

STAT 600 - HW 2

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All Rcpp/RcppArmadillo can be found in my GitHub.

Question 1

(a)

First, consider the likelihood and the log-likelihood function.

$$\begin{aligned} L(\theta) &= \prod_{i=1}^n \frac{1}{\pi(1 + (x_i - \theta)^2)} \\ l(\theta) &= \log(L(\theta)) \\ &= \log\left(\prod_{i=1}^n \frac{1}{\pi(1 + (x_i - \theta)^2)}\right) \\ &= \sum_{i=1}^n \log\left(\frac{1}{\pi(1 + (x_i - \theta)^2)}\right) \\ &= -\sum_{i=1}^n \log(\pi(1 + (x_i - \theta)^2)) \\ &= -n \log(\pi) - \sum_{i=1}^n \log(1 + (x_i - \theta)^2) \end{aligned}$$

Then, consider the derivative of the log-likelihood, $l'(\theta)$.

$$\begin{aligned} l'(\theta) &= \frac{d}{d\theta} l(\theta) \\ &= -\sum_{i=1}^n \frac{1}{1 + (x_i - \theta)^2} \left[\frac{d}{d\theta} (x_i - \theta)^2 \right] \\ &= 2 \sum_{i=1}^n \frac{x_i - \theta}{1 + (x_i - \theta)^2} \end{aligned}$$

```
dat <- c(-8.86, -6.82, -4.03, -2.84, 0.14, 0.19, 0.24, 0.27, 0.49, 0.62, 0.76, 1.09,
        1.18, 1.32, 1.36, 1.58, 1.58, 1.78, 2.13, 2.15, 2.36, 4.05, 4.11, 4.12,
        6.83)
rangeTheta <- seq(-10, 10, 0.01)
data.frame(theta = rangeTheta, dll = sapply(rangeTheta, dloglik, x = dat)) %>%
  ggplot(aes(x = theta, y = dll)) +
  geom_line() +
```

```
theme_bw() +
labs(x = TeX("\\theta"), y = TeX("l'\\theta"),
title = "The plot of the derivative of the log-likelihood based on 25 daat points")
```

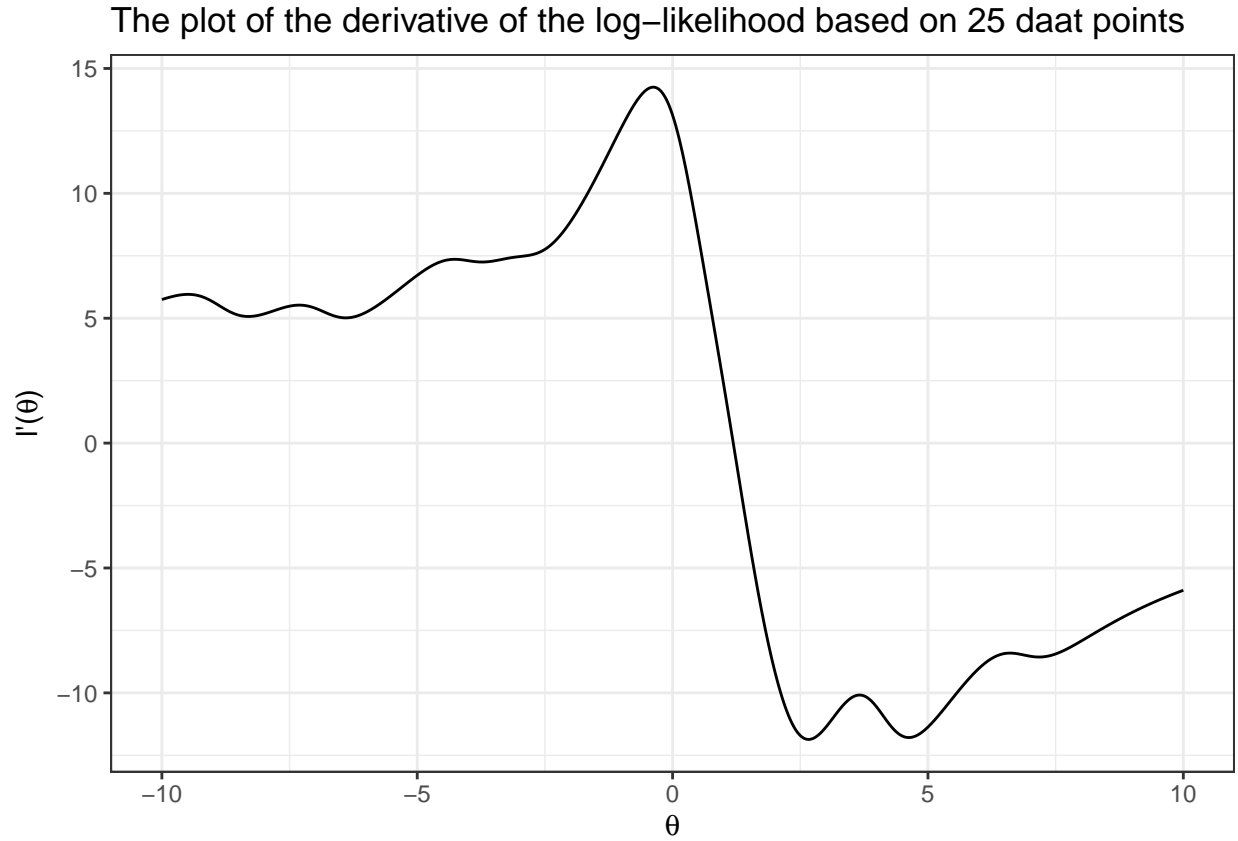


Figure 1: The plot of the derivative of the log-likelihood on the original dataset.

(b)

This is the second derivation for the log-likelihood function.

$$\begin{aligned}
 l''(\theta) &= \frac{d}{d\theta} l'(\theta) \\
 &= \frac{d}{d\theta} 2 \sum_{i=1}^n \frac{x_i - \theta}{1 + (x_i - \theta)^2} \\
 &= 2 \sum_{i=1}^n \frac{d}{d\theta} \frac{x_i - \theta}{1 + (x_i - \theta)^2} \\
 &= 2 \sum_{i=1}^n \frac{-1 + (x_i - \theta)^2}{(1 + (x_i - \theta)^2)^2}
 \end{aligned}$$

(c)

Below are the result from each methods. Note that I have set the ϵ to be 1×10^{-5} .

```
### Run all methods
eps_set <- 1e-5
bs_dat <- bisect_q1(min(dat), max(dat), dat, eps = eps_set)
nr_dat <- nr_q1(x0 = 0, dat = dat, eps = eps_set)
fs_dat <- fs_q1(x0 = 0, dat = dat, eps = eps_set)
sc_dat <- sc_q1(x0 = 0, x1 = 1e-5, dat = dat, eps = eps_set)

### Create the table
data.frame(theta = c(bs_dat$xt, nr_dat$xt, fs_dat$xt, sc_dat$xt),
            iter = c(bs_dat$n_iter, nr_dat$n_iter, fs_dat$n_iter, sc_dat$n_iter)) %>%
  `rownames<-`(c("Bisection", "Newton-Raphson", "Fisher Scoring", "Secant Method")) %>%
  kable(digits = 5, col.names = c("$\\hat{\\theta}$", "Number of iteration"),
        caption = "The result from each methods with only 25 observations.")
```

Table 1: The result from each methods with only 25 observations.

	$\hat{\theta}$	Number of iteration
Bisection	1.18795	20
Newton-Raphson	1.18795	5
Fisher Scoring	1.18795	5
Secant Method	1.18794	5

(d)

For the convergence criteria used in this problem, I decided to employ the absolute convergence criterion, as the $x^{(t)}$ might be close to 0 in some iterations, as indicated by the plot shown in part (a). Additionally, the value of x is neither too tiny nor too huge compared to ϵ .

(e)

(f)

(g)

```
add_dat <- c(-8.34, -1.73, -0.40, -0.24, 0.60, 0.94, 1.05, 1.06, 1.45, 1.50,
            1.54, 1.72, 1.74, 1.88, 2.04, 2.16, 2.39, 3.01, 3.01, 3.08, 4.66,
            4.99, 6.01, 7.06, 25.45)

### Run all methods with complete data
eps_set <- 1e-5
bs_cdat <- bisect_q1(min(c(dat, add_dat)), max(c(dat, add_dat)), dat = c(dat, add_dat), eps = eps_set)
nr_cdat <- nr_q1(x0 = 0.5, dat = c(dat, add_dat), eps = eps_set)
fs_cdat <- fs_q1(x0 = 0.5, dat = c(dat, add_dat), eps = eps_set)
sc_cdat <- sc_q1(x0 = 0, x1 = 0.5, dat = c(dat, add_dat), eps = eps_set)

### Create the table
```

```
data.frame(theta = c(bs_cdat$xt, nr_cdat$xt, fs_cdat$xt, sc_cdat$xt),
            iter = c(bs_cdat$n_iter, nr_cdat$n_iter, fs_cdat$n_iter, sc_cdat$n_iter)) %>%
  `rownames<-`(c("Bisection", "Newton-Raphson", "Fisher Scoring", "Secant Method")) %>%
  kable(digits = 5, col.names = c("$\\hat{\\theta}$", "Number of iteration"),
        caption = "The result from each methods with all 50 observations.")
```

Table 2: The result from each methods with all 50 observations.

	$\hat{\theta}$	Number of iteration
Bisection	1.47131	21
Newton-Raphson	1.47130	5
Fisher Scoring	1.47130	5
Secant Method	1.47130	5

Question 2

Question 3

(a)

I will denote $\beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2}$ as $\mathbf{x}_i \boldsymbol{\beta}$. Since we know that $Y_i \sim \text{Ber} \left(\frac{\exp(\mathbf{x}_i \boldsymbol{\beta})}{1 + \exp(\mathbf{x}_i \boldsymbol{\beta})} \right)$, then the likelihood and the log-likelihood can be derived as below.

$$\begin{aligned}
 L(\boldsymbol{\beta}) &= \prod_{i=1}^n \left[\frac{\exp(\mathbf{x}_i \boldsymbol{\beta})}{1 + \exp(\mathbf{x}_i \boldsymbol{\beta})} \right]^{y_i} \left[1 - \frac{\exp(\mathbf{x}_i \boldsymbol{\beta})}{1 + \exp(\mathbf{x}_i \boldsymbol{\beta})} \right]^{1-y_i} \\
 &= \prod_{i=1}^n \frac{\exp(\mathbf{x}_i \boldsymbol{\beta})^{y_i}}{1 + \exp(\mathbf{x}_i \boldsymbol{\beta})} \\
 l(\boldsymbol{\beta}) &= \log(L(\boldsymbol{\beta})) \\
 &= \sum_{i=1}^n [y_i (\mathbf{x}_i \boldsymbol{\beta}) - \log(1 + \exp(\mathbf{x}_i \boldsymbol{\beta}))]
 \end{aligned}$$

(b)

First, consider the first derivative of the log-likelihood w.r.t. $\boldsymbol{\beta}$, or the gradient.

$$\begin{aligned}
 \nabla_{\boldsymbol{\beta}} l(\boldsymbol{\beta}) &= \frac{d}{d\boldsymbol{\beta}} l(\boldsymbol{\beta}) \\
 &= \sum_{i=1}^n \left[y_i \frac{d}{d\boldsymbol{\beta}} \mathbf{x}_i \boldsymbol{\beta} - \frac{d}{d\boldsymbol{\beta}} \log(1 + \exp(\mathbf{x}_i \boldsymbol{\beta})) \right] \\
 &= \sum_{i=1}^n \left[y_i \mathbf{x}_i^T - \frac{\exp(\mathbf{x}_i \boldsymbol{\beta})}{1 + \exp(\mathbf{x}_i \boldsymbol{\beta})} \mathbf{x}_i^T \right] \\
 &= \sum_{i=1}^n \left[y_i - \frac{\exp(\mathbf{x}_i \boldsymbol{\beta})}{1 + \exp(\mathbf{x}_i \boldsymbol{\beta})} \right] \mathbf{x}_i^T
 \end{aligned}$$

We can rewrite the formula above in a matrix form as $\nabla_{\beta} l(\beta) = \mathbf{X}^T (\mathbf{Y} - \hat{\mathbf{Y}})$, where $\hat{\mathbf{Y}}$ is a vector consisted of $\frac{\exp(\mathbf{x}_i^T \beta)}{1 + \exp(\mathbf{x}_i^T \beta)}$ since we can think this quantity as a predicted probability of success for the observation i .

Now, we will consider the Hessian for the log-likelihood.

$$\begin{aligned} H(\beta) &= \nabla_{\beta} (\nabla_{\beta} l(\beta)) \\ &= \nabla_{\beta} \sum_{i=1}^n \left[y_i - \frac{\exp(\mathbf{x}_i^T \beta)}{1 + \exp(\mathbf{x}_i^T \beta)} \right] \mathbf{x}_i^T \\ &= - \sum_{i=1}^n \mathbf{x}_i^T \mathbf{x}_i \frac{\exp(\mathbf{x}_i^T \beta)}{(1 + \exp(\mathbf{x}_i^T \beta))^2} \end{aligned}$$

Similarly, we can rewrite the Hessian matrix in the matrix form as $H(\beta) = \mathbf{X}^T \mathbf{W} \mathbf{X}$, where \mathbf{W} is a matrix consisted of $-\frac{\exp(\mathbf{x}_i^T \beta)}{(1 + \exp(\mathbf{x}_i^T \beta))^2}$ as a diagonal while the off-diagonal are 0.

By applying the Newton-Raphson, we will update the parameters for the iteration t by using $\beta^{(t)} = \beta^{(t-1)} - \left[H(\beta^{(t-1)}) \right]^{-1} \left[\nabla_{\beta} l(\beta^{(t-1)}) \right] = \beta^{(t-1)} - \left[\mathbf{X}^T \mathbf{W} \mathbf{X} \right]^{-1} \mathbf{X}^T (\mathbf{Y} - \hat{\mathbf{Y}})$.

```
### Data
#### yi, intercept, x1 (coffee assumption), x2 (gender)
designMat <- rbind(c(1, 1, 0, 1), c(0, 1, 0, 1), c(1, 1, 2, 1), c(0, 1, 2, 1),
  c(1, 1, 4, 1), c(0, 1, 4, 1), c(1, 1, 5, 1), c(0, 1, 5, 1),
  c(1, 1, 0, 0), c(0, 1, 0, 0), c(1, 1, 2, 0), c(0, 1, 2, 0),
  c(1, 1, 4, 0), c(0, 1, 4, 0), c(1, 1, 5, 0), c(0, 1, 5, 0)) %>%
  as.matrix()

repTime <- c(9, 41 - 9, 94, 213 - 94, 53, 127 - 53, 60, 142 - 60,
  11, 67 - 11, 59, 211 - 59, 53, 133 - 53, 28, 76 - 28)

designMat <- designMat[rep(1:nrow(designMat), times = repTime), ]

### Run the optimization
resultQ3 <- optimQ3(desMat = designMat[, -1], Y = designMat[, 1],
  b0 = c(0, 0, 0), eps = 1e-10)
resultQ3$bt

##           [,1]
## [1,] -1.1877905
## [2,]  0.1382988
## [3,]  0.3972517
```

(d)

```
est_SD <- sqrt(diag(solve(t(designMat[, -1]) %*% resultQ3$W %*% designMat[, -1])))
z_stat <- (resultQ3$bt - 0)/est_SD

### Compare with z-statistics
(qnorm(0.05/2) <= z_stat) & (z_stat <= qnorm(1 - (0.05/2))) ## If TRUE, FTR H0
```

```
##      [,1]
## [1,] FALSE
## [2,] FALSE
## [3,] FALSE
```