Linear Regression / LDA

Linear Methods for Classification

2019.1.30 임소현

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Classification?

: 입력 벡터 x로부터 이에 대응되는 타겟 클래스 K에 대해 어떤 하나의 클래스에 속하도록 선정하는 작업

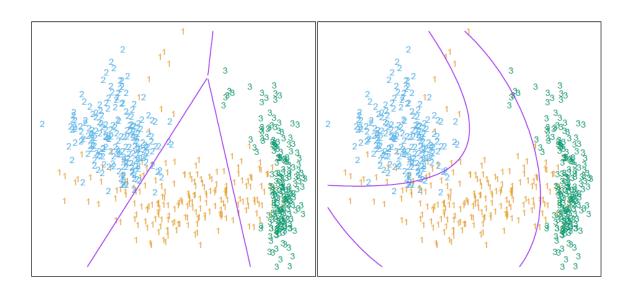
■ 타겟 t에 대해 K의 크기를 가지는 이진 벡터로 정의

ex > K=5, 클래스가 2에 속하는 경우

$$\mathbf{t} = (0, 1, 0, 0, 0)^T$$

4.1 Introduction

First Second **Third Fourth**



- 클래스는 상호 배타적인 관계
- 이렇게 나누어진 지역을 *decision region* 라고 부른다 이를 나누는 경계면을 *decision boundaries* 라고 부른다

Model

- discriminant functions $\delta_k(x)$ for each class
- Posterior probabilities Pr(G = k|X = x)

$$\Pr(G = 1|X = x) = \frac{\exp(\beta_0 + \beta^T x)}{1 + \exp(\beta_0 + \beta^T x)},$$
$$\Pr(G = 2|X = x) = \frac{1}{1 + \exp(\beta_0 + \beta^T x)}.$$

logit transformation

$$\log \frac{\Pr(G = 1 | X = x)}{\Pr(G = 2 | X = x)} = \beta_0 + \beta^T x$$

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• there will be K such indicators Y_k , k = 1, ..., K, with $Y_k = 1$ if G = k else 0

- $Y = (Y_1, ..., Y_K)$, indicator response matrix ($N \times K$)
- Y is a matrix of 0's and 1's, with each row having a single 1
- Fit a linear regression model

$$\hat{\mathbf{Y}} = \mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{Y}$$

Coefficient matrix

$$\hat{\mathbf{B}} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{Y}$$

- A new observation with input x is classified as follows:
 - compute the fitted output $\hat{f}(x)^T = (1, x^T)\hat{\mathbf{B}}$, a K vector;
 - identify the largest component and classify accordingly: $\hat{G}(x) = \operatorname{argmax}_{k \in \mathcal{G}} \hat{f}_k(x)$
- The response vector y_i (ith row of \mathbf{Y}) for observation i has the value $y_i = t_k$ if $g_i = k$
- Fit the linear model by least squares $\min_{\mathbf{B}} \sum_{i=1}^{N} ||y_i [(1, x_i^T)\mathbf{B}]^T||^2$
- The criterion is a sum-of-squared Euclidean distances of the fitted vectors from their targets

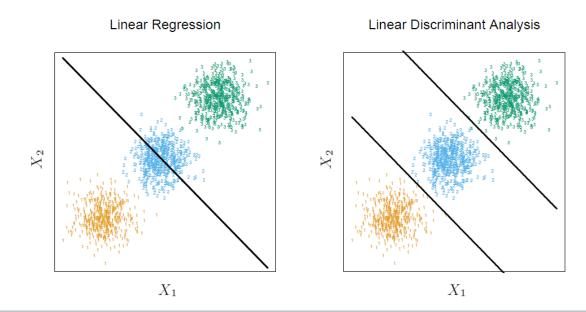
$$\hat{G}(x) = \underset{k}{\operatorname{argmin}} ||\hat{f}(x) - t_k||^2$$

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- First Second
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- There is a serious problem with the regression approach when the number of classes $K \ge 3$
- Figure illustrates an extreme situation when K = 3.
- The three classes are perfectly separated by linear decision boundaries,
 yet linear regression misses the middle class completely.



• Suppose $f_k(x)$ is the class conditional density of X in class G=k

$$f_k(x) = \frac{1}{(2\pi)^{p/2} |\mathbf{\Sigma}_k|^{1/2}} e^{-\frac{1}{2}(x-\mu_k)^T \mathbf{\Sigma}_k^{-1}(x-\mu_k)}$$

- Let π_k be the prior probability of class k, with $\sum_{k=1}^K \pi_k = 1$
- By Bayes theorem,

$$\Pr(G = k | X = x) = \frac{f_k(x)\pi_k}{\sum_{\ell=1}^{K} f_{\ell}(x)\pi_{\ell}}$$

4.3 Linear Discriminant Analysis

LDA

- assume that the classes have a common covariance matrix
- Comparing two classes *k* and *l*, log-ratio

$$\log \frac{\Pr(G = k | X = x)}{\Pr(G = \ell | X = x)} = \log \frac{f_k(x)}{f_\ell(x)} + \log \frac{\pi_k}{\pi_\ell}$$
$$= \log \frac{\pi_k}{\pi_\ell} - \frac{1}{2} (\mu_k + \mu_\ell)^T \mathbf{\Sigma}^{-1} (\mu_k - \mu_\ell)$$
$$+ x^T \mathbf{\Sigma}^{-1} (\mu_k - \mu_\ell),$$

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LDA

linear discriminant functions

$$\delta_k(x) = x^T \mathbf{\Sigma}^{-1} \mu_k - \frac{1}{2} \mu_k^T \mathbf{\Sigma}^{-1} \mu_k + \log \pi_k$$

- We don't know the parameters of the Gaussian distributions, so we need to estimate them
 - $\hat{\pi}_k = N_k/N$, where N_k is the number of class-k observations;
 - $\hat{\mu}_k = \sum_{g_i=k} x_i/N_k;$
 - $\hat{\Sigma} = \sum_{k=1}^{K} \sum_{g_i=k} (x_i \hat{\mu}_k) (x_i \hat{\mu}_k)^T / (N K).$

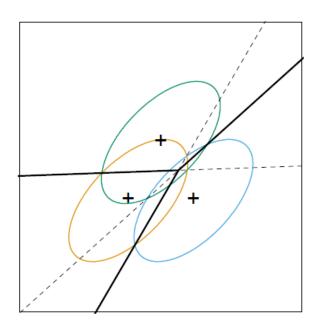
4.3 Linear Discriminant Analysis

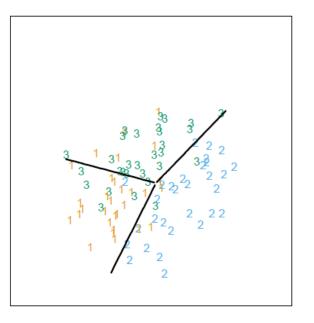
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- shows three Gaussian distributions, with the same covariance and different means
- included are the contours of constant density enclosing 95% of the probability in each case
- the Bayes decision boundaries separating all three classes are the thicker solid lines

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QDA (quadratic discriminant functions)

- if the Σ_k are not assumed to be equal
- linear discriminant functions

$$\delta_k(x) = -\frac{1}{2}\log|\mathbf{\Sigma}_k| - \frac{1}{2}(x - \mu_k)^T \mathbf{\Sigma}_k^{-1}(x - \mu_k) + \log \pi_k$$

■ The decision boundary between each pair of classes k and l is described

by a quadratic equation
$$\{x: \delta_k(x) = \delta_\ell(x)\}$$

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RDA (Regularized Discriminant Analysis)

- to shrink the separate covariances of QDA toward a common covariance as in LDA
- very similar in flavor to ridge regression
- $\alpha \in [0,1]$ allows a continuum of models between LDA and QDA

$$\hat{\Sigma}_k(\alpha) = \alpha \hat{\Sigma}_k + (1 - \alpha) \hat{\Sigma}_k \qquad (4.13)$$

Similar modifications to be shrunk toward the scalar covariance

$$\hat{\mathbf{\Sigma}}(\gamma) = \gamma \hat{\mathbf{\Sigma}} + (1 - \gamma)\hat{\sigma}^2 \mathbf{I}$$

• Replacing $\hat{\Sigma}$ in (4.13) by $\hat{\Sigma}(\gamma)$ leads to a more general family

- http://norman3.github.io/prml/docs/chapter04/1
- https://m.blog.naver.com/PostView.nhn?blogId=sw4r&logNo=221033110991&proxy Referer=https%3A%2F%2Fwww.google.com%2F

Thank you :-)