# ISTA 421/521 – Homework 4

Due: Thursday, November 9, 5pm 20 pts total for Undergrads, 25 pts total for Grads

STUDENT NAME

Undergraduate / Graduate

## Instructions

In this assignment, exercises 3 and 4 require you to write small python scripts; the details for those scripts, along with their .py name are described in the exercises. All of the exercises in this homework require written derivations, so you will also submit a .pdf of your written answers. (You can use LATEX or any other system (including handwritten; plots, of course, must be program-generated) as long as the final version is in PDF.)

The final submission will include (minimally) the two scripts you need to write for problems 3 and 4, and a PDF version of your written part of the assignment. You are required to create either a .zip or tarball (.tar.gz / .tgz) archive of all of the files for your submission and submit your archive to the d2l dropbox by the date/time deadline above.

NOTE: Problem 1 is required for Graduate students only; Undergraduates may complete this problem for extra credit equal to the point value.

(FCMA refers to the course text: Rogers and Girolami (2016), A First Course in Machine Learning, second edition. For general notes on using I<sup>A</sup>T<sub>E</sub>X to typeset math, see:

http://en.wikibooks.org/wiki/LaTeX/Mathematics)

## 1. [5 points; Required only for Graduates] Adapted from Exercise 3.12 of FCMA p.135;

When performing a Bayesian analysis of the Olympics data, we assumed that  $\sigma^2$  was known. If instead we assume that **w** is known and an inverse Gamma prior is placed on  $\sigma^2$ 

$$p(\sigma^2|\alpha,\beta) = \frac{\beta^{\alpha}}{\Gamma(\alpha)}(\sigma^2)^{-\alpha-1} \exp\left\{-\frac{\beta}{\sigma^2}\right\},\,$$

the posterior over  $\sigma^2$  will also be inverse Gamma. Derive the posterior parameters.

#### Solution.

## 2. [6 points] Adapted from Exercise 4.2 of FCMA p.163:

In Chapter 3, we computed the posterior density over r, the probability of a coin giving heads, using a beta prior and a binomial likelihood. Recalling that the beta prior, with parameters  $\alpha$  and  $\beta$ , is given by

$$p(r|\alpha,\beta) = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} r^{\alpha-1} (1-r)^{\beta-1}$$

and the binomial likelihood, assuming y heads in N throws, is given by

$$p(y|r, N) = \binom{N}{y} r^y (1-r)^{N-y}$$

compute the Laplace approximation to the posterior. (Note, you should be able to obtain a closed-form solution for the MAP value,  $\hat{r}$ , by getting the log posterior, differentiating (with respect to r), equating to zero and solving for r.)

#### Solution.

#### 3. [4 points] Adapted from Exercise 4.3 of FCMA p.163:

In the previous exercise you computed the Laplace approximation to the true beta posterior. In this problem, plot both the true beta posterior and the Laplace approximation for the following three parameter settings:

- 1.  $\alpha = 5$ ,  $\beta = 5$ , N = 20, and y = 10,
- 2.  $\alpha = 3$ ,  $\beta = 15$ , N = 10, and y = 3,
- 3.  $\alpha = 1$ ,  $\beta = 30$ , N = 10, and y = 3.

Be sure to clearly indicate the values in your plot captions. Include how the two distributions (the true beta posterior and the Laplace approximation) compare in each case. Include the python script you use to generate these plots; the script should be named plot\_laplace\_approx.py. Suggestion: for plotting the beta and Gaussian (Normal) distributions, you can use scipy.stats.beta and scipy.stats.normal to create the beta and Gaussian random variables, and use the pdf(x) method for each to generate the curves. Note that for scipy.stats.normal, the mean is the location (loc) parameter, and the sigma is the scale parameter. Also, scipy.stats.normal expects the scale parameter to be the standard deviation (i.e., take the square root: math.sqrt(x)) of the variance you'll compute for the Laplace approximation.

## Solution.

#### 4. [4 points] Adapted from Exercise 4.4 of FCMA p.164:

Given the expression for the area of a circle,  $A = \pi r^2$ , and using only uniformly distributed random variates, devise a sampling approach for computing  $\pi$ . Describe your method in detail and provide your

script to do the estimation – this script should be called pi\_sample\_estimate.py. Report your estimate based on 1 million samples to 6 decimal places. (NOTE: You do not need to use Metropolis-Hastings to compute this.)

## Solution.

5. [6 points] Adapted from Exercise 4.6 of FCMA p.164:

Assume that we observe N vectors of attributes,  $\mathbf{x}_1, ..., \mathbf{x}_N$  and associated integer counts  $t_1, ..., t_N$ . A Poisson likelihood would be suitable:

$$p(t_n|\mathbf{x}_n, \mathbf{w}) = \frac{f(\mathbf{x}_n; \mathbf{w})^{t_n} \exp\{-f(\mathbf{x}_n; \mathbf{w})\}}{t_n!},$$

where  $f(\mathbf{x}_n; \mathbf{w}) = \mathbf{w}^{\top} \mathbf{x}_n$ . Assuming a zero-mean Gaussian prior on  $\mathbf{w}$  with constant diagonal covariance of  $\sigma^2$ , derive the gradient and Hessian of the posterior. Using these, express the parameter update rules for (a) gradient *ascent* (because we're maximizing) update (in class we looked at Widrow-Hoff, which is typically expressed for *descent*), and (b) Newton-Raphson.

The following facts will help in the derivation. First, keep in mind that  $\mathbf{w}^{\top}\mathbf{x}_n$  is a *scalar* value. This means you can take the partial derivative of  $\log \mathbf{w}^{\top}\mathbf{x}_n$  with respect to  $\mathbf{w}$ . Also, remember that the Hessian is a matrix representing the second partial derivatives of the gradient with respect to itself (see Comment 2.6 of p.73), and the second derivative will involve the *transpose* of the partial with respect to  $\mathbf{w}$ . So, e.g., as part of taking the second derivative, if you are taking the transpose derivative part of  $\mathbf{w}^{\top}\mathbf{x}_n$ , as follows:

$$\frac{\partial (\mathbf{w}^{\top} \mathbf{x}_n)}{\partial \mathbf{w}} = \mathbf{x}_n \quad \text{ and } \quad \frac{\partial (\mathbf{w}^{\top} \mathbf{x}_n)}{\partial \mathbf{w}^{\top}} = \mathbf{x}_n^{\top}$$

Solution.