

Tail Inequalities

- Given the earlier conditions, it holds that for any $\delta > 0$,

$$P_X(X \geq \mu(1+\delta)) \leq \left(\frac{e^\delta}{(1+\delta)^{1+\delta}} \right)^\mu$$

- So, $P_X(Y \geq e^{t\mu(1+\delta)}) = e^{-\mu(1-e^t) - t\mu(1+\delta)}$
- Since t is a free parameter in the above, we can find a t that minimizes the right hand side.

- To simplify, let $f(t) = \ln e^{-\mu(1-e^t) - t\mu(1+\delta)}$
 $= -\mu(1-e^t) - t\mu(1+\delta)$

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- Given the earlier conditions, it holds that for any $\delta > 0$,

$$\Pr(X \geq \mu(1+\delta)) \leq \left(\frac{e^\delta}{(1+\delta)^{1+\delta}} \right)^\mu$$

- To simplify, let $f(t) = \ln e^{-\mu(1-e^t) - t\mu(1+\delta)}$

$$= -\mu(1-e^t) - t\mu(1+\delta)$$

- Differentiating $f(t)$ wrt t , we get

$$f'(t) = \mu e^t - \mu(1+\delta)$$

- So, $f'(t) = 0$ at $t = \ln(1+\delta)$

- Verify that the above t corresponds to a minima.
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Tail Inequalities

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$$P_X(X \geq \mu(1+\delta)) \leq \left(\frac{e^\delta}{(1+\delta)^{1+\delta}} \right)^\mu$$

- With $t = \ln(1+\delta)$, we get that

$$\begin{aligned} P_X(X \geq \mu(1+\delta)) &\leq \frac{e^{-\mu(1+\delta)t}}{(1+\delta)^{\mu(1+\delta)}} = \frac{e^{\mu\delta}}{(1+\delta)^{1+\delta}} \\ &= \left(\frac{e^\delta}{(1+\delta)^{1+\delta}} \right)^\mu \end{aligned}$$

- completing the proof.

Tail Inequalities

- Given the earlier conditions, it holds that for any $\delta > 0$,

$$\Pr(X \geq \mu(1+\delta)) \leq \left(\frac{e^\delta}{(1+\delta)^{1+\delta}} \right)^\mu$$

- With $t = \ln(1+\delta)$, we get that

$$\Pr(X \geq \mu(1+\delta)) \leq \frac{e^{-\mu(1-(1+\delta))}}{(1+\delta)^{\mu(1+\delta)}} = \frac{e^{\mu\delta}}{(1+\delta)^{1+\delta}}$$

$$= \left(\frac{e^\delta}{(1+\delta)^{1+\delta}} \right)^\mu$$

- A simplification of the RHS gives

$$\Pr(X \geq \mu(1+\delta)) \leq \begin{cases} e^{-\mu\delta^2/4} & \text{if } \delta \leq 1 \\ e^{-\mu\delta \ln \delta} & \text{if } \delta > 1 \end{cases}$$

A Simple Example

- Here is one more practical application of the Chernoff bounds.
- Consider once again counting the number of heads out of tossing n fair coins independently.
 - So, $p = \frac{1}{2}$.
- Let X_i denote the random variable that takes 1 if the i th coin toss results in a head, and 0 otherwise.
 - $E[X_i] = \frac{1}{2}$.
- Let $X = \sum_i X_i$
- X counts the total number of heads over the n tosses.
 - $E[X] = n/2$.
 - With $n = 100$, say, we expect 50 heads over 100 coin tosses.

A Simple Example

- Markov inequality tells that the probability that X takes a value beyond 70 is $\Pr(X \geq 70) = \Pr(X \geq 70/50 \times 50) \leq 5/7 = 0.7$ (approx.)
- To apply Chebyshev's inequality, we need to do some extra work.
- $\text{Var}(X_i)$ for any i is computed as $E[X_i^2] - E[X_i]^2$.
- $E[X_i^2] = 0 \times (1/2) + 1 \times (1/2) = 1/2$.
- $\text{Var}(X_i) = 1/2 - (1/2)^2 = 1/4$.
- $\text{Var}(X) = 100 \times \text{Var}(X_1)$ due to independence.
- So, $\text{Var}(X) = 100/4 = 25$, and $\sigma_X = 5$.
- Now, $\Pr(X \geq 70)$ can be rewritten as $\Pr(|X - 50| \geq 20)$ and further as $\Pr(|X - 50| \geq (20 \times 1/\sigma_X) \sigma_X)$ which is now at most $25/400 = 1/16 = 0.0625$.