

```
# 1.1

library("quantmod")

r <- read.table("http://gamma.mini.pw.edu.pl/~szymanowski/lab1/USPOP.DATA")
head(r,3)

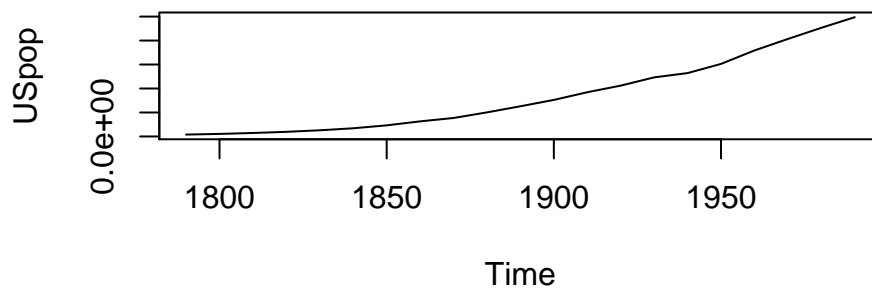
##          V1
## 1 3929214
## 2 5308483
## 3 7239881

# a)

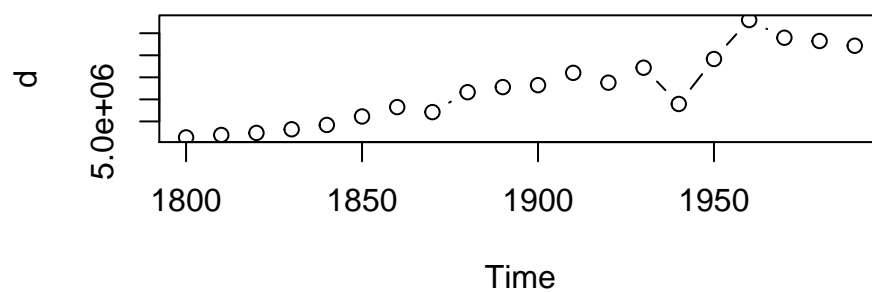
USpop <- ts(data=r, start=1790, end=1990, frequency=0.1)

# b)

ts.plot(USpop)
```



```
d <- diff(USpop)
ts.plot(d,type="b")
```



```
which.min(d) # najmniejszy wzrost w latach 1790-1800

## [1] 1

# c)

czas <- seq(start(USpop), end(USpop), by=1/frequency(USpop))
```

```

m1.lm <- lm(USpop ~ czas + I(czas^2))
summary(m1.lm)

##
## Call:
## lm(formula = USpop ~ czas + I(czas^2))
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -6947521  -358167   436285  1481410  3391761
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  2.10e+10   6.59e+08   31.9   <2e-16 ***
##      czas    -2.34e+07   6.98e+05  -33.5   <2e-16 ***
## I(czas^2)     6.51e+03   1.85e+02   35.2   <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 2770000 on 18 degrees of freedom
## Multiple R-squared:  0.999, Adjusted R-squared:  0.999
## F-statistic: 8.05e+03 on 2 and 18 DF,  p-value: <2e-16

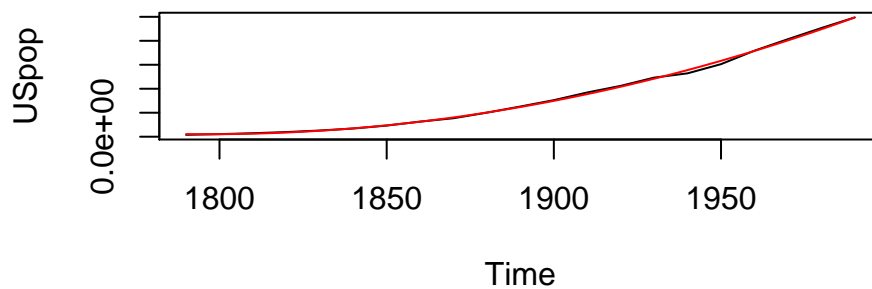
wsp <- summary(m1.lm)$coefficients[,1]

```

```

ts.plot(USpop)
curve(wsp[1]+wsp[2]*x+wsp[3]*x^2,from=1790, to=1990,add=TRUE,col="red")

```

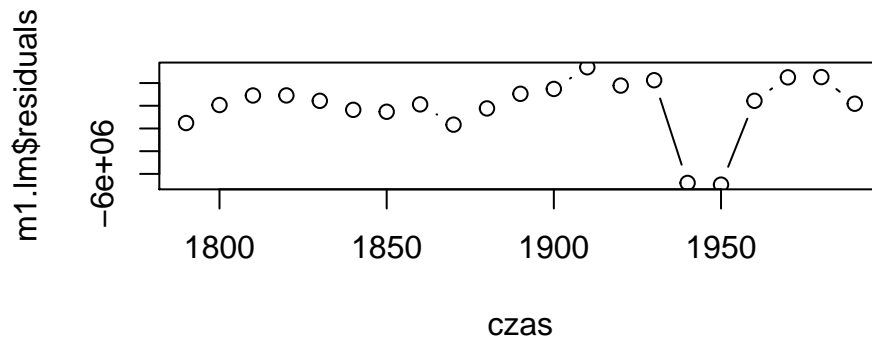


```

# wykres reziduow od czasu:

plot(czas, m1.lm$residuals, type="b")

```



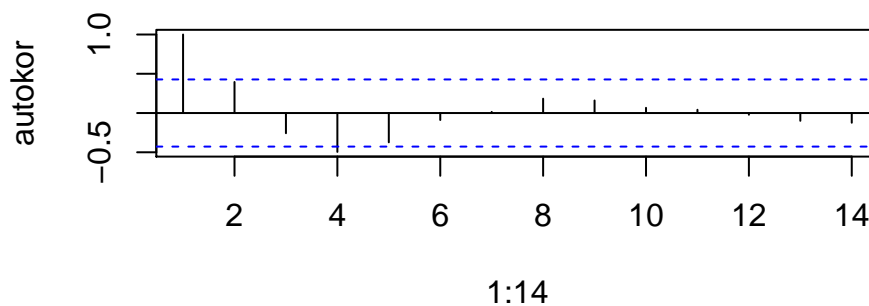
```
# d)

res <- m1.lm$residuals
USres <- ts(res, start=1790, end=1990, frequency=0.1)

n <- length(USres)
sr <- mean(res)
autokor <- numeric(14)
gamma0 <- (sum((res-sr)^2))/n
autokor[1] <- 1

for(i in 1:13){
  gammah <- (sum((res[1:(21-i)]-sr)*(res[(i+1):21]-sr)))/n
  autokor[i+1] <- gammah/gamma0
}

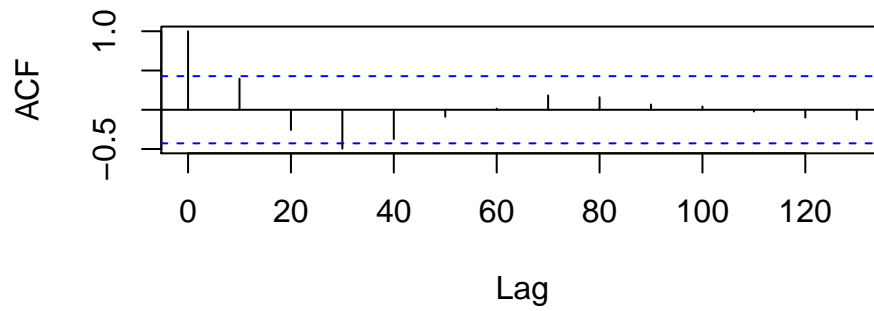
plot(1:14,autokor,type="h")
kw <- qnorm(0.975)/sqrt(n)
abline(b=0,a=kw,lty=2,col="blue")
abline(b=0,a=-kw,lty=2,col="blue")
abline(b=0,a=0)
```



```
# funkcja wbudowana rysujaca ten sam wykres:
```

```
acf(USres) # jest autokorelacja rzędu 3 -> testujemy formalnie, czy jest korelacja
```

## Series USres



```
# e)
# H0: pierwsze 10 korelacji jest rowne zero
# lag := h

Box.test(USres, lag = 10, type="Ljung")

##
## Box-Ljung test
##
## data: USres
## X-squared = 18.6, df = 10, p-value = 0.04565

# czyli rezidua nie sa bialym szumem niestety

# 1.2

data(LakeHuron)
head(LakeHuron)

## [1] 580.4 581.9 581.0 580.8 579.8 580.4

is.ts(LakeHuron)

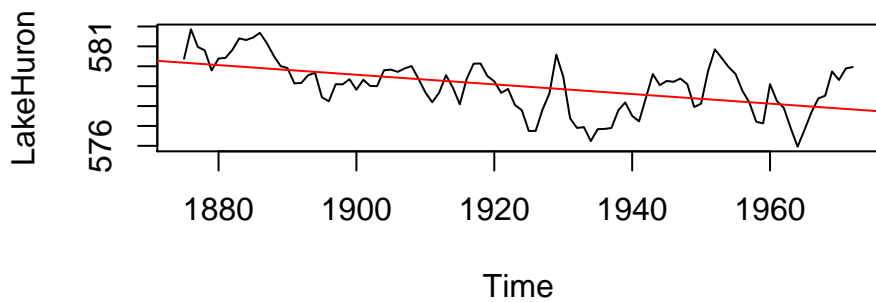
## [1] TRUE
```

```
# a)

ts.plot(LakeHuron) # widac tendencje malejaca

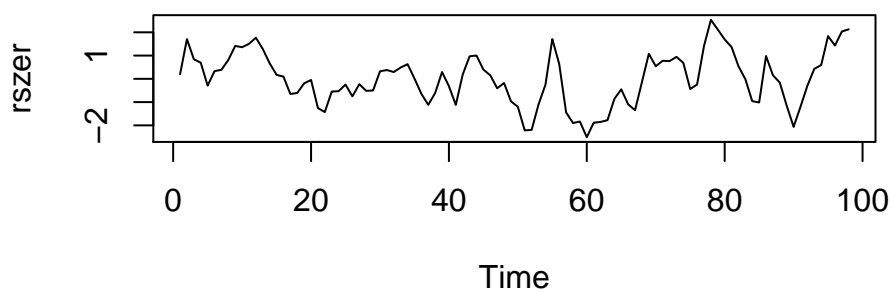
# b)

czas <- as.vector(time(LakeHuron))
lake.lm <- lm(LakeHuron ~ czas, data=LakeHuron)
abline(lake.lm, col="red")
```



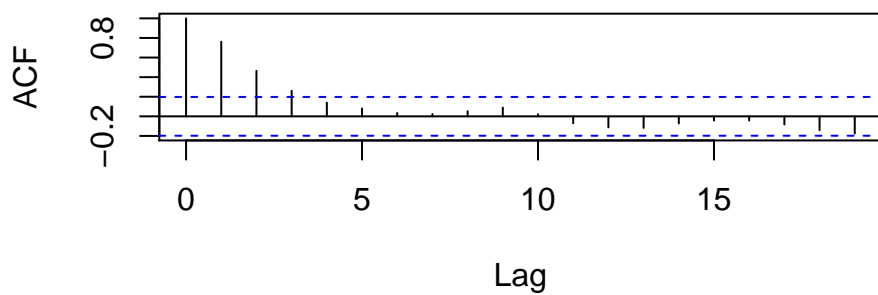
```
# c)

res <- lake.lm$residuals
rszer <- ts(data=res)
ts.plot(rszer)
```



```
acf(rszer) # pasowaloby MA(3)
```

### Series rszer



```
Box.test(rszer, lag = 20, type="Ljung")

##
## Box-Ljung test
##
## data: rszer
## X-squared = 107.8, df = 20, p-value = 4.885e-14
```

```

# to zdecydowanie nie jest biały szum!

# d)

r0 <- res
r1 <- Lag(res,k=1)
r2 <- Lag(res,k=2)

head(r0)

##          1          2          3          4          5          6
##  0.2022  1.7064  0.8406  0.6948 -0.2910  0.3332

head(r1)

##          Lag.1
## [1,]          NA
## [2,]  0.2022
## [3,]  1.7064
## [4,]  0.8406
## [5,]  0.6948
## [6,] -0.2910

head(r2)

##          Lag.2
## [1,]          NA
## [2,]          NA
## [3,]  0.2022
## [4,]  1.7064
## [5,]  0.8406
## [6,]  0.6948

# e)

l <- lm(r0~r1)
summary(l)

##
## Call:
## lm(formula = r0 ~ r1)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -1.9588 -0.4993  0.0017  0.4178  1.8956
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   0.0153     0.0727   0.21    0.83
## r1             0.7911     0.0659  12.00 <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.716 on 95 degrees of freedom
## (1 observation deleted due to missingness)
## Multiple R-squared:  0.603, Adjusted R-squared:  0.599
## F-statistic: 144 on 1 and 95 DF, p-value: <2e-16

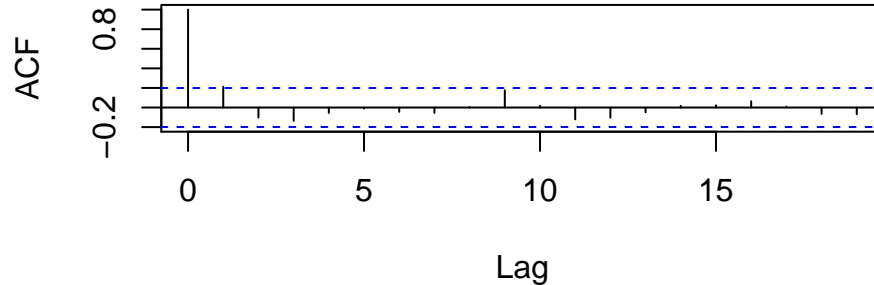
```

```
mean(l$residuals^2)      # MSE
```

```
## [1] 0.5022
```

```
acf(l$res)
```

## Series l\$res



```
Box.test(l$res,lag=20,type="Ljung")
```

```
##
```

```
## Box-Ljung test
```

```
##
```

```
## data: l$res
```

```
## X-squared = 19.71, df = 20, p-value = 0.4765
```

```
# f)
```

```
l2 <- lm(r0~r1+r2)
```

```
summary(l2)
```

```
##
```

```
## Call:
```

```
## lm(formula = r0 ~ r1 + r2)
```

```
##
```

```
## Residuals:
```

```
##      Min       1Q   Median       3Q      Max
```

```
## -1.5843 -0.4525 -0.0162  0.4030  1.7320
```

```
##
```

```
## Coefficients:
```

```
##              Estimate Std. Error t value Pr(>|t|)
```

```
## (Intercept) -0.00785    0.06912   -0.11   0.9098
```

```
## r1           1.00214    0.09722   10.31  <2e-16 ***
```

```
## r2          -0.28380    0.09900   -2.87   0.0051 **
```

```
## ---
```

```
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
##
```

```
## Residual standard error: 0.677 on 93 degrees of freedom
```

```
## (2 observations deleted due to missingness)
```

```
## Multiple R-squared:  0.644, Adjusted R-squared:  0.636
```

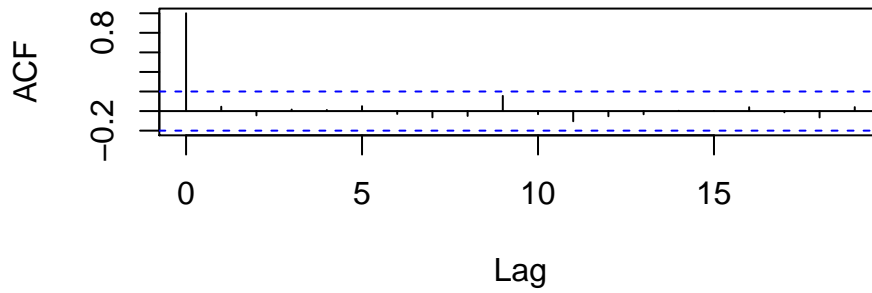
```
## F-statistic: 84.2 on 2 and 93 DF, p-value: <2e-16
```

```
mean(l2$residuals^2)
```

```
## [1] 0.4435
```

```
acf(l2$res)
```

## Series l2\$res



```
Box.test(l2$res,lag=20,type="Ljung") # lepiej dopasowany niz l
```

```
##
## Box-Ljung test
##
## data: l2$res
## X-squared = 9.033, df = 20, p-value = 0.9825
```

```
# g)
```

```
model_ar1 <- arima(r0,c(1,0,0))
model_ar2 <- arima(r0,c(2,0,0))
```

```
Box.test(model_ar1$residuals,lag=20,type="Ljung")
```

```
##
## Box-Ljung test
##
## data: model_ar1$residuals
## X-squared = 19.92, df = 20, p-value = 0.4631
```

```
Box.test(model_ar2$residuals,lag=20,type="Ljung")
```

```
##
## Box-Ljung test
##
## data: model_ar2$residuals
## X-squared = 8.466, df = 20, p-value = 0.9883
```

```
# h)
```

```
model_ar1
```

```
##
## Call:
## arima(x = r0, order = c(1, 0, 0))
##
## Coefficients:
##      ar1  intercept
##    0.783    0.080
## s.e. 0.063    0.318
##
## sigma^2 estimated as 0.497: log likelihood = -105.3, aic = 216.6
```



```

l$coefficients      # ok, w miare podobne

## (Intercept)      r1
##      0.01529      0.79112

# 1.3

d <- read.table("C:\\Users\\Marta\\Desktop\\Marta\\studia\\rok4\\Szeregi czasowe\\DOWJ.DAT")
head(d)

##      V1
## 1 110.9
## 2 110.7
## 3 110.4
## 4 110.6
## 5 110.8
## 6 110.8

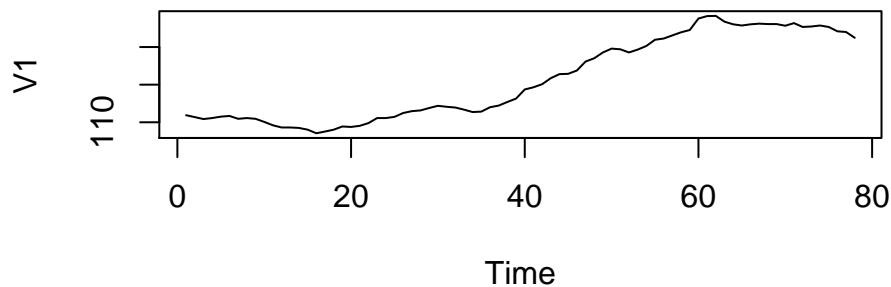
# a)

t <- ts(data=d)

# b)

plot(t)      # widoczny jest trend!!!

```

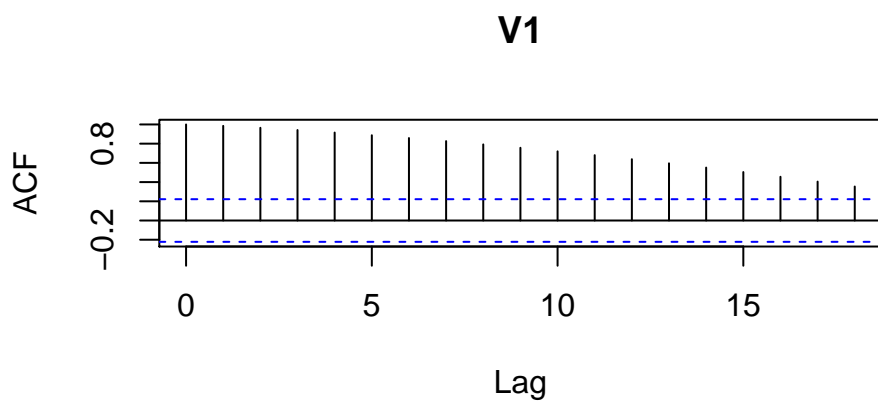


```

# c)

acf(t)      # szereg niestacjonarny -> zeby sie pozbyc trendu roznicujemy

```

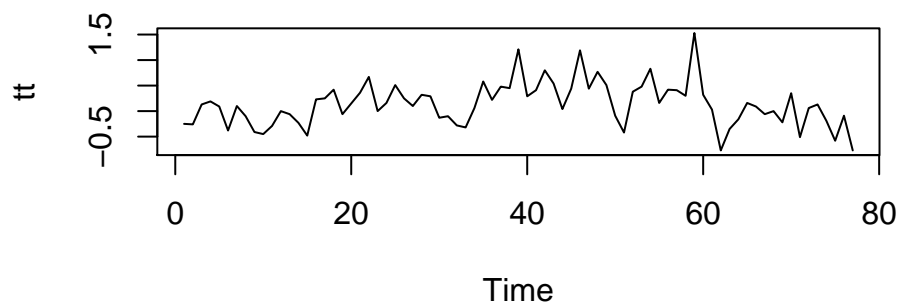


```
# d)

dd <- diff(as.numeric(as.matrix(d)))
tt <- ts(data=dd)

# e)

ts.plot(tt)
```



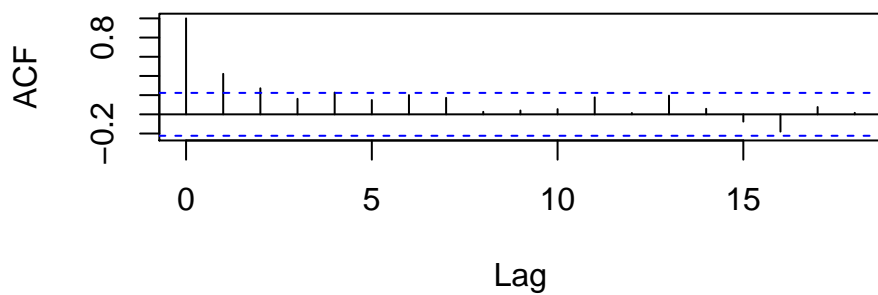
```
Box.test(tt,lag=20,type="Ljung")

##
## Box-Ljung test
##
## data: tt
## X-squared = 46.43, df = 20, p-value = 0.0007039

# f)

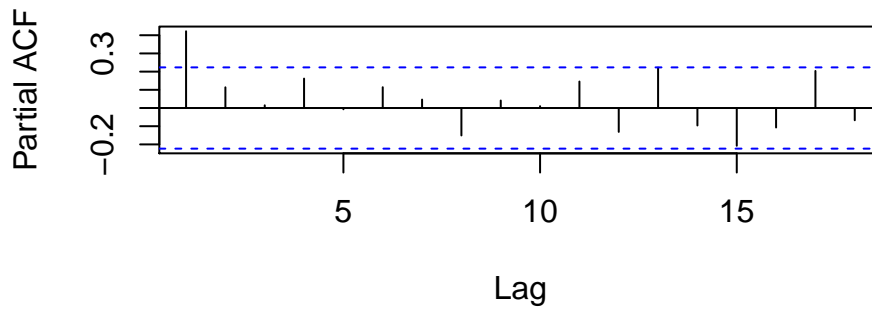
acf(tt) # sugeruje model MA(2)
```

## Series tt



```
pacf(tt) # sugeruje model AR(1)
```

## Series tt



```
ma2 <- arima(dd,c(0,0,2))
ar1 <- arima(dd,c(1,0,0))
arma12 <- arima(dd,c(1,0,2))

Box.test(ma2$res,lag=20,type="Ljung")

##
## Box-Ljung test
##
## data:  ma2$res
## X-squared = 31.4, df = 20, p-value = 0.05011

Box.test(ar1$res,lag=20,type="Ljung")

##
## Box-Ljung test
##
## data:  ar1$res
## X-squared = 32.55, df = 20, p-value = 0.03774

Box.test(arma12$res,lag=20,type="Ljung")

##
## Box-Ljung test
##
## data:  arma12$res
## X-squared = 27.95, df = 20, p-value = 0.1105

acf(arma12$residuals)
```

## Series arma12\$residuals

