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-revision materials for linear regression

Given the linear regression equation:

$$\hat{y} = \beta_0 + \beta_1 x + e$$

Assuming we have $\{(x_i, y_i), i=1, \dots, n\}$
 To derive the least squares we do the following

let	x_i	y_i
	1	1
	2	2
	3	3
	1	1
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$\Rightarrow y = \beta_1 x + (\beta_0 + e)$

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$$= \sum_{i=1}^n (y_i - \beta_0 - \beta_1 x_i - e - \sum_{j=1}^p \beta_j x_{ij})^2$$

$$Y = \beta_0 + \beta_1 x + \dots + \beta_p x_p + e$$

★

$$\rightarrow - \sum_{i=1}^n (y_i - \beta_0 - \beta_1 x_i - e)^2$$

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$\Rightarrow = \sum_{i=1}^n (y_i - \beta_0 - \beta_1 x - \epsilon)^2$

(ii) Including a ridge penalty we have

$\sum_{i=1}^n (y_i - \beta_0 - \beta_1 x - \epsilon) + \lambda \sum_{i=1}^p \beta_i^2$

Last term is the ridge penalty we are adding into a function.
 $\beta_1^2 + \beta_0^2 + \epsilon^2 \leq t$

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In General we have:

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Last term is the ridge penalty we are adding into a function.

$$\beta_1^2 + \beta_0^2 + \varepsilon^2 \leq t$$

$$= \sum_{i=1}^n (y_i - \beta_0 - \beta_1 x_i - \varepsilon)^2 - \sum_{j=1}^p \beta_j^2 x_{ij}^2$$

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In General we have;

$$h_0(x_i) = \beta_0 + \beta_1 x_i + \varepsilon + \lambda \sum_{j=1}^p \beta_j^2$$

(c) Assuming we extend the model to

$$y = \beta_0 + \beta_1 x + \beta_2 x^2 + \varepsilon$$

Where $x = x^2$

Reasons

- The estimator implies that among the possible unbiased estimation, the OLS achieves the minimum variance which gets a non-unique value for the OLS estimator.
- The obtained estimates are highly unstable despite being unbiased thus making the solution to have a non-unique value.

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(c) Assuming we extend the model to
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(d) Now adding penalties to both β_1 and β_2

$$= \sum_{i=1}^n (y_i - \beta_0 x - \sum_{j=1}^p \beta_j x_{ij})^2 + \lambda [\sum_{j=1}^p \beta_j^2 x + \sum \beta_j^2 z]$$

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$$= \sum_{i=1}^n y_i - \beta_0 x + \lambda \sum_{j=1}^p (\beta_j^2 + \beta_j^2) - \left(\sum_{j=1}^p \beta_j x_{ij} \right)^2$$

$$= \sum_{i=1}^n (y_i - \beta_0 - \beta_1 x - \beta_2 x - \epsilon) + \lambda \sum_{j=1}^p \beta_j^2$$

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Discussion

- These estimates and predictors are in the different set score parametrization.

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Discussion

- These estimates and predictors are in the different set score parametrization.

- This implies that the penalization $\lambda \sum_{j=1}^p \beta_j^2$ would have different impact in the predictors.

- Thus we always make calculations with use of standardized version of the obtained predictors.

- So in particular when there is a high collinearity (high correlation between the predictors) or the

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$$\begin{aligned} &= \sum_{i=1}^n y_i - \beta_0 x + \lambda \sum_{j=1}^p (\beta_j^2 + \beta_0^2) - \left(\sum_{j=1}^p \beta_j x_{ij} \right)^2 \\ &= \sum_{i=1}^n (y_i - \beta_0 - \beta_1 x - \beta_0 x - \varepsilon) + \lambda \sum_{j=1}^p \beta_j^2 \end{aligned}$$

e) Discussion

- These estimates and predictors are in the different set scale parametrization.
- This implies that the penalization $\lambda \sum_{j=1}^p \beta_j^2$ would have different impact in the predictors.
- Thus we always make calculations with use of standardized version of the obtained predictor.
- So in particular, when there is a high collinearity (high correlation between the predictors) or the number of predictors (p) is of similar magnitude to the number of observation (n).
- Note that λ is a tuning parameter that is controlling the amount of penalisation.