

Effects of Great Recession on Income Poverty

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1 Study Aim

The aim of this study is to analyze how households with a working age adult with disability compare with households with no adult with disability, during the *great recession*, using “Income Poverty” as a measure of economic wellbeing, when controlling for demographic factors such as gender, marital status, education, race and origin.

2 Sample

For this analysis data from US Census Bureau’s SIPP 2008 panel survey was used.¹ Questions on whether the households had a working age adult with disability were asked in wave-6 of the survey, which ended in August, 2010. Households that participated in wave-6 were included in our sample. There were a total of 34,850 households in wave six. Survey data upto wave-15 were used in our sample. Survey results from July, 2008 through June, 2013 were included in the analysis. Households whose reference person remained the same throughout the 2008 panel were kept in the sample. The reference persons of the households were also required to be 18 years or older throughout the 2008 panel. The final sample had 33,547 households that satisfied all the inclusion criteria.

3 Methods

Total monthly household income was divided by monthly federal poverty level (FPL) and then averaged over quarters to estimate FPL100-ratio. An FPL100-ratio lower than one in any quarter indicated the household was below 100% Federal poverty level in that quarter. Averaging monthly

¹For more information on the SIPP 2008 panel schedule, please refer to this US Census Bureau website

values over a quarter reduced the noise in the response variable by eliminating the month-by-month variability in the income data. In the sample, the monthly income data ranged from -\$27,180 to \$108,900, the average being \$5240 and median \$3,874. The negative incomes were associated with households owning business that incurred losses in those months. The FPL100-ratio ranged from -17.95 to 89.48, with the average being 3.817 and the median 2.924. In the sample, 7,865 out of 33,547 households (23.44%) had at least one working age adult with disability during the observed time period, as identified in wave-6 of the survey. Data from July 2008 (2008-Q3) through May 2013 (2013-Q1) were analyzed. This period overlapped with twelve of the eighteen months ² of the “Great recession” and its long wake.

3.1 Mixed-effects model for income poverty

A mixed (fixed and random) effects model was fit to estimate the impact of the presence of a working-age adult with disability in a household on its FPL100-ratio. Let Y denote the vector of responses (FPL100-ratio). Let Θ denote the vector of fixed effect factors like gender, marital status, education level, race, origin of household head, along with their interactions. Let β denote another fixed effect of time, represented as quarters, starting from 2008-Q3 and ending in 2013-Q1. Let b denote the household level random effect (random intercept). Then, the mixed-effects ([3]) model for the responses, for each household i can be written as

$$Y_{ij} = \beta_0 + \beta t_j + X_i \Theta + b_i + \epsilon_{ij} \quad (1)$$

where, ϵ_{ij} are regarded as measurement errors, i goes from 1 to H , the number of households, j goes from 1 to T , the total number of quarterly observations for every household. In this model, the response from the i^{th} household at time t_j is assumed to differ from the population mean $X_i \Theta + \beta t_j$ by a household effect b_i and a within household measurement error ϵ_{ij} . The within household and between household errors are assumed to be normal and independent ($b_i \sim \mathcal{N}(0, \sigma_b^2)$, $\epsilon_{ij} \sim \mathcal{N}(0, \sigma^2)$, $b_i \perp \epsilon_{ij}, \forall i, j$). The effect of “time” is a fixed effect and it could be considered part of the fixed effect design matrix X . However, the problem of interest is to test a linear hypothesis about the disability-by-time interaction, to detect the effect of disability on the mean response over the period of the study. Hence, the “time” covariate is denoted separately. During estimation it will be estimated as a fixed effect.

²NBER Recession Cycles

The bottom, or trough, of the great recession was reached in the second quarter of 2009 (marking the technical end of the recession, defined as at least two consecutive quarters of declining GDP)³. According to NBER, June 2009 was the final month of the recession. We wanted to check if this was reflected in the FPL100-ratio as a downward trend in the initial quarters followed by an upward trend. A simple linear term in “time” would be insufficient. We added a second order term time^2 to test the change in direction of trend. The second order term was added after centering the original “time” variable, to avoid introducing multicollinearity.

Below is the model that will be fit:

$$Y_{ij} = \beta_0 + \beta_1 t_j + \beta_2 t_j^2 + \delta \mathbb{1}_{D_i} + \gamma(\mathbb{1}_{D_i} * t_j) + X_i \Theta + b_i + \epsilon_{ij} \quad (2)$$

where, $\mathbb{1}_{D_i} = 1$, if household i has a working age adult with disability, else $\mathbb{1}_{D_i} = 0$. The hypotheses of interests are:

$H_0 : \beta_1 = 0$, vs $H_a : \beta_1 \neq 0$ tested if FPL100-ratio changed over time

$H_0 : \beta_2 = 0$, vs $H_a : \beta_2 \neq 0$ tested if the trend of FPL100-ratio changed direction

$H_0 : \delta = 0$, vs $H_a : \delta \neq 0$ tested if disability had any effect on FPL100-ratio

$H_0 : \gamma = 0$, vs $H_a : \gamma \neq 0$ tested if disability had any effect on the slope of FPL100-ratio during the study period

$H_0 : \Theta = 0$, vs $H_a : \Theta \neq 0$ tested if demographic factors had any effect on FPL100-ratio.

In addition, interactions between demographic factors, and between disability and demographic factors were also tested.

Demographic factors

The demographic factors were considered as fixed effects. The factors included in this analysis were *gender*, *marital status*, *education*, and *race/origin* of household head. The *Race/Origin* factor included “non-hispanic white”, “non-hispanic black”, “hispanic” and “others”. For simplicity, “white” and “black” would indicate categories “non-hispanic white” and “non-hispanic black”. Gender of household head had two categories: “male” and “female”. Education of household head has three categories: “high-school or less”, “some college, diploma, associated degrees” and “bachelors or higher”. Marital status of household heads has two categories: “married” and “not married”. Divorced or widowed household heads were considered in the “not married” category.

³Business Cycle Dating Committee, National Bureau of Economic Research (NBER)

Computational software

All analysis were conducted using the statistical software R ([7]), version 3.3.1. The mixed effects models were fit using the R-package “lme4” ([1]) and all hypothesis tests were done using the R package “lmerTest” ([5]). The final model was fit with some of the fixed effect factors along with their interactions after performing “backward elimination” on the full model. Elimination of the fixed effects were done by the principle of marginality, that is: the highest order interactions are tested first: if they are significant, the lower order effects were included in the model without testing for significance. The p-values for the fixed effects are estimated from the F statistics, with “Satterthwaite” approximation ([8]) denominator degrees of freedom. The p-values for the random effect were computed from likelihood ratio tests ([6]).

3.2 Post-hoc tests

Post-hoc tests were conducted between categories of all demographic factors and their interactions, by calculating differences of “Least Squares Means” using R package “lmerTest” ([5]), with “Satterthwaite” approximation ([8]) of the denominator degrees of freedom.

Multiple testing correction

When conducting post-hoc tests for demographic factors and their interactions, due to multiple categories of these factors the size of the tests might be inflated. Sequentially rejective *Bonferroni procedure* ([4]) and *Benjamini-Hochberg procedure* ([2]) remain the two most popular multiple testing correction procedures. Holm’s sequentially rejective Bonferroni procedure controls the family-wise type-I error rate (FWER) and is more powerful than the classical Bonferroni procedure. Benjamini-Hochberg controls the false discovery rate (FDR) which is the expected value of false discovery proportion. Controlling FWER usually proves to be too conservative. Hence, we used the Benjamini-Hochberg procedure, which is less conservative, but more powerful than Bonferroni correction. All post-hoc test p-values reported are Benjamini-Hochberg corrected.

4 Results

Table 1 shows the coefficients of time (measured in year-quarters), and disability and the interaction between them. The model includes the demographic factors, as explained in equation 2. ANOVA table

of these demographic factors is in table 2. The $\beta_{\text{time}} = -0.0515, p < 0.01$ in the model, indicating, FPL100-ratio decreased by 0.05 every year, during the study period. The coefficient of *Disability* is $\beta_D = -0.506, p < 0.01$ indicates that households with a working age adult with disability had their FPL100-ratios 0.506 lower, on an average, compared to households without any working age adult with disability. Next, we observe that the coefficient of interaction between time and disability ($\beta_{t*D} = 0.0150, p < 0.01$) is positive. This implies that the slope of FPL100-ratio for households with disability is $-0.0443(-0.0553 + 0.0150)$, which is less negative than the households without disability. This apparently contradictory finding leads us to conclude that households with disability although had “significantly” worse FPL100-ratio throughout the study period, the households without disability experienced more severe declines in their FPL100-ratios. This could throw some light on the impact of different supplementary coverage programs on households with disability.

For all the demographic factors and for interactions between them, post-hoc tests were conducted to identify the pairwise differences in mean FPL100-ratios of different subgroups.

Gender of household head

Marital status of household head

Race/Origin of household head

In table 4 we can see that regardless of disability, in Model1, households with a “black” ($\mu_{\text{diff}} = -0.53, p = 0.0000$) or “hispanic” ($\mu_{\text{diff}} = -0.46, p = 0.0000$) race/origin as household head is worse off compared to those with “others” race/origin. The discrepancy is higher between “black” and “white” household heads ($\mu_{\text{diff}} = -0.90, p = 0.0000$). When corrected for baseline, the relationships reverse in sign, indicating that during the study period, “black” and “hispanic” households fared marginally better than “white” and “others”. When considering only the households with disability (table ??), there is no significant difference between the different races.

Education of household head

In table 5, we see that regardless of disability, in Model1, households in “high school or less” is worse off than “some college, diploma” ($\mu_{\text{diff}} = -0.47, p = 0.0000$), which is in turn worse off than “bachelors or higher” ($\mu_{\text{diff}} = -1.29, p = 0.0000$). In Model 2, only the difference between “some college, diploma”

and “bachelors or higher” groups are significant ($\mu_{\text{diff}} = -0.08, p = 0.014$). When considering only the households with disability (table ??), there is no significant difference between the different education groups.

Interaction of Gender and Marital status

In table 6, we see that in Model 1 “female, not married” households are worse off than “male, married” ($\mu_{\text{diff}} = -0.92, p = 0.0000$), than “female, married” ($\mu_{\text{diff}} = -0.75, p = 0.0000$) and “male, not married” ($\mu_{\text{diff}} = -0.65, p = 0.0000$). In Model 2, similar relationships hold, with smaller mean difference. Even when considering only the households with disability (table ??), the same relationships hold between the groups. We conclude that “female, not married” households fared the most poorly through our study period.

Interaction of Gender and Education

Focusing on the households In table ??, we see that in Model 1,

Interaction of Race/Origin and Marital status

In table

Interaction of Marital status and Education

In table

Interaction of Race/Origin and Education

In table

5 Limitations

1. Although a linear mixed effects regression model discovered some conventional and some interesting patterns in the relationships between response and demographic factors, along with disability, the trajectory of income poverty over the study period for some households were not linear. This modeling approach does not capture trajectory shapes of individual households. A non-parametric fitting of the income poverty trajectories could be tried as a pre-processing step before testing for differences in behavior between different groups of households.

6 Tables

| Continuous Indep Variables | Model coefficients |
|----------------------------|------------------------|
| Intercept | 2.779*** (0.0706) |
| time (Year-Quarter) | -0.0553*** (0.0021) |
| time ² | 0.0073*** (0.0013) |
| Disability | -0.5061*** (0.0564) |
| Year-Quarter*Disability | 0.0150*** (0.0043) |

Note: *p<0.1; **p<0.05; ***p<0.01

Table 1: Coefficients of Fixed effects in regression models

| | Sum Sq | Mean Sq | NumDF | F.value | p.value |
|-----------------------|----------|---------|-------|---------|---------|
| Time | 5380.93 | 5380.93 | 1 | 446.58 | 0.0000 |
| adult_disb | 2280.29 | 2280.29 | 1 | 189.25 | 0.0000 |
| gender | 3143.82 | 3143.82 | 1 | 260.92 | 0.0000 |
| ms | 5788.34 | 5788.34 | 1 | 480.40 | 0.0000 |
| race_origin | 8743.90 | 2914.63 | 3 | 241.90 | 0.0000 |
| education | 15873.18 | 7936.59 | 2 | 658.69 | 0.0000 |
| gender:ms | 3184.05 | 3184.05 | 1 | 264.26 | 0.0000 |
| ms:race_origin | 1670.73 | 556.91 | 3 | 46.22 | 0.0000 |
| race_origin:education | 652.73 | 108.79 | 6 | 9.03 | 0.0000 |
| I(Time^2) | 414.33 | 414.33 | 1 | 34.39 | 0.0000 |
| ms:education | 450.29 | 225.15 | 2 | 18.69 | 0.0000 |
| gender:education | 310.12 | 155.06 | 2 | 12.87 | 0.0000 |
| adult_disb:gender | 182.78 | 182.78 | 1 | 15.17 | 0.0001 |
| Time:adult_disb | 143.66 | 143.66 | 1 | 11.92 | 0.0006 |
| adult_disb:education | 116.32 | 58.16 | 2 | 4.83 | 0.0080 |

Table 2: Model 1: FPL100 vs demographic factors, time and disability

| | Sum Sq | Mean Sq | NumDF | F.value | p.value |
|-----------------------|--------|---------|-------|---------|---------|
| ms | 66.83 | 66.83 | 1 | 316.81 | 0.0000 |
| education | 65.21 | 32.60 | 2 | 154.56 | 0.0000 |
| ms:race_origin | 18.18 | 6.06 | 3 | 28.73 | 0.0000 |
| race_origin | 14.34 | 4.78 | 3 | 22.66 | 0.0000 |
| gender | 9.09 | 9.09 | 1 | 43.08 | 0.0000 |
| race_origin:education | 11.36 | 1.89 | 6 | 8.98 | 0.0000 |
| ms:education | 8.03 | 4.01 | 2 | 19.03 | 0.0000 |
| gender:ms | 4.40 | 4.40 | 1 | 20.86 | 0.0000 |
| Time | 3.39 | 3.39 | 1 | 16.08 | 0.0001 |
| I(Time^2) | 0.06 | 0.06 | 1 | 0.26 | 0.6074 |

Table 3: Model 2: FPL100 vs demographic factors and time for households with a working age adult with disability

| Factor Levels | Estimate | Standard Error | t-value | p-value |
|-------------------|----------|----------------|---------|---------|
| Black - Others | -0.54 | 0.07 | -8.02 | 0.0000 |
| Black - White | -0.94 | 0.05 | -20.16 | 0.0000 |
| Hispanic - Others | -0.54 | 0.07 | -7.78 | 0.0000 |
| Hispanic - White | -0.94 | 0.05 | -19.60 | 0.0000 |
| Others - White | -0.40 | 0.05 | -7.43 | 0.0000 |
| Black - Hispanic | -0.00 | 0.06 | -0.05 | 0.9610 |

Table 4: Post-hoc test of ethnicity

| Factor Levels | Estimate | Standard Error | t-value | p-value |
|--|----------|----------------|---------|---------|
| High School or less - Some college, diploma, assoc | -0.47 | 0.04 | -12.47 | 0.0000 |
| High School or less - Bachelors or higher | -1.69 | 0.05 | -36.04 | 0.0000 |
| Some college, diploma, assoc - Bachelors or higher | -1.21 | 0.04 | -27.03 | 0.0000 |

Table 5: Post-hoc test of education

| Factor Levels | Estimate | Standard Error | t-value | p-value |
|---------------------------------------|----------|----------------|---------|---------|
| Male Married - Female Married | 0.15 | 0.03 | 5.03 | 0.0000 |
| Male Married - Male Not married | 0.26 | 0.03 | 9.00 | 0.0000 |
| Male Married - Female Not married | 0.94 | 0.03 | 27.73 | 0.0000 |
| Female Married - Female Not married | 0.78 | 0.03 | 27.52 | 0.0000 |
| Male Not married - Female Not married | 0.67 | 0.03 | 22.30 | 0.0000 |
| Female Married - Male Not married | 0.11 | 0.04 | 3.03 | 0.0024 |

Table 6: Post-hoc test of gender and marital status

| Factor Levels | Estimate | Standard Error | t-value |
|---|----------|----------------|---------|
| Male High School or less - Female High School or less | 0.28 | 0.03 | 7.98 |
| Male High School or less - Male Some college, diploma, assoc | -0.55 | 0.05 | -12.14 |
| Male High School or less - Male Bachelors or higher | -1.81 | 0.05 | -33.25 |
| Male High School or less - Female Bachelors or higher | -1.28 | 0.05 | -23.46 |
| Female High School or less - Male Some college, diploma, assoc | -0.82 | 0.05 | -17.49 |
| Female High School or less - Female Some college, diploma, assoc | -0.39 | 0.04 | -9.01 |
| Female High School or less - Male Bachelors or higher | -2.09 | 0.06 | -37.35 |
| Female High School or less - Female Bachelors or higher | -1.56 | 0.05 | -29.82 |
| Male Some college, diploma, assoc - Female Some college, diploma, assoc | 0.43 | 0.04 | 11.59 |
| Male Some college, diploma, assoc - Male Bachelors or higher | -1.26 | 0.05 | -23.80 |
| Male Some college, diploma, assoc - Female Bachelors or higher | -0.73 | 0.05 | -13.44 |
| Female Some college, diploma, assoc - Male Bachelors or higher | -1.69 | 0.05 | -30.91 |
| Female Some college, diploma, assoc - Female Bachelors or higher | -1.16 | 0.05 | -23.24 |
| Male Bachelors or higher - Female Bachelors or higher | 0.53 | 0.04 | 12.35 |
| Male High School or less - Female Some college, diploma, assoc | -0.12 | 0.05 | -2.56 |

Table 7: Post-hoc test of gender and education

| Factor Levels | Estimate | Standard Error | t-value | p-value |
|---|----------|----------------|---------|---------|
| Married Black - Not married Black | 0.54 | 0.05 | 11.34 | 0.0000 |
| Married Black - Not married Hispanic | 0.38 | 0.07 | 5.12 | 0.0000 |
| Married Black - Married Others | -0.55 | 0.08 | -6.83 | 0.0000 |
| Married Black - Married White | -1.05 | 0.06 | -18.44 | 0.0000 |
| Married Black - Not married White | -0.30 | 0.06 | -5.18 | 0.0000 |
| Not married Black - Married Hispanic | -0.39 | 0.07 | -5.70 | 0.0000 |
| Not married Black - Married Others | -1.10 | 0.08 | -14.39 | 0.0000 |
| Not married Black - Not married Others | -0.53 | 0.08 | -6.88 | 0.0000 |
| Not married Black - Married White | -1.59 | 0.05 | -32.10 | 0.0000 |
| Not married Black - Not married White | -0.84 | 0.05 | -16.91 | 0.0000 |
| Married Hispanic - Not married Hispanic | 0.22 | 0.04 | 5.25 | 0.0000 |
| Married Hispanic - Married Others | -0.71 | 0.08 | -8.99 | 0.0000 |
| Married Hispanic - Married White | -1.21 | 0.05 | -22.92 | 0.0000 |
| Married Hispanic - Not married White | -0.45 | 0.05 | -8.57 | 0.0000 |
| Not married Hispanic - Married Others | -0.93 | 0.08 | -11.72 | 0.0000 |
| Not married Hispanic - Not married Others | -0.36 | 0.08 | -4.55 | 0.0000 |
| Not married Hispanic - Married White | -1.43 | 0.05 | -26.45 | 0.0000 |
| Not married Hispanic - Not married White | -0.68 | 0.05 | -12.54 | 0.0000 |
| Married Others - Not married Others | 0.57 | 0.07 | 8.59 | 0.0000 |
| Married Others - Married White | -0.50 | 0.06 | -7.74 | 0.0000 |
| Not married Others - Married White | -1.07 | 0.06 | -16.50 | 0.0000 |
| Not married Others - Not married White | -0.31 | 0.06 | -4.84 | 0.0000 |
| Married White - Not married White | 0.75 | 0.02 | 40.59 | 0.0000 |
| Married Others - Not married White | 0.26 | 0.06 | 4.01 | 0.0001 |
| Not married Black - Not married Hispanic | -0.16 | 0.07 | -2.38 | 0.0193 |
| Married Black - Married Hispanic | 0.16 | 0.07 | 2.14 | 0.0347 |
| Married Hispanic - Not married Others | -0.14 | 0.08 | -1.76 | 0.0820 |
| Married Black - Not married Others | 0.02 | 0.08 | 0.21 | 0.8318 |

Table 8: Post-hoc test of Marital status and Ethnicity

| Factor Levels | Estimate | Standard Error | t |
|---|----------|----------------|---|
| Married High School or less - Not married High School or less | 0.41 | 0.03 | |
| Married High School or less - Married Some college, diploma, assoc | -0.53 | 0.04 | |
| Married High School or less - Married Bachelors or higher | -1.80 | 0.05 | |
| Married High School or less - Not married Bachelors or higher | -1.16 | 0.05 | |
| Not married High School or less - Married Some college, diploma, assoc | -0.93 | 0.04 | |
| Not married High School or less - Not married Some college, diploma, assoc | -0.42 | 0.04 | |
| Not married High School or less - Married Bachelors or higher | -2.21 | 0.05 | |
| Not married High School or less - Not married Bachelors or higher | -1.57 | 0.05 | |
| Married Some college, diploma, assoc - Not married Some college, diploma, assoc | 0.52 | 0.03 | |
| Married Some college, diploma, assoc - Married Bachelors or higher | -1.28 | 0.05 | |
| Married Some college, diploma, assoc - Not married Bachelors or higher | -0.63 | 0.05 | |
| Not married Some college, diploma, assoc - Married Bachelors or higher | -1.79 | 0.05 | |
| Not married Some college, diploma, assoc - Not married Bachelors or higher | -1.15 | 0.05 | |
| Married Bachelors or higher - Not married Bachelors or higher | 0.64 | 0.03 | |
| Married High School or less - Not married Some college, diploma, assoc | -0.01 | 0.05 | |

Table 9: Post-hoc test of ms:education

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