

Towards Causal Inference for SpatioTemporal Data

Conflict and Forest Loss in Colombia

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Context for this work

The Colombian Conflict

- Armed conflict for 50 years
- More than 20,000 fatalities
- Peace agreement in 2016
- Increased forest loss attributed to armed conflict
- Pressure on forest reduced due to armed conflict preventing logging
- A data driven approach ->



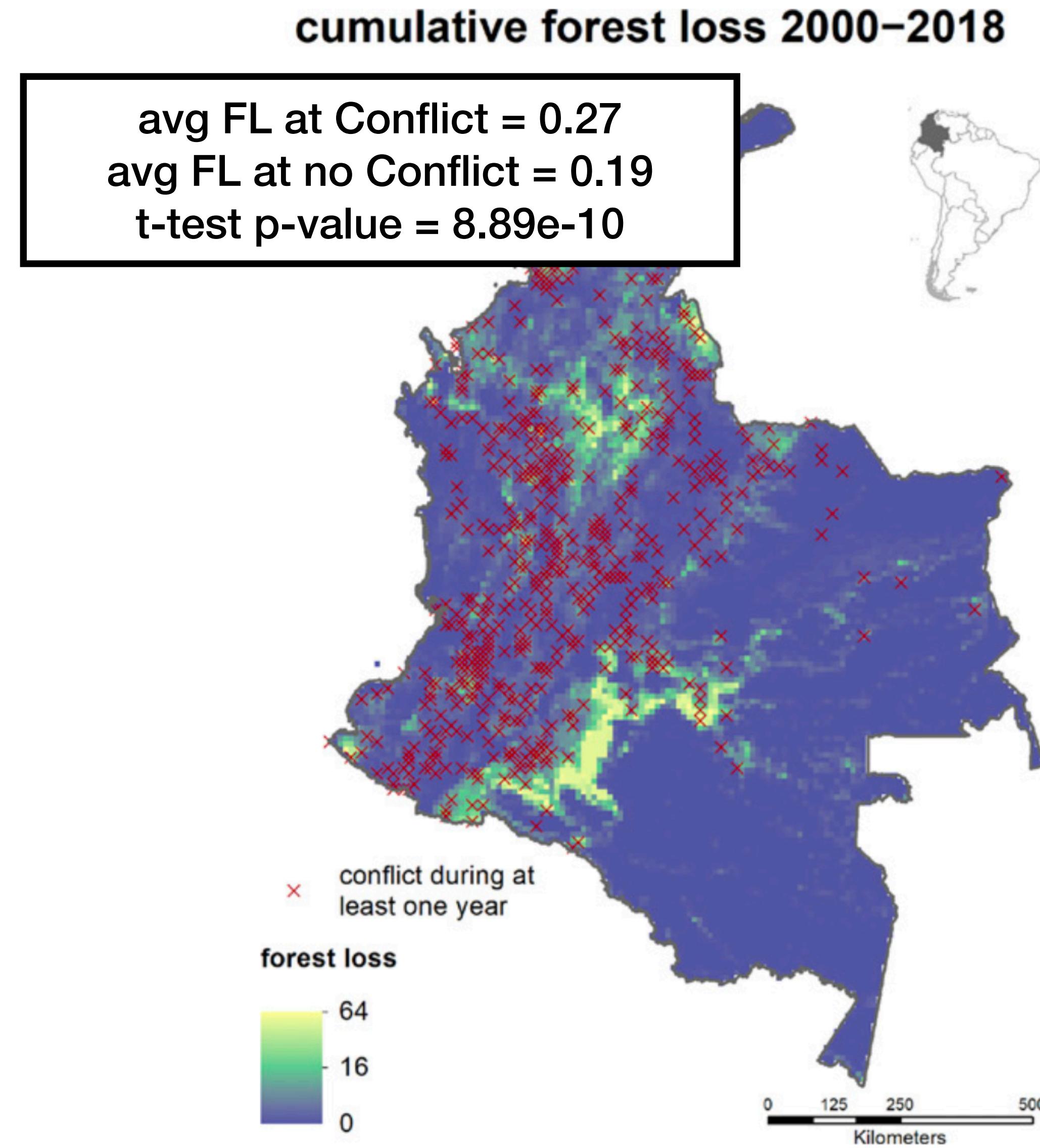
How does conflict affect forest loss?

Data Sources

- Forest loss dataset (2000-2018) (Source: Landsat Satellite Imagery, Resolution 30m x 30m), Complete canopy removal)
- Conflict event dataset (Source: Georeference Events Dataset (GED) from Uppsala Conflict Data Program (UCPD), Armed force by an organised actor resulting in at least one death)
- Accessibility proxy in terms of road distance (Source: <https://diva-gis.org>)
- Additional data sources for population, industrial presence etc (Source: unavailable)

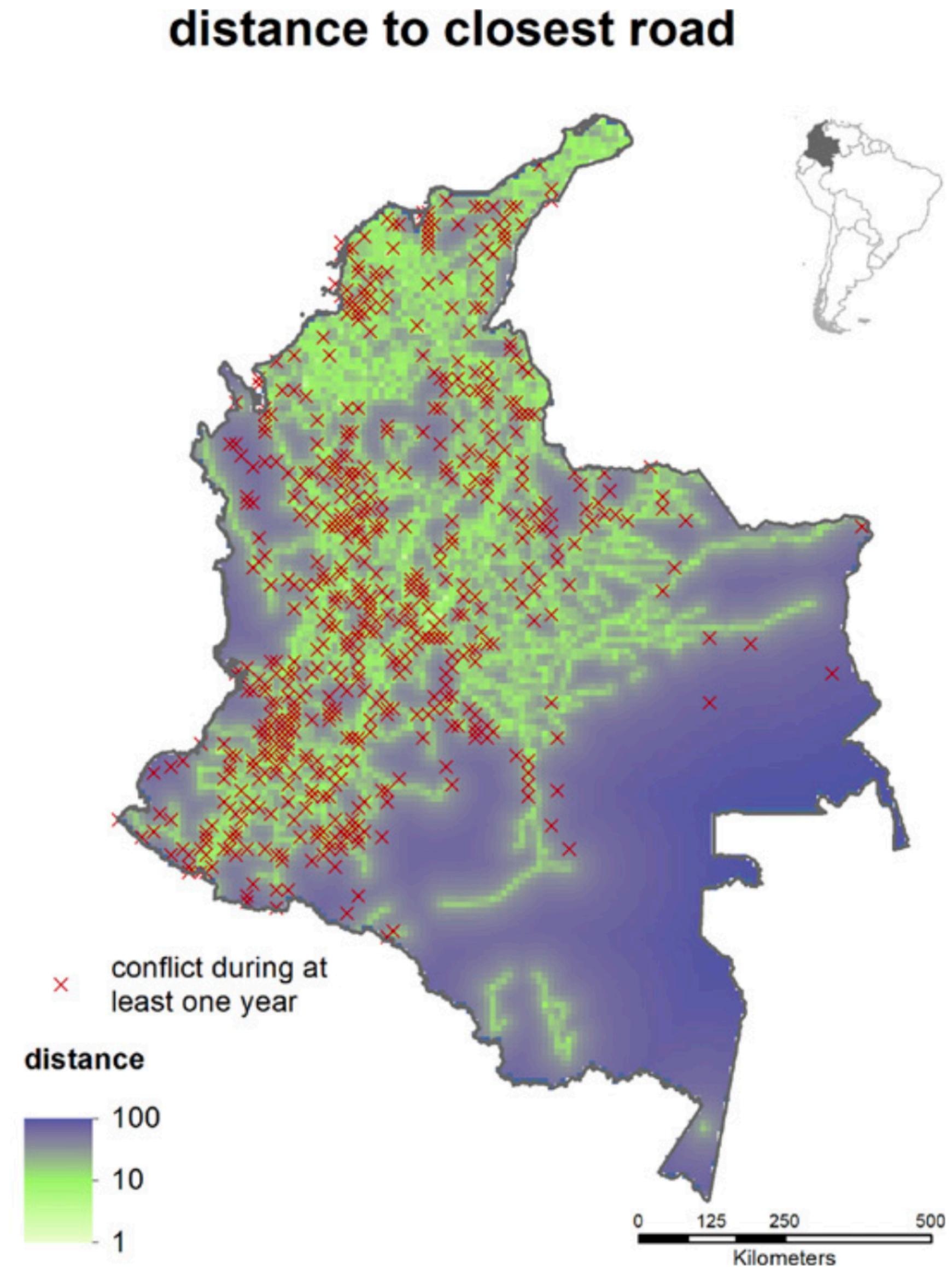
Summary Statistics

- Conflict events aggregated by Counting all events per year in each grids
- Forest loss aggregated by averaging annual loss (and spatial resolution adjusted to 10km x 10km)
- X_s^t : Binary conflict indicator at location s at time t
- Y_s^t : Absolute forest loss in location s from year $t - 1$ to t
- Average forest loss in areas of conflict **exceeds** non conflict by almost 50%



The Fault in our Analysis

- Both conflict and forest loss happen in areas of high accessibility
- W_s^t : Distance from s to the closest road in km
- Other confounders - (population density, market infrastructure etc...)
- Exaggerated significance of t-test due to strong spatial dependencies in X and Y



A method suitable for
Spatiotemporal Data

Causal Model for Spatiotemporal Process

Preliminaries

- Every dataset is considered to be a realisation of a spatiotemporal process
- \mathbf{Z} is a p-dimensional spatiotemporal process taking values in \mathcal{Z}_P
- Z_s^t marginalising \mathbf{Z} at s and t
- $\mathbf{Z}_s : (Z_s^t)_{t \in \mathbb{N}}$ for time series and $\mathbf{Z}^t : (Z_s^t)_{s \in \mathbb{R}^2}$ for spatial process
- *Weakly stationary* if marginal distribution for Z_s^t is same for all s and t
- *Time Invariant* if $\mathbb{P}(Z^1 = Z^2 = \dots) = 1$

Causal Model for Spatiotemporal Process

Causal Graphical Models for Spatiotemporal Processes

- Multivariate Spatiotemporal Processes for joint modelling of phenomenon for eg. Accessibility, population density, market infrastructure, forest loss, conflict
- Modelling causal relations among “disjoint bundles”
- $Z = (\mathcal{S}, \mathcal{G}, \mathcal{P})$
- $\mathcal{S} = (S_j)_{j=1}^k$ for non empty disjoint sets $S_1, \dots, S_k \subseteq \{1\dots p\}$ and $\cup_{j=1}^k S_j = \{1\dots p\}$
- A Direct Acyclic Graph \mathcal{G} with vertices S_1, \dots, S_k
- Where $\mathcal{P} = \{\mathcal{P}^j\}_{j=1}^k$ is $\mathcal{P}^j = \{\mathbb{P}_z^j\}_{z \in Z_{|PA_j|}}$

Causal Model for Spatiotemporal Process Observations and Interventions

- $\mathbb{P}(F) = \int_{F_1} \dots \int_{F_k} \mathbb{P}_{\mathcal{Z}^{PA_k}}^k(dz^{(S_k)}) \dots \mathbb{P}^1(dz^{(S_1)})$ is the observational distribution
- The conditional distribution of $\mathbf{Z}^{(S_j)}$ given $\mathbf{Z}^{(PA_j)}$ as induced by \mathbb{P} is \mathcal{P}_j
- Can therefore be written as $[\mathbf{Z}^{(S_k)} | \mathbf{Z}^{(PA_k)}] \dots [\mathbf{Z}^{(S_1)}]$
- Intervention is defined as replacing \mathcal{P}_j by $\tilde{\mathcal{P}}_j$
- The new graphical model is therefore $(\mathcal{S}, \mathcal{G}, \tilde{\mathcal{P}})$

Latent Spatial Confounder Model

Definition of an LSCM

- $(\mathbf{X}, \mathbf{Y}, \mathbf{H}) = (X_s^t, Y_s^t, H_s^t)_{(s,t) \in \mathbb{R}^2 \times \mathbb{N}}$ where $X_s^t \in \mathbb{R}^d$, $H_s^t \in \mathbb{R}^l$ and real valued Y_s^t
- Causal Structure $[\mathbf{Y} \mid \mathbf{X}, \mathbf{H}] [\mathbf{X} \mid \mathbf{H}] [\mathbf{H}]$
- Assumptions
 - Latent process \mathbf{H} is weakly stationary and time invariant
 - IID sequence $\epsilon^1, \epsilon^2, \dots$ of weakly stationary spatial error processes and measurable function $f: \mathbb{R}^{d+l+1} \rightarrow \mathbb{R}$ st $Y_s^t = f(X_s^t, H_s^t, \epsilon_s^t)$

Latent Spatial Confounder Model

Average Treatment Effect and Causal Interpretation

- Average effect $f_{AVE(X \rightarrow Y)}(x) := \mathbb{E}[f(x, H_0^1, \epsilon_0^1)]$ expectation over both noise and latent variables
- For fixed x and s, t if an intervention is applied s.t. $X_s^t = x$ holds almost surely in the interventional distribution \mathbb{P}_x then $\mathbb{E}_{\mathbb{P}_x}[Y_s^t] = f_{AVE(X \rightarrow Y)}(x)$
- When graph is known we can compute interventional distribution from observational distribution
- $f_{AVE(X \rightarrow Y)}(x) := \mathbb{E}[f_{Y|(X,H)}(x, H_0^1)]$ where $f_{Y|(X,H)}$ is the regression function $(x, h) \rightarrow \mathbb{E}[Y_s^t | X_s^t = x, H_s^t = h]$

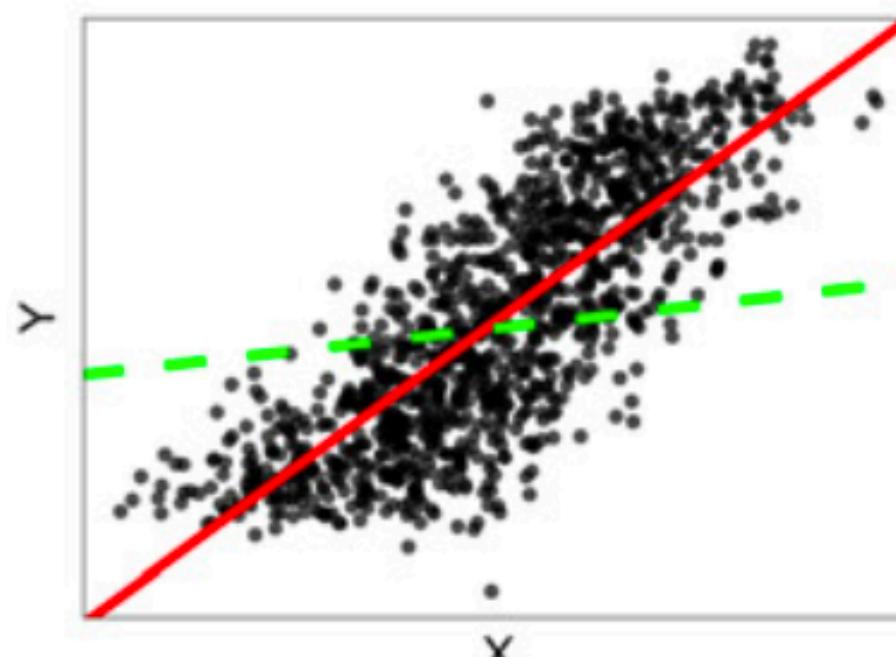
Estimating Average Causal Effect

- We have dataset $(\mathbf{X}_n^m, \mathbf{Y}_n^m) = (X_s^t, Y_s^t)_{(s,t) \in \{s_1, \dots, s_n\} \times \{1, \dots, m\}}$
- For every $s \in \{s_1, \dots, s_n\}$ several time instances with $t \in \{1, \dots, m\}$ with the same conditional $Y_s^t | (X_s^t, H_s^t)$
- The latent realisation h_s of H_s^1 is not observed but since \mathbf{H} is static for every s we can estimate $f_{Y|(\mathbf{X}, \mathbf{H})}(\cdot, h_s)$
- We need to specify a model class for $f_{Y|(\mathbf{X}, \mathbf{H})}(\cdot, h)$ and a suitable estimator
 $\hat{f}_{Y|\mathbf{X}} = (\hat{f}_{Y|\mathbf{X}}^m)_{m \in \mathbb{N}}$
- $\hat{f}_{AVE(X \rightarrow Y)}^{nm}(\mathbf{X}_n^m, \mathbf{Y}_n^m)(x) := \frac{1}{n} \sum_{i=1}^n \hat{f}_{Y|\mathbf{X}}^m(\mathbf{X}_{s_i}^m, \mathbf{Y}_{s_i}^m)(x)$

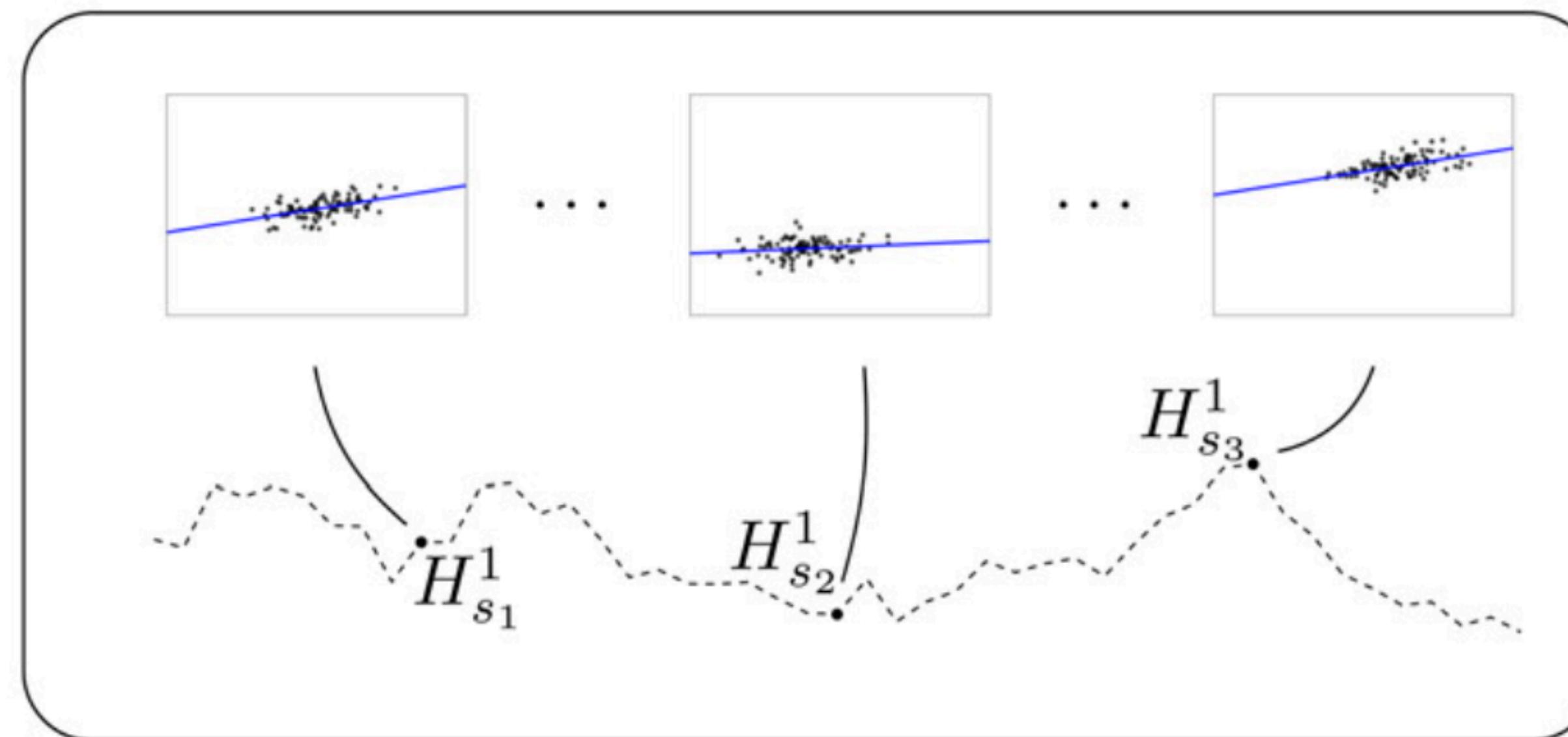
Estimating Average Causal Effect from Data

Averaging localised models

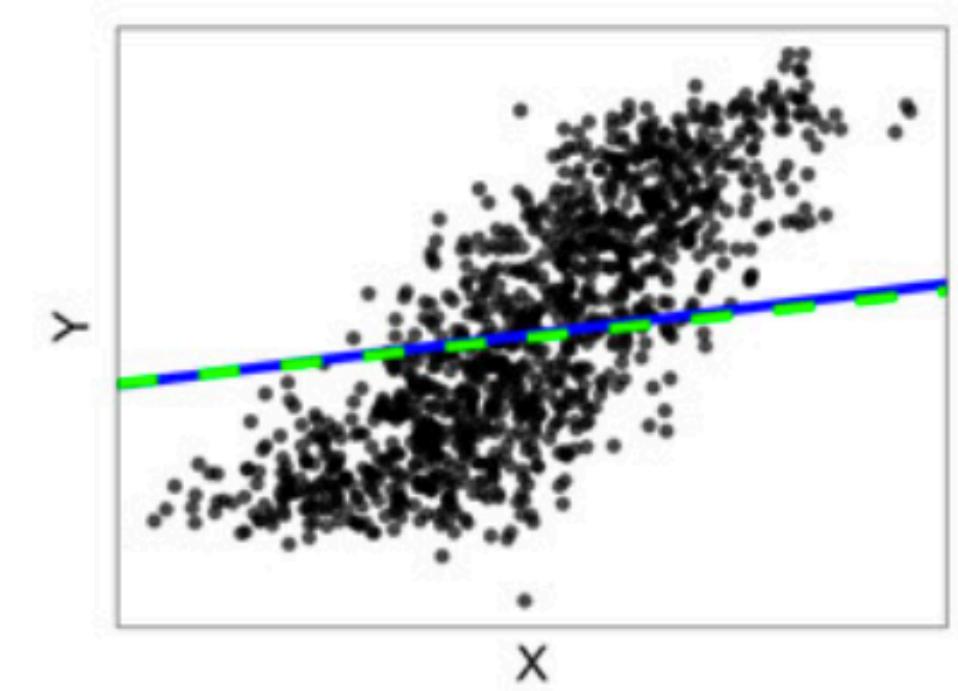
confounded effect



spatially disaggr.



avg causal effect



average

for all s , produce estimate $\hat{f}_{Y|X}^m(\mathbf{X}_s^m, \mathbf{Y}_s^m)(\cdot)$

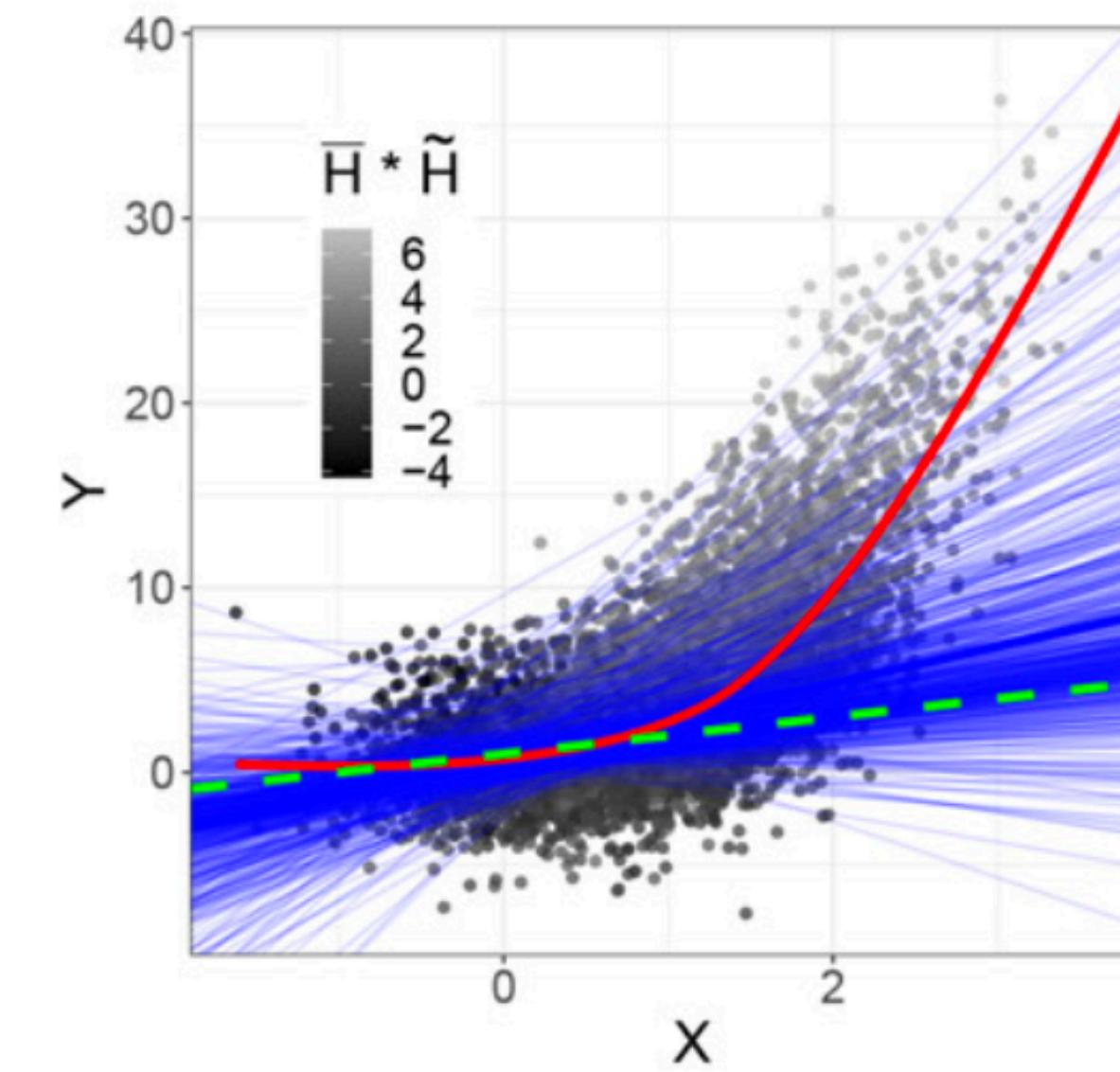
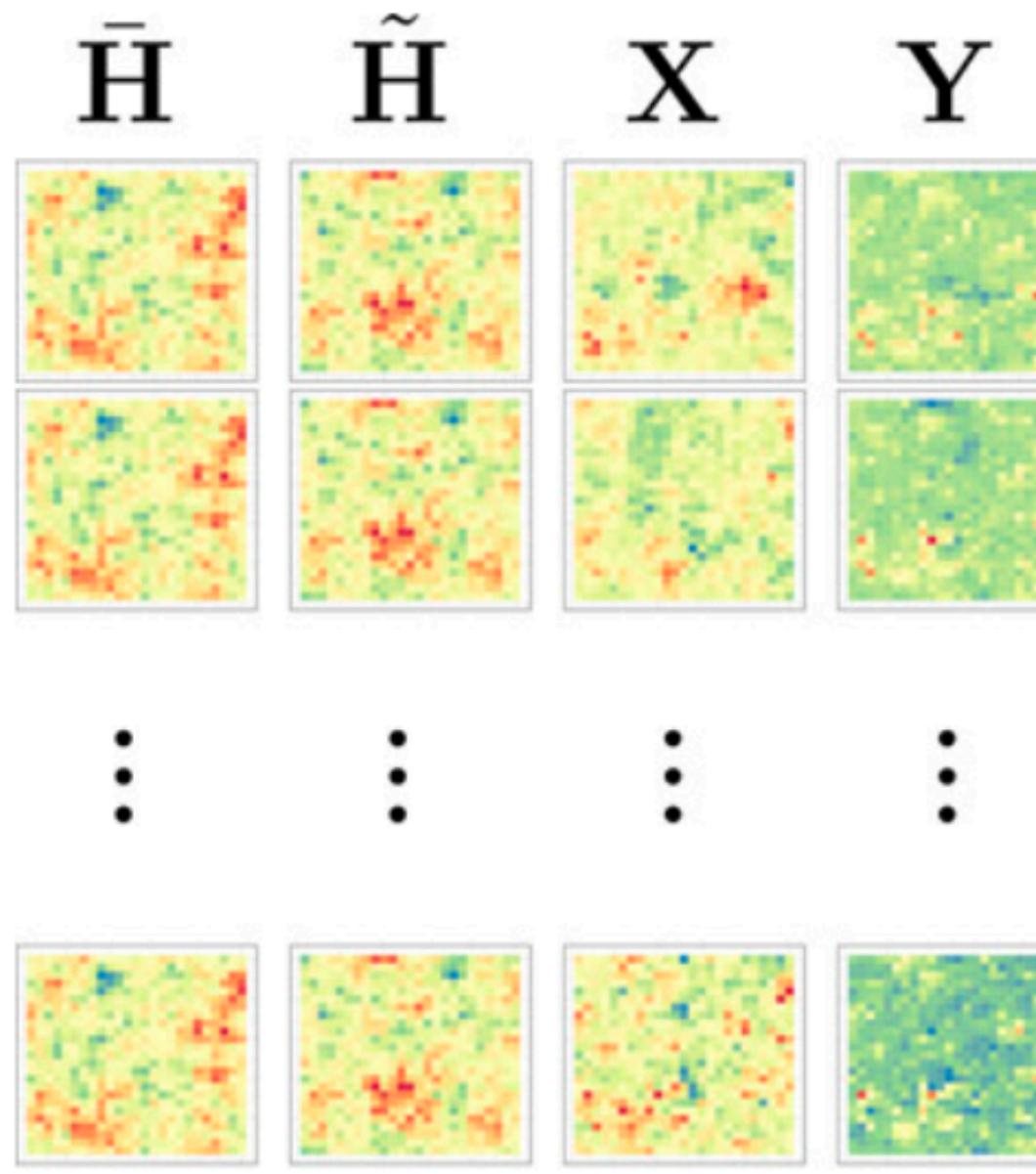
Estimating Average Causal Effect from Data Assumptions

- Law of Large Numbers for LSCM : With increasing number of spatial locations
$$\hat{f}_{AVE(X \rightarrow Y)}^{nm}(\mathbf{X}_n^m, \mathbf{Y}_n^m)(x) := \frac{1}{n} \sum_{i=1}^n \hat{f}_{Y|X}^m(\mathbf{X}_{s_i}^m, \mathbf{Y}_{s_i}^m)(x)$$
 approximates
$$f_{AVE(X \rightarrow Y)}(x) := \mathbb{E}[f_{Y|(X,H)}(x, H_0^1)]$$
 for a stationary Gaussian process H^1 sampled regularly in space
- Consistent estimators of the conditional expectations
$$\hat{f}_{Y|X}^m(\mathbf{X}_s^m, \mathbf{Y}_s^m)(x) - f_{Y|(X,H)}(x, H_s^1) \rightarrow 0 \text{ as } m \rightarrow \infty$$
- there exists $N \in \mathbb{N}$ s.t. for all $n \geq N$ we can find $M_n \in \mathbb{N}$ s.t. for all $m \geq M_n$
$$\mathbb{P}\left(|\hat{f}_{AVE(X \rightarrow Y)}^{nm}(\mathbf{X}_n^m, \mathbf{Y}_n^m)(x) - f_{AVE(X \rightarrow Y)}(x)| > \delta\right) \leq \alpha$$

Asymptotic consistency: An Example LSCM

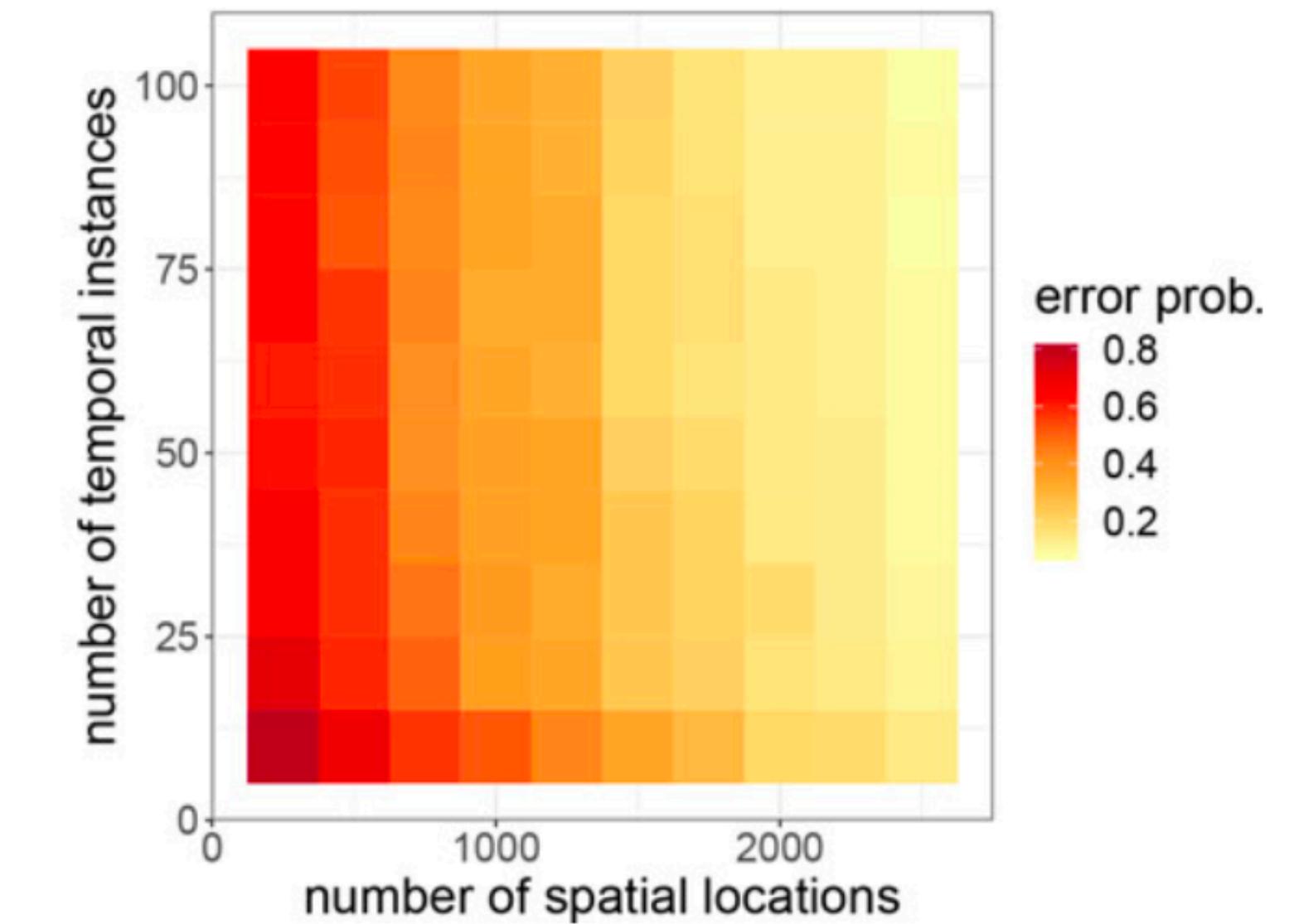
$$\xi, \psi, \xi^t, \varepsilon^t, t \in \mathbb{N}$$

Are independent versions of univariate Gaussian Processes with mean 0 and covariance $u \rightarrow \exp(-\frac{1}{2} \|u\|_2)$



$$\begin{aligned}
 H_s^t &= (\bar{H}_s^t, \tilde{H}_s^t) = (\zeta_s, 1 + \frac{1}{2}\zeta_s + \frac{\sqrt{3}}{2}\psi_s), \\
 X_s^t &= \exp(-\|s\|_2^2/1000) + (0.2 + 0.1 \cdot \sin(2\pi t/100)) \\
 &\quad \cdot \bar{H}_s^t \cdot \tilde{H}_s^t + 0.5 \cdot \xi_s^t, \\
 Y_s^t &= (1.5 + \bar{H}_s^t \cdot \tilde{H}_s^t) \cdot X_s^t + (\bar{H}_s^t)^2 + |\tilde{H}_s^t| \cdot \varepsilon_s^t.
 \end{aligned}$$

$$Y_s^t = f_1(H_s^t) \cdot X_s^t + f_2(H_s^t, \varepsilon_s^t)$$



Testing for Existence of Causal Effects

Exchangeability and Resampling

- $H_0 : (\mathbf{X}, \mathbf{Y})$ is generated from an LSCM with a function f constant wrt X_s^t
- Formalising “No causal effect of X on Y ” within LSCMs
- For a resampling test we use a permutation scheme s.t. for the null hypothesis the distribution of data remains unaffected
- for every $(\mathbf{x}, \mathbf{y}) \in \mathbb{R}^{(d+1) \times n \times m}$ and permutation σ of elements $\{1, \dots, m\}$
 $\sigma(\mathbf{x}, \mathbf{y})$ is the permuted array with entries $(\sigma(x, y))_s^t = (x_s^t, y_s^{\sigma(t)})$
- Exchangeability is under H_0 $\sigma(\mathbf{X}_n^m, \mathbf{Y}_n^m)$ has the same distribution as $\mathbf{X}_n^m, \mathbf{Y}_n^m$

How to perform a test

- For $B \in \mathbb{N}$ uniform draws k_1, \dots, k_b from $\{1, \dots, M\}$ where $M := m!$
- $p_{\hat{T}}(\mathbf{x}, \mathbf{y}) := \frac{1 + |\{b \in \{1, \dots, B\} : \hat{T}(\sigma_{k_b}(\mathbf{x}, \mathbf{y})) \geq \hat{T}(\mathbf{x}, \mathbf{y})\}|}{1 + B}$
- $\phi_{\hat{T}}^\alpha = 1 \Leftrightarrow p_{\hat{T}} \leq \alpha$
- $\hat{T}(\mathbf{X}_n^m, \mathbf{Y}_n^m) = \psi(\hat{f}_{AVE}^{nm}(\mathbf{X}_n^m, \mathbf{Y}_n^m))$
- What is a block permutation scheme and why we need it

Applying the model to Data

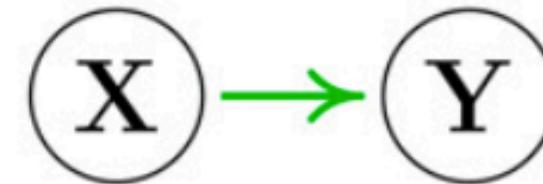
- $T := f_{AVE(X \rightarrow Y)}(1) - f_{AVE(X \rightarrow Y)}(0)$ difference in forest loss intervening on conflict
- How to test $H_0 : T = 0$
- Omit all locations which do not have data for both conflict and no conflict

$$\hat{f}_{AVE(X \rightarrow Y)}^{nm}(\mathbf{X}_n^m, \mathbf{Y}_n^m)(x) = \frac{1}{|\mathcal{J}_n^m|} \sum_{i \in \mathcal{J}_n^m} \frac{1}{|\{t : X_{s_i}^t = x\}|} \sum_{t : X_{s_i}^t = x} Y_{s_i}^t$$

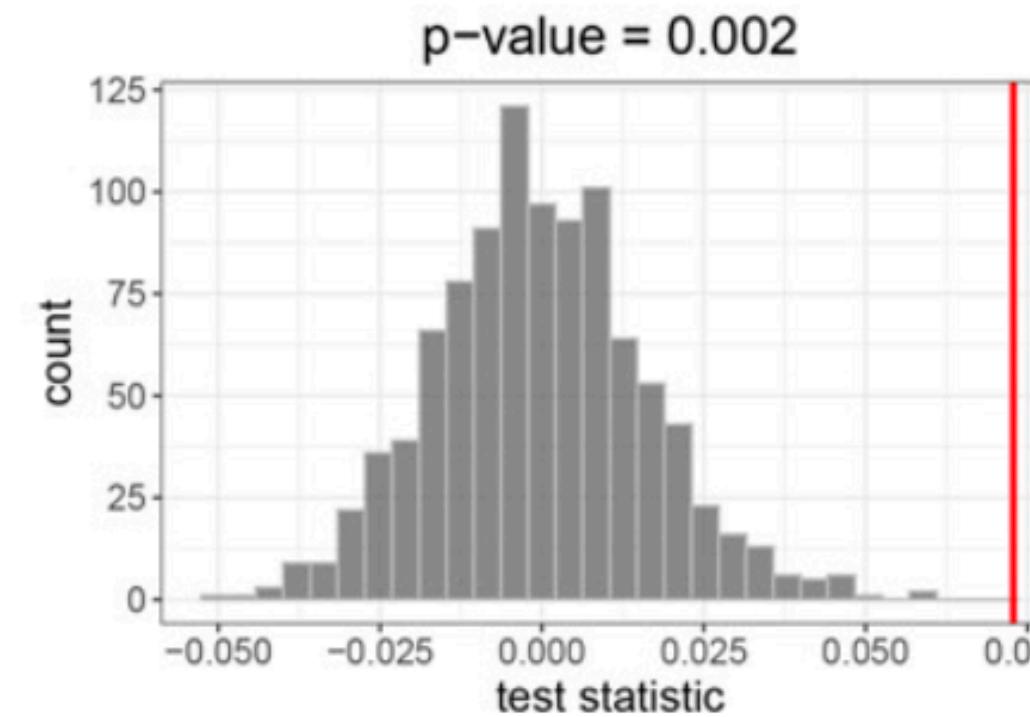
Alternative assumptions on Causal Structure

Quantifying Causal Influence of Conflict on Forest Loss

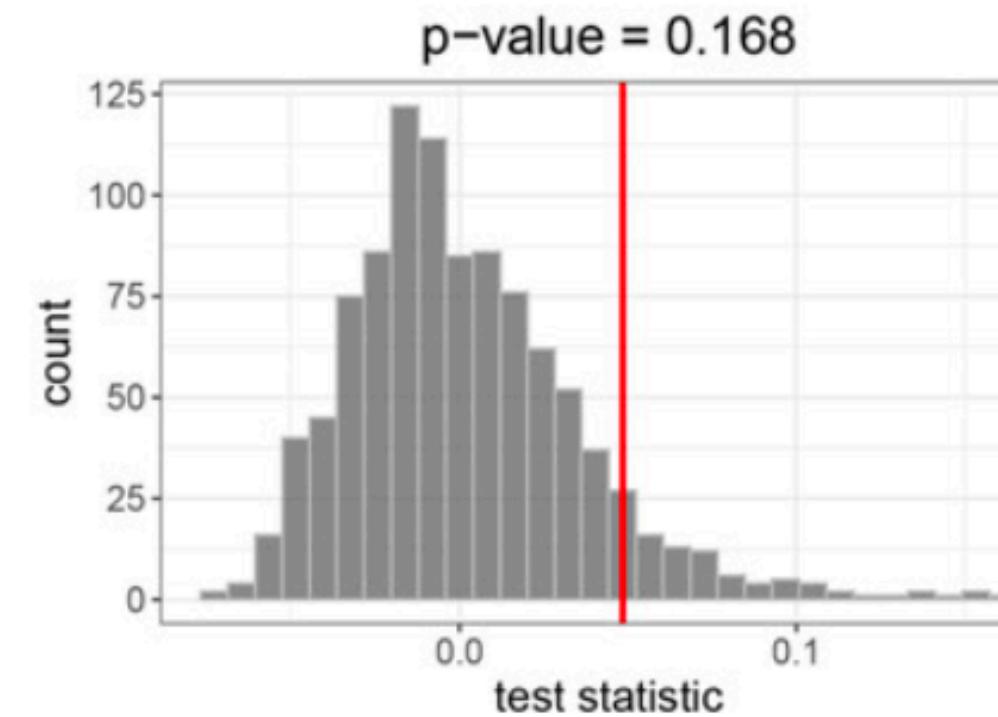
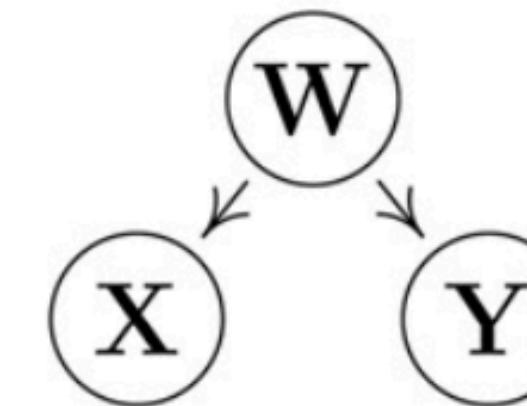
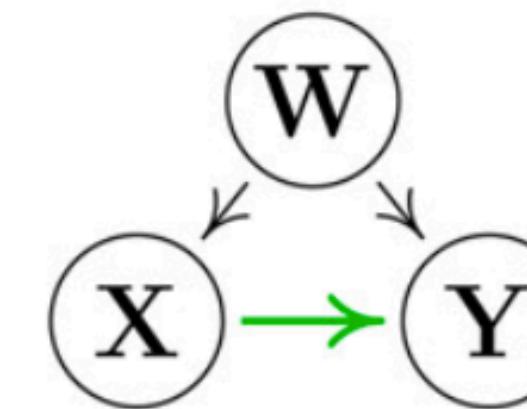
Model 1



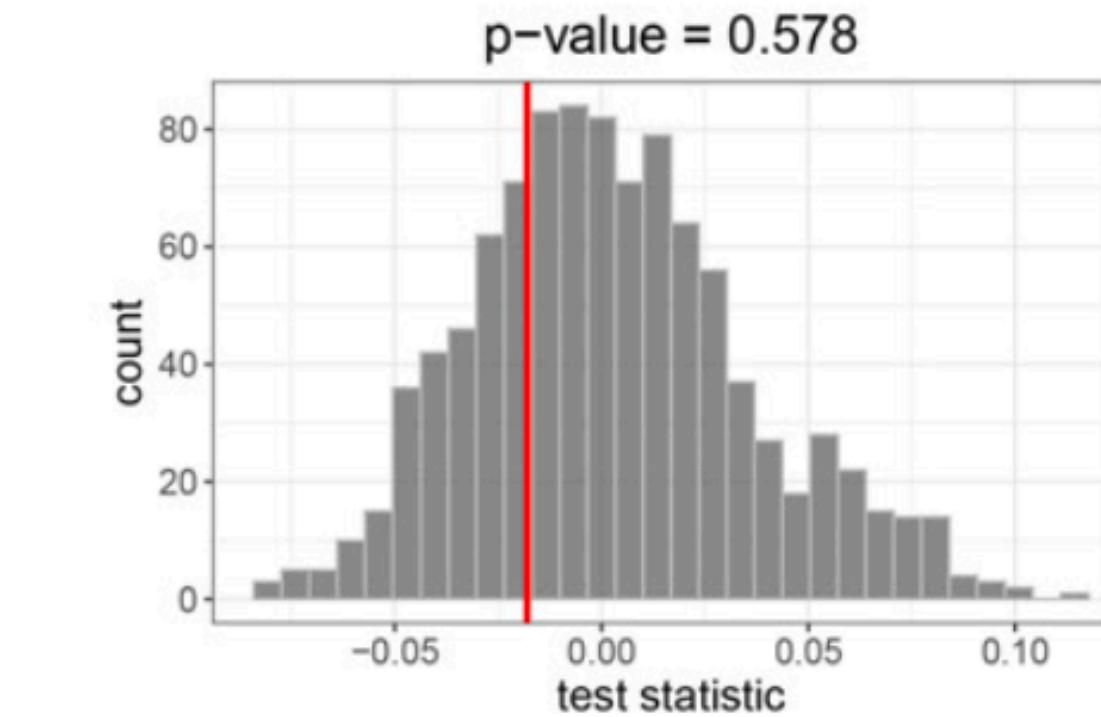
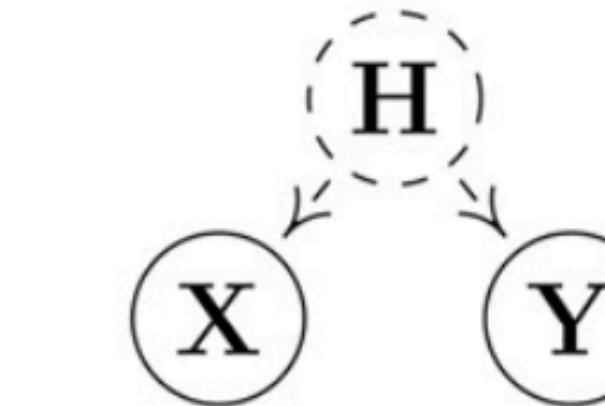
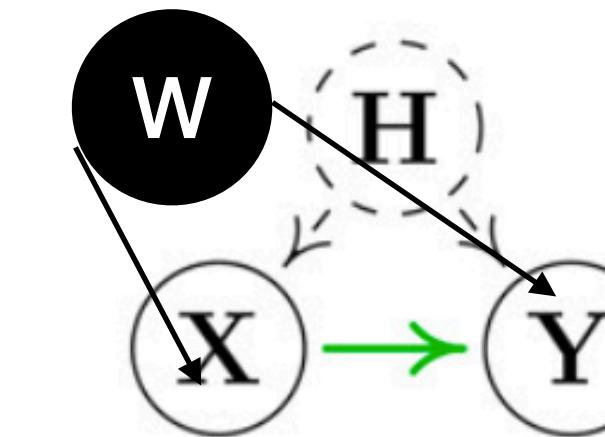
$H_0 :$



Model 2



LSCM

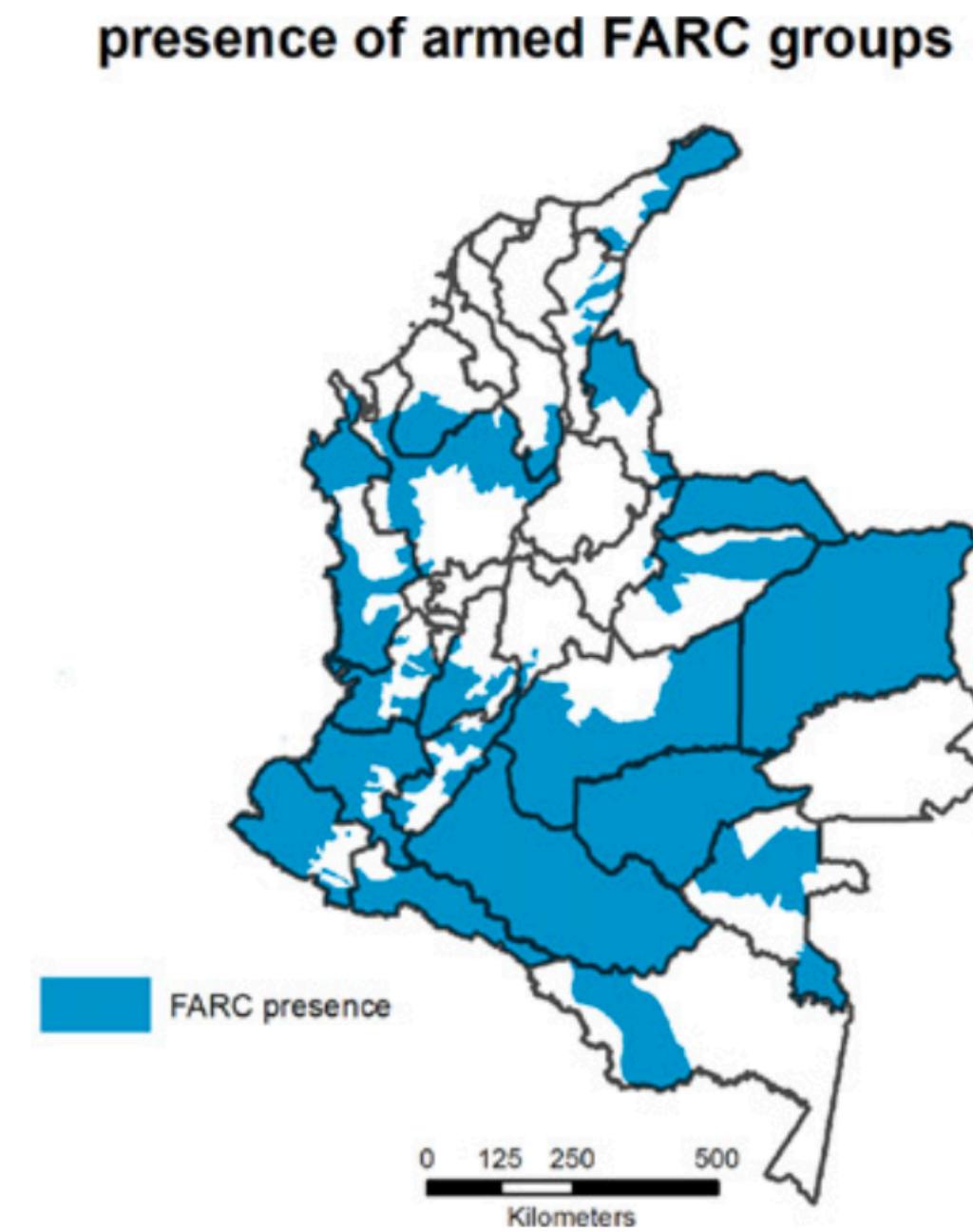
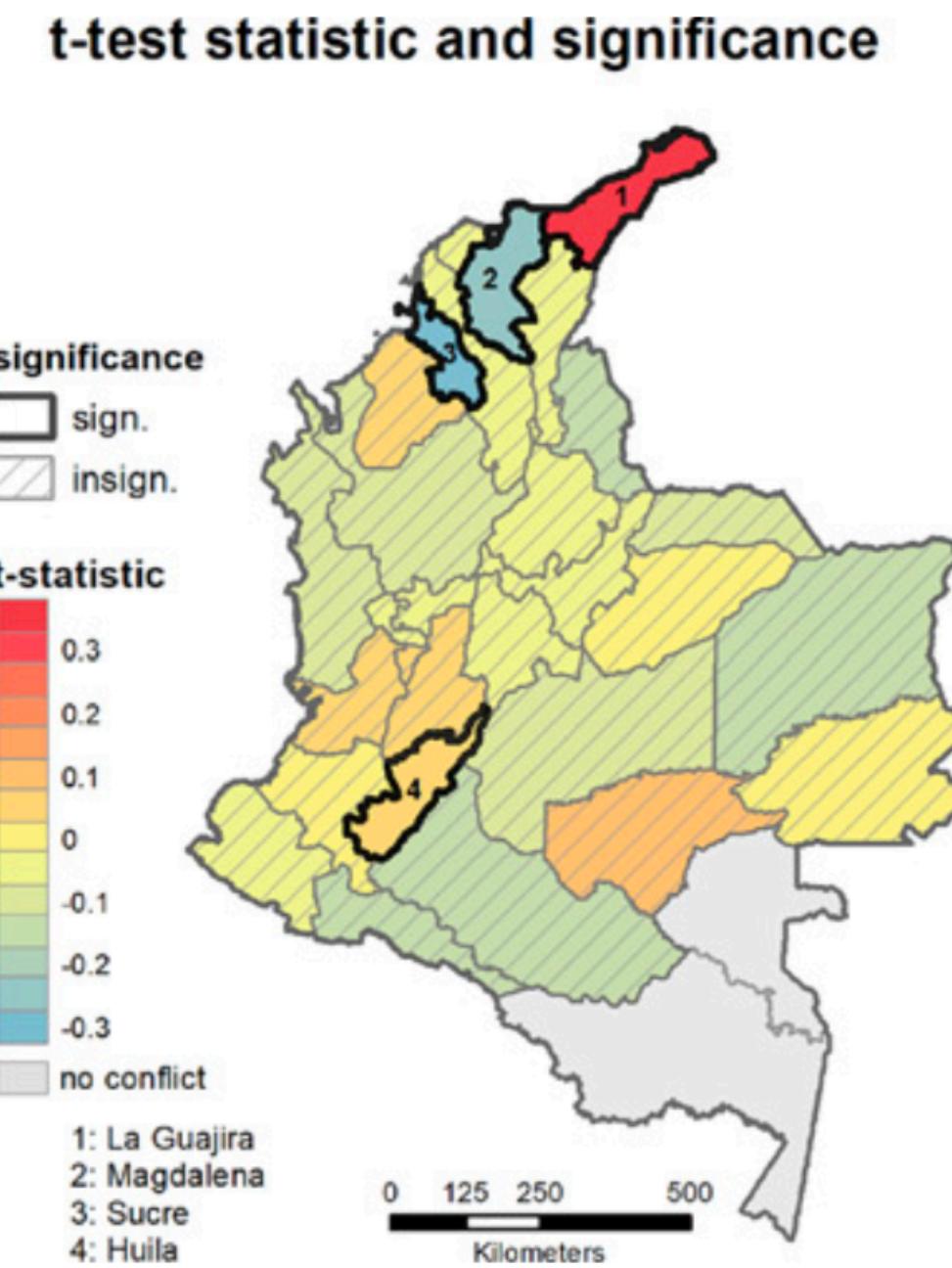
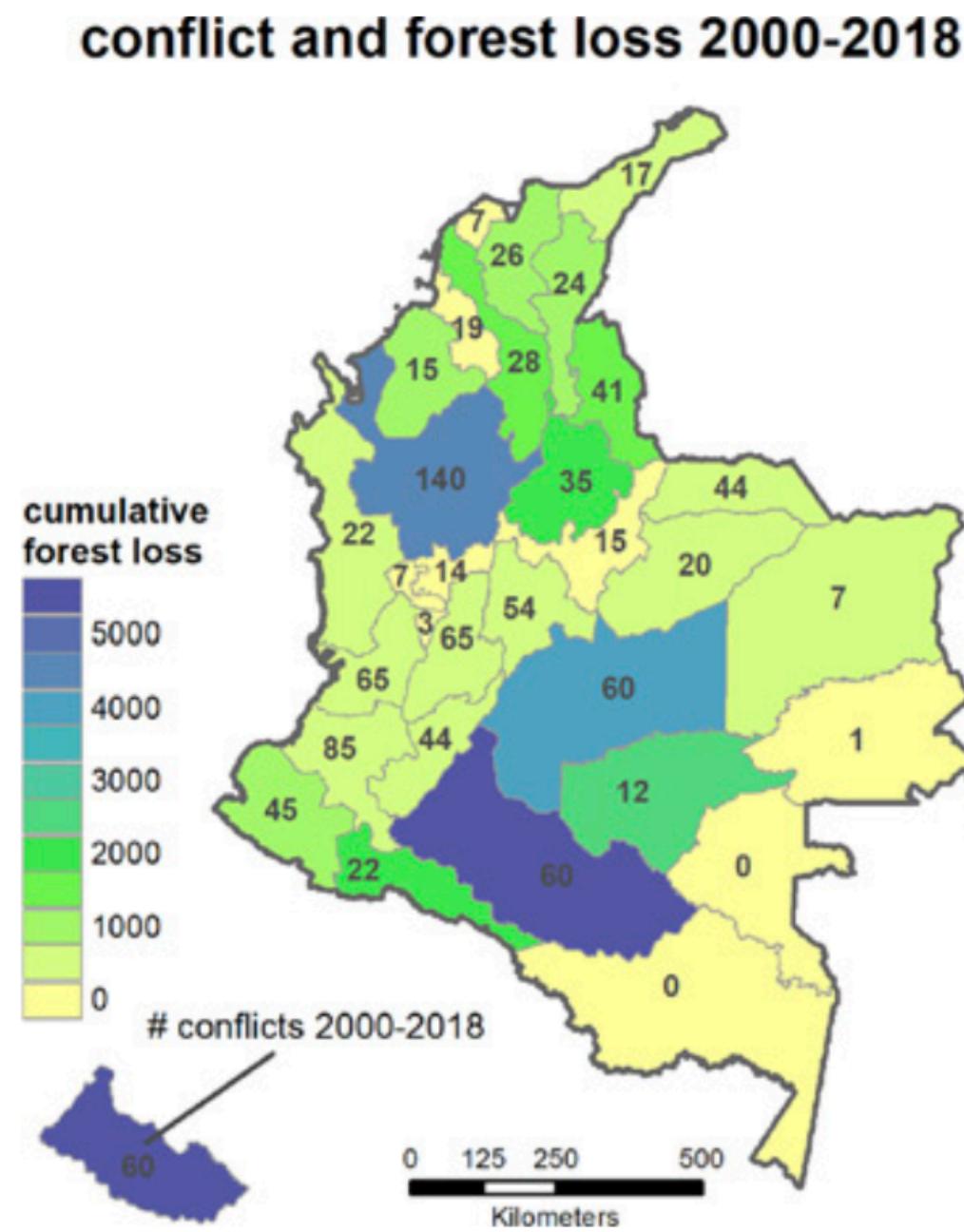


Causal Effects based on Different Models

- Baseline Model: Conflict has a significant positive effect on forest loss ($\hat{T} = 0.073, P = 0.002$)
- Adjusting for Confounders:
 - Accessibility ($\hat{T} = 0.049, P = 0.168$) and Population Density ($\hat{T} = 0.038, P = 0.214$) reduce effect size and remove significance.
 - Accounting for all time-invariant confounders reverses the sign ($\hat{T} = -0.018, P = 0.578$) but remains insignificant.
- Spatial and Temporal Adjustments
 - **No evidence** for spatial spill-over effects
 - No significant effect when accounting for time delay ($\hat{T} = -0.0293, P = 0.354$).
 - Block-permutation tests confirm non-significance.

Regional Analysis of Conflict

Departmental Policies Matter



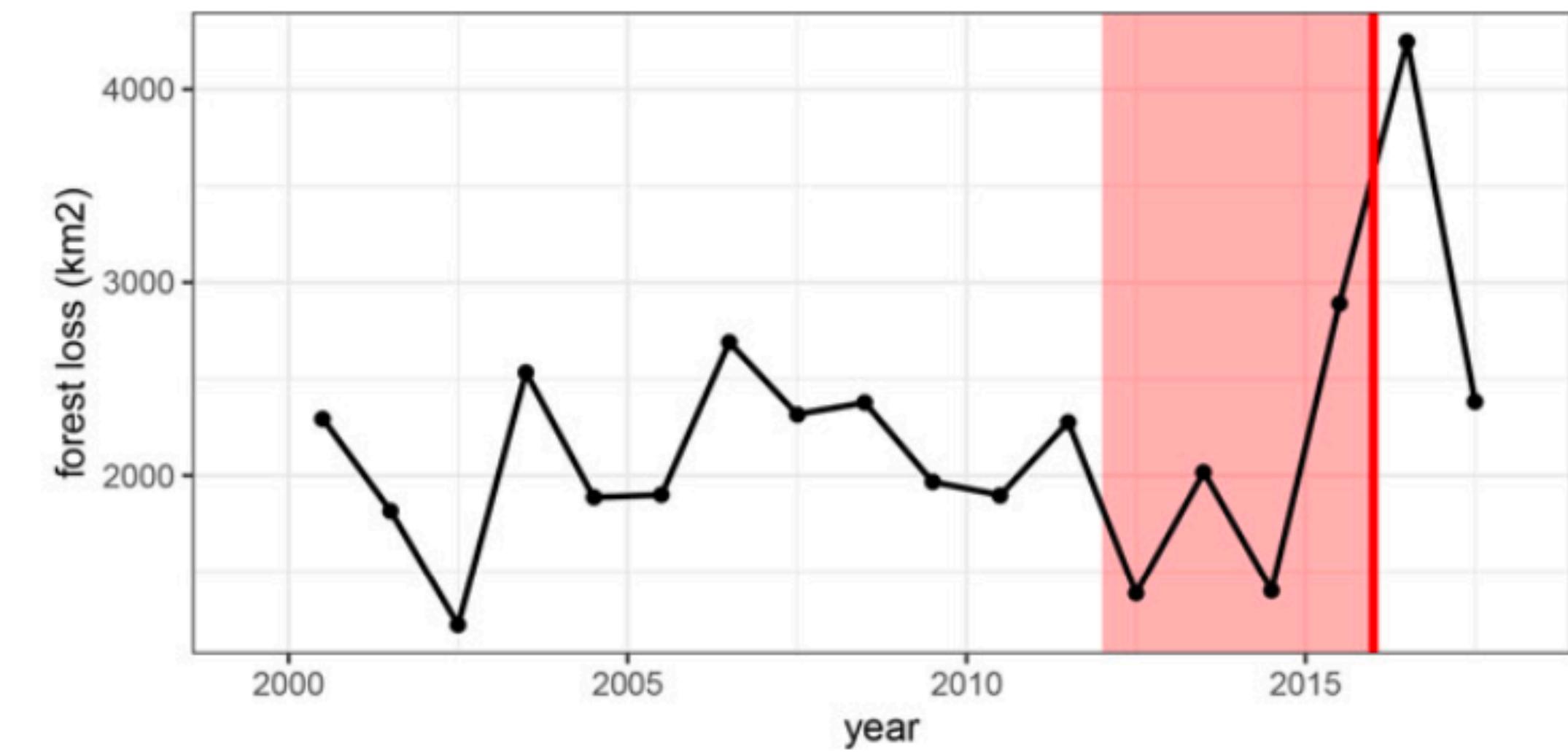
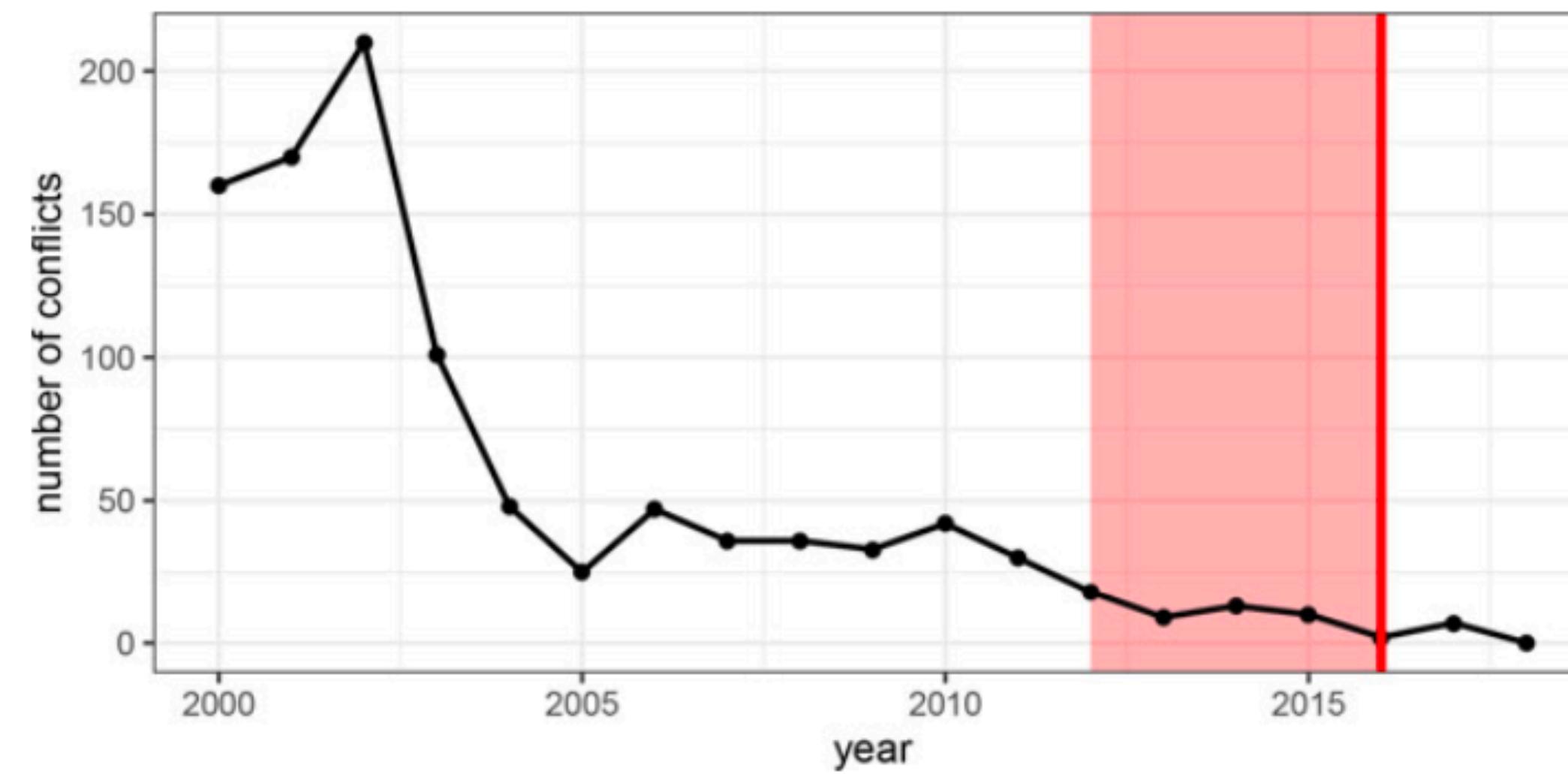
Adjusting for Spatial Heterogeneity

Policy changes across borders result in heterogeneity

- High heterogeneity across departments.
- La Guajira ($\hat{T} = 0.398$, $P = 0.047$): Strongest positive effect on deforestation.
- Magdalena ($\hat{T} = -0.218$, $P = 0.004$): Significant negative effect.
- Huila ($\hat{T} = 0.095$, $P = 0.023$): Moderate positive effect.
- FARC-Controlled Areas:
 - 6 out of 8 regions show negative effects of conflict on deforestation.
 - Explanation for positive effect: Forest cover was a **strategic resource** for internal governance and drug production.

Verifying Intervention Effects

Government Interventions leads to resolution of local tensions



Recap

- Causal vs Predictive Analysis
- New causal framework particularly designed for multivariate spate temporal processes
- Estimation and testing of causal effects (non parametric hypothesis test for causal relationships, asymptotic consistency proven through simulations)
- Empirical findings (no country wide effect, regional variability, sociopolitical influence)
- Interpretations of the method (finding align with post conflict deforestation surge, potential bias from time invariant confounders, relative role of conflict)

Extension and Future Directions

Points for discussion

- Combining time variant confounders with unobserved confounders
- Temporally lagged causal effects?
- Space time interchangeability and analysis based on that
- Slowly varying unobserved confounders
- Smoothness Assumptions in space might help with hypothesis testing
- Datasets and other applications where such an analysis might be useful
- Counterfactuals based on the method proposed

**Thanks for participating
Volunteers for presenting on 2nd April?**