

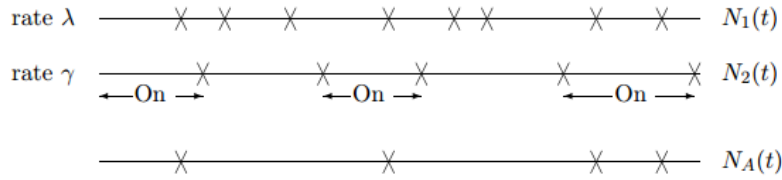
Homework 4 of Stochastic Processes

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1 Exercise 2.23

Exercise 2.23 Let $\{N_1(t); t > 0\}$ be a Poisson counting process of rate λ . Assume that the arrivals from this process are switched on and off by arrivals from a second independent Poisson process $\{N_2(t); t > 0\}$ of rate γ .



Let $\{N_A(t); t \geq 0\}$ be the switched process, i.e., $N_A(t)$ includes the arrivals from $\{N_1(t); t > 0\}$ during periods when $N_2(t)$ is even and excludes the arrivals from $\{N_1(t); t > 0\}$ while $N_2(t)$ is odd.

(a) Find the PMF for the number of arrivals of the first process, $\{N_1(t); t > 0\}$, during the n th period when the switch is on.

(b) Given that the first arrival for the second process occurs at epoch τ , find the conditional PMF for the number of arrivals of the first process up to τ .

(c) Given that the number of arrivals of the first process, up to the first arrival for the second process, is n , find the density for the epoch of the first arrival from the second process.

(d) Find the density of the interarrival time for $\{N_A(t); t \geq 0\}$. Note: This part is quite messy and is done most easily via Laplace transforms.

Solutions

- a) The combined process $\{N_1(t) + N_2(t)\}$ is also a Poisson process with rate $\lambda + \gamma$. During the n th period when the switch is on, once there is an arrival of process $\{N_2(t); t > 0\}$, the n th period will end. Therefore, the PMF for the number of arrivals of the first process $\{N_1(t); t > 0\}$ is

$$p_{N_1(t)}(k) = \left(\frac{\lambda}{\lambda + \gamma}\right)^k \frac{\gamma}{\lambda + \gamma}$$

where k denotes the number of arrivals of the first process $\{N_1(t); t > 0\}$ and n denotes the n th period when the switch is on.

- b) According to the Theorem 2.2.10, for a Poisson of rate λ , and for any $t > 0$, the PMF for $N(t)$, i.e., the number of arrivals in $(0, t]$, is given by the Poisson PMF,

$$p_{N(t)}(n) = \frac{(\lambda t)^n \exp(-\lambda t)}{n!}.$$

Since $\{N_1(t); t > 0\}$ and $\{N_2(t); t > 0\}$ are independent, given that the first arrival for process $\{N_2(t); t > 0\}$ at epoch τ , the conditional PMF for the number of arrivals of $\{N_1(t); t > 0\}$ up to τ is

$$p_{N_1(\tau)}(n) = \frac{(\lambda \tau)^n \exp(-\lambda \tau)}{n!}.$$

- c) Suppose that the first arrival of process $\{N_2(t); t > 0\}$ is at epoch τ , according to the Bayes' Law, we can obtain that

$$p_{N_1(\tau)}(n) f_{S_1^2|N_1(\tau)}(\tau|n) = f_{S_1^2}(\tau) p_{N_1(\tau)|S_1^2}(n|\tau)$$

where S_1^2 denotes the 1st arrival of process 2.

From a) we can know that

$$p_{N_1(\tau)}(n) = \left(\frac{\lambda}{\lambda + \gamma}\right)^n \frac{\gamma}{\lambda + \gamma}$$

From b) we can know that

$$p_{N_1(\tau)|S_1^2}(n|\tau) = \frac{(\lambda \tau)^n \exp(-\lambda \tau)}{n!}$$

And from the equation (2.13) $f_{S_n}(t) = \frac{\lambda^n t^{n-1} \exp(-\lambda t)}{(n-1)!}$ in the textbook we can know that

$$f_{S_1^2}(\tau) = \gamma \exp(-\gamma \tau)$$

Therefore,

$$\begin{aligned} f_{S_1^2|N_1(\tau)}(\tau|n) &= f_{S_1^2}(\tau) \frac{p_{N_1(\tau)|S_1^2}(n|\tau)}{p_{N_1(\tau)}(n)} \\ &= \gamma e^{-\gamma \tau} \cdot \frac{(\lambda \tau)^n e^{-\lambda \tau}}{n!} \cdot \frac{(\lambda + \gamma)^{n+1}}{\lambda^n \gamma} \\ &= \frac{(\lambda + \gamma)^{n+1} \tau^n e^{-(\lambda + \gamma)\tau}}{n!} \end{aligned}$$

- d) Since $N_1(t); t > 0$ and $N_2(t); t > 0$ are both independent and both renewal process, $N_A(t); t > 0$ should be also a renewal process. Therefore, each interval of $N_A(t); t > 0$ is independent and identically distributed.

Let X_A denote the interval of process $\{N_A(t); t \geq 0\}$, X denote the interval of the combined process $\{N_1(t) + N_2(t); t > 0\}$, and X_2 denote the interval of process $\{N_2(t); t > 0\}$.

There exists two cases:

1. If the next arrival belongs to process $\{N_1(t); t > 0\}$, $X_A = X$
2. If the next arrivals belongs to process $\{N_2(t); t > 0\}$, then X_A is the sum of three rv's: X , X_2 and X_A .

Therefore, the density of the interarrival time for $\{N_A(t); t \geq 0\}$ is

$$\begin{aligned}
f_{X_A}(x) &= \frac{\lambda}{\lambda + \gamma} f_X(x) + \left(\frac{\gamma}{\lambda + \gamma} f_X(x) \right) \otimes f_{X_2}(x) \otimes f_{X_A}(x) \\
&= \frac{\lambda}{\lambda + \gamma} \cdot (\lambda + \gamma) e^{-(\lambda + \gamma)x} + \left[\frac{\gamma}{\lambda + \gamma} \cdot (\lambda + \gamma) e^{-(\lambda + \gamma)x} \right] \otimes (\gamma e^{-\gamma x}) \otimes f_{X_A}(x) \\
&= \lambda e^{-(\lambda + \gamma)x} + (\gamma e^{-(\lambda + \gamma)x}) \otimes (\gamma e^{-\gamma x}) \otimes f_{X_A}(x)
\end{aligned}$$

where \otimes denotes the convolution operator and all functions satisfy $f(x) = 0$ for $x < 0$.

Let

$$\begin{aligned}
f_1(x) &= \lambda e^{-(\lambda + \gamma)x} \\
f_2(x) &= \gamma e^{-(\lambda + \gamma)x} \\
f_3(x) &= \gamma e^{-\gamma x}
\end{aligned}$$

Then we can obtain that:

$$f_{X_A}(x) = f_1(x) + f_2(x) \otimes f_3(x) \otimes f_{X_A}(x)$$

Apply the Laplace transforms

$$\begin{aligned}
\mathcal{L}[f_{X_A}] &= \mathcal{L}[f_1] + \mathcal{L}[f_2] \cdot \mathcal{L}[f_3] \cdot \mathcal{L}[f_{X_A}] \\
&\Downarrow \\
\mathcal{L}[f_{X_A}] &= \frac{\mathcal{L}[f_1]}{1 - \mathcal{L}[f_2]\mathcal{L}[f_3]}
\end{aligned}$$

Since

$$\begin{aligned}
\mathcal{L}[f_1] &= \int_0^\infty \lambda e^{-(\lambda + \gamma)t} \cdot e^{-st} dt = \frac{\lambda}{\lambda + \gamma + s} \\
\mathcal{L}[f_2] &= \int_0^\infty \gamma e^{-(\lambda + \gamma)t} \cdot e^{-st} dt = \frac{\gamma}{\lambda + \gamma + s} \\
\mathcal{L}[f_3] &= \int_0^\infty \gamma e^{-\gamma t} \cdot e^{-st} dt = \frac{\gamma}{\gamma + s}
\end{aligned}$$

$$\begin{aligned}
\mathcal{L}[f_{X_A}] &= \frac{\lambda}{\lambda + \gamma + s} \cdot \frac{1}{1 - \frac{\gamma}{\lambda + \gamma + s} \cdot \frac{\gamma}{\gamma + s}} \\
&= \frac{\lambda}{(\lambda + \gamma + s) - \frac{\gamma^2}{\gamma + s}} \\
&= \frac{\lambda(\gamma + s)}{(\lambda + \gamma + s)(\gamma + s) - \gamma^2} \\
&= \frac{\lambda(\gamma + s)}{\lambda\gamma + (\lambda + 2\gamma)s + s^2} \\
&= \frac{F_1(s)}{F_2(s)}
\end{aligned}$$

To obtain f_{X_A} , it needs to perform inverse Laplace transforms.

For $F_2(s)$, $(\lambda + 2\gamma)^2 - 4\lambda\gamma = \lambda^2 + 4\gamma^2 > 0$ and it has two roots with real value. Then we can rewrite $F_2(s)$ as

$$F_2(s) = \left(s - \frac{-(\lambda + 2\gamma) - \sqrt{\lambda^2 + 4\gamma^2}}{2} \right) \left(s - \frac{-(\lambda + 2\gamma) + \sqrt{\lambda^2 + 4\gamma^2}}{2} \right) = (s - s_1)(s - s_2)$$

And $\mathcal{L}[f_{X_A}]$ can be rewritten as

$$\mathcal{L}[f_{X_A}] = \frac{k_1}{s - s_1} + \frac{k_2}{s - s_2}$$

The value of k_1 and k_2 can be solved by:

$$\begin{aligned} k_1 &= (s - s_1) \mathcal{L}[f_{X_A}] \Big|_{s=s_1} \\ &= \frac{\lambda(\gamma + s)}{s - s_2} \Big|_{s=s_1} \\ &= \frac{\lambda(\gamma + \frac{-(\lambda+2\gamma)-\sqrt{\lambda^2+4\gamma^2}}{2})}{\frac{-(\lambda+2\gamma)-\sqrt{\lambda^2+4\gamma^2}}{2} - \frac{-(\lambda+2\gamma)+\sqrt{\lambda^2+4\gamma^2}}{2}} \\ &= \frac{\lambda \frac{-\lambda-\sqrt{\lambda^2+4\gamma^2}}{2}}{-\sqrt{\lambda^2+4\gamma^2}} \\ &= \frac{\lambda(\lambda + \sqrt{\lambda^2+4\gamma^2})}{2\sqrt{\lambda^2+4\gamma^2}} \\ &= \frac{\lambda}{2} \left(1 + \frac{\lambda}{\sqrt{\lambda^2+4\gamma^2}} \right) \\ k_2 &= (s - s_2) \mathcal{L}[f_{X_A}] \Big|_{s=s_2} \\ &= \frac{\lambda(\gamma + s)}{s - s_1} \Big|_{s=s_2} \\ &= \frac{\lambda(\gamma + \frac{-(\lambda+2\gamma)+\sqrt{\lambda^2+4\gamma^2}}{2})}{\frac{-(\lambda+2\gamma)+\sqrt{\lambda^2+4\gamma^2}}{2} - \frac{-(\lambda+2\gamma)-\sqrt{\lambda^2+4\gamma^2}}{2}} \\ &= \frac{\lambda \frac{-\lambda+\sqrt{\lambda^2+4\gamma^2}}{2}}{\sqrt{\lambda^2+4\gamma^2}} \\ &= \frac{\lambda(-\lambda + \sqrt{\lambda^2+4\gamma^2})}{2\sqrt{\lambda^2+4\gamma^2}} \\ &= \frac{\lambda}{2} \left(1 - \frac{\lambda}{\sqrt{\lambda^2+4\gamma^2}} \right) \end{aligned}$$

According to the formula of inverse Laplace transform:

$$f(t) = \mathcal{L}^{-1}[F(s)] = \mathcal{L}^{-1}\left[\frac{k_1}{s - s_1} + \frac{k_2}{s - s_2} + \cdots + \frac{k_n}{s - s_n}\right] = \sum_{i=1}^n k_i e^{s_i t}$$

we can obtain f_{X_A} as

$$\begin{aligned} f_{X_A}(x) &= k_1 \cdot e^{\frac{-(\lambda+2\gamma)-\sqrt{\lambda^2+4\gamma^2}}{2}x} + k_2 \cdot e^{\frac{-(\lambda+2\gamma)+\sqrt{\lambda^2+4\gamma^2}}{2}x} \\ &= k_1 \cdot \exp\left[\frac{-(\lambda+2\gamma)-\sqrt{\lambda^2+4\gamma^2}}{2}x\right] + k_2 \cdot \exp\left[\frac{-(\lambda+2\gamma)+\sqrt{\lambda^2+4\gamma^2}}{2}x\right] \end{aligned}$$

2 Exercise 2.25

Exercise 2.25 (a) Find the conditional density of S_{i+1} , conditional on $N(t) = n$ and $S_i = s_i$.

(b) Use (a) to find the joint density of S_1, \dots, S_n conditional on $N(t) = n$. Verify that your answer agrees with (2.38).

Solutions

a) From the equation (2.41) in the textbook,

$$\Pr\{S_1 > \tau | N(t) = n\} = \left[\frac{t-\tau}{t}\right]^n \quad \text{for } 0 < \tau \leq t$$

we can obtain that

$$\begin{aligned} f_{S_1|N(t)}(\tau|n) &= (1 - \Pr\{S_1 > \tau | N(t) = n\})' \\ &= \left(1 - \left[\frac{t-\tau}{t}\right]^n\right)' \\ &= -n \cdot \left[\frac{t-\tau}{t}\right]^{n-1} \cdot -\frac{1}{t} \\ &= \frac{n(t-\tau)^{n-1}}{t^n} \end{aligned}$$

Then, given that $N(t) = n$ and $S_i = s_i$, we can obtain that

$$\begin{aligned} f_{S_{i+1}|N(t), S_i}(s_{i+1}|n, s_i) &= f_{X_{i+1}|N(t), S_i}(s_{i+1} - s_i|n, s_i) \\ &= f_{X_{i+1}}|\tilde{N}(s_i, t)(s_{i+1} - s_i|n - i) \\ &= \frac{(n-i)[t - s_i - (s - i + 1 - s_i)]^{n-i-1}}{(t - s_i)^{n-i}} \\ &= \frac{(n-i)(t - s_{i+1})^{n-i-1}}{(t - s_i)^{n-i}} \end{aligned}$$

b) For S_i , the conditional probability is independent of S_1, S_2, \dots, S_{i-2} , and thus

$$\begin{aligned} f_{S^{(n)}|N(t)=n}(s^{(n)}|n) &= f_{S_1|N(t)} \cdot f_{S_2|N(t), S_1} \cdots f_{S_n|N(t), \dots, S_{n-1}} \\ &= f_{S_1|N(t)} \cdot f_{S_2|N(t), S_1} \cdots f_{S_n|N(t), S_{n-1}} \\ &= \frac{n(t-s_1)^{n-1}}{t^n} \cdot \frac{(n-1)(t-s_2)^{n-2}}{(t-s_1)^{n-1}} \cdots \frac{(t-s_n)^0}{t-s_{n-1}} \\ &= \frac{n!}{t^n} \end{aligned}$$