients. For instance, suppose clients sort across tax professionals in a way that is correlated with their knowledge of the EITC. Then the heterogeneity in treatment effects across tax professionals could be driven by heterogeneity in clients' knowledge. Complying tax professionals could be those who serve clients with "flat" priors as in Figure 1, panel B, while noncomplying tax professionals could be those whose clients think that the phase-out rate is higher than it actually is. Note that such client heterogeneity explanations require substantial sorting of clients purely on unobserved characteristics. While we cannot rule out such sorting, we believe that the sharp differences in treatment effects across complying and noncomplying tax professionals are more likely to be driven by the tax professionals themselves.

Self-Employment Income versus Wage Earnings Responses.—Next, we explore the extent to which the treatment effects documented above are driven by changes in self-employment income versus wage earnings. This distinction is important to determine whether the information treatment changed labor supply or simply led to changes in reported income in order to maximize EITC refunds.

In column 3 of Table 3, we examine the self-employment income response by focusing on the subsample of tax filers with positive self-employment income in base year. Row 1 shows a marginally significant effect on this subsample, even without cutting the sample by tax professional complying status as we documented row 5 of Table 2. Row 2 shows that complying tax professionals increase their treated clients' EITC amounts by almost \$130 relative to the control group. This treatment effect for the self-employed is twice as large as those reported in the full sample (row 2, column 1). In contrast, row 3 shows that noncomplying tax professionals induce no significant treatment effects on their self-employed clients' EITC amounts or fraction with middle income. Rows 4 and 5 corroborate the substantial differences in year 2 outcomes between clients treated by compliers and noncompliers, even after controlling for observed client heterogeneity.

We next study the effect of the treatment on wage earnings. Column 4 of Table 3 considers the sample of pure wage earners in year 1 and estimates the effect of EITC changes. Row 2 shows that complying tax professional do increase EITC amounts by \$49 (standard error 21) in that subsample. In contrast, noncomplying tax professionals slightly reduce EITC amounts.

The increase in EITC refunds among clients of complying tax professionals could in principle be due to self-employment responses on the extensive margin, i.e., treated wage earners who start reporting self-employment income to increase their EITC refunds. However, we find no significant increase in the likelihood to report self-employment income in this subsample. As an alternative method to quantify the impact on wage earnings itself, we compute EITC amounts based solely on wage earnings.³¹ We report such coefficients in column 5 of Table 3, again for the subsample of those with no self-employment income in base year. Row 1 shows that there is no significant difference in wage-based EITC amounts between the treatment and control groups in the full sample pooling compliers and noncompliers. Row 2 shows that clients treated by complying tax professionals have a \$55 increase in their wage-based EITC amounts relative to control clients (p < 0.05). Noncomplying tax professionals, in contrast, reduce their treated clients' wage-based EITC amounts by \$57 (p < 0.05). Finally, rows 4 and 5 confirm that there are highly significant (p < 0.01) differences in year 2 outcomes between clients treated by compliers and noncompliers, even after controlling for observed client heterogeneity.

Online Appendix Figures A1 and A2 show the counterpart of Figures 6 and 7 using wage earnings instead of total earnings. Figure A1 shows that complying tax professionals increase the mass of the wage earnings distribution around the first kink point for treated clients. This increase in mass is slightly smaller than the

³¹More precisely, we compute the EITC amount that the tax filer would have obtained if her self-employment income were zero (and her wage income was left unchanged). For pure wage earners, actual EITC amounts and wage based EITC amounts naturally coincide.

change in the distribution of total income shown in Figure 6, confirming that part of the treatment effect is driven by the self-employment margin. In contrast, Figure A2 shows that clients given the information treatment by noncomplying tax professionals are more likely to have wage earnings that place them in the phase-out range.³²

The finding that noncompliers increase wage earnings but induce no change in reported self-employment income suggests that they did not explain how to maximize EITC refunds. Conversely, the fact that compliers induce stronger responses in self-employment income—which is easier to manipulate via reporting effects—than wage income (Internal Revenue Service 1996, 8, table 2) suggests that they emphasized the behaviors relevant for maximizing the EITC refund.

V. Conclusion

This paper has reported the results of an experiment testing the effects of providing information about the structure of the EITC on earnings decisions. We find that the information treatment did not induce significant changes in earnings, on average. We find some evidence of heterogeneous responses to the information treatment across the H&R Block tax professionals who implemented the experiment. Half of the tax professionals increase their treated clients' EITC amounts and the concentration of their wage earnings distribution around the first kink point of the EITC schedule. The remaining tax professionals do not induce a significant change in EITC amounts, but increase their clients' probabilities of having high-wage earnings that place them in the phase-out range. We speculate that this heterogeneity in treatment effects arises from the different ways in which tax professionals used the information to advise their clients.

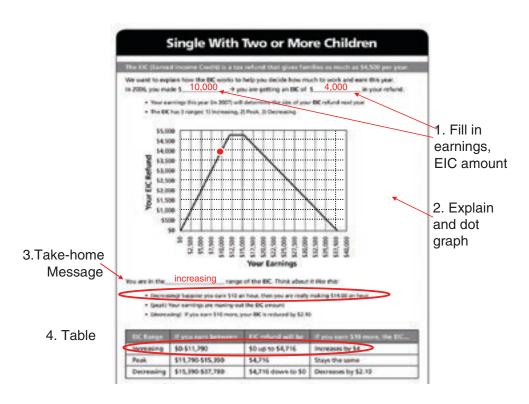
The heterogeneous treatment effects we document are modest in absolute terms, but are fairly large in comparison with intensive margin responses to other policies. Previous studies suggest that the intensive margin elasticity of earnings with respect to the net-of-tax rate is approximately 0.25 (e.g., Chetty 2012). Using this elasticity, a simple calibration exercise (see online Appendix A3) shows that complying tax professionals generate the same labor supply response along the intensive margin as a 33 percent expansion of the EITC. Noncomplying tax professionals increase earnings by an amount equivalent to the response to a 5 percentage point tax rate cut. These findings suggest that tax professionals can influence their clients' earnings choices significantly, and that such advice may have more of an impact on behavior than the pure information provided on the EITC handouts themselves. Unfortunately, we are unable to characterize the mechanisms through which such advice affects behavior. The decentralized implementation of our experiment makes it difficult to define the "treatment" that was provided by each of the tax professionals. In particular, we do not have measures of the informational content, clarity, or salience of the treatment provided by each tax professional.³³

³²In column 2 of online Appendix Table A1, we report the results of KS-tests for a difference between treatment and control groups in the distribution of *wage-based* EITC amounts. These tests confirm that both complying and non-complying tax professionals significantly change their treated clients' distribution of wage earnings.

³³Bhargava and Manoli (2012) conduct a randomized experiment on EITC take-up that implements variation along these dimensions, and shows that each of them matters significantly.

We conclude that providing information about marginal income tax incentives does not have systematic impacts on earnings in the short run. However, recent work by Chetty, Friedman, and Saez (2012) suggests that local knowledge among peers does affect EITC claimants' and affects both self-employment and wage earnings significantly. Chetty, Friedman, and Saez (2012) show that the EITC has very different impacts on earnings behavior across neighborhoods in the United States, and that these differences are likely driven by variation in knowledge about the shape of the EITC schedule. Together, these results suggest that knowledge may have to be manipulated more organically and persistently—e.g., by changing peers' behavior—rather than via one-time provision of information to influence behavior. Investigating the process through which knowledge about government policy diffuses and understanding how it can be shaped by policy would be a very valuable direction for future work.

APPENDIX—EXHIBIT 1. EITC HANDOUT: 4 STEP EXPLANATION



REFERENCES

Angrist, Joshua D., and Guido W. Imbens. 1995. "Two-Stage Least Squares Estimation of Average Causal Effects in Models with Variable Treatment Intensity." *Journal of the American Statistical Association* 90 (430): 431–42.

Beshears, John, James Choi, David Laibson, and Brigitte C. Madrian. 2011. "How Does Simplified Disclosure Affect Individuals' Mutual Fund Choices?" In *Explorations in the Economics of Aging*, edited by David A. Wise, 75–100. Chicago: University of Chicago Press.

- Bettinger, Eric P., Bridget Terry Long, Philip Oreopoulos, and Lisa Sanbonmatsu. 2009. "The Role of Simplification and Information in College Decisions: Results from the H&R Block FAFSA Experiment." National Bureau of Economic Research (NBER) Working Paper 15361.
- **Bhargava, Saurabh, and Day Manoli.** 2012. "Why Are Benefits Left on the Table? Assessing the Role of Information, Complexity, and Stigma on Take-up with an IRS Field Experiment." http://elsa.berkeley.edu/~saez/course/report-paper1.pdf.
- **Bises, Bruno.** 1990. "Income Tax Perception and Labour Supply in a Sample of Industry Workers." *Public Finance* 45 (1): 3–17.
- **Brannas, Kurt, and Niklas Karlsson.** 1996. "Estimating the Perceived Tax Scale within a Labor Supply Model." *Economic Letters* 52 (1): 75–77.
- **Brown, Charles V.** 1968. "Misconceptions about Income Tax and Incentives." *Scottish Journal of Political Economy* 15: 1–21.
- **Chetty, Raj.** 2012. "Bounds on Elasticities with Optimization Frictions: A Synthesis of Micro and Macro Evidence on Labor Supply." *Econometrica* 80 (30): 969–1018.
- Chetty, Raj, and Emmanuel Saez. 2013. "Teaching the Tax Code: Earnings Responses to an Experiment with EITC Recipients: Dataset." American Economic Journal: Applied Economics. http://dx.doi.org/10.1257/app.5.1.1.
- Chetty, Raj, John Friedman, and Emmanuel Saez. 2012. "Using Differences in Knowledge Across Neighborhoods to Uncover the Impacts of the EITC on Earnings." National Bureau of Economic Research (NBER) Working Paper 18232.
- Chetty, Raj, Adam Looney, and Kory Kroft. 2009. "Salience and Taxation: Theory and Evidence." American Economic Review 99 (4): 1145–77.
- de Bartolome, Charles. 1995. "Which Tax Rate Do People Use: Average or Marginal?" *Journal of Public Economics* 56 (1): 79–96.
- Duflo, Esther, William Gale, Jeffrey Liebman, Peter Orszag, and Emmanuel Saez. 2006. "Saving Incentives for Low- and Middle-Income Families: Evidence from a Field Experiment with H&R Block." Quarterly Journal of Economics 121 (4): 1311–46.
- **Eissa, Nada, and Hilary W. Hoynes.** 2004. "Taxes and the Labor Market Participation of Married Couples: The Earned Income Tax Credit." *Journal of Public Economics* 88 (9–10): 1931–58.
- **Eissa, Nada, and Hilary W. Hoynes.** 2006. "Behavioral Responses to Taxes: Lessons from the EITC and Labor Supply." In *Tax Policy and the Economy*, Vol. 20, edited by James Poterba, 73–110. Cambridge, MA: MIT Press.
- **Eissa, Nada, and Jeffrey B. Liebman.** 1996. "Labor Supply Response to the Earned Income Tax Credit." *Quarterly Journal of Economics* 111 (2): 605–37.
- Fujii, Edwin T., and Clifford B. Hawley. 1988. "On the Accuracy of Tax Perceptions." *Review of Economics and Statistics* 70 (2): 344–47.
- Government Accountability Office. 2007. Advance Earned Income Tax Credit: Low Use and Small Dollars Paid Impede IRS's Effort to Reduce High Noncompliance GAO-07-1110. Government Accountability Office (GAO). Washington, DC, August.
- Hotz, V. Joseph, and John Karl Scholz. 2003. "The Earned Income Tax Credit." In Means-Tested Transfer Programs in the United States, edited by Robert Moffitt, 141–98. Chicago: University of Chicago Press.
- Internal Revenue Service. 1996. Federal Tax Compliance Research: Individual Income Tax Gap Estimates for 1985, 1988, and 1992 Publication 1415 (Rev. 4-96). Department of Treasury. Washington, DC, April.
- Internal Revenue Service. 2007. Earned Income Credit (EIC): For Use in Preparing 2007 Returns Publication 596. Department of Treasury. Washington, DC, January.
- Internal Revenue Service. 2012. Statistics of Income: Individual Income Tax Returns, 2009 Publication 1304. Department of Treasury. Washington, DC, June.
- Jones, Damon. 2010. "Information, Preferences, and Public Benefit Participation: Experimental Evidence from the Advance EITC and 401(k) Savings." American Economic Journal: Applied Economics 2 (2): 147–63.
- **Liebman, Jeffrey.** 1998. "The Impact of the Earned Income Tax Credit on Incentives and the Income Distribution." In *Tax Policy and the Economy*, Vol. 12, edited by James Poterba, 83–123. Cambridge, MA: MIT Press.
- **Liebman, Jeffrey, and Erzo Luttmer.** 2011. "Would People Behave Differently If They Better Understood Social Security? Evidence From a Field Experiment." National Bureau of Economic Research (NBER) Working Paper 17287.
- Maag, Elaine. 2005. Paying the Price? Low-Income Parents and the Use of Paid Tax Preparers. Urban Institute New Federalism: National Survey of America's Families B-64. Washington, DC, February.

- Meyer, Bruce, and Dan Rosenbaum. 1999. "Welfare, the Earned Income Tax Credit, and the Labor Supply of Single Mothers." National Bureau of Economic Research (NBER) Working Paper 7363.
- Meyer, Bruce, and Dan Rosenbaum. 2001. "Welfare, the Earned Income Tax Credit, and the Labor Supply of Single Mothers." *Quarterly Journal of Economics* 116 (3): 1063–1114.
- Olson, Lynn M., and Audrey Davis. 1994. "The Earned Income Tax Credit: Views from the Street Level." Northwestern University Working Paper WP-94-1.
- Romich, Jennifer L., and Thomas S. Weisner. 2002. "How Families View and Use the Earned Income Tax Credit: Advance Payment Versus Lump-Sum Delivery." In *Making Work Pay*, edited by Bruce Meyer and Douglas Holtz-Eakin, 366–92. New York: Russell Sage Foundation.
- Ross Phillips, Katherin. 2001. Who Knows About the Earned Income Tax Credit? Urban Institute Policy Brief B-27. Washington, DC, January.
- Rothstein, Jesse. 2010. "Is the EITC as Good as an NIT? Conditional Cash Transfers and Tax Incidence." *American Economic Journal: Economic Policy* 2 (1): 177–208.
- Saez, Emmanuel. 2010. "Do Taxpayers Bunch at Kink Points?" American Economic Journal: Economic Policy 2 (3): 180–212.
- Smeeding, Timothy M., Katherin Ross Phillips, and Michael A. O'Connor. 2002. "The Earned Income Tax Credit: Expectation, Knowledge, Use, and Economic and Social Mobility." In *National Tax Journal* 53 (4): 1187–1209.

This article has been cited by:

- 1. Tuomas Matikka, Tuuli Paukkeri. 2022. Does sending letters increase the take-up of social benefits? Evidence from a new benefit program. *Empirical Economics* 10. . [Crossref]
- Paul Lanier, Whitney Afonso, Gerard Chung, Katherine Bryant, Danny Ellis, Alexandria Coffey, Anita Brown-Graham, Sarah Verbiest. 2022. Demographic, Economic, and Geographic Factors Associated with Uptake of the Earned Income Tax Credit. *Journal of Policy Practice and Research* 36. [Crossref]
- 3. Kai Barron, Christina Gravert. 2022. Confidence and career choices: an experiment*. *The Scandinavian Journal of Economics* 124:1, 35-68. [Crossref]
- 4. Andreas R. Kostøl, Andreas S. Myhre. 2021. Labor Supply Responses to Learning the Tax and Benefit Schedule. American Economic Review 111:11, 3733-3766. [Abstract] [View PDF article] [PDF with links]
- 5. Rik Peeters, Sergio A. Campos. 2021. Taking the bite out of administrative burdens: How beneficiaries of a Mexican social program ease administrative burdens in street-level interactions. *Governance* 34:4, 1001-1018. [Crossref]
- 6. Mariana Chudnovsky, Rik Peeters. 2021. The unequal distribution of administrative burden: A framework and an illustrative case study for understanding variation in people's experience of burdens. *Social Policy & Administration* 55:4, 527-542. [Crossref]
- 7. Didier Boost, Peter Raeymaeckers, Koen Hermans, Sara Elloukmani. 2021. Overcoming non-take-up of rights: A realist evaluation of Integrated Rights-Practices. *Journal of Social Work* 21:4, 831-852. [Crossref]
- 8. John Giles, Xin Meng, Sen Xue, Guochang Zhao. 2021. Can information influence the social insurance participation decision of China's rural migrants?. *Journal of Development Economics* **150**, 102645. [Crossref]
- 9. Jan-Emmanuel De Neve, Clément Imbert, Johannes Spinnewijn, Teodora Tsankova, Maarten Luts. 2021. How to Improve Tax Compliance? Evidence from Population-Wide Experiments in Belgium. *Journal of Political Economy* 129:5, 1425-1463. [Crossref]
- 10. Richard Domurat, Isaac Menashe, Wesley Yin. 2021. The Role of Behavioral Frictions in Health Insurance Marketplace Enrollment and Risk: Evidence from a Field Experiment. American Economic Review 111:5, 1549-1574. [Abstract] [View PDF article] [PDF with links]
- 11. Jonas Hjort, Diana Moreira, Gautam Rao, Juan Francisco Santini. 2021. How Research Affects Policy: Experimental Evidence from 2,150 Brazilian Municipalities. *American Economic Review* 111:5, 1442-1480. [Abstract] [View PDF article] [PDF with links]
- 12. Eric Zwick. 2021. The Costs of Corporate Tax Complexity. *American Economic Journal: Economic Policy* 13:2, 467-500. [Abstract] [View PDF article] [PDF with links]
- 13. Sydnee Caldwell, Scott Nelson, Daniel C. Waldinger. 2021. Tax Refund Uncertainty: Evidence and Welfare Implications. SSRN Electronic Journal . [Crossref]
- 14. Chong-En Bai, Wei Chi, Tracy Xiao Liu, Chao Tang, Jian Xu. 2021. Boosting pension enrollment and household consumption by example: A field experiment on information provision. *Journal of Development Economics* 102622. [Crossref]
- 15. SAKSHI BHARDWAJ. 2021. Effect of the 1996 Welfare Reform on Undocumented Immigrants. SSRN Electronic Journal 49. . [Crossref]
- 16. Roberto Iacono, Marco Ranaldi. 2020. The nexus between perceptions of inequality and preferences for redistribution. *The Journal of Economic Inequality* 116. . [Crossref]

- 17. Sebastien Bradley, Naomi E. Feldman. 2020. Hidden Baggage: Behavioral Responses to Changes in Airline Ticket Tax Disclosure. *American Economic Journal: Economic Policy* 12:4, 58-87. [Abstract] [View PDF article] [PDF with links]
- 18. William C. Boning, John Guyton, Ronald Hodge, Joel Slemrod. 2020. Heard it through the grapevine: The direct and network effects of a tax enforcement field experiment on firms. *Journal of Public Economics* 190, 104261. [Crossref]
- 19. Peter Shirley. 2020. First-time mothers and the labor market effects of the earned income tax credit. *IZA Journal of Labor Policy* 10:1. . [Crossref]
- 20. Jacob A. Mortenson, Andrew Whitten. 2020. Bunching to Maximize Tax Credits: Evidence from Kinks in the US Tax Schedule. American Economic Journal: Economic Policy 12:3, 402-432. [Abstract] [View PDF article] [PDF with links]
- 21. Klaus Armingeon, Reto Bürgisser. 2020. Trade-offs between redistribution and environmental protection: the role of information, ideology, and self-interest. *Journal of European Public Policy* **76**, 1-21. [Crossref]
- 22. Elliott Isaac. 2020. Marriage, Divorce, and Social Safety Net Policy. *Southern Economic Journal* **86**:4, 1576-1612. [Crossref]
- 23. Oronde Small, Leanora Brown. 2020. Taxpayer Service Provision and Tax Compliance: Evidence for Large Taxpayers in Jamaica. *Public Finance Review* **48**:2, 250-277. [Crossref]
- Jessica Leight, Nicholas Wilson. 2020. Framing Flexible Spending Accounts: A Large-Scale Field Experiment on Communicating the Return on Medical Savings Accounts. *Health Economics* 29:2, 195-208. [Crossref]
- 25. Hilary Hoynes. 2019. The Earned Income Tax Credit. *The ANNALS of the American Academy of Political and Social Science* **686**:1, 180-203. [Crossref]
- 26. Johannes Becker, Jonas Fooken, Melanie Steinhoff. 2019. Behavioral Effects of Withholding Taxes on Labor Supply*. *The Scandinavian Journal of Economics* **121**:4, 1417-1440. [Crossref]
- 27. Karen Z. Kramer, Flávia Cristina Drumond Andrade, Andrew J. Greenlee, Ruby Mendenhall, Dylan Bellisle, Renee Lemons Blanks. 2019. Periodic Earned Income Tax Credit (EITC) Payment, Financial Stress and Wellbeing: A Longitudinal Study. *Journal of Family and Economic Issues* 40:3, 511-523. [Crossref]
- 28. Hilary Hoynes, Mark Stabile. 2019. How Do the US and Canadian Social Safety Nets Compare for Women and Children?. *Journal of Labor Economics* 37:S2, S253-S288. [Crossref]
- 29. Stefan Klößner, Gregor Pfeifer. 2019. The Importance of Tax Adjustments When Evaluating Wage Expectations*. *The Scandinavian Journal of Economics* 121:2, 578-605. [Crossref]
- 30. Davide Cerruti, Claudio Daminato, Massimo Filippini. 2019. The Impact of Policy Awareness: Evidence from Vehicle Choices Response to Fiscal Incentives. SSRN Electronic Journal. [Crossref]
- 31. Ghazala Azmat. 2019. Incidence, salience, and spillovers: The direct and indirect effects of tax credits on wages. *Quantitative Economics* **10**:1, 239-273. [Crossref]
- 32. Marc K. Chan, Robert Moffitt. 2018. Welfare Reform and the Labor Market. *Annual Review of Economics* **10**:1, 347-381. [Crossref]
- 33. Philip Armour. 2018. The Role of Information in Disability Insurance Application: An Analysis of the Social Security Statement Phase-In. *American Economic Journal: Economic Policy* 10:3, 1-41. [Abstract] [View PDF article] [PDF with links]

- 34. Steffen Altmann, Armin Falk, Simon Jäger, Florian Zimmermann. 2018. Learning about job search: A field experiment with job seekers in Germany. *Journal of Public Economics* **164**, 33-49. [Crossref]
- 35. Alex Rees-Jones, Dmitry Taubinsky. 2018. Taxing Humans: Pitfalls of the Mechanism Design Approach and Potential Resolutions. *Tax Policy and the Economy* **32**:1, 107-133. [Crossref]
- 36. Dominic P. Parker, Walter N. Thurman. 2018. Tax Incentives and the Price of Conservation. Journal of the Association of Environmental and Resource Economists 5:2, 331-369. [Crossref]
- 37. Hanako Ohmura. 2018. Economic voting using egotropic evaluation as an information cue: how absolute and relative income affects socio-economic evaluations and retrospection. Behaviormetrika 45:1, 57-89. [Crossref]
- 38. Koichiro Ito, Takanori Ida, Makoto Tanaka. 2018. Moral Suasion and Economic Incentives: Field Experimental Evidence from Energy Demand. *American Economic Journal: Economic Policy* 10:1, 240-267. [Abstract] [View PDF article] [PDF with links]
- 39. Jacob Mortenson, Andrew Whitten. 2018. Bunching to Maximize Tax Credits: Evidence from Kinks in the U.S. Tax Schedule. SSRN Electronic Journal. [Crossref]
- 40. Riley Wilson. 2018. The EITC and Employment Transitions: Labor Force Attachment, Annual Exit, and the Role of Information. SSRN Electronic Journal. [Crossref]
- 41. J.J. Prescott. 2018. Assessing Access-to-Justice Outreach Strategies. SSRN Electronic Journal . [Crossref]
- 42. Koichiro Ito, Takanori Ida, Makoto Tanaka. 2018. Moral Suasion and Economic Incentives: Field Experimental Evidence from Energy Demand. SSRN Electronic Journal . [Crossref]
- 43. Nicolas Duquette, Alexandra Graddy-Reed, Mark Phillips. 2018. The Extent and Efficacy of State Charitable Contribution Income Tax Credits. SSRN Electronic Journal. [Crossref]
- 44. Taylor Cranor, Jacob Goldin. 2018. Does Informing Employees About Tax Benefits Increase Take-Up?: Evidence From EITC Notification Laws. SSRN Electronic Journal . [Crossref]
- 45. Kai mname Barron, Christina mname Gravert. 2018. Confidence and Career Choices: An Experiment. SSRN Electronic Journal 106. . [Crossref]
- 46. Shanthi P. Ramnath, Patricia K. Tong. 2017. The Persistent Reduction in Poverty from Filing a Tax Return. *American Economic Journal: Economic Policy* **9**:4, 367-394. [Abstract] [View PDF article] [PDF with links]
- 47. John Beshears, James J. Choi, David Laibson, Brigitte C. Madrian. 2017. Does front-loading taxation increase savings? Evidence from Roth 401(k) introductions. *Journal of Public Economics* 151, 84-95. [Crossref]
- 48. Sebastien Bradley. 2017. Inattention to Deferred Increases in Tax Bases: How Michigan Home Buyers Are Paying for Assessment Limits. *Review of Economics and Statistics* **99**:1, 53-66. [Crossref]
- 49. Florian Buhlmann, Benjamin Elsner, Andreas Peichl. 2017. Tax Refunds and Income Manipulation Evidence from the EITC. SSRN Electronic Journal. [Crossref]
- 50. Maggie R. Jones, Amy B. O'Hara. 2016. Do Doubled-Up Families Minimize Household-Level Tax Burden?. *National Tax Journal* **69**:3, 613-640. [Crossref]
- 51. Nadja Dwenger, Henrik Kleven, Imran Rasul, Johannes Rincke. 2016. Extrinsic and Intrinsic Motivations for Tax Compliance: Evidence from a Field Experiment in Germany. American Economic Journal: Economic Policy 8:3, 203-232. [Abstract] [View PDF article] [PDF with links]

- 52. Olivier Armantier, Scott Nelson, Giorgio Topa, Wilbert van der Klaauw, Basit Zafar. 2016. The Price Is Right: Updating Inflation Expectations in a Randomized Price Information Experiment. *Review of Economics and Statistics* 98:3, 503-523. [Crossref]
- 53. Jeremy G. Moulton, Alexandra Graddy-Reed, Lauren Lanahan. 2016. Beyond the EITC: The Effect of Reducing the Earned Income Tax Credit on Labor Force Participation. *National Tax Journal* 69:2, 261-284. [Crossref]
- 54. Brandon Lehr. 2016. Information and Inflation: An Analysis of Grading Behavior. *The B.E. Journal of Economic Analysis & Policy* **16**:2, 755-783. [Crossref]
- 55. Caroline Weber. 2016. Does the Earned Income Tax Credit Reduce Saving by Low-Income Households?. *National Tax Journal* 69:1, 41-76. [Crossref]
- 56. Julie Shi. 2016. Income Responses to Health Insurance Subsidies: Evidence from Massachusetts. American Journal of Health Economics 2:1, 96-124. [Crossref]
- 57. Hirofumi Kurokawa, Tomoharu Mori, Fumio Ohtake. 2016. A Choice Experiment on Taxes: Are Income and Consumption Taxes Equivalent?. SSRN Electronic Journal. [Crossref]
- 58. Andrew T. Hayashi. 2016. The Effects of Refund Anticipation Loans on Tax Filing and EITC Takeup. SSRN Electronic Journal. [Crossref]
- 59. Saurabh Bhargava, Dayanand Manoli. 2015. Psychological Frictions and the Incomplete Take-Up of Social Benefits: Evidence from an IRS Field Experiment. *American Economic Review* **105**:11, 3489-3529. [Abstract] [View PDF article] [PDF with links]
- 60. Andrea Szabó, Gergely Ujhelyi. 2015. Reducing nonpayment for public utilities: Experimental evidence from South Africa. *Journal of Development Economics* 117, 20-31. [Crossref]
- 61. Jacob Goldin. 2015. Optimal tax salience. Journal of Public Economics 131, 115-123. [Crossref]
- 62. Johannes Abeler, Simon Jäger. 2015. Complex Tax Incentives. *American Economic Journal: Economic Policy* 7:3, 1-28. [Abstract] [View PDF article] [PDF with links]
- 63. Raj Chetty. 2015. Behavioral Economics and Public Policy: A Pragmatic Perspective. *American Economic Review* 105:5, 1-33. [Abstract] [View PDF article] [PDF with links]
- 64. Jeffrey B. Liebman, Erzo F. P. Luttmer. 2015. Would People Behave Differently If They Better Understood Social Security? Evidence from a Field Experiment. *American Economic Journal: Economic Policy* 7:1, 275-299. [Abstract] [View PDF article] [PDF with links]
- 65. Jianbo Luo. 2015. In-Work Benefit, Labor Supply and Optimal Labor Income Taxation: The Fixed Cost of Living and Its Applications. SSRN Electronic Journal . [Crossref]
- 66. Alisa Tazhitdinova. 2015. Behavioral Responses to Payroll and Income Taxes in the UK. SSRN Electronic Journal. [Crossref]
- 67. Sebastien Bradley, Naomi E. Feldman. 2015. Hidden Baggage: Behavioral Responses to Changes in Airline Ticket Tax Disclosure. *SSRN Electronic Journal*. [Crossref]
- 68. George Loewenstein, Cass R. Sunstein, Russell Golman. 2014. Disclosure: Psychology Changes Everything. *Annual Review of Economics* **6**:1, 391-419. [Crossref]
- Koichiro Ito. 2014. Do Consumers Respond to Marginal or Average Price? Evidence from Nonlinear Electricity Pricing. American Economic Review 104:2, 537-563. [Abstract] [View PDF article] [PDF with links]
- 70. Sungmun Choi. 2014. Does Past Experience Affect Future Behavior? Evidence from Estate Tax Avoidance Behavior. SSRN Electronic Journal . [Crossref]
- 71. Kartik Athreya, Devin Reilly, Nicole B. Simpson. 2014. Single Mothers and the Earned Income Tax Credit: Insurance Without Disincentives?. SSRN Electronic Journal. [Crossref]

- 72. Charles Chuan He. 2014. Innovating to Equality: The Egalitarian Distribution of Welfare from the Rise of Consumer Electronics. *SSRN Electronic Journal* . [Crossref]
- 73. Raj Chetty,, John N. Friedman,, Emmanuel Saez. 2013. Using Differences in Knowledge Across Neighborhoods to Uncover the Impacts of the EITC on Earnings. *American Economic Review* 103:7, 2683-2721. [Abstract] [View PDF article] [PDF with links]
- 74. Tuomas Kosonen, Olli Tapani Ropponen. 2013. The Role of Information in Tax Compliance: Evidence from a Natural Field Experiment. SSRN Electronic Journal . [Crossref]
- 75. Cait Poynor Lamberton, Jan-Emmanuel De Neve, Michael I. Norton. 2013. Eliciting Taxpayer Preferences Increases Tax Compliance. SSRN Electronic Journal. [Crossref]
- 76. Jacob Goldin. 2012. Optimal Tax Salience. SSRN Electronic Journal . [Crossref]