

An introduction to the bayesc package

September 24, 2025

- 1 Introduction
 - 1.1 Background and statistical details
 - 1.1.1 Bayesian latent class analysis
 - 1.1.2 Outcome modeling
 - 1.1.3 Accounting for survey design
 - 1.2 Package installation
- 2 An example of functionality using NHANES data
 - 2.1 Prepare NHANES data
 - 2.1.1 Load NHANES dataset
 - 2.1.2 Define observed exposure and outcome data
 - 2.1.3 Define survey design variables
 - 2.1.4 Define additional regression variables
 - 2.2 Run main Bayesian clustering functions for survey data
 - 2.2.1 Use `wolca()` and `wolca_var_adjust()` to derive dietary exposure clusters
 - 2.2.2 Use `wolca_svyglm()` to implement two-step approach relating pattern profiles to a binary hypertension outcome
 - 2.2.3 Use `swolca()` and `swolca_var_adjust()` to implement alternative one-step approach for deriving dietary exposure clusters informed by a binary hypertension outcome
 - 2.2.4 Option to adapt `wolca()` and `swolca()` for unweighted analyses when no survey weights are available
 - 2.3 Summarize and plot output from running `swolca()`
 - 2.3.1 Plot the pattern profiles
 - 2.3.2 Obtain individual class assignments
 - 2.3.3 Reorder classes
 - 2.3.4 Plot distribution of classes
 - 2.3.5 Display distribution of additional variables across classes
 - 2.3.6 Summarize and plot the exposure-outcome relationship
 - 2.4 Obtain performance diagnostics and additional output
- 3 References
- 4 Supplementary installation instructions

1 Introduction

The `bayesc` package provides functionality for running clustering analyses incorporating survey sampling weights and other survey design elements, implemented through Bayesian techniques that allow the number of clusters to be determined by the data while parameter estimation occurs. The clustering analysis implemented is a Bayesian latent class analysis (LCA) that enables elicitation of underlying pattern profiles in the setting where the correlated variables are categorical, while accounting for survey

design. We use the terms “latent class” or “class” to refer to the underlying, unknown clusters that are discovered through the latent class analysis.

When conducting Bayesian LCA using survey data, it is important to account for complex survey design elements such as stratification, clustered sampling, and unequal sampling probabilities. Otherwise, biased parameter estimation and inference can occur and results will lack generalizability to the broader population. To account for survey data, we implement the weighted pseudo-likelihood approaches described in Stephenson, Wu, and Dominici (2023) and Wu et al. (2024), based on work by Savitsky and Toth (2016) and Williams and Savitsky (2021).

The `wolca()` function (Section 2.2.1) implements a survey-weighted Bayesian LCA and can be followed by `wolca_var_adjust()` for correct uncertainty estimation. Options also exist for relating the pattern profiles to a binary outcome. This can either be done through a two-step approach after clustering is completed using the `wolca_svyglm()` function described in Section 2.2.2), or in a one-step supervised approach where creation of the pattern profiles is directly informed by the outcome using the `swolca()` function described in Section 2.2.3, followed by `swolca_var_adjust()` for correct uncertainty estimation. The two-step approach is advantageous for relating the same set of patterns to a variety of different outcomes and is computationally faster. The one-step approach is advantageous for propagating classification measurement error, detecting smaller exposure-outcome effects, and discovering patterns where a small subset of foods are highly influential for the outcome. Summary and plotting functions for visualizing output are available, as are diagnostic functions for examining convergence of the MCMC sampler (Sections 2.3 and 2.4). Full statistical details for the models are provided in Wu et al. (2024).

We provide an example dataset from the National Health and Nutrition Examination Survey (NHANES) that includes dietary intake and hypertension data for low-income women in the United States, as well as survey sampling weights and information on stratification and clustered sampling in the survey design (National Center for Health Statistics 2023).

1.1 Background and statistical details

1.1.1 Bayesian latent class analysis

Latent class analysis (LCA), also known as a finite mixture model for categorical data, summarizes data from correlated categorical variables into a few underlying patterns that capture key differences between individuals in the data. Each individual is assigned to exactly one latent class (i.e., cluster). For each class, a pattern profile is created from the most probable category of each input variable. For instance, if the multivariate categorical input is food items consumed at levels “None”, “Low”, “Medium”, and “High”, then each class has a pattern profile consisting of the most probable consumption level for every food item among those belonging to that class, and these pattern profiles can be interpreted as dietary patterns.

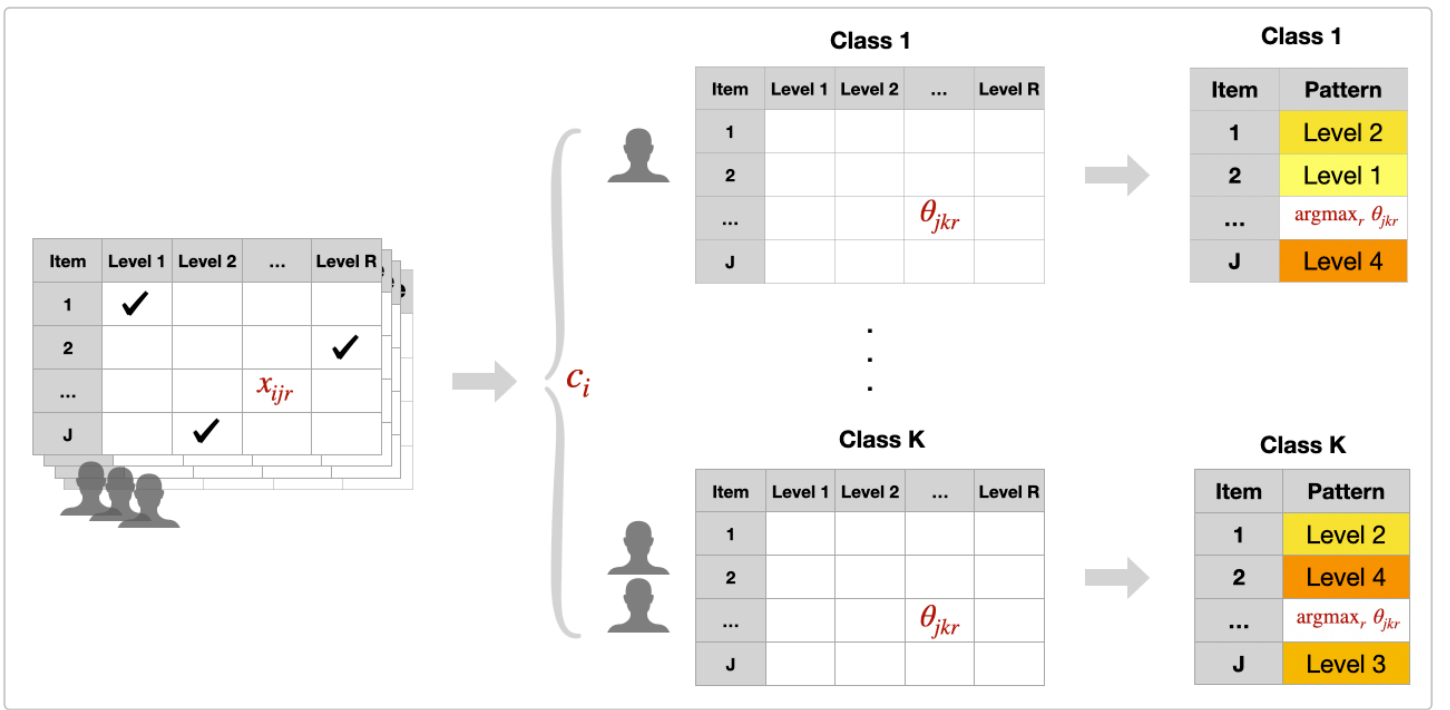


Figure: Diagram of latent class analysis

Let $i \in \{1, \dots, n\}$ index individuals in the sample, $j \in \{1, \dots, J\}$ index items, and $r \in \{1, \dots, R_j\}$ index levels for item j . Latent class membership is characterized by $c_i \in \{1, \dots, K\}$ for individual i , where K denotes the number of latent classes. Parameter π_k is the probability of membership in class k , with $\sum_{k=1}^K \pi_k = 1$. The exposure is a multivariate categorical variable with J items, $(x_{i1}, \dots, x_{iJ})^T$, where each item follows a Multinomial distribution with R_j levels, each observed with item level probability θ_{jkr} , with $\sum_{r=1}^{R_j} \theta_{jkr} = 1$ for all $j \in \{1, \dots, J\}$ and $k \in \{1, \dots, K\}$. LCA relies on the local independence assumption that items are independent conditional on dietary pattern assignment, and the global clustering assumption that individuals assigned to the same pattern share behaviors for all food items. The individual complete likelihood is given by

$$\prod_{k=1}^K \left\{ \pi_k \prod_{j=1}^J \prod_{r=1}^{R_j} \theta_{jkr}^{I(x_{ij}=r)} \right\}^{I(c_i=k)}.$$

Bayesian LCA differs from traditional LCA by allowing the number of latent classes to be chosen by the data, rather than determined by the researcher after multiple model fits. The Bayesian formulation also offers advantages when it comes to accounting for survey design with sparse data. We refer to the Bayesian LCA in this package as an overfitted LCA (OLCA), where a conservatively high number of classes is specified and then empty and unnecessary classes drop out via a sparsity-inducing Dirichlet prior. Suppose K_{\max} is a conservative upper bound for the number of classes (e.g., 30). Let $\pi_1, \dots, \pi_{K_{\max}}$ denote the class assignment probabilities for the K_{\max} classes. We set

$$(\pi_1, \dots, \pi_{K_{\max}}) \sim \text{Dir}(\alpha_1, \dots, \alpha_{K_{\max}})$$

where hyperparameters α_k moderate the rate of growth for nonempty classes, with smaller values yielding a slower rate of growth and more sparsity. By default, we set $\alpha_k = 1/K_{\max}$ for all k , but users can specify other hyperparameter values if desired by using the `alpha_adapt` argument in the `wolca()` and `swolca()` functions. When finalizing the number of classes, we drop any classes that are duplicates, and retain any classes that are above a prespecified threshold in size, using the argument `class_cutoff`. By default, we have `class_cutoff = 0.05`, which drops any classes that are observed in less than 5% of the

population.

1.1.2 Outcome modeling

Interest may also lie in examining the relationship between the latent class pattern profiles and a binary outcome. This can be done in two ways: 1) a two-step “classify-analyze” approach that performs a regression after clustering is completed (Section 2.2.2), or 2) a one-step approach where creation of the pattern profiles is directly informed by the outcome (Section 2.2.3).

We provide brief statistical details for the one-step approach, also referred to as a supervised approach. Let y_i denote the binary outcome of interest and $\mathbf{v}_{i\cdot} = (v_{i1}, \dots, v_{iQ})$ denote a $Q \times 1$ vector of additional regression covariates for individual i , corresponding to a $Q \times 1$ vector of regression coefficients $\boldsymbol{\xi}_{k\cdot} = (\xi_{k1}, \dots, \xi_{kQ})^T$ given assignment to latent class k . The full set of regression coefficients can be written as a $K \times Q$ matrix that follows mixture reference coding. Further details can be found in Wu et al. (2024). The individual complete likelihood in the supervised setting is given by

$$\prod_{k=1}^K \left\{ \pi_k \prod_{j=1}^J \prod_{r=1}^{R_j} \theta_{jkr}^{I(x_{ij}=r)} \Phi(\mathbf{v}_{i\cdot}^T \boldsymbol{\xi}_{k\cdot})^{y_i} [1 - \Phi(\mathbf{v}_{i\cdot}^T \boldsymbol{\xi}_{k\cdot})]^{(1-y_i)} \right\}^{I(c_i=k)}.$$

1.1.3 Accounting for survey design

Survey weights are incorporated into parameter estimation by up-weighting each individual’s likelihood contribution proportional to the number of individuals they represent in the target population, and then using the weighted pseudo-likelihood in the Bayesian posterior update. Let \tilde{w}_i denote the $n \times 1$ vector of survey sampling weights, normalized to sum to the sample size to account for sampling uncertainty.

$$p_{wid}(\theta|D) \propto p(\theta) \prod_{i=1}^n p(D_i|\theta)^{\tilde{w}_i}$$

Using the sampling weights to exponentiate the unit likelihood contributions constructs a pseudo or composite likelihood that approximates the population generating likelihood on the observed sample (Savitsky and Toth 2016). The posterior distributions created by this will create variance estimates that are smaller than their true value. In order to obtain accurate variance estimates, a post-processing variance adjustment proposed by Williams and Savitsky (2021) is used post-hoc to recover the true variance based on pseudo-MLE theory. A rescaling and reshaping transformation to the posterior estimates is applied to obtain adjusted estimates that have approximately correct variance.

1.2 Package installation

The `baysc` package can be installed and loaded by running the following code.

```
# If R package `devtools` is not installed, uncomment the following line:
# `install.packages(devtools)`

# Load devtools and install baysc from GitHub
library(devtools)
devtools::install_github("smwu/baysc")
```

```
library(bayesc)
```

```
# Load additional packages for this vignette
```

```
library(dplyr)
```

```
library(ggplot2)
```

```
library(knitr)
```

During installation, the following errors may arise:

- *No package called ‘rstantools’*: Please install the `rstantools` package using `install.packages("rstantools")`.
- *Library ‘gfortran’ not found*: This is a compiler configuration issue that can arise when using Rcpp on Mac computers with Apple silicon (e.g., M1, M2, M3). Users may need to install Xcode, GNU Fortran, and OpenMP, and edit the `~/R/Makevars` file. For more details, see the [Supplementary section](#).

2 An example of functionality using NHANES data

2.1 Prepare NHANES data

2.1.1 Load NHANES dataset

The example NHANES dataset contains dietary intake as a multivariate categorical exposure consisting of 28 food item groups consumed at levels “None”, “Low”, “Medium”, and “High”, and hypertension as a binary outcome, for $n = 2004$ low-income women in the United States. More information on the dataset preparation can be found in the data documentation and in Wu et al. (2024). The package functions can also handle the scenario where the number of levels differs for different food item groups.

To load the dataset, run the following code:

```
# Load NHANES dataset
```

```
data("data_nhanes")
```

```
# Number of individuals
```

```
nrow(data_nhanes)
```

```
## [1] 2004
```

2.1.2 Define observed exposure and outcome data

```
# Exposure matrix of categorical intake for each individual across 28 food groups
```

```
x_mat <- data_nhanes[, c("citrus", "oth_fruit", "fruit_juice", "dark_green", "tomatoes",
  "oth_red", "potatoes", "oth_starchy", "oth_veg", "whole_grain", "ref_grain",
  "meat", "cured_meats", "organ", "poultry", "seafood_high", "seafood_low", "eggs",
  "soybean", "nuts", "leg_protein", "milk", "yogurt", "cheese", "oils",
  "solid_fats", "add_sugars", "drinks")]
x_mat <- as.matrix(x_mat)
# Binary outcome of hypertension
y_all <- data_nhanes$BP_flag
```

2.1.3 Define survey design variables

Define survey design variables on stratification, clustered sampling, and sampling weights. If there is no survey design information, these variables can be left as NULL (default). For the cluster indicators, use indicators for the first stage of sampling and do not nest the indicators within strata. The variance computation will only use the first stage of clustering and will assume sampling to be with replacement at the first stage.

```
# Survey stratum indicators
stratum_id <- data_nhanes$stratum_id
# Survey cluster indicators
cluster_id <- data_nhanes$cluster_id
# Survey sampling weights
sampling_wt <- data_nhanes$sample_wt
```

2.1.4 Define additional regression variables

Create a dataframe of additional confounders to adjust for when evaluating the diet-hypertension exposure-outcome association, and define a regression formula for including these additional confounders. In this example, we adjust for age, race/ethnicity, smoking status, and physical activity.

```
# Create dataframe of additional confounders
V_data <- data_nhanes %>% select(age_cat, racethnic, smoker, physactive)
# Regression formula for additional confounders
glm_form <- "~ age_cat + racethnic + smoker + physactive"
```

2.2 Run main Bayesian clustering functions for survey data

The following sections describe how to use the main functions in the `baysc` package for running survey-weighted Bayesian LCA. If no outcome information is necessary when clustering, use the unsupervised weighted OLCA, implemented in the `wolca()` function ([Section 2.2.1](#)). If clustering is to be informed by a binary outcome so that latent classes are found that specifically drive the outcome of interest, use the supervised weighted OLCA, implemented in the `swolca()` function ([Section 2.2.3](#)).

2.2.1 Use `wolca()` and `wolca_var_adjust()` to derive dietary exposure clusters

We use the `wolca()` function to run an unsupervised weighted overfitted latent class analysis and obtain

pattern profiles. The `wolca_var_adjust()` function is used after to apply a post-processing adjustment that allows for correct estimation of the variability of the parameter estimates. If this adjustment is not applied, the estimated variances will be too small. We run the sampler for 300 iterations as an example, with a total runtime of around 5 minutes for `wolca()` and `wolca_var_adjust()`. Typically, at least 20000 iterations should be used for `wolca()`, corresponding to a runtime of around 2 hours.

When running the model, there are two samplers that can be run: an adaptive sampler and a fixed sampler. The adaptive sampler is used to determine the number of latent classes using a data-driven process with a sparsity-inducing Dirichlet prior. Once the number of latent classes is determined, the fixed sampler is run using the set number of classes and produces output for estimation and inference. Note that the final number of classes determined by the fixed sampler may be smaller than the number of classes obtained from the adaptive sampler.

```
# Run WOLCA
```

```
res_wolca <- wolca(x_mat = x_mat, sampling_wt = sampling_wt,  
                  cluster_id = cluster_id, stratum_id = stratum_id,  
                  run_sampler = "both", K_max = 30, adapt_seed = 999,  
                  class_cutoff = 0.05, n_runs = 300, burn = 150, thin = 3,  
                  update = 20, save_res = FALSE)
```

```
# Apply variance adjustment
```

```
res_wolca_adjust <- wolca_var_adjust(res = res_wolca, adjust_seed = 999,  
                                    num_reps = 100, save_res = FALSE)
```

Parameters used for `wolca()`:

- `x_mat`: $n \times J$ matrix of J categorical exposure variables for sampled individuals.
- `sampling_wt`: $n \times 1$ vector of survey sampling weights.
- `cluster_id`: $n \times 1$ vector of cluster indicators from clustered sampling.
- `stratum_id`: $n \times 1$ vector of stratum indicators.
- `run_sampler = "both"`: Run both the adaptive and fixed sampler (default). Other options are "adapt" for only the adaptive sampler, or "fixed" for only the fixed sampler when the number of latent classes is known a priori.
- `K_max = 30`: The maximum number of latent classes is set to 30 (default). Typically, the true number will be much smaller.
- `adapt_seed = 999`: Random seed for adaptive sampler, set to 999 for this example but with default NULL.
- `class_cutoff = 0.05`: The minimum size of each latent class, with default set to 5% of the population.
- `n_runs = 300`: Number of MCMC sampler iterations to run. Set to 300 for this example but with default 20000. *Typically, at least 20000 iterations should be used to allow the sampler to converge.*
- `burn = 150`: Number of iterations for burn-in. Set to 150 for this example but with default 10000.
- `thin = 5`: Thin the sampler to keep every 5th iteration (default).
- `update = 20`: Print out a MCMC progress update every 20 iterations (default is 10).
- `save_res = FALSE`: Results do not need to be saved. If `save_res = TRUE` (default), `save_path` must also be specified and include both the desired directory and the beginning of the file name for the results.

Parameters used for `wolca_var_adjust()`:

- `res = res_wolca_unadj`: Results from calling `wolca()`.
- `adjust_seed = 999`: Random seed for the adjustment procedure, set to 999 for this example but with default NULL.

- `num_reps = 100`: Use 100 bootstrap replicates (default) to estimate variability in the sampling process.
- `save_res = FALSE`: Results do not need to be saved. If `save_res = TRUE` (default), `save_path` must also be specified and include both the desired directory and the beginning of the file name for the results.

Additional parameters are available for prior specification and other customization and are described in the documentation help pages for the functions.

Users can ignore the following warning messages: “the design is sampled with replacement and only the first stage is used” and “the number of chains being less than 1 and sampling not done.” Messages about negative variances and large (>1) eigenvalue differences to the nearest positive definite matrix indicate instability in the variance adjustment matrix inversions. Methods to address this issue include running the sampler for more iterations or reducing the number of additional covariate terms in the outcome regression. One could also choose to forgo the variance adjustment and simply use the posterior distributions produced by `wolca()`, acknowledging that estimated variances will be slight underestimates.

2.2.2 Use `wolca_svyglm()` to implement two-step approach relating pattern profiles to a binary hypertension outcome

After the pattern profiles have been created, if interest lies in relating these pattern profiles to a binary outcome, such as presence or absence of hypertension (i.e., high blood pressure), then a crude approach is to apply a two-step approach, also known as the “classify-analyze” paradigm. Step 1 involves an unsupervised clustering as described in [Section 2.2.1](#). Step 2 involves a survey-weighted regression model using the `svyglm()` function in the `survey` package (Lumley 2023). One word of caution: this approach can result in imprecise estimates and estimates that are attenuated due to the measurement error in the first classification step that is unaccounted for. See Nylund-Gibson, Grimm, and Masyn (2019) for more discussion of this issue.

The following code can be used to find the association between latent classes and a binary outcome:

```
# Run weighted outcome regression model
res_wolca_svyglm <- wolca_svyglm(res = res_wolca_adjust, y_all = y_all,
                                V_data = V_data, glm_form = glm_form,
                                ci_level = 0.95, save_res = FALSE)
```

Parameters used:

- `res = res_wolca_adjust`: Results from calling `wolca()` and `wolca_var_adjust()`.
- `y_all`: $n \times 1$ vector of binary outcomes for sampled individuals.
- `V_data`: Dataframe of additional covariates to adjust for when evaluating the exposure-outcome association. Default is `NULL`, but we use age, race/ethnicity, smoking status, and physical activity for this example.
- `glm_form`: String specifying the for additional covariates in the outcome regression. *Interaction terms between class membership and all additional covariates are automatically included as regression terms.*
- `ci_level = 0.95`: Survey-weighted confidence intervals should be evaluated at the 95% confidence level (default).
- `save_res = FALSE`: Results do not need to be saved. If `save_res = TRUE` (default), `save_path` must also be specified and include both the desired directory and the beginning of the file name for the results.

A warning about negative residual degrees of freedom may arise when the survey design degrees of freedom is smaller than the number of covariates in the model. Confidence intervals and p-values are by default not reported in the `svyglm()` function. However, when the covariates are not cluster-level covariates, it is approximately reasonable to use the survey design degrees of freedom instead, which is what the `baysc` package does to calculate confidence intervals and p-values. Please see the `survey` package documentation for more discussion on this issue.

2.2.3 Use `swolca()` and `swolca_var_adjust()` to implement alternative one-step approach for deriving dietary exposure clusters informed by a binary hypertension outcome

As an alternative to the two-step approach in [Section 2.2.2](#), a one-step “supervised” approach allows information about the binary outcome to directly inform the creation of the pattern profiles. We use the `swolca()` function to obtain dietary pattern profiles while simultaneously relating them to the hypertension outcome. The `swolca_var_adjust()` function is used after to apply a post-processing adjustment that allows for correct estimation of the variability of the parameter estimates. If this adjustment is not applied, the estimated variances will be too small. We run the sampler for 100 iterations as an example, with a total runtime of around 8 minutes for `swolca()` and `swolca_var_adjust()`. Typically, at least 20000 iterations should be used for `swolca()`, corresponding to a runtime of around 3 hours.

If continuous variables are incorporated into the binary outcome model, any such variables with standard deviation greater than 5 may result in errors in the variance adjustment. To avoid these errors, consider standardizing the variable to have mean 0 and standard deviation 1 or converting the variable into a categorical form.

```
# Run SWOLCA
```

```
res_swolca <- swolca(x_mat = x_mat, y_all = y_all, V_data = V_data,  
                    glm_form = glm_form, sampling_wt = sampling_wt,  
                    cluster_id = cluster_id, stratum_id = stratum_id,  
                    run_sampler = "both", K_max = 30, adapt_seed = 888,  
                    class_cutoff = 0.05, n_runs = 300, burn = 150, thin = 3,  
                    update = 20, save_res = FALSE)
```

```
# Apply variance adjustment
```

```
res_swolca_adjust <- swolca_var_adjust(res = res_swolca, adjust_seed = 888,  
                                       num_reps = 100, save_res = FALSE)
```

Parameters used for `swolca()`:

- `x_mat`: nxJ matrix of J categorical exposure variables for sampled individuals.
- `y_all`: nx1 vector of binary outcomes.
- `V_data`: nxQ dataframe of additional regression covariates to adjust for when evaluating the exposure-outcome association. Default is NULL, but we use age, race/ethnicity, smoking status, and physical activity for this example.
- `glm_form`: String specifying regression formula for additional covariates. *Interaction terms between class membership and all additional covariates are automatically included as regression terms.*
- `sampling_wt`: nx1 vector of survey sampling weights.
- `cluster_id`: nx1 vector of cluster indicators from clustered sampling.
- `stratum_id`: nx1 vector of stratum indicators.
- `run_sampler = "both"`: Run both the adaptive and fixed sampler (default). Other options are "adapt" for only the adaptive sampler, or "fixed" for only the fixed sampler when the number of

latent classes is known a priori.

- `K_max = 30`: The maximum number of latent classes is set to 30 (default). Typically, the true number will be much smaller.
- `adapt_seed = 888`: Random seed for adaptive sampler, set to 888 for this example but with default NULL.
- `class_cutoff = 0.05`: The minimum size of each latent class, with default set to 5% of the population.
- `n_runs = 300`: Number of MCMC sampler iterations to run. Set to 300 for this example but with default 20000. *Typically, at least 20000 iterations should be used to allow the sampler to converge.*
- `burn = 150`: Number of iterations for burn-in. Set to 150 for this example but with default 10000.
- `thin = 5`: Thin the sampler to keep every 5th iteration (default).
- `update = 20`: Print out a MCMC progress update every 20 iterations (default is 10).
- `save_res = FALSE`: Results do not need to be saved. If `save_res = TRUE` (default), `save_path` must also be specified and include both the desired directory and the beginning of the file name for the results.

Parameters used for `swolca_var_adjust()`:

- `res = res_swolca_unadj`: Results from calling `swolca()`.
- `adjust_seed = 888`: Random seed for the adjustment procedure, set to 888 for this example but with default NULL.
- `num_reps = 100`: Use 100 bootstrap replicates (default) to estimate variability in the sampling process.
- `save_res = FALSE`: Results do not need to be saved. If `save_res = TRUE` (default), `save_path` must also be specified and include both the desired directory and the beginning of the file name for the results.

Additional parameters are available for prior specification and other customization and are described in the documentation help pages for the functions.

When running the `swolca_var_adjust()` variance adjustment, some warning messages may arise to indicate instability in the variance adjustment matrix inversions. Possible solutions include running the sampler for more iterations or reducing the number of additional covariate terms in the outcome regression. If continuous covariates are included in the probit regression model, these errors may also be due to the variance of the covariate being too large. In such cases, it is recommended to reduce the variance of the covariate, either by standardizing the variable to have mean 0 and standard deviation 1, or by transforming the variable into a categorical variable. One could also choose to forgo the variance adjustment and simply use the posterior distributions produced by `swolca()`, acknowledging that estimated variances will be slight underestimates.

2.2.4 Option to adapt `wolca()` and `swolca()` for unweighted analyses when no survey weights are available

If no survey weights are available and we would like to perform analyses without accounting for survey design, we can simply leave the input arguments of `sampling_wt`, `cluster_id`, and `stratum_id` blank, since they are by default set to NULL. This will automatically set all weights equal to 1, corresponding to an unweighted model. Since no survey design information is available, the `wolca_var_adjust()` and `swolca_var_adjust()` functions cannot be run and the estimates will not be correct for the broader population, only for the study sample. Variances may also be underestimated due to unaccounted-for clustered sampling.

```
# Run SWOLCA without survey weights
```

```
res_swolca_unwt <- swolca(x_mat = x_mat, y_all = y_all, V_data = V_data,  
  glm_form = glm_form, run_sampler = "both", K_max = 30,  
  adapt_seed = 888, class_cutoff = 0.05,  
  n_runs = 300, burn = 150, thin = 3, save_res = FALSE)
```

2.3 Summarize and plot output from running `swolca()`

We demonstrate plotting and summarization functions using the `res_swolca_adjust` results from running `swolca()` and `swolca_var_adjust()`, but the same functions can also be used for summarizing output from `wolca()` and `wolca_var_adjust()`.

2.3.1 Plot the pattern profiles

For each latent class, there is a pattern profile defined by the category with the highest posterior probability, also referred to as the modal θ , for each item. To visualize these pattern profiles, use the `plot_pattern_profiles` function.

```
plot_pattern_profiles(res = res_swolca_adjust)
```

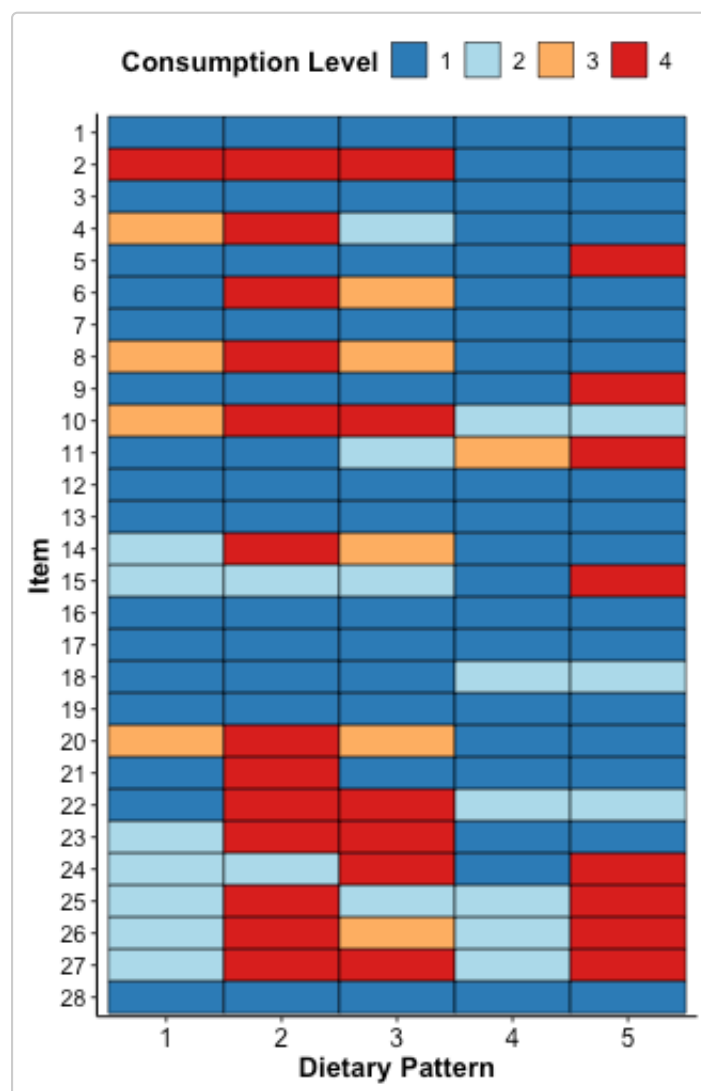


Figure: Latent class pattern profiles

We can also specify labels for the items, categories, and latent classes

```
# Item labels
item_title <- "Item"
item_labels <- c("Citrus/Melon/Berries", "Other Fruits", "Fruit Juice",
  "Dark Green Veggies", "Tomatoes", "Other Red/Orange Veggies",
  "Potatoes", "Other Starchy Veggies", "Other Vegetables",
  "Whole Grains", "Refined Grains", "Meat", "Cured Meats",
  "Organ Meat", "Poultry", "Seafood (High n-3)", "Seafood (Low n-3)",
  "Eggs", "Soybean Products", "Nuts and Seeds", "Legumes (Protein)",
  "Milk", "Yogurt", "Cheese", "Oils", "Solid Fats", "Added Sugar",
  "Alcoholic Drinks")

# Category level labels
categ_title <- "Consumption Level"
categ_labels <- c("None", "Low", "Med", "High")

# Class labels
class_title <- "Dietary Pattern"
class_labels <- 1:5

# Get pattern profiles
plot_pattern_profiles(res = res_swolca_adjust,
```

```

item_labels = item_labels, item_title = item_title,
categ_labels = categ_labels, categ_title = categ_title,
class_labels = class_labels, class_title = class_title)

```

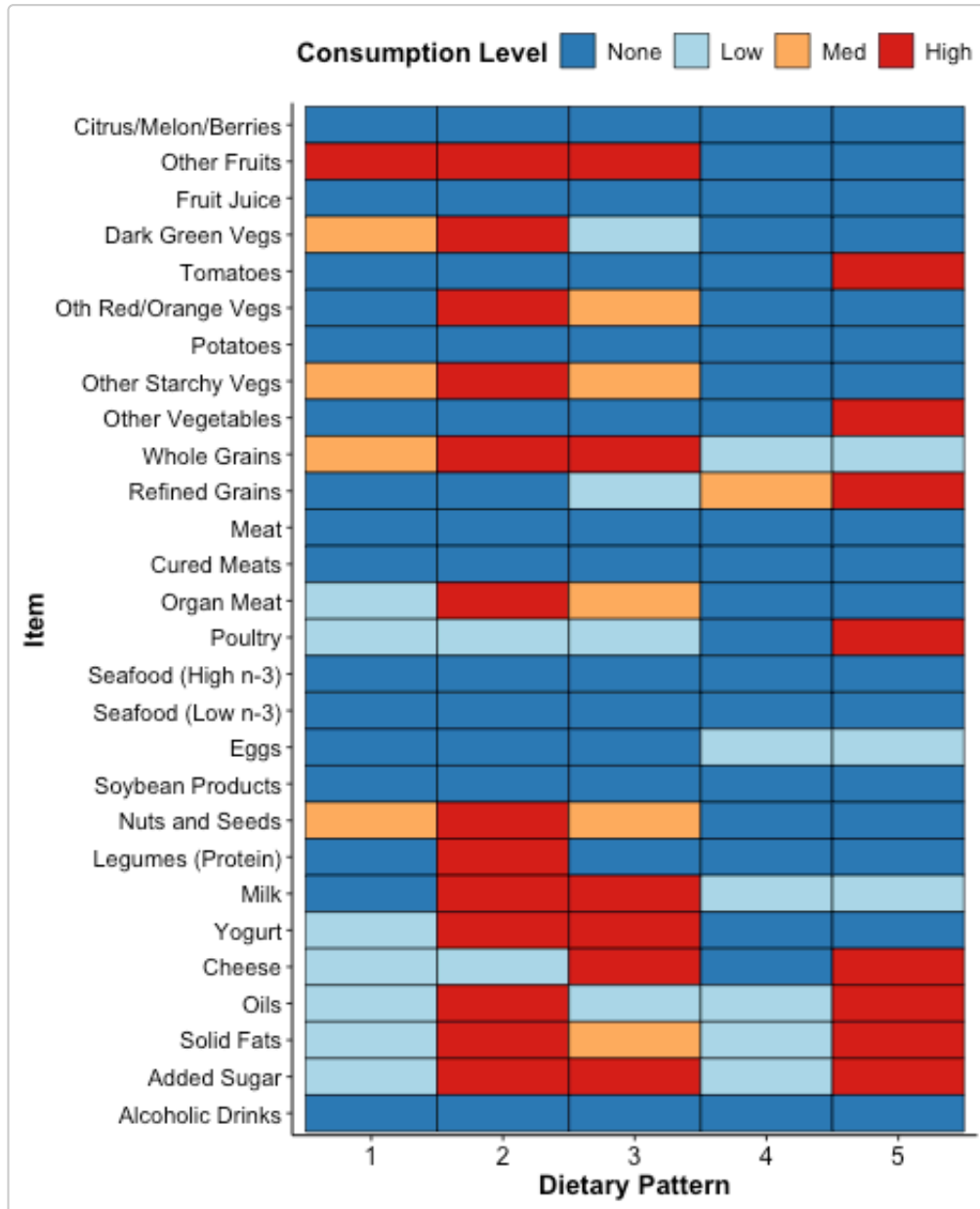


Figure: Latent class pattern profiles with custom labels

For the NHANES example using the `swolca()` supervised clustering function, we can interpret these 5 patterns as diet-hypertension patterns where individuals following the same pattern share similar dietary intake and hypertension status. If the `wolca()` unsupervised clustering function had been used, these patterns would be interpreted as solely diet patterns where individuals following the same pattern share similar dietary intake. Since survey design was accounted for, these patterns should be representative of the patterns in the broader population.

To examine to what extent the data support the pattern profiles, we can view the specific distribution of posterior probabilities across the item categories for each item, for each class. This allows us to examine how much evidence supports the category with the highest posterior probability compared to the other categories. If there is no evidence supporting one category over the others, then all the probabilities will be equal to $1/R_j$, where R_j is the number of categories for item j . For example, the distribution for Item 2

for class 1 is pretty even across the 4 consumption levels, indicating that class 1 still has a fair amount of heterogeneity in consumption of Item 2. If the data strongly support a category, the posterior probability for that category will be much closer to 1 (e.g., 0.8). For example, the probability Item 14 is consumed at level “None” is very high for those in classes 4 and 5.

```
plot_pattern_probs(res = res_swolca_adjust, item_labels = item_labels,
                  categ_labels = categ_labels) +
  theme(strip.text.x = element_text(size = 7)) # font size for item labels
```

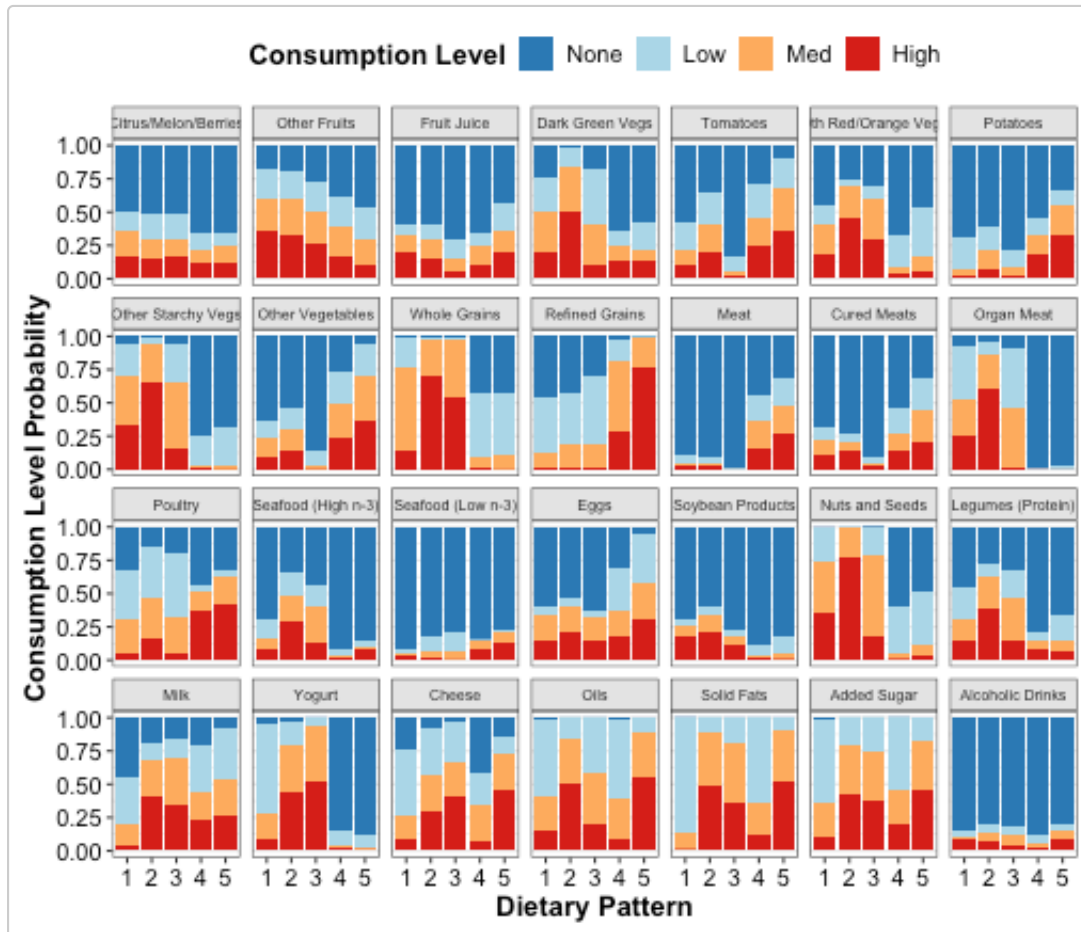


Figure: Item probabilities for each latent class pattern

2.3.2 Obtain individual class assignments

For each individual in the sample, we can obtain their final class assignment as well the posterior probability of them belonging to each class. This information can be used to check class separation metrics such as entropy (Wang et al. 2017) or the average latent class posterior probability (Muthén and Muthén 2000), both recommended to be above 0.8 for adequate class separation (Weller, Bowen, and Faubert 2020). Note that these metrics are NOT generalizable to the broader population and are limited to describing individuals included in the dataset.

```
# Class assignments for first 3 individuals
res_swolca_adjust$estimates_adjust$c_all[1:3]
```

```
## [1] 1 2 3
```

```
# Posterior class probabilities for first 3 individuals, rounded to 3 decimals  
round(res_swolca_adjust$estimates_adjust$pred_class_probs[1:3, ], 3)
```

```
##      [,1] [,2] [,3] [,4] [,5]  
## [1,] 0.999 0.000 0.001    0    0  
## [2,] 0.000 0.958 0.042    0    0  
## [3,] 0.001 0.147 0.851    0    0
```

A dendrogram displaying the distances between classes can also be created. The distance is the number of MCMC iterations with differing class assignments (after discarding burn-in and thinning). The dendrogram displays the hierarchical clustering setup, where individuals first are all assigned to their own class, and then classes are systematically merged together starting from classes that have the smallest distance between them. The preliminary number of latent classes selected is displayed by the rectangles, with each individual assigned to a class. A final merging step between classes that have the exact same modal pattern was done to merge any duplicate classes.

```
# Get dendrogram  
dend <- res_swolca_adjust$post_MCMC_out$dendrogram  
# Get preliminary number of latent classes  
K_med <- res_swolca_adjust$post_MCMC_out$K_med  
# Remove labels for individuals to declutter the plot  
dend$labels <- ""  
# Plot dendrogram  
plot(as.dendrogram(dend), ylab = "Distance")  
# Rectangles around the latent classes  
rect.hclust(dend, k = K_med, border = 2:(2 + K_med - 1))
```

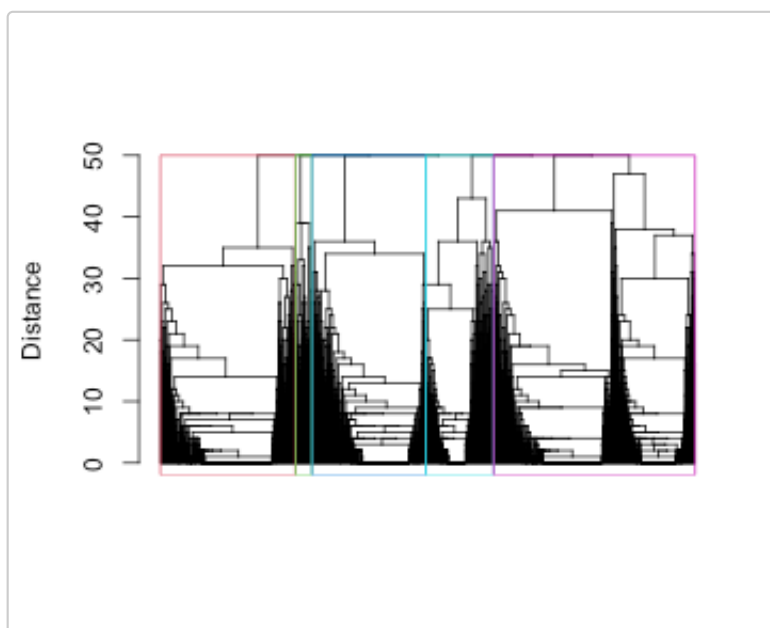


Figure: Dendrogram of class separation

2.3.3 Reorder classes

If desired, we can also change the order of the classes. For example, we may want to swap the first and fifth dietary patterns so that the patterns with more “high” consumption (red) are visually displayed together. One might also want to change the order of classes to define a new baseline class for the regression model. To switch the order so that the first class is now fifth and the fifth class is now first, we can use the `reorder_classes()` function with `new_order = c(5, 2, 3, 4, 1)`. This will change the order so that all subsequent plotting functions will follow the new order, with the previously fifth class now displayed first. Individual class assignments will also follow the new order. That is, individuals who were previously assigned to class 5 will now be assigned to class 1, and vice versa.

```
# Reorder latent classes
res_swolca_adjust <- reorder_classes(res = res_swolca_adjust,
                                     new_order = c(5, 2, 3, 4, 1))

# Class assignments for first 3 individuals
res_swolca_adjust$estimates_adjust$c_all[1:3]

## [1] 5 2 3

# Pattern profiles with first and third classes swapped
plot_pattern_profiles(res = res_swolca_adjust)
```

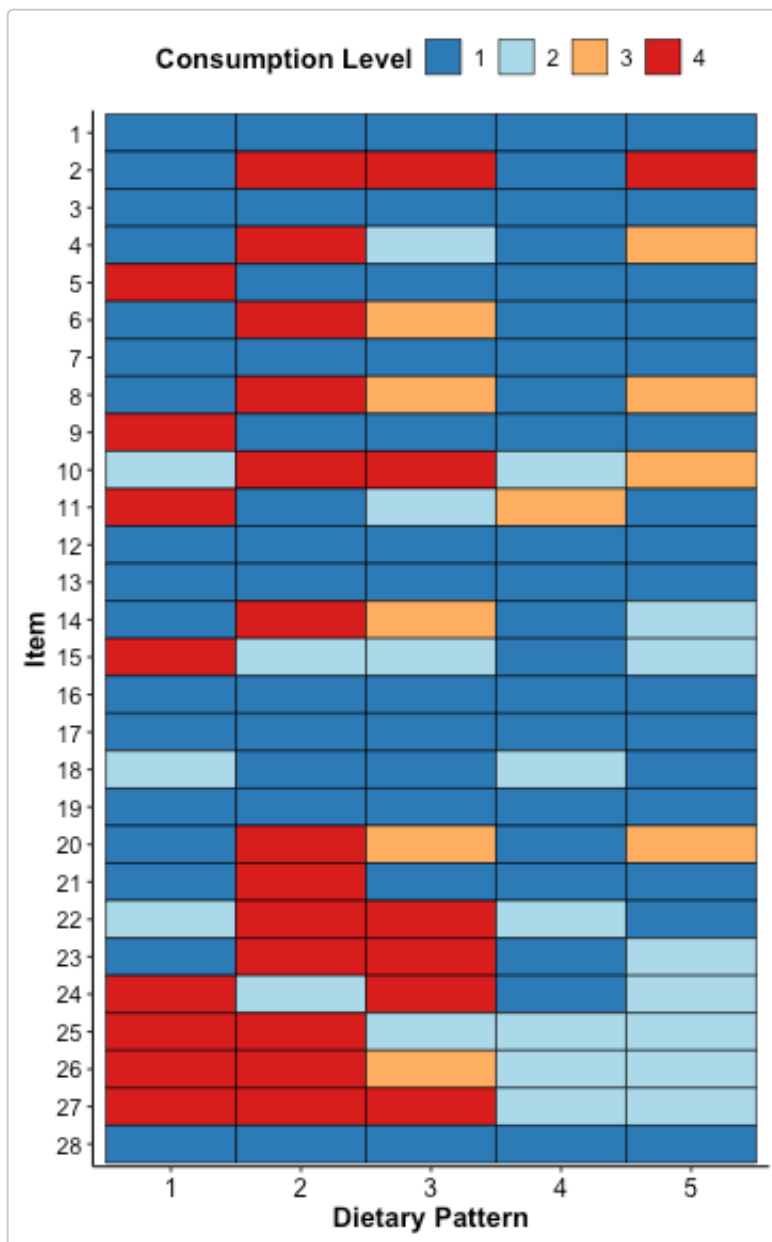


Figure: Latent class pattern profiles with reordered classes

2.3.4 Plot distribution of classes

To examine how common the classes are in the broader population, the `plot_class_dist()` function displays boxplots for the estimated proportion of the population belonging to each class (i.e., latent class membership probabilities, π) across the iterations of the MCMC sampler.

```
plot_class_dist(res = res_swolca_adjust)
```

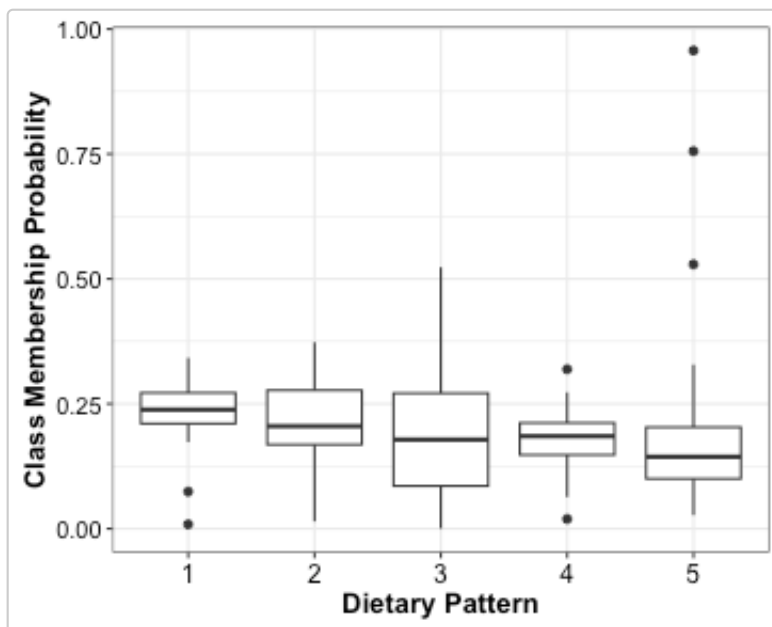


Figure: Distribution of latent classes in the population

```
# Median estimated population class proportions
round(res_swolca_adjust$estimates_adjust$pi_med, 3)
```

```
## [1] 0.251 0.216 0.188 0.195 0.151
```

2.3.5 Display distribution of additional variables across classes

We can create a table to examine the distribution of additional variables across the derived classes. In this example, we create a table of the distribution of age, race and ethnicity, physical activity, and smoking status across five derived dietary patterns. For these distributions, the data are weighted so that the proportions are representative of the distribution of the variables across classes in the population. To see the proportions in the sample, specify `population = FALSE`. The proportion of the classes in the population is also displayed in this table, using the same data as the `plot_class_dist()` function.

```
# Convert class assignment variable from model output to a factor
c_all <- factor(res_swolca_adjust$estimates_adjust$c_all,
               levels = 1:5, labels = paste0("C", 1:5))
# Create dataframe containing covariates to view distribution across classes,
# editing the labels for the variable names and levels
cov_df <- data_nhanes %>%
  select(RIDAGEYR, age_cat, racethnic, smoker, physactive) %>%
  rename(Age = RIDAGEYR) %>%
  mutate(Age_Group = factor(age_cat, levels = c(1, 2, 3),
                           labels = c("[20, 40)", "[40, 60)", ">=60)"),
         Race_Ethnicity = factor(racethnic, c(1, 2, 3, 4, 5),
                                labels = c("NH White", "NH Black", "NH Asian",
                                             "Hispanic/Latino", "Other/Mixed")),
         Current_Smoker = factor(smoker, levels = c(0, 1), labels = c("No", "Yes")),
         Physical_Activity = factor(physactive, levels = c("Inactive", "Active")))
```

```

    .keep = "unused")
# Obtain table displaying distribution of variables across classes
output_df <- vars_across_class(c_all = c_all, cov_df = cov_df,
                              sampling_wt = sampling_wt, stratum_id = stratum_id,
                              cluster_id = cluster_id, digits = 1,
                              col_props = TRUE, res = res_swolca_adjust)

knitr::kable(output_df)

```

Variable	Level	C1	C2	C3	C4	C5	Overall
N: % (posterior SD %)		23.8 (5.9)	21.1 (8.1)	18.9 (13.4)	18.1 (5.1)	18.0 (16.6)	
n: %		25.2	23.1	13.3	21.5	17.0	
Age: Mean		44.0	43.8	47.0	50.1	47.4	46.2
Age_Group: %	[20, 40)	46.3	49.5	37.8	31.1	41.9	41.9
	[40, 60)	30.7	26.7	38.6	40.5	27.4	32.6
	>=60	23.0	23.8	23.5	28.5	30.7	25.5
Race_Ethnicity: %	NH White	51.1	44.1	64.0	46.7	41.9	49.3
	NH Black	17.6	16.4	10.7	15.5	20.8	16.3
	NH Asian	3.1	5.2	2.3	7.5	11.4	5.6
	Hispanic/Latino	23.0	28.7	20.6	21.8	22.7	23.6
	Other/Mixed	5.1	5.6	2.3	8.5	3.2	5.2
Current_Smoker: %	No	74.4	78.4	64.5	71.1	77.3	73.5
	Yes	25.6	21.6	35.5	28.9	22.7	26.5
Physical_Activity: %	Inactive	45.5	44.5	44.4	47.1	42.9	45.1
	Active	54.5	55.5	55.6	52.9	57.1	54.9

Parameters used for `vars_across_class()`:

- `c_all`: nx1 factor vector of latent class assignments, typically obtained from the `swolca()`, `wolca()`, `swolca_var_adjust()` or `wolca_var_adjust()` functions. Factor levels should be labeled with the names that are to appear in the output table.
- `cov_df`: Dataframe with *n* rows containing sociodemographic variables. Factors should be labeled with the names and levels that are to appear in the output table.
- `sampling_wt`: nx1 vector of survey sampling weights.
- `cluster_id`: nx1 vector of cluster indicators for clustered sampling.
- `stratum_id`: nx1 vector of stratum indicators.
- `digits = 1`: Round table output to 1 decimal place (default).
- `col_props = TRUE`: Factor variables should have percentages reported as column totals (default), displaying the population in each category for a given class. If `FALSE`, row totals are used instead and display the population in each class for a given category.
- `res = res_swolca_adjust`: Results from `swolca_var_adjust()` to be used to calculate the

percentage of individuals falling into each class in the population, accounting for survey design.

2.3.6 Summarize and plot the exposure-outcome relationship

To examine how the latent classes relate to the binary outcome of interest, we can first convert the regression coefficients from factor reference coding to the more commonly used reference cell coding and display a summary of the parameter estimates. One advantage of factor reference coding is it allows the reference class to be determined post-hoc, so any class can be chosen as the reference level for which to display regression coefficients without having to re-run the model. Simply run `reorder_classes()` with the desired reference level as the first class in `new_order`, and then run `get_regr_coefs()` as shown below. Note that this is *not* possible with the two-step `wolca_svyglm()` method. Please re-run `wolca_svyglm()` with the desired reference level refactoring

In this example, we display a summary of the probit regression coefficients using class 1 as the reference. We also display the 95% credible intervals (`ci_level = 0.95`) and round the table values to two decimal points (`digits = 2`). To save space, we only display the first 10 regression coefficients in the output.

```
# Get table of regression coefficients
```

```
regr_coefs <- get_regr_coefs(res = res_swolca_adjust, ci_level = 0.95, digits = 2)
regr_coefs[1:10, ]
```

##	Covariate	Estimate	LB	UB	P(xi > 0)
## 1	(Intercept)	-1.13	-1.48	-0.72	<0.01
## 2	c_all2	-0.10	-1.19	1.24	0.42
## 3	c_all3	0.08	-0.75	1.26	0.62
## 4	c_all4	0.28	-0.57	1.09	0.66
## 5	c_all5	0.24	-2.36	1.76	0.62
## 6	age_cat2	1.09	0.82	1.38	1.00
## 7	age_cat3	2.13	1.47	2.58	1.00
## 8	racethnic2	0.44	-0.13	1.08	0.94
## 9	racethnic3	0.06	-0.59	1.12	0.56
## 10	racethnic4	-0.16	-0.48	0.20	0.20

To visualize this in the form of a boxplot, use the `plot_regr_coefs()` function.

```
plot_regr_coefs(regr_coefs = regr_coefs, res = res_swolca_adjust)
```

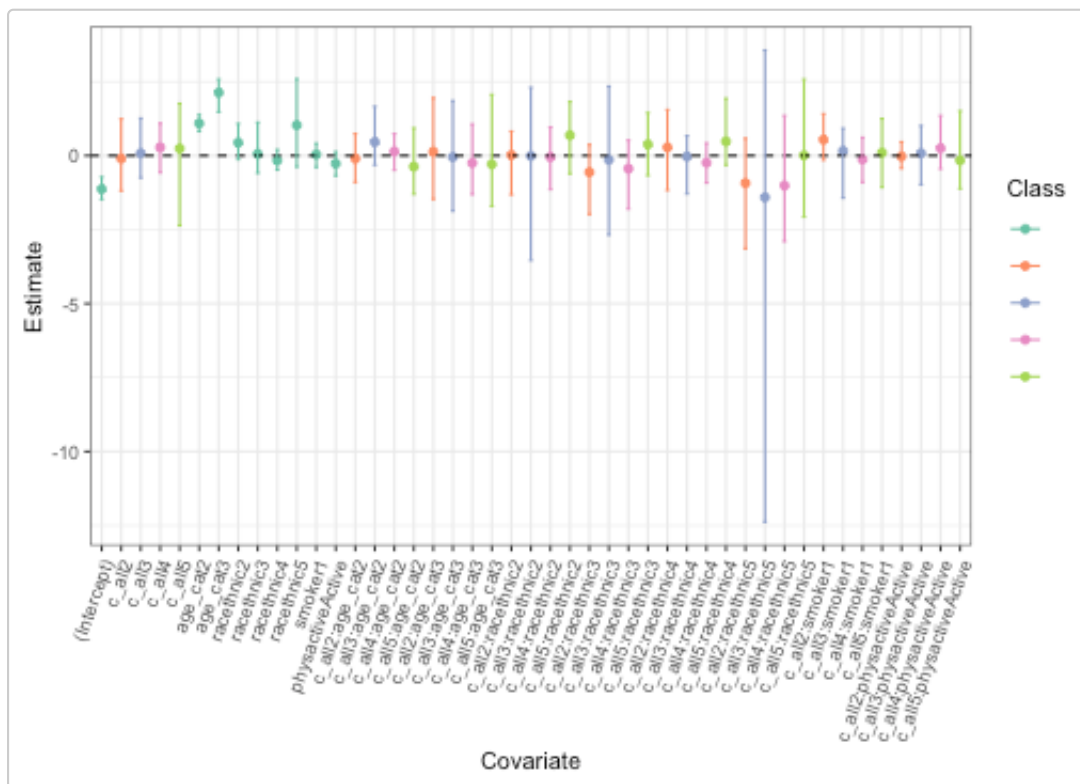


Figure: Outcome regression coefficients

If we want to change the labels for the covariates, we can do so with the `cov_labels` argument. We can also modify other aspects of the plot, such as moving the legend position, by using typical arguments from the `ggplot2` package.

```
# Define new vector of covariate labels
class_dummies <- paste0("C", 2:5)
age_dummies <- paste0("Age", c("40_60", "60"))
race_dummies <- paste0("RaceEth", c("NH_Black", "NH_Asian", "Hisp", "Other"))
cov_labels <- c("Intercept", class_dummies, age_dummies, race_dummies,
               "SmokerYes", "PhysActive",
               paste0(class_dummies, ":", rep(age_dummies, each = 4)),
               paste0(class_dummies, ":", rep(race_dummies, each = 4)),
               paste0(class_dummies, ":", rep("SmokerYes", each = 4)),
               paste0(class_dummies, ":", rep("PhysActive", each = 4)))

# Get plot of regression coefficient estimates and error bars
regr_plot <- plot_regr_coefs(regr_coefs = regr_coefs, res = res_swolca_adjust,
                           cov_labels = cov_labels)

# Display plot with legend above
regr_plot + ggplot2::theme(legend.position = "top")
```

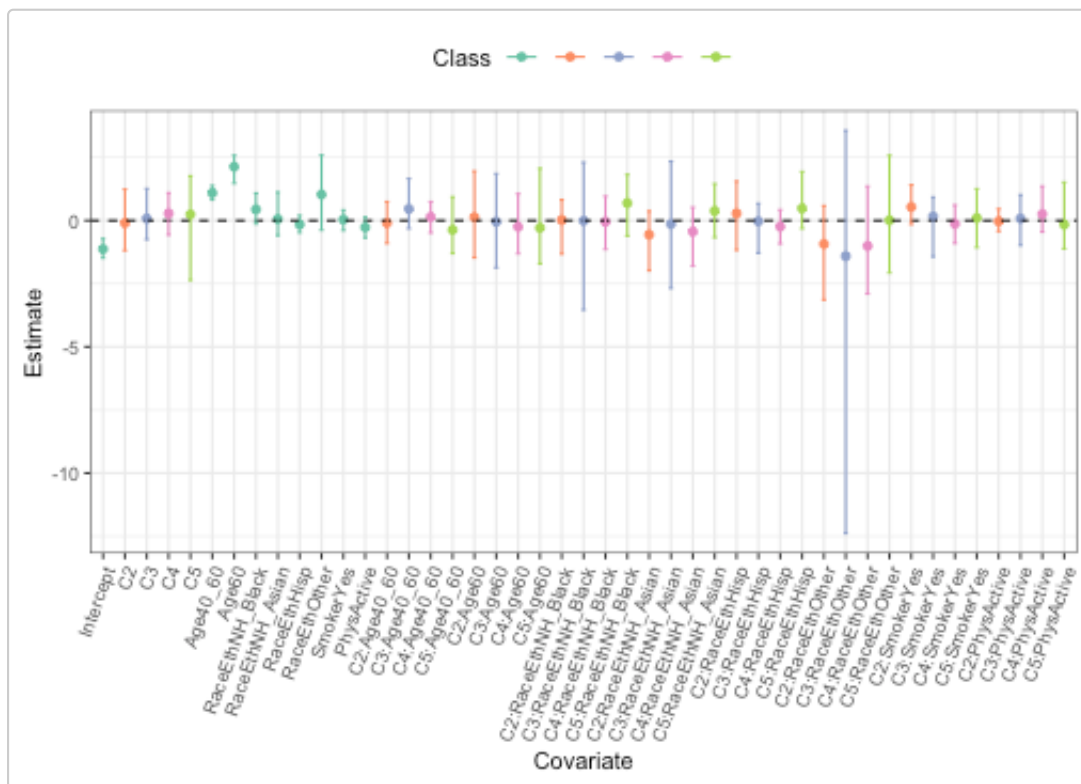


Figure: Outcome regression coefficients with custom labels

Alternatively, for categorical covariates, rather than plotting the regression coefficients, we can evaluate the probit regression model at all levels of the covariate and then plot the conditional probability of the outcome using the `plot_outcome_probs()` function. This displays the probability of the outcome for the different classes, evaluated at all levels of a categorical covariate, with all other covariates evaluated at their reference level. In this example, we plot the probability of hypertension for the five dietary patterns across levels of age, with error bars for the 95% credible intervals.

```
class_labels <- paste0("C", 1:5)
age_cat_categs <- c("[20,40)", "[40,60)", ">=60")
plot_outcome_probs(res = res_swolca_adjust, cov_name = "age_cat",
  cov_labels = age_cat_categs, class_labels = class_labels,
  x_title = "Age Group")
```

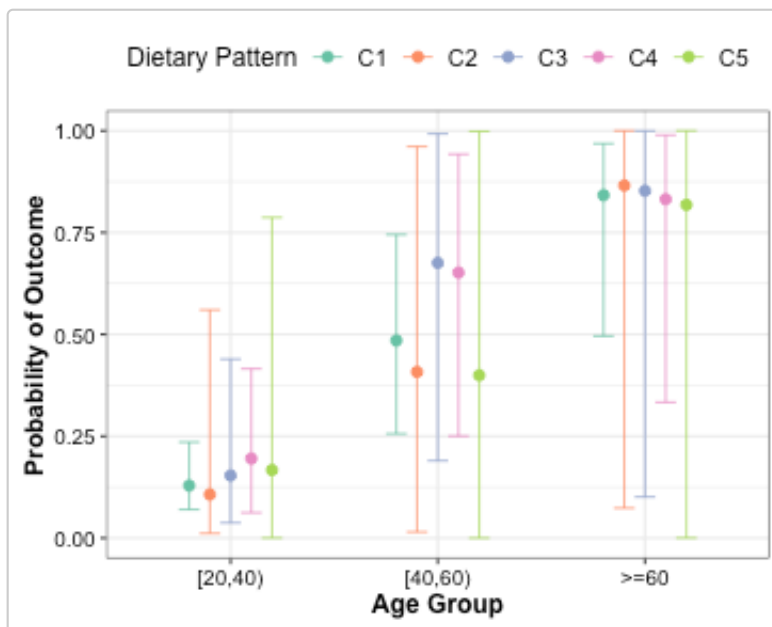



Figure: Outcome probabilities for levels of a categorical covariate

We can also display the probability of outcome plots for two categorical covariates at the same time, assessing for any interactions. We can add lines connecting the points between levels of “Age” using the `add_lines = TRUE` parameter, and remove the error bars by setting `ci_level = NULL`.

```
racethnic_categs <- c("NH White", "NH Black", "NH Asian", "Hispanic/Latino",
  "Other/Mixed")
p <- plot_outcome_probs(res = res_swolca_adjust, cov_name = c("age_cat", "racethnic"),
  cov_labels = list(age_cat_categs, racethnic_categs),
  class_labels = class_labels, x_title = "Age Group",
  ci_level = NULL, add_lines = TRUE)
print(p + ggplot2::theme(axis.text.x = ggplot2::element_text(size = 9)))
```

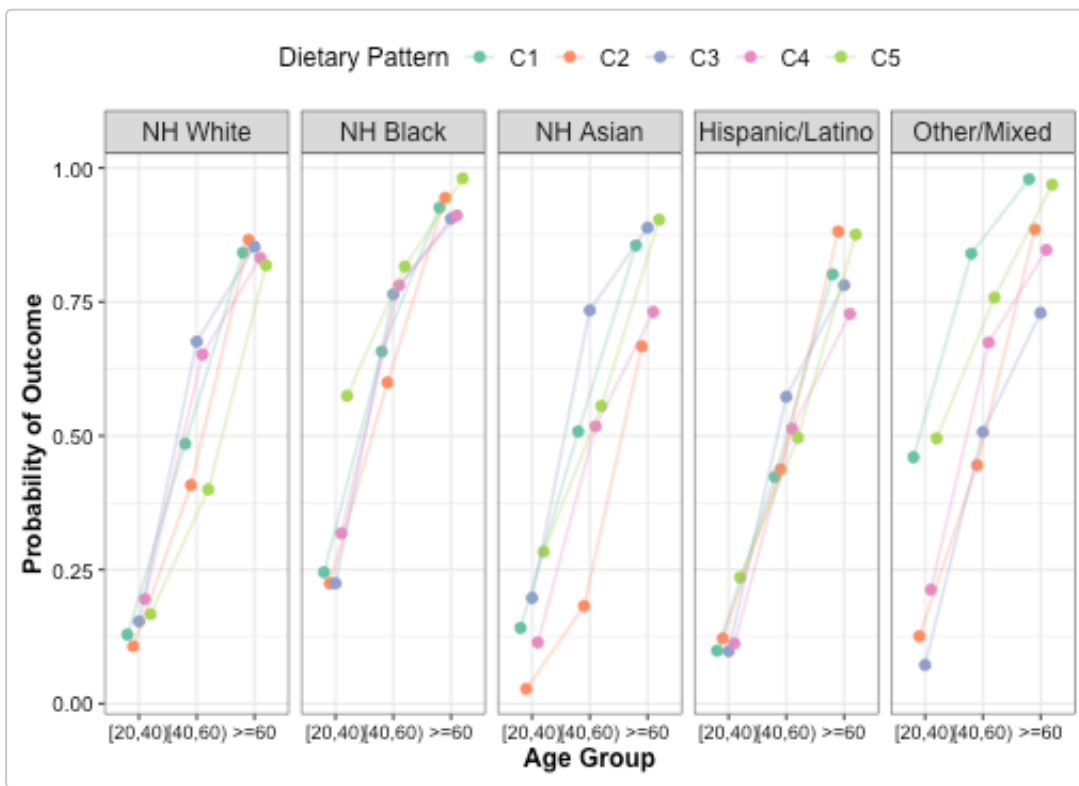


Figure: Outcome probabilities for levels of two categorical covariates

2.4 Obtain performance diagnostics and additional output

Total runtime and the deviance information criteria (DIC) goodness-of-fit measure for model selection are available within the package using the following functions.

```
# Runtime
```

```
res_swolca_adjust$runtime
```

```
## Time difference of 10.97209 mins
```

```
# Deviance information criteria for model goodness-of-fit
```

```
get_dic(res_swolca_adjust)
```

```
## [1] 125463.1
```

A list of the raw posterior point and interval estimates for all parameters can be obtained. We set a credible interval level of 95% and round the estimates to two digits. The parameters consist of:

- π_1, \dots, π_K : latent class membership probabilities for each of the K underlying classes. π parameters are returned in a vector of length K .
- $\theta_{jk1}, \dots, \theta_{jkR_j}$: exposure level probabilities ranging from 1 to R_j , where R_j is the number of exposure levels for item j out of a total of J items, and $k = 1, \dots, K$. θ parameters are returned in a $J \times K \times R$ -dimensional array.
- $\xi_{k1}, \dots, \xi_{kQ}$: coefficients in the binary outcome probit regression with Q covariates. ξ parameters

are returned in a $K \times Q$ -dimensional matrix.

Thus, a total of $K + (J * K * R) + (K * Q) = 610$ estimates are available for this example. To conserve space, we only display the first three estimates for π , θ , and ξ here. π estimates are labeled with “pi_k”, θ estimates are labeled with “theta_j_k_r”, and ξ estimates are labeled with “xi_k_q”. k ranges from 1 to K , j from 1 to J , r from 1 to R , and q from 1 to Q .

```
# Extract summary of parameter estimates
```

```
estimates_df <- summarize_res(res = res_swolca_adjust, ci_level = 0.95, digits = 2)
```

```
# First few pi estimates
```

```
estimates_df[1:3, ]
```

##	Parameter	Estimate	Lower Bound	Upper Bound
## 1	pi_1	0.25	0.10	0.33
## 2	pi_2	0.22	0.03	0.34
## 3	pi_3	0.19	0.02	0.47

```
# First few theta estimates
```

```
estimates_df[6:9, ]
```

##	Parameter	Estimate	Lower Bound	Upper Bound
## 6	theta_1_1_1	0.65	0.43	0.76
## 7	theta_2_1_1	0.46	0.29	0.59
## 8	theta_3_1_1	0.44	0.38	0.52
## 9	theta_4_1_1	0.58	0.41	0.70

```
# First few xi estimates
```

```
estimates_df[566:569, ]
```

##	Parameter	Estimate	Lower Bound	Upper Bound
## 566	xi_1_1	-1.13	-1.48	-0.72
## 567	xi_2_1	-1.24	-2.27	0.15
## 568	xi_3_1	-1.02	-1.78	-0.15
## 569	xi_4_1	-0.86	-1.54	-0.21

We can also extract the MCMC iteration values for all parameters that were obtained through the MCMC sampler. For SWOLCA, this will be π , θ , and ξ . For WOLCA, this will be only π and θ .

```
# Extract parameter MCMC iteration values
```

```
param_mcmc <- get_param_mcmc(res_swolca_adjust)
```

```
# Display first 10 iterations of pi
```

```
param_mcmc$pi_mcmc[1:10, ]
```

##	pi_1	pi_2	pi_3	pi_4	pi_5
## 1	0.199337680	0.29798426	0.1279449684	0.2206540	0.15407908

```
## 2  0.256967542 0.30446362 0.2021383390 0.1572870 0.07914354
## 3  0.263425322 0.08413522 0.4648721743 0.1231027 0.06446456
## 4  0.221394758 0.23644027 0.1722512104 0.2167745 0.15313927
## 5  0.269240234 0.18652288 0.2918798644 0.1791225 0.07323449
## 6  0.311260177 0.19111240 0.2805098911 0.1457299 0.07138768
## 7  0.008695988 0.01466584 0.0003048905 0.0195115 0.95682179
## 8  0.196751600 0.24555358 0.0486972000 0.2042066 0.30479104
## 9  0.277954303 0.22683991 0.1138511948 0.2062676 0.17508698
## 10 0.245779276 0.27931692 0.0645889321 0.1847761 0.22553878
```

Bayesian model diagnostics such as traceplots and autocorrelation plots can be created for all estimates to examine whether the MCMC sampler has adequately converged. For example, we display code to create traceplots and autocorrelation plots for the five class membership probability parameters, π_1, \dots, π_5 .

```
# Specify selection of pi_1 to pi_5
param_names <- colnames(param_mcmc$pi_mcmc)
# Create traceplots
create_traceplot(param_mcmc = param_mcmc, param_names = param_names)
```

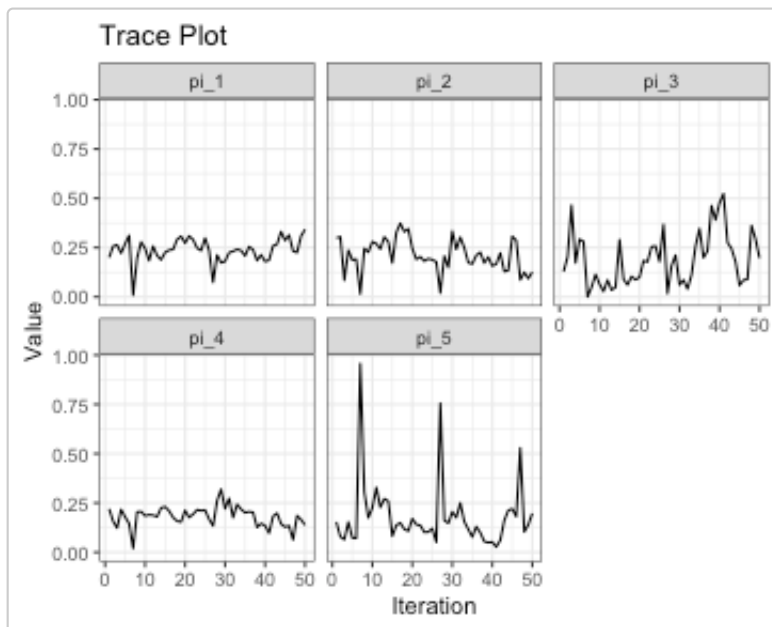


Figure: Traceplots for pi

```
# Create ACF plots
create_acfplot(param_mcmc = param_mcmc, param_names = param_names)
```

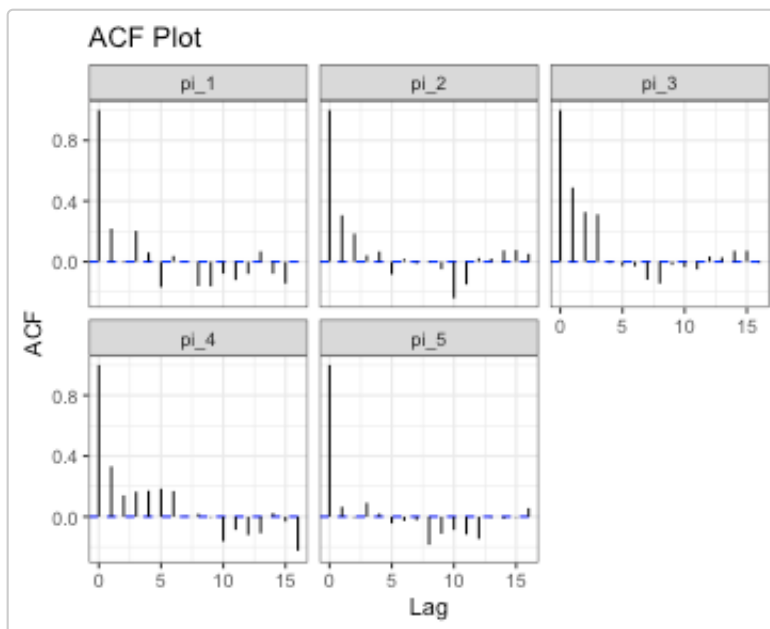


Figure: ACF plots for pi

3 References

- Lumley, Thomas. 2023. "Survey: Analysis of Complex Survey Samples."
- Muthén, Bengt, and Linda K Muthén. 2000. "Integrating Person-Centered and Variable-Centered Analyses: Growth Mixture Modeling with Latent Trajectory Classes." *Alcoholism: Clinical and Experimental Research* 24 (6): 882–91.
- National Center for Health Statistics. 2023. "National Health and Nutrition Examination Survey Home Page." Centers for Disease Control; Prevention, National Center for Health Statistics, Atlanta, GA. <https://www.cdc.gov/nchs/nhanes.htm>.
- Nylund-Gibson, Karen, Ryan P Grimm, and Katherine E Masyn. 2019. "Prediction from Latent Classes: A Demonstration of Different Approaches to Include Distal Outcomes in Mixture Models." *Structural Equation Modeling: A Multidisciplinary Journal* 26 (6): 967–85.
- Savitsky, Terrance D, and Daniell Toth. 2016. "Bayesian Estimation Under Informative Sampling." *Electronic Journal of Statistics* 10 (1): 1677–1708.
- Stephenson, Briana JK, Stephanie M Wu, and Francesca Dominici. 2023. "Identifying dietary consumption patterns from survey data: a Bayesian nonparametric latent class model." *Journal of the Royal Statistical Society Series A: Statistics in Society*, December, qnad135. <https://doi.org/10.1093/jrssa/qnad135>.
- Wang, Meng-Cheng, Qiaowen Deng, Xiangyang Bi, Haosheng Ye, and Wendeng Yang. 2017. "Performance of the Entropy as an Index of Classification Accuracy in Latent Profile Analysis: A Monte Carlo Simulation Study." *Acta Psychologica Sinica*.
- Weller, Bridget E, Natasha K Bowen, and Sarah J Faubert. 2020. "Latent Class Analysis: A Guide to Best Practice." *Journal of Black Psychology* 46 (4): 287–311.
- Williams, Matthew R, and Terrance D Savitsky. 2021. "Uncertainty Estimation for Pseudo-Bayesian Inference Under Complex Sampling." *International Statistical Review* 89 (1): 72–107.
- Wu, Stephanie M, Matthew R Williams, Terrance D Savitsky, and Briana JK Stephenson. 2024. "Derivation of Outcome-Dependent Dietary Patterns for Low-Income Women Obtained from Survey Data Using a Supervised Weighted Overfitted Latent Class Analysis." *Biometrics* 80 (4): ujae122.

4 Supplementary installation

instructions

For users experiencing a “library ‘gfortran’ not found” error message during installation, additional steps are needed to install the `bayesc` package. This is a compiler configuration issue that can arise when using Rcpp-dependent packages on Mac computers with Apple silicon (e.g., M1, M2, M3). Please follow the instructions listed below, adapted from instructions posted at

<https://stackoverflow.com/questions/70638118/configuring-compilers-on-apple-silicon-m1-m2-m3-for-rcpp-and-other-tool>).

1. Download an R binary from CRAN at this link: <https://cran.r-project.org/bin/macosx/>. Select the binary built for Apple silicon (M1-M3), which will typically be the top link on the left under “Latest release”, and install.
2. Go through the instructions to install R, entering in passwords if necessary.
3. Install Xcode by opening Terminal and running the following in the command line: `sudo xcode-select --install`. This will install the latest release version of Apple’s Command Line Tools for Xcode, which includes Apple Clang.
4. Download the GNU Fortran binary .tar.xz file at this link: <https://github.com/R-macos/gcc-12-branch/releases/tag/12.2-darwin-r0>. Install GNU fortran by running the following code in the command line:

```
curl -LO https://github.com/R-macos/gcc-12-branch/releases/download/12.2-darwin-r0/gfortran-12.2-darwin20-r0-universal.tar.xz
sudo tar xvf gfortran-12.2-darwin20-r0-universal.tar.xz -C /
sudo ln -sf $(xcrun --show-sdk-path) /opt/gfortran/SDK
```

5. Check your Apple Clang version by running the following in the command line: `clang --version`.
6. Download OpenMP at <https://mac.r-project.org/openmp/> by clicking on the Release.tar.gz link corresponding to your Apple Clang version.
7. Install OpenMP by running the following in the command line, making sure to include the correct version. For example, for Apple clang version 15.0.0 (i.e., clang 1500.x), the corresponding OpenMP version is 16.0.4, so the commands would be:

```
curl -LO https://mac.r-project.org/openmp/openmp-16.0.4-darwin20-Release.tar.gz
sudo mkdir -p /opt/R/$(uname -m)
sudo tar -xvf openmp-16.0.4-darwin20-Release.tar.gz --strip-components=2 -C /opt/R/$(uname -m)
```

8. Navigate to the R Makevars file by running in command line: `cd ~/.R/Makevars`. If the file does not exist, create it by running `mkdir ~/.R/Makevars`.
9. In command line, run `vi`, add the below lines to the file, then save and exit out of the file by typing ESC followed by `:wq`.

```
CPPFLAGS += -Xclang -fopenmp
LDFLAGS += -lomp
```

10. Retry the `bayesc` package installation by running `devtools::install_github("smwu/bayesc")` in R.