Part I – Theoretical Foundation

1. (Monte Carlo sampling)

The Central Theorem states that the standard deviation of the finite sample mean \bar{f} is

$$\sqrt{\frac{\overline{f^2} - (\overline{f})^2}{M - 1}}$$

where

$$\bar{f} = \langle f(x) \rangle = \frac{1}{M} \sum_{n=1}^{N} f(r_n)$$

$$\bar{f}^2 = \langle f^2(x) \rangle = \frac{1}{M} \sum_{n=1}^{M} f^2(r_n)$$

$$M = \text{no. of } r \text{and } om \text{ } n \text{umbers}$$

Proof: To find the variance of the sample mean -

$$Var\{\overline{f(x)}\} = Var\left\{\frac{1}{M}\sum_{n=1}^{N}f(r_n)\right\} = \frac{1}{M^2}\sum_{m=1}^{M}\sum_{n=1}^{M}\langle [f(r_m) - \langle f(x)\rangle][f(r_n) - \langle f(x)\rangle]\rangle$$

Assumption – the random variables associated with f are uncorrelated. For any uncorrelated random variables A and B, we have $\langle AB \rangle = \langle A \rangle \langle B \rangle$. Thus,

$$\langle [f(r_m) - \langle f(x) \rangle] [f(r_n) - \langle f(x) \rangle] \rangle = \begin{cases} [f(r_n) - \langle f(x) \rangle]^2 = Var\{f(x)\} & \text{if } m = n \\ \langle f(r_m) - \langle f(x) \rangle \rangle \langle f(r_n) - \langle f(x) \rangle \rangle = 0 & \text{otherwise} \end{cases}$$

which gives

$$Var\{\overline{f(x)}\} = \frac{Var\{f(x)\}}{M}$$

However, there is no way to to obtain a precise value for the population variance $Var\{f(x)\}$. We will define a new variable s_M . Then we will show $\langle s_M \rangle$ estimates $Var\{f(x)\}$.

$$s_M \equiv \frac{1}{M} \sum_{n=1}^{M} f^2(r_n) - \left[\frac{1}{M} \sum_{n=1}^{M} f(r_n) \right]^2$$

Hence,

$$\langle s_{M} \rangle = \frac{1}{M} \sum_{n=1}^{M} \langle f^{2}(x) \rangle - \frac{1}{M^{2}} \sum_{m=1}^{M} \sum_{n=1}^{M} \langle f(r_{m}) f(r_{n}) \rangle$$

$$= \langle f^{2}(x) \rangle - \frac{1}{M^{2}} \sum_{m=1}^{M} \sum_{\substack{n=1 \ n \neq m}}^{M} \langle f(r_{m}) f(r_{n}) \rangle - \frac{1}{M^{2}} \sum_{m=1}^{M} \langle f^{2}(x) \rangle$$

Again, imposing the uncorrelated assumption, we have $\langle f(r_m)f(r_n)\rangle = \langle f(x)\rangle\langle f(x)\rangle$. Therefore,

$$\begin{split} \langle s_M \rangle &= \left(1 - \frac{1}{M} \right) \langle f^2(x) \rangle - \frac{1}{M^2} M \langle f(x) \rangle (M - 1) \langle f(x) \rangle \\ &= \frac{M - 1}{M} (\langle f^2(x) \rangle - \langle f(x) \rangle^2) = \frac{M - 1}{M} Var\{f(x)\} \end{split}$$

Thus,

$$Std\{\overline{f(x)}\} = \sqrt{Var\{\overline{f(x)}\}} = \sqrt{\frac{Var\{f(x)\}}{M}} = \sqrt{\frac{\langle s_M \rangle}{M-1}} = \sqrt{\frac{\overline{f^2} - (\overline{f})^2}{M-1}}$$

2. (Non-uniform random number generation: transformation method)

Since the number of samples in a fixed area under the Box-Muller transformation should remain the same, we have

 $N \times probability\ density \times area = N\ P(r_1, r_2)\ dr_1 dr_2 = N\ P(\zeta_1, \zeta_2)\ S(r_1, r_2, dr_1, dr_2)$

where

$$\zeta_1 = \sqrt{-2\ln r_1} \cos(2\pi r_2)$$

$$\zeta_2 = \sqrt{-2\ln r_1} \sin(2\pi r_2)$$

$$S = \left| \frac{\partial(\zeta_1, \zeta_2)}{\partial(r_1, r_2)} \right| dr_1 dr_2$$

so, by inverting the equations above,

$$r_1 = \exp\left(-\frac{\zeta_1^2 + \zeta_2^2}{2}\right)$$

and

$$P(\zeta_{1}, \zeta_{2}) = \left| \frac{\partial(\zeta_{1}, \zeta_{2})}{\partial(r_{1}, r_{2})} \right|^{-1}$$

$$= \left| \frac{\partial\zeta_{1}}{\partial r_{1}} \frac{\partial\zeta_{2}}{\partial r_{2}} - \frac{\partial\zeta_{1}}{\partial r_{2}} \frac{\partial\zeta_{2}}{\partial r_{1}} \right|^{-1}$$

$$= \left| -\frac{r_{1}^{-1} \cos(2\pi r_{2})}{\sqrt{-2 \ln r_{1}}} 2\pi \sqrt{-2 \ln r_{1}} \cos(2\pi r_{2}) - 2\pi \sqrt{-2 \ln r_{1}} \sin(2\pi r_{2}) \frac{r_{1}^{-1} \sin(2\pi r_{2})}{\sqrt{-2 \ln r_{1}}} \right|^{-1}$$

$$= \left| -\frac{2\pi}{r_{1}} \right|^{-1} = \frac{r_{1}}{2\pi} = \frac{1}{2\pi} \exp\left(-\frac{\zeta_{1}^{2} + \zeta_{2}^{2}}{2}\right)$$

$$= \sqrt{\frac{1}{2\pi}} \exp\left(-\frac{\zeta_{1}^{2}}{2}\right) \sqrt{\frac{1}{2\pi}} \exp\left(-\frac{\zeta_{2}^{2}}{2}\right) = P(\zeta_{1})P(\zeta_{2})$$

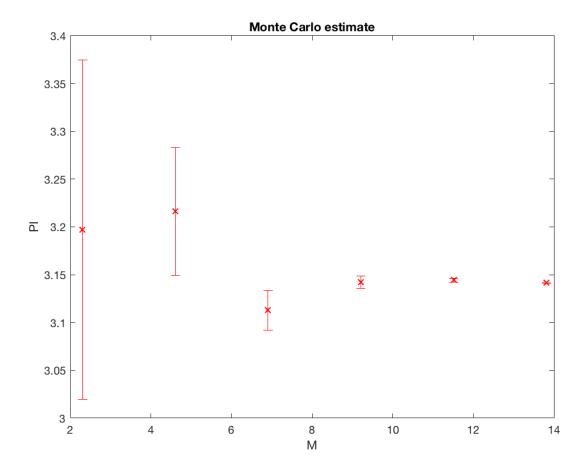
i.e. separable.

$$P(\zeta_{1 \text{ or } 2}) = \sqrt{\frac{1}{2\pi}} \exp\left(-\frac{\zeta_{1 \text{ or } 2}^2}{2}\right)$$

Therefore, the Box-Muller algorithm generates (two) normally distributed random number(s) $\zeta_{1 \text{ or } 2}$ with unit variance.

Plot on next page.

```
Part II – Numerical Test
1. (Monte Carlo estimate)
Source code:
/* Monte Carlo integration of PI by sample mean */
#include <stdio.h>
#include <stdlib.h>
#include <time.h>
#include <math.h>
int main() {
  double x, pi, sum = 0.0, pi2, sum2 = 0.0, stdv, fx;
  int try, ntry;
  printf("Input the number of MC trials\n");
  scanf("%d",&ntry);
  srand((unsigned)time((long *)0));
  for (try=0; try<ntry; try++) {
    x = rand()/(double)RAND MAX;
    fx = 4.0/(1.0 + x*x);
    sum += fx;
    sum2 += fx*fx;
  }
  pi = sum/ntry;
  pi2 = sum2/ntry;
  stdv = sqrt((pi2-pi*pi)/(ntry-1));
  printf("MC estimate for PI = \%f += \%e \ n", pi, stdv);
  return 0;
Results:
M = [10\ 10^2\ 10^3\ 10^4\ 10^5\ 10^6];
PI = [3.197029 \ 3.216067 \ 3.112710 \ 3.142317 \ 3.144274 \ 3.141827];
STDV = [1.776670e-01 6.674424e-02 2.055114e-02 6.455923e-03 2.033267e-03 6.433500e-04];
```



2. (Monte Carlo error)

```
Source code:
/* Monte Carlo integration of PI by sample mean */
#include <stdio.h>
#include <stdlib.h>
#include <time.h>
#include <math.h>
#define NSEED 100
int main() {
  double x, pi, sum = 0.0, pi av = 0.0, pi2 av = 0.0, stdv;
  int try, ntry, outer;
  printf("Input the number of MC trials\n");
  scanf("%d",&ntry);
  srand((unsigned)time((long *)0));
  for (outer=1; outer <= NSEED; outer++) {
    sum = 0.0;
    for (try=0; try<ntry; try++) {
    x = rand()/(double)RAND_MAX;
    sum += 4.0/(1.0 + x*x);
    pi = sum/ntry;
    pi av += pi;
    pi2 av += pi*pi;
  pi av /= NSEED;
  pi2 av /= NSEED;
  stdv = sqrt(pi2 av-pi av*pi av);
  printf("Ntry = %d: Stdv estimate for PI = %e\n", ntry, stdv);
  return 0;
Results:
M = [10\ 10^2\ 10^3\ 10^4\ 10^5\ 10^6];
sM = [1.964829e-01\ 5.781135e-02\ 2.187134e-02\ 5.564133e-03\ 1.788560e-03\ 6.341625e-04];
Estimated value of power: -0.5022
```

Plot on next page.

