# STAT 528 - Advanced Regression Analysis II

#### GLMM and GEE

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# Learning Objectives Today

- ► GLMM examples
- ► GEE theory
- ► GEE examples

We load in necessary packages.

```
library(faraway)
library(tidyverse)
library(ggplot2)
library(MASS)
library(lme4)
library(INLA)
library(glmm)
```

In this example, we have data from a clinical trial of 59 epileptics.

For a baseline, patients were observed for 8 weeks and the number of seizures recorded. The patients were then randomized to treatment by the drug Progabide (31 patients) or to the placebo group (28 patients).

They were observed for four 2-week periods and the number of seizures recorded. We are interested in determining whether Progabide reduces the rate of seizures.

We first perform some data manipulations and then look at the first few observations:

```
data(epilepsy, package="faraway")
epilepsy$period <- rep(0:4, 59)
epilepsy$drug <- factor(c("placebo","treatment")[epilepsy$treat+1])
epilepsy$phase <- factor(c("baseline","experiment")[epilepsy$expind +1])
epilepsy %>% filter(id < 2.5) %>% head(3)
```

The variables are expind variable indicates the baseline phase by 0 and the treatment phase by 1. The length of these time phases is recorded in the timeadj variable.

Three new convenience variables are created: period, denoting the 2- or 8- week periods, drug recording the type of treatment in nonnumeric form and phase indicating the phase of the experiment.

We now compute the mean number of seizures per week broken down by the treatment and baseline vs. experimental period.

```
epilepsy %>%
  group_by(drug, phase) %>%
  summarise(rate=mean(seizures/timeadj)) %>%
  xtabs(formula=rate ~ phase + drug)

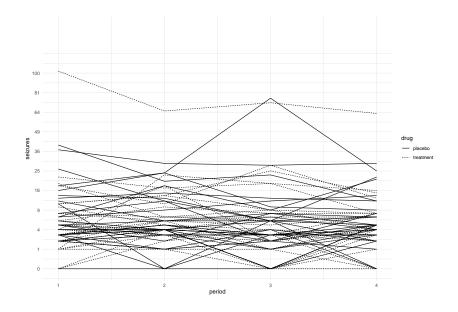
## drug
## objector treatment
```

baseline 3.848214 3.955645 experiment 4.303571 3.983871

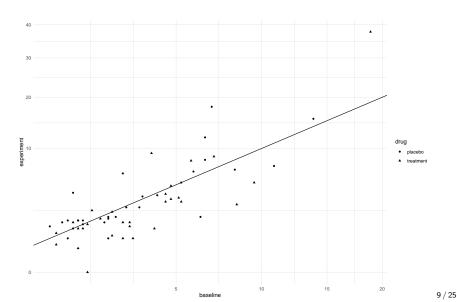
We see that the rate of seizures in the treatment group actually increases during the period in which the drug was taken. The rate of seizures increases even more in the placebo group.

Perhaps some other factor is causing the rate of seizures to increase during the treatment period and the drug is actually having a beneficial effect.

Now we make some plots to show the difference between the treatment and the control. The first plot shows the difference between the two groups during the experimental period only:



We now compare the average seizure rate to the baseline for the two groups. The square-root transform is used to stabilize the variance; this is often used with count data.



A treatment effect, if one exists, is not readily apparent. Now we fit GLMM models. Patient #49 is unusual because of the high rate of seizures observed. We exclude it:

```
epilo <- filter(epilepsy, id != 49)
```

Excluding a case should not be taken lightly. For projects where the analyst works with producers of the data, it will be possible to discuss substantive reasons for excluding cases.

It is worth starting with a GLM even though the model is not correct due to the grouping of the observations. We must use an offset to allow for the difference in lengths in the baseline and treatment periods.

```
I(expind * treat), family = poisson, data = epilo)
##
##
## Deviance Residuals:
              10 Median
##
      Min
                              3Q
                                     Max
## -5.4725 -2.3605 -1.0290 0.9001 14.0104
##
## Coefficients:
                 Estimate Std. Error z value Pr(>|z|)
##
## (Intercept)
                 1.34761 0.03406 39.566 < 2e-16 ***
## expind
                 0.11184 0.04688 2.386 0.017 *
                -0.10682 0.04863 -2.197 0.028 *
## treat
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for poisson family taken to be 1)
##
      Null deviance: 2485.1 on 289 degrees of freedom
##
## Residual deviance: 2411.5 on 286 degrees of freedom
## ATC: 3449 7
## Number of Fisher Scoring iterations: 5
```

The interaction term is the primary parameter of interest. All the subjects were untreated in the baseline. This means that the main effect for treatment does not properly measure the response to treatment because it includes the baseline period.

As we have observed already, we suspect the response may have been different during the baseline time and the active period of the experiment. The interaction term represents the effect of the treatment during the baseline period after adjustment. In the output above we see that this interaction seems highly significant and negative (which is good since we want to reduce seizures).

But this inference is suspect because we have made no allowance for the correlated responses within individuals. The p-value is far smaller than it should be.

### PQL methods

## Number of Groups: 58

```
modpql <- glmmPQL(seizures ~offset(log(timeadj)) + expind + treat +
 I(expind*treat), random = ~1|id, family=poisson, data=epilo)
summary(modpql)
## Linear mixed-effects model fit by maximum likelihood
## Data: epilo
## AIC BIC logLik
   NA NA NA
##
## Random effects:
## Formula: ~1 | id
         (Intercept) Residual
## StdDev: 0.6820012 1.605385
## Variance function:
## Structure: fixed weights
## Formula: ~invwt
## Fixed effects: seizures ~ offset(log(timeadj)) + expind + treat + I(expind *
                                                                                    treat)
                         Value Std.Error DF t-value p-value
##
## (Intercept) 1.0761832 0.09990514 230 10.772050 0.0000
                     0.1125119 0.07412152 230 1.517939 0.1304
## expind
## I(expind * treat) -0.3037615 0.10819095 230 -2.807642 0.0054
## Correlation:
##
                    (Intr) expind
## expind
                    -0.198
## I(expind * treat) -0.014 -0.656
##
## Standardized Within-Group Residuals:
##
                     01
                               Med
                                           03
         Min
                                                    Max
## -2.2934834 -0.5649468 -0.1492931 0.3224895 6.3123337
##
## Number of Observations: 290
```

The parameter estimates from the PQL fit are comparable to the GLM fit. However, the standard errors are larger in the PQL fit as might be expected given that the correlated responses have been allowed for.

As with the binary response example, we still have some doubts about the accuracy of the inference. This is a particular concern when some count responses are small.

## Numerical integration

Numerical quadrature can also be used. We use Gauss-Hermite in preference to Laplace as the model random effect structure is simple and so the computation is fast even though we have used the most expensive nAGQ=25 setting.

```
modgh <- glmer(seizures ~offset(log(timeadj)) + expind + treat +
I(expind*treat)+ (1|id), nAGQ=25, family=poisson, data=epilo)</pre>
```

```
## Generalized linear mixed model fit by maximum likelihood (Adaptive
    Gauss-Hermite Quadrature, nAGQ = 25) [glmerMod]
## Family: poisson (log)
## Formula: seizures ~ offset(log(timeadj)) + expind + treat + I(expind *
      treat) + (1 | id)
##
##
     Data: epilo
##
##
       ATC
               BIC
                   logLik deviance df.resid
##
     877.7
             896 1 -433 9 867 7
                                         285
##
## Scaled residuals:
      Min
             10 Median
                             30
                                   Max
## -3.8724 -0.8482 -0.1722 0.5697 9.8941
##
## Random effects:
## Groups Name
                 Variance Std.Dev.
## id (Intercept) 0.515 0.7176
## Number of obs: 290, groups: id, 58
##
## Fixed effects:
                    Estimate Std. Error z value Pr(>|z|)
##
## (Intercept)
                   1.035998 0.141256 7.334 2.23e-13 ***
## expind
                  0.111838 0.046877 2.386 0.017 *
## treat
                   -0.008152 0.196525 -0.041 0.967
## I(expind * treat) -0.302387  0.069714 -4.338 1.44e-05 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##
            (Intr) expind treat
## expind -0.175
## treat -0.718 0.126
## I(xpnd*trt) 0.118 -0.672 -0.173
```

We see that the interaction effect is significant. Notice that the estimate of this effect has been quite consistent over all the estimation methods so we draw some confidence from this. We have

```
exp(-0.302)
```

## [1] 0.7393381

So the drug is estimated to reduce the rate of seizures by about 26%. However, the subject SD is more than twice the drug effect of -0.3 at 0.718. This indicates that the expected improvement in the drug is substantially less than the variation between individuals.

Interpretation of the main effect terms is problematic in the presence of an interaction. For example, the treatment effect reported here represents the predicted difference in the response during the baseline period (i.e., expind=0).

Since none of the subjects are treated during the baseline period, we are reassured to see that this effect is not significant.

However, this does illustrate the danger in naively presuming that this is the treatment effect.

# Bayesian methods

## quant0.975 1.314419 0.2036471

We can also take a Bayesian approach using INLA.

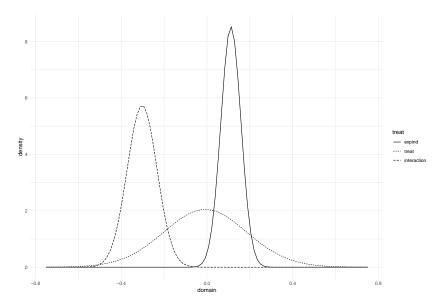
```
formula <- seizures ~ offset(log(timeadj)) + expind + treat +
   I(expind*treat) + f(id,model="iid")
result <- inla(formula, family="poisson", data = epilo)</pre>
```

### We obtain a summary of the posteriors as:

```
sigmaalpha <- inla.tmarginal(function(x) 1/sqrt(x),
 result$marginals.hyperpar$"Precision for id")
restab <- sapply(result$marginals.fixed,
  function(x) inla.zmarginal(x, silent=TRUE))
restab <- cbind(restab.
  inla.zmarginal(sigmaalpha, silent=TRUE))
colnames(restab) = c("mu", "expind", "treat",
  "interaction", "alpha")
data.frame(restab)
                           expind
                                         treat interaction
                                                                alpha
                                                            0.7254792
## mean
             1.036115
                        0.1118965
                                   -0.00825552 -0.3025702
## sd
             0.1421983 0.04685179
                                     0.1978247 0.06967626 0.07183047
## quant0.025 0.7550146 0.01995068
                                    -0.3978168 -0.4393086
                                                              0.59977
## quant0.25 0.9410767 0.08018646
                                     -0.140758 -0.3497282 0.6746726
## quant0.5 1.036183 0.1117989 -0.008691486 -0.3027154
                                                            0.719905
## quant0.75 1.130938 0.1434113
                                      0.123375 -0.2557025
                                                            0.7702153
```

0.380427 -0.1661221

0.8818008



# Monte Carlo likelihood approximation

We can use the glmm package to implement the MCLA approach to fitting GLMM models with Poisson responses.

```
epilo$idF <- as.factor(epilo$id)
epilo$seizures <- as.integer(epilo$seizures)
set.seed(13)
clust <- makeCluster(8)
system.time(m1 <- glmm(seizures -
expind + treat + I(expind*treat), random = list(-0+idF),
family.glmm = poisson.glmm, m = 7e4,
varcomps.names = c("idF"), cluster = clust, data=epilo))</pre>
## user system elapsed
## 6.186 0.632 68.164
```

We obtain summary information. However, the fit is buggy. The Monte Carlo standard error is not returned and the summary table estimates are not depicted.

```
summary(m1)
##
## Call:
## glmm(fixed = seizures ~ expind + treat + I(expind * treat), random = list(~0 +
      idF), varcomps.names = c("idF"), data = epilo, family.glmm = poisson.glmm,
##
      m = 70000. cluster = clust)
##
##
## Link is: "log"
##
## Fixed Effects:
                    Estimate Std. Error z value Pr(>|z|)
##
## (Intercept)
                                     0 31.010 < 2e-16 ***
## expind
                                    0 -27 187 < 2e-16 ***
## treat
                                   0 0.018
                                                  0.986
## I(expind * treat)
                                    0 -4.338 1.44e-05 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
##
## Variance Components for Random Effects (P-values are one-tailed):
      Estimate Std. Error z value Pr(>|z|)/2
## idF
                            5.098 1.72e-07 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

```
# Monte Carlo standard errors
mcse_glmm <- mcse(m1)
mcse_glmm</pre>
```

##	(Intercept)	expind	treat I(expind	* treat)
##	NaN	NaN	NaN	NaN
##	idF			
##	NaN			

That being said, we can obtain estimates of fixed effects and their standard errors from objects in the glmm object.

```
# standard errors
se glmm <- se(m1)
# table for fixed effects
tab <- cbind(m1$beta, se_glmm[-5], m1$beta/se_glmm[-5])
colnames(tab) <- c("Estimate", "Std. Error", "z value")
round(tab, 3)
##
                   Estimate Std. Error z value
## (Intercept)
                     3.106
                              0.100 31.010
## expind
                   -1.274 0.047 -27.187
## treat
                     0.003 0.185 0.018
## I(expind * treat) -0.302 0.070 -4.338
```

We can also obtain estimates of random effect parameters and their standard errors from objects in the glmm object.

```
c(m1$nu, se_glmm[5])

## idF idF
## 0.5152511 0.1010674
```

Parameter estimates are similar to the other fitting techniques which instills confidence.