

CAT 1

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Q1

$$\binom{10-1}{5-1} \binom{1}{2}^5 \binom{1}{2}^5 = \frac{126}{1024} = \frac{63}{512}$$

Q2

a

Equating $E(X) = \left(\frac{1}{n} \sum_{i=1}^n x_i \right) = \bar{X}$

$$\bar{X} = \theta \sqrt{\frac{\pi}{2}}$$

Making θ the subject of the formula:

$$\hat{\theta} = \bar{X} \sqrt{\frac{2}{\pi}}$$

Checking for Bias $E(\hat{\theta}) = \theta$

$$\begin{aligned} E(\hat{\theta}) &= \bar{X} \sqrt{\frac{2}{\pi}} = \sqrt{\frac{2}{\pi}} E(\bar{X}) = \frac{1}{n} \sqrt{\frac{2}{\pi}} \sum_{i=1}^n E(x_i) \\ &= \frac{1}{n} \sqrt{\frac{2}{\pi}} n E(x_i) \\ &= \frac{1}{n} \sqrt{\frac{2}{\pi}} \theta \sqrt{\frac{\pi}{2}} = \theta \end{aligned}$$

The estimator is unbiased

b

$$\begin{aligned} \widehat{\sigma^2} &= \frac{\hat{\theta}^2}{2} (4 - \pi) = \frac{2\bar{X}^2}{2\pi} (4 - \pi) = \frac{\bar{X}^2}{\pi} (4 - \pi) \\ E(\widehat{\sigma^2}) &= \left(\frac{4 - \pi}{\pi} \right) E(\bar{X}^2) \\ &= \left(\frac{4 - \pi}{\pi} \right) \left(\text{Var}(\bar{X}) + (E(\bar{X}))^2 \right) \\ &= \left(\frac{4 - \pi}{\pi} \right) \left(\frac{\theta^2(4 - \pi)}{2n} + \frac{\theta^2\pi}{2} \right) \end{aligned}$$

Q3

a

The population mean and variance of gamma distribution is $\mu = \alpha\beta$ and $\sigma^2 = \alpha\beta^2$ respectively. The sample equivalent of the population sample mean and variance is $\mu = \bar{X}$ and $\sigma^2 = S^2$

getting the value of β :

$$\alpha\beta = \bar{X}$$

$$\alpha = \frac{\bar{X}}{\beta}$$

Substituting α in S^2

$$S^2 = \alpha\beta^2$$

$$S^2 = \frac{\bar{X}}{\beta}\beta^2 = \bar{X}\beta$$

Get β estimate by making it the subject of the formula

$$\hat{\beta} = \frac{S^2}{\bar{X}}$$

Getting the estimator of β

$$\alpha = \frac{\bar{X}}{\beta}$$

Substituting β with the estimate $\hat{\beta}$

$$\alpha = \frac{\bar{X}}{\frac{S^2}{\bar{X}}} = \frac{\bar{X}^2}{S^2}$$

b

Gamma distribution pdf is given by:

$$f(x|\alpha, \beta) = \frac{\beta^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\beta x}$$

The likelihood function is defined by:

$$L(f(x|\alpha, \beta)) = \ln \left(\prod_{i=1}^n \frac{\beta^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\beta x} \right)$$

$$L(f(x|\alpha, \beta)) = \ln \left(\frac{\beta^{n\alpha}}{(\Gamma(\alpha))^n} \prod_{i=1}^n x_i^{\alpha-1} e^{-\sum_{i=1}^n x_i \beta} \right)$$

$$\ln L = n\alpha \ln \beta - n \ln(\Gamma(\alpha)) + (\alpha - 1) \sum_{x=i}^n \ln x - \sum_{x=i}^n x_i \beta$$

First partial derivative with respect to β

$$\frac{d \ln L}{d \beta} = \frac{n\alpha}{\beta} - \sum_{x=i}^n x_i = 0$$

Solve for β

$$\begin{aligned}\frac{\alpha n}{\beta} &= \sum_{x=i}^n x_i \\ \frac{1}{\beta} &= \frac{\sum_{x=i}^n x_i}{\alpha n} \\ \hat{\beta} &= \frac{\alpha}{\frac{1}{n} \sum_{x=i}^n x_i} = \frac{\alpha}{\bar{X}}\end{aligned}$$

First partial derivative with respect to α

$$\begin{aligned}\frac{d \ln L}{d \alpha} &= n \ln \beta - \frac{n \Gamma \alpha'}{\Gamma \alpha} + \sum_{i=1}^n \ln x_i = 0 \\ &= \frac{n \Gamma \alpha'}{\Gamma \alpha} = n \ln \beta + \sum_{i=1}^n \ln x_i\end{aligned}$$

Dividing every term by n:

$$= \frac{\Gamma \alpha'}{\Gamma \alpha} = \ln \beta + \frac{\sum_{i=1}^n \ln x_i}{n}$$

c

$$\hat{\beta} = \frac{S^2}{\bar{X}}$$

Calculate \bar{X} :

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n x_i = \frac{168.5}{25} = 6.74$$

$$\begin{aligned}S^2 &= 0.46166667 \\ \hat{\beta} &= \frac{0.46166667}{6.74} = 0.06849\end{aligned}$$

$$\begin{aligned}\alpha &= \left(\frac{\bar{X}}{S} \right)^2 \\ \alpha &= \left(\frac{6.74}{0.46166667} \right)^2 = 213.1389\end{aligned}$$

Q4

a

To predict weight, Weight is placed on the y-axis and height is placed in the x-axis. The values of weight are varied as compared to the height of a person.

b

```
dataset = read.csv(file="dataset/body_composition.csv")
```

```
plot(dataset$height,dataset$weight,main ="Scatterplot of Men Weight and Height ",xlab = "Height (inches)
```



c

```
plot(
  dataset$height,
  dataset$weight,
  main ="Scatterplot of Weight of Men Against Their Height ",
  xlab = "Height (inches)",
  ylab = "Weight (lbs)"
)

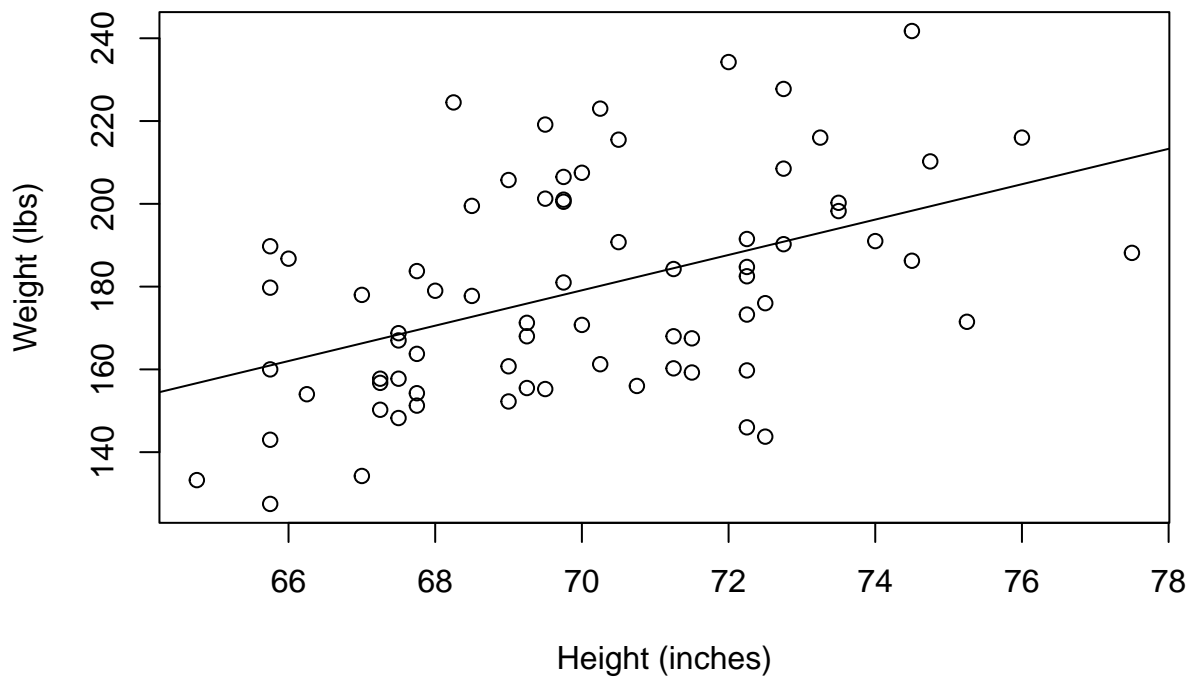
lm_model = lm(weight~ height , data=dataset)
summary(lm_model)
```

```
##
## Call:
## lm(formula = weight ~ height, data = dataset)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -46.04  -17.97   -4.31   17.93   52.87
##
```

```
## Coefficients:
##             Estimate Std. Error t value Pr(>|t|)
## (Intercept) -119.9317   68.3942  -1.754   0.0839 .
## height       4.2720    0.9758   4.378 4.1e-05 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 23.19 on 70 degrees of freedom
## Multiple R-squared:  0.2149, Adjusted R-squared:  0.2037
## F-statistic: 19.17 on 1 and 70 DF,  p-value: 4.105e-05

abline(lm_model, v=0)
```

Scatterplot of Weight of Men Against Their Height



The Linear Regression equation for height and weight is:

$$Weight = -119.9317 + 4.272 \times Height + \epsilon$$

d

A slope of 4.272 represent the estimated change in weight for every one inch of height. The slope is positive hence there is a positive linear relationship between weight and height. it is not appropriate to make a interpretation when the height is zero since the no person that has a height of 0.

e

```
predict(lm_model,newdata = data.frame(height=c(76)))
```

```
##          1  
## 204.7376
```

Q5

$$MSE(\overline{Y}_{str}) = 0.36 \frac{\sigma_A^2}{60} + 0.16 \frac{\sigma_B^2}{40}$$
$$MSE(\overline{Y}) = 0.6 \frac{\sigma_A^2}{100} + 0.4 \frac{\sigma_b^2}{40} + 0.24 \frac{(\mu_A - \mu_B)^2}{100}$$

Q6

Given

$$X_i \sim N(\theta, \sigma^2)$$
$$E(X_i) = \theta$$
$$V(X_i) = \sigma^2$$

a

Y is an unbiased estimator of θ , the $E(Y) = \theta$

$$E(Y) = \left(\frac{E(X_1 + X_2)}{2} \right)$$
$$= \frac{1}{2}E(X_1) + \frac{1}{2}E(X_2)$$
$$= \frac{1}{2}\theta + \frac{1}{2}\theta$$
$$= \theta$$

Y is an unbiased estimator of θ

b

$$Var(Y) \geq \frac{-1}{n \int_{-\infty}^{+\infty} \left(\frac{\partial^2 \ln f(x; \theta)}{\partial \theta^2} \right) f(x; \theta) dx}$$

T The normal p.d.f with $\mu = \theta$

$$f(x; \theta) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x - \theta)^2}{2\sigma^2}\right)$$

Take the logarithm of the probability density function

$$\ln f(x; \theta) = \ln \left(\frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x - \theta)^2}{2\sigma^2}\right) \right)$$

$$= -\frac{1}{2} \ln(2\pi\sigma^2) - \frac{(x-\theta)^2}{2\sigma^2}$$

Determine the second derivative with respect to θ

$$\begin{aligned}\frac{\partial \ln f(x; \theta)}{\partial \theta} &= \frac{-2(x-\theta)}{2\sigma^2} \\ &= \frac{x-\theta}{\sigma^2} \\ \frac{\partial^2 \ln f(x; \theta)}{\partial^2 \theta} &= \frac{1}{\sigma^2}\end{aligned}$$

Determining the Rao-Cramer lower bound:

$$\begin{aligned}\int_{-\infty}^{+\infty} \left(\frac{\partial^2 \ln f(x; \theta)}{\partial^2 \theta} \right) f(x; \theta) dx &= \int_{-\infty}^{+\infty} \frac{1}{\sigma^2} f(x; \theta) dx \\ &= -\frac{1}{\sigma^2} \int_{-\infty}^{+\infty} f(x; \theta) dx \\ &= -\frac{1}{\sigma^2} \cdot 1 = -\frac{1}{\sigma^2}\end{aligned}$$

Solving Rao-Cramer lower bound:

$$\begin{aligned}Var(Y) &\geq \frac{-1}{n \int_{-\infty}^{+\infty} \left(\frac{\partial^2 \ln f(x; \theta)}{\partial^2 \theta} \right) f(x; \theta) dx} \\ &= \frac{-1}{n \left(-\frac{1}{\sigma^2} \right)} \\ &= \frac{\sigma^2}{n} \\ Var(Y) &\geq \frac{\sigma^2}{n}\end{aligned}$$

c

Determining the variance of Y

$$\begin{aligned}Var(Y) &= Var\left(\frac{X_1 + X_2}{2}\right) \\ &= \left(\frac{1}{2}\right) Var(X_1) + \left(\frac{1}{2}\right) Var(X_2) \\ &= \frac{1}{4}\sigma^2 + \frac{1}{4}\sigma^2 \\ &= \frac{\sigma^2}{2}\end{aligned}$$

Efficiency is the Rao-Cramer lower bound to the actual variance of the random variable

$$\begin{aligned}Efficiency &= \frac{Rao - Cramer lower bound of Y}{Var(Y)} \\ &= \frac{\sigma^2/n}{\sigma^2/2} \\ &= \frac{2}{n}\end{aligned}$$