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**Problem Set 5 *Solutions***

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**Question 1.** This problem will use regression difference-in-differences to estimate the impact of a breakfast in the classroom (BIC) program on school meals program participation in New York City. BIC was not implemented under random assignment; rather, schools voluntarily adopted the program. We do, however, have data for these and other schools before and after adoption. (See Corcoran, Elbel, & Schwartz 2015 for details). **(40 points)**

- (a) In Stata, open the panel dataset *NYCbklunch.dta*. This file consists of school-level data in which the rows are elementary or middle schools observed in year  $t$  ( $t=2005$  to 2012). The outcome variables of interest are *bklunch* and *bkbkfast*, which are average daily participation rates in the school breakfast and lunch programs. Provide some descriptive statistics for these two variables. On what scale are they measured? **(2 points)**

**See attached log. Each of the outcome variables ranges between 0-1 and can be interpreted as the proportion of students in attendance on the average school day in school  $i$  who received a breakfast or lunch. The average daily participation was 0.780 (or 78.0%) for lunch and only 0.258 (or 25.8%) for breakfast.**

- (b) Use `xtset` to declare the data as a panel. Which variable is the cross-sectional unit and which is the time dimension? Is this a balanced panel? (Show how you answer the last question). How many schools are observed in all 8 years? **(2 points)**

**See attached log. *schoolid* is the cross-sectional unit, while *year* is the time dimension. This is a balanced panel, as easily seen in `xtdescribe` which shows the pattern of data availability. 907 schools have data in all 8 years.**

- (c) This dataset contains a time-varying treatment variable called *bicpost* that is equal to 1 in years in which the school offered BIC and 0 otherwise. For this part, limit the sample to schools that adopted BIC in 2010 and never-adopters. Estimate a simple “2x2” difference-in-differences regression comparing mean breakfast participation rates for two groups (treated and untreated schools) in two time periods (pre and post). Do the same for lunch participation rates, and interpret your results. (Note: ignore the panel nature of your data and just use the `regress` command for this part). Is the BIC effect statistically significant? Practically significant? What assumptions must be satisfied for this difference-in-differences to be considered a causal effect? **(6 points)**

**See attached log. The `do` file first identifies the first year of treatment for each school and then flags schools that adopted BIC in 2010. The**

2x2 DD coefficient is positive and statistically significant for breakfast, and statistically insignificant for lunch. In the case of breakfast, average daily participation increased 23.9 percentage points more in schools that adopted BIC in 2010 (the treated schools) relative to schools that never adopted BIC (the untreated schools).

For this estimate to be considered causal, the parallel trends assumption must be satisfied. That is, the pre-post change in breakfast participation among non-BIC schools must represent what would have occurred in BIC schools had they not been “treated” by adopting the program. As a panel model, the strict exogeneity assumption must also hold. This means the error term in every year  $t$  cannot be related to treatment adoption. This would be violated if there were anticipation effects, for example.

- (d) Repeat part (c) but include year dummies in the regressions. How does this affect your impact estimates for BIC, if at all? Continue to use the `regress` command, and explain why one of the post-2010 year effects is not estimable. (4 points)

See attached log. There is only a trivial change to the estimated DD coefficient in both the breakfast and lunch regressions. One of the post-2010 year effects is omitted. This is because of perfect collinearity between the *post2010* variable and the three dummies for 2010, 2011, and 2012. (If you know three of these variables, the fourth is also known).

- (e) In parts (c)-(d) you did not take advantage of the panel nature of your data. Re-estimate these regressions using `xtreg` and school fixed effects. How do your estimates compare? (4 points)

See attached log. Again there is only a trivial change to the estimated DD coefficient in both regressions.

- (f) Compare the mean baseline characteristics of treated and untreated schools in parts (c)-(e). (Use 2009 as the baseline year). Look at the following: total enrollment, % ELL, % special education, % Asian, % black, % Hispanic, % female, % free lunch eligible (*free1*), % reduced price lunch eligible (*redu1*). How do schools that adopted BIC in 2010 compare to those that never adopted? (4 points)

See attached log. It is clear that schools which adopted BIC in 2010 differ systematically from those that didn't. They are more disadvantaged on a number of dimensions, including %ELL, %sped, and percent with family income low enough to be eligible for free meals. These differences suggest we should include these covariates in the regressions. Others—such as the race or gender composition of the school—may also help to reduce unexplained variation in the outcome.

- (g) Now estimate the same regression models in part (e), but include the school covariates listed above in your regressions. How do your estimates of the “BIC effect” change, if at all? (And how are these covariates related to meal participation?) (4 points)

See attached log. It appears the inclusion of covariates has little effect on the difference-in-differences estimates for breakfast and lunch participation. Many of the included covariates are related to meals program participation. For example, *%ELL*, *%sped*, and percent eligible for free meals are all positively related to breakfast participation. *Total enrollment*, *Asian*, *Hispanic*, and *female* shares are all negatively related to breakfast participation.

- (h) Repeat part (e)—the model without covariates—but try using the Stata command `xtddidregress`. Compare your results. (4 points)

See attached log. The results are identical, although the standard errors are different (unless you—correctly—adjusted your standard errors for clustering by school).

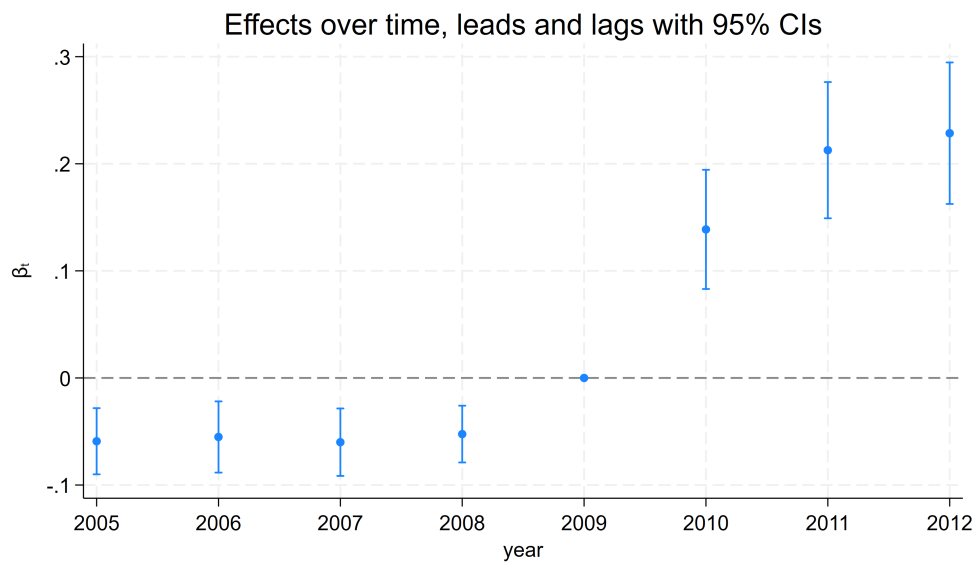
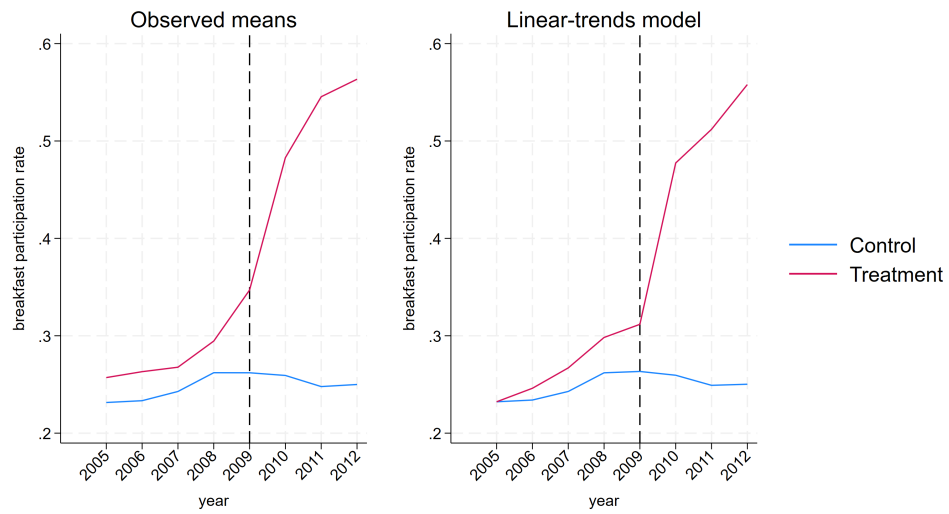
- (i) State the assumption(s) that are required to hold in order to interpret the breakfast participation DD coefficient in part (h) as causal. Use the post-estimation commands `estat trendplots` and `estat granger` to probe these assumptions, and interpret the results. (5 points)

Again, the parallel trends and no anticipation assumptions must be satisfied (see part c). See attached log for the post-estimation commands (for the breakfast model only). The Granger causality test rejects the null hypothesis of “no anticipation.” The graphical output is below. These results are *not* consistent with a common pre-trend for the treated and untreated schools. The second plot is a Granger plot (`estat grangerplot`) which shows more clearly that the trends are consistent until 2009—the year before treatment. This could be an anticipation effect, or an indication that our information about the timing of treatment is incorrect for one or more schools.

- (j) Finally, incorporate all treated schools into the analysis regardless of when they adopted BIC. Estimate the two-way fixed effects model using `xtreg` (as in part e) and `xtddidregress` (as in part h). Try these without and with school covariates. How do your results compare? (5 points)

See attached log. The results are very similar to the previous ones, in all cases, when including the full set of treated schools. Note we have not addressed the potential problems introduced by staggered treatment timing.

## Graphical diagnostics for parallel trends



**Question 2.** For these questions, refer to the recent article by Cellini and Turner (2019), “Gainfully Employed? Assessing the Employment and Earnings of For-Profit College Students Using Administrative Data.” You can find the article here: <http://jhr.uwpress.org/content/54/2/342.abstract>. **(32 points)**

- (a) Cellini and Turner use a generalized difference-in-differences regression model to estimate the causal effect of attending a for-profit certificate program on labor market outcomes. What specific outcome variables do they examine, and what dataset(s) do they use? **(4 points)**

**The main outcome variables include employment (0/1) and earnings (levels and logs). US DOE data identify all federally-aided students who exited a for-profit post-secondary institution or public community college certificate program between 2006 and 2008. Income and employment status are taken from tax data from the IRS (1999-2014).**

- (b) How is the “treatment” variable defined here and what are the possible “pre” and “post” years? How many potential pre and post years are there? **(4 points)**

**“Treatment” is attendance at a for-profit post-secondary institution. The potential “pre-treatment” years are 1999-2007, depending on when students started and exited their program. The potential “post” years are 2007-2014, again depending on when students left their program. Thus, there are approximately 6 pre-treatment years and 5-6 post-treatment years available.**

- (c) The authors use three different groups of “untreated” individuals as comparison groups. What were they, and what was their rationale for looking at each? Which comparison group is their “preferred” one, and why? **(4 points)**

**Their comparison groups include: (1) public community college students; (2) public community college students matched by demographics, prior earnings, field of study, geography, and age group; and (3) a matched sample of individuals with no post-secondary education. The first two comparison groups are used to estimate the effect of attending a for-profit institution relative to a public institution. The second is preferred by the authors since it accounts for differences in the composition of these two populations (for-profit and public). The third comparison group is used to estimate the effect of attending a for-profit institution relative to no college. One challenge with the “no college” comparison group is the lack of a defined pre- and post- period. Unlike the community college comparison group, the no college group also cannot be matched based on, say, field of study.**

- (d) Equation (1) on page 350 shows their regression specification. Carefully explain what each term represents, and how the causal effect of attending a for-profit certificate program is being identified. Why is there not a main effect for “For-Profit” in the model? **(5 points)**

$$y_{it} = \alpha_0 + \alpha_1(Post_{it}) + \alpha_2(Post_{it} * ForProfit_i) + d_t + d_a + d_i + \epsilon_{it}$$

- $Post_{it} = 1$  in all years following exit from the post-secondary program
- $ForProfit_i = 1$  for individuals  $i$  that attended a for-profit program
- $d_t$  are year effects to capture mean differences in the outcome due to macroeconomic conditions
- $d_a$  are age fixed effects to capture mean differences in the outcome due to, say, work experience
- $d_i$  are individual fixed effects
- $\alpha_1$  is the “first difference”: the mean pre-to-post change in the outcome for individuals who attended public certificate programs
- $\alpha_2$  is the “second difference” (or DD): the differential pre-to-post change in the outcome for individuals who attended a for-profit program

There are individual fixed effects in this regression, which means coefficient estimates are identified using *within-person* changes over time. There is no  $ForProfit_i$  main effect since it is collinear with the individual effect. (It does not vary over time, but rather just indicates whether or not the student ultimately enrolled in a for-profit institution).

- (e) Carefully explain the main assumption necessary to interpret the difference-in-differences estimate here as a causal effect. What evidence do the authors provide that this assumption holds for their three different comparison groups? **(5 points)**

The assumption is that the pre-to-post change in the outcome for the comparison group (e.g., public college attendees) represents what would have occurred for the treatment group (for-profit attendees), had they not been treated. Figure 1 displays the trend in employment and earnings for each group, prior to their enrollment in a post-secondary program. While there are differences in the levels of earnings and employment, the time trends look very similar. This is particularly true for the matched sample. This provides some confidence that the trends would have remained the same in the absence of the treatment.

- (f) The paper’s main results are reported in Table 3. Carefully interpret the coefficients reported in Panel B. What additional evidence does Figure 2 provide? **(5 points)**

Panel B of Table 3 reports the estimated pre-post change for the matched public sample, and the DD. For individuals who enrolled in public community college programs, annual earnings increased by \$1,069, on average, in years after exit from the program. For individuals in for-profit programs,

this change was \$2,144 *lower*. Taken together, the annual earnings of the treatment group *declined* by an average of \$1,075.

The log earnings column excludes individuals with zero earnings, and thus should be interpreted as conditional on working. For individuals in public community college programs, annual earnings increased by about 16.8%. For individuals in for-profit programs, this change was 11.3 percentage points *lower*. Taken together, their annual earnings—conditional on working—increased by 5.5%.

The first column shows the effect on employment. For individuals in public programs, there was no significant change in the probability of employment. However, for individuals in for-profit programs, the probability of employment *declined* by 1.5 percentage points after exiting their program.

Figure 2 is an event study—it shows the differential effect of attending a for-profit program separately for each year after exiting the program.

- (g) Finally, Figure 4 shows the distribution of earnings effects *by school* for public and for-profit institutions. Cellini and Turner describe these as the result of “single-difference” regressions. Briefly explain what they mean by this, and why these should not be interpreted as the causal effects of attending specific institutions. (5 points)

It would be difficult to estimate a separate difference-in-differences model for each school since it is not obvious who the comparison group should be. (I.e., who are the students exiting public programs that might have attended *that specific for-profit program*?) Instead, Cellini & Turner estimate mean within-person, pre-post changes in earnings for each for-profit institution. These are single differences because they represent only the mean pre-post change of one group. As with any single difference, they should not be interpreted as causal, as it is difficult to separate the treatment effect from any change in the outcome over time that might have occurred in the absence of the treatment.

```

. // *****
. // LPO-8852 Problem set 5 solutions
. // Last updated: October 7, 2024
. // *****

. // ****
. // (a)
. // ****
. // Get data

.       use https://github.com/spcorcor18/LPO-8852/raw/main/data/NYCbkgfastlunc
> h-v2.dta, clear

.       summ bkfast_part lunch_part

      Variable |           Obs       Mean   Std. dev.       Min       Max
-----+-----
bkfast_part |       7,255   .2580688   .1377646   .0054945   .9725
lunch_part  |       7,005   .7803097   .1838588   .0122249         1

.
.
. // ****
. // (b)
. // ****
. // Set panel

.       xtset schoolid year

Panel variable: schoolid (strongly balanced)
Time variable: year, 2005 to 2012
Delta: 1 unit

.       xtdescribe

schoolid:  1, 2, ..., 1179          n =          907
year:     2005, 2006, ..., 2012      T =           8
Delta(year) = 1 unit
Span(year) = 8 periods
(schoolid*year uniquely identifies each observation)

Distribution of T_i:   min      5%      25%      50%      75%      95%      max
                     8         8         8         8         8         8         8

      Freq.  Percent   Cum. | Pattern
-----+-----
      907    100.00  100.00 | 11111111
-----+-----
      907    100.00         | XXXXXXXX

.
.
. // ****
. // (c)
. // ****
. // Simple 2x2 DD for limited sample

.       tabulate year bicpost

```



=1 if post period for			
BIC school			
year	0	1	Total
2005	907	0	907
2006	907	0	907
2007	906	1	907
2008	902	5	907
2009	878	29	907
2010	848	59	907
2011	828	79	907
2012	800	107	907
Total	6,976	280	7,256

```
.
. // identify first year that school was treated
. egen temp1=min(year) if bicpost==1, by(schoolid)
(6,976 missing values generated)

. egen firstyear=max(temp1), by(schoolid)
(6,400 missing values generated)

. drop temp1

. table year firstyear
```

firstyear							
	2007	2008	2009	2010	2011	2012	Total
year							
2005	1	4	24	30	20	28	107
2006	1	4	24	30	20	28	107
2007	1	4	24	30	20	28	107
2008	1	4	24	30	20	28	107
2009	1	4	24	30	20	28	107
2010	1	4	24	30	20	28	107
2011	1	4	24	30	20	28	107
2012	1	4	24	30	20	28	107
Total	8	32	192	240	160	224	856

```
.
. // simple DD: treated schools adopting in 2010 vs never adopted
. // create "post" period variable
. gen post2010 = (year>=2010 & year~=.)

.
. // create "ever treated" variable
. gen everbic2010 = firstyear==2010

.
. _eststo bk1: reg bkfast_part i.everbic2010##i.post2010 if (firstyear==
> 2010 | firstyear==.)
```

Source	SS	df	MS	Number of obs	=	6,639
Model	7.27543865	3	2.42514622	F(3, 6635)	=	152.58
Residual	105.458173	6,635	.015894224	Prob > F	=	0.0000
				R-squared	=	0.0645
				Adj R-squared	=	0.0641
Total	112.733612	6,638	.016983069	Root MSE	=	.12607

bkfast_part	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
1.everbic2010	.0395468	.010485	3.77	0.000	.0189927	.0601008
1.post2010	.0060172	.0032553	1.85	0.065	-.0003643	.0123986
everbic2010# post2010						
1 1	.2386908	.0171219	13.94	0.000	.2051262	.2722553
_cons	.2463439	.0019936	123.57	0.000	.2424358	.2502521

```
. _eststo lul1: reg lunch_part i.everbic2010##i.post2010 if (firstyear==2
> 010 | firstyear==.)
```

Source	SS	df	MS	Number of obs	=	6,423
Model	1.07347052	3	.357823508	F(3, 6419)	=	10.27
Residual	223.584325	6,419	.034831644	Prob > F	=	0.0000
Total	224.657795	6,422	.034982528	R-squared	=	0.0048
				Adj R-squared	=	0.0043
				Root MSE	=	.18663

lunch_part	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
1.everbic2010	.0664884	.016285	4.08	0.000	.0345645	.0984124
1.post2010	-.006335	.0048792	-1.30	0.194	-.0158999	.0032299
everbic2010# post2010						
1 1	.0057016	.0260823	0.22	0.827	-.0454284	.0568315
_cons	.7758114	.0030141	257.39	0.000	.7699027	.7817201

```
.
.
. // ****
. // (d)
. // ****
. // Add year effects
.
. _eststo bk2: reg bkfast_part i.everbic2010##i.post2010 i.year if (firs
> tyear==2010 | firstyear==.)
note: 2012.year omitted because of collinearity.
```

Source	SS	df	MS	Number of obs	=	6,639
Model	8.11373426	9	.901526029	F(9, 6629)	=	57.12
Residual	104.619878	6,629	.015782151	Prob > F	=	0.0000
Total	112.733612	6,638	.016983069	R-squared	=	0.0720
				Adj R-squared	=	0.0707
				Root MSE	=	.12563

bkfast_part	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
1.everbic2010	.0395506	.010448	3.79	0.000	.0190692	.0600321
1.post2010	.0202802	.0061994	3.27	0.001	.0081274	.0324331
everbic2010# post2010						
1 1	.2386869	.0170615	13.99	0.000	.2052409	.2721329
year						
2006	.0019981	.0061686	0.32	0.746	-.0100944	.0140906
2007	.011271	.0061686	1.83	0.068	-.0008216	.0233635
2008	.0308351	.0061686	5.00	0.000	.0187425	.0429276
2009	.0326727	.0061686	5.30	0.000	.0205802	.0447652
2010	.0060077	.0061668	0.97	0.330	-.0060812	.0180966
2011	-.0027192	.0061668	-0.44	0.659	-.0148081	.0093697
2012	0	(omitted)				

_cons	.2309847	.0043796	52.74	0.000	.2223994	.2395701
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```
. _eststo lu2: reg lunch_part i.everbic2010##i.post2010 i.year if (first
> year==2010 | firstyear==.)
note: 2012.year omitted because of collinearity.
```

Source	SS	df	MS	Number of obs	=	6,423
Model	1.55895067	9	.173216741	F(9, 6413)	=	4.98
Residual	223.098845	6,413	.03478853	Prob > F	=	0.0000
				R-squared	=	0.0069
				Adj R-squared	=	0.0055
Total	224.657795	6,422	.034982528	Root MSE	=	.18652

lunch_part	Coefficient	Std. err.	t	P> t	[95% conf. interval]
1.everbic2010	.0666754	.0162752	4.10	0.000	.0347706 .0985801
1.post2010	-.0026473	.0093423	-0.28	0.777	-.0209614 .0156667
everbic2010# post2010					
1 1	.0057818	.0260666	0.22	0.824	-.0453174 .0568811
year					
2006	.0092882	.0093145	1.00	0.319	-.0089713 .0275477
2007	.01739	.0093231	1.87	0.062	-.0008864 .0356663
2008	.0228403	.009385	2.43	0.015	.0044426 .0412379
2009	.0235812	.0094064	2.51	0.012	.0051416 .0420208
2010	.0203108	.0092229	2.20	0.028	.0022309 .0383908
2011	.0121711	.009234	1.32	0.188	-.0059306 .0302729
2012	0	(omitted)			
_cons	.7612627	.0066357	114.72	0.000	.7482546 .7742708

```
.
.
. // ****
. // (e)
. // ****
. // Use xtreg and school fixed effects
.
. _eststo bk3: xtreg bkfast_part i.bicpost i.year if (firstyear==2010 |
> firstyear==.), fe
```

Fixed-effects (within) regression	Number of obs	=	6,639
Group variable: schoolid	Number of groups	=	830

R-squared:	Obs per group:
Within = 0.2143	min = 7
Between = 0.0414	avg = 8.0
Overall = 0.0697	max = 8

corr(u_i, Xb) = 0.0349	F(8, 5801)	=	197.80
	Prob > F	=	0.0000

bkfast_part	Coefficient	Std. err.	t	P> t	[95% conf. interval]
1.bicpost	.2386411	.0070527	33.84	0.000	.2248152 .2524671
year					
2006	.0022187	.00255	0.87	0.384	-.0027804 .0072177
2007	.0114915	.00255	4.51	0.000	.0064925 .0164905
2008	.0310556	.00255	12.18	0.000	.0260566 .0360547
2009	.0328933	.00255	12.90	0.000	.0278942 .0378923
2010	.0265101	.0025628	10.34	0.000	.0214861 .0315341
2011	.0177832	.0025628	6.94	0.000	.0127593 .0228072
2012	.0205024	.0025628	8.00	0.000	.0154785 .0255264

```

      _cons | .2322215 .0018037 128.74 0.000 .2286855 .2357575
-----+-----
      sigma_u | .11608572
      sigma_e | .05193033
      rho | .8332519 (fraction of variance due to u_i)
-----+-----
F test that all u_i=0: F(829, 5801) = 39.90 Prob > F = 0.0000

.      _eststo lu3: xtreg lunch_part i.bicpost i.year if (firstyear==2010 | f
> firstyear==.), fe

Fixed-effects (within) regression              Number of obs   =       6,423
Group variable: schoolid                      Number of groups  =       830

R-squared:                                   Obs per group:
    Within = 0.0324                               min =           2
    Between = 0.0216                              avg  =          7.7
    Overall = 0.0023                               max  =           8

corr(u_i, Xb) = -0.0113                      F(8, 5585)         =       23.41
                                              Prob > F           =       0.0000

```

```

-----+-----
lunch_part | Coefficient Std. err. t P>|t| [95% conf. interval]
-----+-----
1.bicpost | .0013516 .0090263 0.15 0.881 -.0163435 .0190467
year |
2006 | .0063772 .0031937 2.00 0.046 .0001163 .0126381
2007 | .0158261 .003197 4.95 0.000 .0095588 .0220935
2008 | .0252871 .0032223 7.85 0.000 .0189702 .031604
2009 | .0282942 .0032308 8.76 0.000 .0219605 .0346279
2010 | .0129657 .0031958 4.06 0.000 .0067006 .0192307
2011 | .0054245 .0032009 1.69 0.090 -.0008505 .0116995
2012 | -.0052814 .0032046 -1.65 0.099 -.0115636 .0010008
-----+-----
      _cons | .7647871 .0022679 337.22 0.000 .7603412 .7692331
-----+-----
      sigma_u | .1764832
      sigma_e | .06381114
      rho | .88438184 (fraction of variance due to u_i)
-----+-----
F test that all u_i=0: F(829, 5585) = 59.53 Prob > F = 0.0000

```

```

.
.
. // ****
. // (f)
. // ****
. // Descriptive statistics for BIC 2010 adopters and those who never adopt BIC
.
.      sum totalenrollment- pctfemale free1 redul if firstyear==2010 & year==
> 2009, sep(0)

```

Variable	Obs	Mean	Std. dev.	Min	Max
totalenrol~t	30	628.2	263.4843	246	1244
pctell	30	14.79	11.17585	1.1	44
pctsped	30	18.31	7.017483	5.7	36.5
pctasian	30	2.7	4.906294	0	20.8
pctblack	30	45.85	29.04892	1.7	95.2
pcthispan	30	46.34	27.65217	3.5	91.8
pctwhite	30	4.046667	11.81582	.2	64.6
pctfemale	30	49	3.054336	43.3	54.9
free1	30	82.38852	12.33706	33.07393	100
redul	30	7.525835	3.735603	.8522727	15.38461

```
.      sum totalenrollment- pctfemale free1 redul if firstyear==. & year==200
> 9, sep(0)
```

Variable	Obs	Mean	Std. dev.	Min	Max
totalenrol~t	800	689.1938	321.8793	110	2145
pctell	800	12.53675	10.90083	0	77.3
pctsped	800	15.92213	5.769476	0	44
pctasian	800	14.09225	18.71298	0	91.8
pctblack	800	30.62975	29.66412	0	96.8
pcthispanic	800	37.9295	26.02649	1.6	98.7
pctwhite	800	16.33788	22.33898	0	92.8
pctfemale	800	49.09438	3.938139	36.2	100
free1	800	67.48113	22.88177	5.67201	100
redul	800	10.20999	5.023362	0	26.29717

```
.
.
. // ****
. // (g)
. // ****
. // fixed effects regression with covariates
.
.      global covars "pctell pctsped totalenrollment pctasian pctblack pcthis
> p"
.
.      global covars "$covars pctfemale free1 redul"
.
.      _eststo bk4: xtreg bkfast_part i.bicpost $covars if (firstyear==2010 |
> firstyear==.),fe
```

```
Fixed-effects (within) regression      Number of obs      =      6,639
Group variable: schoolid               Number of groups    =      830
```

```
R-squared:                               Obs per group:
  Within = 0.2149                        min =      7
  Between = 0.2035                       avg  =     8.0
  Overall = 0.2040                       max  =      8
```

```
corr(u_i, Xb) = 0.0871                  F(10, 5799)         =     158.69
                                         Prob > F           =      0.0000
```

bkfast_part	Coefficient	Std. err.	t	P> t	[95% conf. interval]
1.bicpost	.2345413	.006981	33.60	0.000	.2208558 .2482268
pctell	.0005993	.0002543	2.36	0.018	.0001007 .0010978
pctsped	.0020669	.0003058	6.76	0.000	.0014673 .0026665
totalenrol~t	-.000104	8.65e-06	-12.02	0.000	-.0001209 -.000087
pctasian	.0009075	.0003813	2.38	0.017	.00016 .0016549
pctblack	-.0000417	.0003147	-0.13	0.895	-.0006587 .0005753
pcthispanic	.0003415	.000309	1.11	0.269	-.0002643 .0009473
pctfemale	-.0005469	.0003687	-1.48	0.138	-.0012697 .0001758
free1	.0002421	.0000988	2.45	0.014	.0000484 .0004358
redul	.0004033	.0002282	1.77	0.077	-.000044 .0008507
_cons	.2661046	.0303249	8.78	0.000	.2066566 .3255527
sigma_u	.10614395				
sigma_e	.05192171				
rho	.80691974	(fraction of variance due to u_i)			

```
F test that all u_i=0: F(829, 5799) = 26.88      Prob > F = 0.0000
```

```
. _eststo lu4: xtreg lunch_part i.bicpost $covars if (firstyear==2010 |
> firstyear==.),fe
```

```
Fixed-effects (within) regression                Number of obs    =      6,423
Group variable: schoolid                        Number of groups  =      830
```

```
R-squared:                                     Obs per group:
    Within = 0.0777                             min =          2
    Between = 0.3368                             avg  =         7.7
    Overall = 0.3055                             max  =          8
```

```
corr(u_i, Xb) = -0.0535                        F(10, 5583)       =      47.01
                                                Prob > F          =      0.0000
```

lunch_part	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
1.bicpost	-.0060766	.0087263	-0.70	0.486	-.0231836	.0110304
pctell	-.0002715	.000313	-0.87	0.386	-.0008852	.0003421
pctsped	.000745	.0003748	1.99	0.047	.0000103	.0014797
totalenroll~t	-.0000864	.0000106	-8.12	0.000	-.0001073	-.0000655
pctasian	.0042328	.0004605	9.19	0.000	.00333	.0051356
pctblack	.0043441	.0003814	11.39	0.000	.0035964	.0050918
pcthispanic	.0033796	.000376	8.99	0.000	.0026425	.0041168
pctfemale	-.0008161	.0004515	-1.81	0.071	-.0017013	.000069
free1	.0009428	.000121	7.79	0.000	.0007056	.0011801
redul	.001884	.0002761	6.82	0.000	.0013426	.0024253
_cons	.4647235	.0369153	12.59	0.000	.3923551	.537092
sigma_u	.1437498					
sigma_e	.06231349					
rho	.84181471	(fraction of variance due to u_i)				

```
F test that all u_i=0: F(829, 5583) = 25.69                Prob > F = 0.0000
```

```
.
.
. // ****
. // (h)
. // ****
. // try using xtdidregress
.
.       xtdidregress (bkfast_part) (bicpost) if (firstyear==2010 | firstyear==
> .), group(schoolid) time(year)
```

Treatment and time information

```
Time variable: year
Control:      bicpost = 0
Treatment:    bicpost = 1
```

	Control	Treatment
Group		
schoolid	800	30
Time		
Minimum	2005	2010
Maximum	2006	2010

```
Difference-in-differences regression                Number of obs = 6,639
Data type: Longitudinal
```

```

                                (Std. err. adjusted for 830 clusters in schoolid)
-----+-----
bkfast_part | Coefficient   Robust      t      P>|t|      [95% conf. interval]
-----+-----
ATET bicpost |
      (1 vs 0) |      .2386411   .0265832    8.98   0.000    .1864628    .2908195
-----+-----
Note: ATET estimate adjusted for panel effects and time effects.
Note: Treatment occurs at different times.

```

```

.      xtdidregress (lunch_part) (bicpost) if (firstyear==2010 | firstyear==.
> ), group(schoolid) time(year)

```

Treatment and time information

```

Time variable: year
Control:      bicpost = 0
Treatment:    bicpost = 1
-----+-----
Group | Control Treatment
-----+-----
schoolid |      800      30
-----+-----
Time |
  Minimum |      2005      2010
  Maximum |      2008      2012
-----+-----

```

Difference-in-differences regression Number of obs = 6,423  
Data type: Longitudinal

```

                                (Std. err. adjusted for 830 clusters in schoolid)
-----+-----
lunch_part | Coefficient   Robust      t      P>|t|      [95% conf. interval]
-----+-----
ATET bicpost |
      (1 vs 0) |      .0013516   .013313    0.10   0.919    -.0247795    .0274827
-----+-----
Note: ATET estimate adjusted for panel effects and time effects.
Note: Treatment occurs at different times and estimation sample contains units
      that switch in and out of treatment.

```

```

.
.
. // ****
. // (i)
. // ****
. // post-estimation commands
.
.      xtdidregress (bkfast_part) (bicpost) if (firstyear==2010 | firstyear==
> .), group(schoolid) time(year)

```

Treatment and time information

```

Time variable: year
Control:      bicpost = 0
Treatment:    bicpost = 1
-----+-----
Group | Control Treatment
-----+-----
schoolid |      800      30
-----+-----
Time |
  Minimum |      2005      2010
  Maximum |      2006      2010
-----+-----

```

Difference-in-differences regression  
Data type: Longitudinal

Number of obs = 6,639

(Std. err. adjusted for 830 clusters in schoolid)

bkfast_part	Coefficient	Robust std. err.	t	P> t	[95% conf. interval]	
ATET						
bicpost						
(1 vs 0)	.2386411	.0265832	8.98	0.000	.1864628	.2908195

Note: ATET estimate adjusted for panel effects and time effects.

Note: Treatment occurs at different times.

.           estat trendplots

.           graph export "trendplots.png", as(png) name("Graph") replace  
file trendplots.png saved as PNG format

.           estat granger

Granger causality test

H0: No effect in anticipation of treatment

F(4, 829) = 4.88

Prob > F = 0.0007

.  
.           // see also grangerplot command  
.           estat grangerplot

.           graph export "grangerplot.png", as(png) name("Graph") replace  
file grangerplot.png saved as PNG format

.  
.           // \*\*\*\*  
.           // (j)  
.           // \*\*\*\*  
.           // incorporate all schools into the analysis  
.           //  
.           // without covariates (xtreg)  
.           \_eststo bk5: xtreg bkfast\_part i.bicpost i.year, fe

Fixed-effects (within) regression                      Number of obs       =       7,255  
Group variable: schoolid                              Number of groups   =       907

R-squared:    Obs per group:  
    Within = 0.3087    min =       7  
    Between = 0.0667   avg =       8.0  
    Overall = 0.1256   max =       8

corr(u\_i, Xb) = 0.0196                               F(8, 6340)           =       353.93  
  Prob > F           =       0.0000

bkfast_part	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
1.bicpost	.2282953	.0048888	46.70	0.000	.2187115	.2378791
year						
2006	.0028761	.0028522	1.01	0.313	-.0027151	.0084673
2007	.0116797	.0028522	4.10	0.000	.0060884	.0172709
2008	.0307599	.0028523	10.78	0.000	.0251684	.0363514
2009	.0306009	.0028565	10.71	0.000	.0250012	.0362005
2010	.0297214	.0028699	10.36	0.000	.0240955	.0353473
2011	.020442	.0028838	7.09	0.000	.0147888	.0260953
2012	.0219312	.0029099	7.54	0.000	.0162268	.0276357
_cons	.230754	.0020174	114.38	0.000	.2267992	.2347088



```

-----+-----
sigma_u | .11573186
sigma_e | .06071925
rho     | .78415233      (fraction of variance due to u_i)
-----+-----

```

F test that all u\_i=0: F(906, 6340) = 29.04                      Prob > F = 0.0000

```

.      _eststo lu5: xtreg lunch_part i.bicpost i.year, fe

```

```

Fixed-effects (within) regression      Number of obs   =      7,005
Group variable: schoolid               Number of groups =      907

```

```

R-squared:                            Obs per group:
    Within = 0.0305                      min =          2
    Between = 0.0357                     avg  =         7.7
    Overall  = 0.0018                     max  =          8

```

```

corr(u_i, Xb) = -0.0168                F(8, 6090)      =      23.95
                                         Prob > F       =      0.0000

```

```

-----+-----
lunch_part | Coefficient   Std. err.      t    P>|t|     [95% conf. interval]
-----+-----
    1.bicpost |   -.0022084   .0053663    -0.41   0.681    - .0127282   .0083114
        year |
    2006      |   .0060649   .0030923     1.96   0.050     2.94e-06   .0121268
    2007      |   .0150901   .0030987     4.87   0.000     .0090156   .0211645
    2008      |   .024707    .0031243     7.91   0.000     .0185822   .0308317
    2009      |   .0265117   .0031399     8.44   0.000     .0203565   .0326669
    2010      |   .0108692   .0031003     3.51   0.000     .0047915   .016947
    2011      |   .0040942   .0031218     1.31   0.190    - .0020256   .0102141
    2012      |  -.0060021   .0031511    -1.90   0.057    - .0121793   .0001751
        _cons |   .7703875   .0021972   350.63   0.000     .7660803   .7746947
-----+-----

```

```

sigma_u | .17289228
sigma_e | .06459134
rho     | .87752268      (fraction of variance due to u_i)
-----+-----

```

F test that all u\_i=0: F(906, 6090) = 55.50                      Prob > F = 0.0000

```

.      // with covariates (xtreg)
.      _eststo bk6: xtreg bkfast_part i.bicpost i.year $covars, fe

```

```

Fixed-effects (within) regression      Number of obs   =      7,255
Group variable: schoolid               Number of groups =      907

```

```

R-squared:                            Obs per group:
    Within = 0.3318                      min =          7
    Between = 0.2925                     avg  =         8.0
    Overall  = 0.2963                     max  =          8

```

```

corr(u_i, Xb) = 0.1284                F(17, 6331)    =     184.93
                                         Prob > F       =      0.0000

```

```

-----+-----
bkfast_part | Coefficient   Std. err.      t    P>|t|     [95% conf. interval]
-----+-----
    1.bicpost |   .2219539   .0048512    45.75   0.000     .2124438   .2314639
        year |
    2006      |   .0015038   .0028377     0.53   0.596    - .004059   .0070666
    2007      |   .0084261   .0028691     2.94   0.003     .0028017   .0140505
    2008      |   .0249954   .0029865     8.37   0.000     .0191408   .0308499
    2009      |   .0248059   .0030573     8.11   0.000     .0188125   .0307993
    2010      |   .0234917   .003163      7.43   0.000     .0172912   .0296922
    2011      |   .0153152   .0031559     4.85   0.000     .0091286   .0215018
    2012      |   .0173182   .0031775     5.45   0.000     .0110892   .0235472
    pctell    |   .0003668   .0002937     1.25   0.212    - .0002089   .0009424
-----+-----

```

pctsped		.0009591	.0003656	2.62	0.009	.0002423	.0016758
totalenroll~t		-.0000933	9.50e-06	-9.82	0.000	-.0001119	-.0000747
pctasian		.0003216	.0004358	0.74	0.461	-.0005328	.001176
pctblack		.0002158	.0003691	0.58	0.559	-.0005078	.0009394
pcthispanic		.0004245	.0003504	1.21	0.226	-.0002624	.0011114
pctfemale		-.0009431	.0004012	-2.35	0.019	-.0017296	-.0001567
free1		.0005054	.0001067	4.74	0.000	.0002963	.0007145
redul		.0009194	.0002621	3.51	0.000	.0004055	.0014332
_cons		.2549082	.034423	7.41	0.000	.1874275	.3223888
-----							
sigma_u		.10212137					
sigma_e		.05973931					
rho		.7450423	(fraction of variance due to u_i)				

F test that all u\_i=0: F(906, 6331) = 20.55 Prob > F = 0.0000

. \_eststo lu6: xtreg lunch\_part i.bicpost i.year \$covars, fe

Fixed-effects (within) regression	Number of obs	=	7,005
Group variable: schoolid	Number of groups	=	907

R-squared:	Obs per group:	
Within = 0.1003	min =	2
Between = 0.3706	avg =	7.7
Overall = 0.3365	max =	8

corr(u_i, Xb) = -0.0196	F(17, 6081)	=	39.88
	Prob > F	=	0.0000

lunch_part	Coefficient	Std. err.	t	P> t	[95% conf. interval]		
1.bicpost	-.004206	.0052159	-0.81	0.420	-.0144309	.006019	
year							
2006	.0094849	.0030108	3.15	0.002	.0035827	.0153872	
2007	.0174954	.0030516	5.73	0.000	.0115132	.0234777	
2008	.0281332	.003202	8.79	0.000	.0218561	.0344102	
2009	.0313211	.0032831	9.54	0.000	.0248852	.0377571	
2010	.0183397	.0033532	5.47	0.000	.0117662	.0249133	
2011	.0104505	.0033547	3.12	0.002	.0038741	.0170268	
2012	.0004976	.0033807	0.15	0.883	-.0061297	.0071249	
pctell	-.0003212	.0003142	-1.02	0.307	-.0009372	.0002948	
pctsped	-.0005976	.0003886	-1.54	0.124	-.0013594	.0001643	
totalenroll~t	-.0000728	.0000101	-7.19	0.000	-.0000927	-.000053	
pctasian	.0041843	.0004575	9.15	0.000	.0032874	.0050812	
pctblack	.0042077	.0003889	10.82	0.000	.0034453	.0049701	
pcthispanic	.0036614	.0003702	9.89	0.000	.0029357	.0043871	
pctfemale	-.0011776	.0004275	-2.75	0.006	-.0020156	-.0003395	
free1	.0011066	.0001135	9.75	0.000	.0008841	.0013291	
redul	.0016304	.0002758	5.91	0.000	.0010897	.0021711	
_cons	.4648011	.03636	12.78	0.000	.3935227	.5360796	
-----							
sigma_u		.13699664					
sigma_e		.06226828					
rho		.82878041	(fraction of variance due to u_i)				

F test that all u\_i=0: F(906, 6081) = 25.38 Prob > F = 0.0000

```
.
.      // without covariates (xtdidreg)
.      xtdidregress (bkfast_part) (bicpost) , group(schoolid) time(year)
```

Treatment and time information

Time variable: year

Control: bicpost = 0

Treatment: bicpost = 1

		Control	Treatment
Group			
schoolid		800	107
Time			
Minimum		2005	2007
Maximum		2006	2012

Difference-in-differences regression  
Data type: Longitudinal

Number of obs = 7,255

(Std. err. adjusted for 907 clusters in schoolid)						
		Robust				
		std. err.	t	P> t	[95% conf. interval]	
ATET						
bicpost						
(1 vs 0)		.2282953	.0169869	13.44	0.000	.194957 .2616335

Note: ATET estimate adjusted for panel effects and time effects.

Note: Treatment occurs at different times.

```
.      xtdidregress (lunch_part) (bicpost) , group(schoolid) time(year)
```

Treatment and time information

Time variable: year

Control: bicpost = 0

Treatment: bicpost = 1

		Control	Treatment
Group			
schoolid		800	107
Time			
Minimum		2005	2007
Maximum		2008	2012

Difference-in-differences regression  
Data type: Longitudinal

Number of obs = 7,005

(Std. err. adjusted for 907 clusters in schoolid)						
		Robust				
		std. err.	t	P> t	[95% conf. interval]	
ATET						
bicpost						
(1 vs 0)		-.0022084	.0080415	-0.27	0.784	-.0179905 .0135736

Note: ATET estimate adjusted for panel effects and time effects.

Note: Treatment occurs at different times and estimation sample contains units that switch in and out of treatment.

```
.
. // with covariates (xtdidreg)
. xtdidregress (bkfast_part $covars) (bicpost) , group(schoolid) time(ye
> ar)
```

Treatment and time information

```
Time variable: year
Control:      bicpost = 0
Treatment:    bicpost = 1
```

	Control	Treatment
Group		
schoolid	800	107
Time		
Minimum	2005	2007
Maximum	2006	2012

Difference-in-differences regression  
Data type: Longitudinal

Number of obs = 7,255

(Std. err. adjusted for 907 clusters in schoolid)

	Coefficient	Robust std. err.	t	P> t	[95% conf. interval]
ATET					
bicpost					
(1 vs 0)	.2219539	.016661	13.32	0.000	.1892551 .2546526

Note: ATET estimate adjusted for covariates, panel effects, and time effects.  
Note: Treatment occurs at different times.

```
. xtdidregress (lunch_part $covars) (bicpost) , group(schoolid) time(yea
> r)
```

Treatment and time information

```
Time variable: year
Control:      bicpost = 0
Treatment:    bicpost = 1
```

	Control	Treatment
Group		
schoolid	800	107
Time		
Minimum	2005	2007
Maximum	2008	2012

Difference-in-differences regression  
Data type: Longitudinal

Number of obs = 7,005

(Std. err. adjusted for 907 clusters in schoolid)

	Coefficient	Robust std. err.	t	P> t	[95% conf. interval]
ATET					
bicpost					
(1 vs 0)	-.004206	.0079187	-0.53	0.595	-.0197472 .0113352

Note: ATET estimate adjusted for covariates, panel effects, and time effects.  
Note: Treatment occurs at different times and estimation sample contains units that switch in and out of treatment.

```
.  
.   
. // Close log and convert to PDF  
.      capture log close
```