# Midterm report EE5111 Course Project

# Objectives

- Obtaining principled uncertainty estimates from deep networks.
- Using Dropout to obtain uncertainty estimates
- How can uncertainty estimates be used for adversarial robustness?

# Why should I care about uncertainty?

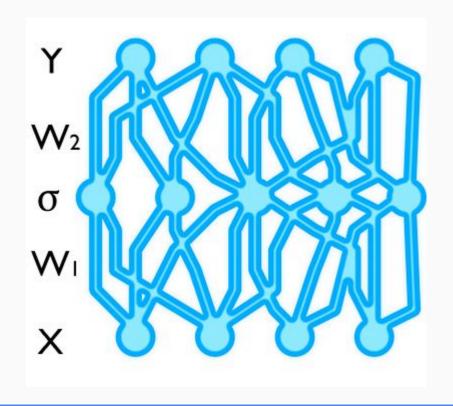
- Uncertainty information is critical in areas such as the life sciences and autonomous driving, where decisions taken are high-stakes
- Model uncertainty is also important in understanding the shortcomings of the model.
- Current deep learning techniques rely on point estimates for the weights.
   Softmax scores are not enough to obtain true uncertainty!

# How do you obtain uncertainty information?

- Bayesian methods are traditionally used for this!
- However, these models are typically harder to optimise.
- Gaussian processes model distributions over possible functions that fit the data
- Neural Networks can be related to a Gaussian process in the limit of infinite width of the hidden layers.
- Dropout allows us to obtain uncertainty estimates from standard NNs without changing a thing!

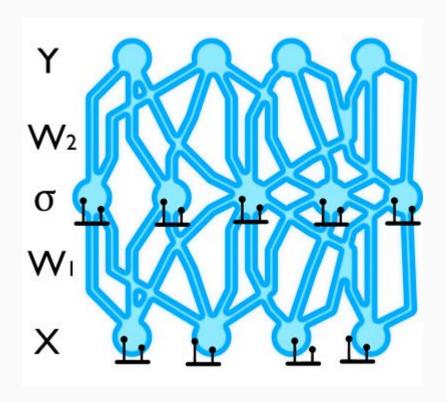
- Consider a neural network with a single hidden layer
- Let W<sub>1</sub> and W<sub>2</sub> be the weights of the network as shown.
- Assume that Input dimensions = Q;
   Output = D and hidden layer = K
- The output for an input x would be

$$\hat{\mathbf{y}} = \sigma(\mathbf{x}\mathbf{W}_1 + \mathbf{b})\mathbf{W}_2$$



- Dropout[Hinton et al.] is a technique to avoid over-fitting the data.
- Sample two binary vectors b<sub>1</sub>, b<sub>2</sub> of dimensions Q and K respectively.
- The elements of b<sub>i</sub> take a value 1 with probability p<sub>i</sub> (i=1,2)
- Essentially, we set a fraction of inputs to each layer to zero.
- The dropout model's output is

$$\widehat{\mathbf{y}} = \sigma(\mathbf{x}\mathbf{b}_1\mathbf{W}_1 + \mathbf{b})\mathbf{b}_2\mathbf{W}_2.$$



• Suppose we have  $\{y_1, \dots, y_N\}$  as N observed outputs for inputs

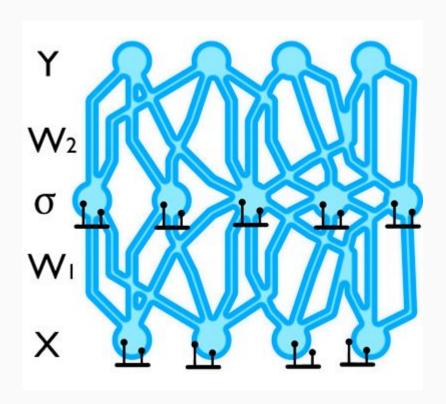
$$\{\mathbf{x}_1,\ldots,\mathbf{x}_N\}$$

 For regression, we can use the Euclidean loss

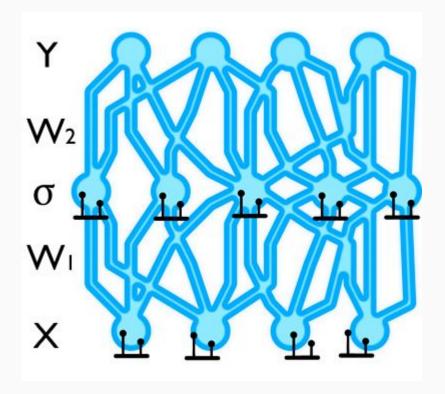
$$E = rac{1}{2N} \sum_{n=1}^{N} ||\mathbf{y}_n - \widehat{\mathbf{y}}_n||_2^2$$

Typically we also add L<sub>2</sub> regularisation :

$$\mathcal{L}_{dropout} := E + \lambda_1 ||\mathbf{W}_1||_2^2 + \lambda_2 ||\mathbf{W}_2||_2^2 + \lambda_3 ||\mathbf{b}||_2^2.$$



- During training, we sample random binary vectors b<sub>i</sub> for each input and for each forward pass
- Weights are typically scaled by 1/p<sub>i</sub> maintain constant output scale
- At test time, we do a single forward pass.

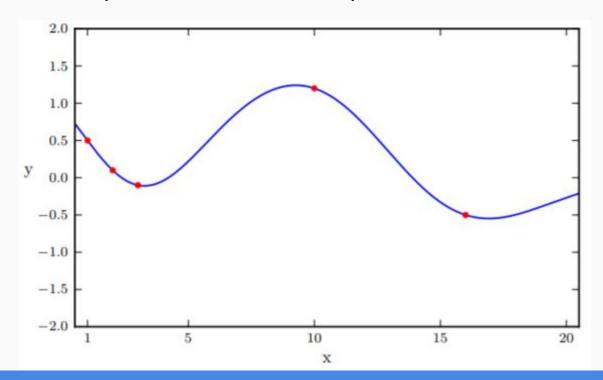


# Gaussian Process

#### **Textbook Definition:**

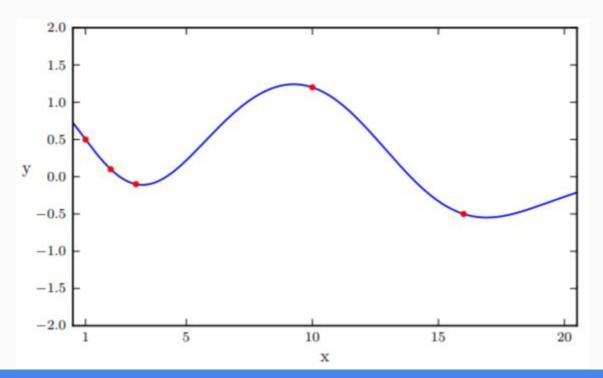
A Gaussian process is a collection of random variables, any finite number of which have consistent Gaussian distributions.

The problem: What are the possible functions that could have generated our data?



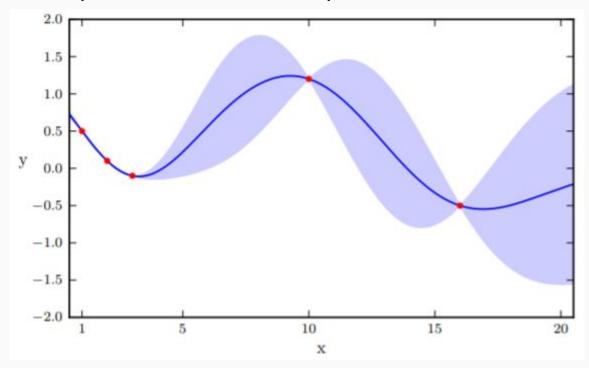
Consider a typical non-linear regression problem. One fit for a given set of points is given here

The problem: What are the possible functions that could have generated our data?



What other curves can explain our data?
With what certainty can we predict values for test samples?

The problem: What are the possible functions that could have generated our data?



Gaussian Processes allow us to model distributions over functions that fit our data.

- Let the inputs be denoted by  $\mathbf{X} \in \mathbb{R}^{N \times Q}$  and outputs  $\mathbf{Y} \in \mathbb{R}^{N \times D}$
- We first place a prior  $p(\mathbf{f})$  over the space of functions.
- The posterior distribution of the space of functions given the data is

$$p(\mathbf{f}|\mathbf{X}, \mathbf{Y}) \propto p(\mathbf{Y}|\mathbf{X}, \mathbf{f})p(\mathbf{f})$$

In a Gaussian process, we model the posteriors as :

$$\mathbf{F} \mid \mathbf{X} \sim \mathcal{N}(\mathbf{0}, \mathbf{K}(\mathbf{X}, \mathbf{X}))$$
  
 $\mathbf{Y} \mid \mathbf{F} \sim \mathcal{N}(\mathbf{F}, \tau^{-1}\mathbf{I}_N)$ 

Where K(.,.) - Covariance function

- Precision parameter

### Variational Inference

- Inference in a Gaussian process requires computing of inverse of an NxN matrix
   Intractable!
- An approximation would be to condition the model on some random variables w
- For a new datapoint  $\mathbf{x}^*$ :  $p(\mathbf{y}^*|\mathbf{x}^*, \mathbf{X}, \mathbf{Y}) = \int p(\mathbf{y}^*|\mathbf{x}^*, \boldsymbol{\omega}) p(\boldsymbol{\omega}|\mathbf{X}, \mathbf{Y}) \, d\boldsymbol{\omega},$
- Computing the posterior of w is generally intractable.
- Use an approximate variational distribution q(w)
- We minimize the KL divergence :

$$KL(q(\omega) \mid p(\omega|X, Y)),$$

• Approximate predictive distribution :  $q(\mathbf{y}^*|\mathbf{x}^*) = \int p(\mathbf{y}^*|\mathbf{x}^*, \boldsymbol{\omega}) q(\boldsymbol{\omega}) d\boldsymbol{\omega}$ .

### Variational Inference

Minimising the KL divergence <---> Maximising the log evidence lower bound:

$$\mathcal{L}_{ ext{VI}} := \int q(\boldsymbol{\omega}) \log p(\mathbf{Y}|\mathbf{X}, \boldsymbol{\omega}) d\boldsymbol{\omega} - \text{KL}(q(\boldsymbol{\omega})||p(\boldsymbol{\omega}))$$

- Explain the data well + Capture the prior!
- What follows: A variational approximation to the Gaussian Process posterior
- This turns out to be equivalent to use of dropout in standard DNNs!

# Dropout as a Bayesian Approximation

#### 3.1 A Gaussian Process Approximation

We begin by defining our covariance function. Let  $\sigma(\cdot)$  be some non-linear function such as the rectified linear (ReLU) or the hyperbolic tangent function (TanH). We define K(x, y) to be

$$\mathbf{K}(\mathbf{x}, \mathbf{y}) = \int p(\mathbf{w})p(b)\sigma(\mathbf{w}^T\mathbf{x} + b)\sigma(\mathbf{w}^T\mathbf{y} + b)d\mathbf{w}db$$

with  $p(\mathbf{w})$  a standard multivariate normal distribution of dimensionality Q and some distribution p(b). It is trivial to show that this defines a valid covariance function following [Tsuda et al., 2002].

We use Monte Carlo integration with K terms to approximate the integral above. This results in the finite rank covariance function

$$\widehat{\mathbf{K}}(\mathbf{x}, \mathbf{y}) = \frac{1}{K} \sum_{k=1}^{K} \sigma(\mathbf{w}_k^T \mathbf{x} + b_k) \sigma(\mathbf{w}_k^T \mathbf{y} + b_k)$$

with  $\mathbf{w}_k \sim p(\mathbf{w})$  and  $b_k \sim p(b)$ . K will be the number of hidden units in our single hidden layer NN approximation.

Using  $\widehat{\mathbf{K}}$  instead of  $\mathbf{K}$  as the covariance function of the Gaussian process yields the following generative model:

$$\mathbf{w}_{k} \sim p(\mathbf{w}), \ b_{k} \sim p(b),$$

$$\mathbf{W}_{1} = [\mathbf{w}_{k}]_{k=1}^{K}, \mathbf{b} = [b_{k}]_{k=1}^{K}$$

$$\widehat{\mathbf{K}}(\mathbf{x}, \mathbf{y}) = \frac{1}{K} \sum_{k=1}^{K} \sigma(\mathbf{w}_{k}^{T} \mathbf{x} + b_{k}) \sigma(\mathbf{w}_{k}^{T} \mathbf{y} + b_{k})$$

$$\mathbf{F} \mid \mathbf{X}, \mathbf{W}_{1}, \mathbf{b} \sim \mathcal{N}(\mathbf{0}, \widehat{\mathbf{K}}(\mathbf{X}, \mathbf{X}))$$

$$\mathbf{Y} \mid \mathbf{F} \sim \mathcal{N}(\mathbf{F}, \tau^{-1} \mathbf{I}_{N}),$$
(8)

with  $\mathbf{W}_1$  a  $Q \times K$  matrix parametrising our covariance function.

Integrating over the covariance function parameters results in the following predictive distribution:

$$p(\mathbf{Y}|\mathbf{X}) = \int p(\mathbf{Y}|\mathbf{F})p(\mathbf{F}|\mathbf{W}_1, \mathbf{b}, \mathbf{X})p(\mathbf{W}_1)p(\mathbf{b})$$

where the integration is with respect to F,  $W_1$ , and b.

Denoting the  $1 \times K$  row vector

$$\phi(\mathbf{x}, \mathbf{W}_1, \mathbf{b}) = \sqrt{\frac{1}{K}} \sigma(\mathbf{W}_1^T \mathbf{x} + \mathbf{b})$$

and the  $N \times K$  feature matrix  $\Phi = [\phi(\mathbf{x}_n, \mathbf{W}_1, \mathbf{b})]_{n=1}^N$ , we have  $\widehat{\mathbf{K}}(\mathbf{X}, \mathbf{X}) = \Phi\Phi^T$ . We rewrite  $p(\mathbf{Y}|\mathbf{X})$  as

$$p(\mathbf{Y}|\mathbf{X}) = \int \mathcal{N}(\mathbf{Y}; \mathbf{0}, \Phi \Phi^T + \tau^{-1} \mathbf{I}_N) p(\mathbf{W}_1) p(\mathbf{b}) d\mathbf{W}_1 d\mathbf{b},$$

analytically integrating with respect to F.

The normal distribution of  $\mathbf{Y}$  inside the integral above can be written as a joint normal distribution over  $\mathbf{y}_d$ , the d'th columns of the  $N \times D$  matrix  $\mathbf{Y}$ , for d = 1, ..., D. For each term in the joint distribution, following identity [Bishop, 2006, page 93], we introduce a  $K \times 1$  auxiliary random variable  $\mathbf{w}_d \sim \mathcal{N}(0, \mathbf{I}_K)$ ,

$$\mathcal{N}(\mathbf{y}_d; 0, \Phi\Phi^T + \tau^{-1}\mathbf{I}_N) = \int \mathcal{N}(\mathbf{y}_d; \Phi\mathbf{w}_d, \tau^{-1}\mathbf{I}_N) \mathcal{N}(\mathbf{w}_d; 0, \mathbf{I}_K) d\mathbf{w}_d.$$
(9)

Writing  $\mathbf{W}_2 = [\mathbf{w}_d]_{d=1}^D$  a  $K \times D$  matrix, the above is equivalent to

$$p(\mathbf{Y}|\mathbf{X}) = \int p(\mathbf{Y}|\mathbf{X}, \mathbf{W}_1, \mathbf{W}_2, \mathbf{b}) p(\mathbf{W}_1) p(\mathbf{W}_2) p(\mathbf{b})$$

where the integration is with respect to  $W_1$ ,  $W_2$ , and b.

#### 3.2 Variational Inference in the Approximate Model

Our sufficient statistics are  $\mathbf{W}_1$ ,  $\mathbf{W}_2$ , and  $\mathbf{b}$ . To perform variational inference in our approximate model we need to define a variational distribution  $q(\mathbf{W}_1, \mathbf{W}_2, \mathbf{b}) := q(\mathbf{W}_1)q(\mathbf{W}_2)q(\mathbf{b})$ . We define  $q(\mathbf{W}_1)$  to be a Gaussian mixture distribution with two components, factorised over Q:

$$q(\mathbf{W}_1) = \prod_{q=1}^{Q} q(\mathbf{w}_q), \tag{10}$$

$$q(\mathbf{w}_q) = p_1 \mathcal{N}(\mathbf{m}_q, \boldsymbol{\sigma}^2 \mathbf{I}_K) + (1 - p_1) \mathcal{N}(0, \boldsymbol{\sigma}^2 \mathbf{I}_K)$$

with some probability  $p_1 \in [0, 1]$ , scalar  $\sigma > 0$  and  $\mathbf{m}_q \in \mathbb{R}^K$ . We put a similar approximating distribution over  $\mathbf{W}_2$ :

$$q(\mathbf{W}_2) = \prod_{k=1}^K q(\mathbf{w}_k),\tag{11}$$

$$q(\mathbf{w}_k) = p_2 \mathcal{N}(\mathbf{m}_k, \boldsymbol{\sigma}^2 \mathbf{I}_D) + (1 - p_2) \mathcal{N}(0, \boldsymbol{\sigma}^2 \mathbf{I}_D)$$

with some probability  $p_2 \in [0, 1]$ .

We put a simple Gaussian approximating distribution over b:

$$q(\mathbf{b}) = \mathcal{N}(\mathbf{m}, \sigma^2 \mathbf{I}_K). \tag{12}$$

#### 3.3 Evaluating the Log Evidence Lower Bound for Regression

We need to evaluate the log evidence lower bound:

$$\mathcal{L}_{GP-VI} := \int q(\mathbf{W}_1, \mathbf{W}_2, \mathbf{b}) \log p(\mathbf{Y}|\mathbf{X}, \mathbf{W}_1, \mathbf{W}_2, \mathbf{b}) - KL(q(\mathbf{W}_1, \mathbf{W}_2, \mathbf{b})||p(\mathbf{W}_1, \mathbf{W}_2, \mathbf{b})),$$
(13)

where the integration is with respect to  $W_1$ ,  $W_2$ , and b.

For the task of regression we can rewrite the integrand as a sum:

$$\log p(\mathbf{Y}|\mathbf{X}, \mathbf{W}_1, \mathbf{W}_2, \mathbf{b}) = \sum_{d=1}^{D} \log \mathcal{N}(\mathbf{y}_d; \Phi \mathbf{w}_d, \tau^{-1} \mathbf{I}_N)$$
$$= -\frac{ND}{2} \log(2\pi) + \frac{ND}{2} \log(\tau) - \sum_{d=1}^{D} \frac{\tau}{2} ||\mathbf{y}_d - \Phi \mathbf{w}_d||_2^2,$$

as the output dimensions of a multi-output Gaussian process are assumed to be independent. Denote  $\hat{\mathbf{Y}} = \Phi \mathbf{W}_2$ . We can then sum over the rows instead of the columns of  $\hat{\mathbf{Y}}$  and write

$$\sum_{d=1}^{D} \frac{\tau}{2} ||\mathbf{y}_{d} - \widehat{\mathbf{y}}_{d}||_{2}^{2} = \sum_{n=1}^{N} \frac{\tau}{2} ||\mathbf{y}_{n} - \widehat{\mathbf{y}}_{n}||_{2}^{2}.$$

Here  $\hat{\mathbf{y}}_n = \phi(\mathbf{x}_n, \mathbf{W}_1, \mathbf{b})\mathbf{W}_2 = \sqrt{\frac{1}{K}}\sigma(\mathbf{x}_n\mathbf{W}_1 + \mathbf{b})\mathbf{W}_2$ , resulting in the integrand

$$\log p(\mathbf{Y}|\mathbf{X}, \mathbf{W}_1, \mathbf{W}_2, \mathbf{b}) = \sum_{n=1}^{N} \log \mathcal{N}(\mathbf{y}_n; \boldsymbol{\phi}(\mathbf{x}_n, \mathbf{W}_1, \mathbf{b}) \mathbf{W}_2, \tau^{-1} \mathbf{I}_D).$$

This allows us to write the log evidence lower bound as

$$\sum_{n=1}^{N} \int q(\mathbf{W}_1, \mathbf{W}_2, \mathbf{b}) \log p(\mathbf{y}_n | \mathbf{x}_n, \mathbf{W}_1, \mathbf{W}_2, \mathbf{b}) - \text{KL}(q(\mathbf{W}_1, \mathbf{W}_2, \mathbf{b}) || p(\mathbf{W}_1, \mathbf{W}_2, \mathbf{b})).$$

We re-parametrise the integrands in the sum to not depend on  $\mathbf{W}_1, \mathbf{W}_2$ , and b directly, but instead on the standard normal distribution and the Bernoulli distribution. Let  $q(\boldsymbol{\epsilon}_1) = \mathcal{N}(\mathbf{0}, \mathbf{I}_{Q \times K})$  and  $q(\mathbf{z}_{1,q}) = \text{Bernoulli}(p_1)$  for q = 1, ..., Q, and  $q(\boldsymbol{\epsilon}_2) = \mathcal{N}(\mathbf{0}, \mathbf{I}_{K \times D})$  and  $q(\mathbf{z}_{2,k}) = \text{Bernoulli}(p_2)$  for k = 1, ..., K. Finally let  $q(\boldsymbol{\epsilon}) = \mathcal{N}(0, \mathbf{I}_K)$ . We write

$$\mathbf{W}_{1} = \mathbf{z}_{1}(\mathbf{M}_{1} + \boldsymbol{\sigma}\boldsymbol{\epsilon}_{1}) + (1 - \mathbf{z}_{1})\boldsymbol{\sigma}\boldsymbol{\epsilon}_{1},$$

$$\mathbf{W}_{2} = \mathbf{z}_{2}(\mathbf{M}_{2} + \boldsymbol{\sigma}\boldsymbol{\epsilon}_{2}) + (1 - \mathbf{z}_{2})\boldsymbol{\sigma}\boldsymbol{\epsilon}_{2},$$

$$\mathbf{b} = \mathbf{m} + \boldsymbol{\sigma}\boldsymbol{\epsilon},$$
(14)

allowing us to re-write the sum over the integrals in the above equation as

$$\sum_{n=1}^{N} \int q(\mathbf{W}_{1}, \mathbf{W}_{2}, \mathbf{b}) \log p(\mathbf{y}_{n} | \mathbf{x}_{n}, \mathbf{W}_{1}, \mathbf{W}_{2}, \mathbf{b}) d\mathbf{W}_{1} d\mathbf{W}_{2} d\mathbf{b}$$

$$= \sum_{n=1}^{N} \int q(\mathbf{z}_{1}, \boldsymbol{\epsilon}_{1}, \mathbf{z}_{2}, \boldsymbol{\epsilon}_{2}, \boldsymbol{\epsilon}) \log p(\mathbf{y}_{n} | \mathbf{x}_{n}, \mathbf{W}_{1}(\mathbf{z}_{1}, \boldsymbol{\epsilon}_{1}), \mathbf{W}_{2}(\mathbf{z}_{2}, \boldsymbol{\epsilon}_{2}), \mathbf{b}(\boldsymbol{\epsilon}))$$

where each integration is over  $\epsilon_1, z_1, \epsilon_2, z_2$ , and  $\epsilon$ .

$$\mathcal{L}_{\text{GP-MC}} := \sum_{n=1}^{N} \log p(\mathbf{y}_n | \mathbf{x}_n, \widehat{\mathbf{W}}_1^n, \widehat{\mathbf{W}}_2^n, \widehat{\mathbf{b}}^n) - \text{KL}(q(\mathbf{W}_1, \mathbf{W}_2, \mathbf{b}) | | p(\mathbf{W}_1, \mathbf{W}_2, \mathbf{b}))$$

with realisations  $\widehat{\mathbf{W}}_{1}^{n}$ ,  $\widehat{\mathbf{W}}_{2}^{n}$ ,  $\widehat{\mathbf{b}}^{n}$  defined following eq. (14) with  $\widehat{\boldsymbol{\epsilon}}_{1}^{n} \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_{Q \times K})$ ,  $\widehat{\mathbf{z}}_{1,q}^{n} \sim$  Bernoulli $(p_{1})$ ,  $\widehat{\boldsymbol{\epsilon}}_{2}^{n} \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_{K \times D})$ , and  $\widehat{\mathbf{z}}_{2,k}^{n} \sim$  Bernoulli $(p_{2})$ . Following [Blei et al., 2012; Hoffman et al., 2013; Kingma and Welling, 2013; Rezende et al., 2014; Titsias and Lázaro-Gredilla, 2014], optimising the *stochastic* objective  $\mathcal{L}_{\text{GP-MC}}$  we would converge to the same limit as  $\mathcal{L}_{\text{GP-VI}}$ .

This is obtained by Monte-Carlo Integration.

For the second term, we can't calculate the KL divergence between a Gaussian mixture and single Gaussian analytically. So, we do a Monte-Carlo approximation of the same.

$$\mathrm{KL}(q(\mathbf{W}_1)||p(\mathbf{W}_1)) \approx QK(\boldsymbol{\sigma}^2 - \log(\boldsymbol{\sigma}^2) - 1) + \frac{p_1}{2} \sum_{q=1}^{Q} \mathbf{m}_q^T \mathbf{m}_q + C$$

$$\mathrm{KL}(q(\mathbf{b})||p(\mathbf{b})) = \frac{1}{2} (\mathbf{m}^T \mathbf{m} + K(\sigma^2 - \log(\sigma^2) - 1)) + C$$

This is calculated analytically.

#### 3.4 Log Evidence Lower Bound Optimisation

Ignoring the constant terms  $\tau$ ,  $\sigma$  we obtain the maximisation objective

$$\mathcal{L}_{\text{GP-MC}} \propto -\frac{\tau}{2} \sum_{n=1}^{N} ||\mathbf{y}_n - \widehat{\mathbf{y}}_n||_2^2 - \frac{p_1}{2} ||\mathbf{M}_1||_2^2 - \frac{p_2}{2} ||\mathbf{M}_2||_2^2 - \frac{1}{2} ||\mathbf{m}||_2^2.$$
 (15)

Note that in the Gaussian processes literature the terms  $\tau$ ,  $\sigma$  will often be optimised as well.

$$\widehat{\mathbf{W}}_1^n \approx \widehat{\mathbf{z}}_1^n \mathbf{M}_1, \ \widehat{\mathbf{W}}_2^n \approx \widehat{\mathbf{z}}_2^n \mathbf{M}_2, \ \widehat{\mathbf{b}}^n \approx \mathbf{m}.$$

Note that  $\widehat{\mathbf{W}}_1^n$  are not maximum a posteriori (MAP) estimates, but random variable realisations. This gives us

$$\widehat{\mathbf{y}}_n \approx \sqrt{\frac{1}{K}} \sigma(\mathbf{x}_n(\widehat{\mathbf{z}}_1^n \mathbf{M}_1) + \mathbf{m})(\widehat{\mathbf{z}}_2^n \mathbf{M}_2).$$

Scaling the optimisation objective by a positive constant  $\frac{1}{\tau N}$  doesn't change the parameter values at its optimum (as long as we don't optimise with respect to  $\tau$ ). We thus scale the objective to get

$$\mathcal{L}_{\text{GP-MC}} \propto -\frac{1}{2N} \sum_{n=1}^{N} ||\mathbf{y}_n - \widehat{\mathbf{y}}_n||_2^2 - \frac{p_1}{2\tau N} ||\mathbf{M}_1||_2^2 - \frac{p_2}{2\tau N} ||\mathbf{M}_2||_2^2 - \frac{1}{2\tau N} ||\mathbf{m}||_2^2$$
 (16)

# Experiments

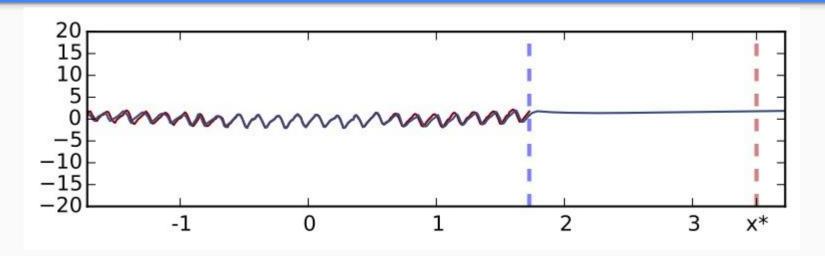
- Need the model to be uncertain in prediction when it sees a point far away from the training set.
- Can be used for dataset augmentation and is of paramount importance in real life high risk applications.
- Experiments on a subset of the CO<sub>2</sub> dataset collected by <u>Keeling et al.</u>, the solar irradiance dataset described in <u>NASA's website</u>, and, Image classification experiments on the MNIST dataset.

# Handling points away from the dataset.

- When asked about a point far away from the dataset, a model must say that it is uncertain in its prediction.
- Standard Neural Networks predict an insensible value with high probability.
- Using softmax probabilities to get uncertainty is not a valid measure (will be shown in an example later)
- Standard dropout networks can be directly used to obtain uncertainty estimates by passing an input sample through it several times and calculating the mean and variance.

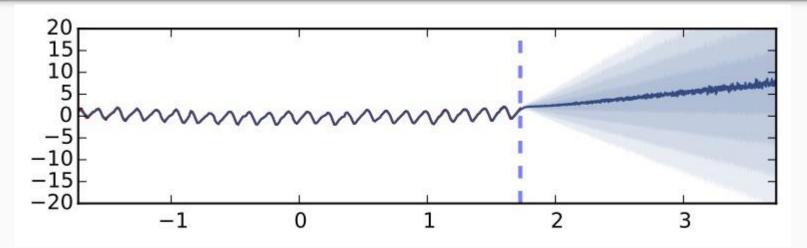
$$egin{aligned} \mathbb{E}(\mathbf{y}^*) &pprox rac{1}{T} \sum_{t=1}^T \widehat{\mathbf{y}}_t^*(\mathbf{x}^*) \ \mathrm{Var}(\mathbf{y}^*) &pprox au^{-1} \mathbf{I}_D \ &+ rac{1}{T} \sum_{t=1}^T \widehat{\mathbf{y}}_t^*(\mathbf{x}^*)^T \widehat{\mathbf{y}}_t^*(\mathbf{x}^*) \ &- \mathbb{E}(\mathbf{y}^*)^T \mathbb{E}(\mathbf{y}^*) \end{aligned}$$

# Results without taking care of uncertainty estimates



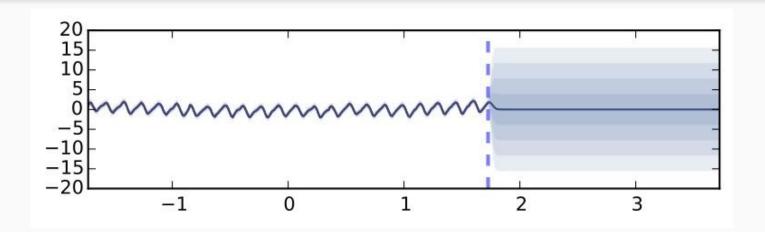
- The training set consists of points only till the blue line. The red line represents a
  point far from the train set.
- No way for us to tell the model is uncertain or not using a point estimate.

# Uncertainty estimates away from the dataset



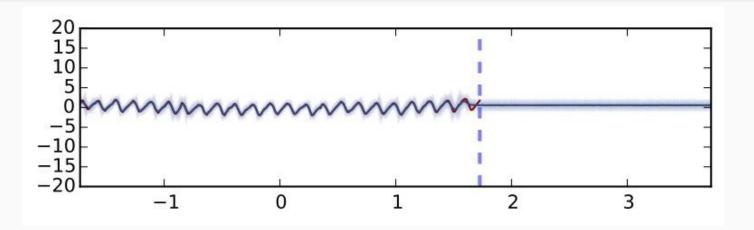
- Exactly the same dropout network performing predictions using uncertainty information and predictive mean (ReLU nonlinearity between layers).
- Clearly the model captures a large amount of uncertain information for the far away test point.
- The uncertainty increases as we go away from the data.

# Comparison to a Gaussian Process



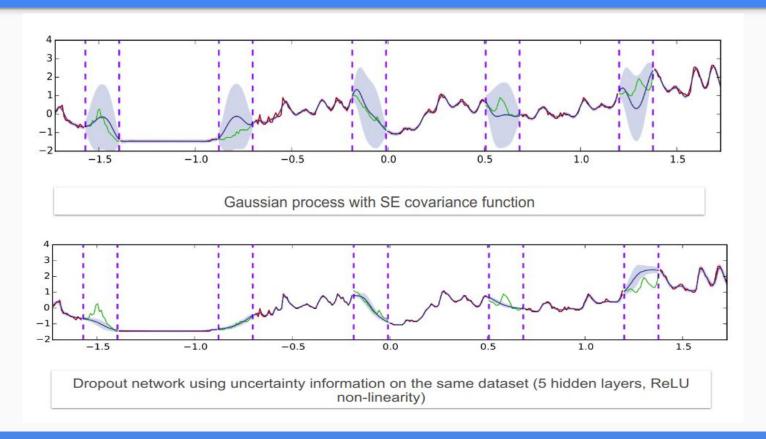
- Uncertainty obtained using a Gaussian Process and a Squared Error(SE) covariance function.
- The uncertainty is clearly higher at points away from the data.
- The estimates look different, however, which should be expected as the ReLU non linearity in the neural network is essentially a different covariance function.

# Effect of changing the non-linearity



- Uncertainty estimates using a dropout network when the nonlinearity used it Tanh.
- The uncertainty doesn't increase far from the data.
- Might be because TanH saturates whereas ReLU does not.
- Not appropriate for tasks where we uncertainty to increase as we go further away from the data.

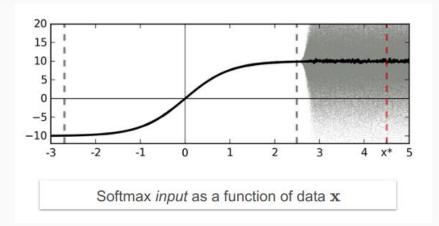
# Experiment on the Solar Irradiance dataset

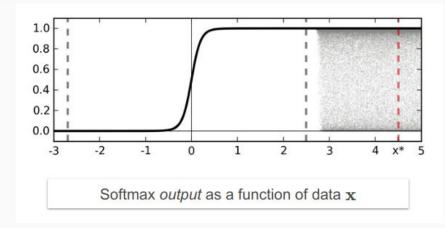


# Experiment on the Solar Irradiance dataset

- Considering the case of interpolation of missing data.
- The green and red lines show the actual data samples and the blue line shows the predictive mean.
- The model fits the data very well as well, but with smaller model uncertainty.
- A well known limitation of variational approximations.
- As dropout can be seen as a variational approximation to the Gaussian process, this is not surprising at all.

# A binary classification example

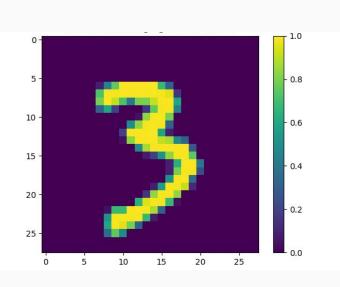




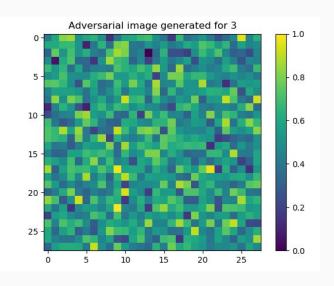
A sketch of softmax input and output for an idealised binary classification problem. Training data is given between the dashed grey lines. Function point estimate is shown with a solid line (TanH for simplicity — left). Function uncertainty is shown with a shaded area. Marked with a dashed red line is a point  $x^*$  far from the training data. Ignoring function uncertainty, point  $x^*$  is classified as class 1 with probability 1.

- The softmax probabilities are not enough to say our model is uncertain or not.
- The standard model would pass the predictive mean through the softmax rather than the entire distribution.
- Passing a point estimate of the mean of a function through a softmax results in highly confident extrapolations with x\* (a point far from the training data) classified as class 1 with probability 1.
- Passing the distribution (shaded area on the left) through a softmax (shaded area on the right) reflects classification uncertainty better far from the training data.
- Taking the mean of this distribution passed through the softmax we get class 1 with probability 0.5 the model's true prediction.

# Adversarial example on the MNIST dataset



```
[[0.05855728]
  [0.77182849]
  [0.13310686]
 [0.00235378]
  [0.02288536]
  ГΘ.
 [0.00095433]
  [0.0103139
0.036066025677021385
Number to generate : 3
[[0.]
 [0.]
  0.7
 [0.]
  [0.]
 [0.]
 [0.]
 [0.]
 [0.]]
Number to generate: 3
```



# Adversarial examples on the MNIST dataset

- Mean probability and variance generated by passing the input through a dropout model 1000 times and taking the mean and variance.
- A clear difference in outputs when a real image was input to the model versus an adversarial example of the digit 3.
- The mean values of the probability clearly tells that the model was uncertain even if we got a high probability.
- The variance, however, is still not high enough for us to conclusively speak about the uncertainty estimate.

# **Future Work**

- Tackle adversarial examples using model uncertainty, making the model robust to adversarial attacks.
  - <u>Lewis et al</u> show failure cases for MC dropout in detecting adv. Examples.
- Look at Uncertainty Information in other areas of computer vision <u>Kendall et al.</u>

# Questions?