

# Stations, trains and small-world networks

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The clustering coefficient, path length and average vertex degree of two urban train line networks have been calculated. The results are compared with theoretical predictions for appropriate random bipartite graphs. They have also been compared with one another to investigate the effect of architecture on the small-world properties.

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## I. INTRODUCTION

Since model small-world networks were first proposed some five years ago [1, 2], to interpolate between the properties of regular and random graphs, many investigations have been carried out which have confirmed that the small-world phenomenon occurs in a variety of real-world settings.

There are two common features which make apparently very different networks all “small-world” [1]. First, the *characteristic path length* between vertices is short, compared to what might be expected based on the total number of vertices. This is the phenomenon which has entered popular culture as “six degrees of separation”. The second feature, the one we are noting when we agree with a new acquaintance that the world is indeed small, is that two vertices both linked to a third vertex are likely to have a direct link also. This feature has been called *clustering*, and while random graphs have the first feature, they do not have the second. Comparisons of real-world networks with random graphs having the same number of vertices and edges indicate much higher clustering occurs in real, small worlds. In a recent review [3] these real-world networks are roughly grouped as social, biological, technological and information networks.

Among the first social networks to be studied in this context are the Hollywood network of actors [1, 2], linked by having appeared in the same movie, and the collaboration network of scientists [4], linked by having published together. By considering not just the people involved but also the media that created their links, the network can be seen to have a *bipartite* structure [5], because nodes of one kind (actors/authors) can only be directly linked to nodes of the other kind (movies/papers). In contrast to the usual random graph, unipartite or one-mode projections of random bipartite graphs onto vertices of the one kind inherit some amount of in-built clustering, from the links to common vertices of the other type [5]. It is thus more appropriate to compare the path length and clustering of such a real world network to the averaged properties of random bipartite graphs with the same number

of vertices of each kind and the same probability distribution of links from one to the other type [5, 6].

The study [7] reported here was motivated by two quite different investigations into small-world properties of railway networks, one type of technological distribution network. The first study (of the Boston subway network) [8] takes as its definition of an edge that a physical piece of track should connect two vertices (stations). This presents immediate difficulties at the end of a line, where the concept of a local clustering coefficient becomes moot and the definition (see (3) below) undefined. (To resolve this problem, the authors propose that different properties should be studied; rather than clustering coefficient and average or greatest path length, they define local and global *efficiency*, and *cost* [8, 9]. Other authors have chosen not to include such vertices in their calculation.)

In the second study [10], of part of the railway network of India, two stations are defined to be linked if it is possible to embark at the first and reach the second without having to change trains, regardless of the number of intermediate stations, i.e., they lie on the same line or route. This definition removes the “end of the line” problem noted above. However, it imposes the same kind of bipartite structure [5] as is present in the original study of the Hollywood actors network [1, 2], and studies of the collaboration network of scientists [4, 5] and of Fortune 1000 directors (sit together on the board of a company) [5, 11]. However, this bipartite structure was not referred to in the analysis provided in [10], unlike the somewhat more whimsical study of the (artificial) Marvel comic book universe [6], our third motivation.

Therefore, we wanted to investigate railway networks from this bipartite viewpoint, comparing the properties of networks having different architecture with each other, and with those of the appropriate random bipartite model. The two networks we have chosen are Boston [16], because of the earlier study, and Vienna [17] because of different features in its layout. In the following section we give the definitions of the network properties we calculate. In Section III A we explain features of the systems that we studied and give our results. The paper closes with discussion and a some conclusions.

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## II. NETWORK PROPERTIES

### A. Definitions

There are two different definitions of the measure of cliquishness or clustering in a network [3, 12] of  $N$  nodes, which attempt to quantify the same property (but which have not always been clearly distinguished). They both take values in the interval  $[0, 1]$ . The first definition [13], which is convenient for analytic calculations, corresponds to what sociologists refer to as transitivity, and will be denoted by  $C^{(T)}$ :

$$C^{(T)} = \frac{3 \times \text{number of triangles in the network}}{\text{number of connected triples of vertices}}. \quad (1)$$

The second definition is easier to apply to an actual network, i.e., numerically. Chronologically, this definition preceded (1) in the small-world literature [1]. If node  $v$  has  $k_v$  neighbors, there are  $\binom{k_v}{2}$  possible connections between them. The local clustering coefficient of node  $v$  is the fraction of these that are actually present:

$$C_v = \frac{2 \times \text{number of links between neighbors of } v}{k_v(k_v - 1)} \quad (2)$$

and then the clustering coefficient of the whole network is the average of the  $C_v$ :

$$C = \frac{1}{N} \sum_v C_v. \quad (3)$$

The number of other vertices to which a vertex  $v$  is linked is its *degree*  $k_v$  and the average degree for the network is  $z$ . If there are  $K$  edges in the network, clearly

$$K = \frac{Nz}{2} \quad (4)$$

since each edge adjoins two of the vertices.

Except in limiting cases, it is probably not important to make too much of the differences between (1) and (3); after all, there are several different measures of network size in common usage. Define the *distance*  $d_{ij}$  between two nodes  $i, j$  to be the length (in number of edges) of the shortest path between them. The *diameter*  $D$  of a network is then the greatest distance between any two randomly chosen nodes. The *average path length*  $L$  is the average of the distances  $d_{ij}$  between all pairs of nodes:

$$L = \frac{\sum_{i \neq j} d_{ij}}{N \cdot (N - 1)}. \quad (5)$$

Analytically, it proves useful to define [5] a *typical path length*  $\ell$  which is determined by finding when the sum of the average number of first  $z_1 (= z)$ , second  $z_2$ , ...,  $z_\ell$  neighbors of a vertex reaches  $N$ :

$$1 + \sum_{i=1}^{\ell} z_i = N. \quad (6)$$

Despite their differences, these various measures of network size scale as  $\log(N)$  for a random graph; since there is no preferential attachment, the clustering coefficient of a random graph (by either measure) is of order  $C_{\text{random}} = O(N^{-1})$  [3]. More precisely, it is a result of some twenty year's standing in random graph theory that  $\ell = \log(N)/\log(z)$ , while the clustering coefficient was shown [2] to be

$$C_{\text{random}} = \frac{2K}{N(N-1)}. \quad (7)$$

The conditions for a network to be small-world are that the characteristic path length is comparable to that of a random graph (with the same  $N$  and  $z$ ) while it displays significantly more clustering.

### B. Formulae for random bipartite networks

For a bipartite network, the number of nodes  $N$  refers to the vertices of the one-mode projection (e.g., the actors or, in this case, the stations). The number of nodes of the other type (train lines) will be denoted  $M$ , the average number of train lines a station lies on is  $\mu$  and the average number of stations per line is  $\nu$  so that

$$N\mu = M\nu. \quad (8)$$

The number of links,  $K$ , refers to the unipartite graph.

Newman et al. [5] have given expressions for average values of the network properties defined above, for various distributions of vertex degree, in terms of their generating functions. In the case of bipartite graphs, the appropriate expressions are summarized below.

For specific real-world graphs, such as these two train networks, one knows the degree of each vertex, so that the generating functions are polynomials. Let  $p_j$  be the probability that a station appears on  $j$  train lines, and  $q_k$  be the probability that a train line has  $k$  stations on it. Then the two generating functions needed are:

$$f_0(x) = \sum_j p_j x^j, \quad g_0(x) = \sum_k q_k x^k. \quad (9)$$

Defining further generating functions in terms of these (their interpretation is given in [5] but is not needed for calculations)

$$f_1(x) = \frac{1}{\mu} f'_0(x), \quad g_1(x) = \frac{1}{\nu} g'_0(x), \quad (10)$$

then the number of first and second neighbors of a station chosen at random are

$$z_1 = f'_0(1)g'_1(1), \quad (11a)$$

$$z_2 = f'_0(1)f'_1(1)[g'_1(1)]^2; \quad (11b)$$

the typical path length (6) is

$$\ell = \frac{\log[(N-1)(z_2 - z_1) + z_1^2] - \log z_1^2}{\log(z_2/z_1)}; \quad (12)$$

and the clustering coefficient (3) is

$$C = \frac{M}{N} \frac{g_0'''(1)}{G_0''(1)}. \quad (13)$$

Here  $G_0(x)$  has interpretation as the generating function of the probability distribution of number of first neighbors (i.e., degree) on the unipartite graph of stations, and is defined by

$$G_0(x) = f_0(g_1(x)). \quad (14)$$

Of course, the role of the nodes could be interchanged, i.e., a projection made onto a one-mode graph of train lines rather than stations. In the case of company directors, studying the *interlock* of Fortune 1000 companies (rather than degree distribution for the network of directors) has provided explanation for the adoption of common corporate policies and practices [11].

### III. THE SMALL WORLDS OF BOSTON AND VIENNA

#### A. This study: stations and trains

In early sociological studies of networks, getting a good definition of connection was sometimes problematic [2]. While for technological networks this is generally less of a problem, recall the definition of a link between two stations used in this study: that one can travel between them without changing train. The railway networks were treated as undirected and unweighted; if more than one train line can be used to travel from a particular station to a second, with no change of train, in either direction, this counted as a connection. Further, no weighting was made for timetable information, such as the number of trains running on a line per day or whether trains ever “run express”.

The subway network of Boston consists of four lines, serving 124 stations [18]. The network is decentralized, in that no station lies on all four lines, thus having every other station as a direct neighbor with our definition of adjacency. In fact, the highest value of  $k_v$  is 97. The diameter of the network is  $D = 3$ . One line splits into two branches and another into four. Consistent with our definition of connectivity, these were treated as separate lines, giving eight in total, some of which have a series of stations in common.

There are five lines serving 76 stations in the U-Bahn network of Vienna, which is also decentralized. These lines interact with each other at single stations, though possibly more than once due to their tendency to follow curved (rather than radial) paths. The highest degree is  $k_v = 43$  and the diameter is also 3.

From the network maps, the value of  $d_{ij}$  for each pair of stations was entered into a spreadsheet. Using various spreadsheet functions, the degree of each station  $k_v$ , the total number of links  $K$ , the average degree  $z = z_1$ ,

the clustering coefficient  $C$  as defined in (3), and the average path length  $L$  defined in (5) were calculated [7]. By counting the instances when  $d_{ij} = 2, 3$ , the average number of second  $z_2$  and third  $z_3$  neighbors could also be determined. These ‘Actual’ values appear in Table I.

Also from the maps, the discrete probability distributions  $p_j$  and  $q_k$  required for using the generating functions (9) were identified. These appear in Tables II and III. The mean properties of random bipartite graphs with the same distributions of links from one type of node to another,  $z_1, z_2, \ell$  and  $C^{(T)}$ , were calculated using equations (11)–(13), and appear in Table I in the ‘Theory’ columns, together with the total number of (predicted) links, calculated using  $z_1$  in (4). Finally, using (7),  $C_{\text{random}}$  was calculated for the “usual” random graph.

Having  $f_0(x)$  and  $g_0(x)$  exactly, MAPLE was used to find the generating function for the probability distribution  $r_z$  of number of stations  $z$  a randomly chosen station is linked to:

$$G_0(x) = f_0(g_1(x)) = \sum_z r_z x^z. \quad (15)$$

By swapping the roles of  $f$  and  $g$  the probability distribution for interactions (i.e., interlock) between train lines could also be determined.

#### B. Discussion

From Table I, both networks satisfy the two basic conditions for being small-world. The size of the networks, measured by diameter ( $D = 3$  for both) or by average path length  $L$ , is small when compared to the number of vertices. The actual clustering coefficients are much greater than  $C_{\text{random}}$  for appropriate  $N, K$ . While in relative terms the ratio  $C/C_{\text{random}}$  is orders of magnitude smaller than, say, that observed for the Hollywood network [1, 5], it is comparable with that observed in studies of networks of similar size as listed in Table IV.

However, in absolute terms, for both networks  $C$  is very much higher than the values reported in other studies. Our definition of connectivity has resulted in  $C$  being so close to 1 because of the proportion of stations in each network ( $\gtrsim 80\%$ , see Table II) which lie on a single train line. All stations to which they are connected are also linked to one another, forming a clique, and each giving  $C_v = 1$  as their contribution to (3). The most useful comparison may be to the railway network of India [10] which has  $C = 0.69$ . A country-wide train network is not constrained as an urban subway network is, and train routes criss-cross between stations in different patterns, giving lower local clustering.

However, some amount of this clustering is expected solely on the basis of the underlying bipartite structure [5], and the calculation of  $C^{(T)}$  gives a measure of it. Although they are not precisely the same quantity, comparing the actual value of  $C$  to the theoretical  $C^{(T)}$  for the same distributions  $p_j$  and  $q_k$  shows that the Boston

	Boston		Vienna	
	Actual	Theory	Actual	Theory
Number of stations, $N$		124		76
Number of train lines, $M$		8		5
Number of links, $K$	1711	1936	785	788
Average degree, $z_1 = z$	27.597	31.226	20.658	20.737
Mean 2nd neighbors, $z_2$	91.048	766.75	43.842	97.213
Mean 3rd neighbors, $z_3$	4.355		10.500	
Network size, $L, \ell$	1.8110	1.4187	1.8646	1.7236
Clustering coefficient, $C, C^{(T)}$	0.9276	0.4808	0.9450	0.7989
$C_{\text{random}}$		0.2244		0.2754

TABLE I: Network properties for each train network. The results in the ‘Actual’ columns were obtained from the network maps (by spreadsheet calculations). Apart from  $C_{\text{random}}$ , the results in the ‘Theory’ columns were obtained from the formulae in Section IIB for random bipartite graphs with appropriate specified probability distributions, using MAPLE.

$j :$	1	2	3	4	5	6
Boston	100	12	3	5	3	1
Vienna	67	8	1			

TABLE II: Number of stations lying on  $j$  lines, i.e.,  $Np_j$ .

$k :$	7	12	14	18	19	20	21	24	33
Boston		1		1	1	1		3	1
Vienna	1		1			1	1	1	

TABLE III: Number of lines containing  $k$  stations, i.e.  $Mq_k$ .

network appear to have more genuine or “excess” clustering. For comparison, in Table V some other values from the literature are listed. In social networks, this clustering is explained by people introducing their friends or collaborators to one another [1, 5] while for the artificial world of Marvel comics it arises from the way that the (6 486) characters are not distributed randomly across all the (12 942) comic books but appear in teams, in series [6]. The Vienna network has a structure similar to that originally proposed to model small-worlds [1, 2]: it consists of cliques of connected nodes (stations on a single line) which interact with one another through certain nodes (stations common to two lines) which connect the network but have very low values of  $C_v$ . On the other

Network	$N$	$K$	$C/C_{\text{random}}$	Reference
Boston	124	1711	4.13	
Vienna	76	785	3.43	
Dolphins	64	159	3.74	[14]
China airports	128	1165	5.13	[15]
<i>C. elegans</i>	282	1974	5.60	[1, 2]
Hollywood	450,000	$2.55 \times 10^7$	789	[5]

TABLE IV: Actual clustering coefficients compared to  $C_{\text{random}}$  for appropriate network size  $N$  and number of links  $K$ .

Network	$C$ (actual)	$C^{(T)}$ (theory)	$C/C^{(T)}$	Ref.
Boston	0.9276	0.4804	1.93	
Vienna	0.9450	0.7989	1.18	
Fortune 1000	0.588	0.590	1.00	[5]
Physics co-authors	0.452	0.019	2.35	[5]
Hollywood	0.199	0.084	2.36	[5]
Marvel	0.192	0.0066	29.09	[6]

TABLE V: Comparison of actual and theoretical clustering for a number of one-mode networks with underlying bipartite structure.

hand, because of the branched lines in the Boston network, whole groups of stations are shared between lines, overlapping the cliques and enhancing  $C_v$  for the shared vertices.

The average degree  $z$  as predicted by the random bipartite theory agrees well with the actual observed values (see Table I). Using  $z$  to calculate the total number of links  $K$  shows that this agreement is better for Vienna than Boston, where the prediction is for an extra 200-odd links between the 124 stations. This is also attributable to the effect of the branched lines, causing stations to interact repeatedly with groups of common stations, rather than having links to as many new stations as might be expected based on the number of lines on which they lie.

The worst discrepancy between the predictions of the random bipartite model and the actual results was in estimating the number of second neighbors  $z_2$ . In both cases, the predicted number exceeded the actual average number of second neighbors. This is a flow-on from the high local clustering; if a node and most of its neighbors form a clique, there are few paths of length two. But the theoretical  $z_2$  also exceeded the total number of nodes in the network, in the case of Boston by an order of magnitude!

There are several limitations to our study which may have affected how the results compare with the bipartite theory. The small-world effect is notable in systems which are not only connected and decentralized (as ours

are) but also *large*, and *sparse* [1]. Though the value of  $N$  is not extremely large, it is comparable to the other studies cited in Table IV. Sparseness refers to the number of links  $K$  in comparison to the number of nodes  $N$ , specifically that

$$K \ll \binom{N}{2} \quad \text{or} \quad z \ll N - 1. \quad (16)$$

We have  $z/(N-1) \approx 0.25$  which is not particularly small, and is certainly larger than the studies cited in Table IV. One feature of our study which appears unique is that  $M/N \approx 1/15$ . For other networks which have been studied from the bipartite point of view this ratio takes values: 0.5 (Marvel universe [6]), 1.0 (Indian railroad [10]), 2.0 (Hollywood [2, 3]), and 8.9 (Fortune 1000 [3, 11]). We speculate that the balance of the various sizes ( $N$ ,  $M$ ,  $K$ ) may be the main problem with the second degree estimate, but since no other authors have calculated or commented on  $z_2$  in their studies we cannot offer further support for this explanation. (We hence do not give too much attention to the calculated value of  $\ell$  since it relies on both  $N$  and  $z_2$ .)

Finally we comment on the distributions of linked stations and interlocked train lines, though we do not give the corresponding long but rather sparse tables of actual values. The distribution  $r_z$  from (15) for both Vienna and Boston corresponded closely to the actual values, with some smoothing and a long, low tail as would be expected. The analogous generating function, for how many lines a randomly-chosen line shares stations with, gave reasonable agreement with the actual values for Vienna. However, the expression for Boston suffered from a similar problem to the  $z_2$  results referred to above; it was a very smeared-out distribution which gave the highest probability to 22 lines, though actually the maximum possible is  $M - 1 = 7$ . Presumably this problem is due not so much to  $M$  being small (since Vienna would have had that problem too), but mainly to the unusual actual distributions which result from the common stretches of line, which the averaged distribution from the random

model struggles to mimic. In this context, it should be noted that the theoretical interlock of company directors [5] was not in close agreement with the actual data, and the proffered explanation is similar (when translated into more human terms).

### C. Conclusion

To quote the paper [5] that introduced analysis based on bipartite structure to the small-world literature: “[It] is perhaps best to regard our [bipartite] random graph as a null model—a baseline from which our expectations about network structure should be measured. It is deviation from the random graph behavior, not agreement with it, that allows us to draw conclusions about real-world networks.” Although there have been other studies of small-world properties of transport networks [15], including railway networks [8], this study appears to be the first that has not only had underlying bipartite structure [10], but has used this in its analysis. Further, because we have studied two networks at once, we have been able to draw comparisons between them based on their different train-line architecture. We have seen properties in common, such as high  $C$ , properties close to prediction, e.g.,  $z$ , and properties which vary from the random bipartite graph model, and we have attempted to explain the behavior in terms of our definition of connectivity and how this works itself out in the actual networks. This study extends the application of bipartite analysis from social affiliation networks [5, 6] to the technological; presumably there are examples to be found also among real-world information and biological networks.

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[18] Note that the new silver line was incomplete at the time

of [8], so we did not include it in our study either.