# HW6

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## Problem 1)

```
data <- read.table('pearls.txt', header=T, colClasses = c('numeric', 'factor', 'factor', 'factor'))</pre>
```

(a)

 $Y_{ijk} = \mu_{..} + \alpha_i + \beta_j + (\alpha\beta)_{ij} + \epsilon_{ijk}$ 

the terms  $\alpha_i, (\alpha \beta)_{ij}, \epsilon_{ijk}$  are random and  $\mu_{..}, \beta_j$  are fixed.

(b)

Mean sq	df	Expected MS
MSA	a - 1 = 2	$\sigma^2 + n\sigma_{\alpha\beta}^2 + \frac{bn}{a-1}\sum_i \alpha_i^2$
MSB	b - 1 = 3	$\sigma^2 + n\sigma_{\alpha\beta}^2 + \frac{bn}{a-1} \sum_j \alpha_i^2$ $\sigma^2 + an\sigma_\beta^2$
MSAB	(a-1)(b-1) = 6	$\sigma^2 + n\sigma_{\alpha\beta}^2$
MSE	ab(n-1) = 36	$\sigma^2$

(c)

```
library(lme4)
```

```
fit <- lmer(value ~ no_coats + (1|batch), data=data)</pre>
summary(fit)
## Linear mixed model fit by REML ['lmerMod']
## Formula: value ~ no_coats + (1 | batch)
##
      Data: data
##
## REML criterion at convergence: 207.9
##
## Scaled residuals:
                1Q Median
                                ЗQ
                                        Max
## -2.1991 -0.6281 0.1033 0.6555 1.3638
##
## Random effects:
## Groups
                         Variance Std.Dev.
             Name
## batch
             (Intercept) 3.898
                                  1.974
## Residual
                         4.178
                                   2.044
```

## Number of obs: 48, groups: batch, 4

## Loading required package: Matrix

```
##
## Fixed effects:
                Estimate Std. Error t value
## (Intercept) 73.1062
                                        65.77
                              1.1116
## no_coats2
                  3.6875
                              0.7227
                                         5.10
## no_coats3
                  3.8188
                              0.7227
                                         5.28
## Correlation of Fixed Effects:
##
              (Intr) n_cts2
## no_coats2 -0.325
## no_coats3 -0.325 0.500
fit2 <- lm(value ~ no_coats + batch + no_coats:batch, data)</pre>
anova(fit2)
## Analysis of Variance Table
##
## Response: value
##
                   Df Sum Sq Mean Sq F value
                                                    Pr(>F)
## no_coats
                    2 150.388 75.194 15.591 1.327e-05 ***
                    3 152.852
                                50.951 10.564 3.984e-05 ***
## batch
## no_coats:batch 6 1.852
                                  0.309
                                          0.064
                                                    0.9988
## Residuals
                   36 173.625
                                  4.823
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
H_0: \sigma^2_{\alpha\beta} = 0
H_a: \sigma_{\alpha\beta}^2 > 0
F* = \frac{MSAB}{MSE} = 0.0627666 < F = 2.363751
concluding null, interaction term is not significant.
```

d)

$$\begin{split} H_0: \sigma_\alpha^2 &= 0 \\ H_a: \sigma_\alpha^2 &> 0 \\ F^* &= \frac{75.194}{0.309} = 243.3462783 \\ F &= 5.1432528 \end{split}$$

 $F^* > F$  concluding  $H_a$  the effect of coat factor is significant.

**e**)

${\rm Mean\ sq}$	df	Expected MS
MSA	a - 1 = 2	$\sigma^2 + bn\sigma_{\alpha}^2$
MSB	b - 1 = 3	$\sigma^2 + n\sigma_{\alpha\beta}^2 + \frac{an}{b-1}\sum_j \beta_j^2$
MSAB	(a-1)(b-1) = 6	$\sigma^2 + n\sigma_{\alpha\beta}^2$
MSE	ab(n-1) = 36	$\sigma^2$

f)

```
mu = tapply(data$value, data$no_coat, mean)
11 = mu[2] - mu[1]
12 = mu[3] - mu[1]
s = sqrt((0.309/(4*4))*(2))
t = qt(1- 0.1/(2*2), 3*4*3)
```

B = 2.028094

$$s\{\hat{D}\} = 0.1965324$$

$$\mu_6 - \mu_8 = 3.6875$$

$$\mu_6 - \mu_{10} = 3.81875$$

$$3.2889137 \le \mu_8 - \mu_6 \le 4.0860863$$

$$3.4201637 \le \mu_8 - \mu_{10} \le 4.2173363$$

## Problem 2)

**a**)

$$Y_{ijk} = \mu_{..} + \alpha_i + \beta_{j(i)} + \epsilon_{ijk}$$

 $\mu$ ...constantvariable

 $\alpha_i$  : constant related to factor A and subject to  $\sum \alpha_i = 0$ 

 $\beta_{j(i)}$ : randomnormalvariable with mean 0

 $epsilon_{ijk}independetN(0, \sigma^2)$ 

$$i = 1, \dots, 2; j = 1, \dots, 3; k = 1, \dots, 5$$

b)

SSA = 0.01825

$$SSB(A) = 0.01153 + 0.44249 = 0.45402$$

SSE=0.29020

Source	SS	df	MS	$E\{MS\}$
Factor A	0.01825	a-1=1	0.01825	$\sigma^2 + bn \frac{\sum \alpha_i^2}{n!} + n\sigma_\beta^2$
Factor B(in	0.45402	a(b-	0.113505	$\sigma^2 + bn \frac{\sum_{\alpha=1}^{\alpha_i^2}}{a-1} + n\sigma_{\beta}^2$ $\sigma^2 + n\sigma_{\beta}^2$
A)		1)=4		,
Error	0.29020	,	0.0120917	$\sigma^2$
		1)=24		
Total	0.47227	abn-	0.0162852	
		1 = 29		

assuming  $\alpha = 0.05$ 

for factor A:

$$F^* = MSA/MSB(A) = 0.1607859$$

F = 7.7086474

$$F^* < F$$

concluding  $H_0$ , there's no considerable variability in different sites.

for factor B:

$$F^* = MSB(A)/MSE = 9.3870434$$

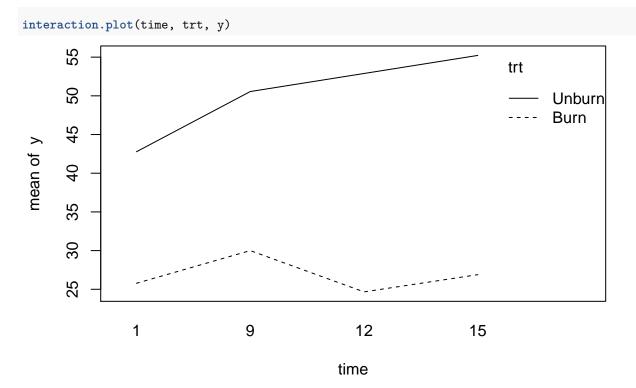
F = 2.7762893

 $F^* > F$  concluding  $H_a$ , there is considerable batch to batch difference.

## Problem 3)

```
data = read.table('floral.dat',header=T,sep='\t')
trt = as.factor(data$trt)
plot = as.factor(data$plot)
time = as.factor(data$time)
y = data$resp
```

a)



The mean of y for unburned region is higher compared to the burned region. The mean of Y increases with time for unburned region. For the burned region the mean of y increases with time between 1-9 and deacreases between 9-12 and again increases between 12-15.

## b)

 $Y_{ijk} = \mu_{...} + \rho_{i(j)} + \alpha_j = \beta_k + (\alpha\beta)_{jk} + \epsilon_{ijk}$ 

 $\mu_{\cdots}$  is constant

 $\rho_{i(j)}$  random ~  $N(0,\sigma_p^2),$  nested with in factor A.

 $\alpha_j$ burn/no burn, fixed

 $\beta_k$  : Time factor, fixed

### **c**)

That they have constant covariance. ## d)

 $\mathbf{e})$