# Equation-free analysis of agent-based models: systematic parameter determination

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#### 1 Time horizon $\tau$

For deterministic systems, an ensemble of microscopic models initialised at the same state, will all be in the same state after a simulation time of  $\tau$ . Stochastic systems however, will result in distribution of final states with a mean  $\bar{x}$  and variance  $\sigma^2 \neq 0$ . The value of  $\sigma^2$  will vary due to the dynamics and the noise of the system, as may  $\bar{x}$ , and thus depends on the size of the time evolution. For increasing time windows,  $\sigma^2$  may either;

- 1. increase indefinitely with  $\tau$  (diffusion in an open system),
- 2. increase to a point then remain approximately constant (point of balance or symmetry in the system), or
- 3. increase to a maximum point then begin to reduce (transition from one state to another).

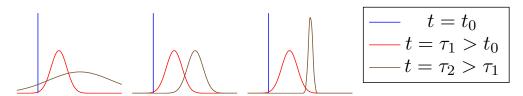


Figure 1: Dependence of  $\sigma^2$  on  $\tau$ . For increasing  $\tau$ ,  $\sigma^2$  increases; indefinitely (left), to a constant (mid), or to a maximum before reducing (right).

This is illustrated in Fig. 1. For an arbitrary system, the distribution of an ensemble of simulation, the realisations, at  $\tau$  is unknown, and although we can simply calculate  $\bar{x}$ , we can not calculate  $\sigma^2$  in general. We can however average a number of ensembles resulting in a distribution of  $\bar{x}$  which, according to the central limit theorem (CLT) is normally distributed. From the distribution of  $\bar{x}$  we can calculate a mean  $\bar{x}$  (the mean of the means) and a standard deviation  $\omega$ . As increasing  $\tau$  leads to an increase, at least initially, in  $\bar{x}$  and  $\sigma^2$ , it will also effect  $\bar{x}$  and  $\omega$ . By defining a a minimum threshold  $\gamma$  for the value of  $\omega$ , which reflects the amount of variation across the ensembles, we can obtain a value of  $\tau$  that enables sufficient time evolution such that the dynamic behaviour in the

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model is observable. That is, increase  $\tau$  until  $\omega \geq \gamma$ . This provides a platform for the systematic determination of  $\tau$  for a specific model, avoiding excessive trial and error testing, or the need of a deterministic equivalent. This process is illustrated in Fig. 2 and the procedure is summarised in Algorithm 1.

#### **Algorithm 1** Time horizon predictor.

```
# obtain initial parameter settings
N = N_{int} \# \text{ initial realisations (default=10)}
\tau = \tau_{int} \# \text{ initial time horizon}
\delta \tau # time horizon increment
M \# number of runs to average over
MaxIter # maximum number of iterations
\gamma # threshold for dynamics
while iterations < MaxIter do
    for j = 0; j < M do
        LIFT(N)
        SIMUALTE(N, \tau)
       \bar{\mathbf{x}}[j] = RESTRICT(N)
       j := j+1
    end for
    # 95 % confidence interval of distribution of means
   \omega_i = 95 \text{ConfInt}(\bar{\mathbf{x}})
    if \omega > \gamma then
        EXIT # \tau large enough to see some dynamics
    end if
    \tau := \tau * \delta \tau
    iterations := iterations + 1
end while
```

# 2 Number of microscopic realisations N

Once we have obtained a value for the time evolution of the microscopic simulations we can then obtain a prediction for the number of microscopic realisations required to perform EF continuation of a generic ABM. The time horizon predictor in Algorithm 1 provides both a prediction of  $\tau$  and a measure of how well defined this value is through  $\omega_N$ , which we can use as our threshold in the realisation predictor. That is, we can again use the CLT to obtain a (Gaussian) distribution of means and calculate the mean of means, variance and confidence interval. Here we can use the value of  $\omega$  from the time horizon predictor as a limit for the variance of the distribution of means for the realisations, i.e.  $\omega_N$ . This process is summaries in Algorithm 2.

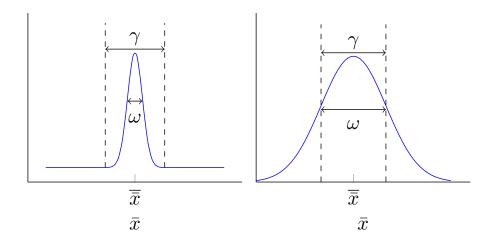


Figure 2: Variance of the distribution of means,  $\omega$  with  $\gamma$  limit. For small  $\tau$  (left)  $\omega < \gamma$ , whereas for larger  $\tau$  (right)  $\omega \approx \gamma$ .

```
Algorithm 2 Realisation predictor.
```

```
# obtain initial parameter settings
N = N_{int} \# initial realisations
\delta N # increment in N
M \# number of runs to average over
N_{Max} # maximum number of realisations
\tau # time horizon from Algorithm 1
\gamma # threshold for dynamics from Algorithm 1
while \omega_N > \gamma do
   for j = 0; j < M do
       LIFT(N)
       SIMUALTE(N, \tau)
       \bar{\mathbf{x}}[j] = RESTRICT(N)
       j := j+1
   end for
    # 95 % confidence interval of distribution of means
   \omega_N = 95 \text{ConfInt}(\bar{\mathbf{x}})
    N := N * \delta N
end while
```

## 3 Continuation step size $\delta s$ and finite differencing

We require a step size  $(\delta s)$ , in the continuation parameter  $\lambda$  such that if  $f(x^*,\lambda) = x^*$ , then  $f(x^*,\lambda+\delta s) \approx x^*$ . That is, we require  $\delta s$  to be small enough so that we obtain a suitable prediction to f under this perturbation. However this cannot be arbitrary small due to finite computational resources, impractical run times, and will depend on the specific problem under investigation. As with N and  $\tau$ , for  $\delta s$  we can determine a value based directly on simulation of the system. In this case by examining the level for stochasticity through the variance using the parameter obtained with Algorithms 1 and 2. We perform the *Lift* and *Simulation* operations for one set of realisations and calculate the standard error in the mean  $(\Delta)$  of the distribution of simulation results. We then

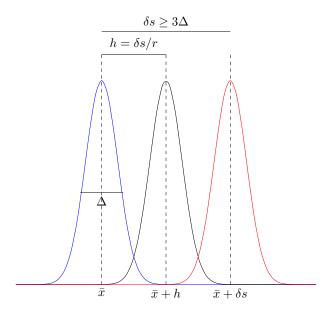


Figure 3: Illustration of the  $\delta s$  predictor. The standard error ( $\Delta$ ) from the the simulation of N realisations (blue) is used to determine the step size for finite differencing (black) and continuation step size (red).

simply define the step size as

$$\delta s = 3\Delta . (3.1)$$

During the predictor-corrector continuation used here, it is common to employ a Newton-Rhapson corrector method. In order to determine the update step size in this procedure we use finite differencing to calculate the Jacobian J,

$$J_{ik} = \frac{f_i(x+h) - f_k(x)}{h} , \qquad (3.2)$$

for i, k = 0...m, where m is the number of orders in f. The result of Eq. (3.2) is an  $m \times m$  matrix. The value h is also problem specific as the accuracy of the determination of J will depend on the magnitude of h compared to x. To ensure that J is accurately calculated for a general system, we define the value of h in terms of the continuation step size,

$$h = \frac{\delta s}{r} \,\,\,\,(3.3)$$

where r=2 in this work.

# 4 Iterative determination and the selection of $\gamma$

It is worth noting that the value of  $\gamma$  in Algorithm 1 and 2 is technically a tunable, problem specific parameter. However, here we have replaced three problem specific parameters with a single parameter, thus reducing the overhead. We note however that  $\gamma$  can easily be obtained for a arbitrary model. By perform the procedure in Algorithm 1 with an arbitrarily large  $\gamma$  (say 10,000) one can monitoring  $\omega$  as a function of  $\tau$ . The profile of this curve can be used to select a new value of  $\gamma$  for the specific model and Algorithm 1 can be repeated. Figure 4 shows the dependence of  $\omega$  on  $\tau$  when using an arbitrarily large  $\gamma$ . Here we can see that we can use these curves to obtain a better guess for  $\gamma$  and obtain our problem specific equation-free parameters.

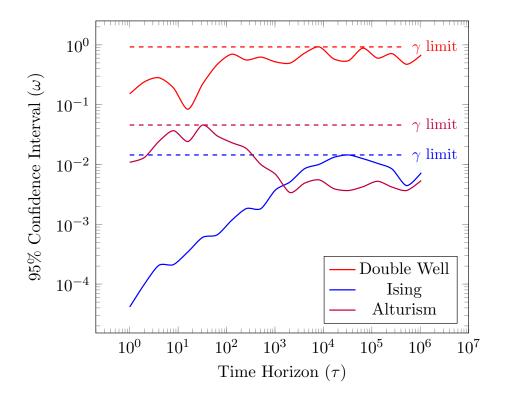


Figure 4: Confidence interval on mean of g(x) for 3 examples from the Repository of Models showing the maximum value of  $\gamma$  in each case.

## Acknowledgements

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