Problem Set 5 Solutions

Question 1. This problem will use regression difference-in-differences to estimate the impact of a breakfast in the classroom (BIC) program on school meals program participation in New York City. BIC was not implemented under random assignment; rather, schools voluntarily adopted the program. We do, however, have data for these and other schools before and after adoption. (See Corcoran, Elbel, & Schwartz 2015 for details). (40 points)

(a) In Stata, open the panel dataset NYCbkfastlunch.dta. This file consists of school-level data in which the rows are elementary or middle schools observed in year t (t=2005 to 2012). The outcome variables of interest are bkfast_part and lunch_part, which are average daily participation rates in the school breakfast and lunch programs. Provide some descriptive statistics for these two variables. On what scale are they measured? (2 points)

See attached log. Each of the outcome variables ranges between 0-1 and can be interpreted as the proportion of students in attendance on the average school day in school i who received a breakfast or lunch. The average daily participation was 0.780 (or 78.0%) for lunch and only 0.258 (or 25.8%) for breakfast.

- (b) Use xtset to declare the data as a panel. Which variable is the cross-sectional unit and which is the time dimension? Is this a balanced panel? (Show how you answer the last question). How many schools are observed in all 8 years? (2 points)
 - See attached log. *schoolid* is the cross-sectional unit, while *year* is the time dimension. This is a balanced panel, as easily seen in xtdescribe which shows the pattern of data availability. 907 schools have data in all 8 years.
- (c) This dataset contains a time-varying treatment variable called *bicpost* that is equal to 1 in years in which the school offered BIC and 0 otherwise. For this part, limit the sample to schools that adopted BIC in 2010 and never-adopters. Estimate a simple "2x2" difference-in-differences regression comparing mean breakfast participation rates for two groups (treated and untreated schools) in two time periods (pre and post). Do the same for lunch participation rates, and interpret your results. (Note: ignore the panel nature of your data and just use the regress command for this part). Is the BIC effect statistically significant? Practically significant? What assumptions must be satisfied for this difference-in-differences to be considered a causal effect? (6 points)

See attached log. The do file first identifies the first year of treatment for each school and then flags schools that adopted BIC in 2010. The

2x2 DD coefficient is positive and statistically significant for breakfast, and statistically insignificant for lunch. In the case of breakfast, average daily participation increased 23.9 percentage points more in schools that adopted BIC in 2010 (the treated schools) relative to schools that never adopted BIC (the untreated schools).

For this estimate to be considered causal, the parallel trends assumption must be satisfied. That is, the pre-post change in breakfast participation among non-BIC schools must represent what would have occurred in BIC schools had they not been "treated" by adopting the program. As a panel model, the strict exogeneity assumption must also hold. This means the error term in every year t cannot be related to treatment adoption. This would be violated if there were anticipation effects, for example.

- (d) Repeat part (c) but include year dummies in the regressions. How does this affect your impact estimates for BIC, if at all? Continue to use the regress command, and explain why one of the post-2010 year effects is not estimable. (4 points)
 - See attached log. There is only a trivial change to the estimated DD coefficient in both the breakfast and lunch regressions. One of the post-2010 year effects is omitted. This is because of perfect collinearity between the *post2010* variable and the three dummies for 2010, 2011, and 2012. (If you know three of these variables, the fourth is also known).
- (e) In parts (c)-(d) you did not take advantage of the panel nature of your data. Reestimate these regressions using xtreg and school fixed effects. How do your estimates compare? (4 points)
 - See attached log. Again there is only a trivial change to the estimated DD coefficient in both regressions.
- (f) Compare the mean baseline characteristics of treated and untreated schools in parts (c)-(e). (Use 2009 as the baseline year). Look at the following: total enrollment, % ELL, % special education, % Asian, % black, % Hispanic, % female, % free lunch eligible (free1), % reduced price lunch eligible (redu1). How do schools that adopted BIC in 2010 compare to those that never adopted? (4 points)
 - See attached log. It is clear that schools which adopted BIC in 2010 differ systematically from those that didn't. They are more disadvantaged on a number of dimensions, including %ELL, %sped, and percent with family income low enough to be eligible for free meals. These differences suggest we should include these covariates in the regressions. Others—such as the race or gender composition of the school—may also help to reduce unexplained variation in the outcome.

- (g) Now estimate the same regression models in part (e), but include the school covariates listed above in your regressions. How do your estimates of the "BIC effect" change, if at all? (And how are these covariates related to meal participation?) (4 points)
 - See attached log. It appears the inclusion of covariates has little effect on the difference-in-differences estimates for breakfast and lunch participation. Many of the included covariates are related to meals program participation. For example, %ELL, %sped, and percent eligible for free meals are all positively related to breakfast participation. Total enrollment, Asian, Hispanic, and female shares are all negatively related to breakfast participation.
- (h) Repeat part (e)—the model without covariates—but try using the Stata command xtdidregress. Compare your results. (4 points)
 - See attached log. The results are identical, although the standard errors are different (unless you—correctly—adjusted your standard errors for clustering by school).
- (i) State the assumption(s) that are required to hold in order to interpret the breakfast participation DD coefficient in part (h) as causal. Use the post-estimation commands estat trendplots and estat granger to probe these assumptions, and interpret the results. (5 points)
 - Again, the parallel trends and no anticipation assumptions must be satisfied (see part c). See attached log for the post-estimation commands (for the breakfast model only). The Granger causality test rejects the null hypothesis of "no anticipation." The graphical output is below. These results are *not* consistent with a common pre-trend for the treated and untreated schools. The second plot is a Granger plot (estat grangerplot) which shows more clearly that the trends are consistent until 2009—the year before treatment. This could be an anticipation effect, or an indication that our information about the timing of treatment is incorrect for one or more schools.
- (j) Finally, incorporate <u>all</u> treated schools into the analysis regardless of when they adopted BIC. Estimate the two-way fixed effects model using xtreg (as in part e) and xtdidregress (as in part h). Try these without and with school covariates. How do your results compare? (5 points)
 - See attached log. The results are very similar to the previous ones, in all cases, when including the full set of treated schools. Note we have not addressed the potential problems introduced by staggered treatment timing.

```
. // *********************
. // ****
. // (a)
. // ****
. // Get data
        use https://github.com/spcorcor18/LPO-8852/raw/main/data/NYCbkfastlunc
> h-v2.dta, clear
        summ bkfast part lunch part
  Variable |
                 Obs
                          Mean
                                  Std. dev.
                                               Min
                                                        Max
______
bkfast_part | 7,255 .2580688 .1377646 .0054945 .9725
lunch_part | 7,005 .7803097 .1838588 .0122249 1
. // ****
. // (b)
. // ****
. // Set panel
        xtset schoolid year
Panel variable: schoolid (strongly balanced)
Time variable: year, 2005 to 2012
Delta: 1 unit
       xtdescribe
schoolid: 1, 2, ..., 1179
year: 2005, 2006, ..., 2012
                                                           907
                                                   n =
                                                   T =
         Delta(year) = 1 unit
         Span(year) = 8 periods
         (schoolid*year uniquely identifies each observation)
                                                 /5% 95%
8
                                                            max
                       5% 25% 50%
8 8 8
                                         JU∜
8
Distribution of T i:
                   min
                                                               8
                    8
    Freq. Percent Cum. | Pattern
    907 100.00 100.00 | 11111111
 ----+----
    907 100.00 | XXXXXXXX
. // ****
. // (c)
. // Simple 2x2 DD for limited sample
        tabulate year bicpost
```

| =1 if post period for BIC school year İ 0 1 | Total ____ 907 0 | 907 0 | 906 1 | 902 5 | 878 29 | 848 59 | 828 79 | 800 107 | 2005 I 907 907 907 907 907 907 907 2006 | 2007 j 2008 2009 | 2010 2011 | 2012 | 107 Total | 6,976 280 | 7,256

// identify first year that school was treated
. egen temp1=min(year) if bicpost==1, by(schoolid)
(6,976 missing values generated)

. egen firstyear=max(temp1), by(schoolid)
(6,400 missing values generated)

. drop temp1

table year firstyear

	2007	2008	2009	firstye 2010	ear 2011	2012	Total
year 2005 2006 2007 2008 2009 2010 2011 2012 Total	1 1 1 1 1 1 1 1 1 8	4 4 4 4 4 4 4 32	24 24 24 24 24 24 24 24 24	30 30 30 30 30 30 30 30 30 30	20 20 20 20 20 20 20 20 20 20	28 28 28 28 28 28 28 28 28	107 107 107 107 107 107 107 107 856

```
// simple DD: treated schools adopting in 2010 vs never adopted
// create "post" period variable
gen post2010 = (year>=2010 & year~=.)

// create "ever treated" variable
gen everbic2010 = firstyear==2010
```

__eststo bk1: reg bkfast_part i.everbic2010##i.post2010 if (firstyear==
> 2010 | firstyear==.)

Source	SS	df	MS	Number of obs	=	6,639
Model	7.27543865		2.42514622	F(3, 6635) Prob > F	=	152.58
	105.458173	-	.015894224	R-squared	=	0.0645
+				Adj R-squared	=	0.0641
Total	112.733612	6,638	.016983069	Root MSE	=	.12607

bkfast_part	Coefficient	Std. err.	t	P> t	[95% conf.	interval]
1.everbic2010 1.post2010	.0395468 .0060172	.010485 .0032553	3.77 1.85	0.000	.0189927 0003643	.0601008
everbic2010# post2010 1 1	.2386908	.0171219	13.94	0.000	.2051262	.2722553
_cons	.2463439	.0019936	123.57	0.000	.2424358	.2502521
0.015	1+0 1111 rog 1	unch nart	i owarhia	2010##;	post2010 if (f	

. __eststo lu1: reg lunch_part i.everbic2010##i.post2010 if (firstyear==2
> 010 | firstyear==.)

Source	SS	df	MS		er of obs	=	6,423 10.27
Model Residual	1.07347052 223.584325		.357823508 .034831644	F(3, 6419) = Prob > F = R-squared = Adj R-squared = Root MSE =		=	0.0000 0.0048 0.0043
Total	224.657795	6,422	.034982528				.18663
lunch_part	Coefficient	Std. err.	t	P> t	[95% (conf.	interval]
1.everbic2010 1.post2010	.0664884 006335	.016285 .0048792	4.08 -1.30	0.000 0.194	.0345		.0984124
everbic2010# post2010							
1 1	.0057016	.0260823	0.22	0.827	04542	284	.0568315
_cons	.7758114	.0030141	257.39	0.000	.7699	027	.7817201

. // ****
. // (d)
. // ****
. // Add year effects

Source | SS

. _eststo bk2: reg bkfast_part i.everbic2010##i.post2010 i.year if (firs > tyear== $2\overline{0}10$ | firstyear==.) note: 2012.year omitted because of collinearity.

df MS Number of obs = 6,639

Model Residual Total	8.11373426 104.619878 112.733612	6,629 6,638	.901526029 .015782151	F(9, 6 Prob 2 R-squa Adj R- Root N	F ared -squared	= = = =	57.12 0.0000 0.0720 0.0707 .12563
bkfast_part	Coefficient	Std. err.	t	P> t	[95% (conf.	interval]
1.everbic2010 1.post2010 everbic2010# post2010	İ	.010448	3.79 3.27	0.000	.0190	274	.0600321
1 1 year 2006 2007 2008 2009 2010 2011 2012	.2386869 .0019981 .011271 .0308351 .0326727 .0060077 0027192	.0170615 .0061686 .0061686 .0061686 .0061668 .0061668 (omitted)	13.99 0.32 1.83 5.00 5.30 0.97 -0.44	0.000 0.746 0.068 0.000 0.000 0.330 0.659	01009 00083 .0187 .02058 00608	944 216 425 802 812	.0140906 .0233635 .0429276 .0447652 .0180966 .0093697

_cons | .2309847 .0043796 52.74 0.000 .2223994 .2395701

eststo lu2: reg lunch_part i.everbic2010##i.post2010 i.year if (first
> year==2010 | firstyear==.)
note: 2012.year omitted because of collinearity.

Source	SS	df	MS		er of obs	=	6,423 4.98
Model Residual	1.55895067 223.098845		.173216741	F(9, 6413) Prob > F R-squared Adj R-squared		=	0.0000 0.0069 0.0055
Total	224.657795	6,422	.034982528	Root		_ 	.18652
lunch_part	Coefficient	Std. err.	t	P> t	[95% (conf.	interval]
1.everbic2010 1.post2010	.0666754 0026473	.0162752	4.10 -0.28	0.000 0.777	.0347		.0985801
everbic2010# post2010 1 1	.0057818	.0260666	0.22	0.824	0453	174	.0568811
year 2006 2007 2008 2009 2010 2011 2012	.0092882 .01739 .0228403 .0235812 .0203108 .0121711	.0093145 .0093231 .009385 .0094064 .0092229 .009234 (omitted)	1.00 1.87 2.43 2.51 2.20 1.32	0.319 0.062 0.015 0.012 0.028 0.188	0089 0008 .0044 .0051 .0022 0059	864 426 416 309	.0275477 .0356663 .0412379 .0420208 .0383908 .0302729
_cons	.7612627 	.0066357	114.72	0.000	.7482	546	.7742708

· // **** · // (e) · // ****

. // Use xtreg and school fixed effects

Fixed-effects (within) regression Group variable: schoolid

. __eststo bk3: xtreg bkfast_part i.bicpost i.year if (firstyear==2010 |
> firstyear==.), fe

Number of obs = 6,639 Number of groups = 830

R-squared:	Obs per group:		
Within = 0.2143 Between = 0.0414 Overall = 0.0697	min avg max	=	8.0 8
corr(u i, Xb) = 0.0349	F(8, 5801) Prob > F	=	197.80

bkfast_part	Coefficient	Std. err.	t	P> t	[95% conf.	interval]
1.bicpost	.2386411	.0070527	33.84	0.000	.2248152	.2524671
year 2006 2007 2008 2009 2010 2011 2012	.0022187 .0114915 .0310556 .0328933 .0265101 .0177832 .0205024	.00255 .00255 .00255 .00255 .0025628 .0025628	0.87 4.51 12.18 12.90 10.34 6.94 8.00	0.384 0.000 0.000 0.000 0.000 0.000	0027804 .0064925 .0260566 .0278942 .0214861 .0127593 .0154785	.0072177 .0164905 .0360547 .0378923 .0315341 .0228072 .0255264

```
_cons | .2322215 .0018037 128.74 0.000 .2286855 .2357575
    sigma_u | .11608572
     sigma_e | .05193033
rho | .8332519 (fraction of variance due to u_i)
F test that all u i=0: F(829, 5801) = 39.90
                                                Prob > F = 0.0000
          eststo lu3: xtreq lunch part i.bicpost i.year if (firstyear == 2010 | f
> irstyear==.), fe
                                                                    6,423
Fixed-effects (within) regression
                                             Number of obs =
Group variable: schoolid
                                             Number of groups =
                                                                       830
R-squared:
                                              Obs per group:
     Within = 0.0324
                                                          min =
     Between = 0.0216
                                                           avq =
                                                                        7.7
     Overall = 0.0023
                                                           max =
                                              F(8, 5585)
                                                                     23.41
                                                               =
                                              Prob > F
                                                                    0.0000
corr(u i, Xb) = -0.0113
______
 lunch part | Coefficient Std. err. t P>|t| [95% conf. interval]
 1.bicpost | .0013516 .0090263 0.15 0.881 -.0163435 .0190467
        year |
      cons | .7647871 .0022679 337.22 0.000
                                                       .7603412 .7692331
_____<del>_</del>____
    sigma_u | .1764832
sigma_e | .06381114
rho | .88438184 (fraction of variance due to u_i)
F test that all u i=0: F(829, 5585) = 59.53
                                                      Prob > F = 0.0000
. // ****
. // (f)
. // ****
. // Descriptive statistics for BIC 2010 adopters and those who never adopt BIC
         sum totalenrollment- pctfemale free1 redu1 if firstyear==2010 & year==
> 2009, sep(0)
   Variable | Obs Mean Std. dev. Min Max
totalenrol~t | 30 628.2 263.4843 246 1244 pctell | 30 14.79 11.17585 1.1 44 pctsped | 30 18.31 7.017483 5.7 36.5 pctasian | 30 2.7 4.906294 0 20.8 pctblack | 30 45.85 29.04892 1.7 95.2 pcthisp | 30 46.34 27.65217 3.5 91.8 pctwhite | 30 4.046667 11.81582 .2 64.6 pctfemale | 30 49 3.054336 43.3 54.9 freel | 30 82.38852 12.33706 33.07393 100 redul | 30 7.525835 3.735603 .8522727 15.38461
```

. sum totalenrollment- pctfemale free1 redu1 if firstyear==. & year==200
> 9, sep(0)

Variable	Obs	Mean	Std. d	ev.	Min	Ma	ax
totalenrol~t pctell pctsped pctasian pctblack pcthisp pctwhite pctfemale freel redul	800 800 800 800 800 800 800 800	689.1938 12.53675 15.92213 14.09225 30.62975 37.9295 16.33788 49.09438 67.48113 10.20999	321.87 10.900 5.7694 18.712 29.664 26.026 22.338 3.9381 22.881 5.0233	83 76 98 12 49 98 39 77 5.6	110 0 0 0 0 1.6 0 36.2 7201 0	21. 77 91 96 98 92 10 26.297	.3 44 .8 .8 .7 .8 00
. globa	ects regressic			rollment	pctasia	an petb	lack pcthis
> p"							
. globa	al covars "\$co	vars pctfem	ale free	1 redu1"			
<pre>. _eststo bk4: xtreg bkfast_part i.bicpost \$covars if (firstyear==2010 > firstyear==.),fe</pre>							
Fixed-effects Group variable		ssion		Number o			6,639 830
R-squared: Within = Between = Overall =	0.2035			Obs per	n	nin = avg = nax =	7 8.0 8
corr(u_i, Xb) =	= 0.0871			F(10, 57 Prob > F		= =	158.69 0.0000
bkfast_part	 Coefficient	Std. err.	t	 P> t	 [95%	conf.	interval]
1.bicpost pctsped totalenroll~t pctasian pctblack pcthisp pctfemale freel redu1 _cons	.0005993 .0020669 000104 .0009075 0000417 .0003415	.006981 .0002543 .0003058 8.65e-06 .0003813 .0003147 .000309 .0003687 .0000988 .0002282 .0303249	33.60 2.36 6.76 -12.02 2.38 -0.13 1.11 -1.48 2.45 1.77 8.78	0.000 0.018 0.000 0.017 0.895 0.269 0.138 0.014 0.077 0.000	.000 .001 000 000 001 .000	2643	.2482268 .0010978 .0026665 000087 .0016549 .0005753 .0009473 .0001758 .0004358 .0004358
sigma_u sigma_e rho	.10614395 .05192171 .80691974	(fraction	of varia	nce due t	o u_i)		
F test that all	l u_i=0: F(829	, 5799) = 2	6.88		Pı	rob > F	= 0.0000

```
> firstyear==.),fe
                                         Number of groups = 6,423
Fixed-effects (within) regression
Group variable: schoolid
R-squared:
                                         Obs per group:
    Within = 0.0777
                                                     min =
                                                               7.7
                                                     avg =
    Between = 0.3368
    Overall = 0.3055
                                                     max =
                                         F(10, 5583) = 47.01
Prob > F = 0.0000
corr(u_i, Xb) = -0.0535
                                         Prob > F
  lunch part | Coefficient Std. err. t P>|t| [95% conf. interval]
sigma_u | .1437498
sigma_e | .06231349
rho | .84181471 (fraction of variance due to u_i)
_____<del>_</del>___<del>_</del>___
F test that all u i=0: F(829, 5583) = 25.69 Prob > F = 0.0000
. // ****
. // (h)
. // ****
. // try using xtdidregress
        xtdidregress (bkfast part) (bicpost) if (firstyear==2010 | firstyear==
> .), group(schoolid) time(year)
Treatment and time information
Time variable: year
Control: bicpost = 0
Treatment: bicpost = 1
          | Control Treatment
Group
  schoolid | 800
   Minimum |
Maximum |
                 2005
                2005 2010 2010
-----
```

Number of obs = 6,639

Difference-in-differences regression

Data type: Longitudinal

eststo lu4: xtreg lunch part i.bicpost \$covars if (firstyear==2010 |

```
(Std. err. adjusted for 830 clusters in schoolid)
                           Robust
bkfast part | Coefficient std. err.
                                        t P>|t| [95% conf. interval]
    bicpost |
  (1 vs 0) | .2386411 .0265832 8.98 0.000 .1864628 .2908195
Note: ATET estimate adjusted for panel effects and time effects.
Note: Treatment occurs at different times.
         xtdidregress (lunch_part) (bicpost) if (firstyear==2010 | firstyear==.
> ), group(schoolid) time(year)
Treatment and time information
Time variable: year
Control: bicpost = 0
Treatment: bicpost = 1
         | Control Treatment
-----
Group
                  800
  schoolid |
Time
   Minimum | 2005 2010
Maximum | 2008 2012
Difference-in-differences regression
                                                       Number of obs = 6,423
Data type: Longitudinal
                           (Std. err. adjusted for 830 clusters in schoolid)
                           Robust
lunch part | Coefficient std. err.
                                        t P>|t| [95% conf. interval]
ATET
    bicpost |
  (1 vs 0) | .0013516 .013313 0.10 0.919 -.0247795 .0274827
Note: ATET estimate adjusted for panel effects and time effects.
Note: Treatment occurs at different times and estimation sample contains units
     that switch in and out of treatment.
. // ****
. // (i)
. // ****
. // post-estimation commands
         xtdidregress (bkfast part) (bicpost) if (firstyear == 2010 | firstyear ==
> .), group(schoolid) time(year)
Treatment and time information
Time variable: year
Control: bicpost = 0
Treatment: bicpost = 1
           | Control Treatment
Group
                  800
  schoolid |
Time
   Minimum | 2005 2010
Maximum | 2006 2010
```

2000 07100 -00	igicaarnar	(Std err	adiuste	d for 830) clusters in	schoolid)
bkfast_part	 Coefficient 	Robust std. err.	t	P> t	[95% conf.	interval]
ATET	.2386411					
	timate adjuste nt occurs at d			and time	effects.	
. esta	at trendplots					
. grap	oh export "tre ts.png saved a	ndplots.png s PNG forma	", as(png t	g) name("G	Graph") repla	ce
. esta	at granger					
Granger causal HO: No effect	lity test in anticipati	on of treat	ment			
F(4, 829) = Prob > F = 0						
	see also grang at grangerplot	erplot comm	and			
	oh export "gra: Lot.png saved			ng) name("	'Graph") repl	ace
. // v	ate all school	ates (xtreg	-)	oost i vos	er fo	
_	(within) regree: schoolid	_	ic i.bicp	Number o	of obs = of groups =	7 , 255
R-squared: Within = Between = Overall =	= 0.0667			Obs per	group: min = avg = max =	8.0
corr(u_i, Xb)	= 0.0196			F(8, 634 Prob > F		353.93 0.0000
bkfast_part	Coefficient	Std. err.	t	P> t	[95% conf.	interval]
1.bicpost	.2282953	.0048888	46.70	0.000	.2187115	.2378791
year 2006 2007 2008 2009 2010 2011 2012	.0028761 .0116797 .0307599 .0306009 .0297214 .020442 .0219312	.0028522 .0028522 .0028523 .0028565 .0028699 .0028838 .0029099	1.01 4.10 10.78 10.71 10.36 7.09 7.54	0.313 0.000 0.000 0.000 0.000 0.000 0.000	0027151 .0060884 .0251684 .0250012 .0240955 .0147888 .0162268	.0084673 .0172709 .0363514 .0362005 .0353473 .0260953 .0276357
_cons	.230754	.0020174	114.38	0.000	.2267992	.2347088

```
sigma_u | .11573186
     sigma_e | .06071925
rho | .78415233 (fraction of variance due to u_i)
         ------
F test that all u i=0: F(906, 6340) = 29.04
                                                        Prob > F = 0.0000
           eststo lu5: xtreg lunch part i.bicpost i.year, fe
                                                         Number of obs = 7,005
Number of groups = 907
Fixed-effects (within) regression
Group variable: schoolid
R-squared:
                                                         Obs per group:
     Within = 0.0305
                                                            min =
     Between = 0.0357
                                                                          avg =
      Overall = 0.0018
                                                                          max =
                                                         F(8, 6090)
                                                                                       23.95
corr(u i, Xb) = -0.0168
                                                         Prob > F
1.bicpost | -.0022084 .0053663 -0.41 0.681 -.0127282 .0083114
         year |

        year
        2006
        .0060649
        .0030923
        1.96
        0.050
        2.94e-06
        .0121268

        2007
        | .0150901
        .0030987
        4.87
        0.000
        .0090156
        .0211645

        2008
        | .024707
        .0031243
        7.91
        0.000
        .0185822
        .0308317

        2009
        | .0265117
        .0031399
        8.44
        0.000
        .0203565
        .0326669

        2010
        | .0108692
        .0031003
        3.51
        0.000
        .0047915
        .016947

        2011
        | .0040942
        .0031218
        1.31
        0.190
        -.0020256
        .0102141

        2012
        | -.0060021
        .0031511
        -1.90
        0.057
        -.0121793
        .0001751

_cons| .7703875 .0021972 350.63 0.000 .7660803 .7746947
     sigma_u | .17289228
     sigma_e | .06459134
rho | .87752268 (fraction of variance due to u i)
_______________
F test that all u i=0: F(906, 6090) = 55.50 Prob > F = 0.0000
          // with covariates (xtreg)
           _eststo bk6: xtreg bkfast_part i.bicpost i.year $covars, fe
                                                         Number of obs = 7,255
Number of groups = 907
Fixed-effects (within) regression
Group variable: schoolid
R-squared:
                                                         Obs per group:
                                                                        min =
     Within = 0.3318
                                                                          avg =
                                                                                       8.0
      Between = 0.2925
      Overall = 0.2963
                                                                         max =
                                                        F(17, 6331) = 184.93

Prob > F = 0.0000
corr(u i, Xb) = 0.1284
 bkfast_part | Coefficient Std. err. t P>|t| [95% conf. interval]
_____
   1.bicpost | .2219539 .0048512 45.75 0.000 .2124438 .2314639
           vear |
         pctell | .0003668 .0002937 1.25 0.212 -.0002089
```

.0009424

```
31) = 20.55
F test that all u i=0: F(906, 6331) = 20.55
                                    Prob > F = 0.0000
      eststo lu6: xtreg lunch part i.bicpost i.year $covars, fe
                             Number of obs = 7,005
Fixed-effects (within) regression
Group variable: schoolid
                             Number of groups =
R-squared:
                             Obs per group:
                                    min = 2
avg = 7.7
max = 8
   Within = 0.1003
   Between = 0.3706
   Overall = 0.3365
                                       = 39.00
0.0000
                             F(17, 6081)
Prob > F
                                       =
corr(u i, Xb) = -0.0196
______
 lunch_part | Coefficient Std. err. t P>|t| [95% conf. interval]
 1.bicpost | -.004206 .0052159 -0.81 0.420 -.0144309 .006019
     year |
    sigma_u | .13699664
sigma_e | .06226828
rho | .82878041 (fraction of variance due to u_i)
```

Prob > F = 0.0000

F test that all u i=0: F(906, 6081) = 25.38

Difference-in-differences regression Number of obs = 7,255

Data type: Longitudinal

(Std. err. adjusted for 907 clusters in schoolid)

bkfast_part	 Coefficient	Robust std. err.	t	P> t	[95% conf.	interval]
ATET bicpost (1 vs 0)		.0169869	13.44	0.000	.194957	.2616335

Note: ATET estimate adjusted for panel effects and time effects.

Note: Treatment occurs at different times.

. xtdidregress (lunch_part) (bicpost) , group(schoolid) time(year)

Treatment and time information

Time variable: year

Control: bicpost = 0
Treatment: bicpost = 1

iicacmene.	Diopose	_
	Control	Treatment
Group schoolid	800	107
Time Minimum Maximum	 2005 2008	2007

Difference-in-differences regression

Data type: Longitudinal

Number of obs = 7,005

		(Std. err.	adjuste	d for 907	clusters in	schoolid)
lunch_part	 Coefficient	Robust std. err.	t	P> t	[95% conf.	interval]
ATET bicpost (1 vs 0)	I .	.0080415	-0.27	0.784	0179905	.0135736

Note: ATET estimate adjusted for panel effects and time effects.

Note: Treatment occurs at different times and estimation sample contains units that switch in and out of treatment.

```
// with covariates (xtdidreg)
        xtdidregress (bkfast part $covars) (bicpost) , group(schoolid) time(ye
Treatment and time information
Time variable: year
Control: bicpost = 0
Treatment: bicpost = 1
_____
             Control Treatment
Group
  schoolid | 800
                          107
   Minimum | 2005 2007
Maximum | 2006 2012
Difference-in-differences regression
                                                 Number of obs = 7,255
Data type: Longitudinal
                         (Std. err. adjusted for 907 clusters in schoolid)
                         Robust
bkfast part | Coefficient std. err. t P>|t| [95% conf. interval]
ATET
    bicpost |
   (1 vs 0) |
               .2219539 .016661 13.32 0.000
                                                  .1892551
                                                             .2546526
______
Note: ATET estimate adjusted for covariates, panel effects, and time effects.
Note: Treatment occurs at different times.
        xtdidregress (lunch part $covars) (bicpost) , group(schoolid) time(yea
> r)
Treatment and time information
Time variable: year
Control: bicpost = 0
Treatment: bicpost = 1
          | Control Treatment
 schoolid | 800
                         107
Time
   Minimum | 2005 2007
Maximum | 2008 2012
Difference-in-differences regression
                                                Number of obs = 7,005
Data type: Longitudinal
                         (Std. err. adjusted for 907 clusters in schoolid)
Robust
  lunch_part | Coefficient std. err. t P>|t| [95% conf. interval]
ATET
    bicpost |
   (1 vs 0) | -.004206 .0079187 -0.53 0.595 -.0197472
Note: ATET estimate adjusted for covariates, panel effects, and time effects.
Note: Treatment occurs at different times and estimation sample contains units
    that switch in and out of treatment.
```

. . // Close log and convert to PDF . capture log close

Graphical diagnostics for parallel trends



