Graduate Industrial Organization

Lecture: Moment Inequalities Part 1, Theory

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Part 1: Building Moment Inequalities

Discrete choice applications

- Today's goal: illustrate how to build moment inequalities that identify the parameters of static single-agent discrete choice models.
- Comments:
 - In theory, moment inequalities can be used beyond discrete choice settings.
 - Most applied papers that have used moment inequalities (to date) aim to estimate parameters of the utility or production function of agents choosing among a finite set of alternatives.
 - Today, we omit dynamic discrete choice (e.g. Morales, Sheu, and Zahler (2018)) and dynamic games (e.g. Ciliberto and Tamer (2009))

^{*}Thanks to Eduardo Morales for providing source material for the lecture content

Static Discrete Choice Problem: Utility

• Utility of agent *i* for alternative *j* is:

$$U_{ij} = \beta \mathbb{E}[\mathsf{x}_{ij}|\mathcal{W}_i] + \nu_{ij}, \qquad j = 1, \dots, \mathcal{J}_i,$$

- ullet $\mathbb{E}[\cdot]$ is the expectation operator with respect to the data generating process
- x_{ii} is a vector of covariates the researcher observes
- W_i is the information set that agent i uses to predict the value of $x_i \equiv \{x_{i1}, \dots, x_{i,7_i}\}.$
- This specification assumes agents have rational expectations: we define agents' expectations with respect to the data generating process.
- If we assume agents have perfect foresight, the utility function simplifies to:

$$U_{ij} = \beta x_{ij} + \nu_{ij}, \qquad j = 1, \dots, \mathcal{J}_i.$$

Static Discrete Choice Problem: Decision

• Define d_{ii} as a dummy variable that takes value 1 if individual i chooses alternative j. We assume that

$$d_{ij} = \mathbb{I}\big\{\beta \mathbb{E}[x_{ij}|\mathcal{W}_i] + \nu_{ij} \geq \max_{j' \in \mathcal{J}_i} \beta \mathbb{E}[x_{ij'}|\mathcal{W}_i] + \nu_{ij'}\big\}.$$

Denote as ε_{ii} the error that agent i makes when predicting x_{ii} :

$$\varepsilon_{ij} = x_{ij} - \mathbb{E}[x_{ij}|\mathcal{W}_i].$$

• Therefore, we can rewrite d_{ii} as

$$d_{ij} = \mathbb{1}\big\{\beta x_{ij} + \nu_{ij} - \beta \varepsilon_{ij} \ge \max_{j' \in \mathcal{J}_i} \beta x_{ij'} + \nu_{ij'} - \beta \varepsilon_{ij'}\big\}.$$

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Static Discrete Choice Problem: Unobserved Components

- ν_{ii} is the structural error.
 - captures elements of the payoff function the agent knows when making a decision $(\nu_{ii} \subseteq \mathcal{W}_i)$ but which the econometrician does not observe.
 - Economic theory generally imposes no restriction on the distribution of the vector $(\nu_{i1}, \dots, \nu_{i,\mathcal{I}_i})$ across individuals. We often impose assumptions on this distribution for identification/convenience purposes.
- ε_{ii} is the expectational error.
 - captures elements affecting the payoff relevant variable x_{ii} that the agent does not know when making a decision.
 - under the assumption of rational expectations:

$$\mathbb{E}[\varepsilon_{ij}|\mathcal{W}_i]=0.$$

 Rational expectations does not imply any additional restriction on the distribution of ε_{ii} conditional on W_i .

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Static Discrete Choice Problem: Example

• Firm i decides whether to enter a market or stay out: $j = \{1, 0\}$, with

$$\pi_{i0} = 0$$
 $\pi_{i1} = \eta^{-1} \mathbb{E}[r_i | \mathcal{W}_i] - f_i = \eta^{-1} \mathbb{E}[r_i | \mathcal{W}_i] - \beta_0 - \beta_1 \operatorname{dist}_i - \nu_i,$

- π_{i1} denotes firm i's expectation of the profits upon entry
- r_i denotes sales revenue conditional on entry
- f_i denotes fixed entry costs; f_i a function of dist_i, the distance between the market and firm i's production.
- ν_i captures unobserved heterogeneity in fixed entry costs.
- The information set W_i captures any variable firm i knows and uses to predict its sales revenue upon entry; i.e. any variable firm i uses to predict r_i

Static Discrete Choice Problem: Data and Parameters

- Data. For a random sample of individuals and a *subset* of all choices in \mathcal{J}_i , the econometrician observes
 - $\bullet \ x_i = \{x_{i1}, \ldots, x_{iJ_i}\},\$
 - $d_i = \{d_{i1}, \ldots, d_{iJ_i}\},\$
 - $z_i = \{z_{i1}, \dots, z_{i\mathcal{J}_i}\}$ such that $z_i \subseteq \mathcal{W}_i$.
- Researchers may be interested in performing counterfactuals with respect to
 - x_i ; e.g. how does entry change if the potential revenue increases in 10%?
 - W_i; e.g. how does entry change if potential entrants become better at predicting ex post revenues in that market?
 - β ; e.g. how does entry change if fixed entry costs are subsidized?
- The parameters of the model (a subset of which may be needed to perform these counterfactuals) are:
 - the vector of preference parameters β ,
 - the joint distribution of $\nu_i = \{\nu_{i1}, \dots, \nu_{i,\mathcal{I}_i}\}$ across individuals,
 - the information set W_i and the conditional density $f(x_i|W_i)$.

Identification Challenge: Unobserved Information Sets

- Researchers generally do not observe agents' information sets, $\{\mathcal{W}_i\}_i$.
- Even if we assume that agents' expectations are rational, we still need to know W_i to correctly define a proxy for the term $\mathbb{E}[x_{ii}|W_i]$ entering U_{ii} .
- Therefore, researchers fail to observe two terms in utility: $\mathbb{E}[x_{ij}|\mathcal{W}_i]$ and ν_{ij}
 - Literature on nonparametric identification provides assumptions needed when ν_{ii} is not observed
 - Literature on moment inequalities provides insights when agents' true information sets are unobserved.

Identification Challenge: Unobserved Information Sets

Pre-Moment Inequality Approach: Manski (1991)

How shall we handle $\mathbb{E}[x_{ij}|\mathcal{W}_i]$ in U_{ij} ?

- Pre-moment inequality approach: assume the researcher observes agents' information sets W_i ex post. That is, the researcher observes all variables ex post that the agent used ex ante.
- Given this assumption, Manski (1991) introduces a two-step estimator:
 - Step 1: Regress the ex-post realization x_{ij} on the observed information set \mathcal{W}_i to obtain a prediction: $\mathbb{E}[\widehat{x_{ij}|\mathcal{W}_i}]$.
 - Step 2: Use restrictions on the distribution of $(\nu_{i1}, \dots, \nu_{i\mathcal{J}_i})$ to estimate β given that:

$$d_{ij} = \mathbb{I}\{\beta \mathbb{E}\widehat{[x_{ij}|\mathcal{W}_i]} + \nu_{ij} \geq \max_{j' \in \mathcal{J}_i} \beta \mathbb{E}\widehat{[x_{ij'}|\mathcal{W}_i]} + \nu_{ij'}\}.$$

Identification Challenge: Unobserved Information Sets

Pre-Moment Inequality Approach: Ahn (1993)

- Ahn (1993) expands this framework by allowing for agents' expectations to have an unobserved component that enters additively.
- Specifically, Ahn (1993) assumes that $\mathbb{E}[x_{ij}|\mathcal{W}_i] = \mathbb{E}[x_{ij}|\mathcal{W}_i^{obs}] + \xi_{ij}$, where the researcher does not observe ξ_{ij}
- Conditional on this separability assumption:

$$d_{ij} = \mathbb{I}\{\beta \widehat{\mathbb{E}[x_{ij}|\mathcal{W}_i^{obs}]} + \xi_{ij} + \nu_{ij} \geq \max_{j' \in \mathcal{J}_i} \beta \widehat{\mathbb{E}[x_{ij'}|\mathcal{W}_i^{obs}]} + \xi_{ij'} + \nu_{ij'}\},$$

and one can identify (and estimate) β as long as one imposes restrictions on the distribution of $(\xi_{i1} + \nu_{i1}, \dots, \xi_{i\mathcal{J}_i} + \nu_{i\mathcal{J}_i})$.

Motivation for Using Moment Inequalities

- When agents' decisions depend on expectations, we'd like to estimate structural parameters of agents' payoff functions while imposing only weak assumptions on agents' information sets.
- Moment inequalities allow us to identify these structural parameters under the assumption that the researcher observes a subset of agents' true information sets.

Motivation for Using Moment Inequalities

- The main objective of using moment inequalities is to relax the assumptions typically needed for point-identification. In today's lecture:
 - Ex. assumptions on agents' information sets
 - Ex. assumptions on agents' consideration sets
- But there are others!
 - Ex. in static entry games with multiple equilibria, want to relax equilibrium selection assumptions (Ciliberto and Tamer, 2009).

Why Not Switch to Moment Inequalities?

Limitations

- (1) While we need only weak assumptions on the distribution of the expectational errors $(\varepsilon_{i1}, \dots, \varepsilon_{i\mathcal{J}_i})$, we need strong assumptions on the distribution of the structural errors $(\nu_{i1}, \dots, \nu_{i\mathcal{J}_i})$
 - Exception: Pakes and Porter (2016) and Illanes (2016) impose weak assumptions on $(\nu_{i1}, \dots, \nu_{i\mathcal{J}_i})$ but assume away expectational errors
- (2) Unclear mapping between assumptions needed for identification and the assumptions needed to perform counterfactuals.
 - Exception (for very specific settings): Dickstein and Morales (2018).
- (3) Severe computational difficulties arise in the estimation of β when the dimensionality of β is relatively large.
 - Standard inference procedures require evaluating a criterion function at each point in a grid covering the parameter space.
 - New: Chen, Christensen, and Tamer (2018) provide a Monte Carlo sampler alternative

Moment Inequalities: Classification

- The applied literature on moment inequalities has so far studied a fairly limited set of models and has resorted to a limited set of "tricks" to derive moment inequalities.
- We will consider two general types of moment inequalities:
 - revealed-preference moment inequalities
 - odds-based moment inequalities
- Within each general type, the form of the inequalities depends on the assumptions imposed on the distribution of the structural errors $(\nu_{i1}, \ldots, \nu_{i\mathcal{J}_i})$ and the expectational errors $(\varepsilon_{i1}, \ldots, \varepsilon_{i\mathcal{J}_i})$.

Revealed-Preference Inequality: Key Insight

The key insight behind all revealed-preference moment inequalities is that

$$d_{ij} = \mathbb{1}\left\{\beta x_{ij} + \nu_{ij} - \varepsilon_{ij} \ge \max_{i' \in \mathcal{J}_i} \beta x_{ij'} + \nu_{ij'} - \varepsilon_{ij'}\right\} \tag{1}$$

implies that, for any $(j, j') \in \mathcal{J}_i$,

$$d_{ij}(\beta(x_{ij}-x_{ij'})+(\nu_{ij}-\nu_{ij'})+(\varepsilon_{ij}-\varepsilon_{ij'}))\geq 0.$$
 (2)

- In other words, equation (2) is a **necessary** condition for equation (1).
- Equation (2) is also sufficient for equation (1) if and only if the cardinality of the choice set \mathcal{J}_i is equal to two.

Revealed-Preference Inequality: Expectational Error

The inequality

$$d_{ij}(\beta(x_{ij}-x_{ij'})+(\nu_{ij}-\nu_{ij'})+(\varepsilon_{ij}-\varepsilon_{ij'}))\geq 0,$$

cannot be used directly for identification of the parameter vector β , as it depends on the unobserved terms ν_{ij} , $\nu_{ij'}$, ε_{ij} , and $\varepsilon_{ij'}$.

ullet Taking expectations conditional on the true information set \mathcal{W}_i , we obtain

$$\mathbb{E}[d_{ij}(\beta(x_{ij}-x_{ij'})+(\nu_{ij}-\nu_{ij'}))|\mathcal{W}_i]\geq 0,$$

as the assumption that agents have rational expectations implies

$$\mathbb{E}[d_{ij}(\varepsilon_{ij}-\varepsilon_{ij'})|\mathcal{W}_i]=0.$$

Proof:

$$egin{aligned} \mathbb{E}[d_{ij}(arepsilon_{ij'})|\mathcal{W}_i] &= \mathbb{E}[\mathbb{E}[d_{ij}(arepsilon_{ij'}-arepsilon_{ij'})|\mathcal{W}_i,
u_i]|\mathcal{W}_i] \ &= \mathbb{E}[d_{ij}\mathbb{E}[(arepsilon_{ii'}-arepsilon_{ii'})|\mathcal{W}_i,
u_i]|\mathcal{W}_i] &= \mathbb{E}[d_{ij} imes0|\mathcal{W}_i] &= 0. \end{aligned}$$

Revealed-Preference Inequality: Structural Error

• The resulting moment inequality is therefore:

$$\mathbb{E}[d_{ij}\beta(x_{ij}-x_{ij'})|\mathcal{W}_i]+\mathbb{E}[d_{ij}(\nu_{ij}-\nu_{ij'})|\mathcal{W}_i]\geq 0.$$

- What is the second term in this inequality? A selection correction. Intuitively, even if we observe that i prefers j over j', we cannot conclude that $\beta \mathbb{E}[x_{ij}|\mathcal{W}_i] > \beta \mathbb{E}[x_{ij'}|\mathcal{W}_i]$ because j may be preferred over j' because $\nu_{ij} >> \nu_{ij'}$.
- For simplicity in the notation, let's define $s_{ii'}(W_i;\beta)$ as

$$s_{jj'}(\mathcal{W}_i; \beta) \equiv \mathbb{E}[d_{ij}(\nu_{ij} - \nu_{ij'})|\mathcal{W}_i].$$

Revealed-Preference Inequality: Structural Error

- Even if the distribution of ν_{ij} is independent of W_i and identical across j, $s_{ii'}(W_i; \beta) \neq 0$. Why? d_{ii} is a function of $(\nu_{ii} \nu_{ii'})$.
- Furthermore, if the distributions of ν_{ij} and $\nu_{ij'}$ conditional on \mathcal{W}_i are identical, $s_{ii'}(\mathcal{W}_i; \beta) \geq 0$.
- Therefore, it could be that, at the true value of the parameter vector β ,

$$\mathbb{E}[d_{ij}\beta(x_{ij}-x_{ij'})|\mathcal{W}_i]+\mathbb{E}[d_{ij}(\nu_{ij}-\nu_{ij'})|\mathcal{W}_i]\geq 0,$$

but

$$\mathbb{E}[d_{ii}\beta(x_{ii}-x_{ii'})|\mathcal{W}_i]<0.$$

• We therefore need to consider the term $s_{ii'}(W_i; \beta)$.

Revealed-Preference Inequality: Structural Error

Two ways researchers handle the selection correction term $s_{jj'}(W_i; \beta)$:

- (1) Impose assumptions on the distribution of the vector ν_i conditional on \mathcal{W}_i such that $s_{ii'}(\mathcal{W}_i; \beta) = 0$.
- (2) Impose assumptions on the distribution of the vector ν_i conditional on \mathcal{W}_i that allow the researcher to derive a function $\bar{s}_{jj'}(\mathcal{W}_i;\beta)$ such that

$$\bar{s}_{jj'}(\mathcal{W}_i;\beta) \geq s_{jj'}(\mathcal{W}_i;\beta).$$

- One may impose different assumptions on the distribution of the vector $\nu_i = (\nu_{i1}, \dots, \nu_{i\mathcal{J}_i})$ conditional on \mathcal{W}_i such that $s_{jj'}(\mathcal{W}_i; \beta) = 0$.
- We note four cases here:
 - No unobserved heterogeneity;
 - @ Group-of-choices fixed effects;
 - Group-of-individuals fixed effects;
 - Ordered-choice model.

(1) No unobserved heterogeneity

- Implies $\nu_{ij} = 0$ for all $j \in \mathcal{J}_i$.
- Applied in Holmes (2011): Walmart choosing where to open stores.
- Discussed in Pakes (2010) and Pakes et al. (2015).
- In most empirical applications, this assumption is too restrictive.

(2) Group-of-choices fixed effects

• Implies ν_{ij} is common to a subset of choices:

$$u_{ij} - \nu_{ij'} = 0$$
 if $g(j) = g(j')$,

where the function $g(\cdot)$ creates a partition of the set of potential choices.

- Applied in Morales et al. (2018): exporters deciding which markets to enter.
- Discussed in Pakes (2010) and Pakes et al. (2015).
- We can only exploit inequalities that compare potential choices j and j' such that g(j) = g(j').

(3) Group-of-individuals fixed effects

• Implies that ν_{ii} is common to a subset of individuals:

$$\nu_{ij} - \nu_{i'j} = 0 \qquad \text{if} \qquad g(i) = g(i'),$$

where the function $g(\cdot)$ creates a partition of the set of individuals.

- Applied in Ho and Pakes (2014): patients deciding which hospital to visit.
- It requires double-differencing. Find two individuals i and i' such that $d_{ii}=d_{i'i'}=1$ and build

$$\begin{split} \mathbb{E}[\beta(x_{ij}-x_{ij'})+(\nu_{ij}-\nu_{ij'})|d_{ij} &=1, d_{i'j'}=1, \mathcal{W}_i, \mathcal{W}_{i'}] \geq 0, \\ \mathbb{E}[\beta(x_{i'j'}-x_{i'j})+(\nu_{i'j'}-\nu_{i'j})|d_{ij} &=1, d_{i'j'}=1, \mathcal{W}_i, \mathcal{W}_{i'}] \geq 0. \end{split}$$

As long as g(i) = g(i'), we can sum these two inequalities and obtain:

$$\mathbb{E}[\beta(x_{ii} - x_{ii'}) + \beta(x_{i'i'} - x_{i'i}) | d_{ii} = 1, d_{i'i'} = 1, \mathcal{W}_i, \mathcal{W}_{i'}] \geq 0.$$

(4) Ordered-choice model

- Ordered-choice model means $\nu_{ij} = j\eta_i$.
- Additionally, assume that $\mathbb{E}[\eta_i|\mathcal{W}_i] = 0$ and $\mathcal{J}_i = \mathcal{J}_{i'} = \mathcal{J}$ for all i and i'.
- Applied in Ishii (2008): banks deciding how many ATMs to install.
- Discussed in Pakes (2010) and Pakes et al. (2015).
- We must build inequalities such that, for every individual i, the alternative choice j' is one unit below the actual choice of i; i.e. j' = j 1. Therefore

$$\mathbb{E}[d_{ij}\beta(x_{ij}-x_{i(j-1)})|\mathcal{W}_i] + \mathbb{E}[d_{ij}(\nu_{ij}-\nu_{i(j-1)})|\mathcal{W}_i] \geq 0,$$

and, summing these inequalities for all j in the choice set \mathcal{J} ,

$$\sum_{j\in\mathcal{J}}\mathbb{E}[d_{ij}\beta(x_{ij}-x_{i(j-1)})|\mathcal{W}_i]+\sum_{j\in\mathcal{J}}\mathbb{E}[d_{ij}(\nu_{ij}-\nu_{i(j-1)})|\mathcal{W}_i]\geq 0.$$

(4) Ordered-choice model (cont.)

• If each individual i chooses j such that j' = j - 1 is in \mathcal{J} , then

$$\begin{split} \sum_{j \in \mathcal{J}} \mathbb{E}[d_{ij}(\nu_{ij} - \nu_{i(j-1)})|\mathcal{W}_i] &= \sum_{j \in \mathcal{J}} \mathbb{E}[d_{ij}(j\eta_i - (j-1)\eta_i)|\mathcal{W}_i] \\ &= \sum_{j \in \mathcal{J}} \mathbb{E}[d_{ij}\eta_i|\mathcal{W}_i] \\ &= \mathbb{E}[\sum_{j \in \mathcal{J}} d_{ij}\eta_i|\mathcal{W}_i] \\ &= \mathbb{E}[\eta_i \sum_{j \in \mathcal{J}} d_{ij}|\mathcal{W}_i] \\ &= \mathbb{E}[\eta_i \times 1|\mathcal{W}_i] \\ &= \mathbb{E}[\eta_i|\mathcal{W}_i] = 0. \end{split}$$

(4) Ordered-choice model (cont.)

The resulting moment inequality is therefore

$$\sum_{j\in\mathcal{J}}\mathbb{E}[d_{ij}\beta(x_{ij}-x_{i(j-1)})|\mathcal{W}_i]\geq 0.$$

• Similarly, we can build inequalities for which the alternative choice j' is one unit above the actual choice of i, i.e. j' = j + 1. In this case

$$\sum_{i \in \mathcal{J}} \mathbb{E}[d_{ij}(\nu_{ij} - \nu_{i(j+1)})|\mathcal{W}_i] = -\mathbb{E}[\eta_i|\mathcal{W}_i] = 0,$$

and the resulting moment inequality becomes

$$\sum_{j\in\mathcal{J}}\mathbb{E}[d_{ij}\beta(x_{ij}-x_{i(j+1)})|\mathcal{W}_i]\geq 0.$$

• If some individuals choose j such that either j-1 or j+1 do not belong to \mathcal{J} , selection issues arise. See Pakes (2010) for a discussion.

Dickstein and Morales (2018)

• The usefulness of deriving an upper bound $\bar{s}_{jj'}(\mathcal{W}_i;\beta)$ on the term

$$s_{jj'}(\mathcal{W}_i; \beta) \equiv \mathbb{E}[d_{ij}(\nu_{ij} - \nu_{ij'}) | \mathcal{W}_i]$$

is that, if it is true that the following inequality holds at the true value of the parameter vector β ,

$$\mathbb{E}[d_{ij}\beta(x_{ij}-x_{ij'})|\mathcal{W}_i]+\mathbb{E}[d_{ij}(\nu_{ij}-\nu_{ij'})|\mathcal{W}_i]\geq 0, \tag{3}$$

then it will also be true that, at the true value of the parameter vector β ,

$$\mathbb{E}[d_{ij}\beta(x_{ij}-x_{ij'})|\mathcal{W}_i]+\bar{s}_{jj'}(\mathcal{W}_i;\beta)\geq 0. \tag{4}$$

• The cost of using equation (4) for identification (instead of equation (3)) is that the identified set will be larger; i.e. the set of values of β consistent with this inequality will be larger.

Dickstein and Morales (2018)

- One can construct such upper bound $\bar{s}_{jj'}(\mathcal{W}_i;\beta)$ as long as the assumed distribution of ν_i conditional on \mathcal{W}_i is such that for any pair of alternatives j and j', the distribution of $\nu_{ij} \nu_{jj'}$ conditional on \mathcal{W}_i verifies:
 - it is known up to a finite parameter vector
 - it has mean zero
 - its truncated expectation is convex in the truncation point.
- ullet Using mathematical notation, for all j and j' included \mathcal{J}_i , it must hold that

$$\mathbb{E}[\nu_{ij} - \nu_{ij'} | \mathcal{W}_i] = 0, \quad \text{and} \quad \mathbb{E}[\nu_{ij} - \nu_{ij'} | \mathcal{W}_i, \nu_{ij} - \nu_{ij'} \ge \lambda]$$
 (5)

is a known function of λ and convex in λ .

• Both the normal and the logistic distribution verify these restrictions.

Odds-Based Inequality: Key Insight

The key insight behind all odds-based moment inequalities is that

$$d_{ij} = \mathbb{1}\left\{\beta \mathbb{E}[x_{ij}|\mathcal{W}_i] + \nu_{ij} \ge \max_{j' \in \mathcal{J}_i} \beta \mathbb{E}[x_{ij'}|\mathcal{W}_i] + \nu_{ij'}\right\}$$
(6)

implies that, for any $(j,j') \in \mathcal{J}_i$,

$$\mathbb{E}\left[\mathbb{1}\left\{\beta\mathbb{E}\left[x_{ij}-x_{ij'}|\mathcal{W}_{i}\right]+\left(\nu_{ij}-\nu_{ij'}\right)\geq0\right\}-d_{ij}|\mathcal{W}_{i}\right]\geq0.$$
 (7)

- In words, the probability that j is preferred over j' must be weakly larger than the probability that j is preferred over any other alternative.
- Equation (7) is a **necessary** condition for equation (6).
- Equation (7) is exactly equal to zero if and only if the cardinality of the choice set \mathcal{J}_i is equal to two.
- Pakes and Porter (2016) assume that, for all j, $\mathbb{E}[x_{ij}|\mathcal{W}_i] = x_{ij}$, and exploit equation (7) and longitudinal data to derive moment inequalities that do not impose any parametric assumption on the distribution of ν_i .

Odds-Based Inequalities: Assumptions

- Dickstein and Morales (2018) introduce the odds-based inequality for the special case of single-agent binary choice models.
- In order to derive odds-based inequalities in Dickstein and Morales (2018), the distribution of $\nu_{ij} \nu_{ij'}$ conditional on W_i must be:
 - known up to a finite parameter vector; and,
 - log-concave.
- The log-concavity of the distribution of $\nu_{ij} \nu_{ij'}$ conditional on \mathcal{W}_i implies:

$$\frac{F_{\nu|\mathcal{W}}(\lambda)}{1 - F_{\nu|\mathcal{W}}(\lambda)} \quad \text{and} \quad \frac{1 - F_{\nu|\mathcal{W}}(\lambda)}{F_{\nu|\mathcal{W}}(\lambda)}$$

are convex in the index λ , where $F_{\nu|\mathcal{W}}(\cdot)$ is the CDF of $\nu_{ij} - \nu_{ij'}$ conditional on the information set \mathcal{W}_i .

• Both the normal and the logistic distribution are log-concave.

Switching Conditioning Sets

All moment inequalities derived above have the following form

$$\mathbb{E}[m(d_{ij},x_{ij},x_{ij'};\beta)|\mathcal{W}_i] \geq 0.$$

- These inequalities cannot be directly used for identification because the researcher does not know the content of the agents' information sets \mathcal{W} .
- However, conditional on observing a vector $z_i \subset W_i$, we can use the Law of Iterated Expectations to derive the inequalities

$$\mathbb{E}[m(d_{ij},x_{ij},x_{ij'};\beta)|z_i]\geq 0.$$

• No matter what the function $m(\cdot)$ is, these inequalities are a function of observed data, $(d_{ii}, x_{ii}, x_{ii'}, z_i)$, and the parameter vector, β . Therefore, they may be used for identification.