

# Data Analysis with Python

Log Transformations + Interactions + More  
on Multiple Regression

*(Class #4)*

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# **Agenda**

1. Log transformations
2. Interactions

# **1. Log transformations**

# Why Log Transform Variables?

# Log Transformations

In order to make some variables “more normal,” or more linear, or to increase interpretability, we often log them.

# Logging

The natural logarithm of a number is the exponent to which we have to raise the *base*(~2.72) to obtain *that number*

original	ln
1	0
10	2.3
10,000	9.2
100,000	11.5
1,000,000	13.8

**N.B., You cannot take the log of 0 ... this can be a problem**

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# Preliminaries...

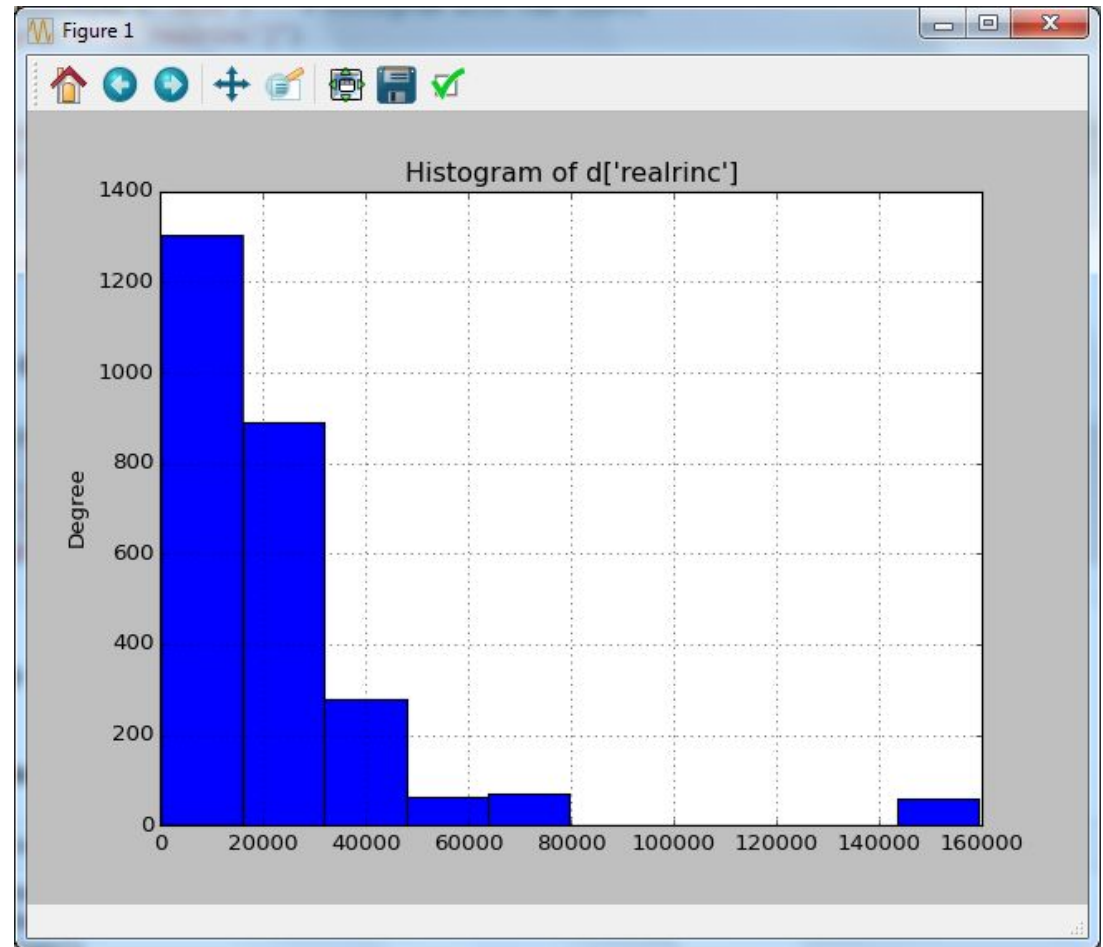
```
from __future__ import division
import pandas as pd
import numpy as np
import statsmodels.api as sm
import statsmodels.formula.api as smf
import os
import matplotlib.pyplot as plt

os.chdir('C:/Users/gme2101/Desktop/Data Analysis Data') # change working directory
d = pd.read_csv("GSS_Cum.csv", usecols=["sex", "educ", "year", "realrinc", "hrs1", "wordsum",
"wrkstat", "race", "trust", "region", "fund", "evolved", "realinc", "sibs", "madeg", "fund",
"marital", "attend", "age", "family16"])
d.head()

sub = d[d["year"] == 2006]
```

# Distribution of *realrinc*

```
sub["realrinc"].plot(kind  
= 'hist')      # histogram  
with raw counts  
plt.title("Histogram of  
d['realrinc']")  
plt.show()
```

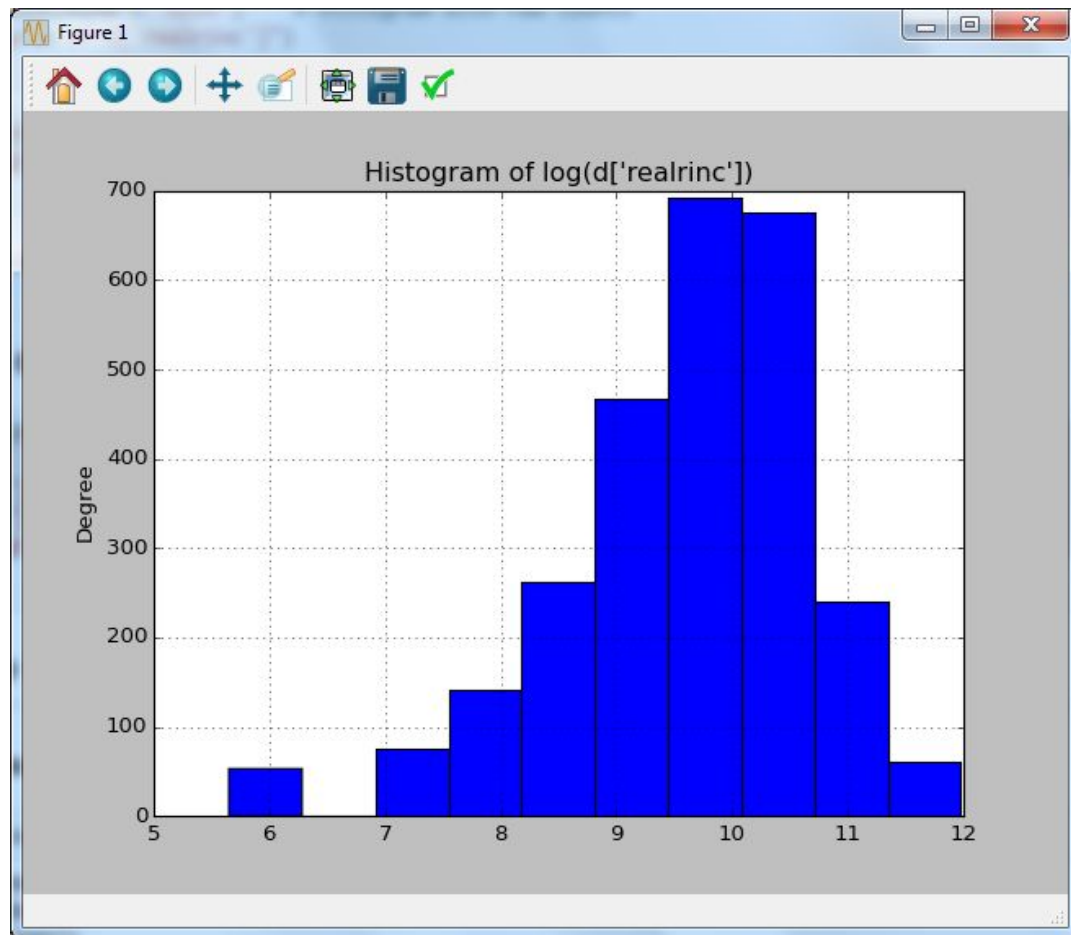




# Distribution of *ln.realrinc*

```
pd.options.mode.chained_assignment = None
```

```
sub["ln_realrinc"] = np.log(sub["realrinc"])  
sub["ln_realrinc"].plot(kind = 'hist')  
plt.title("Histogram of log(d['realrinc'])")  
plt.show()
```



# Looking at the shape of our variables

```
## RAW income ##
```

```
sub["realrinc"].skew()
```

```
3.4025770427387112
```

```
sub["realrinc"].kurtosis()
```

```
14.553164088336334
```

Raw income has skew=3.4 and kurtosis=14.51, while a normal, symmetric distribution will have skew=0, kurtosis=3;

```
## LOGGED income ##
```

```
sub["ln_realrinc"].skew()
```

```
-1.0295369876382472
```

```
sub["ln_realrinc"].kurtosis()
```

```
2.0276359476609085
```

Log(income) has skew of -1.03 and kurtosis of 2.02.

*Which is closer to our ideal normal distribution?*

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# Log Transformations

**TABLE 2.3**

**Summary of Functional Forms Involving Logarithms**

Model	Dependent Variable	Independent Variable	Interpretation of $\beta_1$
Level-level	$y$	$x$	$\Delta y = \beta_1 \Delta x$
Level-log	$y$	$\log(x)$	$\Delta y = (\beta_1/100)\% \Delta x$
Log-level	$\log(y)$	$x$	$\% \Delta y = (100\beta_1) \Delta x$
Log-log	$\log(y)$	$\log(x)$	$\% \Delta y = \beta_1 \% \Delta x$

# A log-log model

```
pd.options.mode.chained_assignment = None
```

```
sub["ln_hrs1"] = np.log(sub["hrs1"])
sub["ln_realrinc"] = np.log(sub["realrinc"])
lm1 = smf.ols(formula = "ln_realrinc ~ ln_hrs1 + C(sex)", data = sub).fit()
print (lm1.summary())
```

## OLS Regression Results

```
=====
Dep. Variable:          ln_realrinc    R-squared:                0.181
Model:                  OLS           Adj. R-squared:           0.180
Method:                 Least Squares  F-statistic:              250.6
Date:                  Wed, 07 Jun 2017  Prob (F-statistic):      4.60e-99
Time:                  16:15:49        Log-Likelihood:           -3070.9
No. Observations:      2275           AIC:                     6148.
Df Residuals:          2272           BIC:                     6165.
Df Model:              2
Covariance Type:       nonrobust
=====
```

	coef	std err	t	P> t	[95.0% Conf. Int.]
Intercept	6.6776	0.176	38.049	0.000	6.333 7.022
C(sex) [T.2]	-0.3383	0.040	-8.479	0.000	-0.417 -0.260
ln_hrs1	0.8645	0.046	18.732	0.000	0.774 0.955

```
=====
Omnibus:                434.513    Durbin-Watson:           1.792
Prob(Omnibus):          0.000      Jarque-Bera (JB):        1354.151
Skew:                   -0.962      Prob(JB):                8.91e-295
Kurtosis:               6.253      Cond. No.                 35.9
=====
```

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Warnings:

# A log-log model

	coef	std err	t	P> t	[95.0% Conf. Int.]	
Intercept	6.6776	0.176	38.049	0.000	6.333	7.022
C(sex) [T.2]	-0.3383	0.040	-8.479	0.000	-0.417	-0.260
ln_hrs1	0.8645	0.046	18.732	0.000	0.774	0.955

Controlling for sex, a 1% increase in work hours leads (on average) to a 0.86% increase in salary

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# A level-log model

```
pd.options.mode.chained_assignment = None
```

```
sub["ln_wordsum"] = np.log(sub["wordsum"])
sub["working"] = sub["wrkstat"].apply(lambda e: 1 if e < 3 else 0)
sub["ln_wordsum"] = sub["ln_wordsum"].map(lambda x: np.nan if x == -float('Inf') else x)
lm2 = smf.ols(formula = "tvhours ~ ln_wordsum + working", data = sub).fit()
print (lm2.summary())
```

## OLS Regression Results

```
=====
Dep. Variable:          tvhours      R-squared:                0.075
Model:                  OLS          Adj. R-squared:           0.073
Method:                 Least Squares  F-statistic:             37.34
Date:                  Mon, 20 May 2019  Prob (F-statistic):       2.57e-16
Time:                  15:29:30       Log-Likelihood:          -2080.8
No. Observations:      921           AIC:                     4168.
Df Residuals:          918           BIC:                     4182.
Df Model:               2
Covariance Type:       nonrobust
=====
```

	coef	std err	t	P> t	[95.0% Conf. Int.]
Intercept	5.4222	0.360	15.046	0.000	4.715 6.130
ln_wordsum	-1.0379	0.194	-5.360	0.000	-1.418 -0.658
working	-1.0174	0.155	-6.571	0.000	-1.321 -0.714

```
=====
Omnibus:                556.467      Durbin-Watson:           2.053
Prob(Omnibus):          0.000        Jarque-Bera (JB):        7021.934
Skew:                   2.543         Prob(JB):                0.00
Kurtosis:               15.535        Cond. No.                 11.4
=====
```

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# A level-log model

	coef	std err	t	P> t	[95.0% Conf. Int.]	
Intercept	5.4222	0.360	15.046	0.000	4.715	6.130
ln_wordsum	-1.0379	0.194	-5.360	0.000	-1.418	-0.658
working	-1.0174	0.155	-6.571	0.000	-1.321	-0.714

Controlling for working status, a 1% increase  
in vocabulary score leads (on average) to a  
-0.0104 hour decrease in TV hours

\*

# A level-log model

	coef	std err	t	P> t	[95.0% Conf. Int.]	
Intercept	5.4222	0.360	15.046	0.000	4.715	6.130
ln_wordsum	-1.0379	0.194	-5.360	0.000	-1.418	-0.658
working	-1.0174	0.155	-6.571	0.000	-1.321	-0.714

Or: Controlling for working status, a 100% increase in vocabulary score leads (on average) to a 1.04 hour decrease in TV hours



# Another log-log model

```
b["tg13"] = b["domgross_2013$"] + b["intgross_2013$"]  
b["tot_gross_13_mil"] = b["tg13"] / (1000000)  
b["budget_13_mil"] = b["budget_2013$"] / (1000000)
```

```
b["ln_bud"] = np.log(b["budget_13_mil"])  
b["ln_tot"] = np.log(b["tot_gross_13_mil"])
```

# Another log-log model

```
lm1 = smf.ols(formula = "ln_tot ~ binary", data = b).fit()  
print(lm1.summary())
```

```

=====
                        OLS Regression Results
=====
Dep. Variable:          ln_tot      R-squared:          0.010
Model:                  OLS        Adj. R-squared:       0.009
Method:                 Least Squares    F-statistic:       17.02
Date:                   Mon, 20 May 2019    Prob (F-statistic): 3.86e-05
Time:                   15:30:54          Log-Likelihood:    -3463.8
No. Observations:      1776             AIC:              6932.
=====

```

	coef	std err	t	P> t	[95.0% Conf. Int.]	
Intercept	4.9290	0.054	90.740	0.000	4.823	5.036
binary[T.PASS]	-0.3352	0.081	-4.126	0.000	-0.495	-0.176

```
=====
Omnibus:                 521.529    Durbin-Watson:          1.988
Prob(Omnibus):            0.000    Jarque-Bera (JB):       1724.285
Skew:                     -1.449    Prob(JB):               0.00
Kurtosis:                 6.861     Cond. No.:              2.51
=====
```

Model 1: Passing the Bechler test reduces the  
predicted total revenues of a movie by  
33.5%

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# Another log-log model

```
lm2 = smf.ols(formula = "ln_tot ~ binary + ln_bud", data = b).fit()  
print (lm2.summary())
```

## OLS Regression Results

=====						
Dep. Variable:	ln_tot	R-squared:	0.448			
Model:	OLS	Adj. R-squared:	0.447			
Method:	Least Squares	F-statistic:	719.5			
Date:	Mon, 20 May 2019	Prob (F-statistic):	1.63e-229			
Time:	15:35:03	Log-Likelihood:	-2944.5			
No. Observations:	1776	AIC:	5895.			
=====						
	coef	std err	t	P> t	[95.0% Conf. Int.]	
-----						
Intercept	1.9699	0.089	22.216	0.000	1.796	2.144
binary[T.PASS]	-0.0610	0.061	-0.998	0.319	-0.181	0.059
ln_bud	0.8304	0.022	37.531	0.000	0.787	0.874
=====						
Omnibus:	488.585	Durbin-Watson:	1.909			
Prob(Omnibus):	0.000	Jarque-Bera (JB):	2818.471			
Skew:	-1.162	Prob(JB):	0.00			
Kurtosis:	8.717	Cond. No.	12.1			
=====						

Model 2: Controlling for the Bechler test, a 1% increase in the budget of a movie increases its predicted revenues by 0.83%

# Another log-log model

	coef	std err	t	P> t	[95.0% Conf. Int.]
Intercept	1.9699	0.089	22.216	0.000	1.796 2.144
binary[T.PASS]	-0.0610	0.061	-0.998	0.319	-0.181 0.059
ln_bud	0.8304	0.022	37.531	0.000	0.787 0.874

Model 2: Controlling for budget, a movie that passes the Bechdel test is predicted to reduce revenues by 6.1% on average (n.s.)

# Remember the original model

```
lm3 = smf.ols(formula = "tot_gross_13_mil ~ binary + budget_13_mil", data = b).fit()
print(lm3.summary()) -- OLS Regression Results
```

```
=====
Dep. Variable:          tot_gross_13_mil    R-squared:                0.316
Model:                  OLS                Adj. R-squared:          0.315
Method:                 Least Squares      F-statistic:            408.7
Date:                  Mon, 20 May 2019    Prob (F-statistic):      1.10e-146
Time:                  15:35:07           Log-Likelihood:          -12839.
No. Observations:      1776              AIC:                    2.568e+04
=====
```

	coef	std err	t	P> t	[95.0% Conf. Int.]	
Intercept	71.4731	14.099	5.069	0.000	43.820	99.126
binary[T.PASS]	-14.9222	16.123	-0.926	0.355	-46.543	16.699
budget_13_mil	4.0963	0.146	28.108	0.000	3.810	4.382

```
=====
Omnibus:                1696.847    Durbin-Watson:            1.942
Prob(Omnibus):           0.000      Jarque-Bera (JB):         104439.408
Skew:                    4.389      Prob(JB):                 0.00
Kurtosis:                39.528     Cond. No.:                 190.
=====
```

# Another log-log model

## Regression Results

	Ln(Total Gross)		Total Gross, Mil
	Model 1	Model 2	Model 3
Pass Bechler	-0.335*** (0.081)	-0.061 (0.061)	-14.922 (16.123)
Ln(Total Budget)		0.830*** (0.022)	
Total Budget, Mil			4.096*** (0.146)
Constant	4.929*** (0.054)	1.970*** (0.089)	71.473*** (14.099)
Observations	1,776	1,776	1,776
Adjusted R2	0.009	0.447	0.315

Note: \*p<0.1; \*\*p<0.05; \*\*\*p<0.01

# How'd I do that (in R)?

```
lm1 = lm(d$ln.tot ~ binary , d)
lm2 = lm(d$ln.tot ~ binary + ln.bud, d)
lm3 = lm(d$tot.gross.13.mil ~ binary + d$budget.13.mil, d)

library(stargazer)
stargazer(lm1, lm2, type = "text")

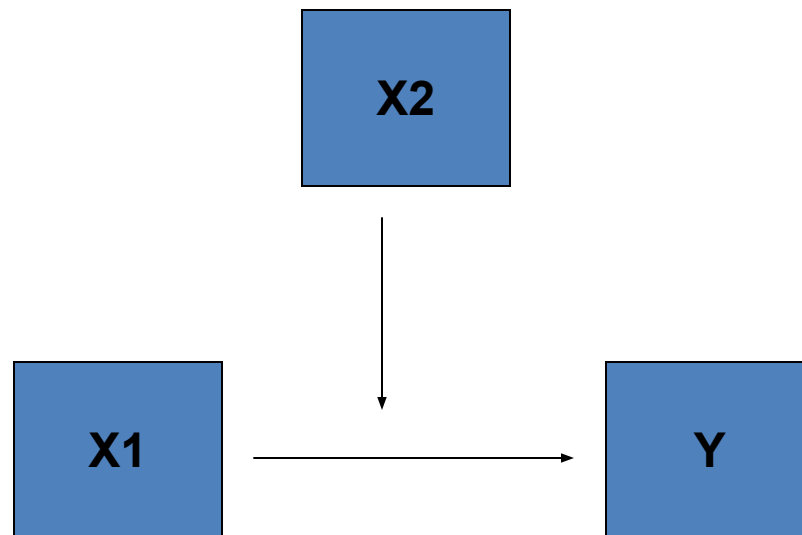
stargazer(lm1, lm2, lm3,
          title="Regression Results",
          align=TRUE,
          dep.var.labels=c("Ln(Total Gross)", "Total Gross, Mil"),
          covariate.labels=c("Pass Bechler", "Ln(Total Budget)", "Total
Budget, Mil"),
          no.space=TRUE,
          omit.stat=c("LL", "ser", "f", "rsq"),
          column.labels=c("Model 1", "Model 2", "Model 3"),
          dep.var.caption="",
          model.numbers=FALSE,
          type = "text")
```

## **2. Interactions**



# Interactions

- Moderation: There is an “interaction” between our X1 and X2
- The association of X1 and Y varies according to levels of X2



# Why would be use interactions?

If we think that 2 of our independent variables may have a relationship that affects the magnitude of our dependent variable

- That the effect of 1 variable may depend on the magnitude of another variable ...

# Moderation: Our simple model

$$Y = a + B_1 X_1 + B_2 X_2 + u$$

$$\text{R's Income} = a + B_1(\text{Gender}) + B_2(\text{Educ})$$

**Why might gender and education interact to predict salary?**

**At higher levels of education, are the differences amplified or minimized between males and females on income?**

# These are the recodes ...

## Income in \$10,000 units

```
sub_new = sub[["realrinc", "educ", "sex"]]
sub_new["female"] = sub_new["sex"] == 2
sub_new.dropna(subset = ["realrinc", "educ"], inplace = True)
sub_new["realrinc10k"] = (sub_new.realrinc) / 10000
sub_new["realrinc10k"].describe()
```

```
count    2663.00
mean       2.36
std        2.60
min        0.03
25%        0.92
50%        1.85
75%        3.13
max       15.93
Name: realrinc10k, dtype: float64
```

# A simple multiple regression

```
pd.options.display.float_format = '{0:1.2f}'.format
lm_income = smf.ols(formula = "realrinc10k ~ educ + female", data = sub_new).fit()
print(lm_income.summary())
sub_new["fitted"] = lm_income.predict()
```

## OLS Regression Results

```
=====
Dep. Variable:          realrinc10k      R-squared:                0.142
Model:                  OLS              Adj. R-squared:           0.141
Method:                 Least Squares    F-statistic:              220.1
Date:                  Tue, 21 May 2019   Prob (F-statistic):       3.64e-89
Time:                  09:28:39          Log-Likelihood:          -6121.0
No. Observations:      2663             AIC:                    1.225e+04
Df Residuals:          2660             BIC:                    1.227e+04
Df Model:               2
Covariance Type:       nonrobust
=====
```

	coef	std err	t	P> t	[95.0% Conf. Int.]
Intercept	-0.6051	0.215	-2.813	0.005	-1.027 -0.183
female[T.True]	-1.1877	0.094	-12.698	0.000	-1.371 -1.004
educ	0.2588	0.015	17.211	0.000	0.229 0.288

```
=====
Omnibus:                1902.623      Durbin-Watson:              1.890
Prob(Omnibus):           0.000        Jarque-Bera (JB):          29109.520
Skew:                   3.292         Prob(JB):                  0.00
Kurtosis:               17.799        Cond. No.                  65.4
=====
```

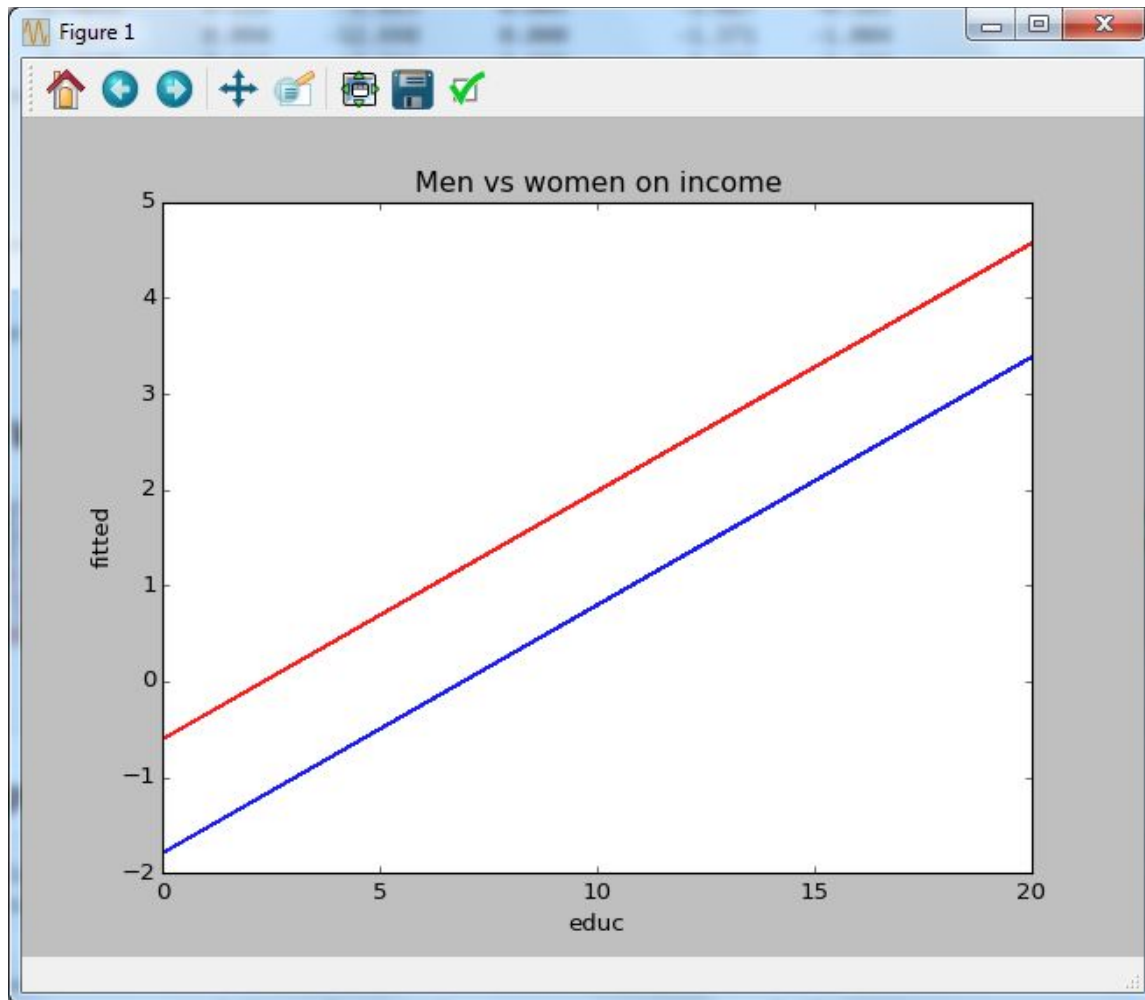
(c) Firich 2012

# A simple multiple regression

For each additional year of education, net of sex, someone earns \$2,588 more (statistically significant) per year; a female – net of education – earns \$11,877 less per year (statistically significant)

	coef	std err	t	P> t	[95.0% Conf. Int.]	
Intercept	-0.6051	0.215	-2.813	0.005	-1.027	-0.183
female[T.True]	-1.1877	0.094	-12.698	0.000	-1.371	-1.004
educ	0.2588	0.015	17.211	0.000	0.229	0.288

# Female vs. male on income





# How did I do that graph?

```
plt.plot(sub_new["educ"], lm_income.params[0] + lm_income.params[1] * 1 +  
lm_income.params[2] * sub_new["educ"], 'b', label = 'female', alpha = 0.9)  
plt.plot(sub_new["educ"], lm_income.params[0] + lm_income.params[1] * 0 +  
lm_income.params[2] * sub_new["educ"], 'r', label = 'male', alpha = 0.9)  
plt.title("Men vs women on income")  
plt.xlabel("educ")  
plt.ylabel("fitted")  
plt.show()
```

# A simple regression for males

For males, each additional year of education, they earn \$2,990 per year (statistically significant)

(A male with no education ( $X=0$ ) has -\$11,522)

```
lm_males = smf.ols(formula = "realrinc10k ~ educ", data = sub_new, subset = sub_new.female == 0).fit()
print(lm_males.summary())
```

## OLS Regression Results

```
=====
Dep. Variable:          realrinc10k      R-squared:                0.106
Model:                  OLS              Adj. R-squared:           0.105
Method:                 Least Squares    F-statistic:             156.9
Date:                   Tue, 21 May 2019  Prob (F-statistic):       4.30e-34
Time:                   09:30:22         Log-Likelihood:          -3296.2
No. Observations:      1329             AIC:                    6596.
Df Residuals:          1327             BIC:                    6607.
Df Model:               1
Covariance Type:       nonrobust
=====
```

	coef	std err	t	P> t	[95.0% Conf. Int.]
Intercept	-1.1522	0.335	-3.444	0.001	-1.808 -0.496
educ	0.2990	0.024	12.524	0.000	0.252 0.346

```
=====
Omnibus:                 796.985      Durbin-Watson:           1.918
Prob(Omnibus):           0.000        Jarque-Bera (JB):        6573.585
Skew:                    2.760        Prob(JB):                0.00
Kurtosis:                12.394        Cond. No.                59.4
=====
```

# A simple regression for females

For females, each additional year of education, they earn \$2,053 per year (statistically significant)

(A female with no education ( $X=0$ ) earns -\$10,516 per year)

```
lm_females = smf.ols(formula = "realrinc10k ~ educ", data = sub_new, subset = sub_new.female == 1).fit()
print(lm_females.summary())
```

## OLS Regression Results

```
=====
Dep. Variable:          realrinc10k      R-squared:                0.097
Model:                  OLS              Adj. R-squared:           0.097
Method:                 Least Squares    F-statistic:              143.9
Date:                   Tue, 21 May 2019  Prob (F-statistic):       1.51e-31
Time:                   09:30:43          Log-Likelihood:           -2675.2
No. Observations:       1334             AIC:                     5354.
Df Residuals:           1332             BIC:                     5365.
Df Model:                1
Covariance Type:        nonrobust
=====
```

	coef	std err	t	P> t	[95.0% Conf. Int.]
Intercept	-1.0516	0.243	-4.336	0.000	-1.527 -0.576
educ	0.2054	0.017	11.994	0.000	0.172 0.239

```
=====
Omnibus:                1221.317      Durbin-Watson:              1.838
Prob(Omnibus):           0.000        Jarque-Bera (JB):           47781.764
Skew:                    4.220         Prob(JB):                   0.00
Kurtosis:                31.079        Cond. No.                   70.1
=====
```

# The interaction model I

1. When female=0 (i.e., for males), at the intercept, they earn -\$11,522 on average with 0 years of education.

```
# Note: the * in the formula means that we want the interaction term in addition to each term separately
```

```
# Note: use ':' instead if you want to include the interaction term only
```

```
lm_income2 = smf.ols(formula = "realrinc10k ~ educ * female", data = sub_new).fit()
```

```
print(lm_income2.summary())
```

## OLS Regression Results

```
=====
Dep. Variable:          realrinc10k      R-squared:                0.145
Model:                  OLS              Adj. R-squared:           0.144
Method:                 Least Squares     F-statistic:               150.4
Date:                  Tue, 21 May 2019    Prob (F-statistic):        5.30e-90
Time:                  09:31:21           Log-Likelihood:            -6116.2
No. Observations:      2663              AIC:                      1.224e+04
Df Residuals:          2659              BIC:                      1.226e+04
Df Model:               3
Covariance Type:       nonrobust
=====
```

	coef	std err	t	P> t	[95.0% Conf. Int.]
Intercept	-1.1522	0.278	-4.138	0.000	-1.698 -0.606
female[T.True]	0.1006	0.428	0.235	0.814	-0.738 0.939
educ	0.2990	0.020	15.046	0.000	0.260 0.338
educ:female[T.True]	-0.0936	0.030	-3.087	0.002	-0.153 -0.034

```
=====
Omnibus:                1893.429      Durbin-Watson:              1.889
Prob(Omnibus):           0.000        Jarque-Bera (JB):           28848.466
Skew:                    3.270         Prob(JB):                   0.00
=====
```

# The interaction model II

2. When female=0 (i.e., for males), they earn \$2,990 on average for each additional year of education.

	coef	std err	t	P> t	[95.0% Conf. Int.]	
Intercept	-1.1522	0.278	-4.138	0.000	-1.698	-0.606
female[T.True]	0.1006	0.428	0.235	0.814	-0.738	0.939
educ	0.2990	0.020	15.046	0.000	0.260	0.338
educ:female[T.True]	-0.0936	0.030	-3.087	0.002	-0.153	-0.034

# The interaction model III

3. When educ=0, for females (female=1), they earn \$1006 on average more than males.

	coef	std err	t	P> t	[95.0% Conf. Int.]	
Intercept	-1.1522	0.278	-4.138	0.000	-1.698	-0.606
female[T.True]	0.1006	0.428	0.235	0.814	-0.738	0.939
educ	0.2990	0.020	15.046	0.000	0.260	0.338
educ:female[T.True]	-0.0936	0.030	-3.087	0.002	-0.153	-0.034

# The interaction model IV

4a. For females, they get \$936 less than males for each year more of education, so males get \$2,990, but females get  $\$2,990 - \$936 = \$2054$ , on average

	coef	std err	t	P> t	[95.0% Conf. Int.]	
Intercept	-1.1522	0.278	-4.138	0.000	-1.698	-0.606
female[T.True]	0.1006	0.428	0.235	0.814	-0.738	0.939
educ	0.2990	0.020	15.046	0.000	0.260	0.338
educ:female[T.True]	-0.0936	0.030	-3.087	0.002	-0.153	-0.034

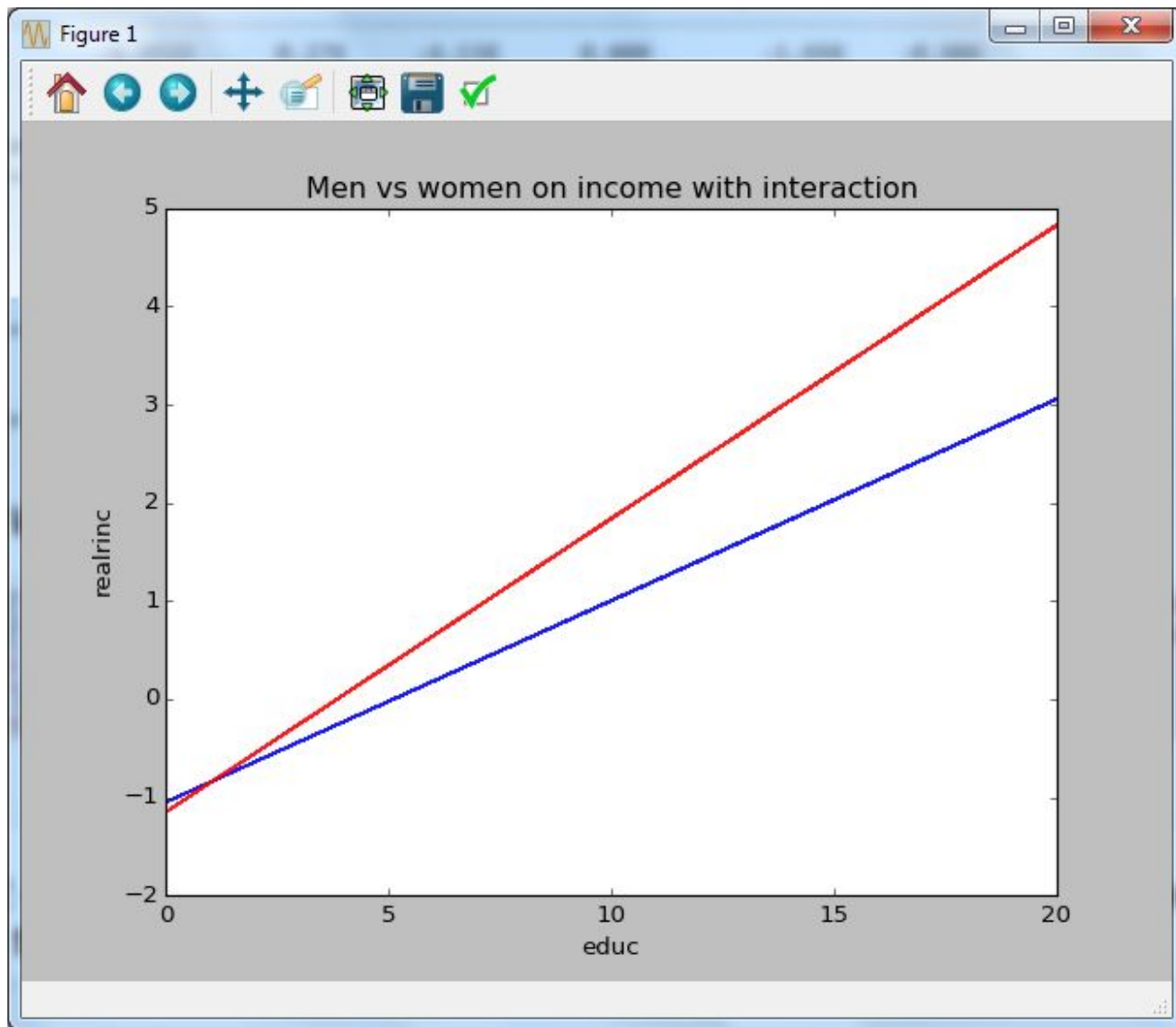
# The interaction model IV

4b. Or: For females on average, the *difference* in rates of return to education is \$936 per year less than for males

	coef	std err	t	P> t	[95.0% Conf. Int.]	
Intercept	-1.1522	0.278	-4.138	0.000	-1.698	-0.606
female[T.True]	0.1006	0.428	0.235	0.814	-0.738	0.939
educ	0.2990	0.020	15.046	0.000	0.260	0.338
educ:female[T.True]	-0.0936	0.030	-3.087	0.002	-0.153	-0.034



# Male vs. female ... with interaction



# How did I do this graph?

```
plt.plot(sub_new["educ"], lm_income2.params[0] + lm_income2.params[1] * 1 +
lm_income2.params[2] * sub_new["educ"] + lm_income2.params[3] * 1 *
sub_new["educ"], 'b', label = 'female', alpha = 0.9)
plt.plot(sub_new["educ"], lm_income2.params[0] + lm_income2.params[1] * 0 +
lm_income2.params[2] * sub_new["educ"] + lm_income2.params[3] * 0 *
sub_new["educ"], 'r', label = 'male', alpha = 0.9)
plt.title("Men vs women on income with interaction")
plt.xlabel("educ")
plt.ylabel("realrinc")
plt.show()
```

## Extra Credit

Try this on the **log** scale and see what difference it makes, if any.

# Notes on Interpretation:

- Interactions are multiplicative in nature
- Must always include  $X1$  and  $X2$  if you are including  $X1 * X2$
- With interactions included, original Bs for  $X1$  and  $X2$  refer to when  $X1=0$  or when  $X2=0$  ... not additive anymore
- Determining statistical significance is trickier

# Another example ...

Do well-off kids suffer educationally the same amount for each additional sibling, as do non-well-off kids?

# A simple multiple regression

```
pd.options.mode.chained_assignment = None

sub_kids = sub[["educ", "sibs", "madeg", "family16", "age"]]
sub_kids["maBA"] = sub_kids['madeg'].isin([3,4])
```

# A simple multiple regression

```
lm_maBA = smf.ols(formula = 'educ ~ sibs + maBA', data = sub_kids).fit()
print (lm_maBA.summary())
```

## OLS Regression Results

```
=====
Dep. Variable:          educ    R-squared:          0.147
Model:                  OLS     Adj. R-squared:      0.146
Method:                 Least Squares    F-statistic:      256.1
Date:                  Tue, 21 May 2019    Prob (F-statistic):  2.31e-103
Time:                  09:39:57    Log-Likelihood:     -7487.4
No. Observations:      2984    AIC:              1.498e+04
Df Residuals:          2981    BIC:              1.500e+04
Df Model:              2
Covariance Type:       nonrobust
=====
```

	coef	std err	t	P> t	[95.0% Conf. Int.]	
Intercept	14.1891	0.090	157.839	0.000	14.013	14.365
maBA[T.True]	2.1164	0.165	12.860	0.000	1.794	2.439
sibs	-0.2859	0.017	-16.493	0.000	-0.320	-0.252

```
=====
Omnibus:              329.893    Durbin-Watson:      1.778
Prob(Omnibus):        0.000    Jarque-Bera (JB):    744.276
Skew:                 -0.664    Prob(JB):            2.41e-162
Kurtosis:             5.055    Cond. No.            15.4
=====
```

## Warnings:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

# A simple multiple regression

For each additional sibling, net of their mom's BA+ degree, a person gets  $-.286$  years less of education (statistically significant)

	coef	std err	t	P> t	[95.0% Conf. Int.]	
Intercept	14.1891	0.090	157.839	0.000	14.013	14.365
maBA[T.True]	2.1164	0.165	12.860	0.000	1.794	2.439
sibs	-0.2859	0.017	-16.493	0.000	-0.320	-0.252



# A simple multiple regression

Net of the number of siblings they have, someone's whose mom has a BA+ degree gets 2.12 years more education than someone's whose mom has less than a BA (statistically significant)

	coef	std err	t	P> t	[95.0% Conf. Int.]	
Intercept	14.1891	0.090	157.839	0.000	14.013	14.365
maBA[T.True]	2.1164	0.165	12.860	0.000	1.794	2.439
sibs	-0.2859	0.017	-16.493	0.000	-0.320	-0.252

# A simple regression for kids of <BA moms

For kids whose mom has less than a BA, each additional sibling reduces their education by -.298 years of schooling (statistically significant)

(A kid with no sibling ( $X=0$ ) has 14.24 years of education)

```
lm_maBA0 = smf.ols(formula = 'educ ~ sibs', data = sub_kids, subset = sub_kids.maBA == 0).fit()
print(lm_maBA0.summary())
```

## OLS Regression Results

```
=====
Dep. Variable:          educ    R-squared:                0.093
Model:                  OLS    Adj. R-squared:            0.092
Method:                 Least Squares    F-statistic:        264.9
Date:                  Wed, 07 Jun 2017    Prob (F-statistic):    8.46e-57
Time:                  17:14:52    Log-Likelihood:        -6590.3
No. Observations:      2600    AIC:                  1.318e+04
Df Residuals:          2598    BIC:                  1.320e+04
Df Model:               1
Covariance Type:       nonrobust
=====
```

	coef	std err	t	P> t	[95.0% Conf. Int.]
Intercept	14.2366	0.094	151.789	0.000	14.053 14.421
sibs	-0.2979	0.018	-16.277	0.000	-0.334 -0.262

```
=====
Omnibus:                278.214    Durbin-Watson:        1.767
Prob (Chi-sq)           0.000      (c) Erich 2012
F-statistic              264.900    Prob (F-stat)          8.46e-57
=====
```

# A simple regression for kids of BA+ moms

For kids whose mom has a BA+, each additional sibling reduces their education only by  $-.084$  years of schooling (not statistically significant)

(A kid with no sibling ( $X=0$ ) has 15.8 years of education)

```
lm_maBA1 = smf.ols(formula = 'educ ~ sibs', data = sub_kids, subset = sub_kids.maBA == 1).fit()
print(lm_maBA1.summary())
```

## OLS Regression Results

```
=====
Dep. Variable:          educ    R-squared:                0.006
Model:                  OLS      Adj. R-squared:           0.003
Method:                 Least Squares    F-statistic:         2.121
Date:                  Wed, 07 Jun 2017    Prob (F-statistic):    0.146
Time:                  17:15:22    Log-Likelihood:       -872.87
No. Observations:      384      AIC:                  1750.
Df Residuals:          382      BIC:                  1758.
Df Model:              1
Covariance Type:       nonrobust
=====
```

	coef	std err	t	P> t	[95.0% Conf. Int.]
Intercept	15.7959	0.189	83.536	0.000	15.424 16.168
sibs	-0.0841	0.058	-1.456	0.146	-0.198 0.029

(c) Erich 2012

```
=====
Omnibus:              38.444    Durbin-Watson:           1.748
```

# The interaction model I

1. When maBA=0 (i.e., for kids of low educated moms), and with zero siblings, the intercept is the predicted amount of schooling of 14.23

```
lm_maBA_Inter = smf.ols(formula = 'educ ~ sibs * maBA', data = sub_kids).fit()
print(lm_maBA_Inter.summary())
```

## OLS Regression Results

```
=====
Dep. Variable:          educ      R-squared:          0.149
Model:                  OLS       Adj. R-squared:       0.148
Method:                 Least Squares   F-statistic:       173.8
Date:                  Wed, 07 Jun 2017   Prob (F-statistic): 7.36e-104
Time:                  17:17:26          Log-Likelihood:    -7483.3
No. Observations:      2984            AIC:              1.497e+04
Df Residuals:          2980            BIC:              1.500e+04
Df Model:               3
Covariance Type:       nonrobust
=====
```

	coef	std err	t	P> t	[95.0% Conf. Int.]
Intercept	14.2366	0.091	155.887	0.000	14.058 14.416
maBA[T.True]	1.5593	0.256	6.102	0.000	1.058 2.060
sibs	-0.2979	0.018	-16.716	0.000	-0.333 -0.263
sibs:maBA[T.True]	0.2138	0.075	2.847	0.004	0.067 0.361

```
=====
Omnibus:                 325.457      Durbin-Watson:          1.773
Prob(Omnibus):            0.000      Jarque-Bera (JB):       734.038
Skewness:                 0.656      Prob (chi-sq):          4.02e-160
=====
```

# The interaction model II

2. When  $\text{maBA}=0$  (i.e., for kids of low educated moms), each additional sibling costs a person  $-.298$  years of education.

	coef	std err	t	P> t	[95.0% Conf. Int.]	
Intercept	14.2366	0.091	155.887	0.000	14.058	14.416
maBA[T.True]	1.5593	0.256	6.102	0.000	1.058	2.060
sibs	-0.2979	0.018	-16.716	0.000	-0.333	-0.263
sibs:maBA[T.True]	0.2138	0.075	2.847	0.004	0.067	0.361

=====

# The interaction model III

3. When sibs=0, for kids with a BA+ mom, they get 1.56 years more of schooling.

	coef	std err	t	P> t	[95.0% Conf. Int.]	
Intercept	14.2366	0.091	155.887	0.000	14.058	14.416
maBA[T.True]	1.5593	0.256	6.102	0.000	1.058	2.060
sibs	-0.2979	0.018	-16.716	0.000	-0.333	-0.263
sibs:maBA[T.True]	0.2138	0.075	2.847	0.004	0.067	0.361
=====						

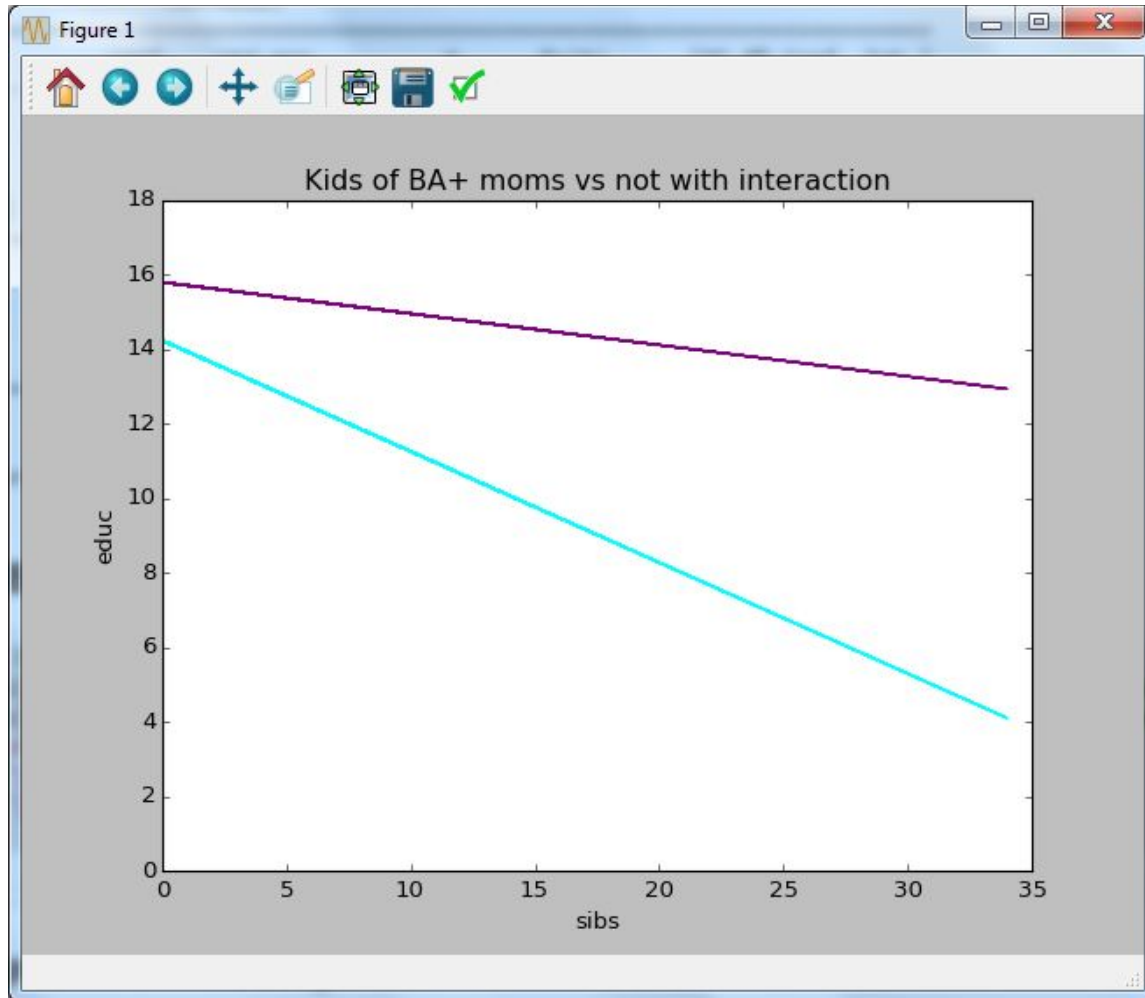
# The interaction model IV

4. For kids of BA+ moms, they gain back .21 for each additional sibling they have. So while they get what kids without BA+ moms get (which is  $-.298$ ), they add back .21, which means they ultimately get  $-.08$  for each additional sibling

	coef	std err	t	P> t	[95.0% Conf. Int.]	
Intercept	14.2366	0.091	155.887	0.000	14.058	14.416
maBA[T.True]	1.5593	0.256	6.102	0.000	1.058	2.060
sibs	-0.2979	0.018	-16.716	0.000	-0.333	-0.263
sibs:maBA[T.True]	0.2138	0.075	2.847	0.004	0.067	0.361

=====

# Kids of BA+ moms vs. not, with interaction





# Here is that graph

```
plt.axis([0, 35, 0, 18])
plt.plot(sub_kids["sibs"], lm_maBA_Inter.params[0] +
lm_maBA_Inter.params[1] * 0 + lm_maBA_Inter.params[2] *
sub_kids["sibs"] + lm_maBA_Inter.params[3] * 0 * sub_kids["sibs"],
'cyan', label = '<BA', alpha = 0.9)
plt.plot(sub_kids["sibs"], lm_maBA_Inter.params[0] +
lm_maBA_Inter.params[1] * 1 + lm_maBA_Inter.params[2] *
sub_kids["sibs"] + lm_maBA_Inter.params[3] * 1 * sub_kids["sibs"],
'purple', label = 'BA+', alpha = 0.9)
plt.title("Kids of BA+ moms vs not with interaction")
plt.xlabel("sibs")
plt.ylabel("educ")
plt.show()
```

# What if we include other variables?

```
sub_kids["twobio"] = sub_kids["family16"] == 1
lm_maBA_twobio = smf.ols("educ ~ sibs * maBA + age + twobio", data = sub_kids).fit()
print(lm_maBA_twobio.summary())
```

## OLS Regression Results

```
=====
Dep. Variable:          educ      R-squared:          0.156
Model:                  OLS       Adj. R-squared:      0.154
Method:                 Least Squares   F-statistic:      109.7
Date:                  Wed, 22 May 2019   Prob (F-statistic): 1.47e-106
Time:                  10:21:41    Log-Likelihood:    -7455.9
No. Observations:      2977        AIC:              1.492e+04
Df Residuals:          2971        BIC:              1.496e+04
Df Model:               5
Covariance Type:       nonrobust
=====
```

	coef	std err	t	P> t	[95.0% Conf. Int.]
Intercept	13.9321	0.185	75.138	0.000	13.569 14.296
maBA[T.True]	1.4527	0.257	5.657	0.000	0.949 1.956
twobio[T.True]	0.5932	0.121	4.892	0.000	0.355 0.831
sibs	-0.2902	0.018	-16.206	0.000	-0.325 -0.255
sibs:maBA[T.True]	0.2251	0.075	3.006	0.003	0.078 0.372
age	-0.0028	0.003	-0.840	0.401	-0.009 0.004

```
=====
Omnibus:                329.509    Durbin-Watson:          1.771
Prob(Omnibus):           0.000    Jarque-Bera (JB):        752.973
Skew:                   -0.661    Prob(JB):                3.12e-164
Kurtosis:                5.079    Cond. No.                 249.
=====
```

Warnings:

(c) Eirich 2012

\*

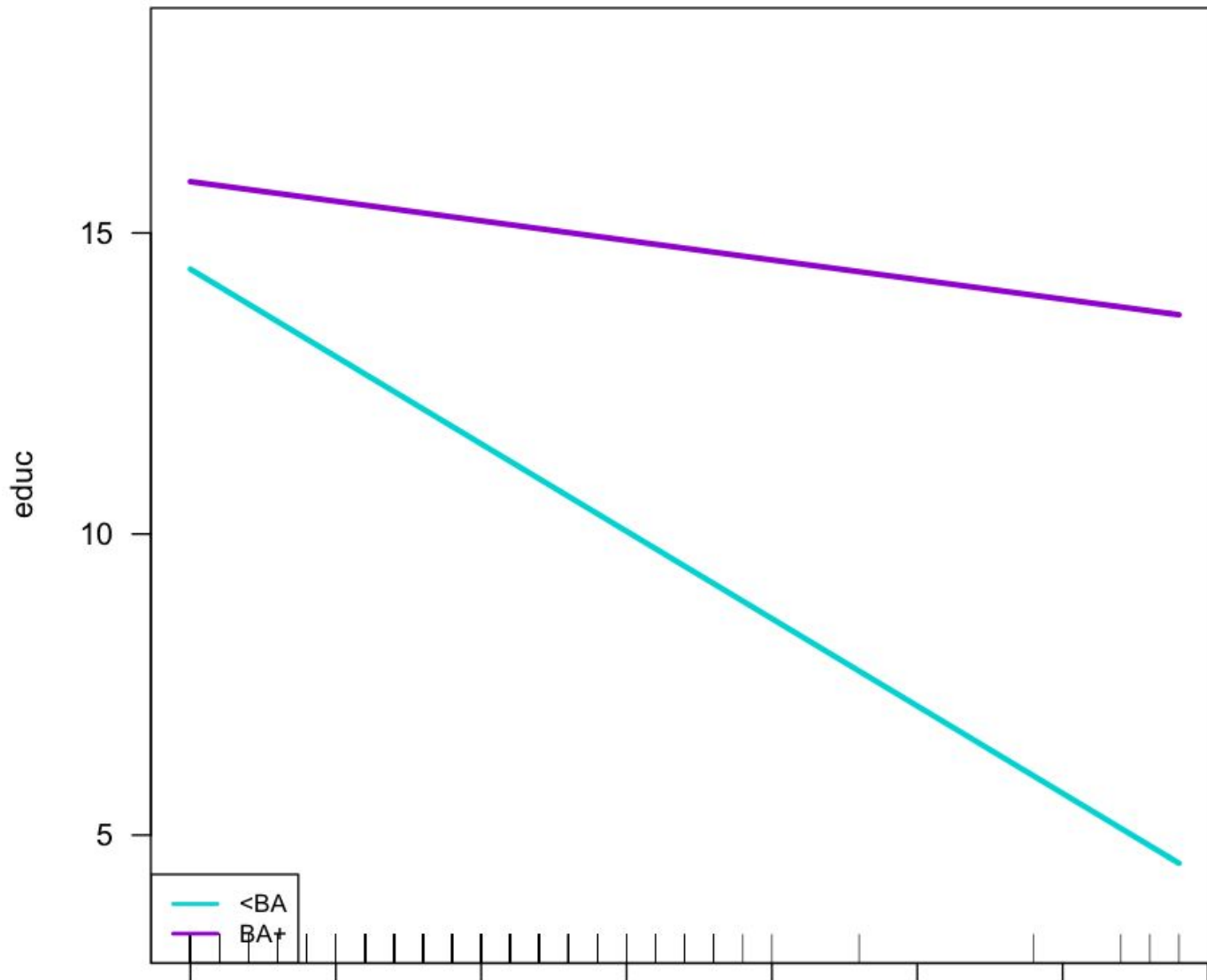
[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

# What if we include other variables?

	coef	std err	t	P> t	[95.0% Conf. Int.]	
Intercept	13.9321	0.185	75.138	0.000	13.569	14.296
maBA[T.True]	1.4527	0.257	5.657	0.000	0.949	1.956
twobio[T.True]	0.5932	0.121	4.892	0.000	0.355	0.831
sibs	-0.2902	0.018	-16.206	0.000	-0.325	-0.255
sibs:maBA[T.True]	0.2251	0.075	3.006	0.003	0.078	0.372
age	-0.0028	0.003	-0.840	0.401	-0.009	0.004

Net of other factors, kids (of BA+ moms) gain back .225 for each additional sibling they have – which means they ultimately lose -.075 for each additional sibling

# Here is the code for that graph



# Code here (in R)

```
# Or using visreg
```

```
visreg(lm(educ ~ sibs + maBA + sibs:maBA + age + twobio, data = sub),  
       xvar = "sibs", by = "maBA", overlay=T, partial = F, band = F, legend =  
F,  
       line = list(col = c("cyan3", "purple3")))  
  
legend("bottomleft", c("<BA", "BA+"), lwd = 2, col = c("cyan3", "purple3"),  
cex = 0.8)
```

# Let's try another one ...

Does education alter fundamentalists opinion on evolution?

# The recodes ...

```
pd.options.mode.chained_assignment = None

sub_evo = sub[["educ", "fund", "evolved", "family16", "age"]]
fund_dummy = {1:1, 2:0, 3:0}
sub_evo["fundamentalist"] = sub_evo["fund"].map(fund_dummy.get)
evolved_dummy = {1:1, 2:0}
sub_evo["evolution"] = sub_evo["evolved"].map(evolved_dummy.get)
```

# A simple multiple regression

A person who belongs to a fundamentalist religion is -.34 points lower in believing in evolution, net of education.

```
lm_evo = smf.ols(formula = 'evolution ~ fundamentalist + educ', data = sub_evo).fit()  
print(lm_evo.summary())
```

## OLS Regression Results

```
=====
```

Dep. Variable:	evolution	R-squared:	0.155
Model:	OLS	Adj. R-squared:	0.154
Method:	Least Squares	F-statistic:	138.2
Date:	Wed, 22 May 2019	Prob (F-statistic):	7.41e-56
Time:	10:22:43	Log-Likelihood:	-969.62
No. Observations:	1512	AIC:	1945.
Df Residuals:	1509	BIC:	1961.
Df Model:	2		
Covariance Type:	nonrobust		

```
=====
```

	coef	std err	t	P> t	[95.0% Conf. Int.]
-----	-----	-----	-----	-----	-----
Intercept	0.1610	0.062	2.580	0.010	0.039 0.283
fundamentalist	-0.3396	0.026	-13.238	0.000	-0.390 -0.289
educ	0.0317	0.004	7.381	0.000	0.023 0.040

```
=====
```

Omnibus:	0.004	Durbin-Watson:	2.018
Prob(Omnibus):	0.998	Jarque-Bera (JB):	117.722
Skew:	-0.004	Prob(JB):	2.73e-26
Kurtosis:	1.633	Cond. No.	75.2

```
=====
```



# A simple multiple regression

For each year more of education someone has, they have .03 more points of believing in evolution, net of religion.

	coef	std err	t	P> t	[95.0% Conf. Int.]	
Intercept	0.1610	0.062	2.580	0.010	0.039	0.283
fundamentalist	-0.3396	0.026	-13.238	0.000	-0.390	-0.289
educ	0.0317	0.004	7.381	0.000	0.023	0.040

# The interaction model I

1. When educ=0 and fundamentalist=0 (meaning for a non-fundamentalist), they are predicted to have 0.055 points of believing in evolution

```
lm_evo_Inter = smf.ols(formula = 'evolution ~ educ * fundamentalist' , data = sub_evo).fit()  
print (lm_evo_Inter.summary())
```

## OLS Regression Results

```
=====
```

Dep. Variable:	evolution	R-squared:	0.159
Model:	OLS	Adj. R-squared:	0.157
Method:	Least Squares	F-statistic:	94.82
Date:	Wed, 22 May 2019	Prob (F-statistic):	3.21e-56
Time:	10:23:09	Log-Likelihood:	-966.17
No. Observations:	1512	AIC:	1940.
Df Residuals:	1508	BIC:	1962.
Df Model:	3		
Covariance Type:	nonrobust		

```
=====
```

	coef	std err	t	P> t	[95.0% Conf. Int.]
-----	-----	-----	-----	-----	-----
Intercept	0.0550	0.074	0.740	0.459	-0.091 0.201
educ	0.0392	0.005	7.610	0.000	0.029 0.049
fundamentalist	-0.0139	0.127	-0.110	0.912	-0.262 0.235
educ:fundamentalist	-0.0244	0.009	-2.625	0.009	-0.043 -0.006

```
=====
```

Omnibus: 7331.985 Durbin-Watson: 2.018

Prob(Omnibus): 0.000 Jarque-Bera (JB): 111.031

Skewness: 0.020 Prob (ID): 7.76e-25

# The interaction model II

2. When educ=0, a fundamentalist is -0.0139 (not stat sig.) points lower on believing in evolution than a non-fundamentalist

	coef	std err	t	P> t	[95.0% Conf. Int.]	
Intercept	0.0550	0.074	0.740	0.459	-0.091	0.201
educ	0.0392	0.005	7.610	0.000	0.029	0.049
fundamentalist	-0.0139	0.127	-0.110	0.912	-0.262	0.235
educ:fundamentalist	-0.0244	0.009	-2.625	0.009	-0.043	-0.006

# The interaction model III

3. When fund=0 (i.e., for a non-fundamentalist), each additional year of education increases someone's belief in evolution by 0.039 points

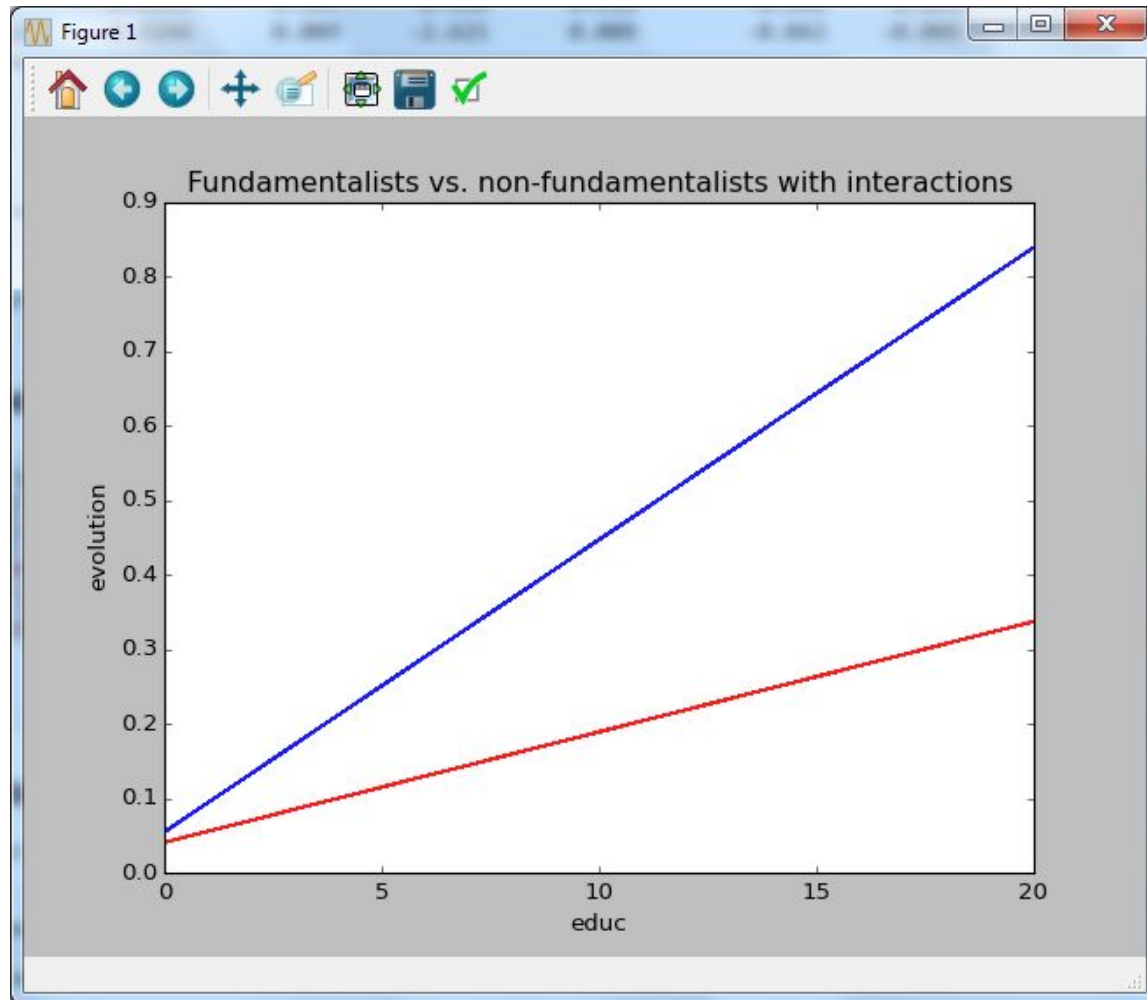
	coef	std err	t	P> t	[95.0% Conf. Int.]	
Intercept	0.0550	0.074	0.740	0.459	-0.091	0.201
educ	0.0392	0.005	7.610	0.000	0.029	0.049
fundamentalist	-0.0139	0.127	-0.110	0.912	-0.262	0.235
educ:fundamentalist	-0.0244	0.009	-2.625	0.009	-0.043	-0.006

# The interaction model IV

4. Fundamentalists get the 0.039 points on the evolution scale that non-fundamentalists get for each year more of education, but then they lose -0.024 points, for a total of 0.015 points for each year of education for fundamentalists.

	coef	std err	t	P> t	[95.0% Conf. Int.]	
Intercept	0.0550	0.074	0.740	0.459	-0.091	0.201
educ	0.0392	0.005	7.610	0.000	0.029	0.049
fundamentalist	-0.0139	0.127	-0.110	0.912	-0.262	0.235
educ:fundamentalist	-0.0244	0.009	-2.625	0.009	-0.043	-0.006

# Fundamentalists vs. non-fundamentalists, with interactions



# Here is that graph's code

```
plt.axis([0, 20, 0, 0.9])
plt.plot(sub_evo["educ"], lm_evo_Inter.params[0] + lm_evo_Inter.params[1] *
sub_evo["educ"] + lm_evo_Inter.params[2] * 0 + lm_evo_Inter.params[3] * 0 *
sub_evo["educ"], 'blue', label = 'Not Fundamentalist', alpha = 0.9)
plt.plot(sub_evo["educ"], lm_evo_Inter.params[0] + lm_evo_Inter.params[1] *
sub_evo["educ"] + lm_evo_Inter.params[2] * 1 + lm_evo_Inter.params[3] * 1 *
sub_evo["educ"], 'red', label = 'Fundamentalist', alpha = 0.9)
plt.title("Fundamentalists vs. non-fundamentalists with interactions")
plt.xlabel("educ")
plt.ylabel("evolution")
plt.show()
```

# But ...

Almost nobody has zero years of education, so why don't we find a better value to set to zero, like – say – the mean.

We can do that through centering:

```
sub_evo["educ"].mean()
```

```
13.293398533007334
```



# I start with this recode ...

```
pd.options.mode.chained_assignment = None
```

```
sub_evo["center_educ"] = sub_evo["educ"] - sub_evo["educ"].mean()
```

```
sub_evo["center_educ"].describe().map(lambda x: round(x, 4))
```

```
count    4499.00
mean         0.00
std         3.23
min        -13.29
25%         -1.29
50%         -0.29
75%          2.71
max          6.71
Name: center_educ, dtype: float64
```

# The interaction model I

1. When educ=13.29 (or centerededuc=0), a fundamentalist is -.34 points lower on believing in evolution than a non-fundamentalist.

```
lm_evo_Inter2 = smf.ols(formula = 'evolution ~ center_educ * fundamentalist', data = sub_evo).fit()
print(lm_evo_Inter2.summary())
```

## OLS Regression Results

```
=====
Dep. Variable:          evolution    R-squared:                0.159
Model:                  OLS         Adj. R-squared:            0.157
Method:                 Least Squares   F-statistic:              94.82
Date:                  Wed, 22 May 2019   Prob (F-statistic):       3.21e-56
Time:                  10:25:14         Log-Likelihood:           -966.17
No. Observations:      1512            AIC:                     1940.
Df Residuals:          1508            BIC:                     1962.
Df Model:              3
Covariance Type:       nonrobust
=====
```

	coef	std err	t	P> t	[95.0% Conf. Int.]
Intercept	0.5763	0.015	38.321	0.000	0.547 0.606
center_educ	0.0392	0.005	7.610	0.000	0.029 0.049
<b>fundamentalist</b>	<b>-0.3382</b>	<b>0.026</b>	<b>-13.204</b>	<b>0.000</b>	<b>-0.388 -0.288</b>
center_educ:fundamentalist	-0.0244	0.009	-2.625	0.009	-0.043 -0.006

```
=====
Omnibus:                7331.985    Durbin-Watson:           2.018
Prob(Omnibus):           0.000      Jarque-Bera (JB):         111.031
Skew:                   0.020      Prob(JB):                 7.76e-25
Kurtosis:               1.673      Cond. No.:                6.96
=====
```

# Remember the original model ...

1. When educ=0, a fundamentalist is -0.014 points lower on believing in evolution than a non-fundamentalist.

	coef	std err	t	P> t	[95.0% Conf. Int.]	
Intercept	0.0550	0.074	0.740	0.459	-0.091	0.201
educ	0.0392	0.005	7.610	0.000	0.029	0.049
<b>fundamentalist</b>	<b>-0.0139</b>	<b>0.127</b>	<b>-0.110</b>	<b>0.912</b>	<b>-0.262</b>	<b>0.235</b>
educ:fundamentalist	-0.0244	0.009	-2.625	0.009	-0.043	-0.006

# About centering ...

Only need to center continuous variables used in the interactions

The statistical significance changes because we are now looking at the change at the mean (where we have lots of data), not at zero (where we had almost no data).

Wait a minute: Why would I think there is an interaction here in the first place?

**These previous examples have all been instances of exacerbating (or amplifying) effects**

## **Now let's look at a diminishing (or redundant) effects example**

# Remember Wordsum & Marriage



# Wordsum, by Marriage & Educ

```
pd.options.mode.chained_assignment = None

sub_word = sub[["marital", "wordsum", "educ", "speduc"]]
sub_word["married"] = sub_word["marital"] == 1
```

# Wordsum, by Marriage & Educ

```
lm_wordsum = smf.ols(formula = 'wordsum ~ married + educ', data = sub_word).fit()
print (lm_wordsum.summary())
```

## OLS Regression Results

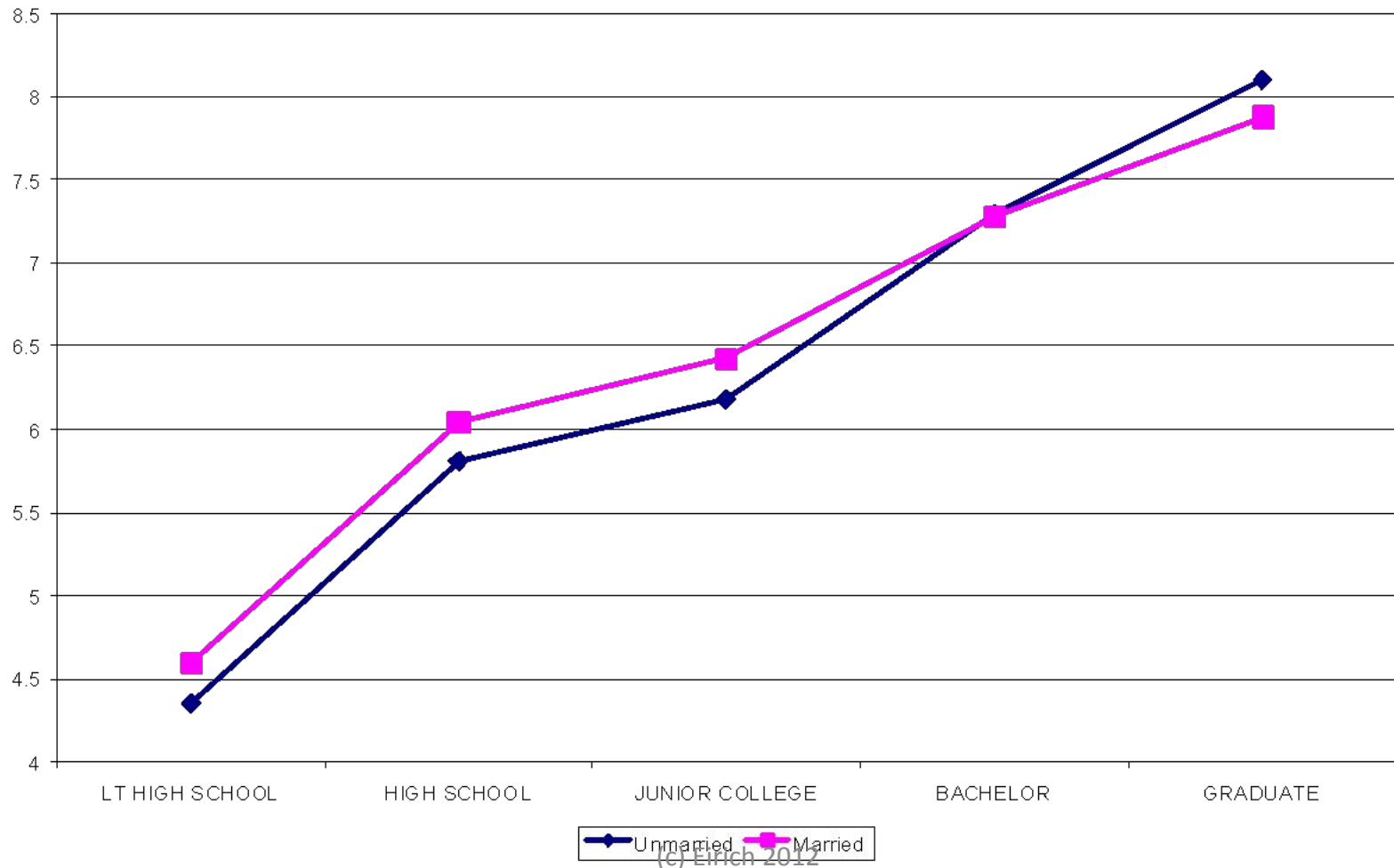
```
=====
Dep. Variable:          wordsum    R-squared:          0.191
Model:                  OLS        Adj. R-squared:       0.190
Method:                 Least Squares    F-statistic:      163.3
Date:                  Wed, 22 May 2019    Prob (F-statistic): 2.15e-64
Time:                  10:27:16          Log-Likelihood:    -2785.0
No. Observations:      1388             AIC:              5576.
Df Residuals:          1385             BIC:              5592.
Df Model:               2
Covariance Type:       nonrobust
=====
```

	coef	std err	t	P> t	[95.0% Conf. Int.]	
Intercept	2.4666	0.212	11.617	0.000	2.050	2.883
married[T.True]	0.0548	0.097	0.566	0.571	-0.135	0.245
educ	0.2721	0.015	18.024	0.000	0.243	0.302

```
=====
Omnibus:                67.260    Durbin-Watson:          1.932
Prob(Omnibus):           0.000    Jarque-Bera (JB):        82.065
Skew:                   -0.496    Prob(JB):                1.51e-18
Kurtosis:               3.661     Cond. No.                61.3
=====
```

# Interaction with 2 Continuous Variables

Wordsum Score



# Wordsum, by Marriage & Educ

```
lm_wordsum2 = smf.ols(formula = 'wordsum ~ married * educ', data = sub_word).fit()
print(lm_wordsum2.summary())
```

## OLS Regression Results

```
=====
Dep. Variable:          wordsum    R-squared:          0.191
Model:                  OLS        Adj. R-squared:       0.189
Method:                 Least Squares    F-statistic:       108.9
Date:                  Wed, 22 May 2019    Prob (F-statistic): 2.67e-63
Time:                  10:28:20          Log-Likelihood:    -2784.9
No. Observations:      1388            AIC:              5578.
Df Residuals:          1384            BIC:              5599.
Df Model:               3
Covariance Type:       nonrobust
=====
```

	coef	std err	t	P> t	[95.0% Conf. Int.]
Intercept	2.3537	0.295	7.985	0.000	1.775 2.932
married[T.True]	0.2794	0.418	0.669	0.504	-0.540 1.099
educ	0.2806	0.022	13.035	0.000	0.238 0.323
married[T.True]:educ	-0.0167	0.030	-0.553	0.581	-0.076 0.043

```
=====
Omnibus:                66.856    Durbin-Watson:          1.930
Prob(Omnibus):           0.000    Jarque-Bera (JB):       81.633
Skew:                   -0.493    Prob(JB):               1.88e-18
Kurtosis:                3.663    Cond. No.                157.
=====
```

Warnings:

(c) Eirich 2012

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

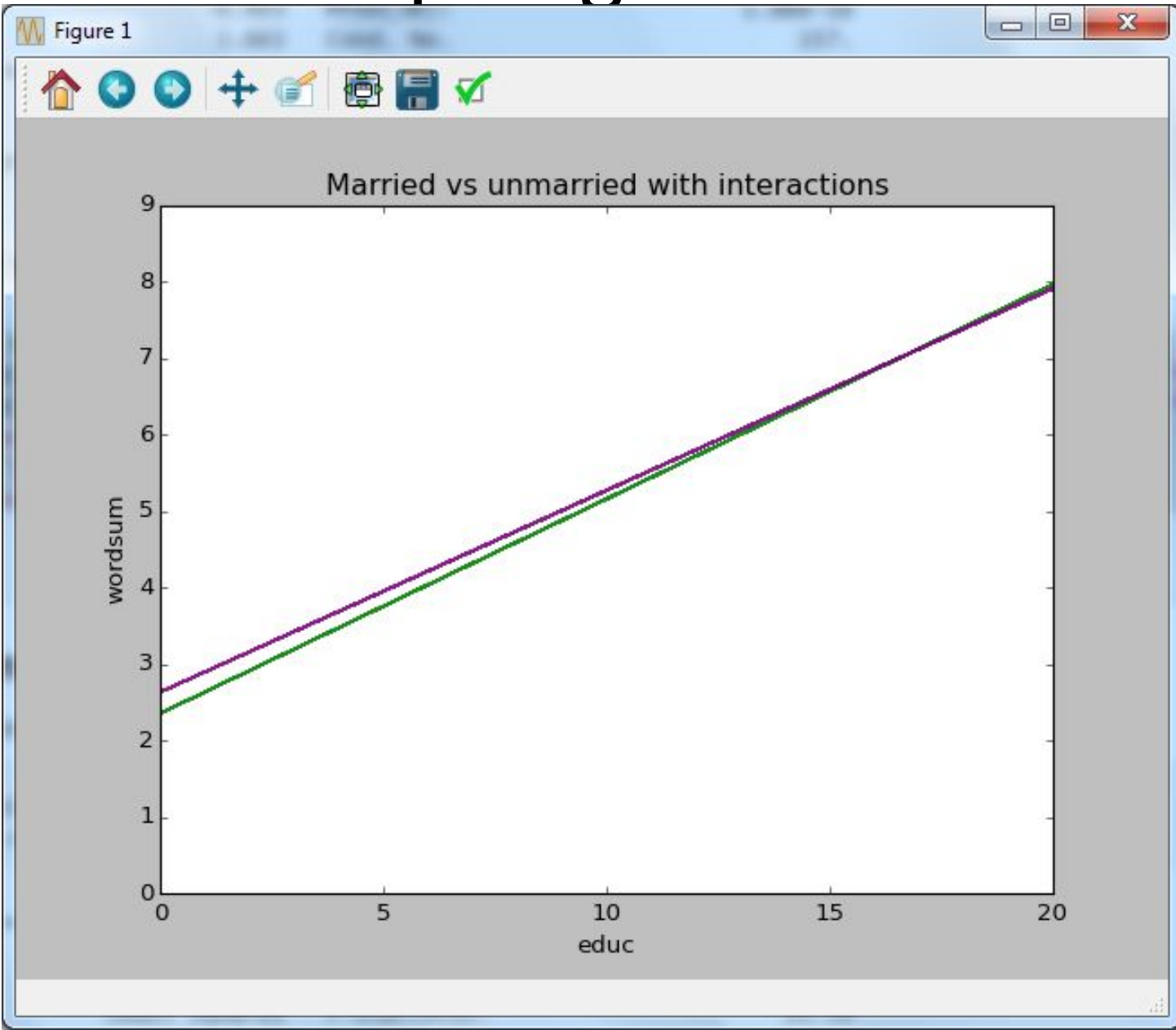
\*

# Wordsum, by Marriage & Educ

	coef	std err	t	P> t	[95.0% Conf. Int.]
Intercept	2.3537	0.295	7.985	0.000	1.775 2.932
married[T.True]	0.2794	0.418	0.669	0.504	-0.540 1.099
educ	0.2806	0.022	13.035	0.000	0.238 0.323
married[T.True]:educ	-0.0167	0.030	-0.553	0.581	-0.076 0.043

For married people, they lose -0.0167 (not statistically significant) Wordsum points for each year more educated they are, relative to non-married people

# Graphing this relationship



# Here is that graph's code

```
plt.axis([0, 20, 0, 9])
plt.plot(sub word["educ"], lm wordsum2.params[0] + lm wordsum2.params[1] * 0
+ lm wordsum2.params[2] * sub word["educ"] + lm wordsum2.params[3] * 0 *
sub word["educ"], 'green', label = 'Unmarried', alpha = 0.9)
plt.plot(sub word["educ"], lm wordsum2.params[0] + lm wordsum2.params[1] * 1
+ lm wordsum2.params[2] * sub word["educ"] + lm wordsum2.params[3] * 1 *
sub word["educ"], 'purple', label = 'Married', alpha = 0.9)
plt.title("Married vs unmarried with interactions")
plt.xlabel("educ")
plt.ylabel("wordsum")
plt.show()
```

A mechanism by which marriage can  
increase WordSum?



# Wordsum, by My Educ & My Spouse's Educ

```
lm_speduc = smf.ols(formula = 'wordsum ~ educ * speduc', data = sub_word).fit()
print (lm_speduc.summary())
```

## OLS Regression Results

```
=====
Dep. Variable:          wordsum      R-squared:          0.198
Model:                  OLS          Adj. R-squared:       0.194
Method:                 Least Squares  F-statistic:         55.38
Date:                  Wed, 22 May 2019  Prob (F-statistic):    5.25e-32
Time:                  10:29:41       Log-Likelihood:      -1352.3
No. Observations:      678           AIC:                2713.
Df Residuals:          674           BIC:                2731.
Df Model:               3
Covariance Type:       nonrobust
=====
```

	coef	std err	t	P> t	[95.0% Conf. Int.]
Intercept	3.5911	0.713	5.039	0.000	2.192 4.990
educ	0.1213	0.058	2.091	0.037	0.007 0.235
speduc	-0.0447	0.061	-0.733	0.464	-0.164 0.075
educ:speduc	0.0084	0.004	1.981	0.048	7.44e-05 0.017

```
=====
Omnibus:                45.184      Durbin-Watson:          1.921
Prob(Omnibus):          0.000      Jarque-Bera (JB):       56.570
Skew:                   -0.588      Prob(JB):               5.20e-13
Kurtosis:               3.788      Cond. No.:              2.14e+03
=====
```

# Wordsum, by My Educ & My Spouse's Educ

	coef	std err	t	P> t	[95.0% Conf. Int.]	
Intercept	3.5911	0.713	5.039	0.000	2.192	4.990
educ	0.1213	0.058	2.091	0.037	0.007	0.235
speduc	-0.0447	0.061	-0.733	0.464	-0.164	0.075
educ:speduc	0.0084	0.004	1.981	0.048	7.44e-05	0.017

The slope on the interaction of R's education and Spouse's education is positive, meaning that it leads to a widening gap as both educ and speduc grow larger

# Interpreting continuous by continuous interactions

$$\text{Wordsum} = 3.59 + 0.12 * \text{Educ} - 0.044 * \text{SpEduc} + 0.008 * \text{Educ} * \text{SpEduc}$$

Set SpEduc=0, then:

$$\text{Wordsum} = 3.59 + 0.12 * \text{Educ} + 0.044 * (0) + 0.008 * \text{Educ} * (0)$$

$$\text{Wordsum} = 3.59 + 0.12 * \text{Educ}$$

# Interpreting continuous by continuous interactions

If you plug in values for  $X_2$ , then you can figure out both the intercept and slope for each line ...

# Interpreting continuous by continuous interactions

Set SpEduc=**10**, then:

$$\text{Wordsum} = 3.59 + 0.12 * \text{Educ} - 0.044 * (10) + 0.008 * \text{Educ} * (10)$$

$$\text{Wordsum} = 3.59 + 0.12 * \text{Educ} - 0.44 + 0.08 * \text{Educ}$$

$$\text{Wordsum} = 3.15 + 0.20 * \text{Educ}$$

# Interpreting continuous by continuous interactions

Set SpEduc=**20**, then:

$$\text{Wordsum} = 3.59 + 0.12 * \text{Educ} - 0.044 * (20) + 0.008 * \text{Educ} * (20)$$

$$\text{Wordsum} = 3.59 + 0.12 * \text{Educ} - 0.88 + 0.16 * \text{Educ}$$

$$\text{Wordsum} = 2.71 + 0.28 * \text{Educ}$$

# Interpreting continuous by continuous interactions

When SpEduc=0, Wordsum = 3.59 + 0.12\*Educ

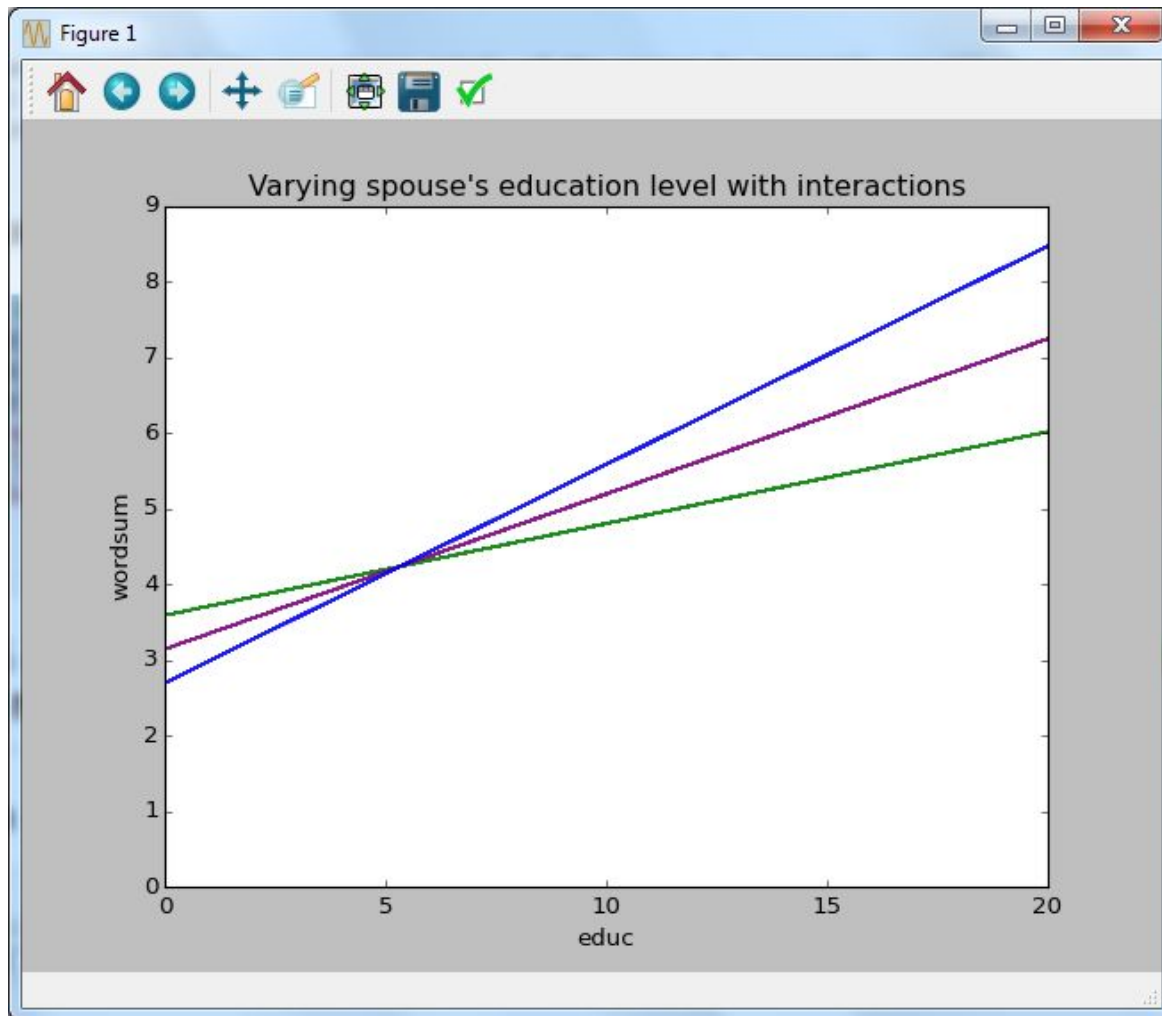
When SpEduc=10, Wordsum = 3.15 + 0.20\*Educ

When SpEduc=20, Wordsum = 2.71 + 0.28\*Educ

As SpEduc increases, the intercept decreases

As SpEduc increases, the slope increases too

# Graphing this relationship





# The code

```
plt.axis([0, 20, 0, 9])
plt.plot(sub word["educ"], lm speduc.params[0] + lm speduc.params[1] *
sub word["educ"] + lm speduc.params[2] * 0 + lm speduc.params[3] * 0 *
sub word["educ"], 'green', label = 'SpEduc = 0', alpha = 0.9)
plt.plot(sub word["educ"], lm speduc.params[0] + lm speduc.params[1] *
sub word["educ"] + lm speduc.params[2] * 10 + lm speduc.params[3] * 10 *
sub word["educ"], 'purple', label = 'SpEduc = 10', alpha = 0.9)
plt.plot(sub word["educ"], lm speduc.params[0] + lm speduc.params[1] *
sub word["educ"] + lm speduc.params[2] * 20 + lm speduc.params[3] * 20 *
sub word["educ"], 'blue', label = 'SpEduc = 20', alpha = 0.9)
plt.title("Varying spouse's education level with interactions")
plt.xlabel("educ")
plt.ylabel("wordsum")
plt.show()
```

**Another example of an interaction**  
**(in STATA, sorry)**

Being more educated is associated with improved health. Going to religious services more is associated with improved health. Does someone get even more out of their education and attendance when they are high on both?

# Simple regression, in STATA

```
. vreverse health, gen(rhealth)
```

```
. reg rhealth educ attend age, beta
```

Source	SS	df	MS	Number of obs = 40522	
Model	4027.32716	3	1342.44239	F( 3, 40518) =	2168.31
Residual	25085.5252	40518	.61912052	Prob > F =	0.0000
Total	29112.8524	40521	.718463325	R-squared =	0.1383
				Adj R-squared =	0.1383
				Root MSE =	.78684

rhealth	Coef.	Std. Err.	t	P> t	Beta
educ	.0673073	.0012567	53.56	0.000	.2533878
attend	.0237323	.0014629	16.22	0.000	.0756068
age	-.0106901	.0002324	-46.00	0.000	-.2197657
_cons	2.549936	.0215238	118.47	0.000	.

Both higher education and higher religious attendance  
are positively predictive of health, net of age

# Interaction model

```
. reg rhealth c.educ#c.attend age
```

Source	SS	df	MS	Number of obs = 40522		
Model	4038.14542	4	1009.53635	F( 4, 40517) = 1631.26		
Residual	25074.707	40517	.618868795	Prob > F = 0.0000		
Total	29112.8524	40521	.718463325	R-squared = 0.1387		
				Adj R-squared = 0.1386		
				Root MSE = .78668		

rhealth	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
educ	.0747059	.0021703	34.42	0.000	.0704522	.0789597
attend	.0473712	.00584	8.11	0.000	.0359247	.0588177
c.educ#						
c.attend	-.0018764	.0004488	-4.18	0.000	-.002756	-.0009967
age	-.0107206	.0002325	-46.12	0.000	-.0111762	-.010265
_cons	2.458203	.0307323	79.99	0.000	2.397967	2.518439

As both education and attendance increase together, they have a diminishing effect on someone's health (note: health was reverse coded), net of age

# Interpreting continuous by continuous interactions

$$\text{Health} = 2.45 + 0.074 * \text{Educ} + 0.047 * \text{Attend} - 0.0018 * \text{Educ} * \text{Attend}$$

(notice: I don't need to include age, because we can set that to anything constant for each line)

Set Attend=0, then:

$$\text{Health} = 2.45 + 0.074 * \text{Educ} + 0.047 * (0) - 0.0018 * \text{Educ} * (0)$$

$$\text{Health} = 2.45 + 0.074 * \text{Educ}$$

# Interpreting continuous by continuous interactions

$$\text{Health} = 2.45 + 0.074 * \text{Educ} + 0.047 * \text{Attend} - 0.0018 * \text{Educ} * \text{Attend}$$

Set Attend=4, then:

$$\text{Health} = 2.45 + 0.074 * \text{Educ} + 0.047 * (4) - 0.0018 * \text{Educ} * (4)$$

$$\text{Health} = 2.65 + 0.066 * \text{Educ}$$

# Interpreting continuous by continuous interactions

$$\text{Health} = 2.45 + 0.074 * \text{Educ} + 0.047 * \text{Attend} - 0.0018 * \text{Educ} * \text{Attend}$$

Set Attend=8, then:

$$\text{Health} = 2.45 + 0.074 * \text{Educ} + 0.047 * (8) - 0.0018 * \text{Educ} * (8)$$

$$\text{Health} = 2.85 + 0.058 * \text{Educ}$$



# Interpreting continuous by continuous interactions

When Attend=0, Attend =  $2.45 + 0.074 * \text{Educ}$

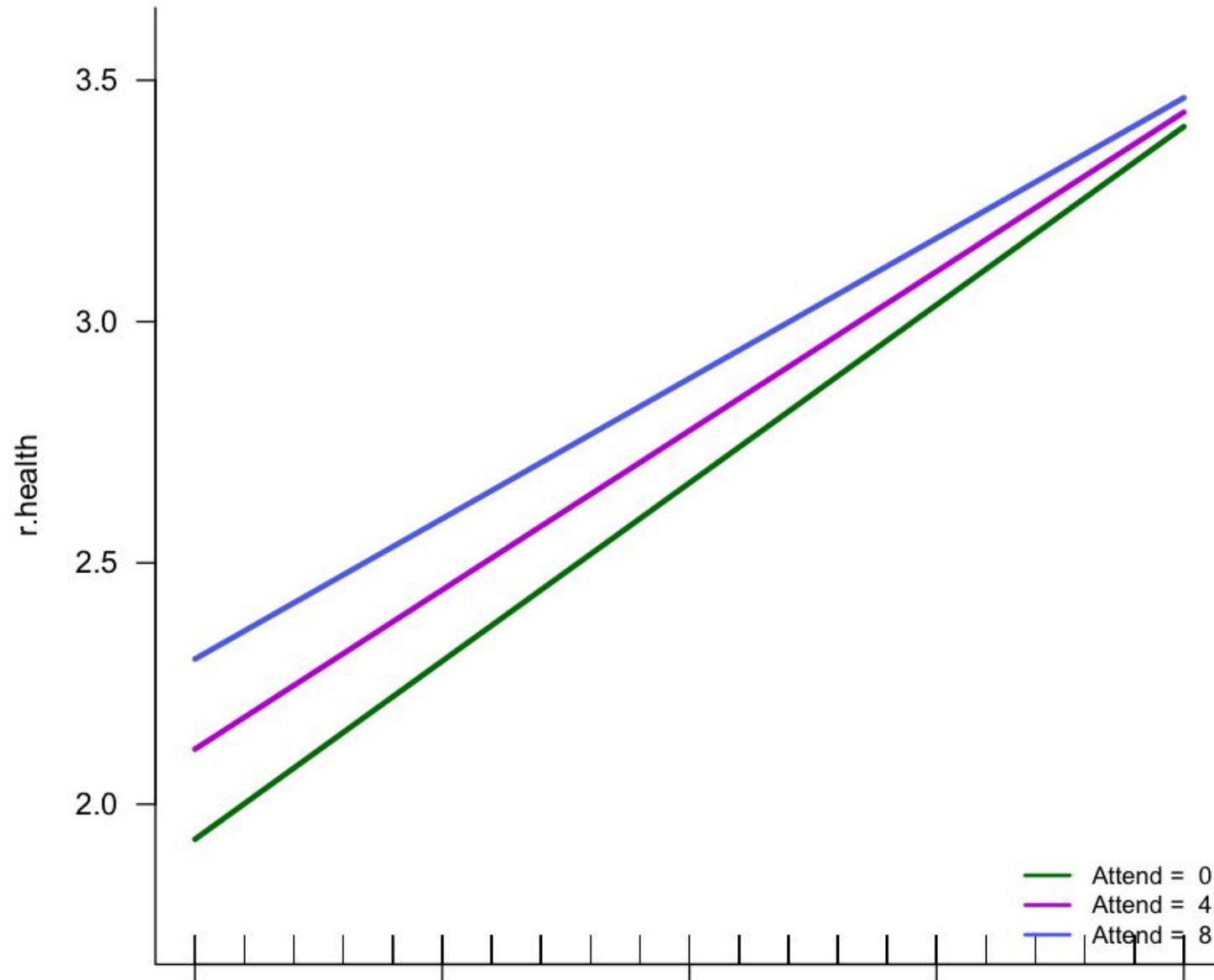
When Attend=4, Health =  $2.65 + 0.066 * \text{Educ}$

When Attend=8, Health =  $2.85 + 0.058 * \text{Educ}$

As Attend increases, the intercept increases too

But as Attend increases, the slope decreases

# Here it is graphed



# Here is that graph's code (in R)

```
lm.health <- lm(r.health ~ educ*attend + age, data = sub)
summary(lm.health)
```

```
# Plotting the relationship (using visreg for practice)
visreg(lm.health, "educ", by = "attend", breaks = c(0,4,8),
       overlay=T, band = F, partial = F, bty = "l", legend = F,
       line = list(col = c("darkgreen", "darkorchid", "royalblue")))
legend("bottomright", paste("Attend = ", c(0,4,8)), bty = "n", lwd = 2,
      col = c("darkgreen", "darkorchid", "royalblue"), cex = 0.8)
```