Kernel Methods

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STAT 4690-Applied Multivariate Analysis

Motivation

- Linearity has been an important assumption for most of the multivariate methods we have discussed.
 - Multivariate Linear Regression
 - PCA, FA, CCA
- This assumption may be more realistic after a transformation of the data.
 - E.g. Log transformation
 - Embedding in a higher dimensional space?

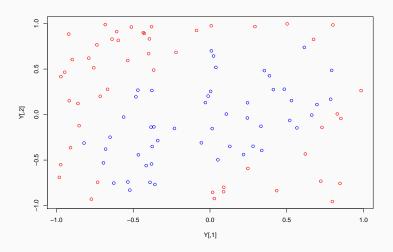
Example i

```
set.seed(1234)
n < -100
# Generate uniform data
Y \leftarrow cbind(runif(n, -1, 1),
            runif(n, -1, 1))
# Check if it falls inside ellipse
Sigma \leftarrow matrix(c(1, 0.5, 0.5, 1), ncol = 2)
dists <- sqrt(diag(Y %*% solve(Sigma) %*%
                      t(Y)))
inside <- dists < 0.85
```

Example ii

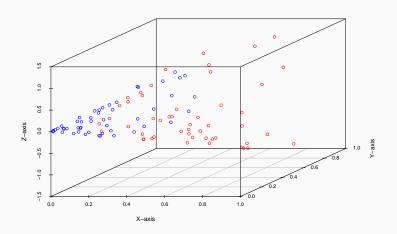
```
# Plot points
colours <- c("red", "blue")[inside + 1]
plot(Y, col = colours)</pre>
```

Example iii

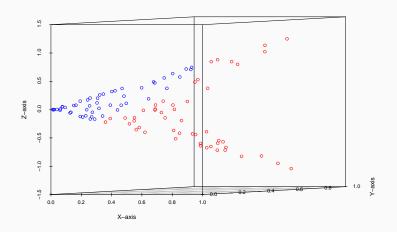


Example iv

Example v



Example vi



Example vii

```
# Linear regression
outcome <- ifelse(inside, 1, -1)
head(outcome)
## [1] -1 1 1 -1 -1 1
model1 <- lm(outcome ~ Y)
pred1 <- sign(predict(model1))</pre>
table(outcome, pred1) # 67%
```

Example viii

```
## pred1
## outcome -1 1
## -1 32 18
## 1 15 35

model2 <- lm(outcome ~ Y_transf)
pred2 <- sign(predict(model2))
table(outcome, pred2) # 92%</pre>
```

Example ix

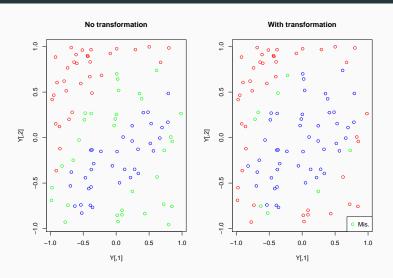
```
## pred2

## outcome -1 1

## -1 44 6

## 1 2 48
```

Example x



Overfitting i

- Overfitting is "the production of an analysis that corresponds too closely or exactly to a particular set of data, and may therefore fail to fit additional data or predict future observations reliably" (OED)
 - In other words, a model is overfitted if it explains the training data very well, but does poorly on test data.
- In regression, this often happens when we have too many covariates
 - Too many parameters for the sample size

Overfitting ii

- When embedding our covariates into a higher dimensional space, we are increasing the number of parameters.
 - There is a danger of overfitting.
- One possible solution: Regularised (or penalised) regression.
 - We constrain the parameter space using a penalty function.

Ridge regression i

- Let (Y_i, \mathbf{X}_i) , $i = 1, \dots, n$ be a sample of outcome with covariates.
- Univariate Linear Regression: Assume that we are interested in the linear model

$$Y_i = \beta^T \mathbf{X}_i + \epsilon_i.$$

 \bullet The Least-Squares estimate of β is given by

$$\hat{\beta}_{LS} = (\mathbb{X}^T \mathbb{X})^{-1} \mathbb{X} \mathbf{Y},$$

where

Ridge regression ii

$$\mathbb{X}^T = \begin{pmatrix} \mathbf{X}_1 & \cdots & \mathbf{X}_n \end{pmatrix},$$
$$\mathbf{Y} = (Y_1, \dots, Y_n).$$

- If the matrix $\mathbb{X}^T\mathbb{X}$ is almost singular, then the least-squares estimate will be unstable.
- **Solution**: Add a small quantity along its diagonal.

 - Bias-Variance trade-off

Ridge regression iii

 \bullet The Ridge estimate of β is given by

$$\hat{\beta}_R = (\mathbb{X}^T \mathbb{X} + \lambda I)^{-1} \mathbb{X}^T \mathbf{Y}.$$

Example i

```
library(ElemStatLearn)
library(tidyverse)
data train <- prostate %>%
  filter(train == TRUE) %>%
  dplyr::select(-train)
data test <- prostate %>%
  filter(train == FALSE) %>%
  dplyr::select(-train)
```

Example ii

Example iii

Example iv

[1] 0.5060843

Dual problem i

 The Ridge estimate actually minimises a regularized least-squares function:

$$RLS(\beta) = \frac{1}{2} \sum_{i=1}^{n} (Y_i - \beta^T \mathbf{X}_i)^2 + \frac{\lambda}{2} \beta^T \beta.$$

• If we take the derivative with respect to β , we get

$$\frac{\partial}{\partial \beta} RLS(\beta) = -\sum_{i=1}^{n} (Y_i - \beta^T \mathbf{X}_i) \mathbf{X}_i + \lambda \beta.$$

Setting it equal to 0 and rearranging, we get

$$\beta = \frac{1}{\lambda} \sum_{i=1}^{n} (Y_i - \beta^T \mathbf{X}_i) \mathbf{X}_i.$$

Dual problem ii

• Define $a_i = \frac{1}{\lambda}(Y_i - \beta^T \mathbf{X}_i)$. We then get

$$\beta = \sum_{i=1}^{n} a_i \mathbf{X}_i = \mathbb{X}^T \alpha,$$

where $\alpha = (a_1, \ldots, a_n)$.

• Why? We can now rewrite $RLS(\beta)$ as a function of α . First note that

$$RLS(\beta) = \frac{1}{2} (\mathbf{Y} - \mathbb{X}\beta)^T (\mathbf{Y} - \mathbb{X}\beta) + \frac{\lambda}{2} \beta^T \beta.$$

Dual problem iii

• Now we can substitute $\beta = \mathbb{X}^T \alpha$:

$$RLS(\alpha) = \frac{1}{2} (\mathbf{Y} - \mathbb{X} \mathbb{X}^T \alpha)^T (\mathbf{Y} - \mathbb{X} \mathbb{X}^T \alpha) + \frac{\lambda}{2} (\mathbb{X}^T \alpha)^T (\mathbb{X}^T \alpha)$$
$$= \frac{1}{2} (\mathbf{Y} - (\mathbb{X} \mathbb{X}^T) \alpha)^T (\mathbf{Y} - (\mathbb{X} \mathbb{X}^T) \alpha) + \frac{\lambda}{2} \alpha^T (\mathbb{X} \mathbb{X}^T) \alpha.$$

- \blacksquare This formulation of regularised least squares in terms of α is called the **dual problem**.
- **Key observation**: $RLS(\alpha)$ depends on X_i only through the Gram matrix $\mathbb{X}\mathbb{X}^T$.
 - If we all we know are the dot products of the covariates X_i , we can still solve the ridge regression problem.

Kernel ridge regression i

- Suppose we have a transformation $\Phi: \mathbb{R}^p \to \mathbb{R}^N$, where N is typically larger than p and can even be infinity.
- Let K be the $n \times n$ matrix whose (i, j)-th entry is the dot product between $\Phi(\mathbf{X}_i)$ and $\Phi(\mathbf{X}_j)$:

$$K_{ij} = \Phi(\mathbf{X}_i)^T \Phi(\mathbf{X}_j).$$

• Important observation: This actually induces a map on pairs of points in \mathbb{R}^p :

$$k(\mathbf{X}_i, \mathbf{X}_j) = \Phi(\mathbf{X}_i)^T \Phi(\mathbf{X}_j).$$

• We will call the function k the **kernel function**.

Kernel ridge regression ii

Now, we can use the dual formulation of ridge regression to fit a linear model between Y_i and the transformed $\Phi(X_i)$:

$$Y_i = \beta^T \Phi(\mathbf{X}_i) + \epsilon_i.$$

• By setting the derivative of $RLS(\alpha)$ equal to zero and solving for α , we see that

$$\hat{\alpha} = (K + \lambda I_n)^{-1} \mathbf{Y}.$$

Kernel ridge regression iii

Note that we would need to know all the images $\Phi(\mathbf{X}_i)$ to recover $\hat{\beta}$ from $\hat{\alpha}$. On the other hand, we don't actually need $\hat{\beta}$ to obtain the *fitted* values:

$$\hat{\mathbf{Y}} = \Phi(\mathbb{X})\hat{\beta} = \Phi(\mathbb{X})\Phi(\mathbb{X})^T\hat{\alpha} = K\hat{\alpha}.$$

• To obtain the predicted value for a new covariate profile $\tilde{\mathbf{X}}$, first compute all the dot products in the feature space:

$$\mathbf{k} = (k(\mathbf{X}_1, \tilde{\mathbf{X}}), \dots, k(\mathbf{X}_n, \tilde{\mathbf{X}})).$$

Kernel ridge regression iv

We can then obtain the predicted value:

$$\tilde{Y} = \hat{\beta}^T \Phi(\tilde{\mathbf{X}})
= \hat{\alpha}^T \Phi(\mathbb{X}) \Phi(\tilde{\mathbf{X}})
= \hat{\alpha}^T \mathbf{k}
= \mathbf{k}^T (K + \lambda I_n)^{-1} \mathbf{Y}.$$

Example (cont'd) i

Example (cont'd) ii

```
# Ridge regression
beta hat <- solve(crossprod(X train) +
                    0.7*diag(ncol(X train))) %*%
 t(X train) ** Y train
beta hat[1:3]
## [1] 0.1323063 0.5709660 0.6160020
```

Example (cont'd) iii

```
# Dual problem
alpha hat <- solve(tcrossprod(X train) +</pre>
                     0.7*diag(nrow(X train))) %*%
  Y train
(t(X train) %*% alpha hat)[1:3]
## [1] 0.1323063 0.5709660 0.6160020
all.equal(beta_hat, t(X_train) %*% alpha hat)
## [1] TRUE
```

Important observation

- We assumed that we had an embedding of the data into a higher dimensional space.
- But our derivation only required the dot products of our observations in the feature space.
- Therefore, we don't need to explicitly define the transformation.
- All we need is to define a kernel function.