Supplemental material: How many faces can it recognize? Performance extrapolation for multi-class classification

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1 Proofs

Theorem 2.1. (i) an OVO recognition systems equipped with a continuous binary classifier is separable; (ii) an OVA recognition systems equipped with a continuous binary scoring rule is separable; (iii) a kNN recognition system with fixed neighborhood size $\alpha \in (0,1)$ is separable; (iv) multinomial logistic regression recognition systems are separable; (v) LDA recognition systems are separable.

Theorem 3.1. Let U be defined as the random variable

$$U = u(F, Y)$$

for X, Y drawn from p(x, y) = p(x)p(y|x), and $\hat{F}(X) = \frac{1}{r_1} \sum_{j=1}^{r_1} \delta Y^j$ with $Y^i \stackrel{iid}{\sim} p(y|X)$ Then $p_k = \mathbf{E}[U^{k-1}]$.

Proof. Write $q^{(i)}(y) = \mathcal{Q}(\hat{F}_i, 0, y)$, and let $Y^{(i),*} \sim p(y|X^{(i)})$ for i = 1, ..., k. Note that by using conditioning and conditional independence, p_k can be

written

$$\begin{split} p_k &= \mathbf{E} \left[\frac{1}{k} \sum_{i=1}^k \Pr[q^{(i)}(Y^{(i),*}) > \max_{j \neq i} q^{(j)}(Y^{(i),*})] \right] \\ &= \mathbf{E} \left[\Pr[q^{(1)}(Y^{(1),*}) > \max_{j \neq 1} q^{(j)}(Y^{(1),*})] \right] \\ &= \mathbf{E}[\Pr[q^{(1)}(Y^{(1),*}) > \max_{j \neq 1} q^{(j)}(Y^{(1),*})|Y^{(1),*}, \hat{F}_1]] \\ &= \mathbf{E}[\Pr[\cap_{j>1} q^{(1)}(Y^{(1),*}) > q^{(j)}(Y^{(1),*})|Y^{(1),*}, \hat{F}_1]] \\ &= \mathbf{E}[\prod_{j>1} \Pr[q^{(1)}(Y^{(1),*}) > q^{(j)}(Y^{(1),*})|Y^{(1),*}, \hat{F}_1]] \\ &= \mathbf{E}[\Pr[q^{(1)}(Y^{(1),*}) > q^{(2)}(Y^{(1),*})|Y^{(1),*}, \hat{F}_1]^{k-1}] \\ &= \mathbf{E}[u(\hat{F}_1, Y^{(1),*})^{k-1}] = \mathbf{E}[U^{k-1}]. \end{split}$$

Theorem 3.2. Let U be defined as in Theorem 2.1, and let ν denote the law of U. Then, for any probability distribution ν' on [0,1], one can construct a joint distribution p(x,y) and a scoring rule Q such that $\nu = \nu'$.

Proof. Let X and Y have the degenerate joint distribution $X = Y \sim Unif[0,1]$. Let G be the cdf of ν , $G(x) = \int_0^x d\nu(x)$, and let $H(u) = \sup_x \{G(x) \leq u\}$. Define \mathcal{Q} by

$$\mathcal{Q}(F,0,y) = \begin{cases} 0 & \text{if } \mu(F) > y + H(y) \\ 0 & \text{if } y + H(y) > 1 \text{ and } \mu(F) \in [H(y) - y, y] \\ 1 + \mu(F) - y & \text{if } \mu(F) \in [y, y + H(y)] \\ 1 + y + \mu(F) & \text{if } \mu(F) + H(y) > 1 \text{ and } \mu(F) \in [0, H(y) - y]. \end{cases}$$

One can verify that $u(\hat{F}(X), X) = H(X)$. Therefore, the cdf of U is equal to G, as needed. \square

Theorem 3.3. Let U be defined as in Theorem 2.2, and let ν denote the law of U. Suppose (X,Y) has a density p(x,y) with respect to Lebesgue measure on $\mathcal{X} \times \mathcal{Y}$, and with probability one, $\mathcal{Q}(y,0,\hat{F}(X))$ satisfies the property of monotonicity

$$p(y|x) > p(y'|x)$$
 implies $Q(\hat{F}(X), 0, y) > Q(\hat{F}(X), 0, y')$

and the property of tie-breaking, then μ has a density $\eta(u)$ on [0,1] which is monotonic in u.

Proof. Choose 0 < u < v < 1 and $0 < \delta < \min(u, 1 - v, v - u)$. For $x \in \mathcal{X}$, define the set

$$\underline{J}_x = \{ y \in \mathcal{Y} : \int_{\mathcal{Y}} I(p(y|x) > p(w|x)) p(w|x) dw \in [u - \delta, u + \delta] \}$$

and

$$\bar{J}_x = \{ y \in \mathcal{Y} : \int_{\mathcal{Y}} I(p(y|x) > p(w|x)) p(w|x) dw \in [v - \delta, v + \delta] \}$$

One can verify that for all $x \in \mathcal{X}$,

$$\int_{J_x} p(y|x)dy \le \int_{\bar{J}_x} p(y|x)dy.$$

Yet, since

$$\Pr[U \in [u-\delta, u+\delta]] = \Pr[\cup_{\mathcal{X}} x \times \underline{J}_x]$$

$$\Pr[U \in [v - \delta, v + \delta]] = \Pr[\cup_{\mathcal{X}} x \times \underline{J}_x].$$

we obtain

$$\Pr[U \in [u - \delta, u + \delta]] \le \Pr[U \in [v - \delta, v + \delta]].$$

Taking $\delta \to 0$, we conclude the theorem. \square