

The Normal Approximation & Binomial Proportion
Confidence Interval

Central Limit Theorem

Suppose $\{X_i\}_{i=1}^n$ are any collection of iid random variables with mean $\mu = \mathbb{E} X_i < \infty$ and standard deviation $\sigma = \text{sd}(X_i) < \infty$.

Then the sum $S_n = \sum_{i=1}^n X_i$ and mean $\frac{S_n}{n}$

Converge in distribution to the normal distribution,

$$\frac{S_n - n\mu}{\sigma\sqrt{n}} \xrightarrow{d} N(0,1) \quad \text{as } n \rightarrow \infty$$

$$\frac{\frac{S_n}{n} - \mu}{\sigma/\sqrt{n}} \xrightarrow{d} N(0,1) \quad \text{as } n \rightarrow \infty$$

Convergence in distribution \xrightarrow{d} means

$$P(a \leq \frac{S_n - n\mu}{\sigma\sqrt{n}} \leq b) \rightarrow P(a \leq Z \leq b) \quad \text{as } n \rightarrow \infty.$$

Tons of quantities in sports are the sum or mean of iid Random Variables! So the normal approximation comes in handy!

Estimating the quality of a free throw shooter

Shaq shoots n free throws, whose results are given by $\{X_i\}_{i=1}^n$, $X_i = \begin{cases} 1 & \text{if } i\text{th free} \\ & \text{throw made} \\ 0 & \text{if missed} \end{cases}$

$S_n = \sum_{i=1}^n X_i$ is his # made free throws.

We model S_n by

$$S_n \sim \text{Binomial}(n, p) = \begin{aligned} &\# \text{ successes in } n \text{ trials} \\ &\text{where each trial is} \\ &\text{independent Bernoulli}(p) \\ &(X_i \text{ iid } \text{Ber}(p)) \end{aligned}$$

We want to estimate Shaq's P from the data $\{X_i\}_{i=1}^n$. Our "best guess" of P is $\hat{P} = \frac{S_n}{n} = \frac{1}{n} \sum_{i=1}^n X_i$. In fact, this is the **MLE** (See the Lab).

But how confident should we be in this guess?

Let $\mu = \mathbb{E}X_i = p$, $\sigma^2 = \text{Var}(X_i) = \sqrt{p(1-p)}$.

By CLT, $\frac{\hat{P}-\mu}{\sigma/\sqrt{n}} \xrightarrow{d} N(0,1)$ as $n \rightarrow \infty$.

So $P(-2 \leq \frac{\hat{P}-\mu}{\sigma/\sqrt{n}} \leq 2) = P(-2 \leq \frac{\hat{P}-\hat{P}}{\sqrt{p(1-p)/n}} \leq 2) \approx 0.95$

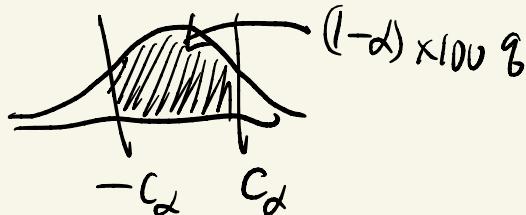
So $p \in \hat{P} \pm 2 \sqrt{p(1-p)/n}$ w.p. ≈ 0.95

p is unknown, so use \hat{P} in place of p

Wald CI: $\hat{P} \pm 2 \sqrt{\frac{\hat{P}(1-\hat{P})}{n}}$

More generally: $\hat{P} \pm c_\alpha \sqrt{\frac{\hat{P}(1-\hat{P})}{n}}$

is a $(1-\alpha) \times 100\%$ CI where c_α is the quantile of $N(0,1)$ taking $(1-\alpha) \times 100\%$ of



You buy an M&M bag with 56 M&Ms and 14 blue ones. Supposing the color of each M&M is randomly drawn from some iid distribution, what is a 95% confidence interval of the probability the company makes an M&M blue?

$$\hat{P} = \frac{14}{56} = \frac{1}{4}$$

$$\hat{P} \pm 2 \sqrt{\frac{\hat{P}(1-\hat{P})}{n}} = \frac{1}{4} \pm 2 \sqrt{\frac{\frac{1}{4} \cdot \frac{3}{4}}{56}} = \frac{1}{4} \pm \frac{\sqrt{3}}{2} \cdot \frac{1}{\sqrt{n}}$$
 $\approx .85$

$$n=56 \Rightarrow CI \approx \frac{1}{4} \pm \frac{1}{2} \quad \text{pretty wide!}$$

$$n=400 \Rightarrow CI \approx \frac{1}{4} \pm .0425$$

The Meaning of a confidence interval

The confidence interval of a Binomial proportion p is

$$\text{Wald CI} = \hat{P} \pm 2 \sqrt{\frac{\hat{P}(1-\hat{P})}{n}}$$

i.e. $P(p \in CI) \approx 0.95$

but what does this probability P actually mean?

Under this model, $S_n = \sum_{i=1}^n X_i \sim \text{Binomial}(n, p)$

$X_i \stackrel{iid}{\sim} \text{Bernoulli}(p)$ (Frequentist)

Under this model, p is an unknown fixed constant.

p either lies in the CI or doesn't lie in it,
so in actuality the probability it lies in the CI
is either 0 or 1. So, what does P mean?

The probability P refers to Randomness in the CI:

The CI itself is a random variable
because \hat{P} is a random variable because
 $\{X_i\}_{i=1}^n$ are random variables.

If we replicated the experiment 100 times,
in each replication p remains the same
but the data $\{X_i\}$ and the CI's change
by Randomness.

→ on average
the CI is expected to contain the true
unobserved p in 95 of these 100 replications

Coverage

The Wald CI is based on 2 approximations:

that $P\left(-2 \leq \frac{\hat{p}-p}{\sqrt{p(1-p)/n}} \leq 2\right) \approx 0.95$ by CLT

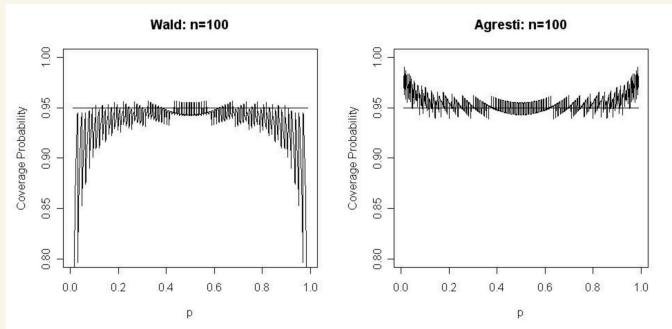
and that we can plug in \hat{p} for p in $\sqrt{p(1-p)}$

so that $P\left(-2 \leq \frac{\hat{p}-p}{\sqrt{\hat{p}(1-\hat{p})/n}} \leq 2\right) \approx 0.95.$

The **actual** coverage probability $P(p \in CI)$ of the 95% Wald interval can be quite far below the **nominal** coverage of 95% as shown by simulations and computations (Brown, Cai, Dasgupta 2001).

If n is several hundred or thousand the Wald interval is tolerably accurate. Otherwise we need to adjust the CI.

Agresti and Coull (1998) recommend introducing **2 artificial successes and failures** into the data before computing \hat{p} and n , which is better than the Wald interval.



→ oscillations from the discreteness of Binomial

$$\text{Agresti-Coull CI} = \hat{p}' \pm 2 \sqrt{\frac{\hat{p}'(1-\hat{p}')}{n+4}}$$

$$\text{where } \hat{p}' = \frac{s_n + 2}{n+4}$$