

CYS 2020/2021

- Invariant distributions
 - Computation in the case of the RW with one reflecting barrier
 - Computation in the case of the success runs
 - Computation in the case of a RW on a graph
 - Detailed balance equation

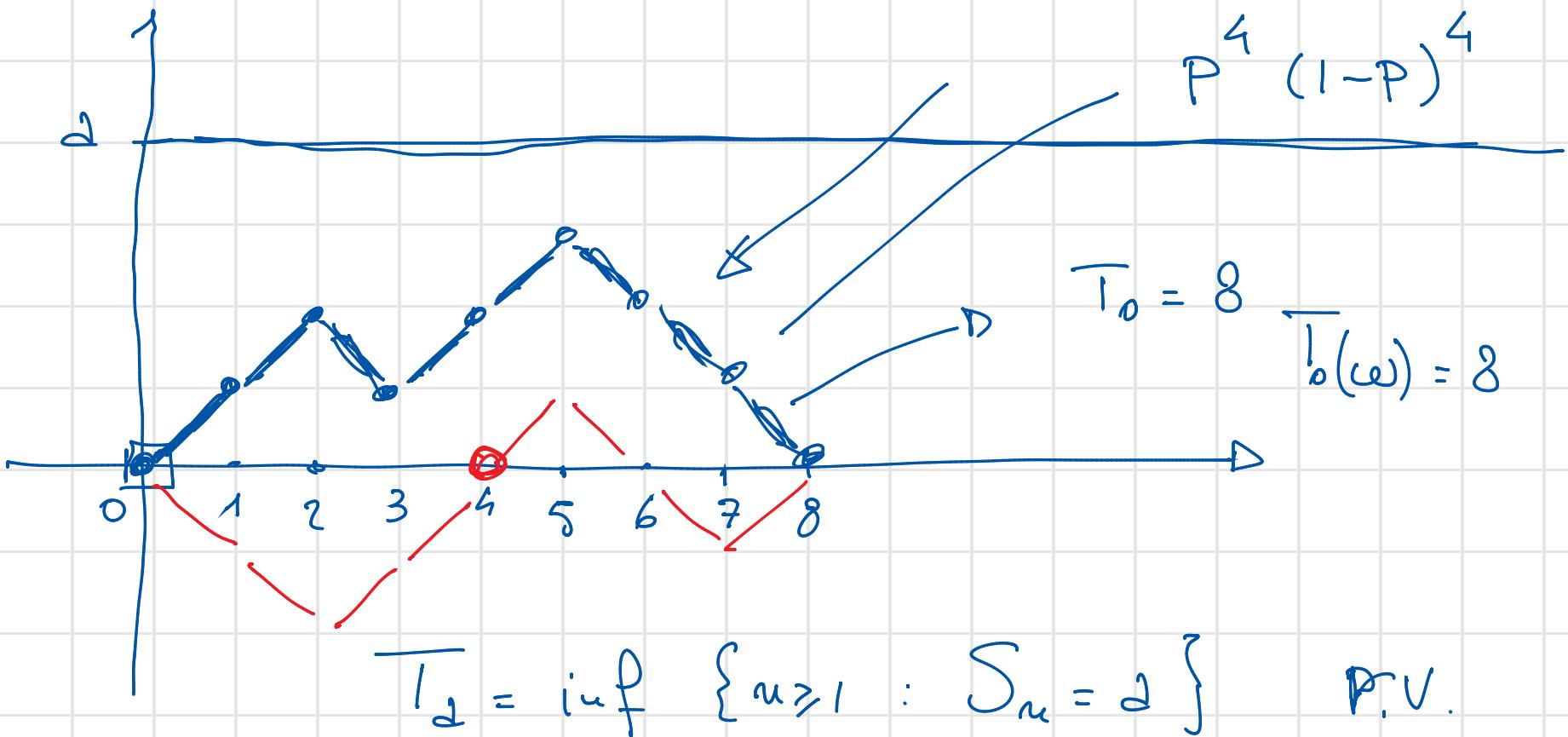


Simple RW on \mathbb{Z}

$$\begin{cases} S_0 = 0 \\ S_n = X_1 + X_2 + \dots + X_n \end{cases}$$

$X_1, n \geq 1$ iid

$$X_1 = \begin{cases} 1 & p \\ -1 & q = 1-p \end{cases}$$



$$\inf \emptyset = +\infty$$

$$Q = 1 - P$$

$$P[\bar{T}_d = \infty] = ? \quad \forall d$$

$$d > 0, \quad d < 0, \quad d = 0$$

$$P[\bar{T}_0 = \infty] = 0 \quad m \text{ is odd}$$

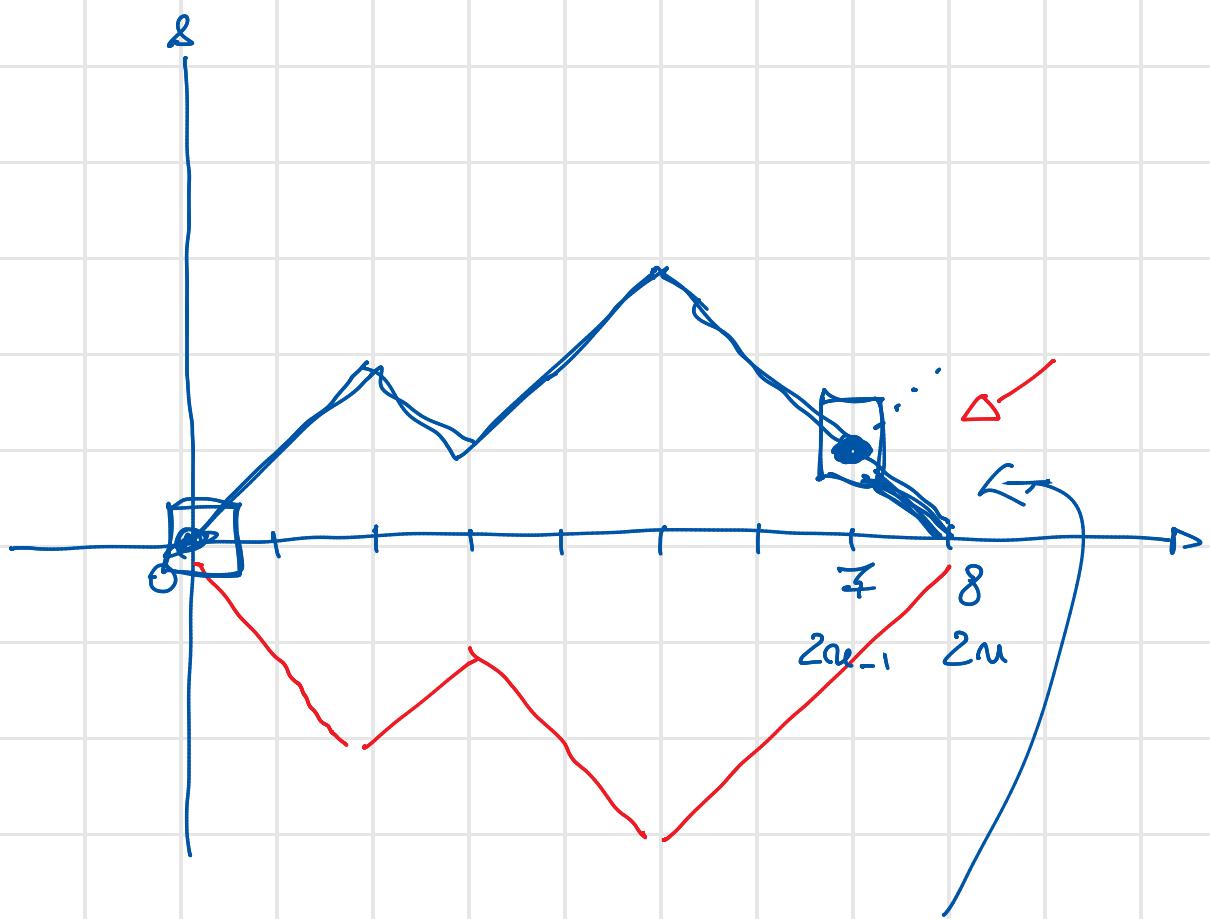
$$P[\bar{T}_0 = 2n] = ?$$

$$[T_0 = 2n]$$

$$P^n Q^n$$

$$P[S_{2n} = 0] = \binom{2n}{n} \cdot P^n \cdot Q^n$$

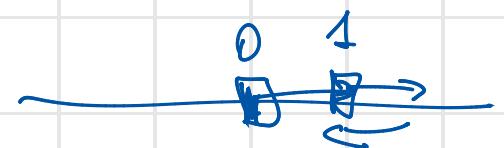
$n (+1)$ among the $2n$ steps



$$n=4$$

$$\overline{T}_{2n} = \overline{T}_8$$

$$2 \quad p^4 q^4$$



$$(0,0) \rightarrow (8,0)$$

without touching 0 before
the 8th step

$$(0,0) \rightarrow (7,1)$$

$$A = \{ \text{all paths that join } (0,0) \text{ with } (2m-1, 1) \}$$

$$B = \{ \text{all paths that join } (0,0) \text{ with } (2m-1, 1) \text{ without touching 0} \}$$

$$A = B \cup C \cup D$$

disjoint

$$|A| = |B| + |C| + |D|$$

$$C = \{ \text{all the paths in } A \setminus B \}$$

$$D = \{ \text{all paths in } A \setminus B \}$$

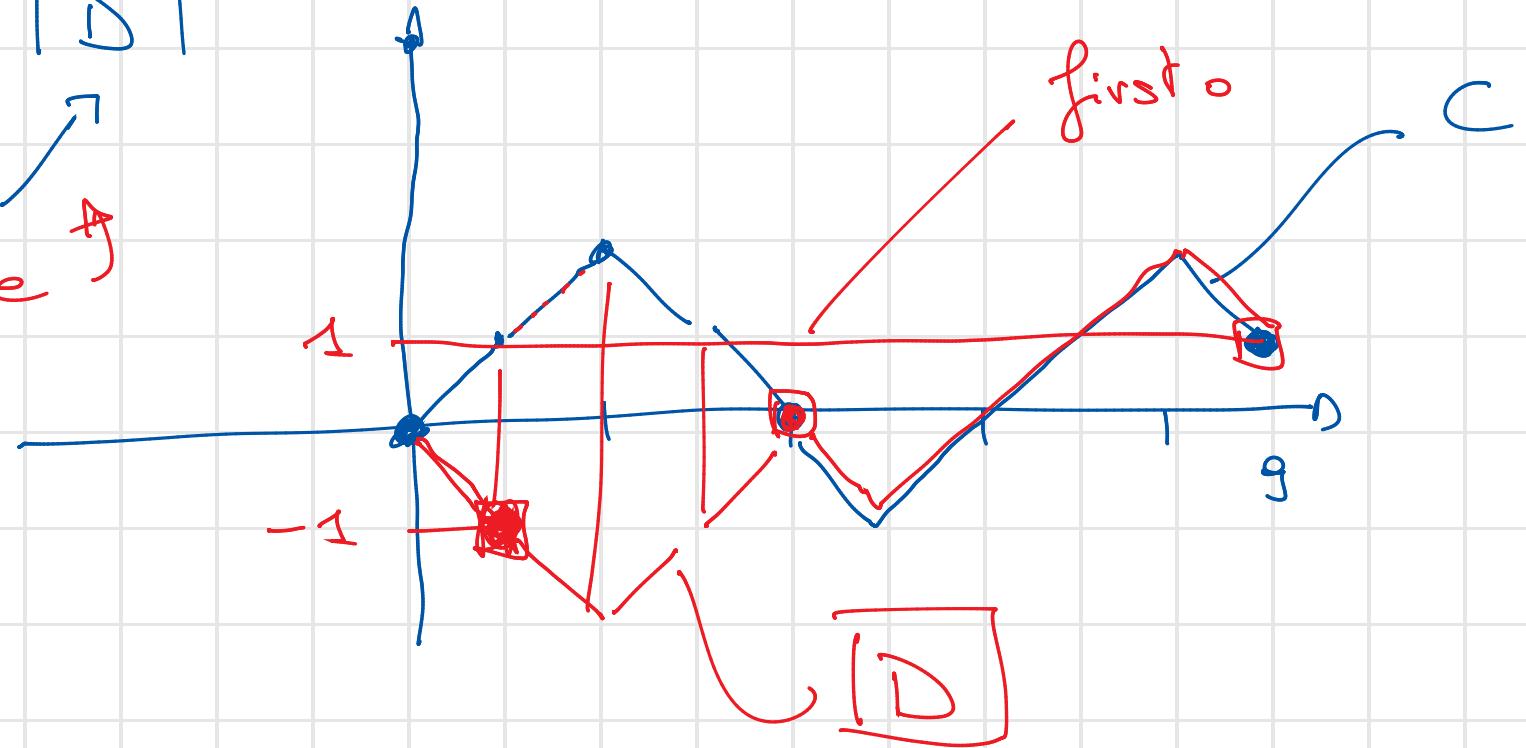
s.t.

$X_1 = 1$

$X_1 = -1$

$$|C| = |D|$$

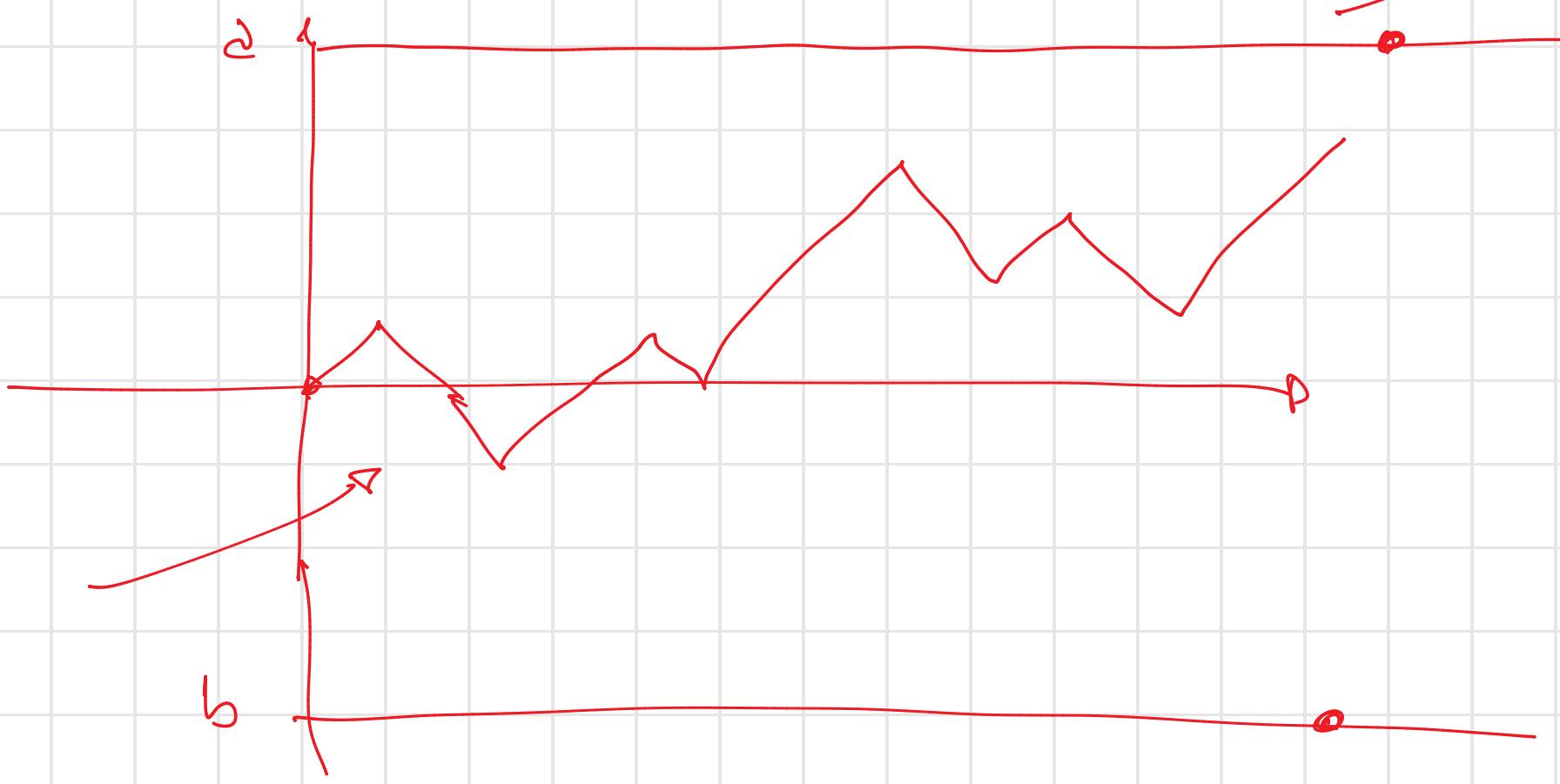
true



$D = \{ \text{paths from } (1, -1) \text{ to } (2n-1, 1) \}$

$|D|$ is easy . . .

$$P[T_0 = 2n] = \frac{P[S_{2n} = 0]}{2n-1}$$



Chpt. 4, Section 4

Theorem 4.2:

$\mu \geq 0$ positive recurrent, aperiodic

class with states $J = 0, 1, 2, \dots$

$$\xrightarrow{l} \lim_{n \rightarrow \infty} P_{JJ}^{(n)} = \pi_J = \sum_{i=0}^{\infty} \pi_i P_{ij},$$

$$\sum_{i=0}^{\infty} \pi_i = 1$$

and the π 's are uniquely determined by
the set of equations

$$\pi_i \geq 0, \quad \sum_{i=0}^{\infty} \pi_i = 1 \quad \text{and} \quad \pi_J = \sum_{i=0}^{\infty} \pi_i P_{ij},$$

$$J = 0, 1, 2, \dots$$

By Theorem 4.1

$$\pi_i = \frac{1}{\mu_i}$$

$$\mu_i = \mathbb{E}[R_i | X_0 = i]$$

$\mu_i < +\infty$ (positive rec.)

$$\pi_i > 0$$

(invariant)

$\pi = (\pi_0, \pi_1, \dots)$ this distribution is called

a stationary probability distribution for the MC

If $P[X_0 = i] = \pi_i$

$$P[X_1 = i] \stackrel{\text{by Markov property}}{=} \sum_{k=0}^{\infty} P[X_0 = k] \cdot P[X_1 = i | X_0 = k]$$

$$= \sum_{k=0}^{\infty} \pi_k P_{ki} = \pi_i$$

$\forall i = 0, 1, 2, \dots$

and in general

$$P[X_n = i] = \pi_i \quad i = 0, 1, 2, \dots \quad \forall n \geq 0$$

Remark : $Y_n = (X_n, X_{n+1})$ $(X_n, n \geq 0)$ is a MC

Prove that $Y_n, n \geq 0$ is a Markov Chain

if $P[X_0 = i] = \pi_i, i = 0, 1, 2, \dots$ initial dist.

$$P[Y_n = (i, j)] = P[X_n = i, X_{n+1} = j]$$



$$= \pi_i P_{ij}$$

$\boxed{\forall n \geq 0}$

Ex: $P = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix}$ irreducible, periodic (2)
MC

$$P^2 = Id = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}, \begin{cases} P^{2n} = Id & \text{for } n \\ P^{2n-1} = P & \text{for } n \end{cases}$$

$$\underset{n \rightarrow \infty}{\overset{\curvearrowleft}{\lim}} P_{ii}^n \neq$$

$$(\pi_0, \pi_1) \quad \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix}$$

$$\boxed{\pi = \pi P}$$

$$\begin{cases} \pi_0 = \pi_1 \\ \pi_1 = \pi_0 \\ \pi_0 + \pi_1 = 1 \end{cases}$$

$$\boxed{\pi_0 = \pi_1 = \frac{1}{2}}$$

$(\frac{1}{2}, \frac{1}{2})$ is the unique invariant distribution to the MC

invariant or stationary distribution

Invariant or stationary distribution

Definition: A distribution π and a transition matrix P are in detailed balance if

$$\boxed{\pi_i P_{ij} = \pi_j P_{ji} \quad \forall i, j \in S}$$

Remark: we have to check only the cases $i \neq j$

Proposition: If P and π are in detailed balance,

then π is invizet for P .

Proof:

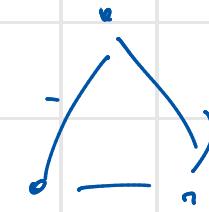
$$(\pi \cdot P)_i = \sum_{j \in S} \pi_j P_{ji} \stackrel{DB}{=} \sum_{j \in S} \pi_i P_{ij}$$

$$= \pi_i \underbrace{\sum_{j \in S} P_{ij}}_{\text{II}} = \pi_i \quad \text{I}$$

■

Ex 1 :

$$\begin{pmatrix} 0 & \frac{2}{3} & \frac{1}{3} \\ \frac{1}{3} & 0 & \frac{2}{3} \\ \frac{2}{3} & \frac{1}{3} & 0 \end{pmatrix}$$

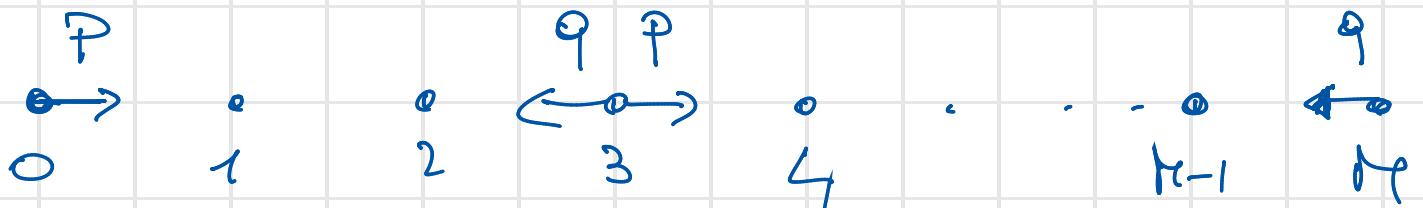


doubly stochastic

$\boxed{\pi_0 = \pi_1 = \pi_2 = \frac{1}{3}}$

$$\left(\frac{1}{3}, \frac{1}{3}, \frac{1}{3} \right)$$

Ex 2 :



$$S = \{0, 1, \dots, n\}$$

RW on S

$$0 < p = 1 - q < 1$$

$$P = \begin{bmatrix} 0 & q & p & 0 & \cdots & \cdots & 0 \\ 1 & q & 0 & p & \cdots & \cdots & 0 \\ 2 & 0 & q & 0 & p & \cdots & \cdots & \cdots \\ \vdots & & & & & & & \\ M-1 & \cdots \\ M & 0 & \cdots & \cdots & \cdots & \cdots & \cdots & 0 \\ & & & & & & & q & 0 & p \\ & & & & & & & 0 & q & p \end{bmatrix}$$

Find the invariant distribution!

Detailed balance equations

$$\pi_i \cdot p_{i,i+1} = \pi_{i+1} \cdot p_{i+1,i}$$

$\forall i=0,1,\dots,M-1$

$$p_{i,i+1} \equiv P, \quad p_{i+1,i} \equiv q$$

$$\pi_{i+1} = \frac{P}{q} \cdot \pi_i = \frac{P^2}{q^2} \pi_{i-1} = \dots$$

$$\pi_i = \left(\frac{P}{q}\right)^i \pi_0, \quad i=0, \dots, M$$

one solution
of $\pi = \pi P$

$$\sum_{i=0}^M \pi_i = 1$$

$$\pi_0 + \left(\frac{P}{q}\right) \pi_0 + \dots + \left(\frac{P}{q}\right)^M \pi_0 = 1$$

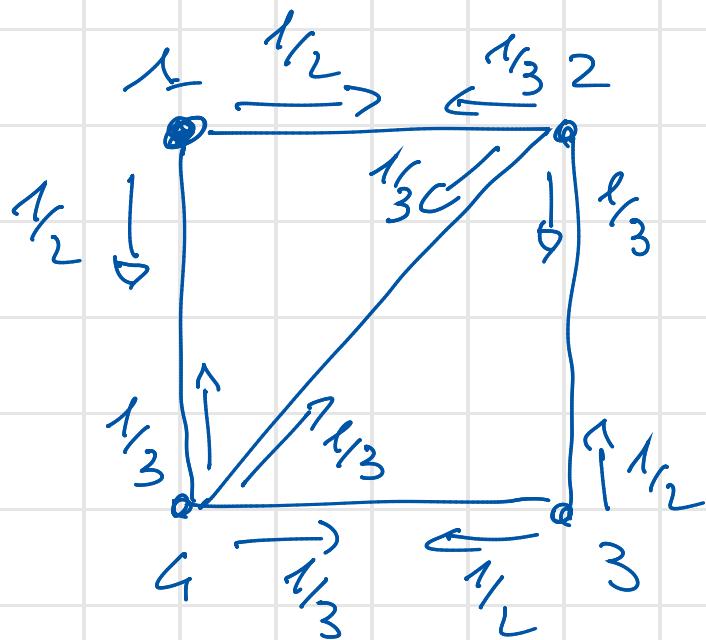
$$\pi_0 \left(1 + \frac{P}{q} + \dots + \left(\frac{P}{q}\right)^M\right) = 1$$

$$\pi_0 = \frac{1}{1 + \frac{P}{q} + \dots + \left(\frac{P}{q}\right)^M} = \frac{1 - \frac{P}{q}}{1 - \left(\frac{P}{q}\right)^{M+1}}$$

Example page 248 + Example page 249

Lemme 4.1 ↗ read!

Random Walk on a Graph



4 nodes
5 edges

$$P = \begin{bmatrix} 1 & 0 & 1/2 & 0 \\ 2 & 1/3 & 0 & 1/3 & 1/3 \\ 3 & 0 & 1/2 & 0 & 1/2 \\ 4 & 1/3 & 1/3 & 1/3 & 0 \end{bmatrix}$$

$\pi = \pi P$

invariant
distribution

$$\pi_i P_{ij} = \pi_j P_{ji}$$

$U_i = \# \text{ of neighbors nodes to } i$

$P_{ij} = \begin{cases} 0 & i, j \text{ are not neighbors} \\ 1/U_i & i, j \text{ are neighbors} \end{cases}$

$$\pi_i P_{ij} = \pi_j P_{ji} \quad \text{if and only if}$$

i and j are neighbors

$$\pi_i \cdot \frac{1}{\sigma_i} = \pi_j \cdot \frac{1}{\sigma_j}$$

$$\sigma = \sum_{i \in S} \sigma_i$$

$$\pi_i = \frac{\sigma_i}{\sigma}$$

$$\pi_i > 0, \quad \sum_i \pi_i = 1$$

$$\frac{\sigma_i}{\sigma} \cdot \frac{1}{\sigma_i} = \frac{1}{\sigma} = \frac{\sigma_j}{\sigma} \cdot \frac{1}{\sigma_j} = \frac{1}{\sigma}$$

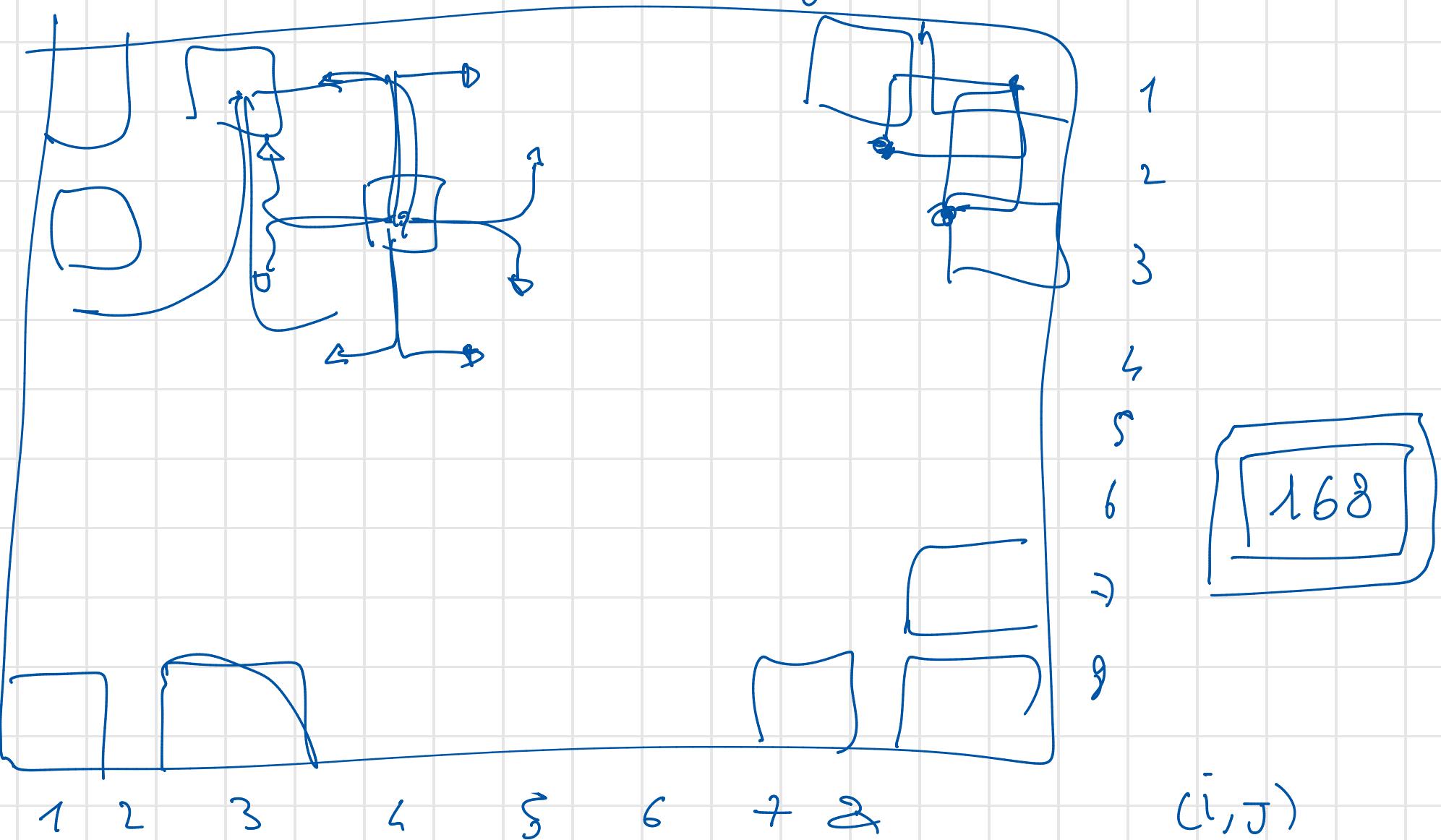
If the Graph is connected $\Rightarrow \pi_i = \frac{\sigma_i}{\sigma}, i \in S$

is the unique
unbiased
distribution

Exercise

Chess

Knight moves randomly



On average how many moves of this

random knight are needed starting

from (1,1) to go back to (1,1).

RW on a Graph

J_i

(1,1)

$\frac{J_i}{5}$