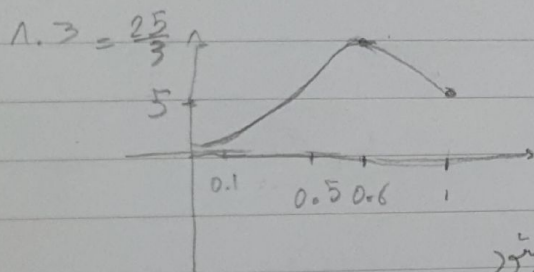


سؤال 2 الف

$$L(\theta) = \text{Arg max}_{\theta} P(x|\theta) = \prod_{i=1}^n \begin{cases} \frac{1}{\theta} & 0 \leq x_i \leq \theta \\ 0 & \text{o.w} \end{cases}$$

$$\Rightarrow \log P(x|\theta) = \sum_{i=1}^n \frac{1}{\theta} \quad 0 \leq x_i \leq \theta \Rightarrow \text{جواب ندارد} \Rightarrow \sum_{i=1}^n -\frac{1}{\theta^2} : \text{منفی}$$

$$\Rightarrow \text{بیشترین حالت} : \sum_{i=1}^n \frac{1}{\theta} \Rightarrow \frac{n}{\theta} \quad 0 \leq x_i \leq \theta \Rightarrow \text{تمام } x_i \text{ ها برابر باشد} \Rightarrow \theta = \max[x_i]$$



نیاز به داشتن بقیه مقادیر نداریم چرا که همان طور که در سؤال قبل

پیدا کردیم $\max \theta = \max[x_i]$ و می دانیم با خط از آن $P(x|\theta)$ کم می شود

$$\text{if } \theta \geq 0.6 \Rightarrow P(x|\theta) = \frac{5}{\theta}$$

$$P(x|\theta) = \prod_{i=1}^n \theta^{x_i} (1-\theta_i)^{1-x_i}$$

سؤال 3:

$$L(\theta) = \text{Arg max}_{\theta} P(x|\theta) \Rightarrow \log L(\theta) = \sum_{i=1}^n \log \theta_i^{x_i} (1-\theta_i)^{1-x_i}$$

$$= \sum_{i=1}^n x_i \log(\theta_i) + (1-x_i) \log(1-\theta_i) \xrightarrow{\sum_{i=1}^n x_i = Y} Y \log \theta + (n-Y) \log(1-\theta)$$

$$\Rightarrow \frac{\partial L(\theta)}{\partial \theta_i} = Y \frac{1}{\theta_i} + (n-Y) \frac{-1}{1-\theta_i} = 0 \Rightarrow \theta_i = \frac{Y}{n} = \frac{\sum_{i=1}^n x_i}{n} \quad \text{میانگین گرفته}$$



$$n_0 = \frac{\sum_{k=1}^n x_k}{\sum_{k=1}^n 1} = \arg \max_{k: n_0+1, \dots, 0}$$

سؤال 4:

$$\text{الف) } \mu_n = \frac{1}{n+n_0} \sum_{k=n_0+1}^n x_k, \quad \sigma_n^2 = \frac{s}{n+n_0} \rightarrow \mu_0 = \frac{1}{n_0} \sum_{i=n_0+1}^0 x_i$$

$$\mu_n = \frac{\sum_{i=1}^n x_i}{n + \frac{s^2}{\sigma_0^2}} + \frac{\frac{s^2}{\sigma_0^2}}{n_0 \left(\frac{s^2}{\sigma_0^2} + n \right)} \sum_{k=n_0+1}^0 x_k = \frac{\sum_{i=1}^n x_i}{n + n_0} + \frac{\sum_{k=n_0+1}^0 x_k}{n_0 + n} = \frac{\sum_{k=n_0+1}^n x_k}{n_0 + n}$$

$$\sigma_n^2 = \frac{s^2 \sigma_0^2}{n \sigma_0^2 + s^2} = \frac{s^2}{n + \frac{s^2}{\sigma_0^2}} = \frac{s^2}{n + n_0}$$

ب)

$$P \sim N(\mu_0, \sigma_0^2) \rightarrow \mu_0 = \frac{1}{n_0} \sum_{i=n_0+1}^0 x_i, \quad \sigma_0^2 = \frac{s^2}{n}$$

در این صورت MLE و Bayes E برای حالتی که ML توابع داده ها و Bayes E و Bayes E برای حالت

سؤال 5:

$$\text{MAP} \rightarrow \arg \max_{\mu} \frac{P(x_i | \mu) P(\mu)}{P(x_i)} = \arg \max_{\mu} P(x_i | \mu) P(\mu) = \hat{\mu}$$

$$\Rightarrow \hat{\mu} = \prod_{i=1}^N P(x_i | \mu) P(\mu) \rightarrow \log \rightarrow \sum_{i=1}^N (\log P(x_i | \mu) + \log P(\mu))$$

$$= \sum_{i=1}^N \left[-\frac{1}{2} \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} (x_i - \mu)^2 + \log \left(\frac{\mu}{\sigma^2} \right) - \frac{\mu^2}{2\sigma^2 \mu} \right] \rightarrow \frac{\partial L}{\partial \mu} = 0$$

$$\Rightarrow \sum_{i=1}^N \left[-\frac{1}{2\sigma^2} (-1) (x_i - \mu) + \frac{1}{\mu} - \frac{2\mu}{2\sigma^2 \mu} \right] = 0 \rightarrow \sum_{i=1}^N \left[\frac{1}{\sigma^2} x_i - \mu \left(\frac{1}{\sigma^2} + \frac{1}{\sigma^2 \mu} \right) + \frac{1}{\mu} \right] = 0$$

$$\mu \text{ در 2 طرف } \rightarrow \mu = \frac{\sum x_i}{\sigma^2} \pm \sqrt{\left(\frac{\sum x_i}{\sigma^2} \right)^2 + 4N^2 \left(\frac{1}{\sigma^2} + \frac{1}{\sigma^2 \mu} \right)} \quad 2N \left(\frac{1}{\sigma^2} + \frac{1}{\sigma^2 \mu} \right)$$

سؤال ٥

$$P(z) = \frac{\lambda^z e^{-\lambda}}{z!}, \log P(z) = z \log \lambda - \lambda - \log z! \quad (2)$$

$$Q(\theta, \theta^{(0)}) = \sum_{j=1}^m \sum_{i=1}^n \log (\lambda_j P_j(x_i | \theta_j)) P(j | x_i, \theta^{(0)})$$

$$= \sum_{j=1}^m \sum_{i=1}^n \log (\lambda_j) P(j | x_i, \theta^{(0)}) + \sum_{j=1}^m \sum_{i=1}^n \log (P_j(x_i | \theta_j)) P(j | x_i, \theta^{(0)})$$

$$\textcircled{1} \rightarrow \hat{\lambda}_j = \frac{1}{n} \sum_{i=1}^n P(j | x_i, \theta^{(0)}), P(j | x_i, \theta^{(0)}) = \frac{\lambda_j^{(0)} P_j(x_i | \theta_j^{(0)})}{\sum_{j=1}^m \lambda_j^{(0)} P_j(x_i | \theta_j^{(0)})}$$

$$\textcircled{2} \frac{\partial LL}{\partial \lambda} = 0 \Rightarrow \frac{x_i}{\lambda_j} - 1 = 0 \Rightarrow \lambda_j = x_i \Rightarrow \hat{\lambda}_j = \frac{\sum_{i=1}^n P(j | x_i, \theta^{(0)}) x_i}{\sum_{i=1}^n P(j | x_i, \theta^{(0)})}$$