



COLLEGE OF ENGINEERING
SCHOOL OF AERONAUTICAL AND ASTRONAUTICAL ENGINEERING

AAE 567: INTRODUCTION TO APPLIED STOCHASTIC PROCESSES

HW4

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Problem 1

Let f be a random vector with values in \mathbb{C}^k . Let g be a random vector with values in \mathbb{C}^n and \mathcal{M} the subspace spanned by g . Assume that R_g is invertible. According to Theorem 2.2.1 in Chapter 2, we have

$$\begin{aligned} P_{\mathcal{M}}f &= R_{fg}R_g^{-1}g \\ E(f - P_{\mathcal{M}}f)(f - P_{\mathcal{M}}f)^* &= R_f - R_{fg}R_g^{-1}R_{gf}. \end{aligned}$$

Now let y be a random vector with values in \mathbb{C}^m , and set $\phi = y - P_{\mathcal{M}}y$. Theorem 2.2.1 in Chapter 2 shows that

$$\begin{aligned} P_{\mathcal{M}}y &= R_{yg}R_g^{-1}g \\ E(y - P_{\mathcal{M}}y)(y - P_{\mathcal{M}}y)^* &= R_{\phi} = R_y - R_{yg}R_g^{-1}R_{gy}. \end{aligned}$$

As in Lemma 3.3.1, consider the space \mathcal{H} spanned by g and y , that is, $\mathcal{H} = g \vee y$. Let h be the random vector defined by $h = [g \ y]^{tr}$ where tr denotes the transpose. Notice that \mathcal{H} equals the span of h . Furthermore, R_h and R_{fh} are given by

$$R_h = \begin{bmatrix} R_g & R_{gy} \\ R_{yg} & R_y \end{bmatrix} \quad \text{and} \quad R_{fh} = [R_{fg} \ R_{fy}].$$

Observe that the Schur complement for R_h is given by $\Delta = R_y - R_{yg}R_g^{-1}R_{gy} = R_{\phi}$. In particular, R_h is invertible if and only if R_{ϕ} is invertible; see Lemma 2.4.1 in Chapter 2.

Now assume that R_{ϕ} is invertible. Then R_h is invertible and Theorem 2.2.1 in Chapter 2 implies that

$$\begin{aligned} P_{\mathcal{H}}f &= R_{fh}R_h^{-1}h \\ E(f - P_{\mathcal{H}}f)(f - P_{\mathcal{H}}f)^* &= R_f - R_{fh}R_h^{-1}R_{hf}. \end{aligned}$$

Using the matrix inversion Lemma 2.4.1 in Chapter 2, give another proof of equations (3.1) and (3.2) in Lemma 3.3.1, that is, show that

$$\begin{aligned} P_{\mathcal{H}}f &= P_{\mathcal{M}}f + R_{f\phi}R_{\phi}^{-1}\phi \\ E(f - P_{\mathcal{H}}f)(f - P_{\mathcal{H}}f)^* &= E(f - P_{\mathcal{M}}f)(f - P_{\mathcal{M}}f)^* - R_{f\phi}R_{\phi}^{-1}R_{\phi f} \end{aligned}$$

Solution:

From what we are given we can and Lemma 2.4.1 we can compute

$$R_h^{-1} = \begin{bmatrix} R_g^{-1} + R_g^{-1}R_{gy}R_{\phi}^{-1}R_{yg}R_g^{-1} & -R_g^{-1}R_{gy}R_{\phi}^{-1} \\ -R_{\phi}^{-1}R_{yg}R_g^{-1} & R_{\phi}^{-1} \end{bmatrix}.$$

Then,

$$\begin{aligned}
R_{fh}R_h^{-1}h &= \begin{bmatrix} R_{fg} & R_{fy} \end{bmatrix} \begin{bmatrix} R_g^{-1} + R_g^{-1}R_{gy}R_\phi^{-1}R_{yg}R_g^{-1} & -R_g^{-1}R_{gy}R_\phi^{-1} \\ -R_\phi^{-1}R_{yg}R_g^{-1} & R_\phi^{-1} \end{bmatrix} \begin{bmatrix} g \\ y \end{bmatrix} \\
&= R_{fg}R_g^{-1}g + R_{fg}R_g^{-1}R_{gy}R_\phi^{-1}R_{yg}R_g^{-1}g - R_{fg}R_\phi^{-1}R_{yg}R_g^{-1}g \\
&\quad - R_{fg}R_g^{-1}R_{gy}R_\phi^{-1}y + R_{fy}R_\phi^{-1}y.
\end{aligned}$$

Since,

$$\begin{aligned}
P_{\mathcal{M}}f &= R_{fg}R_g^{-1}g \\
\phi &= y - P_{\mathcal{M}}y = y - R_{yg}R_g^{-1}g.
\end{aligned}$$

Then the equation above can be rewritten as

$$\begin{aligned}
R_{fh}R_h^{-1}h &= P_{\mathcal{M}}f + R_{fg}R_g^{-1}R_{gy}R_\phi^{-1}(y - R_{yg}R_g^{-1}g) + R_{fy}R_\phi^{-1}(y - R_{yg}R_g^{-1}g) \\
&= P_{\mathcal{M}}f + (R_{fy} - R_{fg}R_g^{-1}R_{gy})R_\phi^{-1}\phi
\end{aligned}$$

and because,

$$\begin{aligned}
\phi &= \text{span}y \\
R_{gy} &= 0 \quad \because g \perp y
\end{aligned}$$

we have

$$R_{fh}R_h^{-1}h = P_{\mathcal{M}}f + R_{f\phi}R_\phi^{-1}\phi$$

$$\therefore P_{\mathcal{H}} = P_{\mathcal{M}}f + R_{f\phi}R_\phi^{-1}\phi$$

Secondly,

$$\begin{aligned}
E(f - P_{\mathcal{H}}f)(f - P_{\mathcal{H}}f)^* &= E(f - P_{\mathcal{M}}f - R_{f\phi}R_\phi^{-1}\phi)(f - P_{\mathcal{M}}f - R_{f\phi}R_\phi^{-1}\phi)^* \\
&= E(f - P_{\mathcal{M}}f)(f - P_{\mathcal{M}}f)^* - R_{f\phi}R_\phi^{-1} \underbrace{E\phi\phi^*}_{R_\phi} R_\phi^{-1}R_{\phi f} \\
&= E(f - P_{\mathcal{M}}f)(f - P_{\mathcal{M}}f)^* - R_{f\phi}R_\phi^{-1}R_{\phi f}
\end{aligned}$$

Thus,

$$E(f - P_{\mathcal{H}}f)(f - P_{\mathcal{H}}f)^* = E(f - P_{\mathcal{M}}f)(f - P_{\mathcal{M}}f)^* - R_{f\phi}R_\phi^{-1}R_{\phi f}$$

q.e.d

Problem 2

As in the Kalman filtering Theorem 3.4.1, consider the state space system

$$\begin{aligned}x(n+1) &= Ax(n) + Bu(n) \\ y(n) &= Cx(n) + Dv(n)\end{aligned}$$

where $u(n)$ and $v(n)$ are independent white noise random processes, which are independent to the initial conditions $x(0)$. Recall that \mathcal{M}_n equals the linear span of $\{y(k(0))\}_0^n$ and the optimal state estimate in the Kalman filter is given by $\hat{x}(n) = P_{\mathcal{M}_{n-1}}x(n)$. Find the state estimate $P_{\mathcal{M}_n}x(n)$ for $x(n)$ in terms of $\hat{x}(n)$ and $y(n)$. Hint, according to Lemma 3.3.1, we have

$$P_{\mathcal{M}_n}f = P_{\mathcal{M}_{n-1}}f + R_{f\phi(n)}R_{\phi(n)}^{-1}\phi(n),$$

where f is any random vector, and $\phi(n) = y(n) - P_{\mathcal{M}_{n-1}}y(n)$.

Solution:

Lemma $\mathcal{H} = \mathcal{M} \vee \mathcal{Y}$, $\mathcal{Y} = \mathcal{Y} - P_{\mathcal{M}}\mathcal{Y}$, $\mathcal{E} = \text{span}\{\mathcal{Y}\}$

$$\mathcal{H} = \mathcal{M} \oplus \mathcal{E} \Rightarrow P_{\mathcal{H}} = P_{\mathcal{M}} + P_{\mathcal{E}}$$

$$P_{\mathcal{H}}f = P_{\mathcal{M}}f + P_{\mathcal{E}}f = P_{\mathcal{M}}f + P_{\mathcal{E}}R_{\mathcal{E}}^{-1}\mathcal{Y}$$

$$E(f - P_{\mathcal{H}}f)(f - P_{\mathcal{H}}f)^*$$

$$= E(f - P_{\mathcal{M}}f)(f - P_{\mathcal{M}}f)^* - P_{\mathcal{E}}R_{\mathcal{E}}^{-1}R_{\mathcal{E}}f$$

$$\mathcal{M}_n \subseteq \text{span}\{1, x(0), u(0), \dots, u(n-1), v(0), \dots, v(n)\}$$

$$\mathcal{M}_n = \mathcal{M}_{n-1} \vee \mathcal{Y}(n)$$

$$\hat{x}(n) = P_{\mathcal{M}_{n-1}}x(n)$$

$$\phi(n) = y(n) - P_{\mathcal{M}_{n-1}}y(n) = y(n) - P_{\mathcal{M}_{n-1}}(Cx(n) + Dv(n))$$

$$\downarrow = y(n) - CP_{\mathcal{M}_{n-1}}x(n) - \cancel{DP_{\mathcal{M}_{n-1}}v(n)}^0 \quad \because v(n) \perp \mathcal{M}_{n-1}$$

$$\phi(n) = y(n) - C\hat{x}(n)$$

$$\text{also } \varphi(n) = Cx(n) + DV(n) - C\hat{x}(n)$$

$$\varphi(n) = C\tilde{x}(n) + DV(n)$$

$$x(n+1) = Ax(n) + Bu(n)$$

$$\hat{x}(n+1) = P_{\mathcal{M}_n} x(n+1) = P_{\mathcal{M}_n} (Ax(n) + Bu(n))$$

$$= AP_{\mathcal{M}_n} x(n) + B \cancel{P_{\mathcal{M}_n}} u(n) \quad \because u(n) \perp \mathcal{M}_n$$

$$= AP_{\mathcal{M}_{n-1}} x(n) + AR_{x(n) \perp \mathcal{M}_n} B u(n) \quad \because \mathcal{M}_n = \mathcal{M}_{n-1} \oplus \text{span}$$

$$\hat{x}(n+1) = A\hat{x}(n) + AR_{x(n) \perp \mathcal{M}_n} R_{\varphi(n)}^{-1} (y(n) - C\hat{x}(n))$$

$$= A\hat{x}(n) + A_n (y(n) - C\hat{x}(n))$$

old
estimate

orthogonal
component

$$Q_n = E(\tilde{x}(n) \tilde{x}(n)^*)$$

$$R_{x(n) \perp \varphi(n)} = E x(n) \varphi(n)^* = E x(n) (C\tilde{x}(n) + DV(n))^*$$

$$= E x(n) \tilde{x}(n) C^* + E \cancel{x(n)} V(n)^* D^*$$

$$x(n) = \sum x(i) + \sum_j u(j)$$

$$x(i) \perp V(n), \quad u(j) \perp V(n)$$

$$x(j) \perp V(n)$$

$$\therefore = E(\tilde{x}(n) + \hat{x}(n)) \tilde{x}(n) C^*$$

$$= E(\tilde{x}(n) \tilde{x}(n)^*) C^* = Q_n C^*$$

$$R_{x(n)}\varphi(n) = Q_n C^*$$

$$R_{e(n)} = E \varphi(n) \varphi(n)^*$$

$$= E(C\tilde{x}(n) + DV(n))(C\tilde{x}(n) + DV(n))^*$$

$$= C E \tilde{x}(n) \tilde{x}(n)^* C^* + \underbrace{DV(n) V(n)^*}_{I} D^*$$

$$R_{e(n)} = C Q_n C^* + D D^*$$

$$\tilde{x} = x(n) - \hat{x}(n)$$

$$= x(n) - P_{m_{n-1}} x(n) \quad x(n) \perp V(n) \quad V(n) \perp m_{n-1}$$

$$\tilde{x}(n) \perp V(n)$$

cross terms go away

$$\hat{x}(n+1) = A \hat{x}(n) + A Q_n C^* (C Q_n C^* + D D^*)^{-1} (y(n) - C \hat{x}(n))$$

$$\Delta_n = A Q_n C^* (C Q_n C^* + D D^*)^{-1}$$

$$= A R_{x(n)} R_{e(n)}^{-1}$$

$$\hat{x}(n+1) = A \hat{x}(n) + \Delta_n (y(n) - C \hat{x}(n))$$

$$\Delta_n = A Q_n C^* (C Q_n C^* + D D^*)^{-1}$$

$$Q_n = E(\tilde{x}(n) \tilde{x}(n)^*)$$

$$Q_{n+1} = A Q_n A^* + B B^* - A Q_n C^* (C Q_n C^* + D D^*)^{-1} C Q_n A^*$$

$$\tilde{x}(n+1) = x(n+1) - \hat{x}(n+1)$$

$$= x(n+1) - P_{m_n} x(n+1)$$

$$= A x(n) + B u(n) - P_{m_n} (A x(n) + B u(n)) \quad \because m_n \perp u(n)$$

$$= A(x(n) - P_{m_n} x(n)) + B u(n)$$

$$Q_{n+1} = E(\tilde{x}(n+1) \tilde{x}(n+1)^*)$$

$$= E[A(x(n) - P_{m_n} x(n)) + B u(n)][A(x(n) - P_{m_n} x(n)) + B u(n)]^*$$

$$= A E \underbrace{(x(n) - P_{m,n} x(n)) (x(n) - P_{m,n} x(n))^*}_{\text{use Lemma}} A^* + B \underbrace{E u(n) u(n)^*}_I B^*$$

$$E (I - P_{m,n}) (I - P_{m,n})^* \\ = E (I - P_{m,n}) (I - P_{m,n})^* - P_{y,n} R_{y,n}^{-1} R_{y,n}$$

$$\Rightarrow A E \underbrace{(x(n) - P_{m,n-1} x(n))}_{\tilde{x}(n)} \underbrace{(x(n) - P_{m,n-1} x(n))^*}_{\tilde{x}(n)^*} A^* + B B^* \\ - A R_{x(n) \oplus (n)} R_{y(n)}^{-1} R_{x(n) \oplus (n)}^*$$

$E \tilde{x}(n) \tilde{x}(n)^* = Q_n$

$$Q_{n+1} = A Q_n A^* + B B^* - A Q_n C^* (C Q_n C^* + D D^*)^{-1} C Q_n A^*$$

$$Q_n = E (\tilde{x}(n) \tilde{x}(n)^*) \geq 0$$

Problem 3

Consider the system

$$\begin{aligned}x(n+1) &= ax(n) + u(n) \\ y(n) &= x(n) + v(n)\end{aligned}$$

where a is a scalar, $u(0), v(0), v(1), x(0)$ are all independent mean zero and variance one Gaussian random variables. Let $\mathcal{M}_0 = \text{span}\{y(0)\}$ and $\mathcal{M}_1 = \text{span}\{y(0), y(1)\}$. Find

- (i) $\hat{x}(0) = P_{\mathcal{M}_1}x(0)$.
- (ii) $E|x(0) - \hat{x}(0)|^2$.
- (iii) Find α and β such that $\phi(1) = y(1) - P_{\mathcal{M}_0}y(1) = \alpha y(1) + \beta y(0)$.

Kalman filtering is not needed to solve this problem.

Solution:

(i) From what we are given we know that

$$\begin{aligned}x(1) &= ax(0) + u(0) \\ y(0) &= x(0) + v(0) \\ y(1) &= x(1) + v(1) = ax(0) + u(0) + v(1)\end{aligned}$$

and if we define

$$g = \begin{bmatrix} y(0) \\ y(1) \end{bmatrix}$$

since

$$\mathcal{M}_1 = \text{span}\{y(0), y(1)\}.$$

Now,

$$\begin{aligned}\hat{x}(0) &= P_{\mathcal{M}_1}x(0) \\ &= R_{x(0)g}R_g^{-1}g.\end{aligned}$$

We can break down the calculations by components

$$\begin{aligned}R_{x(0)g} &= Ex(0)g^* = Ex(0) \begin{bmatrix} y(0) & y(1) \end{bmatrix} \\ &= E \begin{bmatrix} x(0)y(0) & x(0)y(1) \end{bmatrix} \\ &= E \begin{bmatrix} x(0)(x(0) + v(0)) & x(0)(ax(0) + u(0) + v(1)) \end{bmatrix} \\ &= \begin{bmatrix} 1 & a \end{bmatrix}\end{aligned}$$

and

$$\begin{aligned}
R_g &= E g g^* \\
&= \begin{bmatrix} E y(0)^2 & E y(0) y(1) \\ E y(1) y(0) & E y(1)^2 \end{bmatrix} \\
&= \begin{bmatrix} E(x(0) + v(0))^2 & E(x(0) + v(0))(ax(0) + u(0) + v(1)) \\ E(x(0) + v(0))(ax(0) + u(0) + v(1)) & E(ax(0) + u(0) + v(1))^2 \end{bmatrix} \\
&= \begin{bmatrix} 2 & a \\ a & a^2 + 2 \end{bmatrix}.
\end{aligned}$$

Hence,

$$\begin{aligned}
\hat{x}(0) &= \begin{bmatrix} 1 & a \end{bmatrix} \begin{bmatrix} 2 & a \\ a & a^2 + 2 \end{bmatrix}^{-1} \begin{bmatrix} y(0) \\ y(1) \end{bmatrix} \\
&= \frac{1}{a^2 + 4} \begin{bmatrix} 1 & a \end{bmatrix} \begin{bmatrix} a^2 + 2 & -a \\ -a & 2 \end{bmatrix}^{-1} \begin{bmatrix} y(0) \\ y(1) \end{bmatrix} \\
&= \frac{1}{a^2 + 4} \begin{bmatrix} 2 & a \end{bmatrix} \begin{bmatrix} y(0) \\ y(1) \end{bmatrix} \\
&= \frac{2y(0) + ay(1)}{a^2 + 4} \\
&= \frac{(a^2 + 2)x(0) + au(0) + 2v(0) + av(1)}{a^2 + 4}.
\end{aligned}$$

Thus,

$$\hat{x}(0) = \frac{(a^2 + 2)x(0) + au(0) + 2v(0) + av(1)}{a^2 + 4}.$$

(ii) The error is calculated by

$$\begin{aligned}
E|x(0) - \hat{x}(0)|^2 &= E x(0)^2 - R_{x(0)g} R_g^{-1} R_{x(0)g}^* \\
&= 1 - \begin{bmatrix} 1 & a \end{bmatrix} \begin{bmatrix} 2 & a \\ a & a^2 + 2 \end{bmatrix}^{-1} \begin{bmatrix} 1 \\ a \end{bmatrix} \\
&= 1 - \frac{1}{a^2 + 4} \begin{bmatrix} 1 & a \end{bmatrix} \begin{bmatrix} a^2 + 2 & -a \\ -a & 2 \end{bmatrix}^{-1} \begin{bmatrix} 1 \\ a \end{bmatrix} \\
&= 1 - \frac{1}{a^2 + 4} \begin{bmatrix} 2 & a \end{bmatrix} \begin{bmatrix} 1 \\ a \end{bmatrix} \\
&= 1 - \frac{a^2 + 2}{a^2 + 4} \\
&= \frac{2}{a^2 + 4}.
\end{aligned}$$

Thus,

$$E|x(0) - \hat{x}(0)|^2 = \frac{2}{a^2 + 4}$$

(iii) If

$$g = y(0)$$

we calculate

$$\begin{aligned}\phi(1) &= y(1) - P_{\mathcal{M}_0}y(1) \\ &= ax(0) + u(0) + v(1) - R_{y(1)g}R_g^{-1}g \\ &= ax(0) + u(0) + v(1) - \frac{Ey(1)y(0)}{Ey(0)^2}y(0) \\ &= ax(0) + u(0) + v(1) - \frac{a}{2}(x(0) + v(0)) \\ &= \frac{a}{2}x(0) + u(0) - \frac{a}{2}v(0) + v(1) \\ &= y(1) - \frac{a}{2}y(0).\end{aligned}$$

Thus,

$$\alpha = 1, \quad \beta = -\frac{a}{2}.$$