

ISTA 421 + INFO 521 Introduction to Machine Learning

Lecture 9:

Maximum Likelihood –

Uncertainty in Parameters

and Predictions

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Topics

- Maximum Likelihood
- (What to expect from) Expectation
- The generative picture
- Maximum Likelihood Estimation
 - Uncertainty in parameters
 - Uncertainty in predictions



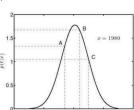
Explicitly Modeling Uncertainty

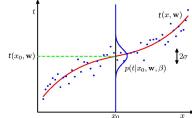
$$t_n = f(\mathbf{x}_n; \mathbf{w}) + \epsilon_n, \ \ \epsilon_n \sim \mathcal{N}(0, \sigma^2)$$
 Adding a R.V to our linear model: ... to model noise/variance/uncertainty in the response (t). $p(t_n | \mathbf{x}_n, \mathbf{w}, \sigma^2) = \mathcal{N}(\mathbf{w}^\mathsf{T} \mathbf{x}_n, \sigma^2)$ Now our "response" is itself a random variable. (Here we assume Gaussian/Normal noise.)

We assume (conditional) **independence**: given a particular input **x** and parameters **w**, the predicted variance at that point is indep. of any other point. However, we also assume the variance form (i.e., what type of distribution) and its parameters will be **identical** at any point (so are constant, not a fn of **x** or **w**). —Hence, i.i.d.

$$L = p(\mathbf{t}|\mathbf{X}, \mathbf{w}, \sigma^2) = \prod_{n=1}^{N} p(t_n|\mathbf{x}_n, \mathbf{w}, \sigma^2) = \prod_{n=1}^{N} \mathcal{N}(\mathbf{w}^{\mathsf{T}}\mathbf{x}_n, \sigma^2)$$

We can now use the data (input and target pairs) to give us a "likelihood" score of how well the model "fits" the mass/density of our model's uncertainty over the data. The more the data all fits under the bulk of the probability mass/density, the more it looks like our model could have generated the data (the more "likely" the data looks given our model).







Maximize the Likelihood

$$L = p(\mathbf{t}|\mathbf{X}, \mathbf{w}, \sigma^2) = \prod_{n=1}^{N} p(t_n|\mathbf{x}_n, \mathbf{w}, \sigma^2) = \prod_{n=1}^{N} \mathcal{N}(\mathbf{w}^{\mathsf{T}}\mathbf{x}_n, \sigma^2)$$

Since we are working with a product of Gaussians, which in turn include The exponential function (e), take the natural log (often just represented generically as $\log(L)$)

$$\log L = \sum_{n=1}^{N} \log \left(\frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{ -\frac{1}{2\sigma^2} (t_n - f(\mathbf{x}_n; \mathbf{w}))^2 \right\} \right)$$

$$= \sum_{n=1}^{N} \left(-\frac{1}{2} \log(2\pi) - \log \sigma - \frac{1}{2\sigma^2} (t_n - f(\mathbf{x}_n; \mathbf{w}))^2 \right)$$

$$= -\frac{N}{2} \log 2\pi - N \log \sigma - \frac{1}{2\sigma^2} \sum_{n=1}^{N} (t_n - f(\mathbf{x}_n; \mathbf{w}))^2.$$



$$\frac{df}{dx} = \frac{df}{da} \cdot \frac{dg}{dx}$$

Maximize the Likelihood: w

$$\log L = -\frac{N}{2} \log 2\pi - N \log \sigma - \frac{1}{2\sigma^2} \sum_{n=1}^{N} (t_n - f(\mathbf{x}_n; \mathbf{w}))^2$$
$$= -\frac{N}{2} \log 2\pi - N \log \sigma - \frac{1}{2\sigma^2} \sum_{n=1}^{N} (t_n - \mathbf{w}^\mathsf{T} \mathbf{x}_n)^2$$

$$\frac{\partial \log L}{\partial \mathbf{w}} = \frac{1}{\sigma^2} \sum_{n=1}^{N} \mathbf{x}_n (t_n - \mathbf{x}_n^\mathsf{T} \mathbf{w}) \\ = \frac{1}{\sigma^2} \sum_{n=1}^{N} \mathbf{x}_n t_n - \mathbf{x}_n \mathbf{x}_n^\mathsf{T} \mathbf{w} = \mathbf{0} \qquad \mathbf{X} = \begin{bmatrix} \mathbf{x}_1^\mathsf{T} \\ \mathbf{x}_2^\mathsf{T} \\ \vdots \\ \mathbf{x}_N^\mathsf{T} \end{bmatrix} = \begin{bmatrix} 1 & x_1 \\ 1 & x_2 \\ \vdots & \vdots \\ 1 & x_N \end{bmatrix}, \quad \mathbf{t} = \begin{bmatrix} t_1 \\ t_2 \\ \vdots \\ t_N \end{bmatrix}$$

$$\frac{\partial \log L}{\partial \mathbf{w}} = \frac{1}{\sigma^2} (\mathbf{X}^\mathsf{T} \mathbf{t} - \mathbf{X}^\mathsf{T} \mathbf{X} \mathbf{w}) = \mathbf{0}$$

$$\frac{1}{\sigma^2} (\mathbf{X}^\mathsf{T} \mathbf{t} - \mathbf{X}^\mathsf{T} \mathbf{X} \mathbf{w}) = 0$$

$$\mathbf{X}^\mathsf{T} \mathbf{t} - \mathbf{X}^\mathsf{T} \mathbf{X} \mathbf{w} = 0$$

$$\mathbf{X}^\mathsf{T} \mathbf{X} \mathbf{w} = \mathbf{X}^\mathsf{T} \mathbf{t}$$

$$\mathbf{w} = (\mathbf{X}^\mathsf{T} \mathbf{X})^{-1} \mathbf{X}^\mathsf{T} \mathbf{t}$$

Maximize the Likelihood: σ

$$\log L = -\frac{N}{2} \log 2\pi - N \log \sigma - \frac{1}{2\sigma^2} \sum_{n=1}^{N} (t_n - \mathbf{w}^\mathsf{T} \mathbf{x}_n)^2$$

$$\frac{\partial L}{\partial \sigma} = -\frac{N}{\sigma} + \frac{1}{\sigma^3} \sum_{n=1}^{N} (t_n - \mathbf{x}^\mathsf{T} \widehat{\mathbf{w}})^2 = 0$$

$$\widehat{\sigma^{2}} = \frac{1}{N} \sum_{n=1}^{N} (t_{n} - \mathbf{x}^{\mathsf{T}} \widehat{\mathbf{w}})^{2}$$

$$\widehat{\sigma^{2}} = \frac{1}{N} (\mathbf{t}^{\mathsf{T}} \mathbf{t} - 2\mathbf{t}^{\mathsf{T}} \mathbf{X} (\mathbf{x}^{\mathsf{T}} \mathbf{x})^{-1} \mathbf{X}^{\mathsf{T}} \mathbf{t} + \mathbf{t}^{\mathsf{T}} \mathbf{X} (\mathbf{x}^{\mathsf{T}} \mathbf{x})^{-1} \mathbf{X}^{\mathsf{T}} \mathbf{t})$$

$$= \frac{1}{N} (\mathbf{t}^{\mathsf{T}} \mathbf{t} - 2\mathbf{t}^{\mathsf{T}} \mathbf{X} (\mathbf{x}^{\mathsf{T}} \mathbf{x})^{-1} \mathbf{X}^{\mathsf{T}} \mathbf{t} + \mathbf{t}^{\mathsf{T}} \mathbf{X} (\mathbf{x}^{\mathsf{T}} \mathbf{x})^{-1} \mathbf{X}^{\mathsf{T}} \mathbf{t})$$

$$= \frac{1}{N} (\mathbf{t}^{\mathsf{T}} \mathbf{t} - \mathbf{X} \widehat{\mathbf{w}})^{\mathsf{T}} (\mathbf{t} - \mathbf{X} \widehat{\mathbf{w}})$$

$$= \frac{1}{N} (\mathbf{t}^{\mathsf{T}} \mathbf{t} - 2\mathbf{t}^{\mathsf{T}} \mathbf{X} \widehat{\mathbf{w}} + \widehat{\mathbf{w}}^{\mathsf{T}} \mathbf{X}^{\mathsf{T}} \mathbf{X} \widehat{\mathbf{w}})$$

$$\widehat{\sigma^{2}} = \frac{1}{N} (\mathbf{t}^{\mathsf{T}} \mathbf{t} - \mathbf{t}^{\mathsf{T}} \mathbf{X} \widehat{\mathbf{w}})$$

$$\sigma^{\mathsf{T}} = \frac{1}{N} (\mathbf{t} - \mathbf{X} \mathbf{w})^{\mathsf{T}} (\mathbf{t} - \mathbf{X} \mathbf{w})$$

$$= \frac{1}{N} (\mathbf{t}^{\mathsf{T}} \mathbf{t} - 2\mathbf{t}^{\mathsf{T}} \mathbf{X} \widehat{\mathbf{w}} + \widehat{\mathbf{w}}^{\mathsf{T}} \mathbf{X}^{\mathsf{T}} \mathbf{X} \widehat{\mathbf{w}})$$

$$\widehat{\sigma^{2}} = \frac{1}{N} (\mathbf{t}^{\mathsf{T}} \mathbf{t} - \mathbf{t}^{\mathsf{T}} \mathbf{X} \widehat{\mathbf{w}})$$

Simplify further by plugging in

$$\widehat{\mathbf{w}} = (\mathbf{X}^\mathsf{T} \mathbf{X})^{-1} \ \mathbf{X}^\mathsf{T} \mathbf{t}$$



Maximum Likelihood

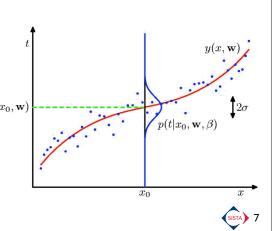
$$L = p(\mathbf{t}|\mathbf{X}, \mathbf{w}, \sigma^2) = \prod_{n=1}^{N} p(t_n|\mathbf{x}_n, \mathbf{w}, \sigma^2) = \prod_{n=1}^{N} \mathcal{N}(\mathbf{w}^{\mathsf{T}}\mathbf{x}_n, \sigma^2)$$

$$\widehat{\mathbf{w}} = \left(\mathbf{X}^\mathsf{T}\mathbf{X}\right)^{-1}\mathbf{X}^\mathsf{T}\mathbf{t}$$

$$\widehat{\mathbf{w}} = \left(\mathbf{X}^{\mathsf{T}} \mathbf{X}\right)^{-1} \mathbf{X}^{\mathsf{T}} \mathbf{t}$$

$$\widehat{\sigma^2} = \frac{1}{N} (\mathbf{t}^{\mathsf{T}} \mathbf{t} - \mathbf{t}^{\mathsf{T}} \mathbf{X} \widehat{\mathbf{w}})$$

Predictive distribution



The equations are unique

$$\widehat{\mathbf{w}} = \left(\mathbf{X}^\mathsf{T}\mathbf{X}\right)^{-1}\mathbf{X}^\mathsf{T}\mathbf{t}$$

$$\widehat{\sigma^2} = \frac{1}{N} (\mathbf{t}^\mathsf{T} \mathbf{t} - \mathbf{t}^\mathsf{T} \mathbf{X} \widehat{\mathbf{w}})$$

Hessian Matrix
$$\widehat{\boldsymbol{w}} = \left(\mathbf{X}^{\mathsf{T}}\mathbf{X}\right)^{-1}\mathbf{X}^{\mathsf{T}}\mathbf{t}$$

$$\widehat{\boldsymbol{\sigma}^2} = \frac{1}{N}(\mathbf{t}^{\mathsf{T}}\mathbf{t} - \mathbf{t}^{\mathsf{T}}\mathbf{X}\widehat{\mathbf{w}})$$

$$\mathbf{H} = \begin{bmatrix} \frac{\partial^2 f}{\partial w_1^2} & \frac{\partial^2 f}{\partial w_1 \partial w_2} & \cdots & \frac{\partial^2 f}{\partial w_1 \partial w_K} \\ \frac{\partial^2 f}{\partial w_2 \partial w_1} & \frac{\partial^2 f}{\partial w_2^2} & \cdots & \frac{\partial^2 f}{\partial w_2 \partial w_K} \\ \vdots & \vdots & \ddots & \vdots \\ \frac{\partial^2 f}{\partial w_K \partial w_1} & \frac{\partial^2 f}{\partial w_K \partial w_2} & \cdots & \frac{\partial^2 f}{\partial w_K^2} \end{bmatrix}$$

$$\log L = -\frac{N}{2} \log 2\pi - N \log \sigma - \frac{1}{2\sigma^2} \sum_{n=1}^{N} (t_n - \mathbf{w}^\mathsf{T} \mathbf{x}_n)^2$$

$$\frac{\partial {\log L}}{\partial \mathbf{w}} = \frac{1}{\sigma^2} (\mathbf{X}^\mathsf{T} \mathbf{t} - \mathbf{X}^\mathsf{T} \mathbf{X} \mathbf{w})$$

$$\frac{\partial^2 \log L}{\partial \mathbf{w} \partial \mathbf{w}^\mathsf{T}} = -\frac{1}{\sigma^2} \mathbf{X}^\mathsf{T} \mathbf{X}$$

 $\frac{\partial^2 \log L}{\partial \mathbf{w} \partial \mathbf{w}^\mathsf{T}} = -\frac{1}{\sigma^2} \mathbf{X}^\mathsf{T} \mathbf{X}$ This matrix is definite negative, so maximum likelihood solution (with linear model with additive Gaussian noise) is unique. $\mathbf{x}^\mathsf{T} \mathbf{H} \mathbf{x} < 0$

Expectation (a reminder)

• The **expected value** of a random variable is the weighted (by probability) average of all possible values $\mathbf{E}_{P(x)} \{f(x)\} = \sum f(x)P(x)$

Possible values
$$\mathbf{E}_{P(\mathbf{x})}\{f(\mathbf{x})\} = \sum_{\mathbf{x}} f(\mathbf{x})P(\mathbf{x})$$

$$\mathbf{E}_{p(\mathbf{x})}\{f(\mathbf{x})\} = \int_{\mathbf{x}} f(\mathbf{x})P(\mathbf{x})d\mathbf{x}$$

$$\mathbf{E}_{p(\mathbf{x})}\{f(\mathbf{x})\} = \int_{\mathbf{x}} f(\mathbf{x})P(\mathbf{x})d\mathbf{x}$$

$$\mathbf{E}_{p(\mathbf{x})}\{f(\mathbf{x})\} = \int_{\mathbf{x}} f(\mathbf{x})P(\mathbf{x})d\mathbf{x}$$

$$\mathbf{E}_{p(\mathbf{x})}\{\mathbf{x}\} = \sum_{\mathbf{x}} \mathbf{x}P(\mathbf{x})$$

$$\text{"(Co)Variance": } \mathbf{x}$$

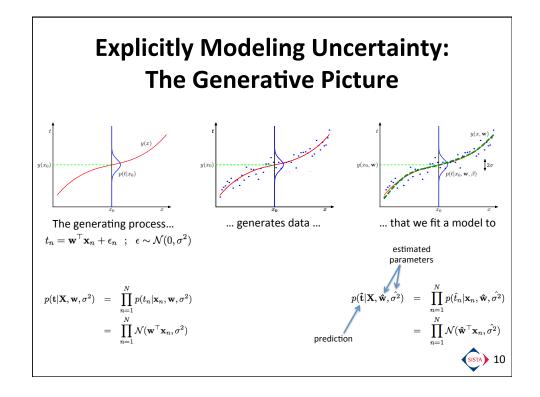
$$\cot(\mathbf{x}) = \mathbf{E}_{P(\mathbf{x})}\{(\mathbf{x} - \mathbf{E}_{P(\mathbf{x})}\{\mathbf{x}\})(\mathbf{x} - \mathbf{E}_{P(\mathbf{x})}\{\mathbf{x}\})^{\top}\}$$

$$\cot(\mathbf{x}) = \mathbf{E}_{P(\mathbf{x})}\{\mathbf{x}\} = \mathbf{E}_{P(\mathbf{x})}\{(\mathbf{x} - \mathbf{E}_{P(\mathbf{x})}\{\mathbf{x}\})(\mathbf{x} - \mathbf{E}_{P(\mathbf{x})}\{\mathbf{x}\})^{\top}\}$$

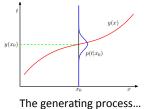
$$\mathbf{E}_{p(\mathbf{x})}\{f(\mathbf{x})\} = \int_{\mathbf{x}} f(\mathbf{x})P(\mathbf{x})d\mathbf{x}$$

$$\cot(\mathbf{x}) = \mathbf{E}_{P(\mathbf{x})}\{\mathbf{x}\} = \mathbf{E}_{P(\mathbf{x})}\{\mathbf{x}\} = \mathbf{E}_{P(\mathbf{x})}\{\mathbf{x}\} = \mathbf{E}_{P(\mathbf{x})}\{\mathbf{x}\}$$

$$\mathbf{E}_{p(\mathbf{x})}\{f(\mathbf{x})\} = \int_{\mathbf{x}} f(\mathbf{x})P(\mathbf{x})d\mathbf{x}$$



Explicitly Modeling Uncertainty: The Generative Picture

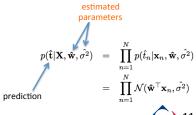


 $y(x_0)$... generates data ...

If we generate new data from the generating process, we get different parameters $\hat{\boldsymbol{w}}$ and $\hat{\boldsymbol{\sigma}}^2$, and a different predictive distribution $p(\hat{\boldsymbol{t}}|...)$

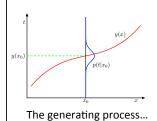
the generating process...
$$t_n = \mathbf{w}^{\top} \mathbf{x}_n + \epsilon_n \; \; ; \; \; \epsilon \sim \mathcal{N}(0, \sigma^2)$$

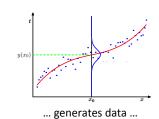
$$\begin{split} p(\mathbf{t}|\mathbf{X}, \mathbf{w}, \sigma^2) &= & \prod_{n=1}^N p(t_n|\mathbf{x}_n, \mathbf{w}, \sigma^2) \\ &= & \prod_{n=1}^N \mathcal{N}(\mathbf{w}^\top \mathbf{x}_n, \sigma^2) \end{split}$$



SISTA 11

Explicitly Modeling Uncertainty: The Generative Picture

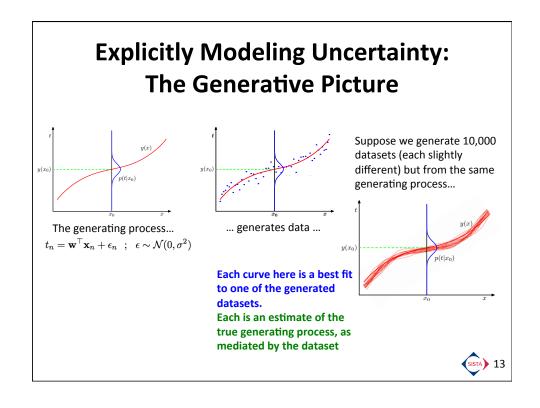


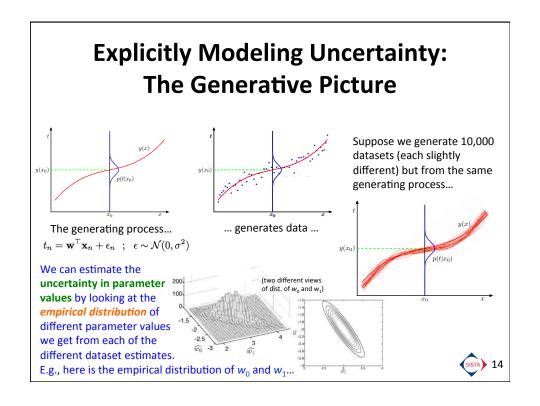


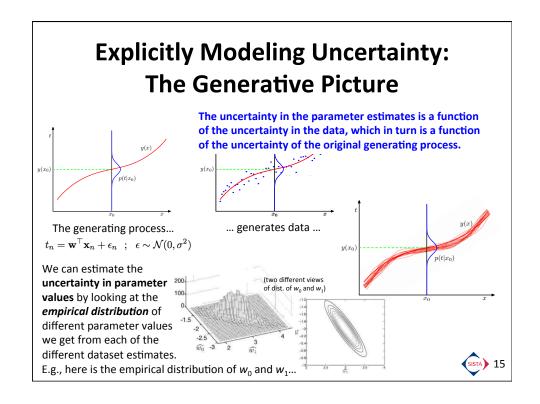
Suppose we generate 10,000 datasets (each slightly different) but from the same generating process...

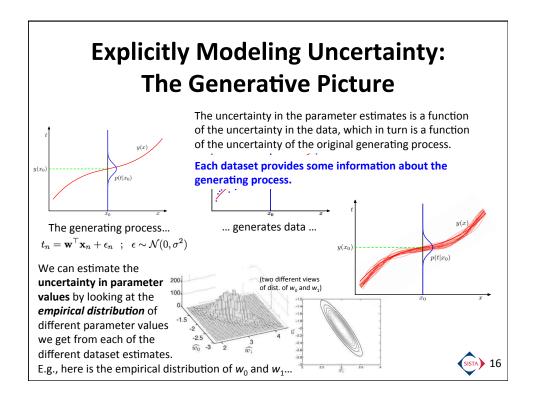
 $t_n = \mathbf{w}^{\top} \mathbf{x}_n + \epsilon_n \; ; \; \epsilon \sim \mathcal{N}(0, \sigma^2)$

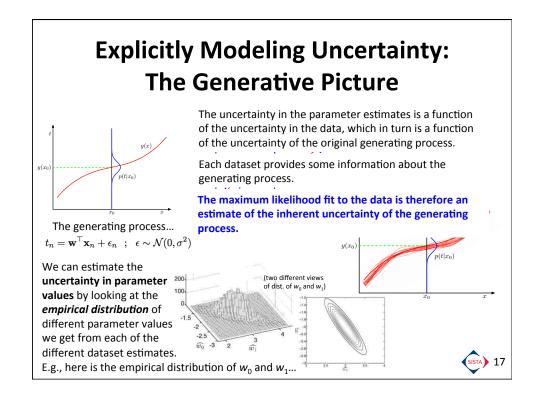
SISTA 12

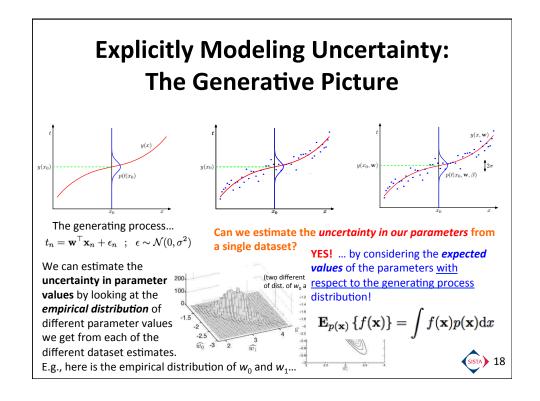












Side Note

$$p(\mathbf{t}|\mathbf{X}, \mathbf{w}, \sigma^2) = \prod_{n=1}^{N} p(t_n|\mathbf{x}_n \mathbf{w}, \sigma^2) = \prod_{n=1}^{N} \mathcal{N}(\mathbf{w}^\mathsf{T} \mathbf{x}_n, \sigma^2)$$

 $p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2) = \mathcal{N}(\mathbf{X}\mathbf{w},\sigma^2\mathbf{I})$ In the following, easier to deal with multivariate Gaussian rather than product of individual Gaussians



Deriving the Expectation of \hat{w} w.r.t. the generating distribution

$$\mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \left\{ \hat{\mathbf{w}} \right\} = \int \hat{\mathbf{w}} p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2) d\mathbf{t}$$

$$\widehat{\mathbf{w}} = (\mathbf{X}^{\mathsf{T}}\mathbf{X})^{-1}\mathbf{X}^{\mathsf{T}}\mathbf{t}$$
 Substitute in $\widehat{\boldsymbol{w}}$

$$\begin{split} \mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \left\{ \widehat{\mathbf{w}} \right\} &= (\mathbf{X}^\mathsf{T}\mathbf{X})^{-1}\mathbf{X}^\mathsf{T} \int \mathbf{t} p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2) d\mathbf{t} \\ \mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \left\{ \widehat{\mathbf{w}} \right\} &= (\mathbf{X}^\mathsf{T}\mathbf{X})^{-1}\mathbf{X}^\mathsf{T} \mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \left\{ \mathbf{t} \right\} \begin{array}{l} p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2) \\ p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2) \end{array} \\ \mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \left\{ \widehat{\mathbf{w}} \right\} &= (\mathbf{X}^\mathsf{T}\mathbf{X})^{-1}\mathbf{X}^\mathsf{T}\mathbf{X}\mathbf{w} \\ \mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \left\{ \widehat{\mathbf{w}} \right\} &= \mathbf{w} \end{split}$$

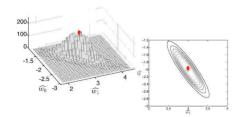
Our estimate $\hat{\boldsymbol{w}}$ of \boldsymbol{w} is inherently unbiased!

This also means any variance in the estimate is encapsulated in its (co)variance (matrix) of $\hat{\mathbf{w}}$

Deriving the Expectation of \hat{w} w.r.t. the generating distribution

$$\mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)}\left\{\hat{\mathbf{w}}
ight\} = \int \hat{\mathbf{w}} p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2) d\mathbf{t}$$

$$\hat{\mathbf{w}} = (\mathbf{X}^{\mathsf{T}}\mathbf{X})^{-1}\mathbf{X}^{\mathsf{T}}\mathbf{t}$$
 Substitute in $\hat{\boldsymbol{w}}$



$$\mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \left\{ \widehat{\mathbf{w}} \right\} = \mathbf{w}$$

Our estimate $\hat{\boldsymbol{w}}$ of \boldsymbol{w} is inherently unbiased!

This also means any variance in the estimate is encapsulated in its (co)variance (matrix) of $\hat{\mathbf{w}}$

Let's be clear on what we mean by "Maximum likelihood of the mean of a linear Gaussian model is unbiased"

- Here we mean: if we take repeated samples of size N from a Gaussian generator distribution with true parameters w, then the collection of maximum likelihood estimated $\hat{\mathbf{w}}$'s will be "centered" (in the sense of the arithmetic mean) around the true w.
- Not to be confused with:
 - "better" estimation by collecting more data (it is true that more data makes better estimation here, but that's a sample size effect)
 - "UMVU" estimator: Uniformly minimum variance unbiased estimator. This is a stronger claim: an UMVU estimator is also an unbiased estimator that is "closest" to the true parameters, for a given sample size.
- And, there could be:
 - Other biased estimators that tend to be closer (from less data)



Derive the Covariance of \hat{w} w.r.t. the generating distribution

$$\begin{split} \mathsf{cov}\{\mathbf{x}\} &= \mathbf{E}_{P(\mathbf{x})} \left\{ \mathbf{x} \mathbf{x}^\mathsf{T} \right\} - \mathbf{E}_{P(\mathbf{x})} \left\{ \mathbf{x} \right\} \mathbf{E}_{P(\mathbf{x})} \left\{ \mathbf{x} \right\}^\mathsf{T} \\ \mathsf{cov}\{\widehat{\mathbf{w}}\} &= \mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \left\{ \widehat{\mathbf{w}} \widehat{\mathbf{w}}^\mathsf{T} \right\} - \mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \left\{ \widehat{\mathbf{w}} \right\} \mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \left\{ \widehat{\mathbf{w}} \right\}^\mathsf{T} \\ &= \mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \left\{ \widehat{\mathbf{w}} \widehat{\mathbf{w}}^\mathsf{T} \right\} - \mathbf{w} \mathbf{w}^\mathsf{T} \\ &= \mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \left\{ \widehat{\mathbf{w}} \right\} = \mathbf{w} \end{split}$$

$$\begin{split} \mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \left\{ \widehat{\mathbf{w}} \widehat{\mathbf{w}}^\mathsf{T} \right\} &= \mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \left\{ ((\mathbf{X}^\mathsf{T}\mathbf{X})^{-1}\mathbf{X}^\mathsf{T}\mathbf{t})((\mathbf{X}^\mathsf{T}\mathbf{X})^{-1}\mathbf{X}^\mathsf{T}\mathbf{t})^\mathsf{T} \right\} \\ &= (\mathbf{X}^\mathsf{T}\mathbf{X})^{-1}\mathbf{X}^\mathsf{T} \ \mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \left\{ \mathbf{t}\mathbf{t}^\mathsf{T} \right\} \ \mathbf{X}(\mathbf{X}^\mathsf{T}\mathbf{X})^{-1}. \end{split}$$

$$= \mathbf{X}\mathbf{w}(\mathbf{X}\mathbf{w})^{\mathsf{T}} + \sigma^{2}\mathbf{I} \qquad \qquad \mathsf{cov}\{\widehat{\mathbf{w}}\} = \mathbf{w}\mathbf{w}^{\mathsf{T}} + \sigma^{2}(\mathbf{X}^{\mathsf{T}}\mathbf{X})^{-1} - \mathbf{w}\mathbf{w}^{\mathsf{T}}$$

$$= \mathbf{w}\mathbf{w}^{\mathsf{T}}\mathbf{X}^{\mathsf{T}} + \sigma^{2}\mathbf{I}. \qquad \qquad = \sigma^{2}(\mathbf{X}^{\mathsf{T}}\mathbf{X})^{-1}$$

$$= \mathbf{w}^{\mathsf{T}} + \sigma^{2}(\mathbf{X}^{\mathsf{T}}\mathbf{X})^{-1}\mathbf{X}^{\mathsf{T}}\mathbf{X}\mathbf{w}\mathbf{w}^{\mathsf{T}}\mathbf{X}^{\mathsf{T}}\mathbf{X}(\mathbf{X}^{\mathsf{T}}\mathbf{X})^{-1}$$

$$= \mathbf{w}\mathbf{w}^{\mathsf{T}} + \sigma^{2}(\mathbf{X}^{\mathsf{T}}\mathbf{X})^{-1}. \qquad \qquad \mathsf{cov}\{\widehat{\mathbf{w}}\} = \sigma^{2}(\mathbf{X}^{\mathsf{T}}\mathbf{X})^{-1} = -\left(\frac{\partial^{2}\log L}{\partial \mathbf{w}\partial \mathbf{w}^{\mathsf{T}}}\right)^{-1} \ge 23$$

The Fisher Information

$$\mathsf{cov}\{\widehat{\mathbf{w}}\} = \sigma^2(\mathbf{X}^\mathsf{T}\mathbf{X})^{-1} = -\left(\frac{\partial^2\!\log L}{\partial\mathbf{w}\partial\mathbf{w}^\mathsf{T}}\right)^{-1}$$

$$\mathcal{I} = \mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \left\{ -\frac{\partial^2 \log p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)}{\partial \mathbf{w} \partial \mathbf{w}^\mathsf{T}} \right\}$$

$$\frac{\partial^2 \log L}{\partial \mathbf{w} \partial \mathbf{w}^{\mathsf{T}}} = -\frac{1}{\sigma^2} \mathbf{X}^{\mathsf{T}} \mathbf{X}$$

$$\mathcal{I} = \mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \left\{ \frac{1}{\sigma^2} \mathbf{X}^\mathsf{T} \mathbf{X} \right\}$$

 $\mathcal{I} = \frac{1}{\sigma^2} \mathbf{X}^\mathsf{T} \mathbf{X}$. The elements of I tell us how much information the data provides about the particular parameter (diagonal elements) or pairs of parameters (off-diagonal elements).



