

I Testing problems relating to a univariate normal distribution.

10th March '23

- i) Parameter of interest is μ / we are to test for $H_0: \mu = \mu_0$
- ~~(a)~~ (a) σ is known, say $= \sigma_0$
 - (b) σ is unknown.

Problems related to a single univariate normal distribution. 24th Feb '23

$$X \sim N(\mu, \sigma^2)$$

1. Testing problems on μ

(a) σ is known say $= \sigma_0$ ($\sigma_0 > 0$)

(i) To test $H_0: \mu = \mu_0$ ag $H_1: \mu > \mu_0$

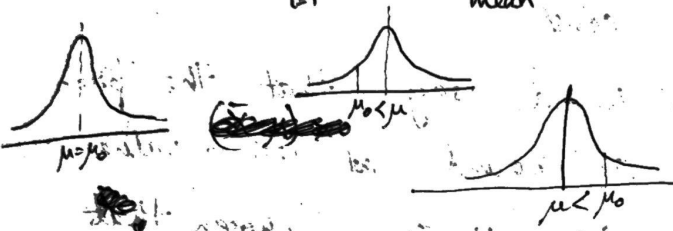
(μ_0 : pre specified value $\in \mathbb{R}$)

Let X_1, \dots, X_n i.i.d of size n from the dist of

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i = \text{sample mean}$$

$(\bar{X} - \mu_0) = 0$ If H_0 is true,

then $(\bar{X} - \mu_0)$ would be close to 0.



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We only reject H_0 if the situation sufficiently favours H_1 else it will favour H_0

$(\bar{X} - \mu_0) > 0$ I reason to favour H_1 over H_0

$(\bar{X} - \mu_0) < 0$ I reason to ~~favour~~ stick to H_0

Thus the difference $(\bar{X} - \mu_0)$ provides us with an idea of favouring or disfavouring H_0 . In particular, a large +ve value of $(\bar{X} - \mu_0)$ indicates a departure from H_0 in favour of H_1 .

Thus the test statistic to test (H_0) should be based on the quantity $(\bar{X} - \mu_0)$

We know from theory that $\bar{X} \sim N(\mu, \frac{\sigma^2}{n})$

Since σ is known $= \sigma_0$ we have $\bar{X} \sim N(\mu_0, \frac{\sigma_0^2}{n})$

Under H_0 (assuming H_0 as true) $\bar{X} \sim N(\mu_0, \frac{\sigma_0^2}{n})$ or $\frac{\bar{X} - \mu_0}{\frac{\sigma_0}{\sqrt{n}}} \sim N(0, 1)$

An appropriate test statistic to test H_0 ag H_1 is given by $Z = \frac{\sqrt{n}(\bar{x} - \mu_0)}{\sigma_0}$ which under

H_0 follows a $N(0,1)$ distⁿ

A high value of Z indicates that there is strong evidence to doubt the validity of H_0 .

Thus the critical region of the test should be

of the form $Z = \frac{\sqrt{n}(\bar{x} - \mu_0)}{\sigma_0} > K$ if choose K is an

appropriately chosen const.

K is chosen in a way that the prob of type I error is bound at desired value

α ($0 < \alpha < 1$) i.e., K is so chosen that

$P_{H_0}(Z = \frac{\sqrt{n}(\bar{x} - \mu_0)}{\sigma_0} > K) = \alpha$ i.e., $K = \text{upper } \alpha \text{ point}$

of a $N(0,1)$ distⁿ Z_{α}

Test Rule: Reject H_0 at level of significance α if

$Z_{\text{obs}} = \frac{\sqrt{n}(\bar{x} - \mu_0)}{\sigma_0} > Z_{\alpha}$

accept H_0

Suppose we want to test

$H_0: \mu = \mu_0$ ag $H_1: \mu < \mu_0$

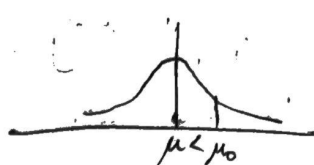
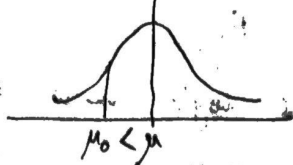
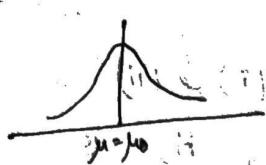
let x_1, \dots, x_n be observations of size n drawn from the distⁿ of X . \bar{x} = sample mean

On line with logic of (i) $\bar{x} = \frac{1}{n} \sum_{i=1}^n x_i$

$(\bar{x} - \mu_0) \approx 0$ H_0 seems to be true

$(\bar{x} - \mu_0) > 0$ Better to stick to H_0

$(\bar{x} - \mu_0) < 0$ Disfavour H_0 in favour of H_1



Thus a large negative value of $(\bar{x} - \mu_0)$ would indicate that \exists sufficient evidence to doubt the validity of H_0 . In line with the logic used in (i) an apt test statistic for testing H_0 is

$$Z = \frac{\sqrt{n}}{\sigma_0} (\bar{x} - \mu_0) \underset{H_0}{\sim} N(0, 1)$$

Critical region: $Z = \frac{\sqrt{n}}{\sigma_0} (\bar{x} - \mu_0) < k'$ where k' is

So chosen that the prob of type I error is bound at a desired level α ($0 < \alpha < 1$)

i.e.

$$P_{\mu_0} \left\{ Z = \frac{\sqrt{n}}{\sigma_0} (\bar{x} - \mu_0) < k' \right\} = \alpha$$

i.e. $k' =$ lower α -pt of $N(0, 1)$

$$= \mathcal{L}_{(1-\alpha)} = -Z_\alpha$$

Test rule: Reject H_0 at α -los

$$\text{iff } Z_{\text{obs}} = \frac{\sqrt{n}}{\sigma_0} (\bar{x} - \mu_0) < -Z_\alpha$$

(iii) Suppose we are to test $H_0: \mu = \mu_0$ ag $H_1: \mu \neq \mu_0$. Let x_1, \dots, x_n be a s.s of size n drawn from the distⁿ of X .

\bar{x} = sample mean = $\frac{1}{n} \sum_{i=1}^n x_i$. In line with logic of (i) & (ii).

$(\bar{x} - \mu_0) \approx 0$: H_0 seems to be true.

$(\bar{x} - \mu_0) <, > 0$: Disfavour H_0 in favour of H_1 .

Thus a high \pm ve value of $(\bar{x} - \mu_0)$ would indicate that \exists sufficient evidence to doubt

the validity of H_0

In line with the logic used in (2) & (7)
an apt test statistic for testing H_0 is

$$Z = \frac{\sqrt{n}}{\sigma_0} (\bar{x} - \mu_0) \sim N(0, 1)$$

Critical Region: $Z = \frac{\sqrt{n}}{\sigma_0} (\bar{x} - \mu_0) > k''$ and

$$\frac{\sqrt{n}}{\sigma_0} (\bar{x} - \mu_0) < -k''$$

To maintain symmetry, we choose k'' & k''' to be equal to k^* , in magnitude.

$$\text{i.e. } Z = \frac{\sqrt{n}}{\sigma_0} (\bar{x} - \mu_0) > k^* \text{ \& } \frac{\sqrt{n}}{\sigma_0} (\bar{x} - \mu_0) < -k^*$$

$$\text{ie, } |Z| = \frac{\sqrt{n} |\bar{x} - \mu_0|}{\sigma_0} > k^*$$

where k^* is so chosen that the prob. of Type I error is bound at α , $0 < \alpha < 1$.

$$\text{Now, } P_{\mu_0} \left\{ \frac{\sqrt{n}}{\sigma_0} |\bar{x} - \mu_0| \geq k^* \right\} = \alpha$$

i.e., $k^* = \text{upper } \alpha/2^{\text{th}} \text{ point of } N(0, 1) \text{ dist}^n$
 $= Z_{\alpha/2}$

Test Rule:

Reject H_0 at α l.o.s iff $|Z_{\text{obs}}| > Z_{\alpha/2}$

Accept

if $|Z_{\text{obs}}| \leq Z_{\alpha/2}$

$$X \sim N(\mu, \sigma^2) \quad | \quad H_0: \mu = \mu_0 \quad \text{ag} \quad H_1: \mu > \mu_0$$

or

$$H_1: \mu < \mu_0$$

or

$$H_1: \mu \neq \mu_0$$

(b) σ is unknown

==

We can not use the test statistic $Z = \frac{\sqrt{n}(\bar{X} - \mu_0)}{\sigma}$ as it involves the unknown ~~sample~~ constant σ .

Thus we replace σ by an estimate obtained from the sample.

$$\hat{\sigma} = s = \sqrt{\frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2}$$

An appropriate test statistic for testing H_0 is given

by $T = \frac{\sqrt{n}(\bar{X} - \mu_0)}{s}$ which under H_0 follows a t -distribution with $(n-1)$ dof.

$$T = \frac{\sqrt{n}(\bar{X} - \mu_0)}{s} = \frac{\frac{\sqrt{n}(\bar{X} - \mu_0)}{\sigma}}{\frac{\sqrt{(n-1) \frac{s^2}{\sigma^2}}}{\sigma}} = \frac{N(0,1)}{\sqrt{\chi_{n-1}^2/n}}$$

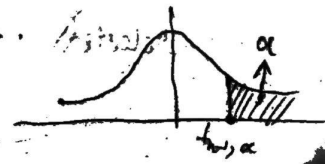
$$T = \frac{N(0,1)}{\sqrt{\chi_{n-1}^2/(n-1)}}$$

- In line with the logic in case when σ is known,
- for $H_1: \mu > \mu_0$ high +ve values of T indicate departure from H_0
 - for $H_1: \mu < \mu_0$ high -ve values of T indicate departure from H_0
 - For $H_1: \mu \neq \mu_0$, high +ve values of T indicate departure from H_0

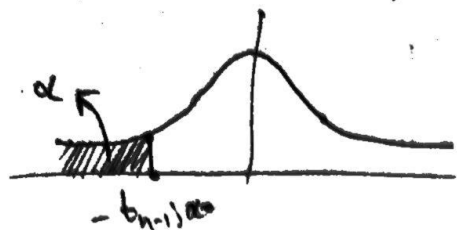
The test rules are given by

1/ when $H_1: \mu > \mu_0$: Reject H_0 iff $T_{obs} > t_{n-1; \alpha}$ (same $t_{\alpha; n-1}$)

where $t_{n-1; \alpha}$ is the upper α -point of a t -distⁿ with $(n-1)$ dof and α is known.



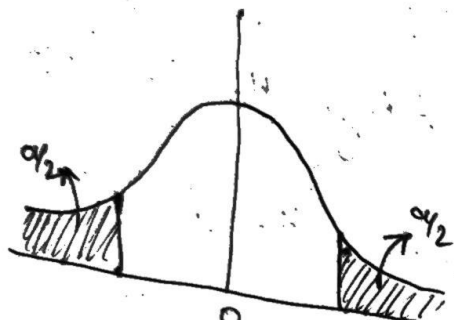
ii) when $H_1: \mu < \mu_0$, Reject H_0 iff $T_{obs} < t_{n-1; 1-\alpha}$



$\equiv T_{obs} < -t_{n-1; \alpha}$
(because t-distribution is
symm about 0)

where α is desired l.o.s.

iii) when $H_1: \mu \neq \mu_0$, Reject H_0 iff $T_{obs} > t_{n-1; \alpha/2}$ ~~or~~
 $T_{obs} < -t_{n-1; \alpha/2} \Rightarrow |T_{obs}| > t_{n-1; \alpha/2}$



where α is the desired
l.o.s.