## **MPEC** notes

Chris Conlon thanks to Che-Lin Su October 27, 2020

Grad IO

## **Unconstrained Optimization**

Basic idea in many estimation problems is to use Newton-type methods to solve FOCs of equilibrium or estimating equations

$$\min f(x) : x \in \mathbb{R}^n$$

- $f: R^n \to R, c: R^n \to R^m$  smooth (typically  $C^2$ )
- $x \in \mathbb{R}^n$  finite dimensional (may be large)

Want to find a local minimizer

$$\nabla f(x^*) = 0$$

Optimization Algorithms generate a sequence  $\boldsymbol{x}^{(k)}$  such that the gradient test

$$||\nabla f(x^{(k)})|| \leq \frac{\tau}{\tau}$$

is satisfied for some tolerance  $\tau = 1e - 6$  or so. Warning!

#### General NLP Problem

A Nonlinear Programming (NLP) problem is defined by:

$$\begin{cases} \min_{x} f(x) & \text{objective} \\ \text{subject to} & c(x) = 0 & \text{constraints} \\ x \ge 0 & \text{variables} \end{cases}$$

#### Typical assumptions

- $f: R^n \to R, c: R^n \to R^m$  smooth (typically  $C^2$ )
- $x \in R^n$  finite dimensional (perhaps large)
- more general  $l \le c(x) \le u$  is possible

## **Optimality Conditions for NLP**

#### Constraint Qualifications (CQ)

Linearization of c(x) = 0 characterizes all feasible permutations,  $x^*$  local minimizer & CQ holds  $\exists$  multipliers  $\lambda^*, \gamma^*$ :

$$\nabla f(x^*) - \nabla c(x^*)\lambda^* - \gamma^* = 0$$

$$c(x^*) = 0$$

$$x^*\lambda^* = 0$$

$$x^* \ge 0, \gamma^* \ge 0$$

Where  $X^* = diag(x^*)$ , thus  $X^*\lambda^* = 0 \Leftrightarrow x_i^*\gamma_i^* = 0$ 

#### **NLP Solvers**

F(w)=0 where  $w=(x,\lambda,\gamma)$  with  $x\geq 0, z\geq 0$ . Optimization Algorithms generate a sequence  $w^{(k)}$  such that the gradient test

$$||\nabla f(w^{(k)})|| \leq \frac{\tau}{T}$$

is satisfied for some tolerance  $\tau = 1e - 6$  or so. (Same warning).

## Optimization

## "Folk Theory" of Optimization in Economics

- Unconstrained Optimization is easier than Constrained Optimization
- More parameters are harder
- Quasi-Newton Methods are unreliable

#### Consequences of Folk Theory

- Rewrite all problems as unconstrained optimization
- Use fixed points and multi-step procedures to reduce parameter space
- Use Nelder-Mead/Simplex methods for optimization

## Optimization

Thanks to recent advances in optimization:

### More Accurate Description of Optimizaiton

- 1. Shape of the problem is what matters convexity is really important
- 2. Constrained Problems are not much more difficult
- 3. More parameters can make the problem easier (or harder)

### Consequences of State of the Art Optimization

- Tested stable Newton-routines are very reliable.
- Good Solvers handle 10,000+ parameters
- Computational burden are Jacobian and Hessian (and storage)

## Recent Advances in Optimization Literature

#### Large Scale Algorithms

- Much focus has been on very large convex optimization problems these have gotten really good.
- Most of these rely on first and second derivatives and quadratic approximations.
- Ways to do derivatives: analytic, numeric, symbolic and automatic (new!)
- Easy to solve 10,000+ parameter constrained problems often in less than 20 major iterations.
- Lots of industrial strength software packages.
- When in doubt express your problem as a convex one.
- Algorithm is polynomial  $\approx O(k^3)$

# Convexity

## Convexity

## An optimization problem is convex if

$$\min_{x} f(x) \quad s.t. \quad h(x) \le 0 \quad Ax = 0$$

- f(x), h(x) are convex (PSD second derivative matrix)
- Equality Constraint is affine

## Some helpful identities about convexity

- Compositions and sums of convex functions are convex.
- $\bullet$  Norms || are convex, max is convex, log is convex
- $\log(\sum_{i=1}^n \exp(x_i))$  is convex.
- Fixed Points can introduce non-convexities.
- Globally convex problems have a unique optimum

## **Nested Logit Model**

#### FIML Nested Logit Model is Non-Convex

$$\min_{\theta} \sum_{j} q_{j} \ln P_{j}(\theta) \quad \text{s.t.} \quad P_{j}(\theta) = \frac{e^{x_{j}\beta/\lambda} (\sum_{k \in g_{j}} e^{x_{j}\beta/\lambda})^{\lambda-1}}{\sum_{\forall l'} (\sum_{k \in g'_{l}} e^{x_{j}\beta/\lambda})^{\lambda}}$$

This is a pain to show but the problem is with the cross term  $\frac{\partial^2 P_j}{\partial \beta \partial \lambda}$  because  $\exp[x_j \beta/\lambda]$  is not convex.

A Simple Substitution Saves the Day: let  $\gamma = \beta/\lambda$ 

$$\min_{\theta} \sum_{j} q_{j} \ln P_{j}(\theta) \quad \text{s.t.} \quad P_{j}(\theta) = \frac{e^{x_{j}\gamma} (\sum_{k \in g_{j}} e^{x_{j}\gamma})^{\lambda - 1}}{\sum_{\forall l'} (\sum_{k \in g_{j}'} e^{x_{j}\gamma})^{\lambda}}$$

This is much better behaved and easier to optimize.

# Nested Logit Model (Conlon Mortimer AEJ 2013)

	$Original^1$	Substitution <sup>2</sup>	No Derivatives <sup>3</sup>
Parameters	49	49	49
Nonlinear $\lambda$	5	5	5
Neg LL	2.279448	2.279448	2.27972
Iterations	197	146	352
Time	59.0 s	10.7 s	192s

Discuss Simplex, Sparsity.

# MPEC

#### **Extremum Estimators**

Often faced with extremum estimator problems in econometrics (ML, GMM, MD, etc.) that look like:

$$\hat{\theta} = \arg \max_{\theta} Q_n(\theta), \quad \theta \in \Theta$$
 (1)

Many economic problems contain constraints, such as: market clearing (supply equals demand), consumer's consume their entire budget set, or firm's first order conditions are satisfied. A natural way to represent these problems is as constrained optimization.

#### **Constrained Problems**

#### **MPEC**

$$\hat{ heta} = rg \max_{ heta,P} Q_n( heta,P), \quad ext{ s.t. } \quad \Psi(P, heta) = 0, \quad heta \in \Theta$$

Fixed Point / Implicit Solution

in much of the literature the tradition has been to express the solutions  $\Psi(P, heta) = 0$  implicitly as P( heta)

$$\hat{ heta} = rg \max_{ heta} \mathcal{Q}_{ extsf{n}}( heta, P( heta)), \quad heta \in \Theta$$

#### **Constrained Problems**

#### **MPEC**

$$\hat{ heta} = rg \max_{ heta,P} Q_n( heta,P), \quad ext{ s.t. } \quad \Psi(P, heta) = 0, \quad heta \in \Theta$$

#### Fixed Point / Implicit Solution

In much of the literature the tradition has been to express the solutions  $\Psi(P,\theta)=0$  implicitly as  $P(\theta)$ :

$$\hat{\theta} = \arg\max_{\theta} Q_n(\theta, P(\theta)), \quad \theta \in \Theta$$

#### Rust Problem

- ullet Bus repairman sees mileage  $x_t$  at time t since last overhaul
- Repairman chooses between overhaul and normal maintenance

$$u(x_t, d_t, \theta^c, RC) = egin{cases} -c(x_t, \theta^c) & ext{if} & d_t = 0 \\ -(RC + c(0, \theta^c)) & ext{if} & d_t = 1 \end{cases}$$

Repairman solves DP:

$$V_{\theta}(x_t) = \sum_{f_t, f_{t+1}, \dots} E\left\{\sum_{j=t}^{\infty} \beta^{j-t} [u(x_j, f_j, \theta) + \varepsilon_j(f_j)] | x_t\right\}$$

- Econometrician
  - Observes mileage  $x_t$  and decision  $d_t$  but not cost.
  - Assumes extreme value distribution for  $\varepsilon_t(d_t)$
- Structural parameters to be estimated  $\theta = (\theta^c, RC, \theta^p)$ .
  - Coefficients of cost function  $c(x, \theta^c) = \theta_1^c x + \theta_2^c x^2$
  - Overhaul cost RC

#### Rust Problem

- Data: time series  $(x_t, d_t)_{t=1}^T$
- Likelihood function

$$\mathcal{L}(\theta) = \prod_{t=2}^{T} P(d_t|x_t, \theta^c, RC) p(x_t|x_{t-1}, d_{t-1}, \theta^p)$$
with  $P(d|x, \theta^c, RC) = \frac{\exp[u(x, d, \theta^c, RC) + \beta EV_{\theta}(x, d)}{\sum_{d' \in \{0, 1\}} \exp[u(x, d', \theta^c, RC) + \beta EV_{\theta}(x', d)]}$ 

$$EV_{\theta}(x, d) = T_{\theta}(EV_{\theta})(x, d)$$

$$\equiv \int_{x'=0}^{\infty} \log \left[ \sum_{d' \in \{0, 1\}} \exp[u(x, d', \theta^c, RC) + \beta EV_{\theta}(x', d)] \right] p(dx'|x, d, \theta^p)$$

#### Rust Problem

• Outer Loop: Solve Likelihood

$$\max_{\theta \geq 0} \mathcal{L}(\theta) = \prod_{t=2}^{T} P(d_t|x_t, \theta^c, RC) p(x_t|x_{t-1}, d_{t-1}, \theta^p)$$

- Convergence test:  $\|\nabla_{\theta} \mathcal{L}(\theta)\| \leq \epsilon_{out}$
- Inner Loop: Compute expected value function  $EV_{\theta}$  for a given  $\theta$
- $EV_{\theta}$  is the implicit expected value function defined by the Bellman equation or the fixed point function

$$EV_{\theta} = T_{\theta}(EV_{\theta})$$

- Convergence test:  $\| EV_{\theta}^{(k+1)} EV_{\theta}^{(k)} \| \le \epsilon_{in}$
- Start with contraction iterations and polish with Newton Steps

#### **NFXP Concerns**

- Inner-loop error propagates into outer-loop function and derivatives
- NFXP needs to solve inner-loop exactly each stage of parameter search
  - to accurately compute the search direction for the outer loop
  - to accurately evaluate derivatives for the outer loop
  - for outer loop to converge!
- Stopping rules: choosing inner-loop and outer-loop tolerance
  - inner loop can be slow: contraction mapping is linearly convergent
  - tempting to loosen inner loop tolerance  $\epsilon_{in}$  (such as 1e-6 or larger!).
  - Outer loop may not converge with loose inner loop tolerance.
    - check solver output message
    - tempting to loosen outer loop tolerance  $\epsilon_{out}$  to promote convergence (1e-3 or larger!).

## **Stopping Rules**

- $\mathcal{L}(EV(\theta, \epsilon_{in}), \theta)$  the programmed outer loop objective function
- L: the Lipschitz constant (like modulus) of inner-loop contraction mapping
- Analytic derivatives  $\nabla_{\theta} \mathcal{L}(EV(\theta, \epsilon_{in}), \theta)$  is provided:  $\epsilon_{out} = O(\frac{L}{1-L}\epsilon_{in})$
- Finite-difference derivatives are used:  $\epsilon_{out} = O(\sqrt{\frac{L}{1-L}}\epsilon_{in})$

## Stopping Rules

• Form the augmented likelihood function for data  $X = (x_t, d_t)_{t=1}^T$ 

$$\mathcal{L}(EV, \theta; X) = \prod_{t=2}^{T} P(d_t | x_t, \theta^c, RC) p(x_t | x_{t-1}, d_{t-1}, \theta^p)$$
with  $P(d | x, \theta^c, RC) = \frac{\exp[u(x, d, \theta^c, RC) + \beta EV(x, d)}{\sum_{d' \in \{0,1\}} \exp[u(x, d', \theta^c, RC) + \beta EV(x', d)}$ 

ullet Rationality and Bellman equation imposes a relationship between heta and  $extit{EV}$ 

$$EV = T(EV, \theta)$$

Solve constrained optimization problem

$$\max_{(\theta, EV)} \mathcal{L}(EV, \frac{\theta}{\theta}; X)$$
 subject to  $EV = T(EV, \frac{\theta}{\theta})$ 

β	Imple.			Para	meters			MSE
		RC	$\theta_{11}$	$\theta_{31}$	$\theta_{32}$	$\theta_{33}$	$\theta_{34}$	
	true	11.726	2.457	0.0937	0.4475	0.4459	0.0127	
0.975	MPEC1	12.212 (1.613)	2.607 (0.500)	0.0943 (0.0036)	0.4473 (0.0057)	0.4454 (0.0060)	0.0127 (0.0015)	3.111
	MPEC2	12.212 (1.613)	2.607 (0.500)	0.0943 (0.0036)	0.4473 (0.0057)	0.4454 (0.0060)	0.0127 (0.0015)	3.111
	NFXP	12.213 (1.617)	2.606 (0.500)	0.0943 (0.0036)	0.4473 (0.0057)	0.4445 (0.0060)	0.0127 (0.0015)	3.123
0.980	MPEC1	12.134 (1.570)	2.578 (0.458)	0.0943 (0.0037)	0.4473 (0.0057)	0.4455 (0.0060)	0.0127 (0.0015)	2.857
	MPEC2	12.134 (1.570)	2.578 (0.458)	0.0943 (0.0037)	0.4473 (0.0057)	0.4455 (0.0060)	0.0127 (0.0015)	2.857
	NFXP	12.139 (1.571)	2.579 (0.459)	0.0943 (0.0037)	0.4473 (0.0057)	0.4455 (0.0060)	0.0127 (0.0015)	2.866 -

β	Imple.			Para	meters			MSE
		RC	$\theta_{11}$	$\theta_{31}$	$\theta_{32}$	$\theta_{33}$	$\theta_{34}$	
	true	11.726	2.457	0.0937	0.4475	0.4459	0.0127	
0.985	MPEC1	12.013	2.541	0.0943	0.4473	0.4455	0.0127	2.140
		(1.371)	(0.413)	(0.0037)	(0.0057)		(0.0015)	-
	MPEC2	12.013	2.541	0.0943	0.4473	0.4455	0.0127	2.140
		(1.371)	(0.413)	(0.0037)	(0.0057)		(0.0015)	-
	NFXP	12.021	2.544	0.0943	0.4473	0.4455	0.0127	2.136
		(1.368)	(0.411)	(0.0037)	(0.0057)		(0.0015)	-
0.990	MPEC1	11.830	2.486	0.0943	0.4473	0.4455	0.0127	1.880
		(1.305)	(0.407)	(0.0036)	(0.0057)		(0.0015)	-
	MPEC2	11.830	2.486	0.0943	0.4473	0.4455	0.0127	1.880
		(1.305)	(0.407)	(0.0036)	(0.0057)		(0.0015)	_
	NFXP	11.830	2.486	0.0943	0.4473	0.4455	0.0127	1.880
		(1.305)	(0.407)	(0.0036)	(0.0057)		(0.0015)	-

β	Imple.			Para	meters			MSE
		RC	$\theta_{11}$	$\theta_{31}$	$\theta_{32}$	$\theta_{33}$	$\theta_{34}$	
	true	11.726	2.457	0.0937	0.4475	0.4459	0.0127	
0.995	MPEC1	11.819 (1.308)	2.492 (0.414)	0.0942 (0.0036)	0.4473 (0.0057)	0.4455 (0.0060)	0.0127 (0.0015)	1.892 -
	MPEC2	11.819 (1.308)	2.492 (0.414)	0.0942 (0.0036)	0.4473 (0.0057)	0.4455 (0.0060)	0.0127 (0.0015)	1.892 -
	NFXP	11.819 (1.308)	2.492 (0.414)	0.0942 (0.0036)	0.4473 (0.0057)	0.4455 (0.0060)	0.0127 (0.0015)	1.892 -

	Imple.	Runs Conv.	CPU Time	# of Major Iter.	# of Func. Eval.	# of Contrac. Mapping Iter.
0.975	MPEC1 MPEC2	1240 1247	0.13 7.9	12.8 53.0	17.6 62.0	
	NFXP MPEC1	998 1236	24.6 0.15	55.9 14.5	189.4 21.8	1.348e + 5
	MPEC2 NFXP	1241 1000	8.1 27.9	57.4 55.0	70.6 183.8	-1.625e + 5
0.985	MPEC1 MPEC2 NFXP	1235 1250 952	0.13 7.5 42.2	13.2 55.0 61.7	19.7 62.3 227.3	- $ 2.658e + 5$
0.990	MPEC1 MPEC2 NFXP	1161 1248 935	0.19 7.5 70.1	18.3 56.5 66.9	42.2 65.8 253.8	- 4.524 $e + 5$
0.995	MPEC1 MPEC2 NFXP	965 1246 950	0.14 7.9 111.6	13.4 59.6 58.8	21.3 70.7 214.7	7.485e + 5

```
# of iterations
# of CG iterations
# of function evaluations = 13
# of gradient evaluations = 13
# of Hessian evaluations
```

## **BLP Demand Example**

#### **BLP 1995**

The estimator solves the following mathematical program:

$$\min_{\theta_2} \qquad g(\xi(\theta_2))'Wg(\xi(\theta_2)) \quad \text{s.t.}$$

$$g(\xi(\theta_2)) = \frac{1}{N} \sum_{\forall j,t} \xi_{jt}(\theta_2)'z_{jt}$$

$$\xi_{jt}(\theta_2) = \delta_j(\theta_2) - x_{jt}\beta - \alpha p_{jt}$$

$$s_{jt}(\delta(\theta_2), \theta_2) = \int \frac{\exp[\delta_j(\theta_2) + \mu_{ij}]}{1 + \sum_k \exp[\delta_j(\theta_2) + \mu_{ik}]} f(\mu|\theta_2)$$

$$\log(S_{jt}) = \log(s_{jt}(\delta(\theta_2), \theta_2)) \quad \forall j, t$$

## **BLP Algorithm**

The estimation algorithm is generally as follows:

- 1. Guess a value of nonlinear parameters  $\theta_2$
- 2. Compute  $s_{jt}(\delta, \theta_2)$  via integration
- 3. Iterate on  $\delta_{jt}^{h+1} = \delta_{jt}^h + \log(S_{jt}) \log(s_{jt}(\delta^h, \theta_2))$  to find the  $\delta$  that satisfies the share equation
- 4. IV Regression  $\delta$  on observable X and instruments Z to get residual  $\xi$ .
- 5. Use  $\xi$  to construct  $g(\xi(\theta_2))$ .
- 6. Possibly construct other errors/instruments from supply side.
- 7. Construct GMM Objective

The idea is that  $\delta(\theta_2)$  is an implicit function of the nonlinear parameters  $\theta_2$ . And for each guess we find that implicit solution for reduce the parameter space of the problem.

#### Dube Fox Su 2009

#### **BLP-MPEC**

The estimator solves the following mathematical program:

$$\min_{\substack{\sigma,\alpha,\beta,\xi \\ \sigma,\alpha,\beta,\xi}} g(\xi)'Wg(\xi) \quad \text{s.t.}$$

$$g(\xi) = \frac{1}{N} \sum_{\forall j,t} \xi'_{jt} z_{jt}$$

$$s_{jt}(\sigma,\alpha,\beta,\xi) = \sum_{i} w_{i} \frac{\exp[x_{jt}\beta + \xi_{jt} - \alpha p_{jt} + \sum_{l} \nu_{il} x'_{jt} \sigma_{l}]}{1 + \sum_{k} \exp[x_{kt}\beta + \xi_{kt} - \alpha p_{kt} + \sum_{l} \nu_{il} x'_{kt} \sigma_{l}]}$$

$$\log(S_{jt}) = \log s_{jt}(\sigma,\alpha,\beta,\xi) \quad \forall j,t$$

- Expand the parameter space of the nonlinear search to include  $\alpha, \beta, \xi$
- ullet Don't have to solve for  $\xi$  except at the end.
- No implicit functions of  $\theta_2$
- Sparsity!

**Empirical Likelihood** 

## **Empirical Likelihood Methods**

## Empirical Likelihood often statistically better than GMM

- Higher order efficiency (gets to semi-parametric efficiency bound faster) (Kitamura 2001, 2006)
- No problems in estimating the weight matrix (Altonji Segal 1995, Newey Smith 2004).
- Likelihood based units for testing
- No problems with scaling of parameters / instruments.
- GMM prone to non-identification in finite-sample (Dominguez Lobato 2004)

## EL as NPMLE (Owen 1990, Kitamura 2006)

## Nonparametric MLE

$$I_{NP}(p_1,\ldots,p_n)=\sum_{i=1}^n\log p_i\quad (p_1,\ldots,p_n)\in\Delta$$
 (2)

- ullet Observed  $z_i$  are IID with measure  $\mu$
- Simplex defined as  $(p_1,p_2,\ldots,p_n)\in\Delta$  such that  $\sum_{i=1}^n p_i=1$  and  $p_i\geq 0$
- Trivial max at  $p_i = \frac{1}{n}$  and  $\mu_n = \frac{1}{n} \sum_{i=1}^n \delta_{z_i}$  and  $I_{NP} = -n \log n$

## Adding moment conditions

(Owen 1990) Shows we can extend NPMLE to moment condition models.

$$E[g(z_i, \theta)] = \int g(z, \theta) d\mu = 0 \in \mathbb{R}^q, \quad \theta \in \Theta \in \mathbb{R}^k$$

This exactly follows our definition of an MPEC problem:

#### **Empirical Likelihood Estimator**

$$arg \max_{ heta, p} I_{NP}$$
  $I_{NP} = \sum_{i=1}^n \log p_i$  s.t.  $\sum_{i=1}^n p_i g(z_i, heta) = 0$  and  $\sum_{i=1}^n p_i = 1$ 

#### Alternative-Generalized Minimum Contrast

Idea is to look for a statistical model  ${\mathcal P}$  close to the true measure  $\mu$ 

$$\mathcal{P}(\theta) = \{P \in M : \int g(z,\theta)dP = 0\}$$
  
 $\mathcal{P} = \cup_{\theta \in \Theta} \mathcal{P}(\theta)$ 

We do this by minimizing the contrast function  $D(P,U)=\int \phi(p)d\mu$  with  $p=\frac{dP}{d\mu}$ .

$$\inf_{\theta \in \Theta} \rho(\theta, \mu)$$
 where  $\rho(\theta, \mu) = \inf_{P \in \mathcal{P}} D(P, \mu)$ 

Produces (infinite dimensional) constrained problem:

$$v(\theta) = \inf_{p} \int \phi(p) d\mu \quad \int g(z,\theta) p d\mu = 0 \quad \int p d\mu = 1$$

Choose a contrast function  $\phi(x) = \log(x)$  (EL) or  $\phi(x) = \frac{1}{2}(x^2 - 1)$  (CUE)

## Interpreting EL

#### Simple Alternate Explanation

- GMM asks what are the parameters θ that minimize the quadratic distance under some metric A
  between my model applied to the observed data, and my model in the "ideal" case. (Model
  generally doesn't hold exactly overidentified).
- EL asks how different a distribution of data would I need to observe in order to meet the implications of my model.
- EL as solutions to systems of equations (which one do we pick?)
- Now I have "likelihood" of different models and can compare structural assumptions.

## Standard Algorithm

### Lagrangian

Construct the dual by differentiating the Lagrangian:

$$\mathcal{L} = \sum_{i=1}^{n} \log p_{i} + \lambda (1 - \sum_{i=1}^{n} p_{i}) - n\gamma' \sum_{i=1}^{n} p_{i}g(z_{i}, \theta)$$

$$\hat{\gamma}(\theta) = \arg \min_{\gamma \in \mathbb{R}^{q}} - \sum_{i=1}^{n} \log(1 + \gamma' g(z_{i}, \theta))$$

$$\hat{p}_{i}(\theta) = \frac{1}{n(1 + \hat{\gamma}(\theta)' g(z_{i}, \theta))} \hat{\lambda} = n$$

$$\hat{\theta}_{EL} = \arg \max_{\theta \in \Theta} I_{NP}(\theta) = \arg \max_{\theta \in \Theta} \min_{\gamma \in \mathbb{R}^{q}} - \sum_{i=1}^{n} \log(1 + \gamma' g(z_{i}, \theta))$$

## Computational Challenges

$$\hat{ heta}_{\mathit{EL}} = rg\max_{ heta \in \Theta} extstyle I_{NP}( heta) = rg\max_{ heta \in \Theta} \min_{\gamma \in \mathbb{R}^q} - \sum_{i=1}^n \log(1 + \gamma' g(z_i, heta))$$

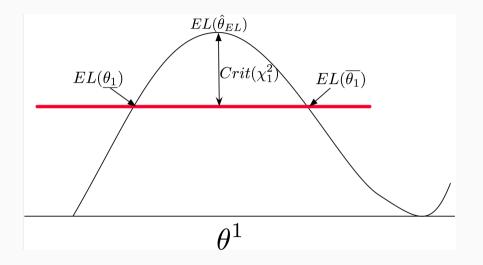
The dual approach presents a number of computational challenges:

- For each guess of  $\theta$  we must find the optimal p (actually  $\gamma$ ), but it may be that  $\nexists p$  s.t  $\sum_{i=1}^{n} p_i g(z_i, \theta) = 0$  at some  $\theta$ .<sup>4</sup>
- At some  $z_i$  we may find that  $\gamma' g(z_i, \theta) \leq -1$ .
- Not much in terms of max min solvers (stuck with nested
- max of a convex function (or the min of a concave function).
- $\nabla_{\theta} \cdot \mathit{I}_{\mathit{NP}}(\theta) = \nabla_{\theta} \left[ \min_{\gamma \in \mathbb{R}^q} \sum_{i=1}^n \log(1 + \gamma' g(z_i, \theta)) \right]$  hard

#### EL Inference

- We can construct likelihood based testing units via Empirical Likelihood Ratio test statistic:  $\frac{\mathit{EL}(\hat{\theta_r})}{\mathit{EL}(\hat{\theta})} \sim \chi_r^2$
- We can invert the test statistic to construct confidence intervals
- MPEC lets us impose the ELR test statistic as an additional restriction
- Solve problems such as  $\max \theta_1 : \mathit{EL}(\theta) > \mathit{EL}(\hat{\theta_0}) \mathit{crit}$
- Nice for nonlinear predictions such as elasticities to test directly without delta method (still have higher order efficiency).

## **EL** Inference



## Altonji Segal 1996 / Abowd Card 1989

The dataset (PSID) tracks 1536 individuals over ten years, and 210 moment conditions are constructed out of all of the possible permutations of variances, covariances, and autocovariances for different lags. A stationary model is estimated with 45 parameters.

	MPEC-EL <sup>5</sup>	$Dual^6$	ELike-M <sup>7</sup>
Moments	210	210	210
Parameters	1591	255 (45)	255(45)
Iterations	65	1023	10,000+
# Converged	50	46	42
Time	$\approx 25s$	pprox 20 m	pprox 45 m

#### Dube Fox Su 2009

An MPEC Algorithm for computing the same estimator as BLP has been suggested:

#### **BLP-MPEC**

The estimator solves the following mathematical program:

$$\min_{\theta, \xi, s, g} g(\xi)'Wg(\xi) \quad \text{s.t.}$$

$$g(\xi) = \frac{1}{N} \sum_{\forall j, t} \xi'_{jt} z_{jt}$$

$$s_{jt}(\theta) = \int \frac{\exp[x_{jt}\beta_i + \xi_{jt} - \alpha_i p_{jt}]}{1 + \sum_k \exp[x_{kt}\beta_i + \xi_{kt} - \alpha_i p_{kt}]} f(\beta_i | \theta)$$

$$\log(S_{jt}) = \log(s_{jt}(\beta, \alpha, \xi, \theta)) \quad \forall j, t$$

Nevo Example

## (Knittel Metaxoglou 2008)

Recent paper (just updated) takes 10 algorithms, 50 starting values and uncovers 100+ parameter estimates and Nevo code/data:

- a local minimum may yield parameter values that are close to the true values but have an objective function value that is very different. Therefore we focus on the economic meaning of the variation in parameter estimates...
- ... researchers will need to use multiple starting values, at least 50 and multiple algorithms
- Mistakes abound
- What weight matrix is used?

## Nevo Results

Nevo	BLP-MPEC	EL
-28.189	-62.726	-61.433
0.330	0.558	0.524
2.453	3.313	3.143
0.016	-0.006	0
0.244	0.093	0.085
15.894	588.206	564.262
-1.200	-30.185	-28.930
2.634	11.058	11.700
5.482	2.29084	2.246
0.2037	1.284	1.37873
29.3611	4.564	
		-17422
28 s	12s	19s
	-28.189 0.330 2.453 0.016 0.244 15.894 -1.200 2.634 5.482 0.2037 29.3611	-28.189       -62.726         0.330       0.558         2.453       3.313         0.016       -0.006         0.244       0.093         15.894       588.206         -1.200       -30.185         2.634       11.058         5.482       2.29084         0.2037       1.284         29.3611       4.564

## Profile Empirical Likelihood

