Chapter 3: Functional Principal components. (FPC)

- · FPC's allow us to reduce the dimension of oo-dimensional functional data to a small (p) finite dimension in an optimal way.
 - a maximizing variabilty
 - @ optimal orthonormal basis

§ 3.1 A maximization problem

Thm 3.1 (Principal axis theorem).

Suppose A is a sym. pxp matrix. Then there is an orthonormal matrix LI= [u1 up] whose columns are the eigenvectors of A. so, UTU = I and A. Uj = Aj. Uj

1/2

> UTAU = 1 = dtag (li - lp)

Thm 3.1 Implies that A = U.A.UT

Question:

Suppose A be symm. and pd. with 17 /2.... > AP.

Find X St XAX is maximum and IIXII = 1.

XAX = XTUNUTX = YT. N.Y where Y = UT.X > X = U.Y

Since || y || = || x || = 1, find unit length vector y first and x = U-y. $y^{T} \wedge y = \sum_{j=1}^{P} \chi_{j} \cdot y_{j}^{2} \Rightarrow y = [1,0...0]^{T} \Rightarrow x = W_{1}$

Extend this idea to a separable Hilbert space.

Suppose 4 is a sym. pd Hilbert-Schmidt operator in L.

RECALL:

$$\langle \psi(x), y \rangle = \langle x, \psi(m) \rangle$$
, $x, y \in \mathcal{L}^{2}$
 $\langle \psi(x), x \rangle > 0$, $x \in \mathcal{L}^{2}$
 $\psi(x) = \sum_{j=1}^{\infty} \lambda_{j} \langle x, v_{j} \rangle v_{j}$, $x \in \mathcal{L}^{2}$ (2.4).

where orthonormal vj and $\psi(v_j) = \lambda_j \cdot v_j$

by (2.4) $\Rightarrow \psi(x) = \overline{\sum_{j=1}^{\infty}} \lambda_j \langle x, v_j \rangle v_j \quad , \quad x \in \mathcal{L}.$ Then,

maximize <ψ(x), x> subj. to ||x||=1

$$\Lambda. \left\langle \psi(x), x \right\rangle = \left\langle \sum_{j=1}^{\infty} \lambda_j \langle x, v_j \rangle v_j, x \right\rangle \\
= \sum_{j=1}^{\infty} \lambda_j \left\langle \langle x, v_j \rangle v_j, x \right\rangle \\
= \sum_{j=1}^{\infty} \lambda_j \langle x, v_j \rangle^2$$

by Parseval's equality b. $\| \| \| \| \|^2 = \sum_{j=1}^{\infty} | \langle \chi, V_j \rangle |^2 = 1$

 \Leftrightarrow maximize $\sum_{j=1}^{\infty} \lambda_j \langle x, v_j \rangle^2$ subj. to $\sum_{j=1}^{\infty} |\langle x, v_j \rangle|^2 = 1$

Suppose 11 > 127 ... 7 ... So, take $\langle X, V_1 \rangle = 1$ and $\langle X, V_j \rangle = 0$ $\exists V_j > 1$ \Rightarrow max $\sum_{j=1}^{\infty} \lambda_j \langle x, V_j \rangle^2 = \lambda_1$ and $x = V_1$ (or $-V_2$) Next, maximize $\langle \Psi(\alpha), \alpha \rangle$ Subj. to $\|x\| = 1$, $\langle \alpha, V_1 \rangle = 0$ then, max $\sum_{j=2}^{\infty} \lambda_j \langle \alpha, V_j \rangle^2$ & $\sum_{j=2}^{\infty} |\langle \alpha, V_j \rangle|^2 = 1$. $\Rightarrow \alpha = V_2$ & maximum value = λ_2 .

Thm 3.2

Suppose Ψ is a symm. pd Hilbert - Schmidt operator with eigenfunctions U_j and eigenvalues λ_j satisfying $\lambda_1 > \lambda_2 > \cdots > \lambda_{p+1}$. Then,

Sup { $\langle \psi(x), x \rangle$: ||x|| = 1, $\langle x, v_5 \rangle = 0$, $|\leq j \leq i-1$, i $and Supremum is reached if <math>x = v_i$

§ 3.2 Optimal Empirical Orthogormal bists

Approach in § 3.1 can be applied to the problem to find an orthonormal basis U1. Up 9, ±

S= ZN | xi - Zk=1 < xi, UK>UK ||2 is minimized.

- · XI ... XN : observed realizations of random functions in L
- · P < N (fixed)
- > Xi~ ∑k=1 < xi, Uk>Uk

U; optimal empirical orthonormal basis.
or natural orthonormal components

The functions U1. - Up minimizing \$2 are equal to the normalized Eigenfunctions of the sample covariance operator C.

proof

Suppose P=1. Find u with ||u||=1.

$$= \sum_{i=1}^{N} ||x_{i}||^{2} - 2\sum_{i=1}^{N} \langle x_{i}, u \rangle^{2} + \sum_{i=1}^{N} \langle x_{i}, u \rangle^{2} ||u||^{2}$$

$$= \sum ||\chi_{\bar{1}}||^2 - \sum \langle \chi_{\bar{1}}, \mu \rangle^2$$

RECALL:

$$\hat{C}(u) = \frac{1}{N} \sum_{i=1}^{N} \langle \chi_i, u \rangle u$$

$$\Rightarrow \langle \hat{C}U, u \rangle = \langle \frac{1}{N} \Sigma_{i=1}^{N} \langle \chi_{i}, u \rangle \chi_{i}, u \rangle$$

$$= \frac{1}{N} \Sigma \langle \chi_{i}, u \rangle^{2}$$

$$\max \langle \partial u, u \rangle = \lambda_1$$
 and $u = U_1$ by Thm 3.2

Generally,

Generally,

min
$$\hat{S}^2 = \sum_{i=1}^N ||\chi_i||^2 - \sum_{i=1}^N \sum_{k=1}^P \langle \chi_i, u_k \rangle^2$$
.

$$\Leftrightarrow$$
 max $\sum_{i=1}^{N} \sum_{k=1}^{P} \langle x_i, u_k \rangle^2$

$$= \max_{k=1}^{\infty} \sum_{k=1}^{\infty} \left\langle \sum_{j=1}^{\infty} \hat{\lambda}_{j} \left\langle u_{k}, \hat{v}_{j} \right\rangle \hat{v}_{j} \right\rangle , \quad u_{k} \rangle$$

=
$$\max \sum_{k=1}^{P} \sum_{j=1}^{\infty} \hat{\lambda}_{j} < u_{k}, \hat{v}_{j} >^{2} \leq \sum_{k=1}^{P} \hat{\lambda}_{k}$$
 by ± 100 3.2

and maximum is attained if up-vi - up=vp

Section 5-3, we identify the funct-princ. score Section 3.3 Functional Principal Components with the eigenfunction of the covariance open of Suppose X1, X2, ..., XN are functional observations. and show how its eigent decompose the variance the functional data. EFPC -> Empirical Functional Principal Components (The eigenfunctions of the sample covariance operator.) Alf these observations have the same have the same distribution as a square integrable Le valued random function X, we define the FPCs as the eigenfunctions of the covariance operator. Recall: If X is square integrable, EllXII2 = E X2(+)dt < 00 and E(X)=0, then C(y) = E(X,y>,X], y e12. M(t)=E(X(t)) (mean fraction) c(tis) = E[(x(t)-m(t)).(x(s)-m(s))] (covariance function) C = E[(X-N), : Y(X-N)] (covariance operator) $\hat{\mu}(t) = \frac{1}{N} \cdot \sum_{i=1}^{N} X_i(t)$ (sample mean function) ê(+,s) = 1 . \frac{1}{N} (\times_{i=1}^{N} (\times_{i}(+) - \hat{\rho}(+)) (\times_{i}(+)) (\t

 $\hat{C}(x) = \frac{1}{N} \cdot \sum_{i=1}^{N} \langle X_i - \hat{n}, x \rangle (X_i - \hat{n}), x \in L^2 \quad (sample cov. operator)$

* Parseval's Eq. =) \(\frac{12}{51} \times \times

Let N; , \(\j, \j \) be the eigenfunctions and eigenvalues of the covariance operator C. The relation $CNj = \lambda j \cdot Nj$ implies that

 \rightarrow

$$\lambda_{j} = \langle C_{N_{j}}, v_{j} \rangle = \langle E[\langle X_{j}, v_{j} \rangle \times], v_{j}]$$

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$$= \int \int E[x(t)x(t)] v_j(t) v_j(t) dt ds$$

$$= E \left[\int (x(t)x(s)v_j(t)v_j(s)dtds \right]$$

Section 2.3 pg.23

If
$$\psi \in I$$
 and X is integrable, then $E[\psi(x)] = \psi[E(x)] = \psi[E(x)] = E[(X, N_j)^2]$

All Section 3.2 explains that the EFPCs can be interpreted as an optimal orthonormal basis.

Recall: Ĉ(x) = 1 . \(\frac{1}{N} \cdot \frac{1}{124} \left(\times_1 - \hat{m}, \times \right) , \times \left(\times_1 - \hat{m} \right) , \times \left(

$$\sim$$

$$\begin{split} \angle \hat{C}x, x \rangle & \text{ or } \langle \hat{C}(x), x \rangle \\ \angle \hat{C}x, x \rangle &= \langle \frac{1}{N} \cdot \sum_{i=1}^{N} \langle X_i - \hat{p}_i, x \rangle \langle X_i - \hat{p}_i), x \rangle \\ + \text{Assume } E(X_i) &= 0 \\ \langle \hat{C}x, x \rangle &= \langle \frac{1}{N} \cdot \sum_{i=1}^{N} \langle X_i, x \rangle X_i, x \rangle \\ &= \frac{1}{N} \langle \sum_{i=1}^{N} \langle X_i(t)x(t)dt | X_i(s), x \rangle \\ &= \frac{1}{N} \cdot \sum_{i=1}^{N} \langle X_i(t)x_i(s)|x(t)dt \rangle \\ &= \frac{1}{N} \cdot \sum_{i=1}^{N} \langle X_i(t)x(t)dt \rangle \\ &= \frac{1}{N} \cdot \sum_{i=1}^{N} \langle X_i(t)x(t)dt \rangle^2 \\ &= \frac{1}{N} \cdot \sum_{i=1}^{N} \langle X_i, x \rangle^2 = \langle \hat{C}x_i, x \rangle \\ &\text{Hence, } \frac{1}{N} \cdot \sum_{i=1}^{N} \langle X_i, x \rangle^2 = \langle \hat{C}x_i, x \rangle \\ &\text{Sumple variance of the Jata} \\ &\text{in the direction } x. \\ &\text{of the function} \end{split}$$

We must find x by maximizing $\langle \hat{C}x, x \rangle$ subject to $||x|| = 1 \Rightarrow By \text{ Th. 3.2}$, $x = \hat{v}_1$. The first EFPC.

Al Observe that since Nj, j=1,..., N form a basis in PN,

$$\frac{1}{N} \sum_{i=1}^{N} ||X_{i}||^{2} = \frac{1}{N} \sum_{i=1}^{N} \langle X_{i}, \hat{n}_{j} \rangle^{2}$$

$$= \sum_{j=1}^{N} \frac{1}{N} \sum_{i=1}^{N} \langle X_{i}, \hat{n}_{j} \rangle^{2}$$

$$= \frac{1}{N} \sum_{j=1}^{N} (\hat{C}\hat{n}_{j}, \hat{n}_{j})$$

$$= \frac{1}{N} \sum_{j=1}^{N} (\hat{C}\hat{n}_{j}, \hat{n}_{j})$$

$$= \frac{1}{N} \sum_{j=1}^{N} \hat{\lambda}_{j}$$

RESULT: Variance in the direction \hat{v}_j is $\hat{\lambda}_j$. $\hat{\mathcal{N}}_j$ explains the fraction of the total sample variance equal to $\frac{\hat{\lambda}_j}{\sum_{k=1}^N \hat{\lambda}_k}$

$$X_n(t) = \sum_{j=1}^{p} a_j \cdot Z_{jn} \cdot e_j(t)$$

95: real numbers for every n

eg (t) orth functions with llej 11=

A Denote by X a random function with the same distribution as Xn(+).

$$X(t) = \sum_{j=1}^{p} a_j \cdot Z_j \cdot e_j(t)$$

By definition;

$$C(x)t = E\left[\left(\int X(s) \cdot x(s)ds\right)X(t)\right]$$

$$= E\left[\int X(t)X(s)x(s)ds\right] = \int E\left[X(t)X(s)\right]x(s)ds$$

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$$C(t,s) = E\left[X(t)X(s)\right] = E\left[\sum_{j=1}^{p} a_{j} \cdot Z_{j} \cdot e_{j}(t) \sum_{i=1}^{p} a_{i} \cdot Z_{i} \cdot e_{j}(s)\right]$$

$$= E\left[\sum_{j=1}^{9} \alpha_j^2 \cdot Z_j^2 \cdot e_j(t)e_j(s)\right]$$

$$= \sum_{j=1}^{p} a_{j}^{2} \cdot E(2j^{2}) \cdot e_{j}(t) e_{j}(s) = \sum_{j=1}^{p} a_{j}^{2} e_{j}(t) e_{j}(s)$$

$$\lambda_j = \alpha_j^2$$
, $N_j = e_j$