Theory for Minimum Norm Interpolation: Regression and Classification in High Dimensions

Tengyuan Liang



Classification: with Pragya Sur (Harvard) Regression: with Sasha Rakhlin (MIT), Xiyu Zhai (MIT)

OUTLINE

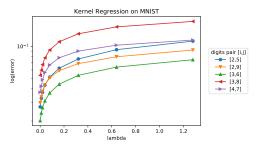
- Motivation: min-norm interpolants
- Regression: multiple descent of risk
- Classification: boosting on separable data

OUTLINE

- Motivation: min-norm interpolants
- Regression: multiple descent of risk
 - application to wide neural networks
 - restricted lower isometry of kernels
 - small-ball property
- Classification: boosting on separable data
 - precise high-dim asymptotics
 - · convex Gaussian min-max theorem
 - · algorithmic implications on boosting

OVER-PARAMETRIZED REGIME OF STAT/ML

Model class complex enough to interpolate the training data.

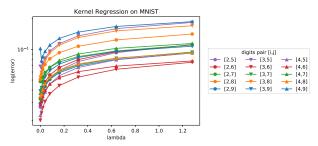


 λ = 0: the interpolants on training data.

MNIST data from LeCun et al. (2010)

OVER-PARAMETRIZED REGIME OF STAT/ML

Model class complex enough to interpolate the training data.



 λ = 0: the interpolants on training data.

MNIST data from LeCun et al. (2010)

Intro. Min-norm Interpolant Regression Classification

OVER-PARAMETRIZED REGIME OF STAT/ML

Model class complex enough to interpolate the training data.

Zhang, Bengio, Hardt, Recht, and Vinyals (2016)

In fact, many models behave the same on training data.



Practical methods or algorithms favor certain functions!

Principle: among the models that **interpolate**, algorithms favor certain form of **minimalism**.

OVER-PARAMETRIZED REGIME OF STAT/ML

Principle: among the models that **interpolate**, algorithms favor certain form of **minimalism**.

- · over-parametrized linear model and matrix factorization
- · kernel machines
- support vector machines
- boosting, AdaBoost
- two-layer ReLU networks

ntro. Min-norm Interpolant Regression Classification

OVER-PARAMETRIZED REGIME OF STAT/ML

Principle: among the models that **interpolate**, algorithms favor certain form of **minimalism**.

- over-parametrized linear model and matrix factorization
- kernel machines
- support vector machines
- boosting, AdaBoost
- two-layer ReLU networks

minimalism typically measured in form of certain norm motivates the study of min-norm interpolants

MIN-NORM INTERPOLANTS

minimalism typically measured in form of certain norm motivates the study of min-norm interpolants

Regression

$$\widehat{f} = \underset{f}{\operatorname{arg \, min}} \|f\|_{\operatorname{norm}}, \text{ s.t. } y_i = f(x_i) \ \forall i \in [n].$$

Classification

$$\widehat{f} = \underset{f}{\operatorname{arg\,min}} \ \|f\|_{\operatorname{norm}}, \ \ \text{s.t.} \ \ y_i \cdot f(x_i) \geq 1 \ \forall i \in [n].$$

REGRESSION

Multiple Descent of Minimum-Norm Interpolants and Restricted Lower Isometry of Kernels

with Sasha Rakhlin (MIT), Xiyu Zhai (MIT)

Classic: U-shape curve

Recent: double descent curve

Belkin, Hsu, Ma, and Mandal (2018); Hastie, Montanari, Rosset, and Tibshirani (2019)

Question: shape of the risk curve w.r.t. "over-parametrization"?

Classic: U-shape curve

Recent: double descent curve

Belkin, Hsu, Ma, and Mandal (2018); Hastie, Montanari, Rosset, and Tibshirani (2019)

Question: shape of the risk curve w.r.t. "over-parametrization"?

We model the **intrinsic dim.** $d = n^{\alpha}$ with $\alpha \in (0, 1)$, with feature cov. $\Sigma_d = I_d$. We consider the **non-linear Kernel Regression** model.

We consider the **intrinsic dim.** $d = n^{\alpha}$ with $\alpha \in (0, 1)$.

A non-linear Kernel Regression model.

DGP.

- $\{x_i\}_{i=1}^n \overset{i.i.d}{\sim} \mu = \mathcal{P}^{\otimes d}$. distribution of each coordinate $\mathbf{x} \sim \mathcal{P}$ satisfies weak moment $\forall t > 0, \mathbb{P}(|\mathbf{x}| > t) \le C(1 + t)^{-\gamma}$.
- target $f_*(x) := \mathbb{E}[Y|X = x]$, with bounded Var[Y|X = x].

Kernel.

- $h \in C^{\infty}(\mathbb{R}), h(t) = \sum_{i=0}^{\infty} \alpha_i t^i \text{ with } \alpha_i \geq 0.$
- inner product kernel $k(x, z) = h(\langle x, z \rangle/d)$.

Target Function.

• Assume $f_{\star}(x) = \int k(x,z) \rho_{\star}(z) \mu(dz)$ with $\|\rho_{\star}\|_{\mu} \leq C$.

We consider the **intrinsic dim.** $d = n^{\alpha}$ with $\alpha \in (0, 1)$.

A non-linear Kernel Regression model.

Given *n* i.i.d. data pairs $(x_i, y_i) \sim \mathcal{P}_{X,Y}$.

Risk curve for minimum RKHS norm $\|\cdot\|_{\mathcal{H}}$ interpolants \widehat{f} ?

$$\widehat{f} = \underset{f}{\operatorname{arg\,min}} \ \|f\|_{\mathcal{H}}, \ \text{ s.t. } y_i = f(x_i) \ \forall i \in [n].$$

Theorem (L., Rakhlin & Zhai, '19).

For any integer $\iota \ge 1$, consider $\frac{d}{d} = n^{\alpha}$ where $\alpha \in \left(\frac{1}{\iota + 1}, \frac{1}{\iota}\right)$.

Theorem (L., Rakhlin & Zhai, '19).

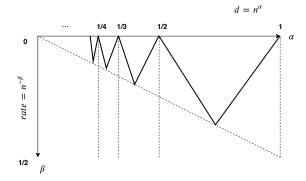
For any integer $\iota \geq 1$, consider $d = n^{\alpha}$ where $\alpha \in (\frac{1}{\iota+1}, \frac{1}{\iota})$.

With probability at least $1 - \delta - e^{-n/d^{1}}$ on the design $\mathbf{X} \in \mathbb{R}^{n \times d}$,

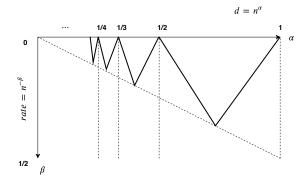
$$\mathbb{E}\left[\|\widehat{f}-f_*\|_{\mu}^2|\mathbf{X}\right] \leq C \cdot \left(\frac{d^{\mathsf{t}}}{n} + \frac{n}{d^{\mathsf{t}+1}}\right) \approx n^{-\beta},$$

$$\beta := \min \left\{ (\iota + 1)\alpha - 1, 1 - \iota \alpha \right\}.$$

Here the constant $C(\delta, \iota, h, P)$ does not depend on d, n.

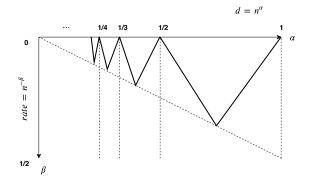


multiple-descent behavior of the rates as the scaling $d = n^{\alpha}$ changes.



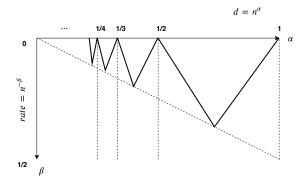
multiple-descent behavior of the rates as the scaling $d = n^{\alpha}$ changes.

• **valley**: "valley" on the rate curve at $d = n^{\frac{1}{\iota + 1/2}}$, $\iota \in \mathbb{N}$



multiple-descent behavior of the rates as the scaling $d = n^{\alpha}$ changes.

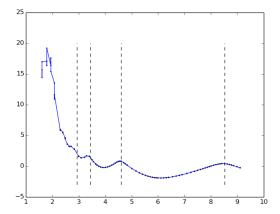
- **valley**: "valley" on the rate curve at $d = n^{\frac{1}{\iota + 1/2}}$, $\iota \in \mathbb{N}$
- **over-parametrization**: towards over-parametrized regime, the good rate at the bottom of the valley is better



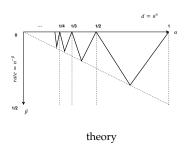
multiple-descent behavior of the rates as the scaling $d = n^{\alpha}$ changes.

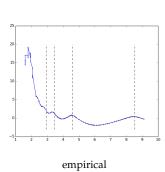
- **valley**: "valley" on the rate curve at $d = n^{\frac{1}{1+1/2}}$, $\iota \in \mathbb{N}$
- over-parametrization: towards over-parametrized regime, the good rate at the bottom of the valley is better
- empirical: preliminary empirical evidence of multiple descent

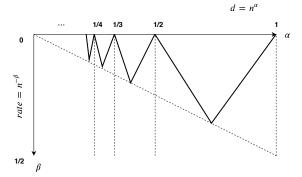
EMPIRICAL EVIDENCE



empirical evidence of **multiple-descent behavior** as the scaling $d = n^{\alpha}$ changes.







multiple-descent behavior of the rates as the scaling $d = n^{\alpha}$ changes.

- $\alpha = 1$: Liang and Rakhlin (2018)
- $\alpha = 0$: Rakhlin and Zhai (2018)
- α = 1 double descent: Belkin, Hsu, Ma, and Mandal (2018); Hastie, Montanari, Rosset, and Tibshirani (2019); Bartlett, Long, Lugosi, and Tsigler (2019)
- general α, stair-case, random fourier feature: Ghorbani, Mei, Misiakiewicz, and Montanari (2019)

APPLICATION TO WIDE NEURAL NETWORKS

Neural Tangent Kernel (NTK)

Jacot, Gabriel, and Hongler (2018); Du, Zhai, Poczos, and Singh (2018).....

$$k_{\mathrm{NTK}}(x,x') = \frac{1}{4\pi} U\Big(\frac{\langle x,x'\rangle}{\|x\|\|x'\|}\Big)$$

$$U(t) = 3t(\pi - \arccos(t)) + \sqrt{1 - t^2}$$

APPLICATION TO WIDE NEURAL NETWORKS

Neural Tangent Kernel (NTK)

Jacot, Gabriel, and Hongler (2018); Du, Zhai, Poczos, and Singh (2018).....

$$k_{\mathrm{NTK}}(x,x') = \frac{1}{4\pi} U\Big(\frac{\left\langle x,x'\right\rangle}{\|x\|\|x'\|}\Big)$$

$$U(t) = 3t(\pi - \arccos(t)) + \sqrt{1 - t^2}$$

Corollary (L., Rakhlin & Zhai, '19).

Our results can be generalized to the following type of kernels

$$k(x,x') = \sum_{i=0}^{\infty} \alpha_i \cdot \left(\frac{\langle x,x' \rangle}{\|x\| \|x'\|}\right)^i$$

that include NTK.

Consider integer ι that satisfies $d^{\iota} \log d \lesssim n \lesssim d^{\iota+1}/\log d$, then

$$\operatorname{Risk} \lesssim \frac{d^{\iota}}{n} + \frac{n \log d}{d^{\iota+1}}$$

IDEAS BEHIND THE PROOF

Proof Idea: on a **filtration of spaces** indexed by polynomial basis, establish restricted lower isometry of the empirical kernel.

filtrated empirical kernel

$$n\mathbf{K}_{ij}^{[\leq t]} := \sum_{\substack{r_1, \dots, r_d \geq 0 \\ r_1 + \dots + r_d \leq t}} c_{r_1 \dots r_d} \alpha_{r_1 + \dots + r_d} p_{r_1 \dots r_d}(x_i) p_{r_1 \dots r_d}(x_j) / d^{r_1 + \dots + r_d}$$

$$n\mathbf{K}^{[\leq t]} = \bigoplus_{n \times \binom{t+d}{t}} \bigoplus_{\substack{t + d \\ t}} \binom{t+d}{t} \times n$$

filtrated sample covariance operator

$$\Theta^{\left[\leq\iota\right]} \coloneqq \frac{1}{n} \underbrace{\Phi^{\top}}_{\left(\begin{smallmatrix}\iota+d\\\iota^{+}d\end{smallmatrix}\right)\times n} \underbrace{n\times \left(\begin{smallmatrix}\iota+d\\\iota^{+}d\end{smallmatrix}\right)}_{n\times \left(\begin{smallmatrix}\iota+d\\\iota^{+}d\end{smallmatrix}\right)}$$

Proof Idea: on a **filtration of spaces** indexed by polynomial basis, **establish restricted lower isometry** of the empirical kernel.

filtrated empirical kernel

$$n\mathbf{K}_{ij}^{[\leq \mathsf{t}]} \coloneqq \sum_{\substack{r_1, \dots, r_d \geq 0 \\ r_1 + \dots + r_d \leq \mathsf{t}}} c_{r_1 \dots r_d} \alpha_{r_1 + \dots + r_d} p_{r_1 \dots r_d}(x_i) p_{r_1 \dots r_d}(x_j) / d^{r_1 + \dots + r_d}$$

$$n\mathbf{K}^{[\leq \mathsf{t}]} = \bigoplus_{n \times \binom{\mathsf{t} + d}{\mathsf{t}}} \bigoplus_{\substack{(\mathsf{t} + d) \\ \mathsf{t}}} \binom{\mathsf{t} + d}{\mathsf{t}} \times n$$

filtrated sample covariance operator

$$\Theta^{[\leq \iota]} := \frac{1}{n} \underbrace{\Phi^{\top}}_{\binom{\iota+d}{\iota} \times n} \cdot \underbrace{\Phi}_{n \times \binom{\iota+d}{\iota}}$$

Restricted Lower Isometry of Kernel: all non-zero eigenvalues of $\mathbf{K}^{[\leq t]}$ is lower bounded by d^{-t}

$$\lambda_{\min}\left(\Theta^{\left[\leq\iota\right]}\right)\gtrsim d^{-\iota}$$

IDEAS BEHIND THE PROOF

small-ball approach rather than standard concentration

lower bound $\lambda_{\min}\left(\frac{1}{n}\Psi^{\top}\Psi\right)$ equiv. $\forall u, \|u\| = 1$, lower bound $\|\Psi u\|^2$ utilize non-negativity

$$\|\Psi u\|^2 = \frac{1}{n} \sum_{i=1}^n \langle \Psi(x_i), u \rangle^2 \geq c_1 \mathbb{E} \big[\langle \Psi(x_i), u \rangle^2 \big] \cdot \frac{1}{n} \sum_{i=1}^n I_{\langle \Psi(x_i), u \rangle^2 \geq c_1 \mathbb{E} \big[\langle \Psi(X), u \rangle^2 \big]}$$

small-ball property, \exists constants c_1, c_2

$$\mathbb{P}\left(\langle \Psi(x_i), u \rangle^2 \ge c_1 \mathbb{E}[\langle \Psi(X), u \rangle^2]\right) \ge c_2$$

Koltchinskii and Mendelson (2015); Mendelson (2014)

which will imply w.p. at least $1 - \exp(-c \cdot n)$

$$\frac{1}{n} \sum_{i=1}^{n} I_{\langle \Psi(x_i), u \rangle^2 \ge c_1 \mathbb{E}[\langle \Psi(X), u \rangle^2]} \ge c_2/2$$

Non-trivial: verify small-ball property for polynomials (weakly dependent) via Paley-Zygmund

CLASSIFICATION

Precise High-Dimensional Asymptotic Theory for Boosting and $Min-L_1$ -Norm Interpolated Classifiers

with Pragya Sur (Harvard)

MIN-L₁-NORM INTERPOLATED CLASSIFIER

Regression so far, what about Classification?

Given n-i.i.d. data pairs $\{x_i, y_i\}_{i=1}^n$ with $y_i \in \{\pm 1\}$ being the labels and $x_i \in \mathbb{R}^p$ being feature vectors.

We consider minimum L_1 -norm interpolated classifier:

$$\hat{\theta} = \min_{\theta} \|\theta\|_1$$
, s.t. $y_i x_i^{\mathsf{T}} \theta \ge 1$.

when data is separable.

MIN-L₁-NORM INTERPOLATED CLASSIFIER

Regression so far, what about Classification?

Given *n*-i.i.d. data pairs $\{x_i, y_i\}_{i=1}^n$ with $y_i \in \{\pm 1\}$ being the labels and $x_i \in \mathbb{R}^p$ being feature vectors.

We consider minimum L_1 -norm interpolated classifier:

$$\hat{\theta} = \min_{\theta} \|\theta\|_1, \text{ s.t. } y_i x_i^{\mathsf{T}} \theta \ge 1.$$

when data is separable.

min- L_1 -norm interpolated classifier agrees with the max- L_1 -margin direction

$$\max_{\|\theta\|_1 \leq 1} \min_{1 \leq i \leq n} \ y_i x_i^\top \theta =: \kappa_{\ell_1}(X, y) \ .$$

WHY L_1 MARGIN?

Algorithmic: on separable data, Boosting algorithm $\hat{\theta}_{boost}^{t,\eta}$ with infinitesimal stepsize η agrees with the *min-L*₁-norm direction asymptotically

$$\lim_{\eta \to 0} \lim_{t \to \infty} \ \hat{\theta}_{boost}^{t,\eta} / \| \, \hat{\theta}_{boost}^{t,\eta} \|_1 = \hat{\theta} \ .$$

Freund and Schapire (1995); Zhang and Yu (2005)

PRECISE HIGH-DIM ASYMPTOTIC THEORY FOR BOOSTING

DGP. $x_i \sim \mathcal{N}(0, \Lambda)$ i.i.d. with cov. $\Lambda \in \mathbb{R}^{p \times p}$, and y_i are generated with some $f : \mathbb{R} \to [0, 1]$,

$$\mathbb{P}(y_i = +1|x_i) = f(x_i^\top \theta_*) ,$$

with some $\theta_{\star} \in \mathbb{R}^p$.

Consider high-dim asymptotic regime with over-parametrized ratio

$$p/n \to \psi \in (0, \infty), \quad p, n \to \infty.$$

ntro. Min-norm Interpolant Regression Classification

PRECISE HIGH-DIM ASYMPTOTIC THEORY FOR BOOSTING

Statistical.

- how large is the empirical L_1 -margin?
- angle between the $\hat{\theta}$ (min- L_1 -norm interpolated classifier) and the truth θ_* ?
- generalization properties of Boosting?

Computational.

- iterations of the Boosting (precisely as a function of over-parametrization p/n) are required for an ϵ -approx. to the max- L_1 -margin?
- proportion of features activated by Boosting (with zero initialization) when the training error vanishes?

PRECISE HIGH-DIM ASYMPTOTIC THEORY FOR BOOSTING

Theorem (L. & Sur, '20).

Under mild conditions, for $\psi \geq \psi^{\star}(0),$ the following sharp asymptotic characterization

$$\lim_{n, y \to \infty} p^{1/2} \cdot \kappa_{\ell_1}(X, y) = \kappa_{\star}(\psi, \mu) , \quad a.s.$$

Generalization error

$$\lim_{n,p\to\infty} \, \mathbb{P}_{\mathbf{x},\mathbf{y}}\left(\mathbf{y}\cdot\mathbf{x}^{\top}\hat{\boldsymbol{\theta}}_{\ell_1}<0\right) = \underline{\mathrm{Err}}_{\star}(\boldsymbol{\psi},\boldsymbol{\mu}) \ , \ a.s.$$

Thrampoulidis et al. (2014, 2015, 2018); Gordon (1988)

Montanari et al. (2019); Deng et al. (2019); Shcherbina and Tirozzi (2003); Gardner (1988)

PRECISE HIGH-DIM ASYMPTOTIC THEORY FOR BOOSTING

 $\kappa_*(\psi, \mu)$ enjoys the analytic characterization: [L. & Sur, '20]

define
$$F_{\kappa}: \mathbb{R} \times \mathbb{R}^{\geq 0} \to \mathbb{R}^{\geq 0}$$

$$F_{\kappa}(c_1, c_2) := \left(\mathbb{E} \left[\left(\kappa - c_1 Y Z_1 - c_2 Z_2 \right)^2 \right] \right)^{\frac{1}{2}} \quad \text{where } \begin{cases} Z_2 \perp (Y, Z_1) \\ Z_i \sim \mathcal{N}(0, 1), \ i = 1, 2 \\ \mathbb{P}(Y = +1 | Z_1) = 1 - \mathbb{P}(Y = -1 | Z_1) = f(\rho \cdot Z_1) \end{cases}.$$

PRECISE HIGH-DIM ASYMPTOTIC THEORY FOR BOOSTING

$\kappa_*(\psi, \mu)$ enjoys the analytic characterization: [L. & Sur, '20]

Fixed point equations for $c_1, c_2, s \in \mathbb{R} \times \mathbb{R}_{>0} \times \mathbb{R}_{>0}$ given $\psi > 0$, where the expectation is over $(\Lambda, W, G) \sim \mu \otimes \mathcal{N}(0, 1) =: \mathcal{Q}$

$$\begin{split} c_1 &= - \underset{(\Lambda,W,G) \sim \mathcal{Q}}{\mathbb{E}} \left(\frac{\Lambda^{1/2}W \cdot \operatorname{prox}_s \left(\Lambda^{1/2}G + \psi^{-1/2} [\partial_1 F_\kappa \left(c_1, c_2 \right) - c_1 c_2^{-1} \partial_2 F_\kappa \left(c_1, c_2 \right)] \Lambda^{1/2}W \right)}{\psi^{-1/2} c_2^{-1} \partial_2 F_\kappa \left(c_1, c_2 \right)} \right) \\ c_1^2 + c_2^2 &= \underset{(\Lambda,W,G) \sim \mathcal{Q}}{\mathbb{E}} \left(\frac{\operatorname{prox}_s \left(\Lambda^{1/2}G + \psi^{-1/2} [\partial_1 F_\kappa \left(c_1, c_2 \right) - c_1 c_2^{-1} \partial_2 F_\kappa \left(c_1, c_2 \right)] \Lambda^{1/2}W \right)}{\psi^{-1/2} c_2^{-1} \partial_2 F_\kappa \left(c_1, c_2 \right)} \right)^2 \\ &= \underset{(\Lambda,W,G) \sim \mathcal{Q}}{\mathbb{E}} \left| \frac{\operatorname{prox}_s \left(\Lambda^{1/2}G + \psi^{-1/2} [\partial_1 F_\kappa \left(c_1, c_2 \right) - c_1 c_2^{-1} \partial_2 F_\kappa \left(c_1, c_2 \right)] \Lambda^{1/2}W \right)}{\psi^{-1/2} c_2^{-1} \partial_2 F_\kappa \left(c_1, c_2 \right)} \right| \\ & \text{with} \quad \operatorname{prox}_\lambda \left(t \right) = \arg \min \left\{ \lambda | s | + \frac{1}{2} \left(s - t \right)^2 \right\} = \operatorname{sgn}(t) \left(|t| - \lambda \right)_+ \end{split}$$

PRECISE HIGH-DIM ASYMPTOTIC THEORY FOR BOOSTING

 $\kappa_{\star}(\psi,\mu)$ enjoys the analytic characterization: [L. & Sur, '20]

Fixed point equations for $c_1, c_2, s \in \mathbb{R} \times \mathbb{R}_{>0} \times \mathbb{R}_{>0}$ given $\psi > 0$, where the expectation is over $(\Lambda, W, G) \sim$ $\mu \otimes \hat{\mathcal{N}}(0,1) =: \mathcal{Q}$

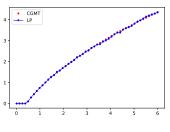
$$\begin{split} c_1 &= - \sum_{(\Lambda, W, G) \sim \mathcal{Q}} \left(\frac{\Lambda^{1/2} W \cdot \operatorname{prox}_s \left(\Lambda^{1/2} G + \psi^{-1/2} [\mathfrak{d}_1 F_\kappa (c_1, c_2) - c_1 c_2^{-1} \mathfrak{d}_2 F_\kappa (c_1, c_2)] \Lambda^{1/2} W \right)}{\psi^{-1/2} c_2^{-1} \mathfrak{d}_2 F_\kappa (c_1, c_2)} \right) \\ c_1^2 + c_2^2 &= \underset{(\Lambda, W, G) \sim \mathcal{Q}}{\mathbb{E}} \left(\frac{\operatorname{prox}_s \left(\Lambda^{1/2} G + \psi^{-1/2} [\mathfrak{d}_1 F_\kappa (c_1, c_2) - c_1 c_2^{-1} \mathfrak{d}_2 F_\kappa (c_1, c_2)] \Lambda^{1/2} W \right)}{\psi^{-1/2} c_2^{-1} \mathfrak{d}_2 F_\kappa (c_1, c_2)} \right)^2 \\ &= \underset{(\Lambda, W, G) \sim \mathcal{Q}}{\mathbb{E}} \left| \frac{\operatorname{prox}_s \left(\Lambda^{1/2} G + \psi^{-1/2} [\mathfrak{d}_1 F_\kappa (c_1, c_2) - c_1 c_2^{-1} \mathfrak{d}_2 F_\kappa (c_1, c_2)] \Lambda^{1/2} W \right)}{\psi^{-1/2} c_2^{-1} \mathfrak{d}_2 F_\kappa (c_1, c_2)} \right| \\ & \text{with } \operatorname{prox}_\lambda (t) = \operatorname{arg\,min} \left\{ \lambda |s| + \frac{1}{s} (s - t)^2 \right\} = \operatorname{sgn}(t) \left(|t| - \lambda \right)_+ \end{split}$$

with
$$\operatorname{prox}_{\lambda}(t) = \arg\min_{s} \left\{ \lambda |s| + \frac{1}{2}(s-t)^{2} \right\} = \operatorname{sgn}(t) (|t| - \lambda)_{+}$$

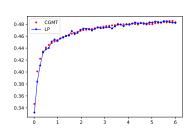
$$T(\psi, \kappa) := \psi^{-1/2} \left[F_{\kappa}(c_1, c_2) - c_1 \partial_1 F_{\kappa}(c_1, c_2) - c_2 \partial_2 F_{\kappa}(c_1, c_2) \right] - s$$
with $c_1(\psi, \kappa), c_2(\psi, \kappa), s(\psi, \kappa)$.

$$\kappa_{\star}(\psi, \mu) := \inf\{\kappa \geq 0 : T(\psi, \kappa) \geq 0\}$$

THEORY VS. EMPIRICAL



 $Max-L_1$ -Margin.



Generalization Error for $Min-L_1$ -Interpolated Classifier.

TECHNICAL REMARKS

Our results builds upon Convex Gaussian Minimax Theorem Thrampoulidis et al. (2014, 2015, 2018); Gordon (1988) and the work on the L_2 -margin by Montanari et al. (2019)

 L_1 case introduce some technical issues to overcome

- we prove a stronger uniform deviation result that suits the L₁ case, self-normalization property.
- different fixed point equation systems.
- (normalized) max L₁ margin much larger than max L₂ margin.

ALGORITHMIC: BOOSTING

Theorem (L. & Sur, '20).

With proper (non-vanishing) learning rate, the sequence $\{\hat{\theta}^t\}_{t=0}^{\infty}$ obtained by the Boosting algorithm satisfy:

for any $0 < \epsilon < 1$, when the number if iterations

$$t \ge T_{\epsilon}(p)$$
 with $\lim_{n,p\to\infty} \frac{T_{\epsilon}(p)}{p \log^2 n} = \frac{12\epsilon^{-2}}{\kappa_{\star}^2(\psi,\mu)}$,

the solution $\hat{\theta}^t/\|\hat{\theta}^t\|_1$ is an $(1-\epsilon)$ -approximation to the Min- L_1 -Interpolated Classifier

$$p^{1/2} \cdot \min_{i \in [n]} \frac{y_i x_i^{\mathsf{T}} \hat{\theta}^t}{\|\hat{\theta}^t\|_1} \in \left[(1 - \epsilon) \cdot \kappa_{\star}(\psi, \mu), \kappa_{\star}(\psi, \mu) \right] .$$

ALGORITHMIC: ACTIVATED FEATURES BY BOOSTING

Theorem (L. & Sur, '20).

Let $S_0(p)$ be the number of features selected when Boosting (for the first time at t) obtains zero training error with $\hat{\theta}^0 = 0$ initialization,

$$\frac{1}{n}\sum_{i=1}^n I_{y_ix_i^\top\hat{\theta}^t\leq 0}=0$$

with

$$S_0(p) \coloneqq \# \left\{ j \in [p] : \hat{\theta}_j^t \neq 0 \right\} \ .$$

We show

$$\limsup_{n,p\to\infty} \frac{S_0(p)}{p\log^2 p} \le \frac{12}{\kappa_{\star}^2(\psi,\mu)} \wedge 1$$

PROOF SKETCH

Step 1: \sqrt{p} -rescaling of L_1 ball

$$\xi_{\psi,\kappa}^{(n,p)} \coloneqq \min_{\|\theta\|_1 \le \sqrt{p}} \max_{\|\lambda\|_2 \le 1, \lambda \ge 0} \frac{1}{\sqrt{p}} \lambda^T (\kappa \mathbf{1} - (y \odot X)\theta)$$

It is not hard to see that

$$\begin{array}{l} \boldsymbol{\xi}_{\boldsymbol{\psi},\kappa}^{(n,p)} = \boldsymbol{0}, \ \text{ if and only if } \ \kappa \leq p^{1/2} \cdot \kappa_{\ell_1} \left(\left\{ \boldsymbol{x}_i, \boldsymbol{y}_i \right\}_{i=1}^n \right) \ , \\ \boldsymbol{\xi}_{\boldsymbol{\psi},\kappa}^{(n,p)} > \boldsymbol{0}, \ \text{ if and only if } \ \kappa > p^{1/2} \cdot \kappa_{\ell_1} \left(\left\{ \boldsymbol{x}_i, \boldsymbol{y}_i \right\}_{i=1}^n \right) \ . \end{array}$$

26 / 37

PROOF SKETCH

Step 1: \sqrt{p} -rescaling of L_1 ball

$$\begin{split} \xi_{\Psi,\,\kappa}^{(n,p)} \coloneqq \min_{\|\theta\|_1 \leq \sqrt{p}} \ \max_{\|\lambda\|_2 \leq 1, \lambda \geq 0} \ \frac{1}{\sqrt{p}} \lambda^T \big(\kappa \mathbf{1} - (y \odot X) \theta \big) \\ \xi_{\Psi,\,\kappa}^{(n,p)} \coloneqq \min_{\|\theta\|_1 \leq \sqrt{p}} \ \max_{\|\lambda\|_2 \leq 1, \lambda \geq 0} \ \frac{1}{\sqrt{p}} \lambda^T \Big(\kappa \mathbf{1} - (y \odot \mathbf{z}) \langle w, \Lambda^{1/2} \theta \rangle \Big) - \frac{1}{\sqrt{p}} \boxed{\lambda^T Z \Pi_{w^{\perp}}(\Lambda^{1/2} \theta)} \end{split}$$

Step 2: reduction via Gordon's comparison (convex Gaussian min-max theorem)

Thrampoulidis et al. (2014, 2015, 2018); Gordon (1988)

$$\begin{split} & \hat{\boldsymbol{\xi}}_{,\psi,\kappa}^{(n,p)} \\ & \coloneqq \min_{\|\boldsymbol{\theta}\|_1 \leq \sqrt{p}} \max_{\|\boldsymbol{\lambda}\|_2 \leq 1, \boldsymbol{\lambda} \geq 0} \frac{1}{\sqrt{p}} \boldsymbol{\lambda}^T \left(\kappa \mathbf{1} - (\boldsymbol{y} \odot \boldsymbol{z}) \langle \boldsymbol{w}, \boldsymbol{\Lambda}^{1/2} \boldsymbol{\theta} \rangle - \tilde{\boldsymbol{z}} \|\boldsymbol{\Pi}_{\boldsymbol{w}^{\perp}} (\boldsymbol{\Lambda}^{1/2} \boldsymbol{\theta}) \|_2 \right) + \frac{1}{\sqrt{p}} \|\boldsymbol{\lambda}\|_2 \langle \boldsymbol{g}, \boldsymbol{\Pi}_{\boldsymbol{w}^{\perp}} (\boldsymbol{\Lambda}^{1/2} \boldsymbol{\theta}) \rangle \\ & = \min_{\|\boldsymbol{\theta}\|_1 \leq \sqrt{p}} \left[\boldsymbol{\psi}^{-1/2} \widehat{\boldsymbol{F}}_{\kappa} \left(\langle \boldsymbol{w}, \boldsymbol{\Lambda}^{1/2} \boldsymbol{\theta} \rangle, \|\boldsymbol{\Pi}_{\boldsymbol{w}^{\perp}} (\boldsymbol{\Lambda}^{1/2} \boldsymbol{\theta}) \|_2 \right) + \frac{1}{\sqrt{p}} \left(\boldsymbol{\Pi}_{\boldsymbol{w}^{\perp}} (\boldsymbol{g}), \boldsymbol{\Lambda}^{1/2} \boldsymbol{\theta} \right) \right] \end{split}$$

TECHNICAL CHALLENGES IN L_1 CASE Step 3: large n, p limit

The empirical problem (finite-dim optimization)

$$\hat{\xi}_{\Psi,\,\kappa}^{(n,p)} = \min_{\parallel\theta\parallel_1 \leq \sqrt{p}} \left[\psi^{-1/2} \widehat{F}_{\kappa} \left(\langle w, \Lambda^{1/2}\theta \rangle, \|\Pi_{w^{\perp}}(\Lambda^{1/2}\theta)\|_2 \right) + \frac{1}{\sqrt{p}} \left\{ \Pi_{w^{\perp}}(g), \Lambda^{1/2}\theta \right\} \right]$$

Let's naively take the limit (infinite-dim optimization)

$$\tilde{\xi}_{\psi,\kappa}^{(\infty,\infty)} \coloneqq \min_{\|h\|_{L_1(\mathcal{Q})} \leq 1} \left[\psi^{-1/2} F_{\kappa} \left(\langle w, \Lambda^{1/2} h \rangle_{L_2(\mathcal{Q})}, \|\Pi_{w^{\perp}}(\Lambda^{1/2} h)\|_{L_2(\mathcal{Q})} \right) + \left\langle \Pi_{w^{\perp}}(G), \Lambda^{1/2} h \rangle_{L_2(\mathcal{Q})} \right]$$

One needs to show

$$\lim_{p\to\infty,p/n(p)\to\psi}\ \hat{\xi}_{\psi,\kappa}^{(n,p)}\stackrel{\mathrm{a.s.}}{=}\ \tilde{\xi}_{\psi,\kappa}^{(\infty,\infty)}$$

TECHNICAL CHALLENGES IN L_1 CASE Step 3: large n, p limit

The empirical problem (finite-dim optimization)

$$\hat{\xi}_{\Psi,\,\kappa}^{(n,p)} = \min_{\parallel\theta\parallel_1 \leq \sqrt{p}} \left[\psi^{-1/2} \widehat{F}_{\kappa} \left(\left\langle w, \Lambda^{1/2} \theta \right\rangle, \|\Pi_{w^{\perp}}(\Lambda^{1/2} \theta)\|_2 \right) + \frac{1}{\sqrt{p}} \left\langle \Pi_{w^{\perp}}(g), \Lambda^{1/2} \theta \right\rangle \right]$$

Let's naively take the limit (infinite-dim optimization)

$$\tilde{\xi}_{\psi,\kappa}^{(\infty,\infty)} \coloneqq \min_{\|h\|_{L_{1}(\mathcal{Q})} \leq 1} \left[\psi^{-1/2} F_{\kappa} \left(\left\langle w, \Lambda^{1/2} h \right\rangle_{L_{2}(\mathcal{Q})}, \|\Pi_{w^{\perp}}(\Lambda^{1/2} h)\|_{L_{2}(\mathcal{Q})} \right) + \left\langle \Pi_{w^{\perp}}(G), \Lambda^{1/2} h \right\rangle_{L_{2}(\mathcal{Q})} \right]$$

One needs to show

$$\lim_{p \to \infty, p/n(p) \to \psi} \hat{\xi}_{\psi, \kappa}^{(n,p)} \stackrel{\text{a.s.}}{=} \tilde{\xi}_{\psi, \kappa}^{(\infty, \infty)}$$

 L_1 vs. L_2 geometry: for the constraint set $\|\theta\|_1 \le \sqrt{p}$, define

$$\begin{split} c_1 = \langle w, \Lambda^{1/2}\theta \rangle, c_2 = \|\Pi_{w^{\perp}}(\Lambda^{1/2}\theta)\|_2 \\ c_2 \text{ could be } \sqrt{p} \to \infty. \end{split}$$

TECHNICAL CHALLENGES IN L_1 CASE

Step 3: large n, p limit

The empirical problem (finite-dim optimization)

$$\hat{\xi}_{\psi,\,\kappa}^{(n,p)} = \min_{\|\theta\|_1 \leq \sqrt{p}} \left[\psi^{-1/2} \widehat{F}_{\kappa} \left((w, \Lambda^{1/2} \theta), \|\Pi_{w^{\perp}} (\Lambda^{1/2} \theta)\|_2 \right) + \frac{1}{\sqrt{p}} \left\{ \Pi_{w^{\perp}} (g), \Lambda^{1/2} \theta \right\} \right]$$

Let's naively take the limit (infinite-dim optimization)

$$\tilde{\xi}_{\psi,\kappa}^{(\infty,\infty)} \coloneqq \min_{\|h\|_{L_{1}(\mathcal{Q})} \leq 1} \left[\psi^{-1/2} F_{\kappa} \left(\left\langle w, \Lambda^{1/2} h \right\rangle_{L_{2}(\mathcal{Q})}, \|\Pi_{w^{\perp}}(\Lambda^{1/2} h)\|_{L_{2}(\mathcal{Q})} \right) + \left\langle \Pi_{w^{\perp}}(G), \Lambda^{1/2} h \right\rangle_{L_{2}(\mathcal{Q})} \right]$$

One needs to show

$$\lim_{p \to \infty, p/n(p) \to \psi} \hat{\xi}_{\psi, \kappa}^{(n,p)} \stackrel{\text{a.s.}}{=} \tilde{\xi}_{\psi, \kappa}^{(\infty,\infty)}$$

 L_1 vs. L_2 geometry: for the constraint set $\|\theta\|_1 \le \sqrt{p}$, define

$$c_1 = \langle w, \Lambda^{1/2} \theta \rangle, c_2 = \|\Pi_{w^{\perp}}(\Lambda^{1/2} \theta)\|_2$$

 $c_2 \text{ could be } \sqrt{p} \to \infty.$

[L. & Sur '20] shows uniform deviation over unbounded domain for the fixed-point equation (KKT), using a key self-normalization property of $\partial_i F_{\kappa}(c_1, c_2)$.

For i = 1, 2, we have w.p. at least $1 - n^{-2}$,

$$\sup_{|c_1| \leq M, \ c_2 > 0} \left| \partial_i \hat{F}_\kappa \left(c_1, c_2 \right) - \partial_i F_\kappa \left(c_1, c_2 \right) \right| \leq \frac{C \log n}{\sqrt{n}}$$

[BACKUP] CONVEX GAUSSIAN MINIMAX THEOREM

Let $C_1 \subset \mathbb{R}^n$, $C_2 \subset \mathbb{R}^p$ be two compact sets and let $R: C_1 \times C_2 \to \mathbb{R}$ be a continuous function. Let $X = (X_{i,j}) \in \mathbb{R}^{n \times p}$, $g \sim \mathcal{N}(0, I_n)$ and $h \sim \mathcal{N}(0, I_p)$ be independent vectors and matrices with standard Gaussian entries. Define

$$\begin{split} Q_1(X) &= \min_{w_1 \in C_1} \max_{w_2 \in C_2} w_1^\top X w_2 + R(w_1, w_2) \\ Q_2(g, h) &= \min_{w_1 \in C_1} \max_{w_2 \in C_2} \|w_2\| g^\top w_1 + \|w_1\| h^\top w_2 + R(w_1, w_2). \end{split}$$

Then

1. For all $t \in \mathbb{R}$,

$$\mathbb{P}(Q_1(X) \le t) \le 2\mathbb{P}(Q_2(g,h) \le t).$$

2. Suppose C_1 and C_2 are both convex, and R is convex concave in (w_1, w_2) . Then, for all $t \in \mathbb{R}$,

$$\mathbb{P}(Q_1(X) \ge t) \le 2\mathbb{P}(Q_2(g,h) \ge t).$$

SUMMARY

Research agenda: statistical or generalization theory for min-norm interpolants

(naive usage of Rademacher complexity, or VC-dim won't explain well)

- Regression: [L. & Rakhlin '18], [L. & Dou '19], [L., Rakhlin & Zhai '19]
- Classification: [L. & Sur '20]

Thank you!

 Liang, T. & Sur, P. (2020). — A Precise High-Dimensional Asymptotic Theory for Boosting and Min-L1-Norm Interpolated Classifiers.

arXiv:2002.01586

 Liang, T., Rakhlin, A. & Zhai, X. (2019). — On the Multiple Descent of Minimum-Norm Interpolants and Restricted Lower Isometry of Kernels.

arXiv:1908.10292

- Liang, T. & Rakhlin, A. (2018). Just Interpolate: Kernel "Ridgeless" Regression Can Generalize.
 The Annals of Statistics, to appear
- Dou, X. & Liang, T. (2019). Training Neural Networks as Learning Data-adaptive Kernels: Provable Representation and Approximation Benefits.

Journal of the American Statistical Association, to appear

PROOF IDEA: RESTRICTED LOWER ISOMETRY

Proof Idea: on a filtration of spaces, establish restricted lower isometry.

Koltchinskii and Mendelson (2015); Mendelson (2014)

PROOF IDEA: RESTRICTED LOWER ISOMETRY

Proof Idea: on a filtration of spaces indexed by polynomial basis, establish restricted lower isometry of the empirical kernel.

Define $n\mathbf{K} := [k(x_i, x_i)]_{i,i \in [n]} \in \mathbb{R}^{n \times n}$

$$n\mathbf{K}_{ij} = \frac{h}{n} \left(\frac{x_i^{\top} x_j}{d} \right) = \sum_{t=0}^{\infty} \alpha_t \left(\frac{x_i^{\top} x_j}{d} \right)^t$$

$$= \sum_{r_1, \dots, r_d \ge 0} c_{r_1 \dots r_d} \alpha_{r_1 + \dots + r_d} p_{r_1 \dots r_d} (x_i) p_{r_1 \dots r_d} (x_j) / d^{r_1 + \dots + r_d}$$

Define filtrated empirical kernel

$$n\mathbf{K}_{ij}^{[\leq \mathfrak{t}]} \coloneqq \sum_{\substack{r_1, \cdots, r_d \geq 0 \\ r_1 + \cdots + r_d \leq \mathfrak{t}}} c_{r_1 \cdots r_d} \, \alpha_{r_1 + \cdots + r_d} p_{r_1 \cdots r_d}(x_i) p_{r_1 \cdots r_d}(x_j) / d^{r_1 + \cdots + r_d}$$

$$c_{r_1 \cdots r_d} = \frac{(r_1 + \cdots + r_d)!}{r_1! \cdots r_d!}, p_{r_1 \cdots r_d}(x_i) = (x_i[1])^{r_1} \cdots (x_i[d])^{r_d} \text{ monomials with multi-index } r_1 \cdots r_d$$

filtrated empirical kernel

$$\begin{split} n\mathbf{K}_{ij}^{\left[\leq t\right]} &:= \sum_{\substack{r_1, \dots, r_d \geq 0 \\ r_1 + \dots + r_d \leq t}} c_{r_1 \cdots r_d} \alpha_{r_1 + \dots + r_d} p_{r_1 \cdots r_d}(x_i) p_{r_1 \cdots r_d}(x_j) / d^{r_1 + \dots + r_d} \\ n\mathbf{K}^{\left[\leq t\right]} &= \bigoplus_{\substack{n \times \binom{t+d}{t}} \binom{t+d}{t} \times n}} \Phi^\top \end{split}$$

filtrated polynomial features

$$\Phi_{i,(r_1\cdots r_d)} = \left(c_{r_1\cdots r_d}\alpha_{r_1+\cdots + r_d}\right)^{1/2} p_{r_1\cdots r_d}(x_i)/d^{(r_1+\cdots + r_d)/2}$$

filtrated sample covariance operator

$$\Theta^{[\leq \iota]} := \frac{1}{n} \underbrace{\Phi^{\top}}_{\binom{\iota+d}{\iota} \times n} \cdot \underbrace{\Phi}_{n \times \binom{\iota+d}{\iota}}$$

filtrated empirical kernel

$$\begin{split} n\mathbf{K}_{ij}^{\left[\leq t\right]} &:= \sum_{\substack{r_1, \cdots, r_d \geq 0 \\ r_1 + \cdots + r_d \leq t}} c_{r_1 \cdots r_d} \alpha_{r_1 + \cdots + r_d} p_{r_1 \cdots r_d}(x_i) p_{r_1 \cdots r_d}(x_j) / d^{r_1 + \cdots + r_d} \\ n\mathbf{K}^{\left[\leq t\right]} &= \bigoplus_{\substack{n \times \binom{t+d}{t} \\ \text{otherwise}}} \bigoplus_{\substack{t \in d \\ t \neq d}} \bigoplus_{\substack{t \in d \\ t \neq d}} \sum_{\substack{t \in d \\ t \neq d}} p_{r_1 \cdots r_d}(x_i) p_{r_1 \cdots r_d}(x_i) / d^{r_1 + \cdots + r_d} \\ \end{split}$$

filtrated polynomial features

$$\Phi_{i,(r_1\cdots r_d)} = \left(c_{r_1\cdots r_d}\alpha_{r_1+\cdots +r_d}\right)^{1/2} p_{r_1\cdots r_d}(x_i)/d^{(r_1+\cdots +r_d)/2}$$

filtrated sample covariance operator

$$\Theta^{[\leq \iota]} := \frac{1}{n} \underbrace{\Phi^{\top}}_{\binom{\iota+d}{\iota} \times n} \underbrace{\Phi}_{n \times \binom{\iota+d}{\iota}}$$

Restricted Lower Isometry of Kernel: all non-zero eigenvalues of $\mathbf{K}^{[\leq t]}$ is lower bounded by d^{-t} , i.e.,

$$\lambda_{\min}\left(\Theta^{[\leq \iota]}\right) \gtrsim d^{-\iota}$$

Lemma (L., Rakhlin & Zhai, '19).

Assume that Taylor coefficients of *h* satisfy $\alpha_i > 0 \ \forall i$.

Consider any positive integer ι that satisfy $d^{\iota} \log d = o(n)$. and $\iota < \nu$. ν is the tail decay of \mathcal{P} .

Then with probability at least $1 - \exp(-C \cdot n/d^{\iota})$,

all non-zero eigenvalues of $\mathbf{K}^{[\leq \iota]} \geq C \cdot d^{-\iota}$.

Lemma (L., Rakhlin & Zhai, '19).

Assume that Taylor coefficients of h satisfy $\alpha_i > 0 \ \forall i$.

Consider any positive integer ι that satisfy $d^{\iota} \log d = o(n)$. and $\iota < \nu$. ν is the tail decay of \mathcal{P} .

Then with probability at least $1 - \exp(-C \cdot n/d^{\iota})$,

all non-zero eigenvalues of $\mathbf{K}^{[\leq \iota]} \geq C \cdot d^{-\iota}$.

Some wrong but useful intuition:

• eigenvalues of $\mathbf{K}^{[\leq \iota]}$ equals that of $\Theta^{[\leq \iota]}$

Lemma (L., Rakhlin & Zhai, '19).

Assume that Taylor coefficients of h satisfy $\alpha_i > 0 \ \forall i$.

Consider any positive integer ι that satisfy $d^{\iota} \log d = o(n)$. and $\iota < \nu$. ν is the tail decay of \mathcal{P} .

Then with probability at least $1 - \exp(-C \cdot n/d^{\iota})$,

all non-zero eigenvalues of $\mathbf{K}^{[\leq \iota]} \geq C \cdot d^{-\iota}$.

Some wrong but useful intuition:

- eigenvalues of $\mathbf{K}^{[\leq \iota]}$ equals that of $\Theta^{[\leq \iota]}$
- suppose monomials $\prod_{i=1}^{d} (x[i])^{r_i}$ are orthogonal (wrong), then

$$\mathbb{E}\left[\Theta^{[\leq \iota]}\right] = \operatorname{diag}(C(0), \dots, C(\iota') \cdot d^{-\iota'}, \dots, \underbrace{C(\iota) \cdot d^{-\iota}}_{d-1}) \text{ such entries}$$

Lemma (L., Rakhlin & Zhai, '19).

Assume that Taylor coefficients of h satisfy $\alpha_i > 0 \ \forall i$.

Consider any positive integer ι that satisfy $d^{\iota} \log d = o(n)$. and $\iota < \nu$. ν is the tail decay of \mathcal{P} .

Then with probability at least $1 - \exp(-C \cdot n/d^{\iota})$,

all non-zero eigenvalues of $\mathbf{K}^{[\leq \iota]} \geq C \cdot d^{-\iota}$.

Some wrong but useful intuition:

- eigenvalues of $\mathbf{K}^{[\leq \iota]}$ equals that of $\Theta^{[\leq \iota]}$
- suppose monomials $\prod_{i=1}^{d} (x[i])^{r_i}$ are orthogonal (wrong), then

$$\mathbb{E}\left[\Theta^{[\leq \iota]}\right] = \operatorname{diag}(C(0), \dots, C(\iota') \cdot d^{-\iota'}, \dots, \underbrace{C(\iota) \cdot d^{-\iota}}_{d-1}) \text{ such entries}$$

even so, standard concentration (fails, at least apply naively)

$$\sup_{u \in B_2^{\binom{d+1}{\iota}}} u^{\top} \left(\Theta^{\left[\leq \iota \right]} - \mathbb{E} \left[\Theta^{\left[\leq \iota \right]} \right] \right) u \leq \frac{1}{\sqrt{n}} \operatorname{Var} \cdots$$

Then, how to make it right? Two Ideas.

Then, how to make it right? Two Ideas.

Idea 1: Gram-Schimdt process on polynomials, weakly-dependent

$$\{1,t,t^2,\cdots\} \to \{1,q_1(t),q_2(t),\cdots\}$$
 q orthogonal polynomial basis on $L^2_{\mathcal{P}}$

Then, how to make it right? **Two Ideas**.

Idea 1: Gram-Schimdt process on polynomials, weakly-dependent

$$\{1,t,t^2,\cdots\} \to \{1,q_1(t),q_2(t),\cdots\} \quad \text{q orthogonal polynomial basis on $L^2_{\mathcal{P}}$}$$

$$\Phi_{i,(r_1\cdots r_d)} \to \Psi_{i,(r_1\cdots r_d)} = \left(c_{r_1\cdots r_d} \alpha_{r_1+\cdots + r_d}\right)^{1/2} \prod_{j\in[d]} q_{r_j}(x_i[j])/d^{(r_1+\cdots + r_d)/2}$$

$$\Phi = \Psi \Lambda$$
, $\Lambda \in \mathbb{R}^{\binom{\iota+d}{\iota} \times \binom{\iota+d}{\iota}}$ upper-triangular

Then, how to make it right? **Two Ideas**.

Idea 1: Gram-Schimdt process on polynomials, weakly-dependent

$$\{1, t, t^2, \dots\} \rightarrow \{1, q_1(t), q_2(t), \dots\}$$
 q orthogonal polynomial basis on $L^2_{\mathcal{P}}$

$$\begin{split} \Phi_{i,(r_1\cdots r_d)} \to \Psi_{i,(r_1\cdots r_d)} &= \left(c_{r_1\cdots r_d}\,\alpha_{r_1+\cdots + r_d}\right)^{1/2} \prod_{j\in[d]} q_{r_j}(x_i[j])/d^{(r_1+\cdots + r_d)/2} \\ \Phi &= \Psi \Lambda, \quad \Lambda \in \mathbb{R}^{\binom{\iota+d}{\iota}} \times \binom{\iota+d}{\iota} \quad \text{upper-triangular} \end{split}$$

Claim: weakly-dependent
$$\Rightarrow$$
 $\|\Lambda\|_{op}$, $\|\Lambda^{-1}\|_{op} \le C(\iota)$

$$u^{\mathsf{T}}\Theta^{\left[\leq\iota\right]}u = \frac{1}{n}\|\Phi u\|^2 = \frac{1}{n}\|\Psi\Lambda u\|^2 \geq \lambda_{\mathsf{min}}\left(\frac{1}{n}\Psi^{\mathsf{T}}\Psi\right)\|\Lambda u\|^2 \asymp \lambda_{\mathsf{min}}\left(\frac{1}{n}\Psi^{\mathsf{T}}\Psi\right)\|u\|^2$$

Then, how to make it right? Two Ideas.

Idea 2: small-ball approach rather than standard concentration

Then, how to make it right? Two Ideas.

Idea 2: small-ball approach rather than standard concentration

lower bound $\lambda_{\min}\left(\frac{1}{n}\Psi^{\mathsf{T}}\Psi\right)$ equiv. $\forall u, \|u\| = 1$, lower bound $\|\Psi u\|^2$ utilize non-negativity

$$\|\Psi u\|^2 = \frac{1}{n} \sum_{i=1}^n \langle \Psi(x_i), u \rangle^2 \ge c_1 \mathbb{E}[\langle \Psi(x_i), u \rangle^2] \cdot \frac{1}{n} \sum_{i=1}^n I_{\langle \Psi(x_i), u \rangle^2 \ge c_1 \mathbb{E}[\langle \Psi(X), u \rangle^2]}$$

small-ball property, \exists constants c_1 , c_2

$$\mathbb{P}\left(\langle \Psi(x_i), u \rangle^2 \ge c_1 \mathbb{E}[\langle \Psi(X), u \rangle^2]\right) \ge c_2$$

which will imply w.p. at least $1 - \exp(-c \cdot n)$

$$\frac{1}{n} \sum_{i=1}^{n} I_{\langle \Psi(x_i), u \rangle^2 \ge c_1 \mathbb{E}[\langle \Psi(X), u \rangle^2]} \ge c_2/2$$

Non-trivial: verify small-ball property for polynomials (weakly dependent) via Paley-Zygmund

Then, how to make it right? Two Ideas.

 $\left(\mathsf{Lemma} \left(\mathsf{L., Rakhlin \& Zhai, '19} \right) \right)$

all non-zero eigenvalues of $\mathbf{K}^{[\leq \iota]} \geq C \cdot d^{-\iota}$.

Mendelson (2014); Liang et al. (2019); Ghorbani et al. (2019)

INTUITION: WEAKLY DEPENDENT

For any three distinct polynomial features indexed by $(r_1 \cdots r_d)$, $(r_1' \cdots r_d')$, $(r_1'' \cdots r_d'')$

$$\prod_{j \in [d]} q_{r_j}(x[j]), \prod_{j \in [d]} q_{r'_j}(x[j]), \prod_{j \in [d]} q_{r''_j}(x[j])$$

Third moment

$$\mathbb{E}\left[q_{r_1\cdots r_d}q_{r_1^{\prime}\cdots r_d^{\prime}}q_{r_1^{\prime\prime}\cdots r_d^{\prime\prime}}\right]\neq 0$$

only if
$$\forall j \in [d], r_j + r'_j \geq r''_j$$
.

Among such triplets, at most $\frac{3^{21}}{d^4} = O(1/d^4)$ fraction has non-zero third moment.

BACK TO MULTIPLE DESCENT PROOF: SKETCH

Decompose Risk to Bias and Variance. Surprisingly, both terms can be bounded by $\mathbb{E}_{x \sim \mathcal{P}^d} \|k(\mathbf{X}, \mathbf{X})^{-1} k(\mathbf{X}, x)\|^2$.

Decompose Risk to Bias and Variance. Surprisingly, both terms can be bounded by $\mathbb{E}_{x \sim \mathcal{P}^d} \|k(\mathbf{X}, \mathbf{X})^{-1} k(\mathbf{X}, x)\|^2$.

Sketch:

$$\begin{split} & \mathbb{E}_{x} \| k(\mathbf{X}, \mathbf{X})^{-1} k(\mathbf{X}, x) \|^{2} \\ & \lesssim \sum_{i=0}^{L} \mathbb{E}_{x} \| \mathbf{K}^{-1} \frac{1}{n} (\mathbf{X} x)^{i} / d^{i} \|^{2} + \mathbb{E}_{x} \| \mathbf{K}^{-1} \frac{1}{n} \sum_{i=\iota+1}^{\infty} (\mathbf{X} x)^{i} / d^{i} \|^{2} \\ & \lesssim \frac{1}{n^{2}} \sum_{i=0}^{L} \mathbb{E}_{x} \| \mathbf{K}^{-1} (\mathbf{X} x)^{i} / d^{i} \|^{2} + \| (n\mathbf{K})^{-1} \|_{\text{op}}^{2} \cdot \mathbb{E}_{x} \| \sum_{i=\iota+1}^{\infty} (\mathbf{X} x)^{i} / d^{i} \|^{2} \\ & \lesssim \frac{1}{n^{2}} \sum_{i=0}^{L} \mathbb{E}_{x} \left[\| (\mathbf{K}^{[\leq i]})^{+} \|_{\text{op}}^{2} \cdot \| (\mathbf{X} x)^{i} / d^{i} \|^{2} \right] + \frac{n}{d^{\iota+1}} \\ & \lesssim \frac{1}{n^{2}} \sum_{i=0}^{L} \mathbb{E}_{x} \left[d^{2i} \cdot \| (\mathbf{X} x)^{i} / d^{i} \|^{2} \right] + \frac{n}{d^{\iota+1}} \quad \text{use restricted lower isometry} \\ & \lesssim \frac{d^{\iota}}{n} + \frac{n}{d^{\iota+1}} \quad . \end{split}$$