Stimulus Identification from fMRI scans

Charles Zheng and Yuval Benjamini

Stanford University, Department of Statistics

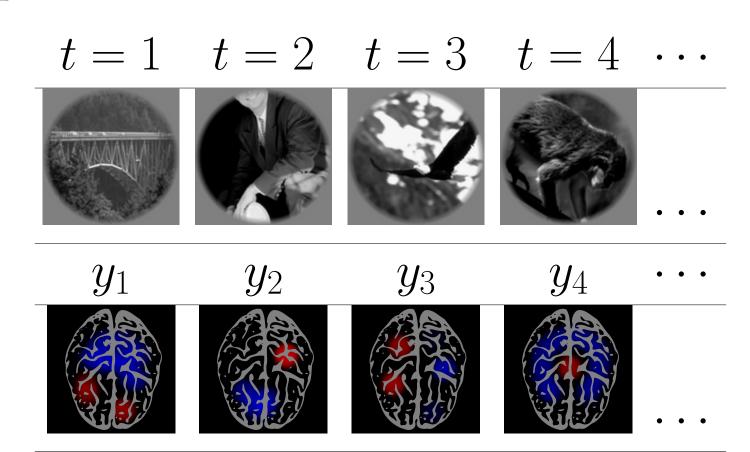
Overview

Seeking to explain the processes behind human perception, scientists employ forward models to model the causal relationship between perception of stimuli and neural activity. But how can we measure the quality of these models? Kay et al (2008) introduced the task of identification as a way to demonstrate the fidelity and generalizability of the model.

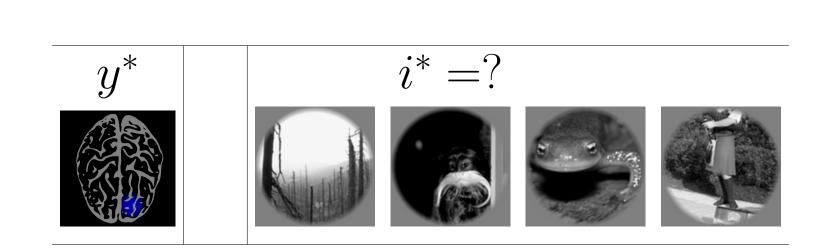
Using the data of Kay et al. as a motivating example, we consider the statistical problem of optimal identification. While identification superficially resembles a classification task (with many classes), it combines the challenge of multivariate regression with high-dimensional discrimination.

Data

- Sequence of stimuli (pictures) shown at time $t=1,\ldots,T=3500$
- Record subject's multivariate response $y_t \in \mathbb{R}^p$, here $p \approx 20000$ voxels



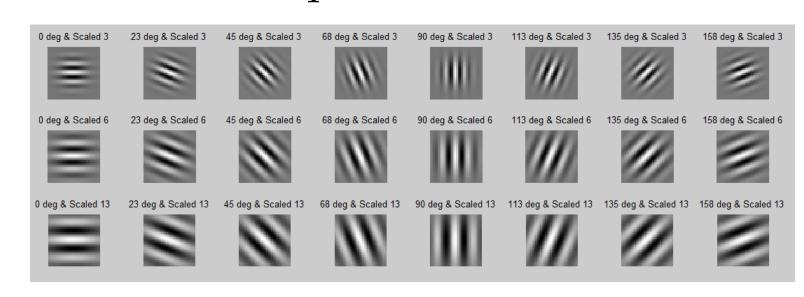
Identification



- Let S be a set of stimuli, possibly outside the training set! |S| can range from 120 to 10000
- Scientist picks a stimulus i^* from S and measures the subject's reponse y^*
- Can the statistician $identify i^* \in S$ from y^* ?
- Objective: minimize misclassification rate

Previous approaches

- In order to generalize to new stimuli, we need to find some quantitative representation
- Kay (2008) uses Gabor filters to describe each picture in terms of q=10000 real-valued features



- $Y_{T \times p}$ matrix containing the T of recorded responses
- $X_{T\times q}$ is the matrix of the *image features* of the corresponding stimuli

Now consider a parametric model

$$Y \sim F_{\theta}(X)$$

Such a *forward model* gives the distribution of the response conditional on the stimuli features; while identification requires the converse.

However, the maximum likelihood (ML) principle can be invoked to identify the stimuli $i \in S$ "most likely" to have produced y^* .

Let $x_i : i \in S$ denote features of the test stimuli, and identify y^* based on the maximum likelihood (ML) principle

$$i^* = \operatorname{argmax}_i \ell_{\theta}(y^* | x_i)$$

Example. We take the following as a representative approach, combining features of [1] and [2]:

- Assume the normal mutivariate linear model

$$Y \sim N(XB, \Sigma_E)$$
, where $B \in \mathbb{R}^{q \times p}$

- Obtain point estimates of B using elastic net [4], and Σ_E using off-diagonal shrinkage
- The ML rule takes the form

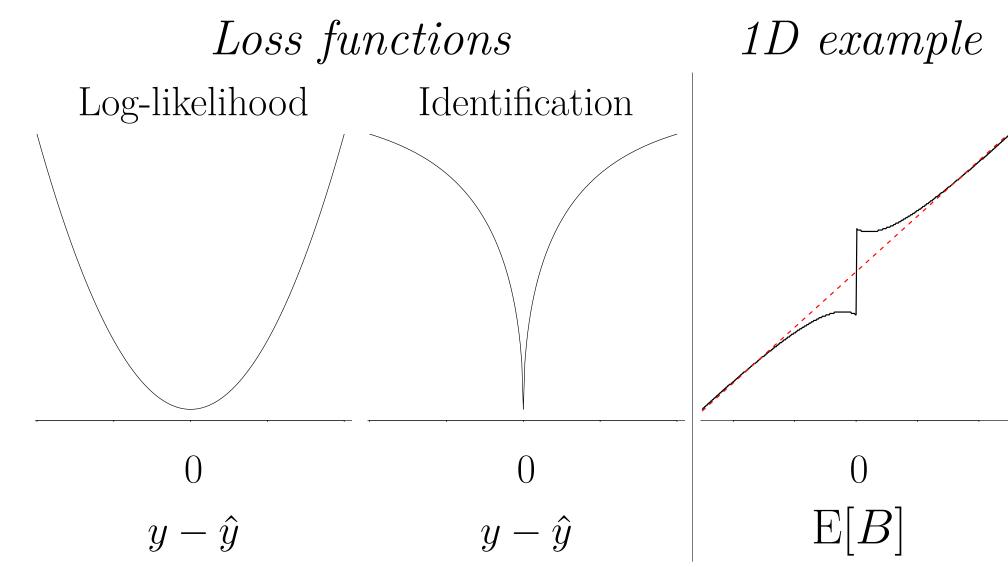
$$i^* = \operatorname{argmin}_i(x_i^T \hat{B} - y^*)^T \hat{\Sigma}_E^{-1}(x_i^T \hat{B} - y^*)$$
 (1)

Initial Questions

- Consider the Gaussian model for now...
- Is ML optimal for identification?
- If not, are there near-optimal rules of the form (1)? Or do we need to consider a richer class of decision rules?

Theory

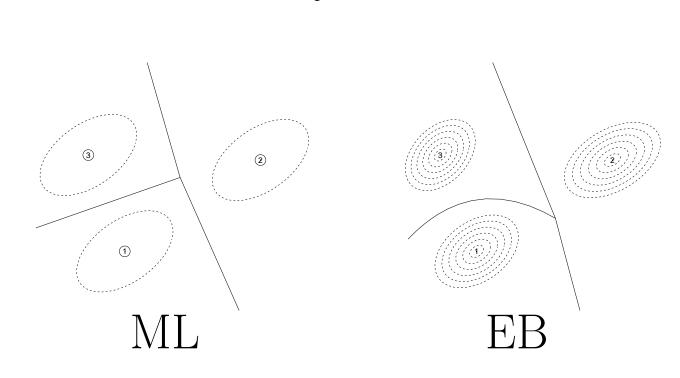
- ML is consistent given the correct model, but can be rather poor in finite samples
- \hat{B} minimizes log-likelihood—but that isn't the loss function for identification!
- We can compute optimal \hat{B} for p = q = 1...
- Optimal rule of form (1) intractable in higher dimensions due to nonconvexity



Right: A p=q=1-dimensional example where ML (actually MAP) fails. The Bayes estimate for identification (black) and the MAP estimate (red) diverge sharply when $B \sim N(0,1)$. The same phenomenon can be found in higher dimensions.

Empirical Bayes

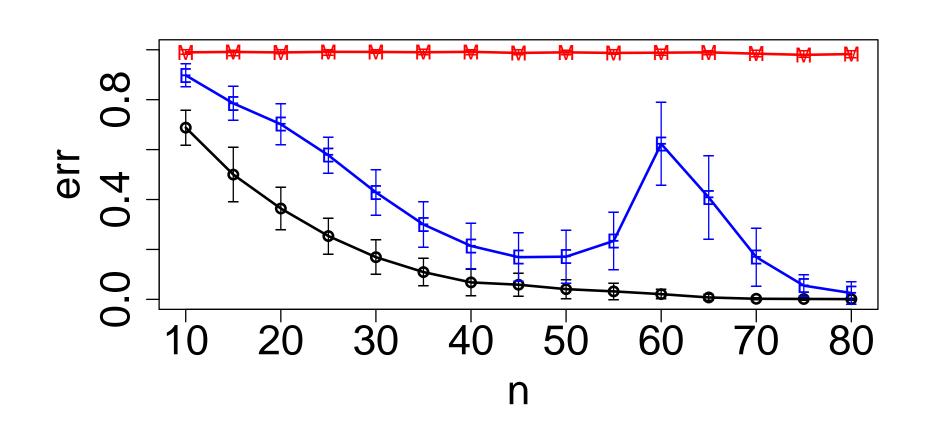
- *Idea*: Unlike ML, the Bayes rule surely optimizes the "correct" objective function. Can we approximate the Bayes rule?
- Empirical Bayes: use the data to estimate the covariances Σ_B and Σ_E , then compute posterior distribution of B
- Assume coefficients of B independent; diagonals of Σ_B can be estimated using any estimate of signal strength, e.g. Eigenprism [3].
- Decision rule is same as (1) but with "added noise" due to uncertainty of B. $\min(x_i^T B y^*)^T (\operatorname{Cov}(x_i^T B) + \hat{\Sigma}_E)^{-1} (x_i^T B y^*)$
- Analogous to LDA vs QDA



• Computation: requires inverting $pq \times pq$ matrix

Simulation Results

- Parameters p=q=60, random B and Σ_E , number of classes |S|=100
- Empirical bayes *does* outperform ML given small sample sizes... however...



(E) Empirical Bayes, (M) Maximum likelihood, (o) Bayes risk (knowing true Σ_B , Σ_E)

Ongoing Work

- Why does error *increase* with sample size!?
- Clearly, we need refine the EB procedure
- Required cost of $O((pq)^3)$ hinders application to real data... develop tractable approximations

Conclusions

- We studied optimal identification under the multivariate linear model
- "LDA"-like rules of the form (1) (such as ML) are too restrictive. But "QDA"-like rules include the optimal Bayes rule!
- We propose EB to approximate the Bayes rule—but we need better theory to do it correctly

References

- [1] Kay et al. *Nature* (2008)
- [2] Vu et al. Annals of Applied Statistics (2011)
- [3] Janson et al. (2015) http://arxiv.org/abs/1505.02097
- [4] Zou et al. J. R. Statist. Soc. B (2005)

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