# Modern Applied Statistics Chap 12: Classification

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# Introduction

#### Introduction

In the statistical literature the word is used in two distinct senses.

- The sense of cluster analysis discussed in Section 11.2
- The other meaning (Ripley, 1997) of allocating future cases to one of g prespecified classes

It is sometimes helpful to distinguish discriminant analysis in the sense of describing the differences between the g groups from classification, allocating new observations to the groups.

- The first provides some measure of explanation
- The second can be a 'black box' that makes a decision without any explanation.

Suppose that we have a set of g classes, and for each case we know the class. We can then use the class information to help reveal the structure of the data.

### The sample covariance matrices

$$W = \frac{(X - GM)^T (X - GM)}{n - g}, \quad B = \frac{(GM - 1\bar{x})^T (GM - 1\bar{x})}{g - 1}$$

- W: the within-class covariance matrix
- B: the between-classes covariance matrix
- M: the  $g \times p$  matrix of class means
- G: the  $n \times g$  matrix of class indicator variables
  - Then the predictions are GM
- $\bar{x}$ : the means of the variables over the whole sample.

Note that B has rank at most min(p, g - 1).

Fisher introduced a **linear discriminant analysis** seeking a linear combination xa of the variables that has a maximal ratio of the separation of the class means to the within-class variance.

- ► Maximizing the ratio  $\mathbf{a}^T B \mathbf{a} / \mathbf{a}^T W \mathbf{a}$ 
  - Choose a sphering xS of the variables
  - The problem is to maximize  $\mathbf{a}^T B \mathbf{a}$  subject to  $\|\mathbf{a}\| = 1$
  - This is solved by taking a to be the eigenvector of B corresponding to the largest eigenvalue.
  - The linear combination *a* is unique up to a change of sign.

As for principal components, we can take further linear components corresponding to the next largest eigenvalues.

- **Eigenvalues**: the proportions of the between classes variance explained by the linear combinations.
- The corresponding transformed variables are called the linear discriminants or canonical variates.
- The linear discriminants are conventionally centred to have mean zero on dataset.

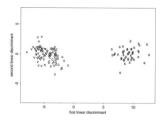


Figure 12.1: The log iris data on the first two discriminant axes.

- ▶ This shows that 99.65% of the between-group variance is on the first discriminant axis.
  - Rao (1948) used the unweighted covariance matrix of the group means.
  - Our approach uses a covariance matrix weighted by the prior probabilities of the classes if these are specified.

### Discrimination for normal populations

An alternative approach to discrimination is via probability models.

▶ Posterior distribution of the classes after observing *x* is

$$p(c \mid \mathbf{x}) = \frac{\pi_c p(\mathbf{x} \mid c)}{p(\mathbf{x})} \propto \pi_c p(\mathbf{x} \mid c)$$

- $\pi_c$ : The prior probabilities of the classes
- $p(x \mid c)$ : The densities of distributions of the observations for each class
- Bayes rule: The allocation rule which makes the smallest expected number of errors chooses the class with maximal p(c | x)

Suppose the distribution for class c is multivariate normal with mean  $\mu_c$  and covariance  $\Sigma_c$ . Then the Bayes rule minimizes

$$Q_c = -2\log p(\mathbf{x} \mid c) - 2\log \pi_c$$
  
=  $(\mathbf{x} - \mu_c) \Sigma_c^{-1} (\mathbf{x} - \mu_c)^T + \log |\Sigma_c| - 2\log \pi_c$ 

- The first term is the squared Mahalanobis distance to the class centre.
- The difference between the  $Q_c$  for two classes is a quadratic function of x.
  - ▶ Quadratic discriminant analysis.
- The boundaries of the decision regions are quadratic surfaces in x space.

Suppose that the classes have a common covariance matrix  $\Sigma$ .

 $\blacktriangleright$  Differences in the  $Q_c$  are then linear functions of x

We can maximize  $-Q_c/2$  or

$$L_c = \mathbf{x} \Sigma^{-1} \boldsymbol{\mu}_c^T - \boldsymbol{\mu}_c \Sigma^{-1} \boldsymbol{\mu}_c^T / 2 + \log \pi_c$$

To use  $Q_c$  or  $L_c$  we have to estimate  $\mu_c$  and  $\Sigma_c$  or  $\Sigma$ .

▶ Using obvious estimates, estimate  $\mu_c$  as the sample mean,  $\Sigma_c$  as covariance matrix, and  $\Sigma$  as W.

How does this relate to Fisher's linear discrimination?

- $ightharpoonup L_c$  gives new variables, the linear discriminants, with unit within-class sample variance.
  - On these variables the Mahalanobis distance is

$$\|x-\mu_c\|^2$$

 $\blacktriangleright$  Only the first r components of the vector depend on c.

$$L_c = \mathbf{x} \boldsymbol{\mu}_c^T - \|\boldsymbol{\mu}_c\|^2 / 2 + \log \pi_c$$

We can work in r dimensions.

$$L_2 - L_1 = \boldsymbol{x} \left( \mu_2 - \mu_1 \right)^T + \text{ const}$$

▶ An affine function of the linear discriminant.

#### Crabs dataset

Construct a rule to predict the sex of a future Leptograpsus crab of unknown colour form.

- Linear discriminant analysis, for what are highly non-normal populations, finds a variable that is essentially CL<sup>3</sup>RW<sup>-2</sup>CW<sup>-1</sup>, a dimensionally neutral quantity.
- Six errors are made, all for the blue form

#### Crabs dataset

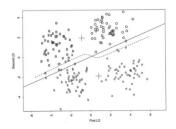


Figure 12.2: Linear discriminants for the crabs data. Males are coded as capitals, females as lower case, colours as the initial letter of blue or orange. The crosses are the group means for a linear discriminant for sex (solid line) and the dashed line is the decision boundary for sex based on four groups.

- ▶ The first two linear discriminants dominate the between-group variation.
- ▶ Using the first two linear discriminants as the data will provide a very good approximation

### Robust estimation of multivariate location and scale

- Apply a robust location estimator to each component of a multivariate mean.
- Consider the estimation of mean and variance simultaneously.
- Two methods for robust covariance estimation
  - ▶ Our function cov.rob ¹
  - ▶ The S-PLUS functions cov.mve and cov.mcd and covRobin in library section robust.
  - MVE(minimum volume ellipsoid method): seeks an ellipsoid containing  $h = \lfloor (n+p+1)/2 \rfloor$  points that is of minimum volume.
  - MCD(minimum covariance determinant method): eeks *h* points whose covariance has minimum determinant.

#### Robust estimation of multivariate location and scale

- The search for an MVE or MCD
  - ▶ provides *h* points whose mean and variance matrix give an initial estimate.
  - ► This is refined by selecting those points whose Mahalanobis distance from the initial mean using the initial covariance is not too large
    - ▶ and returning their mean and variance matrix.
- Alternative approach
  - ▶ Fitting a multivariate  $t_{\nu}$  distribution for a small number  $\nu$  of degrees of freedom.
    - ► This is implemented in our function cov.trob
- Normally cov.trob is faster than cov. rob, but it lacks the latter's extreme resistance.

- In the terminology of pattern recognition
  - ▶ Training set: The given examples together with their classifications.
    - ► Test set: Future cases form.
    - ▶ Our primary measure of success is the error rate.
- The type of errors
  - $\blacktriangleright$  A confusion matrix gives the number of cases with true class i classified as of class j.
    - ▶ We assign costs  $L_{ij}$  to allocating a case of class i to class j.

The average error cost is minimized by the Bayes rule, which is to allocate to the class c minimizing  $\sum_i L_{ic} p(i \mid x)$ .

- p(i | x) is the posterior distribution of the classes after observing x.
- If the costs of all errors are the same, this rule amounts to choosing the class c with the largest posterior probability  $p(c \mid x)$ .

The minimum average cost is known as the Bayes risk.

• Estimate a lower bound for it by the method of Ripley (1996, pp. 196-7).

In Section 12.1,

- How  $p(c \mid x)$  can be computed for normal populations.
- How estimating the Bayes rule with equal error costs leads to linear and quadratic discriminant analysis.

As our functions predict.lda and predict.qda return posterior probabilities.

▶ They can be used for classification with error costs.

The posterior probabilities  $p(c \mid x)$  may also be estimated directly.

- For two classes: we can model  $p(1 \mid x)$  using a logistic regression, fitted by glm.
- For more than two classes: we need a multiple logistic model.
  - ▶ Using a surrogate log-linear Poisson GLM model.
  - ▶ Using the multinom function in library section nnet.

Classification trees model the  $p(c \mid x)$  directly.

► Since the posterior probabilities are given by the predict method it is easy to estimate the Bayes rule for unequal error costs.

### Predictive and 'plug-in' rules

To find the Bayes rule we need to know the posterior probabilities  $p(c \mid x)$ .

- Since these are unknown, we use an parametric family  $p(c \mid \mathbf{x}; \theta)$ .
- We act as if  $p(c \mid x; \hat{\theta})$  were the actual posterior probabilities.
  - $ightharpoonup \hat{ heta}$  is an estimate computed from the training set  $\mathcal{T}$ .
  - ▶ 'plug-in' rule

The 'correct' estimate of  $p(c \mid x)$  is (Ripley, 1996, §2.4) to use the predictive estimates

$$\tilde{p}(c \mid \mathbf{x}) = P(c = c \mid \mathbf{X} = \mathbf{x}, \mathcal{T}) = \int p(c \mid \mathbf{x}; \theta) p(\theta \mid \mathcal{T}) d\theta$$

Non-parametric Rules

# Non-parametric Rules

**Neural Networks** 

# **Neural Networks**

**Support Vector Machine** 

# Support Vector Machine

Forensic Glass Example

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**Calibration Plots** 

# **Calibration Plots**