Soumya Sara John

M.Tech DSP
Department of Avionics
Indian Institute of Space Science and Technology

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Outline

Bayesian Deep Learning Introduction Bayesian Regression

Sampling the posterior Markov Chain Monte Carlo

Approximating the posterior Laplace Approximation Variational Inference

Talking About BDL Advantages Challenges

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Introduction

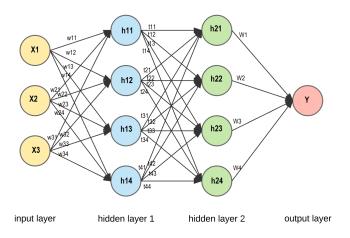


Figure: Two layer DNN : weights are point estimates

Introduction

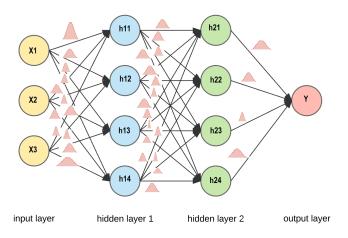


Figure: Two layer DNN: weights are defined using Gaussian distributions

Introduction

Bayes Probability

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$$P(B/A) = \frac{P(A/B)P(B)}{P(A)} \tag{1}$$

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Linear Regression

▶ Dataset :
$$\{\mathbf{x}_i, y_i\}_{i=1}^N$$

 $y(\mathbf{x}) = \mathbf{x}^T \mathbf{w} + \epsilon(\mathbf{x})$
 $\mathbf{w} \in R^d \rightarrow \text{parameters}$
 $\epsilon(\mathbf{x}) \rightarrow \text{residuals}$

• Ordinary Least squares: $\hat{\mathbf{w}} = \underset{\mathbf{w}}{\operatorname{argmin}} \sum_{i=1}^{N} (\mathbf{x}_{i}^{\mathsf{T}} \mathbf{w} - y_{i})^{2}$

$$\Rightarrow \hat{\mathbf{w}} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T y$$

Overfitting might occur

Adding a regularization term

$$\hat{\mathbf{w}} = \underset{\mathbf{w}}{\operatorname{argmin}} \sum_{i=1}^{N} (\mathbf{x}_{i}^{T} \mathbf{w} - y_{i})^{2} + \lambda ||\mathbf{w}||_{2}^{2}$$
$$\Rightarrow \hat{\mathbf{w}} = (\mathbf{X}^{T} \mathbf{X} + \lambda I)^{-1} \mathbf{X}^{T} y$$

$$y(\mathbf{x}) = \mathbf{x}^T \mathbf{w} + \epsilon(\mathbf{x})$$

$$p(\epsilon|\sigma^2) = N(\epsilon; 0, \sigma^2 I)$$

$$p(\mathbf{w}|\mu, \Sigma) = N(\mathbf{w}; \mu, \Sigma)$$

$$\begin{aligned} \mathbf{p}(\mathbf{x}) &= \mathbf{x}^T \mathbf{w} + \epsilon(\mathbf{x}) \\ p(\epsilon | \sigma^2) &= N(\epsilon; 0, \sigma^2 I) \\ p(\mathbf{w} | \mu, \Sigma) &= N(\mathbf{w}; \mu, \Sigma) \\ p(y_i | \mathbf{x}_i, \mathbf{w}) &= \frac{1}{\sqrt{2\pi\sigma^2}} e^{\frac{(y_i - \mathbf{x}_i^T \mathbf{w})^2}{2\sigma^2}} \end{aligned}$$

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- ▶ Likelihood : $p(\mathbf{Y}|\mathbf{X}, \mathbf{w}) = \prod_{i=1}^{N} p(y_i|\mathbf{x}_i, \mathbf{w})$
- ► Prior : p(w)

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- ► Prior : p(**w**)
- ▶ Posterior : $p(\mathbf{w}|\mathbf{X}, \mathbf{Y}) = \frac{p(\mathbf{Y}|\mathbf{X}, \mathbf{w})p(\mathbf{w})}{\int p(\mathbf{Y}|\mathbf{X}, \mathbf{w})p(\mathbf{w})d\mathbf{w}}$

- ► Likelihood : $p(\mathbf{Y}|\mathbf{X}, \mathbf{w}) = \prod_{i=1}^{N} p(y_i|\mathbf{x}_i, \mathbf{w})$ ► Prior : $p(\mathbf{w})$ ► Posterior : $p(\mathbf{w}|\mathbf{X}, \mathbf{Y}) = \frac{p(\mathbf{Y}|\mathbf{X}, \mathbf{w})p(\mathbf{w})}{\int p(\mathbf{Y}|\mathbf{X}, \mathbf{w})p(\mathbf{w})d\mathbf{w}}$
- $\hat{\mathbf{w}} = \underset{\mathbf{w}}{\operatorname{argmax}} p(\mathbf{w}|\mathbf{X}, \mathbf{Y})$

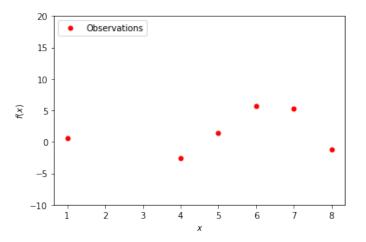


Figure: Observation Points from $f(x) = x\cos(x)$

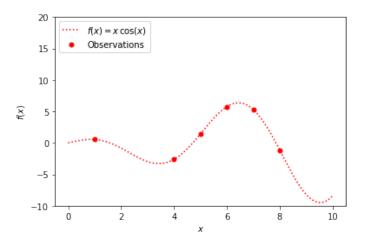


Figure: $f(x) = x\cos(x)$

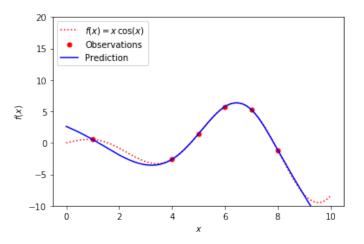


Figure: $f(x) = x\cos(x)$ and predicted function using MLE

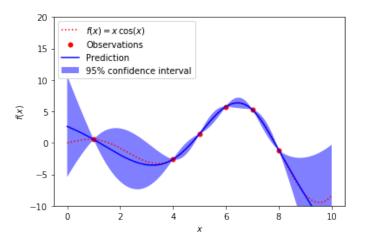


Figure: With uncertainty range

Bayesian Regression - Function Space View

▶ A function f(x) defined such that $p(f) \sim GP(f; \mu, K)$

$$\mu(\mathbf{X}) = [f(\mathbf{X})]$$
 and $K(\mathbf{x}_i, \mathbf{x}_j) = e^{-0.5||\mathbf{x}_i - \mathbf{x}_j||_2^2}$

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▶ Eg: 3 points X_1, X_2, X_3 and

correspondingly,
$$\begin{bmatrix} f_1 \\ f_2 \\ f_3 \end{bmatrix} = \textit{N} \left(\begin{bmatrix} 0 \\ 0 \\ 0 \end{bmatrix}, \begin{bmatrix} K_{11} & K_{12} & K_{13} \\ K_{21} & K_{22} & K_{23} \\ K_{31} & K_{32} & K_{33} \end{bmatrix} \right)$$

Bayesian Regression - Function Space View

- ▶ Posterior : $p(f|D) = \frac{p(D|f)p(f)}{\int p(D|f)p(f)df}$

Bayesian Regression - Function Space View

- ▶ Posterior : $p(f|D) = \frac{p(D|f)p(f)}{\int p(D|f)p(f)df}$
- $\hat{f} = \underset{f}{\operatorname{argmax}} p(f|D)$
- ▶ Likelihood : $p(D|f) \Rightarrow p(\mathbf{Y}|f) = \prod_{i=1}^{N} p(y_i|f_i)$
- $p(y_i = 1|f_i) = \sigma(f_i)$

Bayesian Regression - Function Space View

- And for a new point x^* , get $f(x^*) = f^*$ using the conditional distribution
- Joint distribution :

$$\begin{bmatrix} \mathbf{f} \\ \mathbf{f}^* \end{bmatrix} = N \left(\begin{bmatrix} \mathbf{0} \\ 0 \end{bmatrix}, \begin{bmatrix} K_{11} & K_{12} & K_{13} & K_{1*} \\ K_{21} & K_{22} & K_{23} & K_{2*} \\ K_{31} & K_{32} & K_{33} & K_{3*} \\ K_{*1} & K_{*2} & K_{*3} & K_{**} \end{bmatrix} \right)$$

Bayesian Regression - Function Space View

▶ In actual scenario, mean \neq 0 :

$$\begin{bmatrix} \mathbf{f} \\ \mathbf{f}^* \end{bmatrix} = N \begin{pmatrix} \begin{bmatrix} \mu(\mathbf{X}) \\ \mu^* \end{bmatrix}, \begin{bmatrix} \mathbf{K} & \mathbf{K}^{*T} \\ \mathbf{K}^* & \mathbf{K}^{**} \end{bmatrix} \end{pmatrix}$$

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Conditional Distribution:

$$p(f^*|\mathbf{X^*},\mathbf{X},\mathbf{f}) = N(f^*|\mu^*,\Sigma^*)$$

Bayesian Regression - Function Space View

In actual scenario, mean $\neq 0$: $\begin{bmatrix} \mathbf{f} \\ \mathbf{f}^* \end{bmatrix} = N \begin{pmatrix} \begin{bmatrix} \mu(\mathbf{X}) \\ \mu* \end{bmatrix}, \begin{bmatrix} \mathbf{K} & \mathbf{K}^{*T} \\ \mathbf{K}^* & \mathbf{K}^{**} \end{bmatrix}$

Conditional Distribution:

$$\begin{split} \rho(f^*|\mathbf{X^*},\mathbf{X},\mathbf{f}) &= \textit{N}(f^*|\mu^*,\Sigma^*) \\ \text{such that}: \\ \\ \mu^* &= \mu(\mathbf{X}^*) + \mathbf{K}^*\mathbf{K}^{-1}(f-\mu(\mathbf{X})) \\ \\ \Sigma^* &= \mathbf{K}^{**} - \mathbf{K}^*\mathbf{K}^{-1}\mathbf{K}^* \end{split}$$

► Two steps:

$$p(f^*|\mathbf{X}^*,D) = \int p(f^*|\mathbf{X}^*,\mathbf{X},\mathbf{f})d\mathbf{f}$$

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$$p(f^*|\mathbf{X}^*,D) = \int p(f^*|\mathbf{X}^*,\mathbf{X},\mathbf{f})d\mathbf{f}$$

$$p(y_i^* = 1|f^*) = \int p(y_i^* = 1|f^*)p(f^*|\mathbf{X}^*, D)df^*$$

$$= \int \sigma(f^*)p(f^*|\mathbf{x}^*, D)df^*$$

Difficulties and Solutions

- Difficulties:
 - ► Finding the right prior
 - ► Intractable posterior

Difficulties and Solutions

- Difficulties:
 - Finding the right prior
 - Intractable posterior
- Solutions:
 - Sampling the posterior appropriately
 - Approximating the posterior

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Talking About BDL Advantages

- General Steps:
 - Current parameter w_{current}
 - Propose new parameter w_{new}
 - Accept or reject the proposed value based on probability
- Metrapolis Algorithm uses a Normal distribution to calculate the probability
 - $\mathbf{w}_t \sim \mathcal{N}(\mu, \Sigma)$
 - $ightharpoonup \mathbf{w}_{t+1} \sim N(\mathbf{w}_t, \Sigma)$

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 - $\blacktriangleright \ \mathbf{w}_{t+1} \sim \textit{N}(\mathbf{w}_t, \Sigma)$
 - $r(\mathbf{w}_{t+1}, \mathbf{w}_t) = \frac{post.prob.of \mathbf{w}_{t+1}}{post.prob.of \mathbf{w}_t}$

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 - $r(\mathbf{w}_{t+1}, \mathbf{w}_t) = \frac{post.prob.of \mathbf{w}_{t+1}}{post.prob.of \mathbf{w}_t}$
 - lacktriangleright if $r(\mathbf{w}_{t+1},\mathbf{w}_t)>1$, accept \mathbf{w}_{t+1}

- ► Two issues:
 - Dependent on the starting values
 - Correlation present because of Markov Chain
- Solution:
 - ▶ Burn-in period
 - ► Thinning : increasing the sample size
- Advantage : Accuracy high
- Disadvantage: Slow

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Laplace Approximation
Variational Inference

Talking About BDL Advantages Challenges

Laplace Approximation

- Parameter space w and data D
- ▶ Posterior : $p(w|D) = \frac{1}{Z}p(D|w)p(w)$

$$Z=\int p(D|w)p(w)$$

Laplace Approximation

- Parameter space w and data D
- ▶ Posterior : $p(w|D) = \frac{1}{Z}p(D|w)p(w)$

$$Z = \int p(D|w)p(w)$$

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$$\hat{w} = \underset{w}{\operatorname{argmax}} \ \psi(w)$$

Laplace Approximation

- Parameter space w and data D
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$$Z = \int p(D|w)p(w)$$

$$\hat{w} = \underset{w}{\operatorname{argmax}} \ \psi(w) \to \mathsf{MAP}$$

Laplace Approximation

► Taylors series :

$$\psi(w) = \psi(\hat{w}) + \frac{-1}{2}(w - \hat{w})^T H(w - \hat{w})$$

Laplace Approximation

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$$\psi(w) = \psi(\hat{w}) + \frac{-1}{2}(w - \hat{w})^T H(w - \hat{w})$$

 $p(w|D) = e^{\psi(w)}$

Laplace Approximation

Taylors series :

$$\psi(w) = \psi(\hat{w}) + \frac{-1}{2}(w - \hat{w})^T H(w - \hat{w})$$

$$p(w|D) = e^{\psi(\hat{w})}$$

$$= e^{\psi(\hat{w})} e^{\frac{-1}{2}(w - \hat{w})^T H(w - \hat{w})}$$

Laplace Approximation

► Taylors series :

$$\psi(w) = \psi(\hat{w}) + \frac{-1}{2}(w - \hat{w})^T H(w - \hat{w})$$

$$p(w|D) = e^{\psi(\hat{w})}$$

$$= e^{\psi(\hat{w})} e^{\frac{-1}{2}(w - \hat{w})^T H(w - \hat{w})}$$

$$\approx N(w; \hat{w}, H^{-1})$$

Laplace Approximation

Taylors series :

$$\psi(w) = \psi(\hat{w}) + \frac{-1}{2}(w - \hat{w})^T H(w - \hat{w})$$

$$p(w|D) = e^{\psi(w)}$$

$$= e^{\psi(\hat{w})} e^{\frac{-1}{2}(w - \hat{w})^T H(w - \hat{w})}$$

$$\approx N(w; \hat{w}, H^{-1})$$

$$Z = \int e^{\psi(w)} dw$$

► Taylors series :

$$\psi(w) = \psi(\hat{w}) + \frac{-1}{2}(w - \hat{w})^{T}H(w - \hat{w})$$

$$p(w|D) = e^{\psi(w)}$$

$$= e^{\psi(\hat{w})}e^{\frac{-1}{2}(w - \hat{w})^{T}H(w - \hat{w})}$$

$$\approx N(w; \hat{w}, H^{-1})$$

$$Z = \int e^{\psi(w)}dw$$

$$= e^{\psi(\hat{w})}\frac{(2\pi)^{d/2}}{|H|^{1/2}}$$

Laplace Approximation

► Advantage: Fast

▶ Disadvantage: Less accurate

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▶ Posterior :
$$p(Z|X) = \frac{p(X|Z)p(Z)}{\int p(X|Z)p(Z)dZ}$$

- ▶ Posterior : $p(Z|X) = \frac{p(X|Z)p(Z)}{\int p(X|Z)p(Z)dZ}$
- ▶ Approximate it to $q(Z; \theta)$
- ▶ Distance measurement : KL divergence defined as :

$$KL(q(Z;\theta)||p(Z|X)) = \sum_{z \in Z} q(z;\theta) log(\frac{q(z;\theta)}{p(z|x)})$$

$$KL(q(Z;\theta)||p(Z|X)) = \sum_{z \in Z} q(z;\theta) log(\frac{q(z;\theta)}{p(z|x)})$$

$$= \sum_{z \in Z} q(z;\theta) log(\frac{q(z;\theta)p(x)}{p(z,x)})$$

$$= \sum_{z \in Z} q(z;\theta) (log(\frac{q(z;\theta)}{p(z,x)}) + log(p(x)))$$

$$= log(p(x)) + E_q[log(\frac{q(z;\theta)}{p(z,x)}]$$

$$KL(q(Z;\theta)||p(Z|X)) = \sum_{z \in Z} q(z;\theta) log(\frac{q(z;\theta)}{p(z|x)})$$

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$$= log(p(x)) + E_q[log(\frac{q(z;\theta)}{p(z,x)}]$$

$$Minimize$$

Variational Inference

$$KL(q(Z;\theta)||p(Z|X)) = \sum_{z \in Z} q(z;\theta) log(\frac{q(z;\theta)}{p(z|x)})$$

$$= \sum_{z \in Z} q(z;\theta) log(\frac{q(z;\theta)p(x)}{p(z,x)})$$

$$= \sum_{z \in Z} q(z;\theta) (log(\frac{q(z;\theta)}{p(z,x)}) + log(p(x)))$$

$$= log(p(x)) + E_q[log(\frac{q(z;\theta)}{p(z,x)}]$$

$$Minimize$$

Variational Bound or ELBO(Evidence Lower Bound)

Variational Inference

► Minimizing $E_q[log(\frac{q(z;\theta)}{p(z,x)}]$

- Minimizing $E_q[log(\frac{q(z;\theta)}{p(z,x)}]$
- ► Maximize $-E_q[log(\frac{q(z;\theta)}{p(z,x)}]$

- ▶ Minimizing $E_q[log(\frac{q(z;\theta)}{p(z,x)}]$
- ► Maximize $-E_q[log(\frac{q(z;\theta)}{p(z,x)}]$
 - \Rightarrow Maximize $-E_q[log(\frac{q(z;\theta)}{p(x|z)p(z)}]$

- ▶ Minimizing $E_q[log(\frac{q(z;\theta)}{p(z,x)}]$
- ► Maximize $-E_q[log(\frac{q(z;\theta)}{p(z,x)}]$

$$\Rightarrow$$
 Maximize $-E_q[log(\frac{q(z;\theta)}{p(x|z)p(z)}]$

$$\Rightarrow$$
 Maximize $E_q[log(\frac{p(z)}{q(z;\theta)}] + E_q[log(p(x|z))]$

- Advantages:
 - ► Scalable to large datasets
 - Faster than MCMC
- Disadvantages:
- Does not guarantee a globally optimal $q(Z:\theta)$

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Advantages of Bayesian Deep Learning

- Can avoid overfitting
- It explains the hyperparameters in Deep Learning
- Pruning based on probability
- Robust against adversarial examples

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Challenges with Bayesian Deep Learning

- Modelling multi modal distributions which will help in better inference
- Prior distribution selection
- Posterior approximation

References I

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