Modern Applied Statistics Chap 12: Classification

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Introduction

Introduction

In the statistical literature the word is used in two distinct senses.

- The sense of cluster analysis discussed in Section 11.2
- The other meaning (Ripley, 1997) of allocating future cases to one of g prespecified classes

It is sometimes helpful to distinguish discriminant analysis in the sense of describing the differences between the g groups from classification, allocating new observations to the groups.

- The first provides some measure of explanation
- The second can be a 'black box' that makes a decision without any explanation.

Suppose that we have a set of g classes, and for each case we know the class. We can then use the class information to help reveal the structure of the data.

The sample covariance matrices

$$W = \frac{(X - GM)^T (X - GM)}{n - g}, \quad B = \frac{(GM - 1\bar{x})^T (GM - 1\bar{x})}{g - 1}$$

- W: the within-class covariance matrix.
- B: the between-classes covariance matrix.
- M: the $g \times p$ matrix of class means.
- G: the $n \times g$ matrix of class indicator variables.
 - Then the predictions are GM.
- \bar{x} : the means of the variables over the whole sample.

Note that B has rank at most min(p, g - 1).

Fisher introduced a **linear discriminant analysis** seeking a linear combination xa of the variables that has a maximal ratio of the separation of the class means to the within-class variance.

- ► Maximizing the ratio $\mathbf{a}^T B \mathbf{a} / \mathbf{a}^T W \mathbf{a}$.
 - Choose a sphering xS of the variables.
 - The problem is to maximize $a^T B a$ subject to ||a|| = 1.
 - This is solved by taking *a* to be the eigenvector of *B* corresponding to the largest eigenvalue.
 - The linear combination *a* is unique up to a change of sign.

As for principal components, we can take further linear components corresponding to the next largest eigenvalues.

- **Eigenvalues**: the proportions of the between classes variance explained by the linear combinations.
- The corresponding transformed variables are called the linear discriminants or canonical variates.
- The linear discriminants are conventionally centred to have mean zero on dataset.

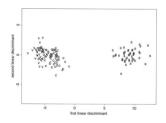


Figure 12.1: The log iris data on the first two discriminant axes.

- ▶ This shows that 99.65% of the between-group variance is on the first discriminant axis.
 - Rao (1948) used the unweighted covariance matrix of the group means.
 - Our approach uses a covariance matrix weighted by the prior probabilities of the classes if these are specified.

Discrimination for normal populations

An alternative approach to discrimination is via probability models.

 \blacktriangleright Posterior distribution of the classes after observing x is

$$p(c \mid \mathbf{x}) = \frac{\pi_c p(\mathbf{x} \mid c)}{p(\mathbf{x})} \propto \pi_c p(\mathbf{x} \mid c)$$

- π_c : The prior probabilities of the classes
- p(x | c): The densities of distributions of the observations for each class
- Bayes rule: The allocation rule which makes the smallest expected number of errors chooses the class with maximal $p(c \mid x)$.

Suppose the distribution for class c is multivariate normal with mean μ_c and covariance Σ_c . Then the Bayes rule minimizes

$$Q_c = -2\log p(\mathbf{x} \mid c) - 2\log \pi_c$$

= $(\mathbf{x} - \mu_c) \Sigma_c^{-1} (\mathbf{x} - \mu_c)^T + \log |\Sigma_c| - 2\log \pi_c$

- The first term is the squared Mahalanobis distance to the class centre.
- The difference between the Q_c for two classes is a quadratic function of x.
 - ► Quadratic discriminant analysis.
- The boundaries of the decision regions are quadratic surfaces in x space.

Suppose that the classes have a common covariance matrix Σ .

▶ Differences in the Q_c are then linear functions of x.

We can maximize $-Q_c/2$ or

$$L_c = \mathbf{x} \Sigma^{-1} \boldsymbol{\mu}_c^T - \boldsymbol{\mu}_c \Sigma^{-1} \boldsymbol{\mu}_c^T / 2 + \log \pi_c$$

To use Q_c or L_c we have to estimate μ_c and Σ_c or Σ .

▶ Using obvious estimates, estimate μ_c as the sample mean, Σ_c as covariance matrix, and Σ as W.

How does this relate to Fisher's linear discrimination?

- $ightharpoonup L_c$ gives new variables, the linear discriminants, with unit within-class sample variance.
 - On these variables the Mahalanobis distance is

$$\|x-\mu_c\|^2$$

 \blacktriangleright Only the first r components of the vector depend on c.

$$L_c = \mathbf{x} \boldsymbol{\mu}_c^T - \|\boldsymbol{\mu}_c\|^2 / 2 + \log \pi_c$$

We can work in r dimensions.

$$L_2 - L_1 = \boldsymbol{x} \left(\boldsymbol{\mu}_2 - \boldsymbol{\mu}_1 \right)^T + \text{ const}$$

▶ An affine function of the linear discriminant.

Crabs dataset

Construct a rule to predict the sex of a future Leptograpsus crab of unknown colour form.

- Linear discriminant analysis, for what are highly non-normal populations, finds a variable that is essentially ${\rm CL^3RW^{-2}CW^{-1}}$, a dimensionally neutral quantity.
- Six errors are made, all for the blue form.

Crabs dataset

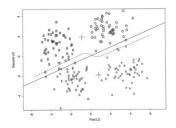


Figure 12.2: Linear discriminants for the crabs data. Males are coded as capitals, females as lower case, colours as the initial letter of blue or orange. The crosses are the group means for a linear discriminant for sex (solid line) and the dashed line is the decision boundary for sex based on four groups.

- ▶ The first two linear discriminants dominate the between-group variation.
- ▶ Using the first two linear discriminants as the data will provide a very good approximation

Robust estimation of multivariate location and scale

- Apply a robust location estimator to each component of a multivariate mean.
- Consider the estimation of mean and variance simultaneously.
- Two methods for robust covariance estimation
 - ► Our function cov.rob
 - ► The S-PLUS functions cov.mve and cov.mcd and covRobin in library section robust.
 - MVE(minimum volume ellipsoid method): seeks an ellipsoid containing $h = \lfloor (n+p+1)/2 \rfloor$ points that is of minimum volume.
 - MCD(minimum covariance determinant method): eeks *h* points whose covariance has minimum determinant.

Robust estimation of multivariate location and scale

- The search for an MVE or MCD
 - ▶ provides *h* points whose mean and variance matrix give an initial estimate.
 - ▶ This is refined by selecting those points whose Mahalanobis distance from the initial mean using the initial covariance is not too large
 - ▶ and returning their mean and variance matrix.
- Alternative approach
 - \blacktriangleright Fitting a multivariate t_{ν} distribution for a small number ν of degrees of freedom.
 - ► This is implemented in our function cov.trob
- Normally cov.trob is faster than cov.rob, but it lacks the latter's extreme resistance.

- In the terminology of pattern recognition
 - ▶ Training set: The given examples together with their classifications.
 - ▶ Test set: Future cases form.
 - ▶ Our primary measure of success is the error rate.
- The type of errors
 - \blacktriangleright A confusion matrix gives the number of cases with true class i classified as of class j.
 - ▶ We assign costs L_{ij} to allocating a case of class i to class j.

The average error cost is minimized by the Bayes rule, which is to allocate to the class c minimizing $\sum_i L_{ic} p(i \mid x)$.

- p(i | x) is the posterior distribution of the classes after observing
 x.
- If the costs of all errors are the same, this rule amounts to choosing the class c with the largest posterior probability $p(c \mid x)$.

The minimum average cost is known as the Bayes risk.

• Estimate a lower bound for it by the method of Ripley (1996, pp. 196-7).

In Section 12.1,

- How $p(c \mid x)$ can be computed for normal populations.
- How estimating the Bayes rule with equal error costs leads to linear and quadratic discriminant analysis.

As our functions predict.lda and predict.qda return posterior probabilities.

▶ They can be used for classification with error costs.

The posterior probabilities $p(c \mid x)$ may also be estimated directly.

- For two classes: we can model $p(1 \mid x)$ using a logistic regression, fitted by glm.
- For more than two classes: we need a multiple logistic model.
 - ▶ Using a surrogate log-linear Poisson GLM model.
 - ▶ Using the multinom function in library section nnet.

Classification trees model the $p(c \mid x)$ directly.

► Since the posterior probabilities are given by the **predict** method it is easy to estimate the Bayes rule for unequal error costs.

Predictive and 'plug-in' rules

To find the Bayes rule we need to know the posterior probabilities $p(c \mid x)$.

- Since these are unknown, we use an parametric family $p(c \mid x; \theta)$.
- We act as if $p(c \mid x; \hat{\theta})$ were the actual posterior probabilities.
 - \blacktriangleright $\hat{\theta}$ is an estimate computed from the training set \mathcal{T} .
 - ▶ 'plug-in' rule.

The 'correct' estimate of $p(c \mid x)$ is (Ripley, 1996, §2.4) to use the predictive estimates

$$\tilde{p}(c \mid \mathbf{x}) = P(c = c \mid \mathbf{X} = \mathbf{x}, \mathcal{T}) = \int p(c \mid \mathbf{x}; \theta) p(\theta \mid \mathcal{T}) d\theta$$

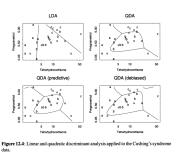
Predictive and 'plug-in' rules

$$\tilde{p}(c \mid \mathbf{x}) = P(c = c \mid \mathbf{X} = \mathbf{x}, \mathcal{T}) = \int p(c \mid \mathbf{x}; \theta) p(\theta \mid \mathcal{T}) d\theta$$

- ▶ It is possible for linear and quadratic discrimination with appropriate 'vague' priors on θ .
- ▶ This estimate is implemented by method = "predictive" of the predict methods for our functions lda and qda.
- ▶ When there are not enough data for a good estimate of the variance matrices, Considerable improvement can be obtained by using an unbiased estimator of $\log p(x \mid c)$, implemented by the argument method = "debiased".

A simple example: Cushing's syndrome

- The data are on diagnostic tests on patients with Cushing's syndrome.
- This dataset has three recognized types of the syndrome represented as a, b, c.
- The observations are urinary excretion rates, and are considered on log scale.



- ightharpoonup There are six patients of unknown type (marked u), one of whom
- was later found to be of a fourth type, and another was measured faultily.
- ▶ Figure 12.4 shows the classifications produced by 1 da and the various options of quadratic discriminant analysis.

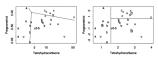


Figure 12.5: Logistic regression and classification trees applied to the Cushing's syndrome data.

▶ When, as here, the classes have quite different variance matrices, linear and logistic discrimination can give quite different answers.

Non-parametric Rules

Non-parametric Rules

- Library section class: The *k*-nearest neighbour classifier and related methods and learning vector quantization.
- These are all based on
 - ▶ finding the *k* **nearest examples** in some reference set.
 - ▶ taking a **majority vote** among the classes of these *k* examples.
 - \blacktriangleright estimating the **posterior probabilities** $p(c \mid x)$ by the proportions of the classes among the k examples.
- These methods almost always measure 'nearest' by Euclidean distance.

Non-parametric Rules

For the **Cushing's syndrome data** we use Euclidean distance on the logged covariates, rather arbitrarily scaling them equally.

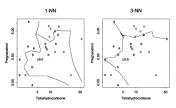


Figure 12.6: k-nearest neighbours applied to the Cushing's syndrome data.

▶ This dataset is too small to try the editing and LVQ methods in library section class.

Neural Networks

Neural Networks

Neural networks provide a flexible non-linear extension of multiple logistic regression, as we saw in Section 8.10.

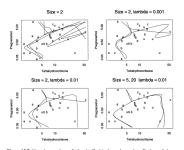


Figure 12.7: Neural networks applied to the Cushing's syndrome data. Each panel shows the fits from two or three local maxima of the (penalized) log-likelihood.

▶ We see that in all cases there are multiple local maxima of the likelihood, since different runs gave different classifiers.

Neural Networks

Once we have a **penalty**, the choice of the number of hidden units is often not critical (see Figure 12.7). The spirit of the predictive approach is to average the predicted $p(c \mid x)$ over the local maxima.

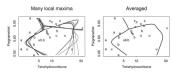


Figure 12.8: Neural networks with three hidden units and $\lambda = 0.01$ applied to the Cushing's syndrome data.

▶ Note that there are two quite different types of local maxima occurring here, and some local maxima occur several times (up to convergence tolerances). An average does better than either type of classifier.

Support vector machines (SVMs)

The method for g = 2 classes

- Logistic regression will fit exactly in separable cases where there
 is a hyperplane that has all class-one points on one side and all
 class-two points on the other.
- **Support vector methods** attempt directly to find a hyperplane in the middle of the gap, that is with maximal margin.
 - ► Such a hyperplane has support vectors, data points that are exactly the margin distance away from the hyperplane.

The problem: Separating hyperplane usually doesn't exist. This difficulty is tackled in two ways.

- We can allow some points to be on the wrong side of their margin subject to a constraint on the total of the 'mis-fit' distances being less than some constant, with Lagrange multiplier C > 0.
- The set of variables is expanded greatly by taking non-linear functions of the original set of variables.
 - ▶ Thus we seek $f(x) = h(x)^T \beta + \beta_0 = 0$ for a vector of $M \gg p$ functions h_i .
 - ▶ Then finding a optimal separating hyperplane is equivalent to solving

$$\min_{\beta_0,\beta} \sum_{i=1}^{n} [1 - y_i f(\mathbf{x}_i)]_+ + \frac{1}{2C} \|\beta\|^2$$

where $y_i = \pm 1$ for the two classes.

There is an implementation of SVMs for R in function svm in package e1071. Here cost is C and gamma is a coefficient of the kernel used to form h.

- ▶ We can try a 10-fold cross-validation
- ▶ The extension to g>2 classes: The svm function uses one attributed to Knerr e al. (1990) in which classifiers are built comparing each pair of classes, and the majority vote amongst the resulting g(g-1)/2 classifiers determines the predicted class.

The forensic glass dataset fgl has 214 points from six classes with nine measurements, and provides a fairly stiff test of classification methods.

- ▶ The types of glass do not form compact well-separated groupings, and the marginal distributions are far from normal.
- ▶ There are some small classes, so we cannot use quadratic discriminant analysis.

- **Performance assessment**: 10 -fold cross-validation, using the same random partition for all the methods.
- Logistic regression provides a suitable benchmark, and in this example linear discriminant analysis does equally well.
- We can use nearest-neighbour methods to estimate the lower bound on the Bayes risk as about 10%
- We need to cross-validate over the choice of tree size, which does vary by group from four to seven.

Neural Networks

- For testing neural network models by V-fold cross-validation, we rescale the dataset so the inputs have range [0,1].
- We want to average across several fits and to choose the number of hidden units and the amount of weight decay by an inner cross-validation.

Neural Networks

- This fits a neural network 1000 times, and so is fairly slow
- An alternative is to use logarithmic scoring.
- Rather than count 0 if the predicted class is correct and 1 otherwise, we count $-\log p(c \mid x)$ for the true class c.

```
> sum(-log(res[cbind(seq(along = truth),
as.numeric(truth))]/nreps))
```

Learning vector quantization

- For LVQ as for k-nearest neighbour methods we have to select a suitable metric.(The rescaled variables or Mahalanobis distance) code
- Our initialization code in lvqinit follows Kohonen's in selecting
 the number of representatives.

One measure that a suitable model for $p(c \mid x)$ has been found is that the predicted probabilities are well calibrated.

• For the forensic glass example we are making six probability forecasts for each case, one for each class. To ensure that they are genuine forecasts, we should use the cross-validation procedure.



Figure 12.9: Calibration plot for multiple logistic fit to the fgl data.

▶ This plot does show substantial over-confidence in the predictions, especially at probabilities dose to one. Indeed, only 22/64 of the events predicted with probability greater than 0.9 occurred.

Where calibration plots are **not straight**, the best solution is to find a better model.

- The over-confidence is minor, and mainly attributable to the use of plug-in rather than predictive estimates.
- Then the plot can be used to adjust the probabilities.