

# Upper and lower bounds on cdf of generalized non-central chi-squared

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## 1 Introduction

Let  $Z \sim N(0, I_p)$ , and let  $\mu \in \mathbb{R}^p$  and  $\Sigma$  a positive semidefinite matrix. Define the generalized noncentral chi-squared distribution with noncentrality  $\mu$  and shape  $\Sigma$  as the distribution of

$$Y = (Z + \mu)^T \Sigma (Z + \mu)$$

Let  $V \Lambda V^T = \Sigma$  be the eigendecomposition of  $\Sigma$ , and let  $\eta = V^T \mu$ . Then

$$Y \stackrel{d}{=} (Z + \eta)^T \Lambda (Z + \eta) = \sum_{i=1}^p \lambda_i W_i$$

where  $W_i \sim \chi_1^2(\eta_i^2)$ . Recall that the mgf of the noncentral chi-squared with one df is given by

$$\mathbf{E}[e^{tW_i}] = \frac{\exp[\frac{\eta_i^2 t}{1-2t}]}{\sqrt{1-2t}}$$

It follows that the moment-generating function of  $Y$  is given by

$$\mathbf{E}[e^{tY}] = \prod_{i=1}^p \mathbf{E}[e^{\lambda_i t W_i}] = \prod_{i=1}^p \frac{\exp[\frac{\eta_i^2 \lambda_i t}{1-2t\lambda_i}]}{\sqrt{1-2t\lambda_i}}$$

## 2 Upper bound

We wish to bound the probability  $\Pr[Y < x]$ . We have

$$\begin{aligned}
\log \Pr[Y < x] &= \log \Pr[e^{tY} > e^{tx}] \text{ for } t < 0 \\
&\leq \log \left( \frac{\mathbf{E}[e^{tY}]}{e^{tx}} \right) \\
&= \log(\mathbf{E}[e^{tY}]) - tx \\
&= \left( \frac{-1}{2} \sum_{i=1}^p \log(1 - 2t\lambda_i) \right) + \left( \sum_{i=1}^p \frac{\eta_i^2 \lambda_i t}{1 - 2t\lambda_i} \right) - tx
\end{aligned}$$

Now consider minimizing the bound over  $t$ . The derivative of the bound wrt  $t$  is

$$-x + \sum_{i=1}^p \frac{\lambda_i(1 + \eta_i)^2}{1 - 2\lambda_i t} + \frac{2\eta_i^2 \lambda_i^2 t}{(1 - 2\lambda_i t)^2}$$

which, to a first-order approximation in  $1/t$ , is

$$-\frac{p}{2t} - x$$

Hence this implies that for small  $x$ , the optimal value of  $t$  is approximately

$$t^* \approx -\frac{p}{2x}.$$

Plugging in this value of  $t$  yields the bound

$$\log \Pr[Y < x] < -\frac{1}{2} \left( \sum_{i=1}^p \log \left( 1 + \frac{\lambda_i p}{x} \right) \right) + \frac{p}{2} \left( 1 - \frac{1}{x} \sum_{i=1}^p \frac{\eta_i^2 \lambda_i}{1 + \frac{\lambda_i p}{x}} \right)$$

## 3 Lower bound

From the previous notation, the probability is written as an integral over an ellipse  $\mathcal{E}$ :

$$\Pr[Y < x] = \int_{\mathcal{E}} \phi(\nu - z) dz$$

where

$$\phi(\nu - z) = \left( \frac{1}{2\pi} \right)^{p/2} \exp^{-\|\nu - z\|^2/2}$$

and

$$\mathcal{E} = \{z \in \mathbb{R}^p : z^T \Lambda z < x\}.$$

Define the height of the ellipse as

$$h = \max_{z \in \mathcal{E}} \left( \frac{\eta}{\|\eta\|} \right)^T z = \sqrt{x\rho},$$

where

$$\rho = \frac{\eta^T \Sigma^{-1} \eta}{\|\eta\|^2},$$

and let  $z^*$  be the point in  $\mathcal{E}$  which attains  $z^T \eta = h\|\eta\|$ , whence

$$z^* = \sqrt{\frac{x}{\rho}} \Sigma^{-1} \frac{\eta}{\|\eta\|}.$$

Define the fractional cap with fraction  $u$  by

$$\mathcal{C}_u = \{z \in \mathcal{E} : \eta^T z < uh\|\eta\|\}.$$

Intuitively, this cap contains points  $z \in \mathcal{E}$  with a large density  $\phi(z - \nu)$ . The volume of the cap is given by

$$x^{p/2} \text{Vol}(B_1) / \sqrt{|\Lambda|} J(u)$$

where

$$J_p(u) = \frac{\int_{1-2u}^1 (1-x^2)^{\frac{p-1}{2}} dx}{\int_{-1}^1 (1-x^2)^{\frac{p-1}{2}} dx} = \frac{\Pr[\text{Beta}(\frac{1}{2}, \frac{p+1}{2}) > (1-2u)^2]}{2} \text{ for } u < \frac{1}{2}.$$

Let us proceed to lower bound the density in  $\mathcal{C}_u$ . This amounts to upper bounding the distance

$$\max_{z \in \mathcal{C}_u} \|z - \eta\|^2.$$

Let us define

$$m(u) = (\|\eta\| - h(1-2u))^2 + \left( \|\eta/\|\eta\| - z^*\|^2 + \sqrt{\frac{x}{\lambda_{\min}}(1 - (1-2u)^2)} \right)$$

where  $\lambda_{\min} = \min_{i=1}^p \lambda_i$ . We claim that

$$\max_{z \in \mathcal{C}_u} \|z - \eta\|^2 \leq m(u).$$

To prove the claim, for any  $z \in \mathcal{C}_u$ , decompose the difference into parallel and orthogonal components,

$$\eta - z = \alpha \frac{\eta}{\|\eta\|} + \beta \eta^\perp$$

where  $\eta^\perp$  is orthogonal to  $\eta$ . Hence

$$\|\eta - z\|^2 = \alpha^2 + \beta^2.$$

It suffices to show that

$$\alpha \leq \|\eta\| - h(1 - 2u)$$

and

$$\beta \leq \|\eta/\|\eta\| - z^*\| + \sqrt{\frac{x}{\lambda_{\min}}(1 - (1 - 2u)^2)}$$

for every  $z \in \mathcal{C}_u$ .

Hence,

$$\log \Pr[Y < x] \leq \log(\phi(\sqrt{m(u)})) + \log \text{Vol}(B_{\sqrt{x}}) - \frac{1}{2} \log |\lambda| + \log J_p(u)$$