

Problem 1

Let us first define some notations. We define $\hat{L}_S(Q) = \mathbb{E}_Q[\hat{L}_S(h)]$ and $L(Q) = \mathbb{E}_Q[L(h)]$.

- (a) Let $g(h) = (m-1)D(\hat{L}_S(h)||L(h))$. We use Donsker-Varadhan inequality on distributions P and Q and obtain:

$$\begin{aligned} \mathbb{E}_Q[g(h)] &\leq D(Q||P) + \log(\mathbb{E}_P[\exp(g(h))]) \\ \implies \mathbb{E}_Q[(m-1)D(\hat{L}_S(h)||L(h))] &\leq D(Q||P) + \log\left(\mathbb{E}_P[\exp((m-1)D(\hat{L}_S(h)||L(h)))]\right) \\ \implies (m-1)D(\hat{L}_S(Q)||L(Q)) &\leq D(Q||P) + \log\left(\mathbb{E}_P[\exp((m-1)D(\hat{L}_S(h)||L(h)))]\right) \end{aligned}$$

The last inequality holds because KL Divergence is convex in both arguments and using Jensen's inequality, we obtain $D(\hat{L}_S(Q)||L(Q)) \leq \mathbb{E}_Q[D(\hat{L}_S(h)||L(h))]$.

- (b) Now, we proceed to prove the upper bound for the second term in the RHS of the expression.

I We have the $f(\cdot)$ is an non-negative non-increasing function. So,

$$P(e^{(m-1)f(X)} \geq e^{(m-1)f(\epsilon)}) = P(f(X) \geq f(\epsilon)) = P(X \leq \epsilon) \leq e^{-mf(\epsilon)} \quad (1)$$

Now, as $e^{(m-1)f(X)}$ is a non-negative random variable, we have

$$\mathbb{E}[e^{(m-1)f(X)}] = \int_0^\infty P(e^{(m-1)f(X)} \geq t) dt$$

Let $t = e^{(m-1)f(\epsilon)} \implies dt = (m-1)e^{(m-1)f(\epsilon)} f'(\epsilon) d\epsilon$. So,

$$\begin{aligned} \mathbb{E}[e^{(m-1)f(X)}] &= \int_0^\infty P(e^{(m-1)f(X)} \geq t) dt \\ &= \int_0^\infty P(e^{(m-1)f(X)} \geq e^{(m-1)f(\epsilon)}) (m-1)e^{(m-1)f(\epsilon)} f'(\epsilon) d\epsilon \\ &\stackrel{(1)}{\leq} (m-1) \int_0^\infty e^{-mf(\epsilon)} e^{(m-1)f(\epsilon)} f'(\epsilon) d\epsilon \\ &= (m-1) \int_0^\infty e^{-f(\epsilon)} f'(\epsilon) d\epsilon \\ &\leq m \left(e^{-f(0)} - e^{-f(\infty)} \right) \\ &\leq m \end{aligned} \quad (2)$$

The last inequality follows as $e^{-f(0)} - e^{-f(\infty)} \leq e^{-f(0)}$ and $f(0) \in [0, \infty) \implies -f(0) \in (-\infty, 0] \implies e^{-f(0)} \in (0, 1]$.

- II Let $f(\epsilon) = D^+(\epsilon||p)$ where X_1, \dots, X_m are i.i.d random variables in $[0, 1]$ and $p = \mathbb{E}[X_1]$. By definition, $f(\epsilon) \geq 0$.

Now, we show that $f(\epsilon)$ is non-increasing in $[0, 1]$. For, $\epsilon \geq p$, $f'(\epsilon) = 0$ by definition. As $f(\epsilon)$ differentiable for $\epsilon \leq p$, we show that $f'(\epsilon) \leq 0$ in that range.

Proof. By definition, for $\epsilon \leq p$, $f(\epsilon) = D(\epsilon||p) = \epsilon \log\left(\frac{\epsilon}{p}\right) + (1 - \epsilon) \log\left(\frac{1-\epsilon}{1-p}\right)$.

$$\begin{aligned} f'(\epsilon) &= 1 + \log \epsilon - \log p - 1 - \log(1 - \epsilon) + \log(1 - p) \\ &= \log\left(\frac{\epsilon}{1 - \epsilon}\right) - \log\left(\frac{p}{1 - p}\right) \\ &= \log\left(\frac{\epsilon - p\epsilon}{p - p\epsilon}\right) \\ &\leq \log\left(\frac{\epsilon}{p}\right) \leq 0 \end{aligned} \quad \text{If } a \leq b, \frac{a - ab}{b - ab} \leq \frac{a}{b}$$

So, $f(\epsilon) = D^+(\epsilon||\mathbb{E}[X_1])$ is a non-increasing and non-negative function. □

Let $X_i = \ell(h, z_i)$, where z_i is the i th training sample, then $\bar{X} = \hat{L}_S(h)$ and $\mathbb{E}[X] = L(h)$. So, using Chernoff-Hoeffding's inequality, we have:

$$P(\hat{L}_S(h) \leq \epsilon) \leq \exp(-mD^+(\epsilon||L(h))) \quad (3)$$

Using result (2), (3) and considering $f(\hat{L}_S(h)) = D^+(\hat{L}_S(h)||L(h))$, we obtain:

$$\mathbb{E}_{S \sim \mathcal{D}^m} [e^{(m-1)D^+(\hat{L}_S(h)||L(h))}] \leq m \quad (4)$$

III Consider the random variable $Y_S = \mathbb{E}_{h \sim P} [e^{(m-1)D^+(\hat{L}_S(h)||L(h))}]$ which depends on the sample S from the distribution. Using Markov's inequality, we get:

$$P\left(Y_S \geq \frac{2\mathbb{E}_{S \sim \mathcal{D}^m} [Y_S]}{\delta}\right) \leq \frac{\delta}{2} \implies P\left(Y_S \leq \frac{2\mathbb{E}_{S \sim \mathcal{D}^m} [Y_S]}{\delta}\right) \geq 1 - \frac{\delta}{2} \quad (5)$$

Now,

$$\begin{aligned} \mathbb{E}_{S \sim \mathcal{D}^m} [Y_S] &= \mathbb{E}_{S \sim \mathcal{D}^m} [\mathbb{E}_{h \sim P} [e^{(m-1)D^+(\hat{L}_S(h)||L(h))}]] && \text{By definition} \\ &= \mathbb{E}_{h \sim P} [\mathbb{E}_{S \sim \mathcal{D}^m} [e^{(m-1)D^+(\hat{L}_S(h)||L(h))}]] && \text{Fubini's theorem} \\ &\stackrel{(4)}{\leq} \mathbb{E}_{h \sim P} [m] = m && (6) \end{aligned}$$

From (5) and (6), with probability at least $1 - \frac{\delta}{2}$, $\mathbb{E}_{h \sim P} [e^{(m-1)D^+(\hat{L}_S(h)||L(h))}] \leq \frac{2m}{\delta}$. ■

Problem 2

We have a function class $S_{\alpha,\gamma}(C_{\max}, L)$ which is defined as:

$$S_{\alpha,\gamma}(C_{\max}, L) = \{f : [0, 1] \rightarrow \mathbb{R} : |f^{(j)}|_{\infty} \leq C_{\max}, \forall 0 \leq j \leq \alpha, \text{ and} \\ |f^{\alpha}(x) - f^{\alpha}(y)| \leq L|x - y|^{\gamma}, \forall x, y \in [0, 1]\}$$

We have $y_i = f^*(x_i) + \epsilon_i, \forall i \in [n]$ where $\epsilon \sim \mathcal{N}(0, 1)$ and \hat{f} is the minimizer of $\sum_{i=1}^n (y_i - f(x_i))^2$ for $f \in S_{\alpha,\gamma}(C_{\max}, L)$. Let $\|\hat{f} - f^*\|_n = \sqrt{\frac{1}{n} \sum_{i=1}^n (\hat{f}(x_i) - f^*(x_i))^2}$. So, $\text{MSE} = \mathbb{E} \left[\|\hat{f} - f^*\|_n^2 \right]$.

Furthermore, for $f \in S_{\alpha,\gamma}(C_{\max}, L)$, we have $\|f\|_n = \sqrt{\frac{1}{n} \sum_{i=1}^n f(x_i)^2} \leq \sqrt{\frac{1}{n} \sum_{i=1}^n C_{\max}^2} = C_{\max}$. For ease of notation, let $S_{\alpha,\gamma}(C_{\max}, L)$ be denoted by \mathcal{F} hereafter.

Lemma 1. $\|\hat{f} - f^*\|_n^2 \leq \frac{2}{n} \sum_{i=1}^n \epsilon_i (\hat{f}(x_i) - f^*(x_i))$

Proof.

$$\begin{aligned} \|\hat{f} - f^*\|_n^2 &= \frac{1}{n} \sum_{i=1}^n (\hat{f}(x_i) - f^*(x_i))^2 \\ &= \frac{1}{n} \sum_{i=1}^n (\hat{f}(x_i) - y_i + y_i - f^*(x_i))^2 \\ &= \frac{1}{n} \sum_{i=1}^n (\hat{f}(x_i) - y_i)^2 + (y_i - f^*(x_i))^2 + 2(\hat{f}(x_i) - y_i)(f^*(x_i) - y_i) \\ &\leq \frac{2}{n} \sum_{i=1}^n (y_i - f^*(x_i))^2 + (\hat{f}(x_i) - y_i)(f^*(x_i) - y_i) \\ &= \frac{2}{n} \sum_{i=1}^n \epsilon_i^2 + \epsilon_i (\hat{f}(x_i) - f^*(x_i) + f^*(x_i) - y_i) \\ &= \frac{2}{n} \sum_{i=1}^n \epsilon_i^2 + \epsilon_i (\hat{f}(x_i) - f^*(x_i) - \epsilon_i) \\ &= \frac{2}{n} \sum_{i=1}^n \epsilon_i (\hat{f}(x_i) - f^*(x_i)) \end{aligned}$$

□

Now, using Lemma 1

$$\begin{aligned} \|\hat{f} - f^*\|_n^2 &\leq \frac{2}{n} \sum_{i=1}^n \epsilon_i (\hat{f}(x_i) - f^*(x_i)) \leq \frac{4}{n} \sup_{f \in \mathcal{F}} \sum_{i=1}^n \epsilon_i f(x_i) \\ \implies \text{MSE} &= \mathbb{E} \left[\|\hat{f} - f^*\|_n^2 \right] \leq 4 \cdot \mathbb{E} \left[\frac{1}{n} \sup_{f \in \mathcal{F}} \sum_{i=1}^n \epsilon_i f(x_i) \right] \end{aligned} \quad (7)$$

Theorem 2 (Dudley's Entropy Integral). *The Gaussian complexity of a function class \mathcal{F} can be upper bounded as the covering number of the function class in sup-norm as:*

$$\mathbb{E} \left[\frac{1}{n} \sup_{f \in \mathcal{F}} \langle \epsilon, f \rangle \right] \leq \frac{12}{\sqrt{n}} \int_0^D \sqrt{\log N(\delta, \mathcal{F}, \|\cdot\|_\infty)} d\delta$$

where D is the diameter of \mathcal{F} .

Proof. Let V_j be a $2^{-j}D$ -cover of \mathcal{F} in ℓ_∞ norm. So, by definition, for any $f \in \mathcal{F}$, $\exists v_f^j \in V_j$, we have $\|f - v_f^j\|_\infty \leq 2^{-j}D$. Now, from the definition of Gaussian complexity, we have (ignoring the factor of $1/n$):

$$\begin{aligned} \mathbb{E} \left[\sup_{f \in \mathcal{F}} \langle \epsilon, f \rangle \right] &= \mathbb{E} \left[\sup_{f \in \mathcal{F}} \left(\langle \epsilon, f - v_f^m \rangle + \sum_{j=1}^m \langle \epsilon, v_f^j - v_f^{j-1} \rangle \right) \right] \\ &\leq \mathbb{E} \left[\sup_{f \in \mathcal{F}} \langle \epsilon, f - v_f^m \rangle \right] + \sum_{j=1}^m \mathbb{E} \left[\sup_{v \in V_j, v' \in V_{j-1}} \langle \epsilon, v - v' \rangle \right] \\ &\stackrel{(a)}{\leq} \mathbb{E} \left[\sup_{f \in \mathcal{F}} \langle \epsilon, f - v_f^m \rangle \right] + \sum_{j=1}^m \sigma_j \sqrt{2 \log |V_j| |V_{j-1}|} \\ &\stackrel{(b)}{\leq} \mathbb{E} \left[\sup_{f \in \mathcal{F}} \langle \epsilon, f - v_f^m \rangle \right] + \sum_{j=1}^m 2\sigma_j \sqrt{\log |V_j|} \end{aligned}$$

The statement (a) follows from the definition of Massart's Lemma

(b) holds because by construction, we have $|V_j| \geq |V_{j-1}|$.

Now, $\sigma_j = \sqrt{\text{Var}(\langle \epsilon, v - v' \rangle)} = \sqrt{n} \|v - v'\|_n = \sqrt{n} \|v - f + f - v'\|_n \leq \sqrt{n} (\|f - v'\|_n + \|f - v\|_n)$. The last inequality follows from the triangle inequality for the $\|\cdot\|_n$ -norm. Furthermore, by definition of v, v' , we have $\sqrt{n} (\|f - v'\|_n + \|f - v\|_n) \leq \sqrt{n} (2^{-j}D + 2^{-(j-1)}D) = 3 \cdot 2^{-j} \sqrt{n} D \implies \sigma_j \leq 3 \cdot 2^{-j} \sqrt{n} D$. So,

$$\begin{aligned} \mathbb{E} \left[\sup_{f \in \mathcal{F}} \langle \epsilon, f \rangle \right] &\leq \mathbb{E} \left[\sup_{f \in \mathcal{F}} \langle \epsilon, f - v_f^m \rangle \right] + 6\sqrt{n}D \sum_{j=1}^m 2^{-j} \sqrt{\log N(2^{-j}D, \mathcal{F}, \|\cdot\|_\infty)} \\ &\leq \mathbb{E} \left[\sup_{f \in \mathcal{F}} \langle \epsilon, f - v_f^m \rangle \right] + 12\sqrt{n}D \sum_{j=1}^m (2^{-j} - 2^{-(j-1)}) \sqrt{\log N(2^{-j}D, \mathcal{F}, \|\cdot\|_\infty)} \\ &\leq \mathbb{E} \left[\sup_{f \in \mathcal{F}} \langle \epsilon, f - v_f^m \rangle \right] + 12\sqrt{n} \sum_{j=1}^m \int_{2^{-j}D}^{2^{-(j-1)}D} \sqrt{\log N(\delta, \mathcal{F}, \|\cdot\|_\infty)} d\delta \\ &\leq \mathbb{E} \left[\sup_{f \in \mathcal{F}} \langle \epsilon, f - v_f^m \rangle \right] + 12\sqrt{n} \int_{2^{-m}D}^D \sqrt{\log N(\delta, \mathcal{F}, \|\cdot\|_\infty)} d\delta \end{aligned}$$

Now, if $m \rightarrow \infty$, $2^{-m} \rightarrow 0$. Further, $\left\|f - v_f^m\right\|_n \leq 2^{-m}D \rightarrow 0$, so, $|f - v_f^m| \rightarrow 0 \implies$ the first term on RHS vanishes. Finally, dividing both sides by n , we get:

$$\mathbb{E} \left[\frac{1}{n} \sup_{f \in \mathcal{F}} \langle \epsilon, f \rangle \right] \leq \frac{12}{\sqrt{n}} \int_0^D \sqrt{\log N(\delta, \mathcal{F}, \|\cdot\|_\infty)} d\delta$$

□

It is known that:

$$\log N(\delta, \mathcal{F}, \|\cdot\|_\infty) \leq C \left(\frac{1}{\delta} \right)^{1/(\alpha+\gamma)} \quad (8)$$

Now, we compute the diameter of \mathcal{F} . $D = \sup_{f, f' \in \mathcal{F}} |f - f'| \leq 2\|f\|_\infty \leq 2C_{\max}$. Using (7), (8) and Theorem 2, we obtain:

$$\begin{aligned} \text{MSE} &= \mathbb{E} \left[\left\| \hat{f} - f^* \right\|_n^2 \right] \\ &\leq 4 \cdot \mathbb{E} \left[\frac{1}{n} \sup_{f \in \mathcal{F}} \sum_{i=1}^n \epsilon_i f(x_i) \right] \\ &\leq \frac{48}{\sqrt{n}} \int_0^D \sqrt{\log N(\delta, \mathcal{F}, \|\cdot\|_\infty)} d\delta \\ &= 48 \sqrt{\frac{C}{n}} \int_0^{2C_{\max}} \delta^{-\frac{1}{2(\alpha+\gamma)}} d\delta \\ &= 48 \sqrt{\frac{C}{n}} (2C_{\max})^{1-\frac{1}{2(\alpha+\gamma)}} = \mathcal{O} \left(\frac{1}{\sqrt{n}} \right) \end{aligned}$$

■

Problem 3

We define w'_{t+1} as follows:

$$\nabla\psi(w'_{t+1}) = \nabla\psi(w_t) - \eta\nabla f_t(w_t) \quad (9)$$

Regarding Bregman divergence, it is to be noted that $D_\psi(u, w) \geq 0$ and equality is attained iff $u = w$. This can be proved as follows: by definition, Bregman divergence is defined for strictly convex function, \implies for $u \neq w$, $\psi(u) > \psi(w) + \langle \nabla\psi(w), u - w \rangle \implies D_\psi(u, w) > 0$. For $u = w$, $D_\psi(u, w) = 0$.

(a) By definition:

$$\begin{aligned} w_{t+1} &= \operatorname{argmin}_{w \in \Omega} \left[\langle w, \nabla f_t(w_t) \rangle + \frac{1}{\eta} D_\psi(w, w_t) \right] \\ &= \operatorname{argmin}_{w \in \Omega} [\langle w, \eta \nabla f_t(w_t) \rangle + D_\psi(w, w_t)] \\ &\stackrel{(9)}{=} \operatorname{argmin}_{w \in \Omega} [\langle w, \nabla\psi(w_t) - \nabla\psi(w'_{t+1}) \rangle + D_\psi(w, w_t)] \\ &= \operatorname{argmin}_{w \in \Omega} [\langle w, \nabla\psi(w_t) - \nabla\psi(w'_{t+1}) \rangle + \psi(w) - \psi(w_t) - \langle \nabla\psi(w_t), w - w_t \rangle] \\ &= \operatorname{argmin}_{w \in \Omega} [-\langle w, \nabla\psi(w'_{t+1}) \rangle + \psi(w)] \\ &= \operatorname{argmin}_{w \in \Omega} [\psi(w) - \psi(w'_{t+1}) - \langle \nabla\psi(w'_{t+1}), w - w'_{t+1} \rangle] \\ &= \operatorname{argmin}_{w \in \Omega} D_\psi(w, w'_{t+1}) \end{aligned} \quad (10)$$

(b) We begin from RHS:

$$\begin{aligned} D_\psi(u, w_t) - D_\psi(u, w'_{t+1}) + D_\psi(w_t, w'_{t+1}) &= \psi(u) - \psi(w_t) - \langle \nabla\psi(w_t), u - w_t \rangle \\ &\quad - \psi(u) + \psi(w'_{t+1}) + \langle \nabla\psi(w'_{t+1}), u - w'_{t+1} \rangle \\ &\quad + \psi(w_t) - \psi(w'_{t+1}) - \langle \nabla\psi(w'_{t+1}), w_t - w'_{t+1} \rangle \\ &= -\langle \nabla\psi(w_t), u - w_t \rangle + \langle \nabla\psi(w'_{t+1}), u - w'_{t+1} \rangle \\ &\quad - \langle \nabla\psi(w'_{t+1}), w_t - w'_{t+1} \rangle \\ &= -\langle \nabla\psi(w_t), u - w_t \rangle + \langle \nabla\psi(w'_{t+1}), u - w_t \rangle \\ &= \langle \nabla\psi(w'_{t+1}) - \nabla\psi(w_t), u - w_t \rangle \\ &\stackrel{(9)}{=} \langle \eta \nabla f_t(w_t), w_t - u \rangle \end{aligned} \quad (11)$$

Dividing both sides by η , we get the desired result.

(c) From (10), we have $w_{t+1} = \operatorname{argmin}_{w \in \Omega} D_\psi(w, w'_{t+1})$. By definition, $w'_{t+1} \in \Omega \implies D_\psi(w, w'_{t+1})$ attains the global minima in Ω and hence $w_{t+1} = w'_{t+1}$. So,

$$D_\psi(u, w_{t+1}) \leq D_\psi(u, w'_{t+1}), \forall u \quad (12)$$

(d) As $f_t(\cdot)$ is convex,

$$\begin{aligned}
 f_t(w_t) - f_t(u) &\leq \langle \nabla f_t(w_t), w_t - u \rangle \\
 &\stackrel{(11)}{=} \frac{1}{\eta} (D_\psi(u, w_t) - D_\psi(u, w'_{t+1}) + D_\psi(w_t, w'_{t+1})) \\
 &\stackrel{(12)}{\leq} \frac{1}{\eta} (D_\psi(u, w_t) - D_\psi(u, w_{t+1}) + D_\psi(w_t, w'_{t+1}))
 \end{aligned}$$

Adding terms from $t = 1$ to $t = T$, we see that the first two terms on the RHS telescope.

$$\begin{aligned}
 \sum_{t=1}^T (f_t(w_t) - f_t(u)) &\leq \sum_{t=1}^T \frac{1}{\eta} (D_\psi(u, w_t) - D_\psi(u, w_{t+1})) + \frac{1}{\eta} \sum_{t=1}^T D_\psi(w_t, w'_{t+1}) \\
 &= \frac{1}{\eta} (D_\psi(u, w_1) - D_\psi(u, w_{T+1})) + \frac{1}{\eta} \sum_{t=1}^T D_\psi(w_t, w'_{t+1}) \\
 &\leq \frac{D_\psi(u, w_1)}{\eta} + \frac{1}{\eta} \sum_{t=1}^T D_\psi(w_t, w'_{t+1})
 \end{aligned} \tag{13}$$

(e) Let $\Omega = \Delta_N$ (N -dimensional simplex). Let $f_t(w_t) = \langle l_t, w_t \rangle$ where l_t is the loss vector for time t and $\psi(w) = \sum_i w_i \log w_i$. Clearly, $\nabla \psi(w) = [1 + \log w_1, \dots, 1 + \log w_N]^\top$ is invertible. So,

$$\begin{aligned}
 D_\psi(u, w) &= \psi(u) - \psi(w) - \langle \nabla \psi(w), u - w \rangle \\
 &= \sum_i u_i \log u_i - \sum_i w_i \log w_i - \sum_i (1 + \log w_i)(u_i - w_i) \\
 &= \sum_i u_i \log \frac{u_i}{w_i} \\
 &= D(u||w)
 \end{aligned}$$

Let $u = e_j$ such that $j = \operatorname{argmin}_{i \in [N]} \sum_{t=1}^T \langle l_t, e_i \rangle$ where e_i 's are N -dimensional unit vectors. Now, we compute w'_{t+1} . Using (9), we get:

$$1 + \log w'_{t+1,i} = 1 + \log w_{t,i} - \eta l_{t,i} \implies w'_{t+1,i} \propto w_{t,i} e^{-\eta l_{t,i}} \implies w'_{t+1,i} = \frac{w_{t,i} e^{-\eta l_{t,i}}}{\sum_{j=1}^N w_{t,j} e^{-\eta l_{t,j}}}$$

Let $w_1 = \frac{1}{N} \mathbf{1}$. So, (13) becomes:

$$\begin{aligned}
 R_T &= \sum_{t=1}^T \langle l_t, w_t \rangle - \min_{i \in [N]} \sum_{t=1}^T \langle l_t, e_i \rangle \leq \frac{D(u||w_1)}{\eta} + \frac{1}{\eta} \sum_{t=1}^T D(w_t||w'_{t+1}) \\
 &= \frac{\log N}{\eta} + \frac{1}{\eta} \sum_{t=1}^T \sum_{i=1}^N w_{t,i} \log \frac{w_{t,i}}{w'_{t+1,i}} \\
 &= \frac{\log N}{\eta} + \frac{1}{\eta} \sum_{t=1}^T \sum_{i=1}^N w_{t,i} \log \frac{w_{t,i} (\sum_{j=1}^N w_{t,j} e^{-\eta l_{t,j}})}{w_{t,i} e^{-\eta l_{t,i}}}
 \end{aligned}$$

$$\begin{aligned}
&= \frac{\log N}{\eta} + \frac{1}{\eta} \sum_{t=1}^T \eta \langle l_t, w_t \rangle + \log \left(\sum_{j=1}^N w_{t,j} e^{-\eta l_{t,j}} \right) \\
&\stackrel{(a)}{\leq} \frac{\log N}{\eta} + \frac{1}{\eta} \sum_{t=1}^T \eta \langle l_t, w_t \rangle + \log \left(\sum_{j=1}^N w_{t,j} (1 - \eta l_{t,j} + \eta^2 l_{t,j}^2) \right) \\
&= \frac{\log N}{\eta} + \frac{1}{\eta} \sum_{t=1}^T \eta \langle l_t, w_t \rangle + \log (1 - \eta \langle l_t, w_t \rangle + \eta^2 \langle w_t, l_t^2 \rangle) \\
&\stackrel{(b)}{\leq} \frac{\log N}{\eta} + \frac{1}{\eta} \sum_{t=1}^T \eta \langle l_t, w_t \rangle - \eta \langle l_t, w_t \rangle + \eta^2 \langle w_t, l_t^2 \rangle \\
&= \frac{\log N}{\eta} + \eta \sum_{t=1}^T \langle w_t, l_t^2 \rangle \\
&\stackrel{(c)}{\leq} \frac{\log N}{\eta} + \eta T
\end{aligned}$$

The inequality (a) follows because $e^{-x} \leq 1 - x + x^2, x \geq 0$,

(b) follows as $\ln(1+x) \leq x, x \in \mathbb{R}$, and

(c) holds as $\langle w_t, l_t^2 \rangle = \sum_{i=1}^N w_{t,i} l_{t,i}^2 \leq \sum_{i=1}^N w_{t,i} = 1$ as $0 \leq l_{t,i} \leq 1$.

Setting $\eta^* = \sqrt{\frac{\log N}{T}}$, we obtain $R_T \leq \mathcal{O}(\sqrt{T \log N})$, the regret bound for Hedge algorithm.

Problem 4

In this question, we compare the performance of three bandit algorithms namely: ϵ -greedy, UCB1 and Exp3. We have four Bernoulli bandits with expected rewards as: $\mathbf{p} = [0.5, 0.95, 0.2, 0.8]$. We have set $\epsilon = 0.05$ for the ϵ -greedy algorithm and $\gamma = 0.05$ for the Exp3 algorithm (γ is the random exploration probability).

- (a) In this part, we plot the average fraction of times any arm a was selected by algorithm π by some time instant t in semi-log scale in Figure 1, 2 and 3. As $T \rightarrow \infty$, we can expect that probability of choosing the arm with maximum reward approaches 1 and the choosing rest of the arms tends to 0.

Furthermore, from these plots we can observe that for $T \gg 1$, the order of $N_a^\pi(t)$ for arms is according to the expected reward of the arm, i.e., $N_{a_2}^\pi(t) > N_{a_4}^\pi(t) > N_{a_1}^\pi(t) > N_{a_3}^\pi(t)$ as $p(a_2) > p(a_4) > p(a_1) > p(a_3)$ where $p(a)$ denotes the average reward of arm a .

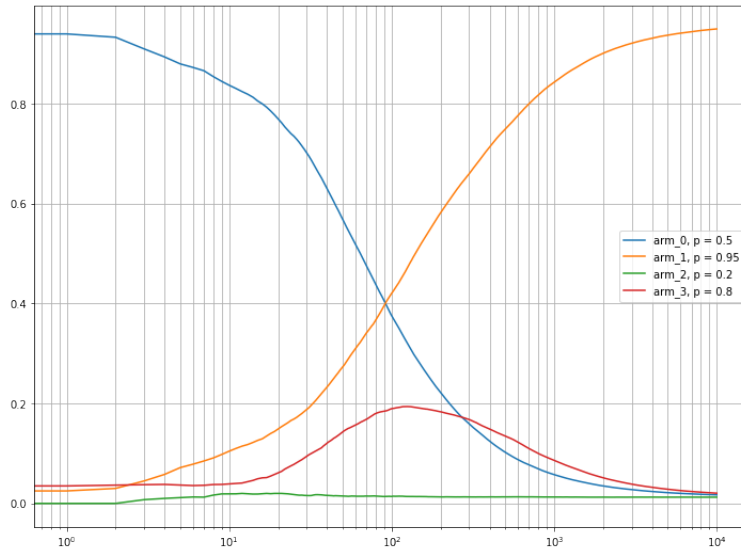


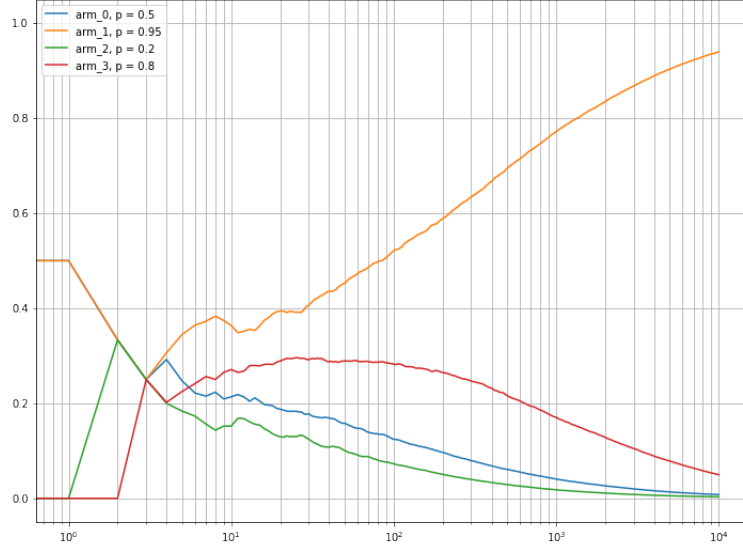
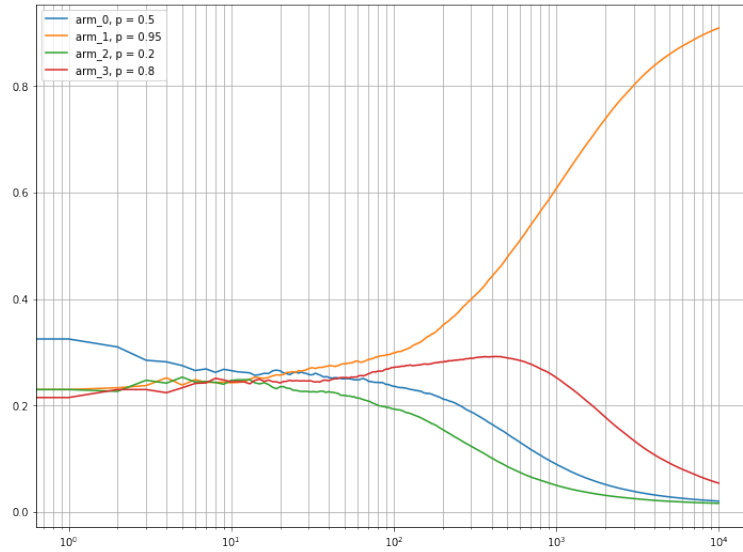
Figure 1: $N_a^\pi(t)$ for ϵ -greedy algorithm

- (b) In this part, we compare the pseudo-regret, $R^\pi(t)$, of each algorithm. Concretely,

$$R^\pi(t) = t \max_a p(a) - \sum_{\tau=1}^t \bar{r}^\pi(\tau) \quad (14)$$

where $\bar{r}^\pi(t)$ is the average reward obtained by algorithm π at time t . The time-evolution of $R^\pi(t)$ is plotted in Figure 4. From the plot we can observe that all the algorithms achieve sublinear regret and UCB1 achieves the best regret bound.

- (c) In this part, we check the sensitivity of the pseudo-regret with hyperparameters of the algorithms.

Figure 2: $N_a^\pi(t)$ for UCB1 algorithmFigure 3: $N_a^\pi(t)$ for Exp3 algorithm

- (i) Variation of ϵ in ϵ -greedy algorithm:

We plot the variation of pseudo-regret with time as ϵ is varied in Figure 5 for five values of $\epsilon = [0.02, 0.05, 0.1, 0.2, 0.5]$. From the plot, we observe that $\epsilon = 0.02$ achieves the best regret bound till $T = 10^4$ steps.

- (ii) Variation of γ in Exp3 algorithm:

We plot the variation of pseudo-regret with time as γ is varied in Figure 6. The best

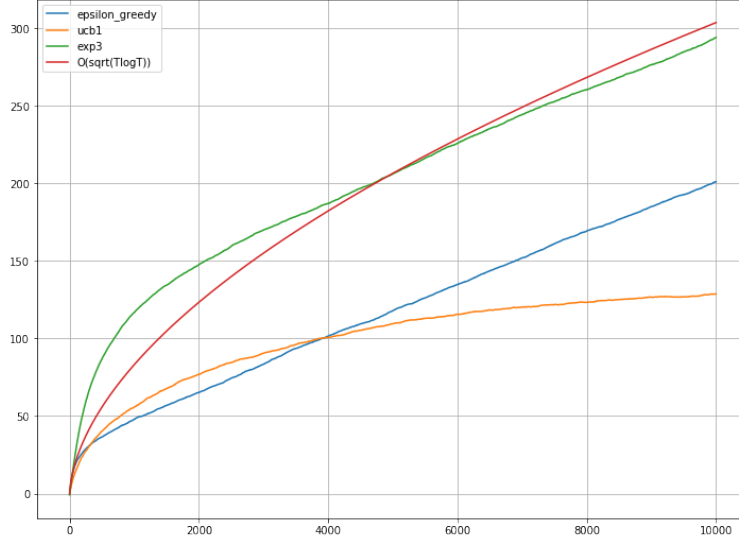


Figure 4: Time-evolution of $R^\pi(t)$ for each algorithm. For comparison, we also plot the theoretical regret bound for most bandit algorithms: $\mathcal{O}(\sqrt{T \log T})$

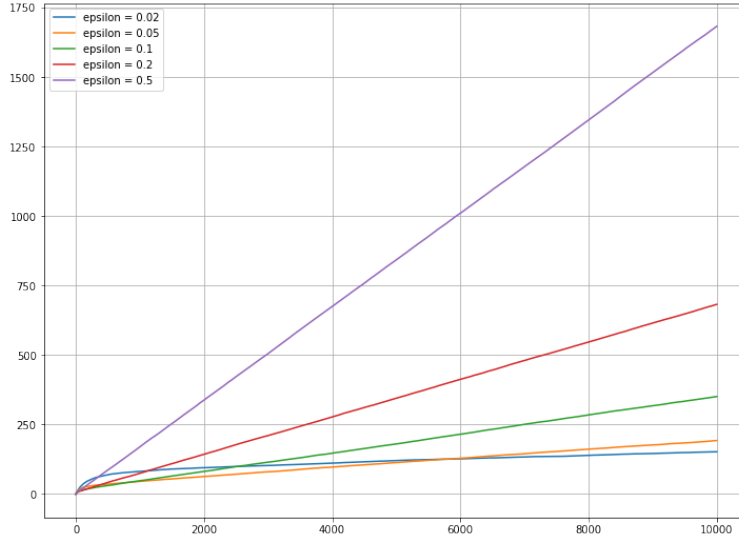


Figure 5: Time-evolution of pseudo-regret as ϵ is varied for ϵ -greedy algorithm.

performing value of γ in Exp3 algorithm is obtained according to¹:

$$\gamma = \min \left\{ 1, \sqrt{\frac{N \log N}{(e-1)G}} \right\} \quad (15)$$

where G is the maximum achievable reward and N is the number of arms. Substituting the values, we obtain $\gamma^* = 0.018$. But, in Figure 6, we observe that $\gamma = 0.05$ attains the

¹<https://www.cs.princeton.edu/courses/archive/fall16/cos402/lectures/402-lec22.pdf>

best regret till time $T = 10^4$. It is to be also noted that the growth rate for $\gamma = 0.02$ is slower than $\gamma = 0.05$. So, as $T \rightarrow \infty$, we can expect $\gamma = 0.02$ achieves the best regret.

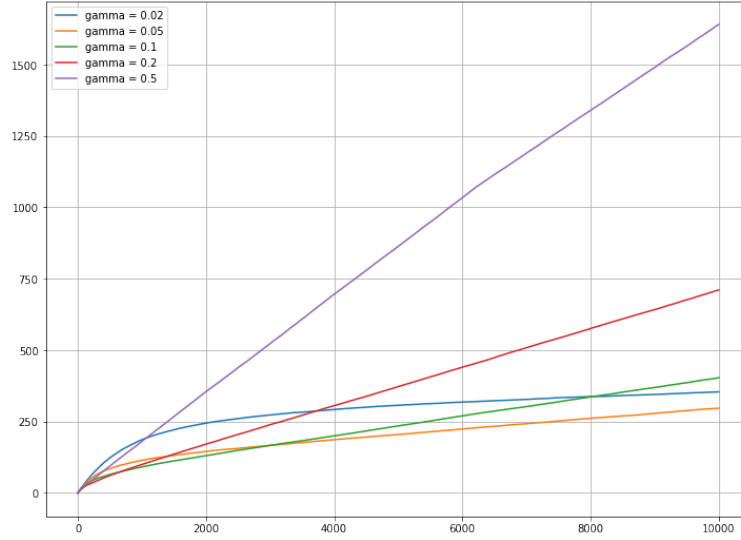


Figure 6: Time-evolution of pseudo-regret as γ is varied for Exp3 algorithm.

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Problem 5

Let $C(a) = \text{cost}(a) = \sum_{t=1}^T c_t(a)$ and $C(a^*) = \min_a \text{cost}(a)$. Furthermore, let $\mathbb{E}[\text{cost}(a)] = Tc(a)$.

- (a) We do an “clean event” analysis by using Hoeffding’s inequality on the cost sequence of each arm. For any arm a , using Hoeffding’s inequality, we have:

$$\begin{aligned} P\left(\left|\frac{1}{T} \sum_{t=1}^T c_t(a) - c(a)\right| \geq \sqrt{\frac{1}{2T} \log \frac{2K}{\delta}}\right) &\leq \frac{\delta}{K} \\ \Rightarrow P\left(\left|\sum_{t=1}^T c_t(a) - \mathbb{E}[\text{cost}(a)]\right| \geq \sqrt{\frac{T}{2} \log \frac{2K}{\delta}}\right) &\leq \frac{\delta}{K} \end{aligned}$$

So, for all the K arms, we get the following tail bound using union bound:

$$P\left(\left|\sum_{t=1}^T c_t(a) - \mathbb{E}[\text{cost}(a)]\right| \geq \sqrt{\frac{T}{2} \log \frac{2K}{\delta}}\right) \leq \delta$$

Choosing $\delta = 2/T^4$, we obtain:

$$P\left(\left|\sum_{t=1}^T c_t(a) - \mathbb{E}[\text{cost}(a)]\right| \leq \sqrt{2T \log KT}\right) \geq 1 - \frac{2}{T^4}$$

So, w.h.p we have:

$$\sum_{t=1}^T c_t(a) \geq \mathbb{E}[\text{cost}(a)] - \sqrt{2T \log KT} \quad (16)$$

$$\sum_{t=1}^T c_t(a^*) \leq \mathbb{E}[\text{cost}(a^*)] + \sqrt{2T \log KT} \quad (17)$$

From (16) and (17), we obtain:

$$\begin{aligned} \mathbb{E}[\text{cost}(a)] &\leq \mathbb{E}[\text{cost}(a^*)] + 2\sqrt{2T \log KT} \\ \Rightarrow \min_a \mathbb{E}[\text{cost}(a)] &\leq \mathbb{E}[\text{cost}(a^*)] + \mathcal{O}(\sqrt{T \log KT}) \end{aligned}$$

- (b) Consider a scenario in which all the arms have 0-1 costs with mean 0.5. In this case,

$$\begin{aligned} \mathbb{E}[\text{cost}(\text{ALG})] &= \mathbb{E}\left[\sum_{t=1}^T c_t(a_t)\right] \\ &= \sum_{t=1}^T \mathbb{E}[c_t(a_t)] = \frac{T}{2} \end{aligned} \quad (18)$$

Now, we consider the random variable $\text{cost}(a)$. It is to be noted that we are dealing the case of deterministic adversary and i.i.d. costs, so, the costs in each step are independent. So,

$$\text{cost}(a) = X_a = X_{a,1} + \dots + X_{a,T}$$

where $X_{a,i}, i \in [T], a \in [K]$ is Bernoulli r.v. with mean 0.5. So, $X_a \sim B(T, 1/2), \forall a \in [K]$. Now, we need to estimate $\mathbb{E}[\min_a \text{cost}(a)] = \mathbb{E}[\min_a X_a]$. For X_a , using Hoeffding's inequality, we have:

$$\begin{aligned} P \left(\left| \frac{1}{T} \sum_{t=1}^T X_{a,t} - \mathbb{E}[X_{a,1}] \right| \geq \sqrt{\frac{1}{2T} \log \frac{2K}{\delta}} \right) &\leq \frac{\delta}{K} \\ \implies P \left(\left| X_a - \frac{T}{2} \right| \geq \sqrt{\frac{T}{2} \log \frac{2K}{\delta}} \right) &\leq \frac{\delta}{K} \end{aligned}$$

So, for all the K arms (random variables), we get the following tail bound using union bound:

$$P \left(\left| X_a - \frac{T}{2} \right| \geq \sqrt{\frac{T}{2} \log \frac{2K}{\delta}} \right) \leq \delta$$

So, with probability $1 - \delta$, we have:

$$X_a \in \left[\frac{T}{2} - \sqrt{\frac{T}{2} \log \frac{2K}{\delta}}, \frac{T}{2} + \sqrt{\frac{T}{2} \log \frac{2K}{\delta}} \right]$$

Let $Y = \min_a X_a$. Clearly, with probability $1 - \delta$, $Y \geq \frac{T}{2} - \sqrt{\frac{T}{2} \log \frac{2K}{\delta}} \implies \mathbb{E}[Y] \geq \frac{T}{2} - \sqrt{\frac{T}{2} \log \frac{2K}{\delta}}$. Now, we can choose δ such that:

$$\frac{T}{2} - \sqrt{\frac{T}{2} \log \frac{2K}{\delta}} \leq \mathbb{E}[Y] \leq \frac{T}{2} - \sqrt{\frac{T}{2} \log \frac{K}{\delta}} \quad (19)$$

As $K \geq 2$, let $\delta = 1/K^{2\gamma-1}$ for some $\gamma \geq 1$. From (19), we get

$$\begin{aligned} \mathbb{E}[\min_a \text{cost}(a)] &= \mathbb{E}[Y] \leq \frac{T}{2} - \sqrt{\frac{T}{2} \log \frac{K}{\delta}} \\ &= \frac{T}{2} - \sqrt{\gamma T \log K} \\ &= \frac{T}{2} - \Omega(\sqrt{T \log K}) \\ &\stackrel{(18)}{=} \mathbb{E}[\text{cost}(\text{ALG})] - \Omega(\sqrt{T \log K}) \\ \implies \mathbb{E}[\text{cost}(\text{ALG}) - \min_a \text{cost}(a)] &\geq \Omega(\sqrt{T \log K}) \end{aligned}$$

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