

Solution to PSET 4

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- **Do not** distribute the solutions outside the class.
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1. **(A PAC-Bayesian Theorem)** (a) We have

$$\begin{aligned}
 (m-1)D(\hat{L}_S(Q)||L(Q)) &= (m-1)D\left(\mathbb{E}_{h \sim Q} \hat{L}_S(h) || \mathbb{E}_{h \sim Q} L(h)\right) \\
 &\stackrel{\text{Jensen}}{\leq} (m-1)\mathbb{E}_{h \sim Q} \left(D(\hat{L}_S(h)||L(h)) \right) \\
 &\stackrel{\text{Donsker-Varadhan}}{\leq} D(Q||P) + \ln \mathbb{E}_{h \sim P} [e^{((m-1)D(\hat{L}_S(h)||L(h)))}].
 \end{aligned} \tag{1}$$

(b)

I Since f is non-negative and non-increasing, we have

$$\mathbb{P}(f(X) \geq f(\epsilon)) \leq \mathbb{P}(X \leq \epsilon) \leq e^{-mf(\epsilon)}.$$

Thus, for any $t \geq 0$, we have:

$$\mathbb{P}(f(X) \geq t) \leq \min(1, e^{-mt}).$$

Hence,

$$\begin{aligned}
 \mathbb{E}[e^{(m-1)f(X)}] &= \int_0^\infty \mathbb{P}(e^{(m-1)f(X)} \geq z) dz \\
 &= \int_0^\infty \mathbb{P}(f(X) \geq \ln z^{\frac{1}{m-1}}) dz \\
 &\leq \int_1^\infty z^{-\frac{m}{m-1}} dz + \int_0^1 1 dz \\
 &= m
 \end{aligned}$$

II Note that, the r.v. $\hat{L}_S(h)$ follows a binomial distribution with mean $L(h)$. Hence, using the standard Chernoff estimate for Binomial distribution, we have

$$\mathbb{P}(\hat{L}_S(h) \leq \epsilon) \leq e^{-mD^+(\epsilon||L(h))}.$$

Hence, using the result from part I., we have

$$\mathbb{E}_{S \sim D^m} [e^{(m-1)D^+(\hat{L}_S(h)||L(h))}] \leq m.$$

III Taking expectation of the above w.r.t. prior distribution $h \sim P$, we have:

$$\mathbb{E}_{h \sim P} \mathbb{E}_{S \sim D^m} [e^{(m-1)D^+(\hat{L}_S(h)||L(h))}] = \mathbb{E}_{S \sim D^m} \mathbb{E}_{h \sim P} [e^{(m-1)D^+(\hat{L}_S(h)||L(h))}] \leq m$$

The result follows from an application of Markov's inequality.

IV The PAC-Bayes bound follows upon combining the above result with Eqn. (1).

2. **(Non-Parametric Least Square Estimation)** We first prove the following basic inequality:

$$\frac{1}{n} \sum_{i=1}^n \left(\hat{f}(x_i) - f^*(x_i) \right)^2 \leq \frac{2}{n} \sum_{i=1}^n \epsilon_i \left(\hat{f}(x_i) - f^*(x_i) \right). \quad (2)$$

This follows immediately from the definition of the least-square estimator. In particular, we have

$$\begin{aligned} \frac{1}{n} \sum_{i=1}^n \left(Y_i - \hat{f}(x_i) \right)^2 &= \frac{1}{n} \sum_{i=1}^n \left(f^*(x_i) + \epsilon_i - \hat{f}(x_i) \right)^2 \\ &\stackrel{(a)}{\leq} \frac{1}{n} \sum_{i=1}^n \left(f^*(x_i) + \epsilon_i - f^*(x_i) \right)^2 \\ &= \frac{1}{n} \sum_{i=1}^n \epsilon_i^2, \end{aligned}$$

where the inequality (a) follows from the least-square optimality of \hat{f} . The above inequality implies that

$$\frac{1}{n} \sum_{i=1}^n \left(\hat{f}(x_i) - f^*(x_i) \right)^2 \leq \frac{2}{n} \sum_{i=1}^n \epsilon_i \left(\hat{f}(x_i) - f^*(x_i) \right),$$

which proves (2). The RHS of (2) may be bounded further as follows:

$$\frac{1}{n} \sum_{i=1}^n \left(\hat{f}(x_i) - f^*(x_i) \right)^2 \leq \frac{4}{n} \left(\sup_{f \in S_{\alpha, \gamma}(C_{\max}, L)} \sum_{i=1}^n \epsilon_i f(x_i) \right). \quad (3)$$

Taking expectation (w.r.t. $\{\epsilon_i\}_{i=1}^n$) of both sides, we have

$$\mathbb{E} \left[\frac{1}{n} \sum_{i=1}^n \left(\hat{f}(x_i) - f^*(x_i) \right)^2 \right] \leq \frac{4}{n} \mathcal{G}(S_{\alpha, \gamma}(C_{\max}, L); \mathbf{x}),$$

where the RHS denotes the Gaussian complexity of the function class $S_{\alpha, \gamma}(C_{\max}, L)$ induced by the points \mathbf{x} .

Note that, the diameter of the class $S_{\alpha, \gamma}(C_{\max}, L)$ is $\delta = \sup_{x \in [0, 1]} |f(x) - g(x)| \leq 2C_{\max}$.

Define the Gaussian process, $X_f = \sum_{i=1}^n \epsilon_i f(x_i), \forall f \in S_{\alpha, \gamma}(C_{\max}, L)$. For any two $f, g \in S_{\alpha, \gamma}(C_{\max}, L)$, we have

$$X_f - X_g = \sum_{i=1}^n \epsilon_i (f(x_i) - g(x_i)).$$

The variance of $X_f - X_g$ is bounded as

$$\text{Var}(X_f - X_g) = \sum_{i=1}^n (f(x_i) - g(x_i))^2 \leq n \|f - g\|_{\infty}^2.$$

Hence, for each $f, g \in S_{\alpha, \gamma}(C_{\max}, L)$, $X_f - X_g \sim \text{SG}(0, n \|f - g\|_{\infty}^2)$. Hence, using **Dudley's entropy integral**, we have

$$\mathbb{E} \left[\frac{1}{n} \sum_{i=1}^n \left(\hat{f}(x_i) - f^*(x_i) \right)^2 \right] \leq \frac{48}{\sqrt{n}} \int_0^{C_{\max}} \sqrt{C \left(\frac{1}{u} \right)^{1/(\alpha+\gamma)}} du = O \left(\frac{C_{\max}^{1-\frac{1}{2(\alpha+\gamma)}}}{\sqrt{n}} \right).$$

3. **(Online Mirror Descent)** (a) Let $w^* \in \arg \min D_{\psi}(w, w'_{t+1})$. Setting the gradient of the objective function $D_{\psi}(w, w'_{t+1})$ to zero, we have

$$0 = \nabla D_{\psi}(w^*, w'_{t+1}) = \nabla \psi(w^*) - \nabla \psi(w'_{t+1}) = \nabla \psi(w^*) - \nabla \psi(w_t) + \eta \nabla f_t(w_t).$$

On the other hand, by definition,

$$0 = \nabla f_t(w_t) + \frac{1}{\eta} (\nabla \psi(w_{t+1}) - \nabla \psi(w_t)).$$

Combining the above two equations, we conclude that

$$\nabla \psi(w^*) = \nabla \psi(w_{t+1}).$$

Assuming that the mapping $\nabla \psi : \Omega \rightarrow \Omega$ is invertible, we have

$$w_{t+1} = w^*.$$

- (b) The result follows upon simplifying the RHS using the definition of Bregman divergence and using the fact that

$$\nabla \psi(w'_{t+1}) = \nabla \psi(w_t) - \eta \nabla f_t(w_t).$$

- (c) Since $w_{t+1} = \arg \min D_{\psi}(w, w'_{t+1})$, using the first order optimality criterion for convex function $D_{\psi}(\cdot, w'_{t+1})$, we have for any $u \in \Omega$:

$$\langle \nabla \psi(w_{t+1}) - \nabla \psi(w'_{t+1}), u - w_{t+1} \rangle \geq 0. \quad (4)$$

Hence,

$$\begin{aligned}
& D_\psi(u, w'_{t+1}) - D_\psi(u, w_{t+1}) \\
&= \underbrace{\psi(w_{t+1}) - \psi(w'_{t+1}) - \langle \nabla \psi(w'_{t+1}), w_{t+1} - w'_{t+1} \rangle}_{\geq 0 \text{ (as } \psi \text{ is convex)}} + \underbrace{\langle \nabla \psi(w_{t+1}) - \nabla \psi(w'_{t+1}), u - w_{t+1} \rangle}_{\geq 0 \text{ from eqn.(4)}} \\
&\geq 0
\end{aligned}$$

(d) The result follows from (b) and (c) upon telescoping the summation.

(e) In the experts setting, we have for any $w_t \in \Delta_N$:

$$f_t(w_t) = \langle w_t, l_t \rangle.$$

Take $\psi : \Omega \rightarrow \mathbb{R}$ to be the negative entropy function, i.e., $\psi(w) = \sum_i w_i \ln w_i$. Hence,

$$D_\psi(w, u) = \sum_i w_i \ln \frac{w_i}{u_i} - \sum_i (w_i - u_i).$$

Also, the mirror action w'_{t+1} satisfies the equation:

$$1 + \ln w'_{t+1,i} = 1 + \ln w_{t,i} - \eta l_{t,i},$$

i.e, $w'_{t+1,i} = w_{t,i} \exp(-\eta l_{t,i})$, which leads to the **Hedge** algorithm. Thus,

$$\begin{aligned}
D_\psi(w_t, w'_{t+1}) &= \sum_i w_{t,i} \ln \frac{w_{t,i}}{w'_{t+1,i}} - 1 + \sum_i w'_{t+1,i} \\
&= \sum_i w_{t,i} (e^{-\eta l_{t,i}} - \eta l_{t,i} - 1) \\
&\leq \eta^2,
\end{aligned}$$

where in the last step, we have used the inequality $e^{-x} \leq 1 - x + x^2, \forall x \geq 0$ and the fact that $w_t \in \Delta_N$. Also, taking w_1 to be uniform over $[N]$, we have $D_\psi(u, w_1) \leq \log N$. Hence, using the regret bound for OMD from part (d), we have

$$R_T^{\text{Hedge}} \leq \frac{\log N}{\eta} + T\eta.$$

Choosing $\eta = \sqrt{\frac{\log N}{T}}$, we have $R_T^{\text{Hedge}} \leq 2\sqrt{T \log N}$.

4. (Foresight and Hindsight Regret for the IID Cost Model)

(1) Consider the clean event \mathcal{E} :

$$\mathcal{E} = \{|\bar{\mu}_T(a) - \mu(a)| \leq r_T, \forall a\},$$

where $r_T = \sqrt{\frac{\log(KT)}{T}}$. Using Hoeffding's inequality and union bound, we have $\mathbb{P}(\mathcal{E}) \geq 1 - \frac{1}{KT^2}$. Assuming the clean event, we have for each arm $a \in \mathcal{A}$:

$$\mathbb{E}(\text{cost}(a)) \leq \text{cost}(a) + \sqrt{T \log(KT)}.$$

Hence,

$$\min_a \mathbb{E}(\text{cost}(a)) \leq \min_a \text{cost}(a) + \sqrt{T \log(KT)}$$

On the other hand, on \mathcal{E}^c , we have the trivial lower bound, $\min_a \text{cost}(a) \geq 0$. Hence,

$$\begin{aligned} & \mathbb{E}(\min_a \text{cost}(a)) \\ & \geq \left(1 - \frac{1}{KT^2}\right) \left(\min_a \mathbb{E}(\text{cost}(a)) - \sqrt{T \log(KT)} \right). \end{aligned}$$

Noting that $\min_a \mathbb{E}(\text{cost}(a)) \leq T$, the above equation implies that

$$\min_a \mathbb{E}(\text{cost}(a)) \leq \mathbb{E}(\min_a \text{cost}(a)) + O(\sqrt{T \log(KT)}).$$

(2) Fix any algorithm. We consider an ensemble of problem instances where all-arms have 0 – 1 costs with mean $\frac{1}{2}$. Hence, irrespective of the choice that the algorithm makes, we have

$$\mathbb{E}(\text{cost}(\text{ALG})) = \frac{T}{2}.$$

Further, from Hoeffding's inequality, we have

$$\mathbb{E}[\min_a \text{cost}(a)] \leq \frac{T}{2} - \Omega(\sqrt{T \log K}).$$

Combining the above two results, we conclude that

$$\mathbb{E}[\text{cost}(\text{ALG}) - \min_a \text{cost}(a)] \geq \Omega(\sqrt{T \log K}),$$

where the expectation is over the random ensemble of problem instances and the randomness of the algorithm. Hence, for any algorithm, there must exist a problem instant for which the above lower bound holds.

(3) Under the clean event, which happens w.h.p., we have for all $a \in \mathcal{A}$:

$$|\text{cost}(a) - T\mu(a)| \leq \sqrt{T \log T}. \quad (5)$$

Next, divide the set of arms in two categories:

- (a) **Category-I** ($a \in \mathcal{A} : T\Delta(a) > 3\sqrt{T \log T}$): Clearly, $\arg \min_a \text{cost}(a) \equiv a^{**}$ does not belong to this category.
- (b) **Category-II** ($a \in \mathcal{A} : T\Delta(a) \leq 3\sqrt{T \log T}$): Clearly, a^{**} belongs to this category. Hence, we have

$$\begin{aligned}
\text{cost}(a^{**}) &\geq T\mu(a^{**}) - \sqrt{T \log T} \\
&= T\Delta(a^{**}) + T\mu(a^*) - \sqrt{T \log T} \\
&= \mathbb{E}(\text{cost}(a^*)) + \log(T) \left(\Delta(a^{**}) \frac{T}{\log T} - \sqrt{\frac{T}{\log T}} \right).
\end{aligned}$$

Next, we lower bound the term in the bracket. Define $x = \sqrt{\frac{T}{\log T}}$. Using calculus, we see that

$$\Delta(a^{**})x^2 - x \geq -\frac{1}{4\Delta(a^{**})}.$$

This yields,

$$\text{cost}(a^{**}) \geq \mathbb{E}(\text{cost}(a^*)) - \frac{\log T}{4\Delta(a^{**})}.$$

Since, the above event (5) happens w.h.p., we have

$$\mathbb{E}(\min_a \text{cost}(a)) \geq \min_a \mathbb{E}(\text{cost}(a)) - O(\log(T)/\Delta(a^{**})).$$

Thus, we see that the instance dependent logarithmic regret bound carries over for hindsight regret for UCB and Successive Elimination algorithms.