

# Econ 142: HW 3

Sofia Guo

2/8/2019

## 1(a). Question

Show that if  $x_i$  contains a constant, then  $\bar{y} = \bar{x}'\hat{\beta}$ , where  $\bar{y} = \frac{1}{N} \sum_{i=1}^N y_i$  and  $\bar{x} = \frac{1}{N} \sum_{i=1}^N x_i$

## 1(a) Proof:

Under implications of defining properties of the PRF, one FOC is  $E[x_i \hat{u}_i] = 0$ . Thus:

$\Rightarrow E[y_i] = E[x_i' \hat{\beta} + \hat{u}_i]$  by taking the expected value of the sample regression equation  $y_i = x_i' \hat{\beta} + \hat{u}_i$

$\Rightarrow E[y_i] = E[x_i' \hat{\beta}] + E[\hat{u}_i]$  by distributing the expectation across sums

$\Rightarrow E[y_i] = E[x_i' \hat{\beta}] + 0$  from the sample FOC  $E[x_i \hat{u}_i] = \frac{1}{N} \sum_{i=1}^N x_i \hat{u}_i = 0$

$\Rightarrow \frac{1}{N} \sum_{i=1}^N y_i = \frac{1}{N} \sum_{i=1}^N x_i' \hat{\beta}$  from taking the expectation

$\therefore \bar{y} = \bar{x}' \hat{\beta}$  using the definitions  $\bar{y} = \frac{1}{N} \sum_{i=1}^N y_i$  and  $\bar{x} = \frac{1}{N} \sum_{i=1}^N x_i$  ■

## 1(b). Question

Show that if  $x_i$  contains a dummy variable for membership in group  $g$  (which has  $N_g$  observations in the sample) then  $\bar{y}_g = \bar{x}_g' \hat{\beta}$  where  $\bar{y}_g = \frac{1}{N_g} \sum_{i \in g} y_i$ , and  $\bar{x}_g = \frac{1}{N_g} \sum_{i \in g} x_i$ .

## 1(b) Proof:

The FOC require  $\sum_{i=1}^N D_i(y_i - x_i' \hat{\beta}) = 0 = \sum_{i \in g} (y_i - x_i' \hat{\beta})$ . Thus:

$\Rightarrow \frac{1}{N_g} \sum_{i \in g} (y_i - x_i' \hat{\beta}) = 0$  just by taking the expected value

$\Rightarrow \frac{1}{N_g} \sum_{i \in g} y_i = \frac{1}{N_g} \sum_{i \in g} x_i' \hat{\beta}$  distributing the sums and bringing second term over

$\therefore \bar{y}_g = \bar{x}_g' \hat{\beta}$  since  $\bar{y}_g = \frac{1}{N_g} \sum_{i \in g} y_i$ , and  $\bar{x}_g = \frac{1}{N_g} \sum_{i \in g} x_i$  ■

## 1(c). Question

Complete the proof of the Frisch-Waugh Theorem for the sample OLS regression coefficients by showing that the  $j^{th}$  row of  $\hat{\beta}$  is:

$$\hat{\beta}_j = \left[ \frac{1}{N} \sum_{i=1}^N \hat{\xi}_i^2 \right]^{-1} \left[ \frac{1}{N} \sum_{i=1}^N \hat{\xi}_i y_i \right] \quad (1)$$

## 1(c) Proof:

$\Rightarrow$  We know from the FOC for the sample regression that  $\frac{1}{N} \sum_{i=1}^N x_i \hat{u}_i = 0$

$\Rightarrow$  Then define  $\hat{u}_i = y_i - x_i' \hat{\beta}$  from the sample OLS regression  $y_i = x_i' \hat{\beta} + \hat{u}_i$

$\Rightarrow$  We also know from the FOC that  $\frac{1}{N} \sum_{i=1}^N x_{(\sim j)i} \hat{\xi}_i = 0$   
 $\Rightarrow$  Then define  $\hat{\xi}_i = x_{ji} - x_{(\sim j)i} \hat{\pi}$  from the auxiliary regression  $x_{ji} = x_{(\sim j)i} \hat{\pi} + \hat{\xi}_i$   
 $\Rightarrow$  We know that the sample regression equation is  $y_i = \hat{\beta}_1 x_{1i} + \hat{\beta}_2 x_{2i} + \cdots + \hat{\beta}_j x_{ji} + \cdots + \hat{\beta}_K x_{Ki} + \hat{u}_i$   
 $\Rightarrow$  Thus we can write

$$\frac{1}{N} \sum_{i=1}^N \hat{\xi}_i y_i = \frac{1}{N} \sum_{i=1}^N \hat{\xi}_i (\hat{\beta}_1 x_{1i} + \hat{\beta}_2 x_{2i} + \cdots + \hat{\beta}_j x_{ji} + \cdots + \hat{\beta}_K x_{Ki} + \hat{u}_i) \quad (2)$$

from the previous step

$\Rightarrow$  Since  $\frac{1}{N} \sum_{i=1}^N x_{(\sim j)i} \hat{\xi}_i = 0$  we know that  $\hat{\xi}_i \perp x_{(\sim j)i}$   
 $\Rightarrow$  Since  $\hat{\xi}_i = x_{ji} - x_{(\sim j)i} \hat{\pi}$  we know that  $\hat{\xi}_i \perp \hat{u}_i$  because  $\frac{1}{N} \sum_{i=1}^N x_i \hat{u}_i = 0$  ( $x_i \perp \hat{u}_i$ )  
 $\Rightarrow$  Thus we know that all the terms in 2 except  $\hat{\beta}_j x_{ji}$  equal 0; Thus:  
 $\Rightarrow \frac{1}{N} \sum_{i=1}^N \hat{\xi}_i y_i = \hat{\beta}_j \frac{1}{N} \sum_{i=1}^N x_{ji}$  by simplification  
 $\Rightarrow$  Using  $x_{ji} = x_{(\sim j)i} \hat{\pi} + \hat{\xi}_i$  we substitute into  $\frac{1}{N} \sum_{i=1}^N \hat{\xi}_i x_{ji} = \frac{1}{N} \sum_{i=1}^N \hat{\xi}_i (x_{(\sim j)i} \hat{\pi} + \hat{\xi}_i) = \frac{1}{N} \sum_{i=1}^N \hat{\xi}_i^2$  because  $\hat{\xi}_i \perp x_{(\sim j)i}$  using the FOC for  $\hat{\pi}$   
 $\Rightarrow \frac{1}{N} \sum_{i=1}^N \hat{\xi}_i y_i = \hat{\beta}_j \frac{1}{N} \sum_{i=1}^N \hat{\xi}_i^2$  by substitution  
 $\therefore \hat{\beta}_j = [\frac{1}{N} \sum_{i=1}^N \hat{\xi}_i^2]^{-1} [\frac{1}{N} \sum_{i=1}^N \hat{\xi}_i y_i]$  ■.

## 2. OVB Dataset

```

#load libraries
library(dplyr)
library(ggplot2)
library(magrittr)
library(xlsx)
library(reshape2)
library(stargazer)
library(lubridate)
library(lmtest)
library(ivpack)
library(kableExtra)

#import dataset
ovb_raw <- read.csv("/Users/sofia/Box/Cal (sofiagu@berkeley.edu)/2018-19/Spring 2019/Econ 142/PSETS/PS1")
  
```

### 2(a). Question:

Write an expression for the OLS estimate of the coefficient on immigrant statues from logwage = constant, immigrant statues, if the true model is logwage = constant, education, immigrant status.

### 2(a) Solution:

$$\hat{\beta}_{imm}^0 = \hat{\beta}_{imm} + \hat{\pi}_2 \hat{\beta}_{educ} \quad (3)$$

where  $\hat{\pi}_2$  is the coefficient on immigration from regressing immigration on education (the omitted variable):

$$education_i = \hat{\pi}_1 + \hat{\pi}_2 immigration_i + \hat{\xi}_i \quad (4)$$

## 2(b). Question:

Estimate the 5 models and show values for (a), first female then male.

1. logwage = constant, immigrant status
2. logwage = constant, education
3. immigrant status = constant, education
4. education = constant, immigrant status
5. logwage = constant, education, immigrant status

## 2(b) Solution:

Values for the female regressions:

$$\hat{\pi}_{2_{female}} = -1.49214$$

$$\hat{\beta}_{imm_{female}} = -0.179986$$

$$\hat{\beta}_{educ_{female}} = 0.113853$$

$$\hat{\beta}_{imm_{female}}^0 = -0.179986 + (-1.49214) * (0.113853) = -0.3498706$$

Values for the male regressions:

$$\hat{\pi}_{2_{male}} = -1.61176$$

$$\hat{\beta}_{imm_{male}} = -0.24477$$

$$\hat{\beta}_{educ_{male}} = 0.105620$$

$$\hat{\beta}_{imm_{male}}^0 = -1.61176 + (-0.24477) * (0.105620) = -1.637613$$

Code:

```
#estimate model 1 for both genders

#filter females only
fem <- ovb_raw %>%
  dplyr::filter(female == 1)

#filter males only
male <- ovb_raw %>%
  dplyr::filter(female == 0)

#run regression 1 for females
reg_fem_1 <- summary(lm(logwage ~ imm, data = fem))
reg_fem_1

##
## Call:
## lm(formula = logwage ~ imm, data = fem)
##
## Residuals:
```

```
##      Min      1Q  Median      3Q      Max
## -1.5001 -0.4407 -0.0206  0.4066  3.2851
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  2.886378   0.007154  403.48  <2e-16 ***
## imm         -0.179986   0.016532  -10.89  <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.664 on 10599 degrees of freedom
## Multiple R-squared:  0.01106,    Adjusted R-squared:  0.01097
## F-statistic: 118.5 on 1 and 10599 DF,  p-value: < 2.2e-16
```

```
#run regression 1 for males
```

```
reg_male_1 <- summary(lm(logwage ~ imm, data = male))
reg_male_1
```

```
##
## Call:
## lm(formula = logwage ~ imm, data = male)
##
## Residuals:
##      Min      1Q  Median      3Q      Max
## -1.76962 -0.42464 -0.00445  0.41578  3.08032
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  3.15592    0.00728  433.51  <2e-16 ***
## imm         -0.24477    0.01558  -15.71  <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.6844 on 11304 degrees of freedom
## Multiple R-squared:  0.02136,    Adjusted R-squared:  0.02127
## F-statistic: 246.7 on 1 and 11304 DF,  p-value: < 2.2e-16
```

```
#estimate model 2 for both genders
```

```
#run regression 2 for females
```

```
reg_fem_2 <- summary(lm(logwage ~ educ, data = fem))
reg_fem_2
```

```
##
## Call:
## lm(formula = logwage ~ educ, data = fem)
##
## Residuals:
##      Min      1Q  Median      3Q      Max
## -2.1316 -0.3413  0.0034  0.3555  3.3868
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  1.234852   0.029811  41.42  <2e-16 ***
## educ         0.114153   0.002064  55.30  <2e-16 ***
```

```
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.5882 on 10599 degrees of freedom
## Multiple R-squared:  0.2239, Adjusted R-squared:  0.2238
## F-statistic: 3058 on 1 and 10599 DF,  p-value: < 2.2e-16

#run regression 2 for males
reg_male_2 <- summary(lm(logwage ~ educ, data = male))
reg_male_2

##
## Call:
## lm(formula = logwage ~ educ, data = male)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -2.38111 -0.35217  0.01881  0.35307  3.08724
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  1.609458   0.027146   59.29  <2e-16 ***
## educ         0.107897   0.001917   56.28  <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.6115 on 11304 degrees of freedom
## Multiple R-squared:  0.2189, Adjusted R-squared:  0.2188
## F-statistic: 3167 on 1 and 11304 DF,  p-value: < 2.2e-16

#estimate model 3 for both genders

#run regression 3 for females
reg_fem_3 <- summary(lm(imm ~ educ, data = fem))
reg_fem_3

##
## Call:
## lm(formula = imm ~ educ, data = fem)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -0.51850 -0.22201 -0.13306 -0.07376  0.98554
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  0.607445   0.019329   31.43  <2e-16 ***
## educ        -0.029649   0.001339  -22.15  <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.3814 on 10599 degrees of freedom
## Multiple R-squared:  0.04424, Adjusted R-squared:  0.04415
## F-statistic: 490.6 on 1 and 10599 DF,  p-value: < 2.2e-16
```

```

#run regression 3 for males
reg_male_3 <- summary(lm(imm ~ educ, data = male))
reg_male_3

##
## Call:
## lm(formula = imm ~ educ, data = male)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -0.54939 -0.27436 -0.15213 -0.09101  0.97011
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  0.641066   0.017880   35.85  <2e-16 ***
## educ        -0.030559   0.001263  -24.20  <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.4028 on 11304 degrees of freedom
## Multiple R-squared:  0.04925,    Adjusted R-squared:  0.04917
## F-statistic: 585.6 on 1 and 11304 DF,  p-value: < 2.2e-16

```

```

#estimate model 4 for both genders

```

```

#run regression 4 for females
reg_fem_4 <- summary(lm(educ ~ imm, data = fem))
reg_fem_4

##
## Call:
## lm(formula = educ ~ imm, data = fem)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -12.9597  -1.4518  -0.4518   1.5482   7.0403
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  14.45183   0.02915  495.76  <2e-16 ***
## imm         -1.49214   0.06737  -22.15  <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 2.706 on 10599 degrees of freedom
## Multiple R-squared:  0.04424,    Adjusted R-squared:  0.04415
## F-statistic: 490.6 on 1 and 10599 DF,  p-value: < 2.2e-16

```

```

#run regression 2 for males
reg_male_4 <- summary(lm(educ ~ imm, data = male))
reg_male_4

```

```

##
## Call:
## lm(formula = educ ~ imm, data = male)

```

```
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -12.5776  -2.1894  -0.5776   1.8106   7.4224
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 14.18939    0.03111   456.1  <2e-16 ***
## imm         -1.61176    0.06660   -24.2  <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 2.925 on 11304 degrees of freedom
## Multiple R-squared:  0.04925, Adjusted R-squared:  0.04917
## F-statistic: 585.6 on 1 and 11304 DF, p-value: < 2.2e-16
```

```
#estimate model 5 for both genders
```

```
#run regression 5 for females
```

```
reg_fem_5 <- summary(lm(logwage ~ educ + imm, data = fem))
reg_fem_5
```

```
##
## Call:
## lm(formula = logwage ~ educ + imm, data = fem)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -2.1318  -0.3439   0.0038   0.3552   3.3943
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  1.240988    0.031170   39.814  <2e-16 ***
## educ         0.113853    0.002112   53.915  <2e-16 ***
## imm        -0.010101    0.014981   -0.674    0.5
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.5883 on 10598 degrees of freedom
## Multiple R-squared:  0.2239, Adjusted R-squared:  0.2238
## F-statistic: 1529 on 2 and 10598 DF, p-value: < 2.2e-16
```

```
#run regression 2 for males
```

```
reg_male_5 <- summary(lm(logwage ~ educ + imm, data = male))
reg_male_5
```

```
##
## Call:
## lm(formula = logwage ~ educ + imm, data = male)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -2.38334  -0.36121   0.01482   0.35657   3.14132
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
```

```
## (Intercept)  1.657239    0.028614   57.917   < 2e-16 ***
## educ        0.105620    0.001964   53.780   < 2e-16 ***
## imm        -0.074534    0.014263   -5.226   1.77e-07 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.6108 on 11303 degrees of freedom
## Multiple R-squared:  0.2208, Adjusted R-squared:  0.2206
## F-statistic: 1601 on 2 and 11303 DF,  p-value: < 2.2e-16
```

## 2(c). Question:

Redo 5 models for Females and Males, distinguishing 3 groups of immigrants. Put and include results in 2 new tables similar to the table in lecture 5.

```
#add dummies for asian, hispanic and other immigrants

#make female data frame
ovb_race_imm_fem <- ovb_raw %>%
  mutate(asian_imm = as.numeric(asian == 1 & imm ==1 & hispanic ==0),
         hispanic_imm = as.numeric(hispanic == 1 & imm ==1 & asian ==0),
         other_imm = as.numeric(imm ==1 & asian ==0 & hispanic==0)) %>%
  dplyr::filter(female ==1)

#make male data frame
ovb_race_imm_male <- ovb_raw %>%
  mutate(asian_imm = as.numeric(asian == 1 & imm ==1 & hispanic ==0),
         hispanic_imm = as.numeric(hispanic == 1 & imm ==1 & asian ==0),
         other_imm = as.numeric(imm ==1 & asian ==0 & hispanic==0)) %>%
  dplyr::filter(female ==0)

#run regression 1 for females
reg_fem_1_imm <- lm(logwage ~ imm + asian_imm + hispanic_imm + other_imm, data = ovb_race_imm_fem)

#run regression 1 for males
reg_male_1_imm <- lm(logwage ~ imm + asian_imm + hispanic_imm + other_imm, data = ovb_race_imm_male)

#estimate model 2 for both genders

#run regression 2 for females
reg_fem_2_imm <- lm(logwage ~ educ + asian_imm + hispanic_imm + other_imm, data = ovb_race_imm_fem)

#run regression 2 for males
reg_male_2_imm <- lm(logwage ~ educ + asian_imm + hispanic_imm + other_imm, data = ovb_race_imm_male)

#estimate model 3 for both genders

#run regression 3 for females
reg_fem_3_imm <- lm(imm ~ educ + asian_imm + hispanic_imm + other_imm, data = ovb_race_imm_fem)

#run regression 3 for males
reg_male_3_imm <- lm(imm ~ educ + asian_imm + hispanic_imm + other_imm, data = ovb_race_imm_male)

#estimate model 4 for both genders
```



```

#run regression 4 for females
reg_fem_4_imm <- lm(educ ~ imm + asian_imm + hispanic_imm + other_imm, data = ovb_race_imm_fem)

#run regression 4 for males
reg_male_4_imm <- lm(educ ~ imm + asian_imm + hispanic_imm + other_imm, data = ovb_race_imm_male)

#estimate model 5 for both genders

#run regression 5 for females
reg_fem_5_imm <- lm(logwage ~ educ + imm + asian_imm + hispanic_imm + other_imm, data = ovb_race_imm_fem)

#run regression 5 for males
reg_male_5_imm <- lm(logwage ~ educ + imm + asian_imm + hispanic_imm + other_imm, data = ovb_race_imm_m

```

2(c) Tables:

```
stargazer(reg_fem_1_imm,
          reg_fem_2_imm,
          reg_fem_3_imm,
          reg_fem_4_imm,
          reg_fem_5_imm,
          type = "latex", title = "Female Immigrant Regression
Results",
          header = F,
          multicolumn = F,
          column.sep.width = '0.1pt',
          single.row = T,
          omit.stat = c("f", "ser"))
```

Table 1: Female Immigrant Regression Results

	<i>Dependent variable:</i>				
	logwage (1)	logwage (2)	imm (3)	educ (4)	logwage (5)
imm	0.106 (0.218)			−2.341*** (0.859)	0.368* (0.196)
educ		0.112*** (0.002)	−0.0003*** (0.0001)		0.112*** (0.002)
asian_imm	−0.020 (0.220)	0.029 (0.027)	0.999*** (0.001)	2.851*** (0.866)	−0.339* (0.198)
hispanic_imm	−0.543** (0.219)	−0.054** (0.021)	0.998*** (0.001)	−1.092 (0.863)	−0.421** (0.197)
other_imm	−0.057 (0.220)	0.016 (0.028)	0.999*** (0.001)	2.624*** (0.867)	−0.351* (0.198)
Constant	2.886*** (0.007)	1.269*** (0.033)	0.005*** (0.002)	14.452*** (0.028)	1.267*** (0.033)
Observations	10,601	10,601	10,601	10,601	10,601
R <sup>2</sup>	0.038	0.224	0.994	0.134	0.225
Adjusted R <sup>2</sup>	0.038	0.224	0.994	0.134	0.224

Note:

\*p<0.1; \*\*p<0.05; \*\*\*p<0.01

```
stargazer(reg_male_1_imm,
          reg_male_2_imm,
          reg_male_3_imm,
          reg_male_4_imm,
          reg_male_5_imm,
          type = "latex", title = "Male Immigrant Regression
Results",
          header = F,
          multicolumn = F,
          column.sep.width = '0.1pt',
          single.row = T,
          omit.stat = c("f", "ser"))
```

Table 2: Male Immigrant Regression Results

	<i>Dependent variable:</i>				
	logwage (1)	logwage (2)	imm (3)	educ (4)	logwage (5)
imm	−0.141 (0.195)			−2.273*** (0.787)	0.097 (0.177)
educ		0.105*** (0.002)	−0.0003*** (0.0001)		0.105*** (0.002)
asian_imm	0.213 (0.197)	−0.060** (0.027)	0.999*** (0.001)	3.529*** (0.795)	−0.156 (0.178)
hispanic_imm	−0.331* (0.195)	−0.092*** (0.019)	0.997*** (0.001)	−1.366* (0.789)	−0.188 (0.177)
other_imm	0.155 (0.197)	−0.057** (0.027)	0.999*** (0.001)	2.947*** (0.795)	−0.154 (0.178)
Constant	3.156*** (0.007)	1.671*** (0.031)	0.006*** (0.002)	14.189*** (0.029)	1.671*** (0.031)
Observations	11,306	11,306	11,306	11,306	11,306
R <sup>2</sup>	0.051	0.221	0.994	0.176	0.221
Adjusted R <sup>2</sup>	0.051	0.221	0.994	0.176	0.221

*Note:*

\*p<0.1; \*\*p<0.05; \*\*\*p<0.01