

# Upper bounds for average Bayes accuracy in terms of mutual information

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These are preliminary notes.

## 1 Introduction

Suppose  $X$  and  $Y$  are continuous random variables (or vectors) which have a joint distribution with density  $p(x, y)$ . Let  $p(x) = \int p(x, y)dy$  and  $p(y) = \int p(x, y)dx$  denote the respective marginal distributions, and  $p(y|x) = p(x, y)/p(x)$  denote the conditional distribution.

Mutual information is defined

$$I[p(x, y)] = \int p(x, y) \log \frac{p(x, y)}{p(x)p(y)} dx dy.$$

$ABE_k$ , or  $k$ -class Average Bayes accuracy is defined as follows. Let  $X_1, \dots, X_K$  be iid from  $p(x)$ , and draw  $Z$  uniformly from  $1, \dots, k$ . Draw  $Y \sim p(y|X_Z)$ . Then, the average Bayes accuracy is defined as

$$ABA_k[p(x, y)] = \sup_f \Pr[f(X_1, \dots, X_k, Y) = Z]$$

where the supremum is taken over all functions  $f$ . A function  $f$  which achieves the supremum is

$$f_{Bayes}(x_1, \dots, x_k, y) = \operatorname{argmax}_{z \in \{1, \dots, k\}} p(y|x_z),$$

where an arbitrary rule can be employed to break ties. Such a function  $f_{Bayes}$  is called a *Bayes classification rule*. It follows that  $ABA_k$  is given explicitly

by

$$\text{ABA}_k = \frac{1}{k} \int \left[ \prod_{i=1}^k p(x_i) dx_i \right] \int dy \max_i p(y|x_i),$$

as stated in the following theorem.

**Theorem 1.1** *For a joint distribution  $p(x, y)$ , define*

$$\text{ABA}_k[p(x, y)] = \sup_f \Pr[f(x_1, \dots, x_k, y) = Z]$$

*where  $X_1, \dots, X_K$  are iid from  $p(x)$ ,  $Z$  is uniform from  $1, \dots, k$ , and  $Y \sim p(y|X_Z)$ , and the supremum is taken over all functions  $f : \mathcal{X}^k \times \mathcal{Y} \rightarrow \{1, \dots, k\}$ . Then,*

$$\text{ABA}_k = \frac{1}{k} \int \left[ \prod_{i=1}^k p(x_i) dx_i \right] \int dy \max_i p(y|x_i).$$

**Proof.** First, we claim that the supremum is attained by choosing

$$f(x_1, \dots, x_k, y) = \operatorname{argmax}_{z \in \{1, \dots, k\}} p(y|x_z).$$

To show this claim, write

$$\sup_f \Pr[f(X_1, \dots, X_k, Y) = Z] = \sup_f \frac{1}{k} \int p_X(x_1) \dots p_X(x_k) p(y|x_{f(x_1, \dots, x_k, y)}) dx_1 \dots dx_k dy$$

We see that maximizing  $\Pr[f(X_1, \dots, X_k, Y) = Z]$  over functions  $f$  additively decomposes into infinitely many subproblems, where in each subproblem we are given  $\{x_1, \dots, x_k, y\} \in \mathcal{X}^k \times \mathcal{Y}$ , and our goal is to choose  $f(x_1, \dots, x_k, y)$  from the set  $\{1, \dots, k\}$  in order to maximize the quantity  $p(y|x_{f(x_1, \dots, x_k, y)})$ . In each subproblem, the maximum is attained by setting  $f(x_1, \dots, x_k, y) = \operatorname{argmax}_z p(y|x_z)$ —and the resulting function  $f$  attains the supremum to the functional optimization problem. This proves the claim.

We therefore have

$$p(y|x_{f(x_1, \dots, x_k, y)}) = \max_{i=1}^k p(y|x_i).$$

Therefore, we can write

$$\begin{aligned}
\text{ABA}_k[p(x, y)] &= \sup_f \Pr[f(X_1, \dots, X_k, Y) = Z] \\
&= \frac{1}{k} \int p_X(x_1) \dots p_X(x_k) p(y | x_{f(x_1, \dots, x_k, y)}) dx_1 \dots dx_k dy. \\
&= \frac{1}{k} \int p_X(x_1) \dots p_X(x_k) \max_{i=1}^k p(y | x_i) dx_1 \dots dx_k dy.
\end{aligned}$$

## 2 Problem formulation

Let  $\mathcal{P}$  denote the collection of all joint densities  $p(x, y)$  on finite-dimensional Euclidean space. For  $\iota \in [0, \infty)$  define  $C_k(\iota)$  to be the largest  $k$ -class average Bayes error attained by any distribution  $p(x, y)$  with mutual information not exceeding  $\iota$ :

$$C_k(\iota) = \sup_{p \in \mathcal{P}: I[p(x, y)] \leq \iota} \text{ABA}_k[p(x, y)].$$

A priori,  $C_k(\iota)$  exists since  $\text{ABA}_k$  is bounded between 0 and 1. Furthermore,  $C_k$  is nondecreasing since the domain of the supremum is monotonically increasing with  $\iota$ .

It follows that for any density  $p(x, y)$ , we have

$$\text{ABA}_k[p(x, y)] \leq C_k(I[p(x, y)]).$$

Hence  $C_k$  provides an upper bound for average Bayes error in terms of mutual information.

Conversely we have

$$I[p(x, y)] \geq C_k^{-1}(\text{ABA}_k[p(x, y)])$$

so that  $C_k^{-1}$  provides a lower bound for mutual information in terms of average Bayes error.

On the other hand, there is no nontrivial *lower* bound for average Bayes error in terms of mutual information, nor upper bound for mutual information in terms of average Bayes error, since

$$\inf_{p \in \mathcal{P}: I[p(x, y)] \leq \iota} \text{ABA}_k[p(x, y)] = \frac{1}{k}.$$

regardless of  $\iota$ .

The goal of this work is to attempt to compute or approximate the functions  $C_k$  and  $C_k^{-1}$ .

## 2.1 Notation

$|\cdot|$  denotes set cardinality.

## 3 Theory

In this section we determine the value of  $C_k(\iota)$ , leading to the following result.

**Theorem 3.1** *For any  $\iota > 0$ , there exists  $c_\iota \geq 0$  such that defining*

$$Q_c(t) = \frac{\exp[ct^{k-1}]}{\int_0^1 \exp[ct^{k-1}]},$$

*we have*

$$\int_0^1 Q_{c_\iota}(t) \log Q_{c_\iota}(t) dt = \iota.$$

*Then,*

$$C_k(\iota) = \int_0^1 Q_{c_\iota}(t) t^{k-1} dt.$$

We obtain this result by first reducing the problem to the case of densities with uniform marginals, then doing the optimization over the reduced space.

### 3.1 Reduction

Let  $p(x, y)$  be a density supported on  $\mathcal{X} \times \mathcal{Y}$ , where  $\mathcal{X}$  is a subset of  $\mathbb{R}^{d_1}$  and  $\mathcal{Y}$  is a subset of  $\mathbb{R}^{d_2}$ , and such that  $p(x)$  is uniform on  $\mathcal{X}$  and  $p(y)$  is uniform on  $\mathcal{Y}$ .

Now let  $\mathcal{P}^{unif}$  denote the set of such distributions: in other words,  $\mathcal{P}^{unif}$  is the space of joint densities in Euclidean space with uniform marginals over the marginal supports. In this section, we prove that

$$C_k(\iota) \inf_{p \in \mathcal{P}: \mathbb{I}[p(x, y)] \leq \iota} \text{ABA}_k[p(x, y)] = \inf_{p \in \mathcal{P}^{unif}: \mathbb{I}[p(x, y)] \leq \iota} \text{ABA}_k[p(x, y)],$$

thus reducing the problem of optimizing over the space of all densities to the problem of optimizing over densities with uniform marginals.

Also define  $\mathcal{P}^{bounded}$  to be the space of all densities  $p(x, y)$  with finite-volume support. Since uniform distributions can only be defined over sets of finite volume, we have

$$\mathcal{P}^{unif} \subset \mathcal{P}^{bounded} \subset \mathcal{P}.$$

Therefore, it is necessary to first show that

$$\inf_{p \in \mathcal{P}: I[p(x, y)] \leq \epsilon} \text{ABA}_k[p(x, y)] = \inf_{p \in \mathcal{P}^{\text{bounded}}: I[p(x, y)] \leq \epsilon} \text{ABA}_k[p(x, y)].$$

This is accomplished via the following lemma.

**Lemma 3.2** (*Truncation*). *Let  $p(x, y)$  be a density on  $\mathbb{R}^{d_x} \times \mathbb{R}^{d_y}$ . For all  $\epsilon > 0$ , there exists a subset  $\mathcal{X} \subset \mathbb{R}^{d_x}$  with finite volume with respect to  $d_x$ -dimensional Lebesgue measure, and a subset  $\mathcal{Y} \subset \mathbb{R}^{d_y}$  with finite volume with respect to  $d_y$ -dimensional Lebesgue measure, such that defining*

$$\tilde{p}(x, y) = \frac{I\{(x, y) \in \mathcal{X} \times \mathcal{Y}\}}{\int_{\mathcal{X} \times \mathcal{Y}} p(x, y) dx dy} p(x, y),$$

we have

$$|I[p] - I[\tilde{p}]| < \epsilon$$

and

$$|\text{ABA}_k[p] - \text{ABA}_k[\tilde{p}]| < \epsilon.$$

**Proof.** Recall the definition of the Shannon entropy  $H$ :

$$H[p(x)] = - \int p(x) \log p(x) dx.$$

It is a well-known in information theory that

$$I[p(x, y)] = H[p(x)] + H[p(y)] - H[p(x, y)].$$

There exists a sequence  $(\mathcal{X}_i, \mathcal{Y}_i)_{i=1}^{\infty}$  where  $(\mathcal{X}_i)_{i=1}^{\infty}$  is an increasing sequence of finite-volume subsets of  $\mathbb{R}^{d_x}$  and  $(\mathcal{Y}_i)_{i=1}^{\infty}$  is an increasing sequence of finite-volume subsets of  $\mathbb{R}^{d_y}$ , and  $\lim_{i \rightarrow \infty} \mathcal{X}_i = \mathbb{R}^{d_x}$ ,  $\lim_{i \rightarrow \infty} \mathcal{Y}_i = \mathbb{R}^{d_y}$ . Define

$$\tilde{p}_i(x, y) = \frac{I\{(x, y) \in \mathcal{X}_i \times \mathcal{Y}_i\}}{\int_{\mathcal{X}_i \times \mathcal{Y}_i} p(x, y) dx dy} p(x, y)$$

Note that  $\tilde{p}_i$  gives the conditional distribution of  $(X, Y)$  conditional on  $(X, Y) \in \mathcal{X}_i \times \mathcal{Y}_i$ . Furthermore, it is convenient to define  $\tilde{p}_{\infty} = p$ . We can find some  $i_1$ , such that for all  $i \geq i_1$ , we have

$$\left| \int_{x \notin \mathcal{X}_i} p(x) \log p(x) dx \right| < \frac{\epsilon}{6}$$

$$\left| \int_{y \notin \mathcal{Y}_i} p(y) \log p(y) dy \right| < \frac{\epsilon}{6}$$

$$\left| \int_{(x,y) \notin \mathcal{X}_i \times \mathcal{Y}_i} p(x,y) \log p(x,y) dxdy \right| < \frac{\epsilon}{6}$$

and also such that

$$-\log \left[ \int_{x,y \in \mathcal{X}_i \times \mathcal{Y}_i} p(x,y) dxdy \right] < \frac{\epsilon}{2}$$

Then, it follows that

$$|\mathbf{I}[p] - \mathbf{I}[\tilde{p}_i]| < \epsilon$$

for all  $i \geq i_1$ .

Now we turn to the analysis of average Bayes error. Let  $f_i$  denote the Bayes  $k$ -class classifier for  $\tilde{p}_i(x,y)$  and  $f_\infty$  the Bayes  $k$ -class classifier for  $p(x,y)$ : recall that by definition,

$$\text{ABA}_k[\tilde{p}_i] = \Pr_{\tilde{p}_i}[f_i(X_1, \dots, X_k, Y) = Z]$$

Define

$$\epsilon_i = 1 - \int_{x,y \in \mathcal{X}_i \times \mathcal{Y}_i} p(x,y) dxdy.$$

We claim that

$$|\text{ABA}_k[\tilde{p}_i] - \text{ABA}_k[p]| \leq \epsilon_i.$$

Given the claim, the proof is completed by finding  $i > i_1$  such that  $\epsilon_i < \epsilon$ , and defining  $\mathcal{X} = \mathcal{X}_i$ ,  $\mathcal{Y} = \mathcal{Y}_i$ .

Consider using  $f_i$  to obtain a classification rule for  $p(x,y)$ : define

$$\tilde{f}_i = \begin{cases} f_i(x_1, \dots, x_k, y) & \text{when } (x_1, \dots, x_k, y) \in \mathcal{X}_i^k \times \mathcal{Y} \\ 0 & \text{otherwise.} \end{cases}$$

We have

$$\begin{aligned}
\text{ABA}_k[p] &= \sup_f \Pr_p[f(X_1, \dots, X_k, Y) = Z] \\
&\geq \\
&= (1 - \epsilon_i) \Pr_p[f_i(X_1, \dots, X_k, Y) = Z | (X_1, \dots, X_k, Y) \in \mathcal{X}_i^k \times \mathcal{Y}_i] \\
&\quad + \epsilon_i \Pr_p[f_i(X_1, \dots, X_k, Y) = Z | (X_1, \dots, X_k, Y) \notin \mathcal{X}_i^k \times \mathcal{Y}_i] \\
&= (1 - \epsilon_i) \Pr_{\tilde{p}}[f_i(X_1, \dots, X_k, Y) = Z] + \epsilon_i 0 \\
&= (1 - \epsilon_i) \text{ABA}_k[\tilde{p}_i] \geq \text{ABA}_k[\tilde{p}_i] - \epsilon_i.
\end{aligned}$$

In other words, when  $\tilde{p}_i$  is close to  $p$ , the Bayes classification rule for  $\tilde{p}_i$  obtains close to the Bayes rate when the data is generated under  $p$ .

Now consider the reverse scenario of using  $f_p$  to perform classification under  $\tilde{p}_i$ . This is equivalent to generating data under  $p(x, y)$ , performing classification using  $f$ , then only evaluating classification accuracy conditional on  $(X_1, \dots, X_k, Y) \in \mathcal{X}_i^k \times \mathcal{Y}_i$ . Therefore,

$$\begin{aligned}
\text{ABA}_k[\tilde{p}_i] &= \sup_f \Pr_{\tilde{p}_i}[f(X_1, \dots, X_k, Y) = Z] \\
&\geq \Pr_{\tilde{p}_i}[f_p(X_1, \dots, X_k, Y) = Z] \\
&= \Pr_p[f_p(X_1, \dots, X_k, Y) = Z | (X_1, \dots, X_k, Y) \in \mathcal{X}_i^k \times \mathcal{Y}_i] \\
&= \frac{1}{1 - \epsilon_i} \Pr_p[I\{(X_1, \dots, X_k, Y) \in \mathcal{X}_i^k \times \mathcal{Y}_i\} \text{ and } f_p(X_1, \dots, X_k, Y) = Z] \\
&\geq \frac{1}{1 - \epsilon_i} \left( 1 - \Pr_p[I\{(X_1, \dots, X_k, Y) \notin \mathcal{X}_i^k \times \mathcal{Y}_i\}] - \Pr_p[f_p(X_1, \dots, X_k, Y) \neq Z] \right) \\
&= \frac{\text{ABA}_k[p] - \epsilon_i}{1 - \epsilon_i} \geq \text{ABA}_k[p] - \epsilon_i.
\end{aligned}$$

In other words, when  $\tilde{p}_i$  is close to  $p$ , the Bayes classification rule for  $p$  obtains close to the Bayes rate when the data is generated under  $\tilde{p}_i$ .

Combining the two directions gives  $|\text{ABA}_k[\tilde{p}_i] - \text{ABA}_k[p]| \leq \epsilon_i$ , as claimed.

□

One can go from bounded-volume sets to uniform distributions by adding auxillary variables. To illustrate the intuition, consider a density  $p(x)$  on a

set of bounded volume,  $\mathcal{X}$ . Introduce a variable  $W$  such that conditional on  $X = x$ , we have  $w$  uniform on  $[0, p(x)]$ . It follows that the joint density  $p(x, w) = 1$  and is supported on a set  $\mathcal{X}' = \mathcal{X} \times [0, \infty]$ . Furthermore,  $\mathcal{X}'$  is of bounded volume (in fact, of volume 1) since

$$\int_{\mathcal{X}'} dx = \int_{\mathcal{X}'} p(x, w) dx = 1.$$

Therefore, to accomplish the reduction from  $\mathcal{P}$  to  $\mathcal{P}^{unif}$ , we start with a density  $p(x, y) \in \mathcal{P}$ , and using Lemma 3.2, find a suitable finite-volume truncation  $\tilde{p}(x, y)$ . Finally, we introduce auxillary variables  $w$  and  $z$  so that the expanded joint distribution  $p(x, w, y, z)$  has uniform marginals  $p(x, w)$  and  $p(y, z)$ . However, we still need to check that the introduction of auxillary variables preserves the mutual information and average Bayes error; this is the content of the next lemma.

**Lemma 3.3** *Suppose  $X, Y, W, Z$  are continuous random variables, and that  $W \perp Y|Z, Z \perp X|Y$ , and  $W \perp Z|(X, Y)$ . Then,*

$$I[p(x, y)] = I[p((x, w), (y, z))]$$

**Proof.** Due to conditional independence relationships, we have

$$p((x, w), (y, z)) = p(x, y)p(w|x)p(z|y).$$

It follows that

$$\begin{aligned} I[p((x, w), (y, z))] &= \int dx dw dy dz p(x, y)p(w|x)p(z|w) \log \frac{p((x, w), (y, z))}{p(x, w)p(y, z)} \\ &= \int dx dw dy dz p(x, y)p(w|x)p(z|w) \log \frac{p(x, y)p(w|x)p(z|y)}{p(x)p(y)p(w|x)p(z|y)} \\ &= \int dx dw dy dz p(x, y)p(w|x)p(z|w) \log \frac{p(x, y)}{p(x)p(y)} \\ &= \int dx dy p(x, y) \log \frac{p(x, y)}{p(x)p(y)} = I[p(x, y)]. \end{aligned}$$



Also,

$$\begin{aligned}
\text{ABA}_k[p((x, w), (y, z))] &= \int \left[ \prod_{i=1}^k p(x_i, w_i) dx_i dw_i \right] \int dy dz \max_i p(y, z | x_i, w_i). \\
&= \int \left[ \prod_{i=1}^k p(x_i, w_i) dx_i dw_i \right] \int dy \max_i p(y | x_i) \int dz p(z | y). \\
&= \int \left[ \prod_{i=1}^k p(x_i) dx_i \right] \left[ \prod_{i=1}^k \int dw_i p(w_i | x_i) \right] \int dy \max_i p(y | x_i) \\
&= \text{ABA}_k[p(x, y)].
\end{aligned}$$

□

Combining these lemmas gives the needed reduction, given by the following theorem.

**Theorem 3.4** (*Reduction.*)

$$\inf_{p \in \mathcal{P}: I[p(x, y)] \leq \iota} \text{ABA}_k[p(x, y)] = \inf_{p \in \mathcal{P}^{unif}: I[p(x, y)] \leq \iota} \text{ABA}_k[p(x, y)].$$

The proof is trivial given the previous two lemmas.

## 3.2 Optimization

Having reduced the problem to an optimization over  $\mathcal{P}^{unif}$ , in this section we use variational calculus to find the global optimum to the optimization problem

$$\text{maximize}_{p \in \mathcal{P}^{unif}: I[p(x, y)] \leq \iota} \text{ABA}_k[p(x, y)]$$

The proof depends on the following lemmas.

**Lemma 3.5** *Let  $f(t)$  be an increasing function from  $[a, b] \rightarrow \mathbb{R}$ , where  $a < b$ , and let  $g(t)$  be a bounded continuous function from  $[a, b] \rightarrow \mathbb{R}$ . Define the set*

$$A = \{t : f(t) \neq g(t)\}.$$

*Then, we can write  $A$  as a countable union of intervals*

$$A = \bigcup_{i=1}^{\infty} A_i$$

*where  $A_i$  are mutually disjoint intervals, with  $\inf A_i < \sup A_i$ , and for each  $i$ , either  $f(t) > g(t)$  for all  $t \in A_i$  or  $f(t) < g(t)$  for all  $t \in A_i$ .*

**Lemma 3.6** *Let  $f(t)$  be a measurable function from  $[a, b] \rightarrow \mathbb{R}$ , where  $a < b$ . Then there exists sets  $\mathcal{B}_0$  and  $\mathcal{B}_1$ , satisfying the following properties:*

- $\mathcal{B} = \mathcal{B}_0 \cup \mathcal{B}_1$  is countable partition of  $[a, b]$ ,
- $f(t)$  is constant on all  $B \in \mathcal{B}_0$ , but not constant on any proper super-interval  $B' \supset B$ , and
- $B \in \mathcal{B}_1$  contains no positive-length subinterval where  $f(t)$  is constant.

**Lemma 3.7** *Define an exponential family on  $[0, 1]$  by the density function*

$$q_\beta(t) = \exp[\beta t^{k-1} - \log Z(\beta)]$$

where

$$Z(\beta) = \int_0^1 \exp[\beta t^{k-1}] dt.$$

Then, the negative entropy

$$I(\beta) = \int_0^1 q_\beta(t) \log q_\beta(t) dt$$

is decreasing in  $\beta$  on the interval  $(-\infty, 0]$ . and increasing on the interval  $[0, \infty)$ .

Furthermore, for any  $\iota \in (0, \infty)$ , there exist two solutions to  $I(\beta) = \iota$ : one positive and one negative.

**Lemma 3.8** *For any measure  $G$  on  $[0, \infty]$ , let  $G^k$  denote the measure defined by*

$$G^k(A) = G(A)^k,$$

and define

$$E[G] = \int x dG(x).$$

$$I[G] = \int x \log x dG(x)$$

and

$$\psi_k[G] = \int x d(G^k)(x).$$

Then, defining  $Q_c$  and  $c_\iota$  as in Theorem 1, we have

$$\sup_{G: E[G]=1, I[G] \leq \iota} \psi_k[G] = \int_0^1 Q_{c_\iota}(t) t^{k-1} dt.$$

Furthermore, the supremum is attained by a measure  $G$  that has cdf equal to  $Q_c^{-1}$ , and thus has a density  $g$  with respect to Lesbegue measure.

**Lemma 3.9** *The map*

$$\iota \rightarrow \int_0^1 Q_{c_\iota}(t) t^{k-1} dt$$

*is concave in  $\iota > 0$ .*

**Proof of Lemma 3.5.** (This will appear in the appendix of the paper.)

The function  $h(t) = f(t) - g(t)$  is measurable, since all increasing functions are measurable. Define  $A^+ = \{t : f(t) > g(t)\}$  and  $A^- = \{t : f(t) < g(t)\}$ . Since  $A^+$  and  $A^-$  are measurable subsets of  $\mathbb{R}$ , they both admit countable partitions consisting of open, closed, or half-open intervals. Let  $\mathcal{H}^+$  be the collection of all partitions of  $A^+$  consisting of such intervals. There exists a least refined partition  $\mathcal{A}^+$  within  $\mathcal{H}^+$ . Define  $\mathcal{A}^-$  analogously, and let

$$\mathcal{A} = \mathcal{A}^+ \cup \mathcal{A}^-$$

and enumerate the elements

$$\mathcal{A} = \{A_i\}_{i=1}^\infty.$$

We claim that the partitions  $\mathcal{A}^+$  and  $\mathcal{A}^-$  have the property that for all  $t \in A^\pm$ , the interval  $I \in \mathcal{A}^\pm$  containing  $t$  has endpoints  $l \leq u$  defined by

$$l = \inf_{x \in [a, b]} \{x : \text{Sign}(h([x, t])) = \{\text{Sign}(h(t))\}\}$$

and

$$u = \sup_{x \in [a, b]} \{x : \text{Sign}(h([t, x])) = \{\text{Sign}(h(t))\}\}.$$

We prove the claim for the partition  $\mathcal{A}^+$ . Take  $t \in A^+$  and define  $l$  and  $u$  as above. It is clear that  $(l, u) \in A^+$ , and furthermore, there is no  $l' < l$  and  $u' > u$  such that  $(l', x) \in A^+$  or  $(x, u') \in A^+$  for any  $x \in I$ . Let  $\mathcal{H}$  be any other partition of  $A^+$ . Some disjoint union of intervals  $H_i \in \mathcal{H}$  necessarily

covers  $I$  for  $i = 1, \dots$ , and we can further require that none of the  $H_i$  are disjoint with  $I$ . Since each  $H_i$  has nonempty intersection with  $I$ , and  $I$  is an interval, this implies that  $\cup_i H_i$  is also an interval. Let  $l'' \leq u''$  be the endpoints of  $\cup_i H_i$ . Since  $I \subseteq \cup_i H_i$ , we have  $l'' \leq l \leq u \leq u''$ . However, since also  $I \in \mathcal{A}^+$ , we must have  $l \leq l'' \leq u'' \leq u$ . This implies that  $l'' = l$  and  $u'' = u$ . Since  $\cup_i H_i = I$ , and this holds for any  $I \in \mathcal{A}^+$ , we conclude that  $\mathcal{H}$  is a refinement of  $\mathcal{A}^+$ . The proof of the claim for  $\mathcal{A}^-$  is similar.

It remains to show that there are not isolated points in  $\mathcal{A}$ , i.e. that for all  $I \in \mathcal{A}$  with endpoints  $l \leq u$ , we have  $l < u$ . Take  $I \in \mathcal{A}$  with endpoints  $l \leq u$  and let  $t = \frac{l+u}{2}$ . By definition, we have  $h(t) \neq 0$ . Consider the two cases  $h(t) > 0$  and  $h(t) < 0$ .

If  $h(t) > 0$ , then  $t' = g^{-1}(h(t)) > t$ , and for all  $x \in [t, t']$  we have  $h(x) > 0$ . Therefore, it follows from definition that  $[t, t'] \in I$ , and since  $l \leq t < t' \leq u$ , this implies that  $l < u$ . The case  $h(t) < 0$  is handled similarly.  $\square$

**Proof of Lemma 3.6.** (This will appear in the appendix of the paper.) To construct the interval, define

$$l(t) = \inf\{x \in [0, 1] : f([x, t]) = \{f(t)\}\}$$

$$u(t) = \sup\{x \in [0, 1] : f([t, x]) = \{f(t)\}\},$$

Let  $B_0$  be the set of all  $t$  such that  $l(t) < u(t)$ , and let  $B_1$  be the set of all  $t$  such that  $l(t) = t = u(t)$ . For all  $t \in B_0$ , define

$$I(t) = (l(t), u(t)) \cup \{x \in \{l(t), u(t)\} : f(x) = f(t)\}.$$

Then we claim

$$\mathcal{B}_0 = \{I(t) : t \in B_0\}$$

is a countable partition of  $B_0$ . The claim follows since the members of  $\mathcal{B}_0$  are disjoint intervals of nonzero length, and  $B_0$  has finite length. It follows from definition that for any  $B \in \mathcal{B}_0$ , that  $f$  is not constant on any proper superinterval  $B' \supset B$ .

Meanwhile, let  $\mathcal{B}_1$  be a countable partition of  $B_1$  into intervals.

Next, we show that for all  $I \in \mathcal{B}_1$ ,  $I$  does not contain a subinterval  $I'$  of nonzero length such that  $f$  is constant on  $I'$ . Suppose to the contrary, we could find such an interval  $I$  and subinterval  $I'$ . Then for any  $t \in I'$ , we have  $t \in B_0$ . However, this implies that  $t \notin B_1$ , a contradiction.

Since  $t \in [a, b]$  belongs to either  $B_0$  or  $B_1$ , letting  $\mathcal{B} = \mathcal{B}_0 \cup \mathcal{B}_1$  yields the desired partition of  $[a, b]$ .  $\square$ .

**Proof of Lemma 3.7.**

Define  $\beta(\mu)$  as the solution to

$$\mu = \int_0^1 t q_\beta(t) dt.$$

By [Wainwright and Jordan 2008], the function  $\beta(\mu)$  is well-defined. Furthermore, since the sufficient statistic  $t^{k-1}$  is increasing in  $t$ , it follows that  $\beta(\mu)$  is increasing.

Define the negative entropy as a function of  $\mu$ ,

$$N(\mu) = \int_0^1 q_{\beta(\mu)}(t) \log q_{\beta(\mu)}(t) dt.$$

By Theorem 3.4 of [Wainwright and Jordan 2008],  $N(\mu)$  is convex in  $\mu$ . We claim that the derivative of  $N(\mu) = 0$  at  $\mu = \frac{1}{2}$ . This implies that  $N(\mu)$  is decreasing in  $\mu$  for  $\mu \leq \frac{1}{2}$  and increasing for  $\mu \geq \frac{1}{2}$ . Since  $I(\beta(\mu)) = N(\mu)$ ,  $\beta$  is increasing in  $\mu$ , and  $\beta(\frac{1}{2}) = 0$ , this implies that  $I(\beta)$  is decreasing in  $\beta$  for  $\beta \leq 0$  and increasing for  $\beta \geq 0$ .

We will now prove the claim. Write

$$\left. \frac{d}{d\mu} N(\mu) \right|_{\mu=1/2} = \left. \frac{d}{d\beta} I(\beta(\mu)) \right|_{\beta=0} \left. \frac{d\beta}{d\mu} \right|_{\mu=1/2}.$$

We have

$$\frac{d}{d\beta} I(\beta) = \beta \int q_\beta t^{k-1} dt - \log Z(\beta).$$

Meanwhile,  $Z(0) = 1$  so  $\log Z(0) = 0$ . Therefore,

$$\left. \frac{d}{d\beta} I(\beta) \right|_{\beta=0} = 0.$$

This implies that  $\left. \frac{d}{d\mu} N(\mu) \right|_{\mu=1/2} = 0$ , as needed.

For the final statement of the lemma, note that  $I(0) = 0$  since  $q_0$  is the uniform distribution. Meanwhile, since  $q_\beta$  tends to a point mass as either  $\beta \rightarrow \infty$  or  $\beta \rightarrow -\infty$ , we have

$$\lim_{\beta \rightarrow \infty} I(\beta) = \lim_{\beta \rightarrow -\infty} I(\beta) = \infty.$$

And, as we can check that  $I(\beta)$  is continuous in  $\beta$ , this means that

$$I((-\infty, 0]) = I([0, \infty)) = [0, \infty)$$

by the mean-value theorem. Combining this fact with the monotonicity of  $I(\beta)$  restricted to either the positive and negative half-line yields the fact that for any  $\iota > 0$ , there exists  $\beta_1 < 0 < \beta_2$  such that  $I(\beta_1) = I(\beta_2) = \iota$ .  $\square$ .

**Proof of Lemma 3.8.** (This will appear in the appendix of the paper.)

Consider the quantile function  $Q(t) = \inf_{x \in [0, 1]} : G((-\infty, x]) \geq t$ .  $Q(t)$  must be a monotonically increasing function from  $[0, 1]$  to  $[0, \infty)$ . Let  $\mathcal{Q}$  denote the collection of all such quantile functions.

We have

$$E[G] = \int_0^1 Q(t) dt$$

$$\psi_k[G] = \int_0^1 Q(t) x^{k-1} dt.$$

and

$$I[G] = \int_0^1 Q(t) \log Q(t) dt.$$

For any given  $\iota$ , let  $P_\iota$  denote the class of probability distributions  $G$  on  $[0, \infty]$  such that  $E[G] = 1$  and  $I[G] \leq \iota$ . From Markov's inequality, for any  $G \in P_\iota$  we have

$$G([x, \infty]) \leq x^{-1}$$

for any  $x \geq 0$ , hence  $P_\iota$  is tight. From tightness, we conclude that  $P_\iota$  is closed under limits with respect to weak convergence. Hence, since  $\psi_k$  is a continuous function, there exists a distribution  $G^* \in P_\iota$  which attains the supremum

$$\sup_{G \in P_\iota} \psi_k[G].$$

Let  $\mathcal{Q}_\iota$  denote the collection of quantile functions of distributions in  $P_\iota$ . Then,  $\mathcal{Q}_\iota$  consists of monotonic functions  $Q : [0, 1] \rightarrow [0, \infty]$  which satisfy

$$E[Q] = \int_0^1 Q(t) dt = 1,$$

and

$$I[Q] = \int_0^1 Q(t) \log Q(t) dt \leq \iota.$$

Let  $\mathcal{Q}$  denote the collection of *all* quantile functions from measures on  $[0, \infty]$ . And letting  $Q^*$  be the quantile function for  $G^*$ , we have that  $Q^*$  attains the supremum

$$\sup_{Q \in \mathcal{Q}_t} \phi_k[Q] = \sup_{Q \in \mathcal{Q}_t} \int_0^1 Q(t) t^{k-1} dt. \quad (1)$$

Therefore, there exist Lagrange multipliers  $\lambda \geq 0$  and  $\nu \geq 0$  such that defining

$$\mathcal{L}[Q] = -\phi_k[Q] + \lambda E[Q] + \nu I[Q] = \int_0^1 Q(t) (-t^{k-1} + \lambda + \nu \log Q(t)) dt,$$

$Q^*$  attains the infimum of  $\mathcal{L}[Q]$  over *all* quantile functions,

$$\mathcal{L}[Q^*] = \inf_{Q \in \mathcal{Q}} \mathcal{L}[Q].$$

The global minimizer  $Q^*$  is also necessarily a stationary point: that is, for any perturbation function  $\xi : [0, 1] \rightarrow \mathbb{R}$  such that  $Q^* + \xi \in \mathcal{Q}$ , we have  $\mathcal{L}[Q^*] \leq \mathcal{L}[Q^* + \xi]$ . For sufficiently small  $\xi$ , we have

$$\mathcal{L}[Q + \xi] \approx \mathcal{L}[Q] + \int_0^1 \xi(t) (-t^{k-1} + \lambda + \nu + \nu \log Q(t)) dt. \quad (2)$$

Define

$$\nabla \mathcal{L}_{Q^*}(t) = -t^{k-1} + \lambda + \nu + \nu \log Q(t). \quad (3)$$

The function  $\nabla \mathcal{L}_{Q^*}(t)$  is a *functional derivative* of the Lagrangian. Note that if we were able to show that  $\nabla \mathcal{L}_{Q^*}(t) = 0$ , this immediately yields

$$Q^*(t) = \exp[-1 - \lambda\nu^{-1} + \nu^{-1}t^{k-1}]. \quad (4)$$

At this point, we know that the right-hand side of (4) gives a stationary point of  $\mathcal{L}$ , but we cannot be sure that it gives the global minimizer. The reason is because the optimization occurs on a constrained space. We will show that (4) indeed gives the global minimizer  $Q^*$ , but we do so by showing that the set of points  $t$  where  $\nabla \mathcal{L}_{Q^*}(t) \neq 0$  is of zero measure. Since sets of zero measure don't affect the integrals defining the optimization problem (1), we conclude there exists a global optimal solution with  $\nabla \mathcal{L}_{Q^*}(t) = 0$  everywhere, which is therefore given explicitly by (4) for some  $\lambda \in \mathbb{R}$ ,  $\nu \geq 0$ .

We will need the following result: that for  $\iota > 0$ , any solution to (1) satisfies  $\phi_k[Q] < 1$ . This follows from the fact that

$$E[Q] - \phi_k[Q] = \int_0^1 (1 - t^{k-1})Q(t)dt,$$

where the term  $(1 - t^{k-1})$  is negative, except for the one point  $t = 1$ . Therefore, in order for  $\phi_k[Q] = 1 = E[Q]$ , we must have  $Q(t) = 0$  for  $t < 1$ . However, this yields a contradiction since  $Q(t) = 0$  for  $t < 1$  implies that  $E[Q] = 0$ , a violation of the hard constraint  $E[Q] = 1$ .

Let us establish that  $\nu > 0$ : in other words, the constraint  $I[Q] = \iota$  is tight. Suppose to the contrary, that for some  $\iota > 0$ , the global optimum  $Q^*$  minimizes a Lagrangian with  $\nu = 0$ . Let  $\phi^* = \phi_k[Q^*] < 1$ . However, if we define  $Q_\kappa(t) = I\{t \geq 1 - \frac{1}{\kappa}\}\kappa$ , we have  $E[Q_\kappa] = 1$ , and also for some sufficiently large  $\kappa > 0$ ,  $\phi_k[Q_\kappa] > \phi^*$ . But since the Lagrangian lacks a term corresponding to  $I[Q]$ , we conclude that  $\mathcal{L}[Q_\kappa] < \mathcal{L}[Q^*]$ , a contradiction.

The rest of the proof proceeds as follows. We will use Lemmas 3.5 and 3.6 to define a decomposition  $A = D_0 \cup D_1 \cup D_2$ , where  $D_2$  is of measure zero. First, we show that assuming the existence of  $t \in D_0$  yields a contradiction, and hence  $D_0 = \emptyset$ . Then, again using argument from contradiction we establish that  $D_1 = \emptyset$ . Finally, since  $D_2$  is a set of zero measure, this allows us to conclude that the  $Q^*(t) = 0$  on all but a set of zero measure.

We will now apply the Lemmas to obtain the necessary ingredients for constructing the sets  $D_i$ . Since  $\nabla \mathcal{L}_{Q^*}(t)$  is a difference between an increasing function and a continuous strictly increasing function, we can apply Lemma 3.5 to conclude that there exists a countable partition  $\mathcal{A}$  of the set  $A : \{t \in [0, 1] : \nabla \mathcal{L}_{Q^*}(t) \neq 0\}$  into intervals such that for all  $J \in \mathcal{A}$ ,  $|\text{Sign}(\nabla Q^*(J))| = 1$  and  $\inf J < \sup J$ . Applying Lemma 3.6 we get a countable partition  $\mathcal{B} = \mathcal{B}_0 \cup \mathcal{B}_1$  of  $[0, 1]$  so that each element  $J \in \mathcal{B}_0$  is an interval such that  $\nabla \mathcal{L}_{Q^*}(t)$  is constant on  $J$ , and furthermore is not properly contained in any interval with the same property, and each element  $J \in \mathcal{B}_1$  is an interval, such that  $J$  contains no positive-length subinterval where  $\nabla \mathcal{L}_{Q^*}(t)$  is constant. Also define  $B_i$  as the union of the sets in  $\mathcal{B}_i$  for  $i = 0, 1$ .

Note that  $B_0$  is necessarily a subset of  $A$ . That is because if  $\nabla \mathcal{L}_{Q^*}(t) = 0$  on any interval  $J$ , then that  $Q^*(t)$  is necessarily not constant on the interval.

We will construct a new countable partition of  $A$ , called  $\mathcal{D}$ . The partition  $\mathcal{D}$  is constructed by taking the union of three families of intervals,

$$\mathcal{D} = \mathcal{D}_0 \cup \mathcal{D}_1 \cup \mathcal{D}_2.$$



Define  $D_i$  to be the union of intervals in  $\mathcal{D}_i$  for  $i = 0, 1, 2$ .

Define  $\mathcal{D}_0 = \mathcal{B}_0$ , Define a countable partition  $\mathcal{D}_1$  by

$$\mathcal{D}_1 = \{J \cap L : J \in \mathcal{A}, L \in \mathcal{B}_1, \text{ and } |L| > 1\},$$

in order words,  $\mathcal{D}_1$  consists of positive-length intervals where  $\nabla Q^*(t)$  is entirely positive or negative and is not constant. Define

$$\mathcal{D}_2 = \{J \in \mathcal{B}_1 : J \subset A \text{ and } |J| = 1\},$$

i.e.  $\mathcal{D}_2$  consists of isolated points in  $A$ .

One verifies that  $\mathcal{D}$  is indeed a partition of  $A$  by checking that  $D_0 = B_0$ ,  $D_1 \cup D_2 = B_1 \cap A$ , so that  $D_0 \cup D_1 \cup D_2 = A$ : it is also easy to check that elements of  $\mathcal{D}$  are disjoint. Furthermore, as we mentioned earlier, the set  $D_2$  is indeed of zero measure, since it consists of countably many isolated points.

Now we will show that the existence of  $t \in D_0$  implies a contradiction. Take  $t \in D$  for  $D \in \mathcal{D}_0$ , and let  $a = \inf D$  and  $b = \sup D$ . Define

$$\xi^+ = I\{t \in D\}(Q^*(b) - Q^*(t))$$

and

$$\xi^- = I\{t \in D\}(Q^*(a) - Q^*(t)).$$

Observe that  $Q + \epsilon \xi^+ \in \mathcal{Q}$  and  $Q + \epsilon \xi^- \in \mathcal{Q}$  for any  $\epsilon \in [0, 1]$ . Now, if  $\nabla \mathcal{L}_{Q^*}(t)$  is strictly positive on  $D$ , then for some  $\epsilon > 0$  we would have  $\mathcal{L}[Q^* + \epsilon \xi^-] < \mathcal{L}[Q^*]$ , a contradiction. A similar argument with  $\xi^+$  shows that  $\nabla \mathcal{L}_{Q^*}(t)$  cannot be strictly negative on  $D$  either. From this perturbation argument, we conclude that  $\nabla \mathcal{L}_{Q^*}(t) = 0$ . Since this argument applies for all  $t \in D_0$ , we conclude that  $D_0 = \emptyset$ : therefore, on the set  $[0, 1] \setminus (D_1 \cup D_2)$ , we have  $\nabla \mathcal{L}_{Q^*}(t) = 0$ .

The following observation is needed for the next stage of the proof. If we look at the function  $Q^*(t)$ , then up to sets of negligible measure, it is given by the expression (4) on the set  $[0, 1] \setminus D_1$ , and it is piecewise constant in-between. But since (4) gives a strictly increasing function, and since  $Q^*$  is increasing, this implies that  $Q^*$  is discontinuous at the boundary of  $D_1$ .

Now we are prepared to show that  $\nabla \mathcal{L}_{Q^*}(t) = 0$  for  $t \in D_1$ . Take  $t \in D$  for  $D \in \mathcal{D}_1$ , and let  $a = \inf D$  and  $b = \sup D$ . From the previous argument, there is a discontinuity at both  $a$  and  $b$ , so that  $\lim_{u \rightarrow a^-} Q(u) < Q(t) < \lim_{u \rightarrow b^+} Q(u)$ . Therefore, for any  $\xi(t)$  which is increasing on  $D$  and zero

elsewhere, there exists  $\epsilon > 0$  such that  $\nabla Q^* + \epsilon \xi \in \mathcal{Q}$ . It remains to find such a perturbation  $\xi$  such that  $\mathcal{L}[Q + \epsilon \xi] < \mathcal{L}[Q]$ .

Also, since by definition  $\nabla \mathcal{L}_{Q^*}(t)$  is constant on  $D$ , follows from (3) that  $\nabla Q^*$  is strictly decreasing, and thus either

- Case 1:  $\nabla \mathcal{L}_{Q^*}(t) \geq 0$  on  $D$ ,
- Case 2:  $\nabla \mathcal{L}_{Q^*}(t) \leq 0$  on  $D$ , or
- Case 3:  $\nabla \mathcal{L}_{Q^*}(t) \geq 0$  for all  $t \in D \cap [a, t_0]$  and  $\nabla \mathcal{L}_{Q^*}(t) \leq 0$  for all  $t \in D \cap [t_0, b]$ .

Depending on the case, we construct a suitable perturbation  $\xi$ :

- Case 1: Construct  $\xi(t) = -I\{t \in D\}$ .
- Case 2: Construct  $\xi(t) = I\{t \in D\}$
- Case 2: Construct

$$\xi(t) = \begin{cases} -1 & \text{for } t \in D \cap [a, t_0], \\ 0 & \text{otherwise.} \end{cases}$$

In all three cases, given the corresponding construction for  $\xi(t)$  we get

$$\int_0^1 \xi(t) \nabla \mathcal{L}_{Q^*}(t) dt < 0.$$

Therefore, from (2), there exists some  $\epsilon > 0$  such that  $\mathcal{L}[Q + \epsilon \xi] < \mathcal{L}[Q]$ , a contradiction. Again, since the contradiction applies for all  $t \in D_1$ , we conclude that  $D_1 = \emptyset$ .

By now we have established that a global optimum for (1) exists, and is given by (4) for some  $\lambda \in \mathbb{R}$ ,  $\nu > 0$ . It remains to determine the values of  $\lambda$  and  $\nu$ .

Reparameterize  $\alpha = \exp[-1 - \lambda \nu^{-1}]$  and  $\beta = \nu^{-1}$ . Therefore,

$$Q^*(t) = \alpha \exp[\beta t^{k-1}]$$

for  $\alpha > 0$ ,  $\beta > 0$ . There is a one-to-one mapping from  $(\alpha, \beta) \in (0, \infty)^2$  to  $(\lambda, \nu) \in \mathbb{R} \times (0, \infty)$ .

Now, from the constraint

$$1 = E[Q^*] = \int_0^1 \alpha \exp[\beta t^{k-1}] dt.$$

we conclude that

$$\alpha = \frac{1}{\int_0^1 \exp[\beta t^{k-1}] dt}.$$

Therefore, we have reduced the set of possible solutions  $Q^*$  to a one-parameter family,

$$Q^*(t) = \frac{\exp[\beta t^{k-1}]}{Z(\beta)}.$$

where

$$Z(\beta) = \int_0^1 \exp[\beta t^{k-1}] dt.$$

Next, note that

$$I[Q^*] = \int_0^1 Q^*(t) \log Q^*(t) dt = \beta \mu_\beta - \log Z(\beta),$$

as a function of  $\beta$ , is completely characterized by Lemma 3.7. Let us define  $c_\iota$  as the unique positive solution to the equation

$$c_\iota \mu_{c_\iota} - \log Z(c_\iota) = \iota$$

given by Lemma 3.7. We therefore have

$$Q^*(t) = \frac{\exp[c_\iota t^{k-1}]}{\int_0^1 \exp[c_\iota t^{k-1}] dt},$$

as needed.  $\square$

### **Proof of theorem 3.1**

Using Theorem 3.4, we have

$$C_k(\iota) = \inf_{p \in \mathcal{P}^{unif}: I[p(x, y)] \leq \iota} \text{ABA}_k[p(x, y)].$$

Define  $f(\iota) = \int_0^1 Q_{c_\iota}(t) t^{k-1} dt$ : our goal is to establish that  $C_k(\iota) = f(\iota)$ . Note that  $f(\iota)$  is the same function which appears in Lemma 3.9 and the same bound as established in Lemma 3.8.

Define the density  $p_\iota(x, y)$  where

$$p_\iota(x, y) = \begin{cases} g_\iota(y - x) & \text{for } x \geq y \\ g_\iota(1 + y - x) & \text{for } x < y \end{cases}$$

where

$$g_\iota(x) = \frac{d}{dx} G_\iota(x)$$

and  $G_\iota$  is the inverse of  $Q_c$ .

One can verify that  $I[p_\iota] = \iota$ , and

$$\text{ABA}_k[p] = \int_0^1 Q_{c_\iota}(t) t^{k-1} dt.$$

This establishes that

$$C_k(\iota) \geq \int_0^1 Q_{c_\iota}(t) t^{k-1} dt.$$

It remains to show that for all  $p \in \mathcal{P}^{unif}$  with  $I[p] \leq \iota$ , that  $\text{ABA}_k[p] \leq \text{ABA}_k[p_\iota]$ .

Take  $p \in \mathcal{P}^{unif}$  such that  $I[p] \leq \iota$ . Letting  $X_1, \dots, X_k \sim \text{Unif}[0, 1]$ , and  $Y \sim \text{Unif}[0, 1]$  define  $Z_i(y) = p(y|X_i)$ . We have  $\mathbf{E}(Z(y)) = 1$  and,

$$I[p(x, y)] = \mathbf{E}(Z(Y) \log Z(Y))$$

while

$$\text{ABA}_k[p(x, y)] = k^{-1} \mathbf{E}(\max_i Z_i(Y)).$$

Letting  $G_y$  be the distribution of  $Z(y)$ , we have

$$E[G_y] = 1$$

$$I[p(x, y)] = \mathbf{E}(I[G_Y])$$

$$\text{ABA}_k[p(x, y)] = \mathbf{E}(\psi_k[G_Y])$$

where the expectation is taken over  $Y \sim \text{Unif}[0, 1]$  and where  $E[G]$ ,  $I[G]$ , and  $\psi_k[G]$  are defined as in Lemma 3.8.

Define the random variable  $J = I[G_Y]$ . We have

$$\begin{aligned}
\text{ABA}_k[p(x, y)] &= \mathbf{E}(\psi_k[G_Y]) \\
&= \int_0^1 \psi_k[G_y] dy \\
&\leq \int_0^1 \left( \sup_{G: I[G] \leq I[G_y]} \psi_k[G] \right) dy \\
&= \int_0^1 f(I[G_y]) dy = \mathbf{E}[f(J)].
\end{aligned}$$

Now, since  $f$  is concave by Lemma 3.9, we can apply Jensen's inequality to conclude that

$$\text{ABA}_k[p(x, y)] = \mathbf{E}[f(J)] \leq f(\mathbf{E}[J]) = f(\iota),$$

which completes the proof.  $\square$