# 1 Distributions

Bernoulli: 
$$\begin{cases} 1 - p & \text{if } x = 0 \\ p & \text{if } x = 1 \end{cases}$$

$$\mu=p, Var=pq$$

$$MGF = q + pe^t$$

Binomial:  $\binom{n}{k} p^k q^{n-k}$ 

$$\mu = np, Var = npq$$

$$MGF = (q + pe^t)^n$$

Geometric:  $p(1-p)^{k-1}$ 

$$\mu = \frac{1}{p}, Var = \frac{1-p}{p^2}$$

$$MGF = \frac{pe^t}{1 - (1 - p)e^t}$$

Poisson:  $\frac{e^{-\lambda}\lambda^k}{k!}$ 

$$\mu = \lambda, Var = \lambda$$

$$MGF = e^{\lambda(e^t-1)}$$

Normal:  $\frac{1}{\sqrt{2\pi\sigma^2}}e^{-\frac{(k-\mu)^2}{2\sigma^2}}$ 

$$\mu = \mu$$
,  $Var = \sigma^2$ 

$$MGF = e^{\mu t + \frac{1}{2}\sigma^2 t^2}$$

Exponential:  $\lambda e^{-\lambda k}$ 

$$\mu = \frac{1}{\lambda}, Var = \frac{1}{\lambda^2}$$

$$MGF = \frac{\lambda}{\lambda - t}, \ t < \lambda$$

Uniform:  $\frac{1}{b-a}$ ,  $a \le k \le b$ 

$$\mu = \frac{a+b}{2}, Var = \frac{(b-a)^2}{12}$$

$$MGF = \frac{e^{bt} - e^{at}}{t(b - a)}$$

Beta: 
$$\frac{k^{\alpha-1}(1-k)^{\beta-1}}{B(\alpha,\beta)}$$
,  $0 \le k \le 1$ 

$$\mu = \frac{\alpha}{\alpha + \beta}, Var = \frac{\alpha\beta}{(\alpha + \beta)^2(\alpha + \beta + 1)}$$

$$MGF = {}_{2}F_{1}(\alpha, \alpha + \beta; \alpha + \beta + 1; t)$$

$$\begin{split} \beta(\alpha,\beta) &= \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha+\beta)} \\ \Gamma(\alpha) &= (\alpha-1)!, \ \alpha \in \mathbb{N} \\ \text{Gamma:} \ \frac{\beta^{\alpha}k^{\alpha-1}e^{-\beta k}}{\Gamma(\alpha)} \end{split}$$

$$\mu = \frac{\alpha}{\beta}, Var = \frac{\alpha}{\beta^2}$$

$$MGF = \left(\frac{\beta}{\beta - t}\right)^{\alpha}, \ t < \beta$$

# Formulas

# 2.1 Probability Formulas

 $A \perp B \Leftrightarrow \mathbb{P}(A||B) = \mathbb{P}(A) \& \mathbb{P}(B||A) = \mathbb{P}(B) \Leftrightarrow$  $\mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$ 

Union Bound:  $\mathbb{P}(\bigcup_{i=1}^n A_i) \leq \sum_{i=1}^n \mathbb{P}(A_i)$ Bayes' Rule:  $\mathbb{P}(A \| B) = \frac{\mathbb{P}(B \| A) \cdot \mathbb{P}(A)}{\mathbb{P}(B)}$ 

Law of Total Probability:  $\mathbb{P}(A) = \sum_{i=1}^{n} \mathbb{P}(A || B_i)$ .

 $\mathbb{P}(B_i)$ Chain Rule:  $\mathbb{P}(A_1 \cap A_2 \cap ... \cap A_n) = \mathbb{P}(A_1)$ .

 $\mathbb{P}(A_2 || A_1) \cdot \ldots \cdot \mathbb{P}(A_n || A_1 \cap \ldots \cap A_{n-1})$ Conditional Independence:  $A \perp B \parallel C \Leftrightarrow$  $\mathbb{P}(A || B \cap C) = \mathbb{P}(A || C)$ 

# Expected Value and Variance

$$\begin{split} \mathbb{E}[aX+b] &= a\mathbb{E}[X] + b \\ \mathrm{Var}(X) &= \mathbb{E}[X^2] - \mathbb{E}[X]^2 = \mathbb{E}((X-\mathbb{E}(X))^2) \\ \mathrm{Var}(X+Y) &= \mathrm{Var}(X) + \mathrm{Var}(Y) \text{ if } X \perp Y \\ \mathrm{Var}(aX) &= a^2 \mathrm{Var}(X) \end{split}$$

# 2.3 Moments and MGF

$$\begin{split} & \frac{\mu_k}{\mu_k} = \mathbb{E}(X^k) \\ & \frac{\mu_k}{\mu_k} = \mathbb{E}((X - \mathbb{E}(X))^k) \\ & M_X(t) = \mathbb{E}(e^{tX}) \\ & M_{X+Y}(t) = M_X(t) \cdot M_Y(t) \text{ if } X \perp Y \\ & \mu_k = M_X^{(k)}(0) \\ & \mathbb{E}[\mathbb{E}[X \| Y]] = \mathbb{E}[X] \end{split}$$

# 2.4 function of random variables

$$\mathbb{E}[g(X)] = \sum_{x} g(x) \cdot \mathbb{P}(X = x)$$

$$f_{Y}(y) = f_{X}(g^{-1}(y)) \cdot \left| \frac{d}{dy} g^{-1}(y) \right|$$

## 2.5 Common Distributions

$$\begin{split} F_{X,Y}(x,y) &= \mathbb{P}(X \leq x, Y \leq y) \\ f_{X,Y}(x,y) &= \frac{\partial^2}{\partial x \partial y} F_{X,Y}(x,y) \\ f_X(x) &= \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy \\ X \perp Y &\Leftrightarrow f_{X,Y}(x,y) = f_X(x) \cdot f_Y(y) \\ \text{Jointly Gaussian:} \end{split}$$

$$\begin{split} f_{X,Y}(x,y) &= \frac{1}{2\pi\sigma_X\sigma_Y\sqrt{1-\rho^2}} \exp\left(-\frac{1}{2(1-\rho^2)}\right) \\ &\left(\left(\frac{x-\mu_X}{\sigma_X}\right)^2\right. \\ &\left. -2\rho\left(\frac{x-\mu_X}{\sigma_X}\right)\left(\frac{y-\mu_Y}{\sigma_Y}\right) \\ &\left. + \left(\frac{y-\mu_Y}{\sigma_Y}\right)^2\right)\right) \end{split}$$

$$f_{X||Y}(x||y) = \frac{f_{X,|Y}(x,y)}{f_Y(y)}$$

$$\mathbb{E}[X||Y = y] = \int_{-\infty}^{\infty} x \cdot f_{X||Y}(x||y) dx$$

## 2.6 Normal Distribution

if  $X \sim \mathcal{N}(\mu_X, \sigma_X^2)$  and  $Y \sim \mathcal{N}(\mu_Y, \sigma_Y^2)$ , then 3.3 Chernoff Bound  $X + Y \sim \mathcal{N}(\mu_X + \mu_Y, \sigma_X^2 + \sigma_Y^2)$ if X and Y are jointly Gaussian, then  $X || Y = y \sim$  $\mathcal{N}(\mu_X + \rho \frac{\sigma_X}{\sigma_Y}(y - \mu_Y), \sigma_X^2(1 - \rho^2))$ 

## 2.7 Covaariance and Correlation

$$\begin{array}{lll} \operatorname{Cov}(X,Y) &=& \mathbb{E}[(X-\mathbb{E}(X))(Y-\mathbb{E}(Y))] &=& \\ \mathbb{E}(XY)-\mathbb{E}(X)\mathbb{E}(Y) & \\ \operatorname{Corr}(X,Y) &=& \frac{\operatorname{Cov}(X,Y)}{\sigma_X\sigma_Y} \\ \text{if two random variables are linearly independent} \end{array}$$

then Cov(X,Y) = 0Cov(aX + b, cY + d) = acCov(X, Y) (independent

of a,b,c,d

$$\mu_{i,j} = \mathbb{E}(X^i Y^j)$$
  
$$\overline{\mu_{i,j}} = \mathbb{E}((X - \mathbb{E}(X))^i (Y - \mathbb{E}(Y))^j)$$

## CLT and LLN

Central Limit Theorem:  $\frac{\overline{X}-\mu}{\sigma(\sqrt{n})} \to \mathcal{N}(0,1) \text{ as } n \to \infty$ 

Law of Large Numbers:  $\overline{X} \to \mu \text{ as } n \to \infty$ 

# 2.9 Functions of multiple Random Variables

$$J = \begin{bmatrix} \frac{\partial g_1}{\partial x_1} & \cdots & \frac{\partial g_1}{\partial x_n} \\ \vdots & \ddots & \vdots \\ \frac{\partial g_m}{\partial x_1} & \cdots & \frac{\partial g_m}{\partial x_n} \end{bmatrix}$$

$$\begin{array}{l} f_{Y_1,...,Y_m}(y_1,\ldots,y_m) = f_{X_1,...,X_n}(g_1^{-1}(y_1,\ldots,y_m) \\ \ldots,g_m^{-1}(y_1,\ldots,y_m)) \cdot det(J) \end{array}$$

# 3 Probability Inequalities

## 3.1 Markov's Inequality

For any non-negative random variable X and any a > 0:

$$\mathbb{P}(X \ge a) \le \frac{\mathbb{E}[X]}{a}$$

# Chebyshev's Inequality

For any random variable X and any a > 0:

$$\mathbb{P}(|X - \mathbb{E}[X]| \ge a) \le \frac{\operatorname{Var}(X)}{a^2}$$

For any random variable X and any t > 0:

$$\mathbb{P}(X \ge (1+\delta)\mathbb{E}[X]) \le \left(\frac{e^{\delta}}{(1+\delta)^{1+\delta}}\right)^{\mathbb{E}[X]}$$

$$\mathbb{P}(X \le (1 - \delta)\mathbb{E}[X]) \le \left(\frac{e^{-\delta}}{(1 - \delta)^{1 - \delta}}\right)^{\mathbb{E}[X]}$$

# 3.4 Jensen's Inequality

For any random variable X and any convex function f:

$$f(\mathbb{E}[X]) \le \mathbb{E}[f(X)]$$

# 3.5 Cauchy-Schwarz Inequality

given two vectors x and y:

$$\left(\sum_{i=1}^n x_i y_i\right)^2 \le \left(\sum_{i=1}^n x_i^2\right) \left(\sum_{i=1}^n y_i^2\right)$$

For any random variables X and Y:

$$\mathbb{E}[XY]^2 \le \mathbb{E}[X^2] \cdot \mathbb{E}[Y^2]$$

and the equality holds if and only if X = aYfor some  $a \in \mathbb{R}$ .

# Sample Statistics

# 4.1 Properties:

Bias:  $\mathbb{E}[\hat{\theta}] - \theta$ 

Variance:  $Var(\hat{\theta}) = \mathbb{E}[(\hat{\theta} - \mathbb{E}[\hat{\theta}])^2]$ 

Mean Squared Error:  $MSE(\hat{\theta}) = \mathbb{E}[(\hat{\theta} - \theta)^2] =$ 

 $Var(\hat{\theta}) + Bias^2$ 

Consistency:  $\lim_{n\to\infty}\hat{\theta}=\theta$ 

# 4.2 Estimators:

Sample Mean:  $\overline{X} = \frac{1}{n} \sum_{i=1}^n X_i$ 

Sample Variance: 
$$S^2 = \frac{1}{n-1} \sum_{i=1}^{n} (X_i - \overline{X})^2$$

## Maximum Likelihood Estimation:

$$\begin{split} \hat{\theta}_{\text{MLE}} &= \arg \max_{\theta} \prod_{i=1}^{n} f(X_i \| \theta) \\ &= \arg \max_{\theta} \sum_{i=1}^{n} \log f(X_i \| \theta) \\ &\Rightarrow \frac{\partial_{\mathcal{L}}(\theta)}{\partial_{\theta}} \end{split}$$

## MMSE:

linear MMSE:  $h(a, b) = \mathbb{E}[(X - aY - b)^2]$ objective: minimize h(a, b)

solution:  $a = \frac{\text{Cov}(X,Y)}{\text{Var}(Y)} = \rho \frac{\sigma_X}{\sigma_Y}, \ b = \mathbb{E}[X] - a\mathbb{E}[Y]$ multiple MMŠÉ:

$$\begin{bmatrix} \sigma_{X_1}^2 & \sigma_{X_1X_2} \\ \sigma_{X_1X_2} & \sigma_{X_2}^2 \end{bmatrix} \begin{bmatrix} a_1^* \\ a_2^* \end{bmatrix} = \begin{bmatrix} \operatorname{Cov}(X_1, Y) \\ \operatorname{Cov}(X_2, Y) \end{bmatrix}$$

Covariance Matrix

$$\hat{Y}(X_1, X_2) = a_1^*(X_1 - \mathbb{E}[X_1]) + a_2^*(X_2 - \mathbb{E}[X_2]) + \mathbb{E}[Y]$$

# 4.5 Hypothesis Testing:

Null Hypothesis:  $H_0: \theta = \theta_0$ 

Alternative Hypothesis:  $H_1: \theta \neq \theta_0$ 

Type I Error: Reject  $H_0$  when it is true

Type II Error: Accept  $H_0$  when it is false

P value: Probability of observing the data or

more extreme data given  $H_0$  is true

# 4.6 Z-Test:

$$Z = \frac{X - \mu}{\sigma / \sqrt{n}}$$
  

$$\Rightarrow \text{p-value} = 2 \cdot (1 - \Phi(|Z|))$$

Confidence Interval for Mean:  $\overline{X} \pm Z_{\alpha/2} \cdot \frac{\sigma}{\sqrt{n}}$ 

#### 4.7 T-test

: Single Sample T-test:

 $T = \frac{\overline{X} - \mu}{S/\sqrt{n}}$ 

 $\Rightarrow$  p-value =  $2 \cdot (1 - t_{n-1}(|T|))$ 

2-Sample T-test:

 $T = \frac{\overline{X}_1 - \overline{X}_2}{\sqrt{\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}}}$ 

 $\Rightarrow$  p-value =  $2 \cdot (1 - t_{n_1+n_2-2}(|T|))$ 

2-Sample T-test (Unequal Variance):

 $\Rightarrow \text{p-value} = 2 \cdot (1 - t_{\nu}(|T|))$ 

# 4.8 Chi-Square Test:

Goodness of Fit Test:  $\chi^2 = \sum_{i=1}^k \tfrac{(O_i - E_i)^2}{E_i}$ 

 $\Rightarrow$  p-value =  $1 - \chi_{k-1}^2(\chi^2)$ 

#### Fisher's Exact Test:

 $H_0$ : The two samples are from the same distribution

 $H_1$ : The two samples are from different distributions

p value: getting the observed data or more ex-

treme data

## Permutation Test:

 $H_0$ : The two samples are from the same distribution

 $H_1$ : The two samples are from different distributions

consider an arbitrary test statistic T

permute the labels of the samples to get the

distribution of T under  $H_0$ 

then calculate the p-value as the probability of observing the data or more extreme data under

 $H_0$ 

# Beysian Inference:

MAP:

 $\hat{\theta}_{MAP} = \arg \max_{\theta} f(\theta || X) = \arg \max_{\theta} f(X || \theta) \cdot f(\theta)$ 

## 4.12 Linear Regression:

Simple Linear Regression:  $Y = \beta_0 + \beta_1 X + \epsilon$ where  $\epsilon \sim \mathcal{N}(0, \sigma^2)$ 

 $\beta_1$  and  $\beta_0$  can be estimated using the Maximum Likelihood Estimation

$$\begin{array}{l} \hat{\beta}_1 = \frac{\sum_{i=1}^n (X_i - \overline{X})(Y_i - \overline{Y})}{\sum_{i=1}^n (X_i - \overline{X})^2} = \frac{\operatorname{Cov}(X,Y)}{\operatorname{Var}(X)} \\ \hat{\beta}_0 = \overline{Y} - \hat{\beta}_1 \overline{X} \\ \hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (Y_i - \hat{\beta}_0 - \hat{\beta}_1 X_i)^2 \end{array}$$